Problem 10.1 (A simple identity)

Define

$$\begin{split} \tau &\equiv \operatorname{E}\left(Y_i(1) - Y_i(0)\right), \\ \tau_{\operatorname{T}} &\equiv \operatorname{E}\left(Y_i(1) - Y_i(0) \mid Z_i = 1\right), \\ \tau_{\operatorname{C}} &\equiv \operatorname{E}\left(Y_i(1) - Y_i(0) \mid Z_i = 0\right), \end{split}$$

where the expectations are taken over the joint distribution of $(Y_i(1), Y_i(0), Z_i)$. Later in the derivation, we omit the subscript i for simplicity.

We are asked to show that

$$\tau = P(Z=1)\tau_{\mathrm{T}} + P(Z=0)\tau_{\mathrm{C}}.$$

By the law of total expectation,

$$\begin{split} \tau &= \operatorname{E}\left(Y(1) - Y(0)\right) \\ &= \operatorname{E}\left(\operatorname{E}\left(Y(1) - Y(0) \mid Z\right)\right) \\ &= P(Z=1)\operatorname{E}\left(Y(1) - Y(0) \mid Z=1\right) + P(Z=0)\operatorname{E}\left(Y(1) - Y(0) \mid Z=0\right) \\ &= P(Z=1)\tau_{\mathrm{T}} + P(Z=0)\tau_{\mathrm{C}}, \end{split}$$

which completes the proof.

Problem 10.2 (Nonparametric identification of other causal effects)

Under ignorability, i.e., $Y(z) \perp \!\!\! \perp Z \mid X$ for z = 0, 1, we are asked to show that

1. the distributional causal effect

$$P(Y(1) > y) - P(Y(0) > y)$$

is nonparametrically identifiable for all y.

2. the quantile causal effect

$$\operatorname{quantile}_q\left(Y(1)\right) - \operatorname{quantile}_q\left(Y(0)\right)$$

is nonparametrically identifiable for all $q \in (0,1)$, where quantile $_q(Y(z)) = \inf\{y : P(Y(z) \le y) \ge q\}$ for z=0,1.

First, we show that P(Y(z) > y) is nonparametrically identifiable:

$$\begin{split} &P(Y(z)>y)\\ &= \operatorname{E}\left(P(Y(z)>y\mid X)\right)\\ &= \operatorname{E}\left(P(Y(z)>y\mid X,Z=z)\right) \end{split} \tag{Iterated Expectation}$$

$$= \mathbb{E}\left(P(Y > y \mid X, Z = z)\right), \tag{Consistency}$$

which suggests that the distributional causal effect is identifiable.

Next, we have

$$\begin{aligned} &\operatorname{quantile}_q(Y(z)) \\ &= \inf\{y: P(Y(z) \leq y) \geq q\} \\ &= \inf\{y: 1 - P(Y(z) > y) \geq q\} \end{aligned} & \text{(Complement Rule)} \\ &= \inf\{y: 1 - \operatorname{E}\left(P(Y > y \mid X, Z = z)\right) \geq q\}, \end{aligned} & \text{(From the previous part)} \end{aligned}$$

which suffices to show that the quantile causal effect is identifiable.

Problem 10.3 (Outcome imputation estimator in the fully interacted logistic model)

Assume that a binary outcome follows a logistic model

$$\mathbf{E}\left(Y\mid Z,X\right) = P(Y=1\mid Z,X) = \frac{e^{\beta_0+\beta_zZ+\beta_xX+\beta_{xz}^{\intercal}XZ}}{1+e^{\beta_0+\beta_zZ+\beta_xX+\beta_{xz}^{\intercal}XZ}},$$

We are asked to derive the outcome regression estimator for the average treatment effect (ATE).

By the identification result of the ATE,

$$\begin{split} \tau &= \operatorname{E}\left(Y(1) - Y(0)\right) \\ &= \operatorname{E}\left(\operatorname{E}\left(Y \mid Z = 1, X\right) - \operatorname{E}\left(Y \mid Z = 0, X\right)\right) \\ &= \operatorname{E}\left(\frac{e^{\beta_0 + \beta_z + (\beta_x + \beta_{xz})^\top X}}{1 + e^{\beta_0 + \beta_z + (\beta_x + \beta_{xz})^\top X}} - \frac{e^{\beta_0 + \beta_x^\top X}}{1 + e^{\beta_0 + \beta_x^\top X}}\right), \end{split} \tag{From the logistic model}$$

which can be estimated by the sample average

$$\hat{\tau} = \frac{1}{n} \sum_{i=1}^n \left(\frac{e^{\hat{\beta}_0 + \hat{\beta}_z + (\hat{\beta}_x + \hat{\beta}_{xz})^\intercal X_i}}{1 + e^{\hat{\beta}_0 + \hat{\beta}_z + (\hat{\beta}_x + \hat{\beta}_{xz})^\intercal X_i}} - \frac{e^{\hat{\beta}_0 + \hat{\beta}_x^\intercal X_i}}{1 + e^{\hat{\beta}_0 + \hat{\beta}_x^\intercal X_i}} \right),$$

where $\hat{\beta}_0$, $\hat{\beta}_z$, $\hat{\beta}_x$, $\hat{\beta}_{xz}$ are the maximum likelihood (ML) estimates.

Problem 10.5 (Ignorability versus strong ignorability)

We are asked to give an example where ignorability holds but strong ignorability does not; that is, $Y(z) \perp \!\!\! \perp Z \mid X$ for z = 0, 1 but $\{Y(0), Y(1)\} \perp \!\!\! \perp Z \mid X$ does not hold.

Without loss of generality, we omit X in the following example. Let $U \sim \text{Bernoulli}(0.5)$, and $V \sim \text{Bernoulli}(0.5)$ be independent of U. Define $Z = U \oplus V$ (i.e., the XOR operation), and

$$Y(0) = U, \quad Y(1) = V.$$

Then, we have

$$P(Z=1) = P(U \neq V) = 0.5, \quad P(Z=0) = P(U=V) = 0.5,$$

We can verify that

$$P(Y(0) = 1 \mid Z = 1) = P(U = 1 \mid U \neq V) = \frac{P(U = 1, V = 0)}{P(U \neq V)} = \frac{0.25}{0.5} = 0.5,$$

and similarly,

$$P(Y(0) = 1 \mid Z = 0) = P(U = 1 \mid U = V) = \frac{P(U = 1, V = 1)}{P(U = V)} = \frac{0.25}{0.5} = 0.5,$$

which implies that $Y(0) \perp \!\!\! \perp Z$. Analogously, we can show that $Y(1) \perp \!\!\! \perp Z$:

$$P(Y(1) = 1 \mid Z = 1) = P(V = 1 \mid U \neq V) = \frac{P(U = 0, V = 1)}{P(U \neq V)} = \frac{0.25}{0.5} = 0.5,$$

and

$$P(Y(1) = 1 \mid Z = 0) = P(V = 1 \mid U = V) = \frac{P(U = 1, V = 1)}{P(U = V)} = \frac{0.25}{0.5} = 0.5.$$

However, we have

$$P(Y(0) = 1, Y(1) = 1 \mid Z = 1) = P(U = 1, V = 1 \mid U \neq V) = 0,$$

and

$$P(Y(0) = 1, Y(1) = 1 \mid Z = 0) = P(U = 1, V = 1 \mid U = V) = 1,$$

which implies that $\{Y(0), Y(1)\}$ is not independent of Z. Thus, we have constructed an example where ignorability holds but strong ignorability does not.

Acronyms

ATE average treatment effect. 2

ML maximum likelihood. 2