

Solutions to Exercises of Chapter 5

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Problem 5.1 (Covariate balance in the CRE)

Consider a completely randomized experiment (CRE) with n units, among which n_1 are assigned to treatment ($Z_i = 1$) and n_0 to control ($Z_i = 0$). Let $X_i \in \{1, \dots, K\}$ be the stratum membership of unit i , where K is the number of strata. Let $n_{[k]} \equiv \sum_{i=1}^n \mathbf{1}(X_i = k)$ be the number of units in stratum k , $\pi_{[k]} \equiv n_{[k]}/n$ be the proportion of units in stratum k , and $n_{[k]z} \equiv \sum_{i=1}^n \mathbf{1}(X_i = k, Z_i = z)$ be the number of units in stratum k assigned to treatment z .

We are asked to show that

$$\mathbb{E} \left(\frac{n_{[k]1}}{n_1} - \frac{n_{[k]0}}{n_0} \right) = 0. \quad (5.2)$$

This can directly be shown by the following calculation:

$$\begin{aligned} & \mathbb{E} \left(\frac{n_{[k]1}}{n_1} - \frac{n_{[k]0}}{n_0} \right) \\ &= \mathbb{E} \left(\frac{n_{[k]1}}{n_1} \right) - \mathbb{E} \left(\frac{n_{[k]0}}{n_0} \right) \quad (\text{Linearity of } \mathbb{E}(\cdot)) \\ &= \frac{1}{n_1} \mathbb{E} \left(\sum_{i=1}^n Z_i \mathbf{1}(X_i = k) \right) - \frac{1}{n_0} \mathbb{E} \left(\sum_{i=1}^n (1 - Z_i) \mathbf{1}(X_i = k) \right) \quad (\text{Definition of } n_{[k]z}) \\ &= \frac{1}{n_1} \sum_{i=1}^n \mathbf{1}(X_i = k) \mathbb{E}(Z_i) - \frac{1}{n_0} \sum_{i=1}^n \mathbf{1}(X_i = k) \mathbb{E}(1 - Z_i) \quad (\text{Linearity of } \mathbb{E}(\cdot)) \\ &= \frac{1}{n_1} \sum_{i=1}^n \mathbf{1}(X_i = k) \frac{n_1}{n} - \frac{1}{n_0} \sum_{i=1}^n \mathbf{1}(X_i = k) \frac{n_0}{n} \quad (\mathbb{E}(Z_i) = n_1/n \text{ in the CRE}) \\ &= \frac{1}{n} \sum_{i=1}^n \mathbf{1}(X_i = k) - \frac{1}{n} \sum_{i=1}^n \mathbf{1}(X_i = k) \quad (\text{Simplification}) \\ &= 0. \end{aligned}$$

Problem 5.2 (Consequence of the constant propensity score)

Following the notation in Problem 5.1, let $e_{[k]} \equiv n_{[k]1}/n_{[k]}$ be the propensity score of stratum k . Define

$$\begin{aligned} \hat{\tau} &\equiv \frac{1}{n_1} \sum_{i=1}^n Z_i Y_i - \frac{1}{n_0} \sum_{i=1}^n (1 - Z_i) Y_i, \\ \hat{\tau}_S &\equiv \sum_{k=1}^K \pi_{[k]} \hat{\tau}_{[k]}, \end{aligned}$$

where

$$\hat{\tau}_{[k]} \equiv \frac{1}{n_{[k]1}} \sum_{i: X_i=k} Z_i Y_i - \frac{1}{n_{[k]0}} \sum_{i: X_i=k} (1 - Z_i) Y_i.$$

We are asked to show that when the propensity score is constant across strata,

$$\hat{\tau} = \hat{\tau}_S. \quad (5.3)$$

Notice that if the propensity score is constant across strata, then $e_{[k]} = n_{[k]1}/n_{[k]} = n_1/n$ for all $k = 1, \dots, K$. Therefore, we have

$$\begin{aligned} & \hat{\tau}_S \\ \equiv & \sum_{k=1}^K \pi_{[k]} \hat{\tau}_{[k]} && \text{(Definition of } \hat{\tau}_S \text{)} \\ = & \sum_{k=1}^K \frac{n_{[k]}}{n} \left(\frac{1}{n_{[k]1}} \sum_{i: X_i=k} Z_i Y_i - \frac{1}{n_{[k]0}} \sum_{i: X_i=k} (1 - Z_i) Y_i \right) && \text{(Definition of } \pi_{[k]} \text{ and } \hat{\tau}_{[k]} \text{)} \\ = & \frac{1}{n} \sum_{k=1}^K \left(\frac{n_{[k]}}{n_{[k]1}} \sum_{i: X_i=k} Z_i Y_i - \frac{n_{[k]}}{n_{[k]0}} \sum_{i: X_i=k} (1 - Z_i) Y_i \right) && \text{(Factor out } 1/n \text{)} \\ = & \frac{1}{n} \sum_{k=1}^K \left(\frac{n}{n_1} \sum_{i: X_i=k} Z_i Y_i - \frac{n}{n_0} \sum_{i: X_i=k} (1 - Z_i) Y_i \right) && \text{(Use } e_{[k]} = n_1/n \text{)} \\ = & \frac{1}{n_1} \sum_{k=1}^K \sum_{i: X_i=k} Z_i Y_i - \frac{1}{n_0} \sum_{k=1}^K \sum_{i: X_i=k} (1 - Z_i) Y_i && \text{(Factor out } n/n_z \text{)} \\ = & \frac{1}{n_1} \sum_{i=1}^n Z_i Y_i - \frac{1}{n_0} \sum_{i=1}^n (1 - Z_i) Y_i && \text{(Combine sums)} \\ \equiv & \hat{\tau}, && \text{(Definition of } \hat{\tau} \text{)} \end{aligned}$$

which completes the proof.

Problem 5.3 (Consequence of constant individual causal effects)

Problem 5.4 (Compare the CRE and SRE)

We are asked to compare the sampling variance of the difference-in-means estimator under the CRE and the stratified randomized experiment (SRE).

Suppose that we have a dataset from a SRE with K strata. Assume that $e_{[k]} = n_{[k]1}/n_{[k]}$ is constant across strata.

We are asked to show that when $n_{[k]}$ is large enough for all $k = 1, \dots, K$, the difference between $\text{Var}_{\text{CRE}}(\hat{\tau})$ and $\text{Var}_{\text{SRE}}(\hat{\tau}_S)$ is always non-negative:

$$\begin{aligned} & \text{Var}_{\text{CRE}}(\hat{\tau}) - \text{Var}_{\text{SRE}}(\hat{\tau}_S) \\ = & \sum_{k=1}^K \left[\frac{\pi_{[k]}}{n_1} (\bar{Y}_{[k]}(1) - \bar{Y}(1))^2 + \frac{\pi_{[k]}}{n_0} (\bar{Y}_{[k]}(0) - \bar{Y}(0))^2 - \frac{\pi_{[k]}}{n} (\tau_{[k]} - \tau)^2 \right] \geq 0. \end{aligned}$$

For the simplicity of notation, let $a_{[k]} \equiv \bar{Y}_{[k]}(1) - \bar{Y}(1)$ and $b_{[k]} \equiv \bar{Y}_{[k]}(0) - \bar{Y}(0)$. Then, $a_{[k]} - b_{[k]} = \tau_{[k]} - \tau$. Therefore, we have

$$\begin{aligned}
& \text{Var}_{\text{CRE}}(\hat{\tau}) - \text{Var}_{\text{SRE}}(\hat{\tau}_{\text{S}}) \\
&= \sum_{k=1}^K \left[\frac{\pi_{[k]}}{n_1} a_{[k]}^2 + \frac{\pi_{[k]}}{n_0} b_{[k]}^2 - \frac{\pi_{[k]}}{n} (a_{[k]} - b_{[k]})^2 \right] && \text{(Substitution)} \\
&= \sum_{k=1}^K \pi_{[k]} \left[\frac{1}{n_1} a_{[k]}^2 + \frac{1}{n_0} b_{[k]}^2 - \frac{1}{n} (a_{[k]}^2 - 2a_{[k]}b_{[k]} + b_{[k]}^2) \right] && \text{(Expand the square)} \\
&= \sum_{k=1}^K \pi_{[k]} \left[\left(\frac{1}{n_1} - \frac{1}{n} \right) a_{[k]}^2 + \left(\frac{1}{n_0} - \frac{1}{n} \right) b_{[k]}^2 + \frac{2}{n} a_{[k]}b_{[k]} \right] && \text{(Rearrangement)} \\
&= \sum_{k=1}^K \pi_{[k]} \left(\frac{n_0}{nn_1} a_{[k]}^2 + \frac{n_1}{nn_0} b_{[k]}^2 + \frac{2}{n} a_{[k]}b_{[k]} \right) && \text{(Simplification)} \\
&= \sum_{k=1}^K \frac{\pi_{[k]}}{n} \left(\frac{n_0}{n_1} a_{[k]}^2 + \frac{n_1}{n_0} b_{[k]}^2 + 2a_{[k]}b_{[k]} \right) && \text{(Factor out } 1/n) \\
&= \sum_{k=1}^K \frac{\pi_{[k]}}{n} \left(\sqrt{\frac{n_0}{n_1}} a_{[k]} + \sqrt{\frac{n_1}{n_0}} b_{[k]} \right)^2 && \text{(Rewrite the first two terms)} \\
&\geq 0. && \text{(Square is non-negative)}
\end{aligned}$$

This completes the proof.

Problem 5.5 (From the CRE to the SRE)

In a CRE with n units, among which n_1 are assigned to treatment ($Z_i = 1$) and n_0 to control ($Z_i = 0$). Suppose that we have a categorical covariate $X_i \in \{1, \dots, K\}$, where K is the number of strata. Now, the quantities $n_{[k]1}$ and $n_{[k]0}$ will be random variables because the number of units in stratum k assigned to treatment z will vary across randomizations. Let $\mathbf{Z} \equiv (Z_1, \dots, Z_n)$ be the treatment assignment vector, and $\mathbf{n} \equiv \{n_{[k]1}, n_{[k]0}\}_{k=1}^K$ be the collection of the numbers of units in each stratum assigned to treatment and control.

We are asked to show that

$$P_{\text{CRE}}(\mathbf{Z} = \mathbf{z} \mid \mathbf{n}) = \frac{1}{\prod_{k=1}^K \binom{n_{[k]}}{n_{[k]1}}}.$$

That is, conditional on \mathbf{n} , the treatment assignment \mathbf{Z} is equivalent to that from a SRE.

Note that for any realization of \mathbf{n} , the number of possible treatment assignments \mathbf{z} that satisfy \mathbf{n} is exactly

$$\prod_{k=1}^K \binom{n_{[k]}}{n_{[k]1}},$$

and each of these treatment assignments has the same probability $1/\binom{n}{n_1}$ under the CRE.

Therefore, we have

$$P_{\text{CRE}}(\mathbf{n}) = \frac{\prod_{k=1}^K \binom{n_{[k]}}{n_{[k]1}}}{\binom{n}{n_1}}.$$

Now, we can calculate $P_{\text{CRE}}(\mathbf{Z} = \mathbf{z} \mid \mathbf{n})$ as follows:

$$\begin{aligned} & P_{\text{CRE}}(\mathbf{Z} = \mathbf{z} \mid \mathbf{n}) \\ &= \frac{P_{\text{CRE}}(\mathbf{Z} = \mathbf{z}, \mathbf{n})}{P_{\text{CRE}}(\mathbf{n})} && \text{(Definition of conditional probability)} \\ &= \frac{P_{\text{CRE}}(\mathbf{Z} = \mathbf{z})}{P_{\text{CRE}}(\mathbf{n})} && (\mathbf{n} \text{ is determined by } \mathbf{Z}) \\ &= \frac{1 / \binom{n}{n_1}}{\prod_{k=1}^K \binom{n_{[k]}}{n_{[k]1}} / \binom{n}{n_1}} && \text{(Substitution)} \\ &= \frac{1}{\prod_{k=1}^K \binom{n_{[k]}}{n_{[k]1}}}, \end{aligned}$$

which completes the proof.

Problem 5.6 (More FRTs for Section 5.2.2)

Acronyms

CRE completely randomized experiment. [1–3](#)
SRE stratified randomized experiment. [2](#), [3](#)