Bayesian_HW1

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3.1 a

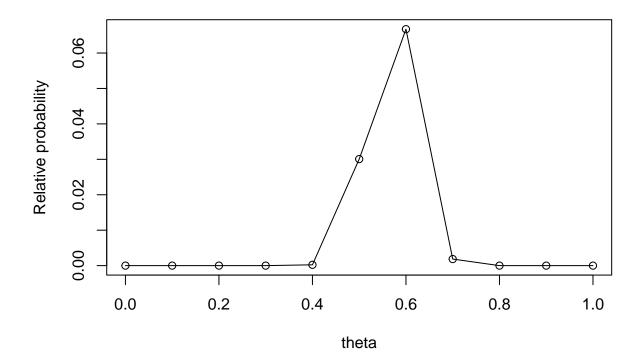
$$\Pr(Y_1 = y_1, ... Y_{100} = y_{100} | \theta) = \prod_{i=1}^{100} \theta^{y_i} (1 - \theta)^{100 - y_i}$$
$$= \theta^{\sum_{i=1}^{100} y_i} (1 - \theta)^{100 - \sum_{i=1}^{100} y_i}$$
$$\Pr(\Sigma Y_i = y | \theta) = \binom{100}{y} \theta^y (1 - \theta)^{100 - y}$$

b

```
prob <- function(n, y, theta){
    x = n-y
    return(factorial(n)/(factorial(y)*factorial(x))* theta**y * (1-theta)**(x))
}
theta = seq(0,1,0.1)
probs = prob(100,57,theta=theta)
data.frame(theta=theta, Relative_probability=round(probs, 4))</pre>
```

```
##
      theta Relative_probability
## 1
        0.0
                           0.0000
## 2
        0.1
                           0.0000
## 3
                           0.0000
        0.2
## 4
        0.3
                           0.0000
## 5
        0.4
                           0.0002
## 6
        0.5
                           0.0301
## 7
        0.6
                           0.0667
## 8
        0.7
                           0.0019
## 9
        0.8
                           0.0000
## 10
        0.9
                           0.0000
## 11
                           0.0000
        1.0
```

```
plot(theta, probs, type = "o",ylab = "Relative probability")
```



 \mathbf{c}

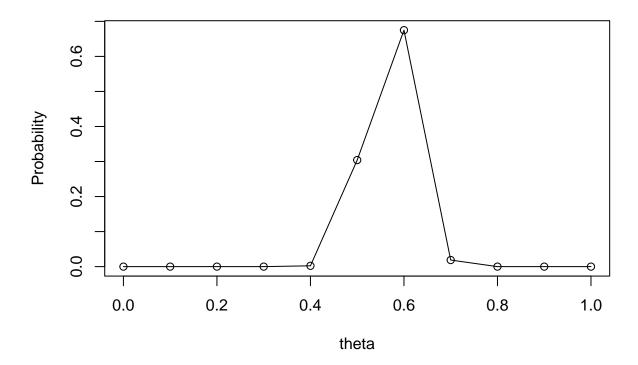
$$p(\theta|\Sigma_{i=1}^{n}Y_{i}=57) = \frac{p(\Sigma_{i=1}^{n}Y_{i}=57|\theta)p(\theta)}{p(\Sigma_{i=1}^{n}Y_{i}=57)} \propto p(\Sigma_{i=1}^{n}Y_{i}=57|\theta)$$

We can get the posterior distribution by the relative probabilities of question 3.1b

```
new_probs = probs/sum(probs)
data.frame(theta=theta, probability=round(new_probs, 4))
```

```
##
      theta probability
## 1
         0.0
                   0.0000
## 2
         0.1
                  0.0000
## 3
         0.2
                   0.0000
         0.3
                   0.0000
##
##
         0.4
                   0.0023
##
        0.5
                  0.3041
##
         0.6
                   0.6749
## 8
         0.7
                   0.0187
## 9
         0.8
                   0.0000
                   0.0000
## 10
         0.9
## 11
         1.0
                   0.0000
```

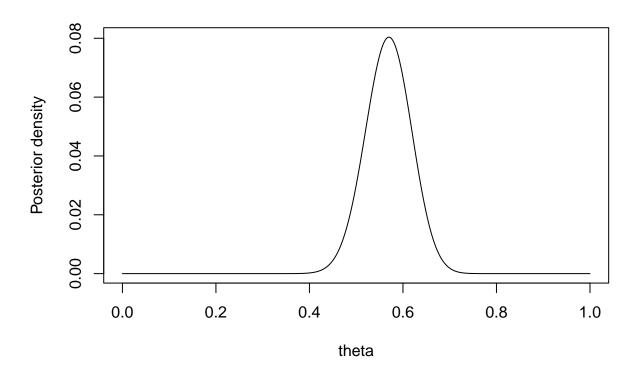
```
plot(theta, new_probs, type = "o",ylab = "Probability")
```



 \mathbf{d}

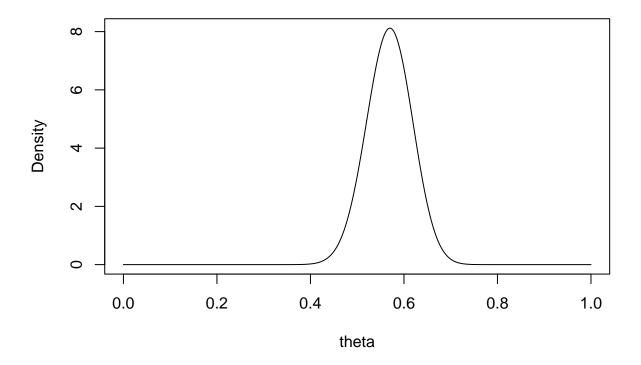
$$p(\theta) * \Pr(\Sigma_{i=1}^{n} Y_i = 57 | \theta) = p(\theta) \binom{100}{57} \theta^{57} (1 - \theta)^{100 - 57} = \binom{100}{57} \theta^{57} (1 - \theta)^{100 - 57}$$

```
theta=seq(0,1,1e-4)
probs = prob(n=100, y=57, theta=theta)
plot(theta, probs, type = "l", ylab = "Posterior density")
```



 \mathbf{e}

```
betas = dbeta(seq(0,1,1e-4), shape1=58, shape2=44)
plot(theta, betas, type="l", ylab="Density")
```



Discussion: 1b and 1e is the posterior before normalization. $\,$

1c is the posterior after scaling

e is the posterior density given a uniform prior Beta(1,1)

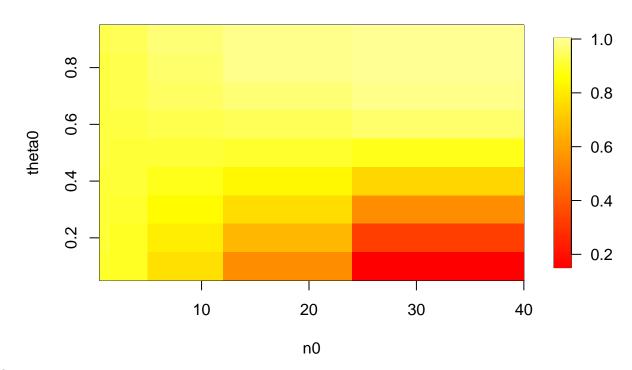
```
theta0s = seq(0.1,0.9, 0.1)
n0s = c(1,2,8,16,32)
w = 0.5
n = 100
y = 57

prior_as <- n0s %*% t(theta0s)
prior_bs <- n0s %*% t((1-theta0s))

post_prob <- matrix(nrow = length(n0s), ncol = length(theta0s))
post_a <- prior_as + y
post_b <- prior_bs + n - y

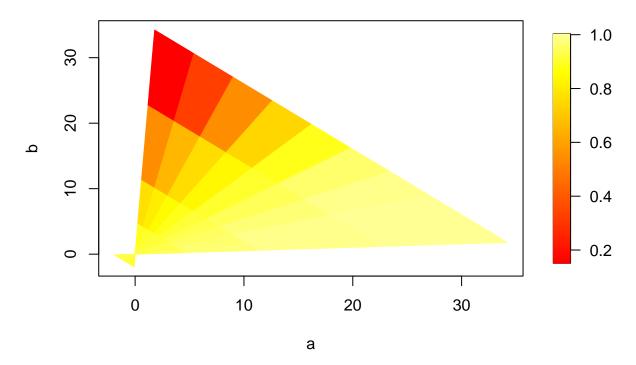
for (i in 1:length(n0s)){
    for (j in 1:length(theta0s)){
        post_prob[i,j] <- 1-pbeta(w,post_a[i,j],post_b[i,j])
    }
}</pre>
```

p(theta > 0.5 | y)



3.2

p(theta > 0.5 | y)



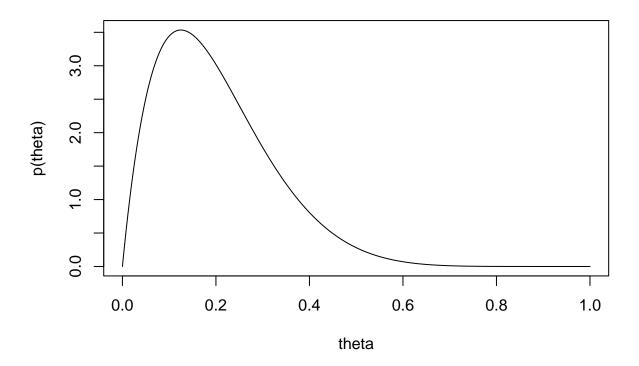
From the 2D contour plot, we could say that most priors result in posterior theta > 0.5 except those with very large n0 and very small theta0.

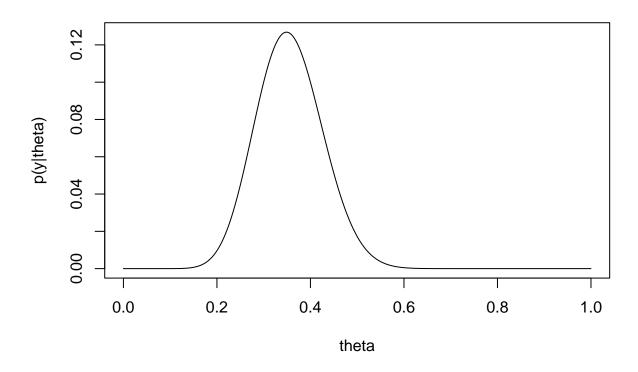
3.4

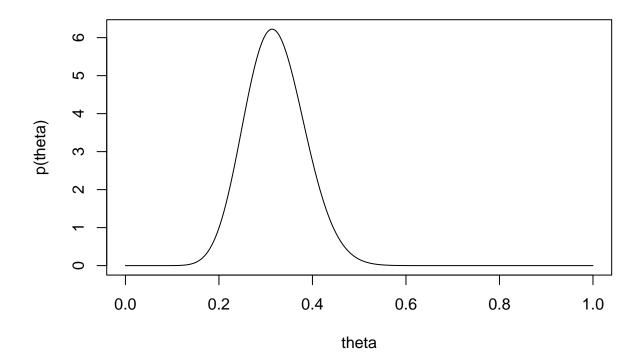
a

```
\begin{split} p(\theta) \sim &Beta(2,8) \\ p(y|\theta) \sim &Binomial(43,\theta) \\ p(\theta|y) \sim &Beta(15+2,43-15+8) \equiv Beta(17,36) \\ E(\theta|y) = &17/(17+36) \approx 0.3207 \\ Mode = &\arg\max f(\theta|y) = &(17-1)/(17+36-2) \approx 0.3137 \\ Var(\theta|y) = &[17/(17+36) \cdot (36/17+36)]/(17+36+1) \approx 0.0635 \\ 95\% \text{Confidence Interval} : &(0.2033,0.4510) \end{split}
```

```
y = 15
n = 43
a = 2
b = 8
theta = seq(0,1,1e-3)
ploting <- function(theta,y,n,a,b){
  new_a = y + a
  new_b = n-y+b
  plot(theta, dbeta(theta,a,b), type="l", ylab="p(theta)")
  plot(theta, dbinom(y,n,theta), type="l", ylab="p(y|theta)")</pre>
```







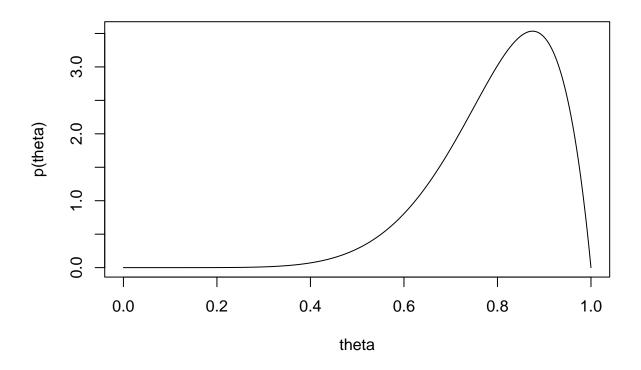
```
## mean mode sd Confidence_Interval ## Beta(2,8) 0.3207547 0.3137255 0.0635189 0.2033,0.451
```

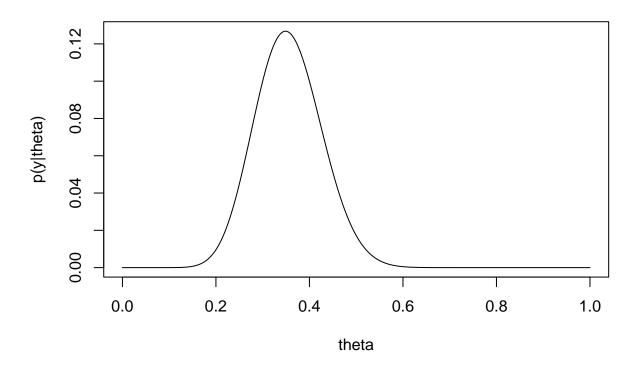
 \mathbf{b}

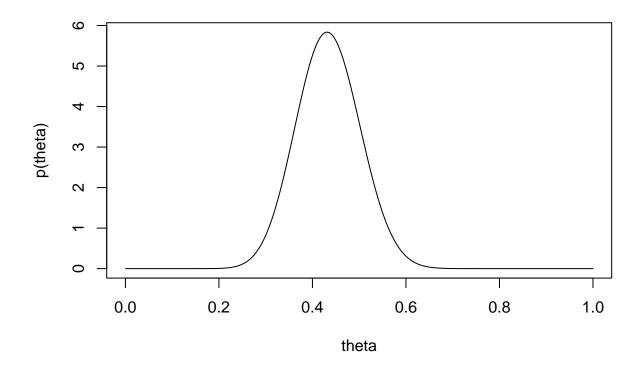
Since they're the same but a & b switched, I used the exact same function created from q3.4a for $p(\theta|y) \sim Beta(23,30)$

$$\begin{split} Mean &\approx 0.4340 \\ Mode &\approx 0.4314 \\ sd &\approx 0.0674 \\ 95\% CI: (0.3047, 0.5680) \end{split}$$

ploting(theta=theta, y=y,n=n,a=b,b=a)



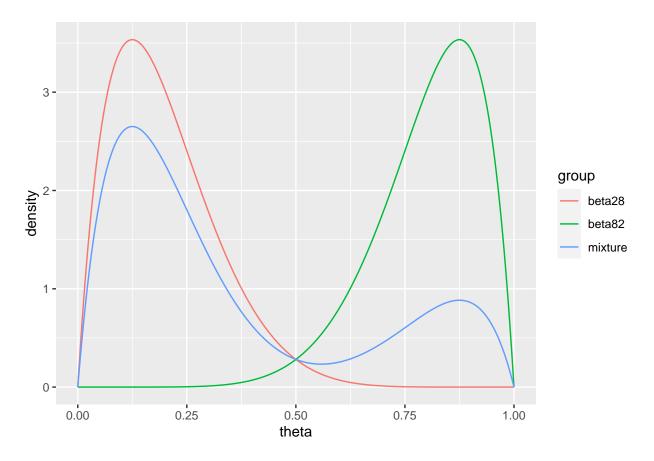




```
## mean mode sd Confidence_Interval ## Beta(8,2) 0.4339623 0.4313725 0.06744532 0.3047,0.568
```

 \mathbf{c}

The mixture shows a split prior opinions at around 0.125 and around 0.875.



 $\frac{\mathbf{d}}{\mathbf{i}}$

gamma(10)/(gamma(2)*gamma(8))/4

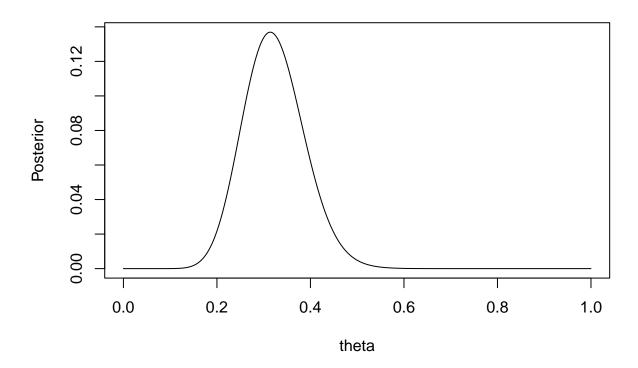
[1] 18

$$p(\theta)*p(y|\theta) = 18[3\theta(1-\theta)^7 + \theta^7(1-\theta)]*[\binom{43}{15}\theta^{15}(1-\theta)^{28}] = 18\binom{43}{15}[3\theta^{16}(1-\theta)^{35} + \theta^{22}(1-\theta)^{29}] = 18\binom{43}{15}\theta^{16}(1-\theta)^{29}[3(1-\theta)^{6} + \theta^{22}(1-\theta)^{29}] = 18\binom{43}{15}\theta^{16}(1-\theta)^{29}[3(1-\theta)^{29}] = 18\binom{43}{15}\theta^{16}[3(1-\theta)^{29}] = 18\binom{15}{15}\theta^{16}[3(1-\theta)^{29}] = 18\binom{15}{15}\theta^{16}[3(1-\theta)^{29}]$$

ii

Since $p(\theta|y) \propto p(\theta) * p(y|\theta)$, the posterior distribution should be a mixture of the posteriors in above question 3.4a and 3.4b.

iii



max(posterior)

[1] 0.136914

(mode = theta[which.max(posterior)]) # approximated mode

[1] 0.314

The posterior mode is approximately 0.314, which is closer to the mode of question a. (with prior Beta(2,8)) e

Let $p(\theta|y) = w_1 * f_1(\theta|y) + (1 - w_1)f_2(\theta|y)$

I have no idea what to do next.

3.9