Real-Time Bayesian Forecasting with State-Space Models



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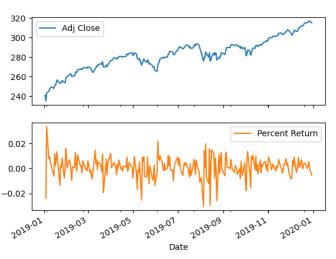
Overview

- 1 Motivation For State-Space Models
- 2 Forecasting in a Bayesian Way
- 3 Option 2: The Two Step Approach
- 4 Option 1: On-Line Sampling of Both States and Parameters
- 5 Option 2 (With My Approach)
- 6 Numerical Experiments

Motivation For State-Space Models

Motivation





Our State-Space Model

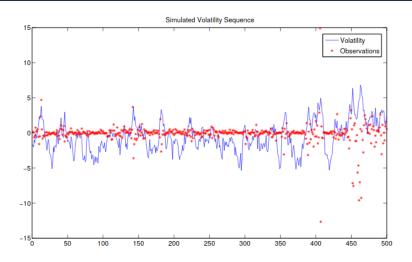
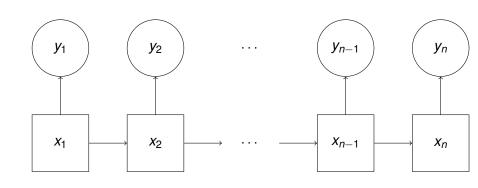


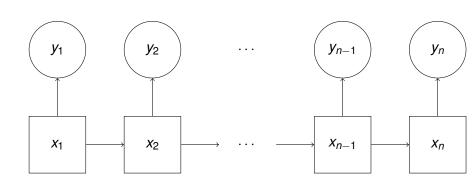
Figure: image from Doucet and Johansen, 2011

State-Space Models



vertical arrows: $g(y_t \mid x_t, \theta)$ horizontal arrows: $f(x_t \mid x_{t-1}, \theta)$

Our State-Space Model



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g(y_t \mid x_t, \theta) = \text{Normal}(0, \beta^2 e^{x_t})

f(x_t \mid x_{t-1}, \theta) = \text{Normal}(\phi x_{t-1}, \sigma^2)

x_t is log-volatility, represents how "active" the market is...

(Taylor, 1982)
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Forecasting in a Bayesian Way

Our Goal

Forecasting in a Bayesian way–the posterior predictive distribution:

$$p(y_{t+1} \mid y_{1:t}) = \int p(y_{t+1} \mid \theta, y_{1:t}) p(\theta \mid y_{1:t}) d\theta$$

- $y_{1:t}$ are the percent returns
- m heta is the vector of model parameters for the assumed true model
- \blacksquare $t \to \infty$
- can't condition on unknown parameters or unknown states
- could predict further into the future, too

Our Goal

With state-space models, there are two general approaches to approximating

$$p(y_{t+1} \mid y_{1:t}) = \int p(y_{t+1} \mid \theta, y_{1:t}) p(\theta \mid y_{1:t}) d\theta$$

- option 1: the one step approach
- option 2: the two step approach

Option 1: A Very Difficult Task!

We want

$$p(y_{t+1} \mid y_{1:t}) = \iint p(y_{t+1} \mid x_t, \theta, y_{1:t}) p(x_t, \theta \mid y_{1:t}) dx_t d\theta$$
$$= \iint p(y_{t+1} \mid x_t, \theta) p(x_t, \theta \mid y_{1:t}) dx_t d\theta$$

- Sampling both states and parameters from one on-line algorithm $p(x_t, \theta \mid y_{1:t})$
- \triangleright $p(y_{t+1} \mid x_t, \theta)$ is easy to deal with

Option 2: Two Difficult Tasks in One

Option 2:

$$p(y_{t+1} \mid y_{1:t}) = \int p(y_{t+1} \mid \theta, y_{1:t}) p(\theta \mid y_{1:t}) d\theta$$

- **1** $p(y_{t+1} | \theta, y_{1:t})$ involves filtering
- $p(\theta \mid y_{1:t})$ is standard posterior inference, but now for every t

We focus on this first, and briefly summarize option 1 later...

Option 2: The Two Step Approach

Option 2: Two Difficult Tasks in One

Option 2 was based on:

$$p(y_{t+1} \mid y_{1:t}) = \int p(y_{t+1} \mid \theta, y_{1:t}) p(\theta \mid y_{1:t}) d\theta$$

filtering:

$$\begin{split} & p(y_{t+1} \mid \theta, y_{1:t}) \\ &= \int g(y_{t+1} \mid x_{t+1}, \theta) f(x_{t+1} \mid x_t, \theta) \underbrace{p(x_t \mid y_{1:t}, \theta)}_{\text{filtering distribution}} dx_{t:t+1} \end{split}$$

 $p(\theta \mid y_{1:t})$ (posterior inference for every t)

We describe filtering first...

A Quick Word on Filtering

$$p(x_t \mid y_{1:t}, \theta) = \frac{g(y_t \mid x_t, \theta)p(x_t \mid y_{1:t-1}, \theta)}{p(y_t \mid y_{1:t-1}, \theta)}$$

where $p(x_t \mid y_{1:t-1}, \theta) = \int f(x_t \mid x_{t-1}, \theta) p(x_{t-1} \mid y_{1:t-1}, \theta) dx_{t-1}$ comes from the previous filtering distribution.

On-line algorithms include

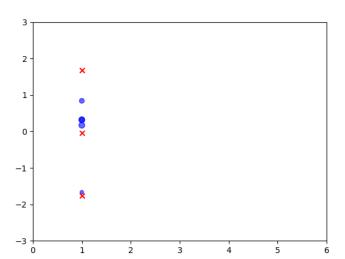
- (Kalman, 1960) filtering for linear Gaussian models
- 2 Hidden Markov Model filter for finite state space models
- approximate filtering otherwise (e.g. extended Kalman filter, unscented Kalman filter, particle filters, etc.)

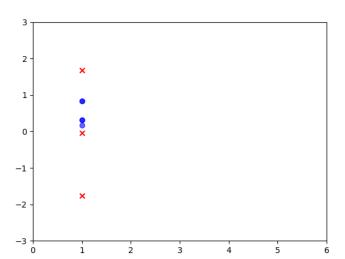
Particle Filtering (The General Strategy)

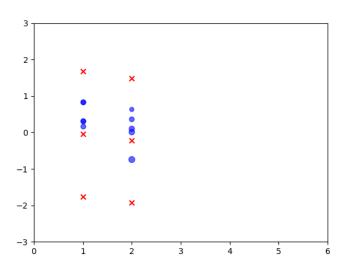
Choose a large N. Then

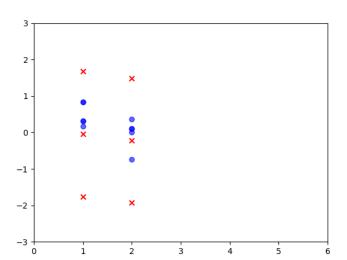
- sample $x_1^1, \dots, x_1^N \stackrel{\text{approx. iid}}{\sim} p(x_1 \mid y_1, \theta)$
- compute weights then resample those into $x_2^1, \ldots, x_2^N \stackrel{\text{approx. iid}}{\sim} p(x_2 \mid y_{1:2}, \theta)$
- compute weights then resample those into $x_3^1, \ldots, x_3^N \stackrel{\text{approx. iid}}{\sim} p(x_3 \mid y_{1:3}, \theta)$
- etc. etc.

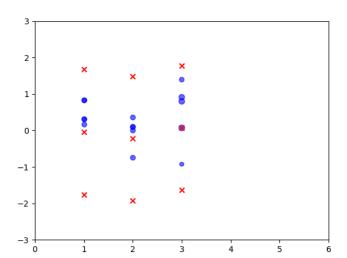


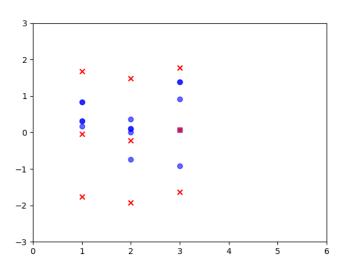


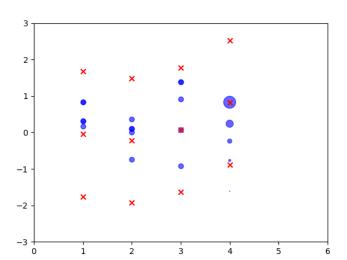


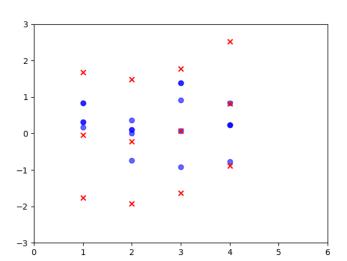


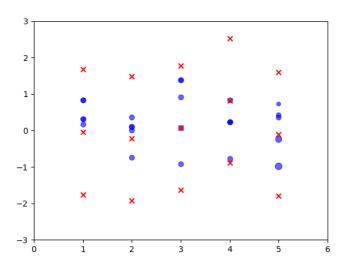


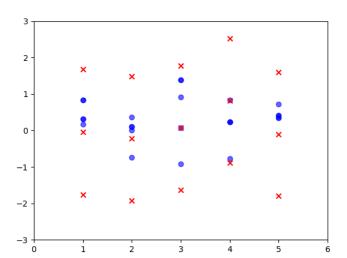












The other half of Option 2: Posteriors for θ

Option 2 was motivated by

$$p(y_{t+1} \mid y_{1:t}) = \int \underbrace{p(y_{t+1} \mid \theta, y_{1:t})}_{2} \underbrace{p(\theta \mid y_{1:t})}_{2} d\theta$$

How do we use particle filtering, but incorporate uncertainty about θ ?

The other half of Option 2: Posteriors for θ

There are many ways to sample from **one** posterior $p(\theta \mid y_{1:t})$

...but they can all be either time-consuming, difficult to tune or program, or impossible to use on data sets with large t.

Is there a way to avoid having to do this at every time step on an increasingly large data set?

Even if there was, how do we use parameter samples together with a particle filter?

This is why so much work has been done on Option 1...

Option 1: On-Line Sampling of Both States and Parameters

Option 1: On-Line Sampling From $p(x_t, \theta \mid y_{1:t})$

This is just a quick review.

It won't describe any of these in detail, but we list out a few names...

Option 1: On-Line Sampling From $p(x_t, \theta \mid y_{1:t})$

- Run one particle filter on a model with extended state $\tilde{x}_t = (x_t, \theta)$ (Kitagawa, 1998)
- 2 Do the above, but exploit sufficient statistics of $p(\theta \mid x_{1:t}, y_{1:t})$ (i.e. the Storvik Filter (Storvik, 2002), (Fearnhead, 2002) and Particle Learning (Carvalho et al., 2010))
- 3 Run one particle filter on a model with extended state $\tilde{x}_t = (x_t, \theta_t)$, assume θ_t is an (artificial) random walk (the Liu-West Filter (Liu and West, 2001) and the original proposal in (Kitagawa, 1998))
- 4 Other recursive (but not on-line) algorithms: The Resample-Move algorithm (Gilks and Berzuini, 2001) and SMC² (Chopin, Jacob, and Papaspiliopoulos, 2013)

Option 1: On-Line Sampling From $p(x_t, \theta \mid y_{1:t})$

Option 1 approaches are all generally "fast", but they suffer from some combination of the following:

- 1 high variance (due to particle degeneracy),
- 2 have a nonquantifiable bias, or they are
- 3 not on-line.

So, what do we do...

Option 2 (With My Approach)

Back To Option 2: The Particle Swarm Filter

My approach: don't always update the posterior, and aggregate particle filters' forecasts.

$$p(y_{t+1} \mid y_{1:t}) = \int p(y_{t+1} \mid \theta, y_{1:t}) p(\theta \mid y_{1:t}) d\theta$$
$$\approx \int p(y_{t+1} \mid \theta, y_{1:t}) p(\theta \mid y_{1:s}) d\theta$$

where $0 \le s < t$. $p(\theta \mid y_{1:s})$ is either a prior, or an "old" posterior that you can't afford to update every moment.

■ Sample $\{\theta^j\}_j$ once (or occasionally); run many particle filters with these plugged in; keep averaging predictions.

The Particle Swarm Filter: Theoretical Guarantees

Summary of some of the theoretical results in (Brown, 2021):

- Consistency *to the proxy distribution* under strong assumptions (compact parameter space, stochastic equicontinuity of $\hat{p}(x_t \mid y_{1:t}, \theta)$);
- asymptotic normality under the same assumptions, and again, to the proxy distribution;
- 3 no uniform-in-time bounds for either the bias or for the asymptotic variances (caveat emptor)

Bias

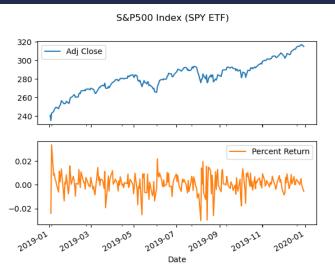
ideal fcast density
$$\overbrace{p(y_{t+1} \mid y_{1:t})}^{\text{our forecast}} - \int p(y_{t+1} \mid \theta, y_{1:t}) p(\theta \mid y_{1:s}) d\theta$$

$$= \int p(y_{t+1} \mid \theta, y_{1:t}) \left\{ \frac{p(\theta \mid y_{1:t})}{p(\theta \mid y_{1:s})} - 1 \right\} p(\theta \mid y_{1:s}) d\theta \qquad (1)$$

- Using a prior is not encouraged (always make sure to examine forecasts produced!).
- For fixed s, $\frac{p(\theta|y_{1:t})}{p(\theta|y_{1:s})}$ (probably) grows in t.
- Occasionaly updating posterior is better than never updating.
- Easy to extend to averaging over models, too? (future work)

Numerical Experiments

The S&P 500 Index



What about options contracts on this?

Options on the S&P 500 Index

- a **call (put) option** is a derivative contract that allows one to buy (sell) a fixed quantity at some **strike price**.
- 2 you can either buy or sell this contract
- It has a finite lifetime—it expires on the **expiration date**.
- 4 you can buy and sell these contracts, or you can **exercise** the privelege of buying (selling) the underlying shares.
- **European style** options will only allow you to exercise on the expiration date (SPX).
- **American style** options allow an exercise at any time before expiration (SPY).

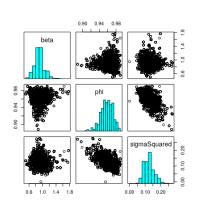
Options: A Few Takeaways From Mathematical Finance

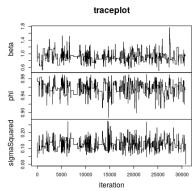
The option's price should be tightly linked with the price of the SPY, but we are not doing **pricing** here. We're not trying to come up with a hedge and identify arbitrages (if we were we would need SDEs and stochastic calculus)—we want to take bets without trading too actively.

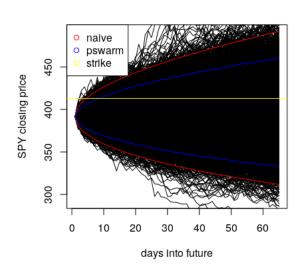
Option prices tend to

- decrease in value over time
- change in value if the underlying moves towards or away from the strike (nonlinearly with respect to the underlying)
- 3 decrease in value if volatility decreases

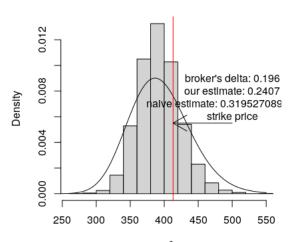
- Use old posterior samples from a pseudo-marginal Metropolis-Hastings algorithm
- 2 Sample 100 parameter vectors "from" those
- 3 Each particle filter maintains 100 state samples
- 4 Each particle filter runs through the past 807 trading days (2018-01-01 until now)
- Each particle filter is used to simulate forward in time 64 days (until expiration day)
- 6 Code: github.com/tbrown122387/pf (Brown, 2020), github.com/tbrown122387/ssme



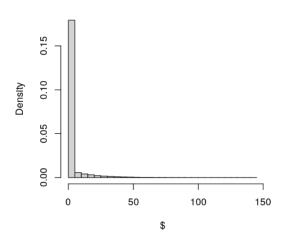




SPY on expiration (May 21 '21)



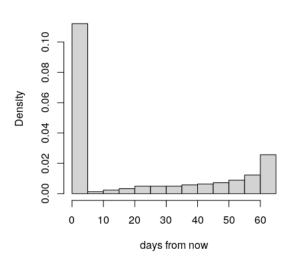
max end-of-day intrinsic value



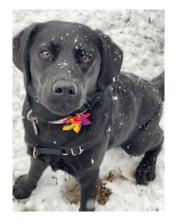
for American-style,

\$2.61-\$2.68 at March 18 close.

day of max intrinsic value



Thanks!



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