

# MED Gear Comparison Analysis

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**Updated:** 11 Oct. 2018, 14:52, using R version 3.4.4 (2018-03-15)

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## Introduction

This document presents full statistical results and describes the R code used to conduct the analysis in Wainwright et al. “*Effect of a mammal excluder device on catches of small pelagic fishes*”. Data and scripts for the analysis are available in the R package ‘GearComparisonAnalysis2018’ available on [GitHub](#).

**NOTE:** “R” is open source software freely available from the [R Project](#).

## Part 1: Read and Summarize data

First, read in the data.

```
#data.dir <- '.' #Directory where the data files resides
MMEDdata <- read.csv(system.file('extdata', 'AllSppCounts&Lengths.csv',
                                package = "GearComparisonAnalysis2018"))
# print(summary(MMEDdata))
```

Fix some data problems. Create consistent haul numbers, filter out “experimental” hauls, and change Chinook and coho salmon “species” names to reflect age groups. Some species are renamed to make nicer plot labels, and months and times are recoded.

```
# Haul ID is last three characters of station code:
MMEDdata$Station <- as.character(MMEDdata$Station) #remove factor levels
MMEDdata$Haul <- with(MMEDdata, substr(Station, nchar(Station)-2, nchar(Station)))
# Excluder code (Y or N) is last character of station code:
MMEDdata$Excluder <- with(MMEDdata, substr(Station, nchar(Station), nchar(Station)))
# Filter out "experimental" hauls:
MMEDdata <- MMEDdata[-grep('*X$', MMEDdata$Haul), ]
# Recode MMED types to Standard short labels
### NOTE: work around database error where some records with Excluder code "N"
### have MMED code "Yes; Up"
MMEDdata$MMED <- as.character(MMEDdata$MMED) #remove factor levels
MMEDdata$MMED[MMEDdata$Excluder == "N"] <- "None"
MMEDdata$MMED[MMEDdata$Excluder == "Y" & grepl("Up", MMEDdata$MMED)] <- "Up"
MMEDdata$MMED[MMEDdata$Excluder == "Y" & grepl("Down", MMEDdata$MMED)] <- "Down"
# Fix a few species names, add age-classes for Chinook & coho
MMEDdata$Species <- toupper(as.character(MMEDdata$Species)) #all upper case
MMEDdata$Species[MMEDdata$Species %in% 'CALIFORNIA MARKET SQUID'] <- 'MARKET SQUID'
MMEDdata$Species[grepl('SMELT', MMEDdata$Species)] <- 'SMELT SPP.'
MMEDdata$SpecAge <- as.character(MMEDdata$SpecAge)
.index <- MMEDdata$SpecAge %in% 'CHINOOK SALMON'
.newnames <- paste('CHINOOK', MMEDdata$AgeGp[.index])
```

```

MMEDdata$SpecAge[.index] <- .newnames
.index <- MMEDdata$SpecAge %in% 'COHO SALMON'
.newnames <- paste('COHO', MMEDdata$AgeGp[.index])
MMEDdata$SpecAge[.index] <- .newnames
MMEDdata$SpecAge[MMEDdata$SpecAge %in%
  c('CHINOOK subadult/adult', 'CHINOOK mixed age juvenile')] <- 'CHINOOK subadult'
MMEDdata$SpecAge[MMEDdata$SpecAge %in% 'COHO subadult/adult'] <- 'COHO subadult'
MMEDdata$SpecAge <- factor(as.character(MMEDdata$SpecAge))
# print(summary(MMEDdata))

```

Divide the hauls into analytic “blocks” which are combinations of Date X Location. There are 12 blocks (“A” through “L”), with the following number of hauls in each:

- May 2011: block A : 10, B : 10
- July 2011: C : 4, D : 4, E : 10
- Jun-Jul 2014: F : 4, G : 4, H : 8
- July 2015: I : 8, J : 8, K : 8, L : 8

```

# sequential along the sorted haul numbers,
blocks <- as.factor(c(rep('A',10), rep('B',10), rep('C', 4), rep('D', 4),
  rep('E',10), rep('F', 4), rep('G', 4), rep('H', 8),
  rep('I', 8), rep('J', 8), rep('K', 8), rep('L', 8)))
sortHauls <- sort(unique(MMEDdata$Haul))
MMEDdata$Block <- blocks[match(MMEDdata$Haul, sortHauls)]
# print(with(MMEDdata, t(apply(as.character(Block), list(Haul,MMED), unique))))

```

Then, create a summary table of total catch by species and gear type (None==“without MED”, Down==“with downward MED”, Up==“with upward MED”), ...

```

cat('\n*** Total Catch By Species and Gear Type ***\n')

```

```
##
```

```
## *** Total Catch By Species and Gear Type ***
```

```

tab1 <- with(MMEDdata, tapply(Number, list(SpecAge, MMED), sum, na.rm=T))
tab1[is.na(tab1)] <- 0 #Missing values are actually zero counts
tab1 <- cbind(tab1, Total=apply(tab1, 1, sum))
print(tab1[, c('Down', 'Up', 'None', 'Total')])

```

```
##
```

	Down	Up	None	Total
## AMERICAN SHAD	6	0	6	12
## BLACK ROCKFISH	0	1	5	6
## BLUE SHARK	0	0	1	1
## CABEZON	0	1	0	1
## CHINOOK subadult	28	36	59	123
## CHINOOK subyearling	191	713	1092	1996
## CHINOOK yearling	90	223	364	677
## CHUM SALMON	54	0	173	227
## COHO subadult	41	25	61	127
## COHO yearling	38	24	194	256
## COMB JELLIES	0	0	0	0
## CUTTHROAT TROUT	0	0	4	4
## EGGYOLK JELLY	29	2	41	72
## FISH	0	2	0	2
## FLATFISHES	0	1	3	4
## HORMIPHORA CUCUMIS	0	0	0	0

## HYBRID STEELHEAD AND CUTTHROAT TROUT	0	0	1	1
## JACK MACKEREL	0	9	13	22
## LINGCOD	0	0	1	1
## LION'S MANE JELLY	0	0	3	3
## MARKET SQUID	2160	258	8285	10703
## MOON JELLY	10259	2	1377	11638
## NORTHERN ANCHOVY	2	200	3073	3275
## OCEAN SUNFISH	0	1	0	1
## PACIFIC CHUB MACKEREL	0	3	2	5
## PACIFIC HERRING	1755	58	563	2376
## PACIFIC POMPAÑO	7	0	5	12
## PACIFIC SANDDAB	0	0	4	4
## PACIFIC SANDFISH	0	0	1	1
## PACIFIC SARDINE	0	45	75	120
## PACIFIC SPINY DOGFISH	0	3	5	8
## PACIFIC STAGHORN SCULPIN	0	1	0	1
## PACIFIC TOMCOD	0	15	22	37
## PINK SALMON	0	1	1	2
## REX SOLE	0	0	1	1
## SALPS	0	0	0	0
## SEA NETTLE	21	1285	1422	2728
## SHINER PERCH	1	0	0	1
## SMELT SPP.	120	464	2221	2805
## SOCKEYE SALMON	0	0	22	22
## STARRY FLOUNDER	0	3	2	5
## STEELHEAD	1	3	6	10
## THRESHER SHARK	0	0	1	1
## TOPE	0	0	7	7
## WATER JELLY	53291	528	71472	125291
## WESTERN RIVER LAMPREY	6	2	16	24
## WOLF-EEL	6	2	29	37
## YELLOWTAIL ROCKFISH	0	0	1	1

... and the same for number measured and subsampling rate (for size-selectivity analysis), ...

```
cat('\n*** Total Number Measured By Species and Gear Type ***\n')
```

```
##
```

```
## *** Total Number Measured By Species and Gear Type ***
```

```
tab4 <- with(MMEDdata[!is.na(MMEDdata$Length), ],
  tapply(Number, list(SpecAge, MMED), sum, na.rm=T))
tab4[is.na(tab4)] <- 0 #Missing values are actually zero counts
tab4 <- cbind(tab4, Total=apply(tab4, 1, sum))
print(tab4[, c('Down', 'Up', 'None', 'Total')])
```

##	Down	Up	None	Total
## AMERICAN SHAD	6	0	6	12
## BLACK ROCKFISH	0	1	5	6
## BLUE SHARK	0	0	1	1
## CABEZON	0	1	0	1
## CHINOOK subadult	28	36	59	123
## CHINOOK subyearling	142	713	1035	1890
## CHINOOK yearling	90	223	363	676
## CHUM SALMON	54	0	173	227
## COHO subadult	41	25	61	127

## COHO yearling	38	24	194	256
## COMB JELLIES	0	0	0	0
## CUTTHROAT TROUT	0	0	4	4
## EGGYOLK JELLY	23	2	15	40
## FISH	0	0	0	0
## FLATFISHES	0	0	2	2
## HORMIPHORA CUCUMIS	0	0	0	0
## HYBRID STEELHEAD AND CUTTHROAT TROUT	0	0	1	1
## JACK MACKEREL	0	9	13	22
## LINGCOD	0	0	1	1
## LION'S MANE JELLY	0	0	3	3
## MARKET SQUID	347	189	582	1118
## MOON JELLY	111	2	149	262
## NORTHERN ANCHOVY	2	152	241	395
## OCEAN SUNFISH	0	1	0	1
## PACIFIC CHUB MACKEREL	0	3	2	5
## PACIFIC HERRING	41	31	73	145
## PACIFIC POMPAÑO	7	0	5	12
## PACIFIC SANDDAB	0	0	4	4
## PACIFIC SANDFISH	0	0	1	1
## PACIFIC SARDINE	0	45	54	99
## PACIFIC SPINY DOGFISH	0	3	5	8
## PACIFIC STAGHORN SCULPIN	0	1	0	1
## PACIFIC TOMCOD	0	15	22	37
## PINK SALMON	0	1	1	2
## REX SOLE	0	0	1	1
## SALPS	0	0	0	0
## SEA NETTLE	21	536	623	1180
## SHINER PERCH	1	0	0	1
## SMELT SPP.	44	60	99	203
## SOCKEYE SALMON	0	0	22	22
## STARRY FLOUNDER	0	3	2	5
## STEELHEAD	1	3	6	10
## THRESHER SHARK	0	0	1	1
## TOPE	0	0	7	7
## WATER JELLY	385	306	639	1330
## WESTERN RIVER LAMPREY	6	2	16	24
## WOLF-EEL	6	2	29	37
## YELLOWTAIL ROCKFISH	0	0	1	1

```
cat('\n*** Average Subsampling Rate By Species and Gear Type ***\n')
```

```
##
```

```
## *** Average Subsampling Rate By Species and Gear Type ***
```

```
tab5 <- round(tab4/tab1,2)
```

```
print(tab5[ , c('Down', 'Up', 'None', 'Total')])
```

##	Down	Up	None	Total
## AMERICAN SHAD	1.00	NaN	1.00	1.00
## BLACK ROCKFISH	NaN	1.00	1.00	1.00
## BLUE SHARK	NaN	NaN	1.00	1.00
## CABEZON	NaN	1.00	NaN	1.00
## CHINOOK subadult	1.00	1.00	1.00	1.00
## CHINOOK subyearling	0.74	1.00	0.95	0.95

## CHINOOK yearling	1.00	1.00	1.00	1.00
## CHUM SALMON	1.00	NaN	1.00	1.00
## COHO subadult	1.00	1.00	1.00	1.00
## COHO yearling	1.00	1.00	1.00	1.00
## COMB JELLIES	NaN	NaN	NaN	NaN
## CUTTHROAT TROUT	NaN	NaN	1.00	1.00
## EGGYOLK JELLY	0.79	1.00	0.37	0.56
## FISH	NaN	0.00	NaN	0.00
## FLATFISHES	NaN	0.00	0.67	0.50
## HORMIPHORA CUCUMIS	NaN	NaN	NaN	NaN
## HYBRID STEELHEAD AND CUTTHROAT TROUT	NaN	NaN	1.00	1.00
## JACK MACKEREL	NaN	1.00	1.00	1.00
## LINGCOD	NaN	NaN	1.00	1.00
## LION'S MANE JELLY	NaN	NaN	1.00	1.00
## MARKET SQUID	0.16	0.73	0.07	0.10
## MOON JELLY	0.01	1.00	0.11	0.02
## NORTHERN ANCHOVY	1.00	0.76	0.08	0.12
## OCEAN SUNFISH	NaN	1.00	NaN	1.00
## PACIFIC CHUB MACKEREL	NaN	1.00	1.00	1.00
## PACIFIC HERRING	0.02	0.53	0.13	0.06
## PACIFIC POMPAÑO	1.00	NaN	1.00	1.00
## PACIFIC SANDDAB	NaN	NaN	1.00	1.00
## PACIFIC SANDFISH	NaN	NaN	1.00	1.00
## PACIFIC SARDINE	NaN	1.00	0.72	0.82
## PACIFIC SPINY DOGFISH	NaN	1.00	1.00	1.00
## PACIFIC STAGHORN SCULPIN	NaN	1.00	NaN	1.00
## PACIFIC TOMCOD	NaN	1.00	1.00	1.00
## PINK SALMON	NaN	1.00	1.00	1.00
## REX SOLE	NaN	NaN	1.00	1.00
## SALPS	NaN	NaN	NaN	NaN
## SEA NETTLE	1.00	0.42	0.44	0.43
## SHINER PERCH	1.00	NaN	NaN	1.00
## SMELT SPP.	0.37	0.13	0.04	0.07
## SOCKEYE SALMON	NaN	NaN	1.00	1.00
## STARRY FLOUNDER	NaN	1.00	1.00	1.00
## STEELHEAD	1.00	1.00	1.00	1.00
## THRESHER SHARK	NaN	NaN	1.00	1.00
## TOPE	NaN	NaN	1.00	1.00
## WATER JELLY	0.01	0.58	0.01	0.01
## WESTERN RIVER LAMPREY	1.00	1.00	1.00	1.00
## WOLF-EEL	1.00	1.00	1.00	1.00
## YELLOWTAIL ROCKFISH	NaN	NaN	1.00	1.00

... and, do the same for frequency of catch.

```
cat('\n*** Frequency of Catch By Species and Gear Type ***\n')
```

```
##
```

```
## *** Frequency of Catch By Species and Gear Type ***
```

```
.tmp.all <- with(MMEDdata, table(SpecAge, Haul))
.tmp.std <- with(MMEDdata[MMEDdata$MMED=='None', ], table(SpecAge, Haul))
.tmp.up <- with(MMEDdata[MMEDdata$MMED=='Up', ], table(SpecAge, Haul))
.tmp.dwn <- with(MMEDdata[MMEDdata$MMED=='Down', ], table(SpecAge, Haul))
tab2 <- cbind(apply(.tmp.dwn>0, 1, sum),      # Num. occurrences in STD
              apply(.tmp.up>0, 1, sum),      # . . .
```

```

      apply(.tmp.std>0, 1, sum),      # . . .
      apply(.tmp.all>0, 1, sum))     # Total Num. occurrences
colnames(tab2) <- c('Down','Up','None','Total')
print(tab2)

```

##	Down	Up	None	Total
## AMERICAN SHAD	2	0	3	5
## BLACK ROCKFISH	0	1	2	3
## BLUE SHARK	0	0	1	1
## CABEZON	0	1	0	1
## CHINOOK subadult	13	19	27	59
## CHINOOK subyearling	14	18	31	63
## CHINOOK yearling	13	16	37	66
## CHUM SALMON	12	0	21	33
## COHO subadult	14	11	25	50
## COHO yearling	13	15	32	60
## COMB JELLIES	0	2	2	4
## CUTTHROAT TROUT	0	0	3	3
## EGGYOLK JELLY	15	2	8	25
## FISH	0	1	0	1
## FLATFISHES	0	1	4	5
## HORMIPHORA CUCUMIS	0	0	1	1
## HYBRID STEELHEAD AND CUTTHROAT TROUT	0	0	1	1
## JACK MACKEREL	0	3	2	5
## LINGCOD	0	0	1	1
## LION'S MANE JELLY	0	0	1	1
## MARKET SQUID	16	16	36	68
## MOON JELLY	16	2	15	33
## NORTHERN ANCHOVY	2	8	10	20
## OCEAN SUNFISH	0	1	0	1
## PACIFIC CHUB MACKEREL	0	2	2	4
## PACIFIC HERRING	4	2	10	16
## PACIFIC POMPAÑO	6	0	1	7
## PACIFIC SANDDAB	0	0	3	3
## PACIFIC SANDFISH	0	0	1	1
## PACIFIC SARDINE	0	3	4	7
## PACIFIC SPINY DOGFISH	0	3	5	8
## PACIFIC STAGHORN SCULPIN	0	1	0	1
## PACIFIC TOMCOD	0	4	5	9
## PINK SALMON	0	1	1	2
## REX SOLE	0	0	1	1
## SALPS	0	6	5	11
## SEA NETTLE	6	17	26	49
## SHINER PERCH	1	0	0	1
## SMELT SPP.	7	2	9	18
## SOCKEYE SALMON	0	0	7	7
## STARRY FLOUNDER	0	2	2	4
## STEELHEAD	1	3	3	7
## THRESHER SHARK	0	0	1	1
## TOPE	0	0	3	3
## WATER JELLY	16	17	29	62
## WESTERN RIVER LAMPREY	5	2	9	16
## WOLF-EEL	4	2	14	20
## YELLOWTAIL ROCKFISH	0	0	1	1

Also, look at species by cruise to see which are “regularly caught.” The final two columns give the number of cruises with nonzero catch (“Ngt0”) and catch > 1 (“Ngt1”) for each species.

```
cat('\n*** Total Catch By Species and Cruise ***\n')
```

```
##
```

```
## *** Total Catch By Species and Cruise ***
```

```
tab3 <- with(MMEDdata, tapply(Number, list(SpecAge, Cruise), sum, na.rm=T))
tab3[is.na(tab3)] <- 0 #Missing values are actually zero counts
ngt0 <- apply(tab3>0, 1, sum)
ngt1 <- apply(tab3>1, 1, sum)
tab3 <- cbind(tab3, Ngt0=ngt0, Ngt1=ngt1)
print(tab3)
```

##	41	43	50	53	Ngt0	Ngt1
## AMERICAN SHAD	0	1	0	11	2	1
## BLACK ROCKFISH	0	1	1	4	3	1
## BLUE SHARK	0	1	0	0	1	0
## CABEZON	0	0	1	0	1	0
## CHINOOK subadult	21	42	11	49	4	4
## CHINOOK subyearling	8	1451	53	484	4	4
## CHINOOK yearling	472	9	25	171	4	4
## CHUM SALMON	19	2	1	205	4	3
## COHO subadult	6	31	34	56	4	4
## COHO yearling	124	16	20	96	4	4
## COMB JELLIES	0	0	0	0	0	0
## CUTTHROAT TROUT	3	1	0	0	2	1
## EGGYOLK JELLY	1	0	1	70	3	1
## FISH	2	0	0	0	1	1
## FLATFISHES	1	0	0	3	2	1
## HORMIPHORA CUCUMIS	0	0	0	0	0	0
## HYBRID STEELHEAD AND CUTTHROAT TROUT	0	0	0	1	1	0
## JACK MACKEREL	0	22	0	0	1	1
## LINGCOD	1	0	0	0	1	0
## LION'S MANE JELLY	0	0	3	0	1	1
## MARKET SQUID	126	899	65	9613	4	4
## MOON JELLY	0	1	1	11636	3	1
## NORTHERN ANCHOVY	0	2827	445	3	3	3
## OCEAN SUNFISH	0	1	0	0	1	0
## PACIFIC CHUB MACKEREL	4	0	1	0	2	1
## PACIFIC HERRING	2	2	608	1764	4	4
## PACIFIC POMPAÑO	0	0	0	12	1	1
## PACIFIC SANDDAB	0	0	0	4	1	1
## PACIFIC SANDFISH	0	0	0	1	1	0
## PACIFIC SARDINE	111	8	0	1	3	2
## PACIFIC SPINY DOGFISH	1	2	2	3	4	3
## PACIFIC STAGHORN SCULPIN	0	0	1	0	1	0
## PACIFIC TOMCOD	0	0	37	0	1	1
## PINK SALMON	0	2	0	0	1	1
## REX SOLE	0	0	0	1	1	0
## SALPS	0	0	0	0	0	0
## SEA NETTLE	0	1570	1111	47	3	3
## SHINER PERCH	0	0	0	1	1	0
## SMELT SPP.	0	0	2582	223	2	2

## SOCKEYE SALMON	22	0	0	0	1	1
## STARRY FLOUNDER	0	0	3	2	2	2
## STEELHEAD	3	2	0	5	3	3
## THRESHER SHARK	0	0	0	1	1	0
## TOPE	0	7	0	0	1	1
## WATER JELLY	0	377	507	124407	3	3
## WESTERN RIVER LAMPREY	0	2	2	20	3	3
## WOLF-EEL	0	6	1	30	3	2
## YELLOWTAIL ROCKFISH	1	0	0	0	1	0

Examining these tables, for analysis we select the species that have total catch of at least 100 and occurred more than once in at least 3 cruises.

```
cat('\n*** Species Selected for Analysis ***\n')

##
## *** Species Selected for Analysis ***

sel.spec <- rownames(tab1)[tab1[, "Total"] >= 100]
sel.spec <- sel.spec[sel.spec %in% rownames(tab3[tab3[, "Ngt1"]>=3, ])]
print(sel.spec)

## [1] "CHINOOK subadult"      "CHINOOK subyearling" "CHINOOK yearling"
## [4] "CHUM SALMON"          "COHO subadult"       "COHO yearling"
## [7] "MARKET SQUID"         "NORTHERN ANCHOVY"    "PACIFIC HERRING"
## [10] "SEA NETTLE"           "WATER JELLY"
```

## Part2: Summary plots - CPUE by time

To estimate CPUE for each species in each haul, we need to convert the raw data that has counts by size into total counts for each species in each haul, then we need to summarize effort.

Generating the total catch summary is just a tabulation of numbers by size:

```
# Total number by Haul (rows) and Species
MMEDcnt <- with(MMEDdata, tapply(Number, list(Haul, SpecAge), FUN=sum, simplify=T))
MMEDcnt[is.na(MMEDcnt)] <- 0
MMEDcnt <- as.data.frame(MMEDcnt)
## print(summary(MMEDcnt)) ### DEBUG ###
```

Then, build a parallel structure of the haul data. This includes location, date, time, effort (as distance towed), and gear information.

```
MMEDhauls <- data.frame(HaulID = rownames(MMEDcnt))
MMEDhauls$Station <- with(MMEDdata, tapply(as.character(Station), list(Haul), FUN=unique))
MMEDhauls$Date <- with(MMEDdata, tapply(as.character(Date), list(Haul), FUN=unique))
MMEDhauls$Effort <- with(MMEDdata, tapply(Distance, list(Haul), FUN=unique))
MMEDhauls$Gear <- factor(with(MMEDdata, tapply(as.character(MMED), list(Haul), FUN=unique)),
                          levels=c('None', 'Up', 'Down'))
MMEDhauls$Block <- with(MMEDdata, tapply(as.character(Block), list(Haul), FUN=unique))
```

Finally, generate CPUE plots for selected species.

For this, we create an artificial “time” scale for plotting, creating equal intervals within blocks. Then, CPUE is computed as counts for each species divided by effort.

```
MMEDhauls$PlotTime <- match(MMEDhauls$Block, LETTERS[1:13]) - 1 +
  c((1:10)/11, (1:10)/11, (1:4)/5, (1:4)/5, (1:10)/11, (1:4)/5, (1:4)/5, (1:8)/9,
```



```

      (1:8)/9, (1:8)/9, (1:4)/5, (1:4)/5, (1:8)/9)
## print(summary(MMEDhauls)) ### DEBUG ###
MMEDcpue=sweep(MMEDcnt, 1, MMEDhauls$Effort, '/')
## print(summary(MMEDcpue)) ### DEBUG ###

```

Before running the analysis, set up some plotting configurations:

```

bw <- FALSE          #Flag for black-and-white figures
BLACK <- 'black'
BLUE <- if(bw) 'black' else 'blue'  #color code for blue
RED <- if(bw) 'black' else 'red'    #color code for red
lndscp <- FALSE      #flag for landscape figures
plcol <- if(lndscp) 3 else 2       #number of columns for multi-plots

```

A special plotting function is created, so it can be re-used later.

```

cpue.plot <- function(t, y, dot.col=1, log.zero=FALSE, ...) {
  minpos <- 0
  if (log.zero) {
    minpos <- min(y[y>0]) #minimum positive value
    if(any(y<=0)) {
      y[y<=0] <- minpos/2 # recode
    } # if(any...)
  } # if(log.zero)
  plot(t, y, log=ifelse(log.zero,"y",""), type='p',
       axes=F, ...)
  # label blocks at midpoint:
  blocks <- seq(round(min(t)), round(max(t)))
  axis(side=1, at=blocks, labels=NA) # Ticks at day boundaries
  abline(v=blocks, col='blue')
  axis(side=1, at=blocks[-1]-0.5,
       labels=sort(unique(as.character(MMEDhauls$Block))),
       tick=FALSE, cex.axis=0.8)
  if (log.zero && any(y<minpos)) {
    tck<- axisTicks(range(y),log=FALSE) #default tick locations
    axis(side=2, at=c(minpos/2,tck), labels=c(0,tck))
  } else {
    axis(side=2)
  } # if (log.zero && ...)
  box()
} # cpue.plot()

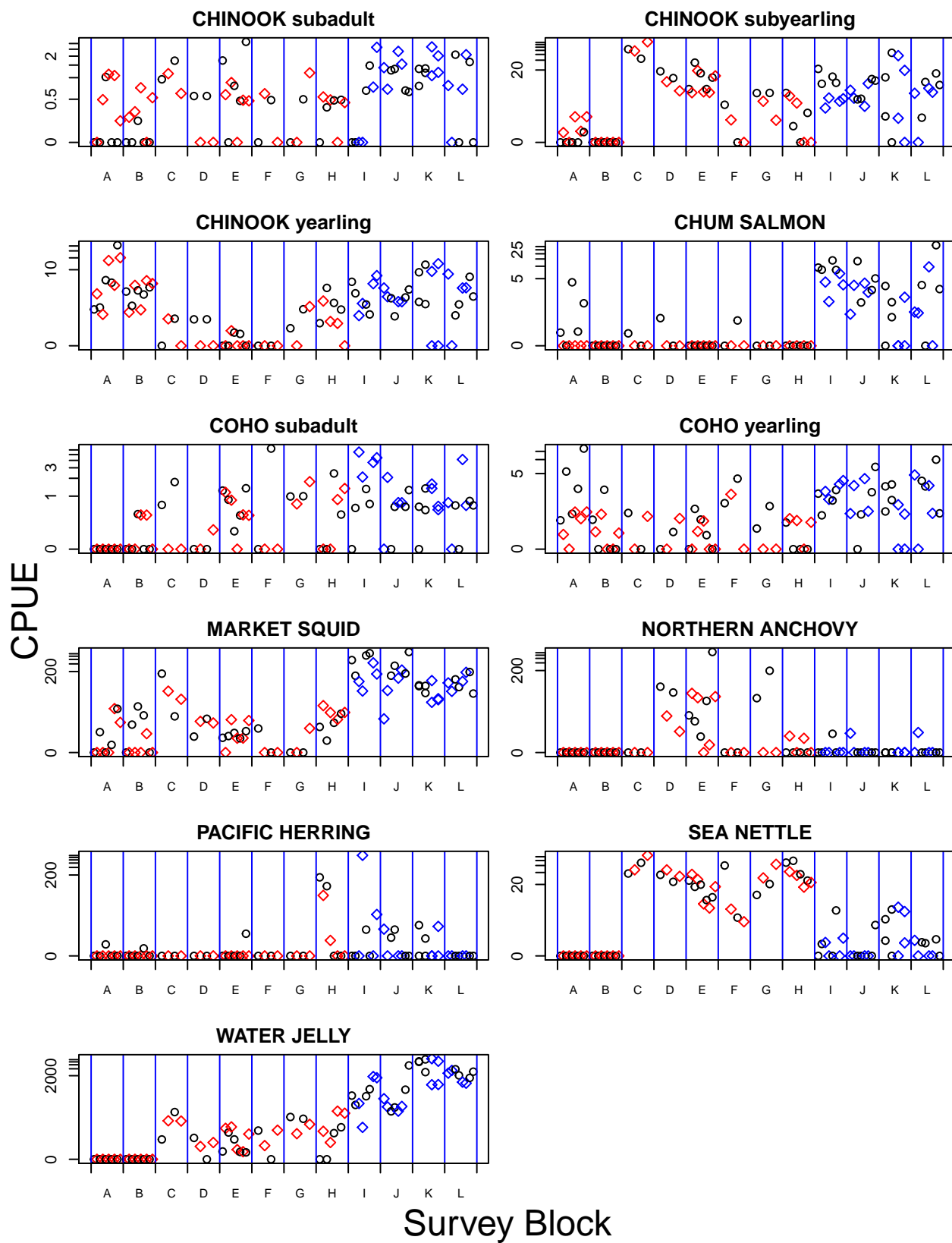
```

Generate the individual plots:

```

par(mfrow=c(ceiling(length(sel.spec)/plcol), plcol), omi=c(0.5,0.5,0,0), mar=c(3,2,2,1))
for (sp in sel.spec) {
  cpue.plot(MMEDhauls$PlotTime, MMEDcpue[[sp]], log.zero=TRUE,
           col=c(BLACK,RED,BLUE)[as.numeric(MMEDhauls$Gear)],
           pch=1+4*(MMEDhauls$Gear!='None'), xlab="", ylab="", main=sp)
} # for (sp)
mtext('Survey Block', side=1, outer=TRUE, cex=1.5)
mtext('CPUE', side=2, outer=TRUE, cex=1.5)

```



## Part 3: Overall Catch Ratio Statistics

In preliminary analyses, we considered a number of methods, including:

- *CPUE ratio estimate*. Ratio of mean CPUEs (Wilderbuer et al. 1998, North. Am. J. Fish. Manage. 18:11-18) using bootstrapped quantiles rather than a normal approximation.
- *Paired sample differences*. This method uses paired t-test statistics with a  $\log(x+\text{offset})$  transform on CPUE.
- *Nonparametric paired differences*. This method computes the median catch ratio and approximate binomial quantiles based on eq. 10.3 in Efron (1982, “*The Jackknife, the Bootstrap, and Other Resampling Plans*”) which finds the nearest observation toward the tail from the nominal quantile.
- *GLM analysis of deviance*. This method computes a GLM blocked ANODEV model of `Catch ~ Block + Gear + offset()` with either a Poisson or a negative binomial distribution, where the offset is the effort (km towed) for each sample. (Because the model uses log links, including effort as an offset is similar to using a linear model  $\log(\text{CPUE}) \sim \text{Block} + \text{Gear}$ .)

After reviewing the methods, for the final analysis, we used only the GLM method.

### 3.1 The analysis

To make the coding easier, we define standard names for the gear types, and set the rounding parameters for easy-to-read results tables.

```
assign('gears', c('None', 'Up', 'Down'), 1)
options(scipen=3) # Favor non-scientific notation
```

Then, create a function to compute the GLM estimates. The *R* `glm()` method does not support negative binomial distributions, so we use the MASS library (Venables & Ripley 2002, “*Modern Applied Statistics with S*, 4th ed.”). The `glm.nb()` function is used to estimate the  $\theta$  parameter; if that estimation fails, it reverts to the value specified in “init.theta”. Then (because `glm.nb()` fails for some species) `glm()` is used with the estimated  $\theta$  to get the final statistics. This means that the distribution of the final catch ratio estimate does not include error in the estimation of  $\theta$ .

The GLM model is on a log scale, so exponential transformations are needed to get the catch-ratio estimate and quantiles. Predictions are returned in `$Pred` for later plotting.

**NOTE** that we want the ratio MED/STD, so we use the negative of the log-scale `GearStd` coefficient.

```
GLMAnoDevEst <- function(dat, qprobs=0.5, nb=TRUE, init.theta=1,
                          diag.plt=FALSE, plt.lab='') {
  rslt <- matrix(NA, ncol=length(gears)-1, nrow=1+length(qprobs),
                dimnames=list(c("Mean", as.character(qprobs)), gears[2:length(gears)]))
  library(MASS)
  fit.data <- data.frame(Count=as.vector(dat[,1]),
                        Date=as.factor(dat$Date),
                        Station=as.factor(dat$Station),
                        Block=as.factor(dat$Block),
                        Gear=as.factor(dat$Gear),
                        Offset=as.vector(dat$Effort))
  if (nb) { # Negative Binomial fit
    cat('\n\tInitial fit to estimate theta\n')
    fit.init <- try(glm.nb(Count ~ Block + Gear + offset(log(Offset)),
                          data=fit.data, init.theta=init.theta,
                          control=list(epsilon=1e-03, maxit=500, trace=0)))
    if (inherits(fit.init, 'try-error')) {
```

```

print(fit.init)
warning("glm.nb failed to estimate theta; using default value")
theta.init=init.theta # initial estimate for herring from theta.ml
} else {
cat('Estimated theta: ', fit.init$theta, ', SE: ', fit.init$SE.theta, '\n')
if(is.finite(fit.init$SE.theta)) {
theta.init <- fit.init$theta
} else {
warning("glm.nb failed to estimate theta; using default value")
theta.init <- init.theta
}
} # if 'try-error'
cat('\n\tFinal fit with theta = ', theta.init, '\n')
fit.fin <- glm(Count ~ Block + Gear + offset(log(Offset)),
data=fit.data, family=negative.binomial(theta.init),
control=list(epsilon=1e-08, maxit=500, trace=FALSE))
} else { # Poisson fit
fit.fin <- glm(Count ~ Block + Gear + offset(log(Offset)),
data=fit.data, family=poisson,
control=list(epsilon=1e-08, maxit=500, trace=FALSE))
} # if (nb)
cat("\nFIT STATISTICS:")
# print(summary(fit.fin))
fit.anova <- anova(fit.fin, test="Chisq")
print(fit.anova)
geareffects <- paste('Gear',gears[2:length(gears)], sep='')
.lmn <- summary(fit.fin)$coefficients[geareffects, "Estimate"]
.lsd <- summary(fit.fin)$coefficients[geareffects, "Std. Error"]
.df <- fit.fin$df.residual
for (g in 1:length(.lmn)) {
if (.lsd[g] > 1000) { # Estimate blew up, just use the mean value
.mn <- exp(.lmn[g])
.qnt <- rep(NA, length(qprobs))
} else {
.mn <- exp(.lmn[g] + .lsd[g]^2 / 2)
.qnt <- exp(qt(qprobs, .df)*.lsd[g]+.lmn[g])
} # if (.lsd[g])
rslt[, g] <- c(.mn, .qnt)
} # for (g)
predCatch <- predict(fit.fin, type="response", se.fit=TRUE)
if (diag.plt) {
rs <- resid(fit.fin, type="deviance")
op <- par(omi=c(0,0,0.25,0), mfrow=c(1,2), mar=c(4,4,1,1))
plot(predCatch$fit, rs, xlab="Prediction", ylab="Deviance Resids")
qqnorm(rs, ylab="Deviance Resids")
qqline(rs)
mtext(paste(plt.lab, ifelse(nb, "Neg. Binomial", "Poisson"), sep=' - '),
side=3, outer=TRUE)
par(op)
} # if (diat.plt)
return(list(Smry=rslt, Pred=predCatch))
} # GLMAnoDevEst()

```

Before running the analysis, create two summary lists: one for the means and quantiles of all methods by

species, and one for storing GLM model results for plotting. The first list contains one element for each species, holding a matrix of summary results for estimates of the ratio  $A$  by each of the various estimation methods. Rows are the methods, columns are a six-number summary: Mean, Median, and quantiles (0.05, 0.25, 0.75, 0.95).

The routine then loops through the species, computing first the GLM Poisson catch ratio estimates, then the GLM negative-binomial estimates. The Poisson results are used to get a crude initial estimate of the neg-bin  $\Theta$  parameter, which is used to initiate the refined estimate via `glm.nb()`. If the refined estimate fails, the initial estimate of  $\Theta$  is used instead.

```
stat.sum <- list() # structure for storing summary results
GLM.pred <- list() # structure for storing GLM predictions
for (sp in sel.spec) {
  cat('\n***** ', sp, ' *****\n')
  .sumtbl <- array(NA, dim=c(2, 6, length(gears)-1),
    dimnames=list(c('GLM.Po', 'GLM.nb'),
      c('Mean', 'q0.05', 'q0.25', 'Median', 'q0.75', 'q0.95'),
      gears[2:length(gears)]))
  .qprobs <- c(0.05, 0.25, 0.50, 0.75, 0.95)
  cat("\n*** METHOD 1: GLM ANODEV, Poisson ***\n")
  .est <- GLMAnoDevEst(cbind(MMEDcnt[[sp]], MMEDhauls), .qprobs,
    nb=FALSE, diag.plt=TRUE, plt.lab=sp)
  print(.est$Smry)
  .sumtbl['GLM.Po', , ] <- .est$Smry

  cat("\n*** METHOD 2: GLM ANODEV, negative binomial ***\n")
  init.theta <- theta.ml(y=MMEDcnt[[sp]], mu=.est$Pred$fit,
    n=length(MMEDcnt[[sp]]), limit=100, trace=FALSE)
  cat('Initial Theta: ', init.theta, '\n')
  .est <- GLMAnoDevEst(cbind(MMEDcnt[[sp]], MMEDhauls), .qprobs,
    nb=TRUE, init.theta=init.theta, diag.plt=TRUE, plt.lab=sp)
  print(.est$Smry)
  .sumtbl['GLM.nb', , ] <- .est$Smry

  # Add species to summary lists
  GLM.pred[[sp]] <- .est$Pred
  .sumtbl <- round(.sumtbl, 3) # round statistical results
  stat.sum[[sp]] <- .sumtbl
} # for(sp)
```

```
##
## ***** CHINOOK subadult *****
##
## *** METHOD 1: GLM ANODEV, Poisson ***
##
## FIT STATISTICS:Analysis of Deviance Table
##
## Model: poisson, link: log
##
## Response: Count
##
## Terms added sequentially (first to last)
##
##
##      Df Deviance Resid. Df Resid. Dev    Pr(>Chi)
```

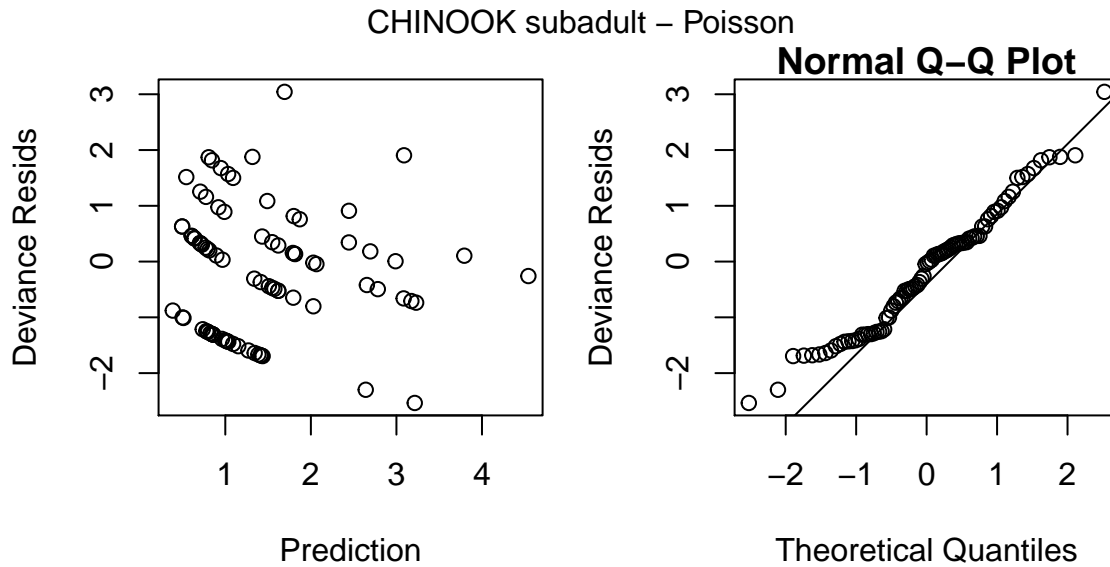
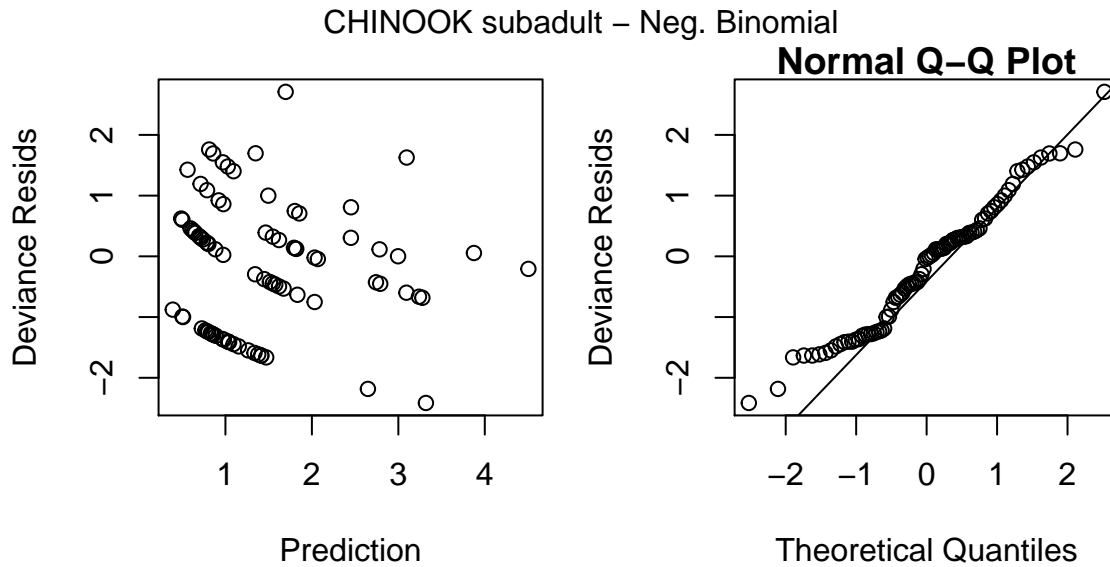


Figure 1:

```
## NULL                85      157.50
## Block 11    47.887      74      109.61 0.000001496 ***
## Gear   2      1.179      72      108.43      0.5545
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##           Up      Down
## Mean 1.0303668 1.4245452
## 0.05 0.6805217 0.8441632
## 0.25 0.8564966 1.1232247
## 0.5  1.0028455 1.3662921
## 0.75 1.1742010 1.6619597
## 0.95 1.4778356 2.2113664
##
## *** METHOD 2: GLM ANODEV, negative binomial ***
## Initial Theta: 11.84953
##
## Initial fit to estimate theta
## Estimated theta: 11.4745 , SE: 15.34951
##
## Final fit with theta = 11.4745
##
## FIT STATISTICS:Analysis of Deviance Table
##
## Model: Negative Binomial(11.4745), link: log
##
## Response: Count
##
## Terms added sequentially (first to last)
##
##
##           Df Deviance Resid. Df Resid. Dev Pr(>Chi)
## NULL                85      141.391
```



```
## Block 11    42.336         74    99.055 0.0003957 ***
## Gear   2     1.051         72    98.004 0.6575545
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##           Up      Down
## Mean 1.0744294 1.4517892
## 0.05 0.6471796 0.7688163
## 0.25 0.8540238 1.0819175
## 0.5  1.0329542 1.3676013
## 0.75 1.2493731 1.7287207
## 0.95 1.6486834 2.4327441
##
## ***** CHINOOK subyearling *****
##
## *** METHOD 1: GLM ANODEV, Poisson ***
##
## FIT STATISTICS:Analysis of Deviance Table
##
## Model: poisson, link: log
##
## Response: Count
##
## Terms added sequentially (first to last)
##
##
##      Df Deviance Resid. Df Resid. Dev   Pr(>Chi)
## NULL              85      5379.7
## Block 11    4427.6      74      952.1 < 2.2e-16 ***
## Gear   2      22.5      72      929.5 0.00001275 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##           Up      Down
```

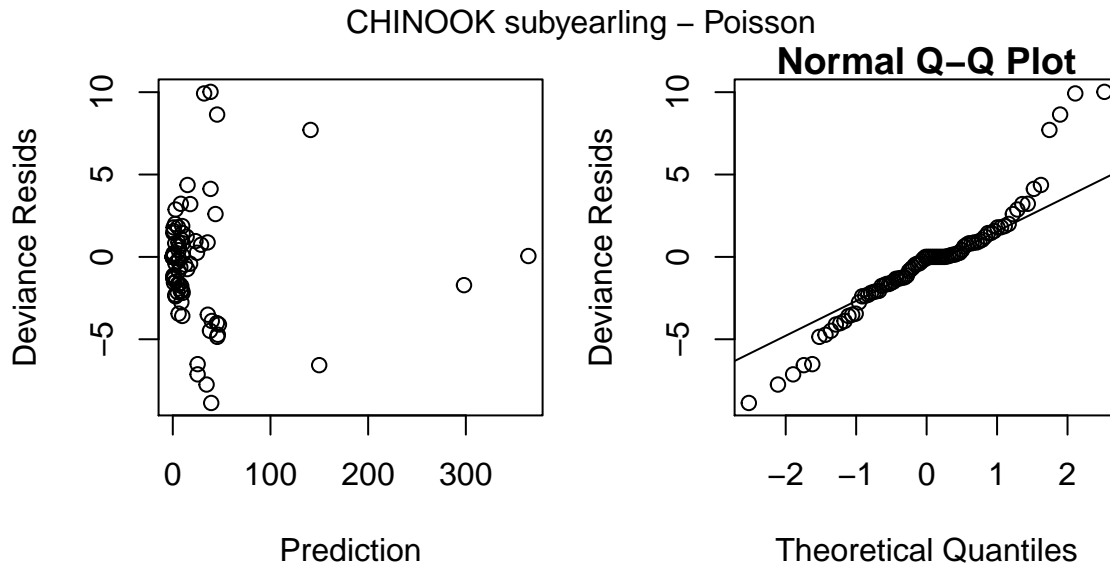
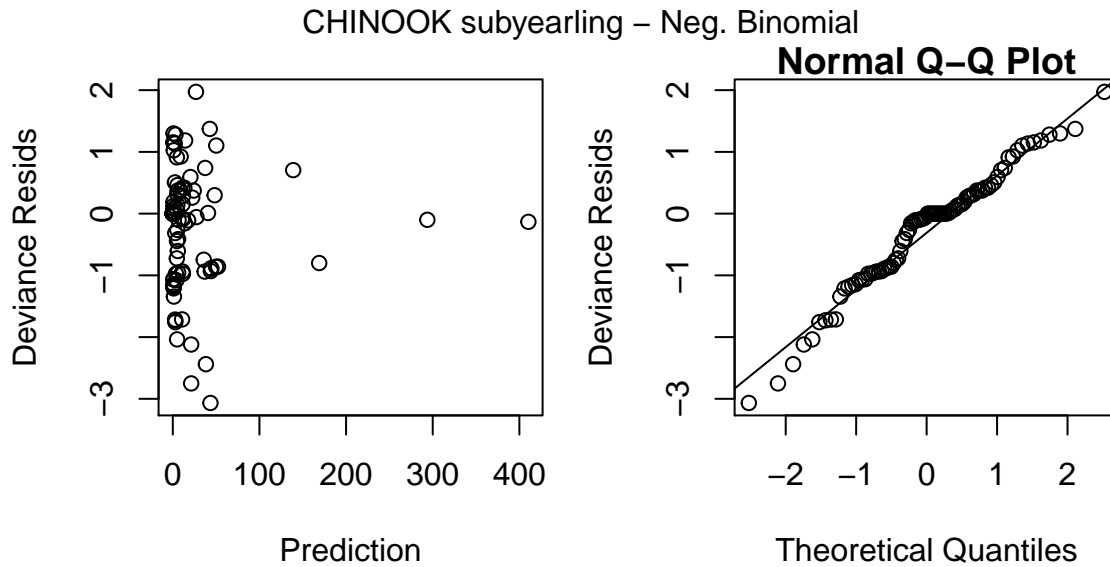


Figure 3:

```
## Mean 0.9848031 0.6494879
## 0.05 0.9025556 0.5537613
## 0.25 0.9497254 0.6071303
## 0.5 0.9834959 0.6466797
## 0.75 1.0184673 0.6888055
## 0.95 1.0716949 0.7551895
##
## *** METHOD 2: GLM ANODEV, negative binomial ***
## Initial Theta: 1.310679
##
## Initial fit to estimate theta
## Estimated theta: 1.340598 , SE: 0.2734801
##
## Final fit with theta = 1.340598
##
## FIT STATISTICS:Analysis of Deviance Table
##
## Model: Negative Binomial(1.3406), link: log
##
## Response: Count
##
## Terms added sequentially (first to last)
##
##      Df Deviance Resid. Df Resid. Dev Pr(>Chi)
## NULL                85    357.53
## Block 11  267.174      74    90.36 < 2e-16 ***
## Gear    2    4.698      72    85.66 0.06016 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
##           Up      Down
## Mean 0.8942809 0.5101777
```





```
## 0.05 0.5381177 0.2964244
## 0.25 0.7104730 0.3983538
## 0.5 0.8596332 0.4878714
## 0.75 1.0401087 0.5975053
## 0.95 1.3732482 0.8029653
##
## ***** CHINOOK yearling *****
##
## *** METHOD 1: GLM ANODEV, Poisson ***
##
## FIT STATISTICS:Analysis of Deviance Table
##
## Model: poisson, link: log
##
## Response: Count
##
## Terms added sequentially (first to last)
##
##      Df Deviance Resid. Df Resid. Dev Pr(>Chi)
## NULL                85    1456.63
## Block 11      766.09      74     690.54 < 2e-16 ***
## Gear   2        6.53      72     684.01 0.03822 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##           Up      Down
## Mean 0.8063947 1.1275647
## 0.05 0.6918107 0.8631587
## 0.25 0.7558489 1.0043845
## 0.5 0.8031664 1.1143906
## 0.75 0.8534461 1.2364452
## 0.95 0.9324461 1.4387464
```

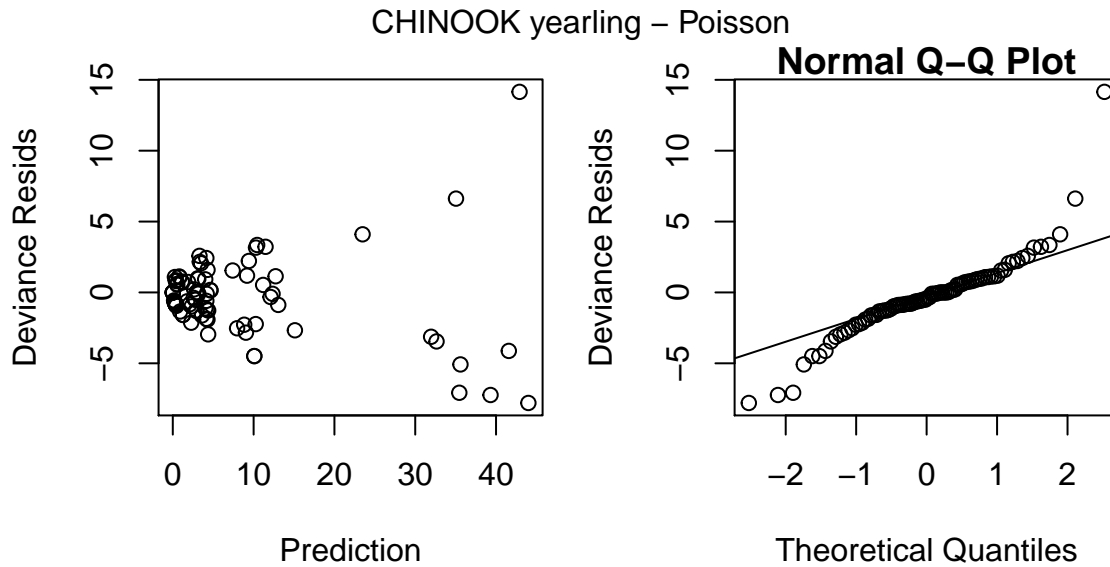


Figure 5:

```
##
## *** METHOD 2: GLM ANODEV, negative binomial ***
## Initial Theta: 1.5042
##
## Initial fit to estimate theta
## Estimated theta: 1.503116 , SE: 0.3096866
##
## Final fit with theta = 1.503116
##
## FIT STATISTICS:Analysis of Deviance Table
##
## Model: Negative Binomial(1.5031), link: log
##
## Response: Count
##
## Terms added sequentially (first to last)
##
##
```

	Df	Deviance	Resid.	Df	Resid.	Dev	Pr(>Chi)
NULL				85		219.731	
Block 11	138.594			74		81.138	<2e-16 ***
Gear 2	1.083			72		80.055	0.5341

```
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
##           Up      Down
## Mean 0.7867095 1.2389568
## 0.05 0.4741009 0.7050806
## 0.25 0.6254745 0.9574429
## 0.5  0.7563934 1.1809969
## 0.75 0.9147153 1.4567488
## 0.95 1.2067706 1.9781479
##
```

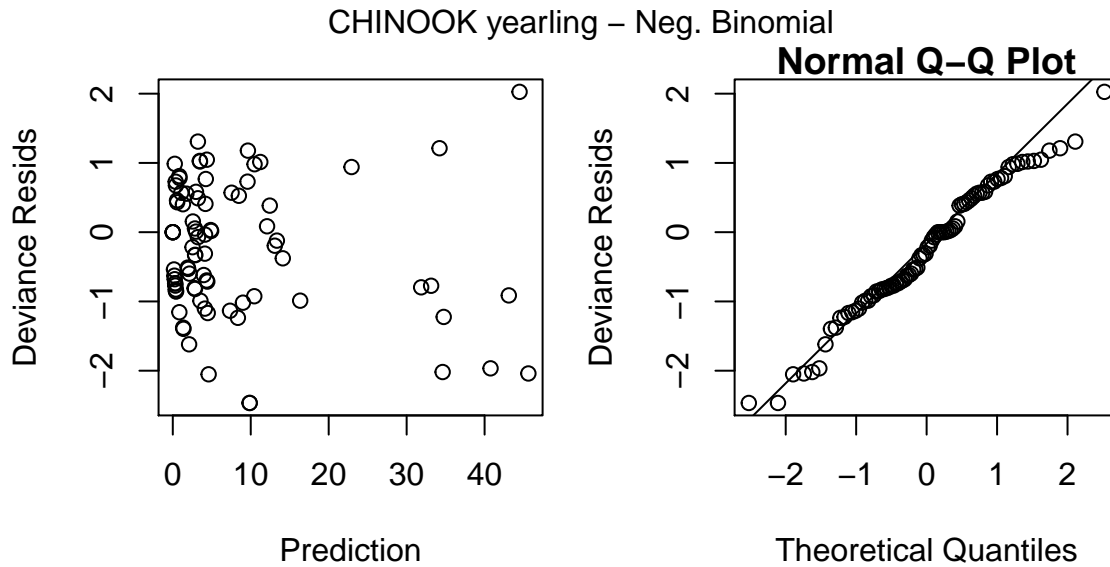
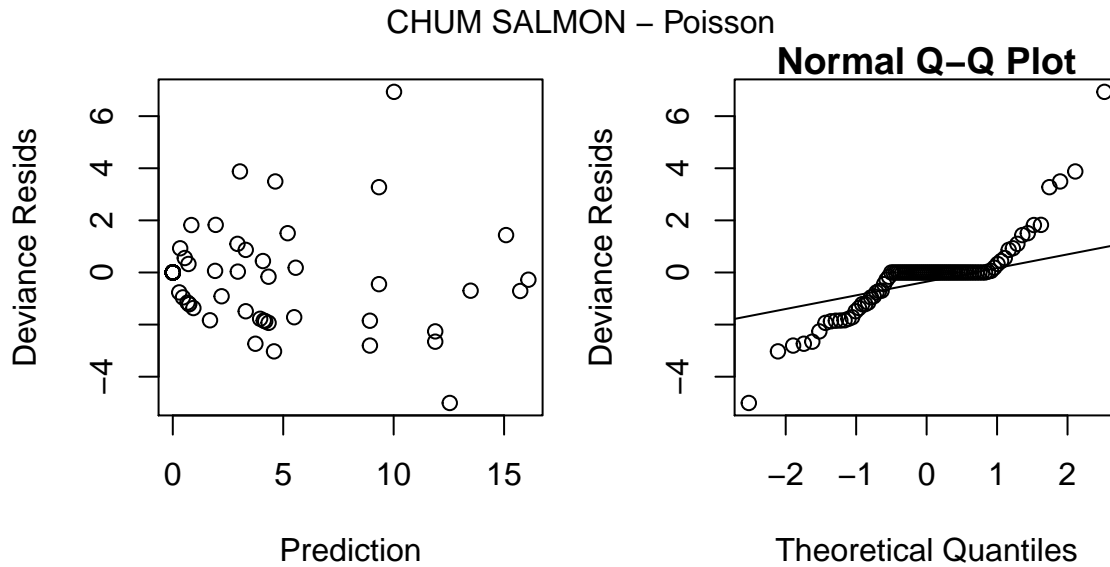


Figure 6:

```
## ***** CHUM SALMON *****
##
## *** METHOD 1: GLM ANODEV, Poisson ***
## Warning: glm.fit: fitted rates numerically 0 occurred
##
## FIT STATISTICS:Analysis of Deviance Table
##
## Model: poisson, link: log
##
## Response: Count
##
## Terms added sequentially (first to last)
##
##      Df Deviance Resid. Df Resid. Dev  Pr(>Chi)
## NULL                        85      882.42
## Block 11      613.58         74      268.84 < 2.2e-16 ***
## Gear   2       71.07         72      197.77 3.696e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##              Up      Down
## Mean 8.724456e-09 0.3905847
## 0.05          NA 0.2961159
## 0.25          NA 0.3463782
## 0.5           NA 0.3857016
## 0.75          NA 0.4294894
## 0.95          NA 0.5023904
##
## *** METHOD 2: GLM ANODEV, negative binomial ***
## Initial Theta: 1.259582
##
```



```
## Initial fit to estimate theta
## Estimated theta: 1.259586 , SE: 0.3875694
##
## Final fit with theta = 1.259586
##
## FIT STATISTICS:Analysis of Deviance Table
##
## Model: Negative Binomial(1.2596), link: log
##
## Response: Count
##
## Terms added sequentially (first to last)
##
##      Df Deviance Resid. Df Resid. Dev Pr(>Chi)
## NULL                        85      292.744
## Block 11  221.763          74       70.981 < 2.2e-16 ***
## Gear    2   25.384          72       45.597 1.216e-10 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##              Up      Down
## Mean 2.679148e-09 0.3863633
## 0.05          NA 0.2359931
## 0.25          NA 0.3092175
## 0.5           NA 0.3721880
## 0.75          NA 0.4479821
## 0.95          NA 0.5869828
##
## ***** COHO subadult *****
##
## *** METHOD 1: GLM ANODEV, Poisson ***
##
```

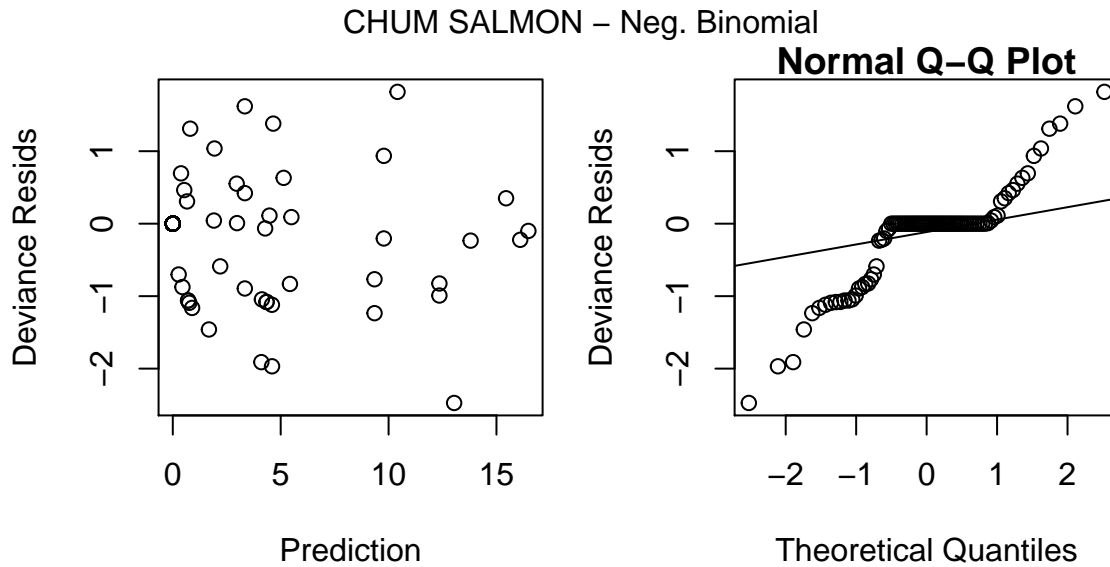


Figure 8:

```
## FIT STATISTICS:Analysis of Deviance Table
##
## Model: poisson, link: log
##
## Response: Count
##
## Terms added sequentially (first to last)
##
##      Df Deviance Resid. Df Resid. Dev   Pr(>Chi)
## NULL                85      251.80
## Block 11   118.686      74      133.11 < 2.2e-16 ***
## Gear    2    18.565      72      114.55 0.00009303 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##           Up      Down
## Mean 0.6194021 3.043432
## 0.05 0.3968357 1.758135
## 0.25 0.5073904 2.369548
## 0.5  0.6005476 2.907803
## 0.75 0.7108086 3.568325
## 0.95 0.9088332 4.809254
##
## *** METHOD 2: GLM ANODEV, negative binomial ***
## Initial Theta:  3.830574
##
## Initial fit to estimate theta
## Estimated theta:  3.686582 , SE:  2.051033
##
## Final fit with theta =  3.686582
##
## FIT STATISTICS:Analysis of Deviance Table
```

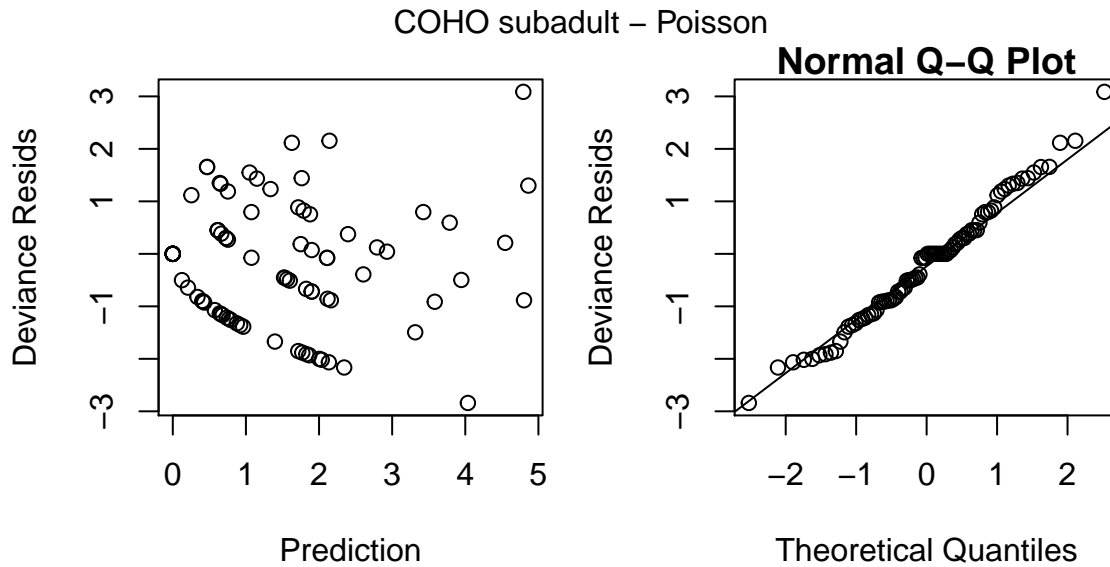


Figure 9:

```
##
## Model: Negative Binomial(3.6866), link: log
##
## Response: Count
##
## Terms added sequentially (first to last)
##
##      Df Deviance Resid. Df Resid. Dev  Pr(>Chi)
## NULL                85    180.195
## Block 11    87.323      74    92.872 5.602e-14 ***
## Gear    2    10.248      72    82.624 0.005957 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##      Up      Down
## Mean 0.6639356 2.872783
## 0.05 0.3817730 1.484663
## 0.25 0.5157354 2.114405
## 0.5  0.6338960 2.694712
## 0.75 0.7791285 3.434287
## 0.95 1.0525212 4.890992
##
## ***** COHO yearling *****
##
## *** METHOD 1: GLM ANODEV, Poisson ***
##
## FIT STATISTICS: Analysis of Deviance Table
##
## Model: poisson, link: log
##
## Response: Count
##
```

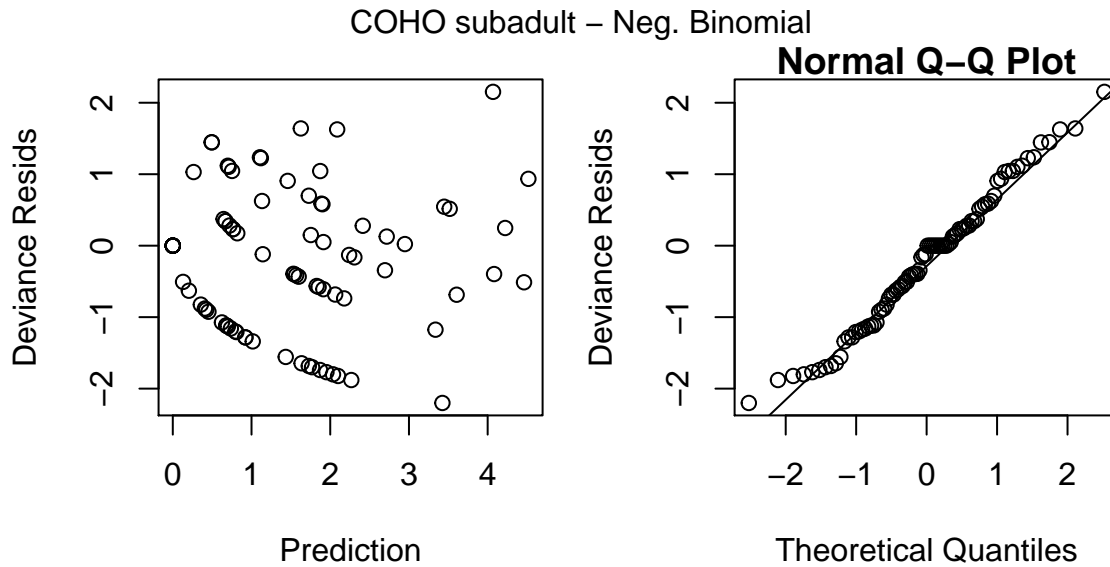


Figure 10:

```
## Terms added sequentially (first to last)
##
##
##      Df Deviance Resid. Df Resid. Dev  Pr(>Chi)
## NULL                85      588.09
## Block 11  220.599      74      367.49 < 2.2e-16 ***
## Gear    2   86.535      72      280.96 < 2.2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##           Up      Down
## Mean 0.1860394 0.7047559
## 0.05 0.1255099 0.4868875
## 0.25 0.1562238 0.5985192
## 0.5  0.1815321 0.6895519
## 0.75 0.2109403 0.7944304
## 0.95 0.2625601 0.9765742
##
## *** METHOD 2: GLM ANODEV, negative binomial ***
## Initial Theta:  1.554399
##
## Initial fit to estimate theta
## Estimated theta:  1.504939 , SE:  0.3966753
##
## Final fit with theta =  1.504939
##
## FIT STATISTICS:Analysis of Deviance Table
##
## Model: Negative Binomial(1.5049), link: log
##
## Response: Count
##
## Terms added sequentially (first to last)
```

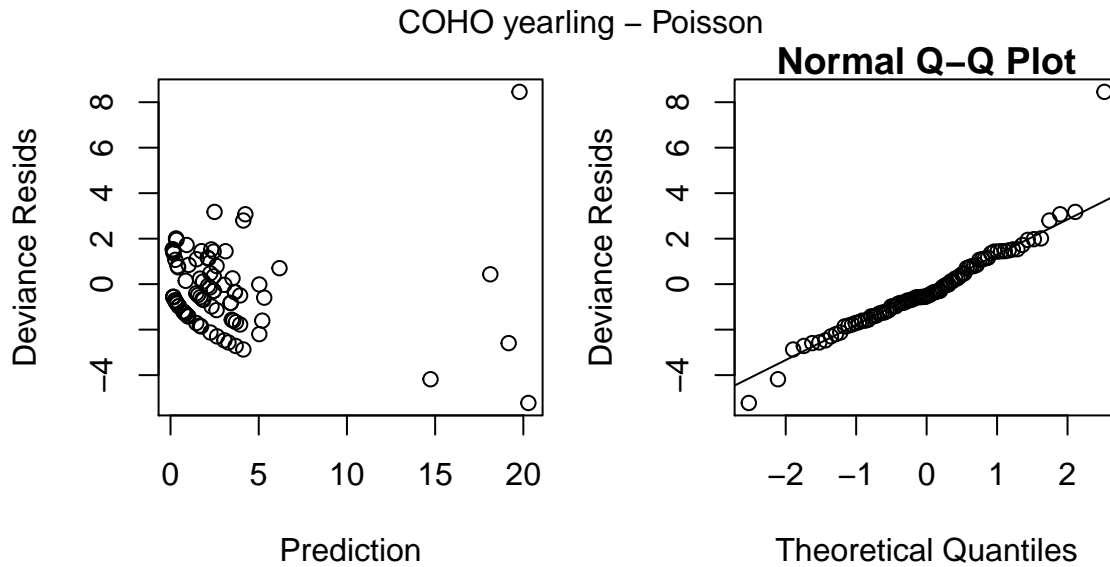


Figure 11:

```
##
##
##      Df Deviance Resid. Df Resid. Dev  Pr(>Chi)
## NULL                        85    171.099
## Block 11    72.115         74     98.983 3.228e-09 ***
## Gear   2    11.723         72     87.260 0.006241 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##           Up      Down
## Mean 0.3397104 0.7189217
## 0.05 0.1718611 0.3510543
## 0.25 0.2473126 0.5139029
## 0.5  0.3174407 0.6674197
## 0.75 0.4074544 0.8667963
## 0.95 0.5863376 1.2688893
##
## ***** MARKET SQUID *****
##
## *** METHOD 1: GLM ANODEV, Poisson ***
##
## FIT STATISTICS:Analysis of Deviance Table
##
## Model: poisson, link: log
##
## Response: Count
##
## Terms added sequentially (first to last)
##
##
##      Df Deviance Resid. Df Resid. Dev  Pr(>Chi)
## NULL                        85    41497
## Block 11    31995         74     9502 < 2.2e-16 ***
```



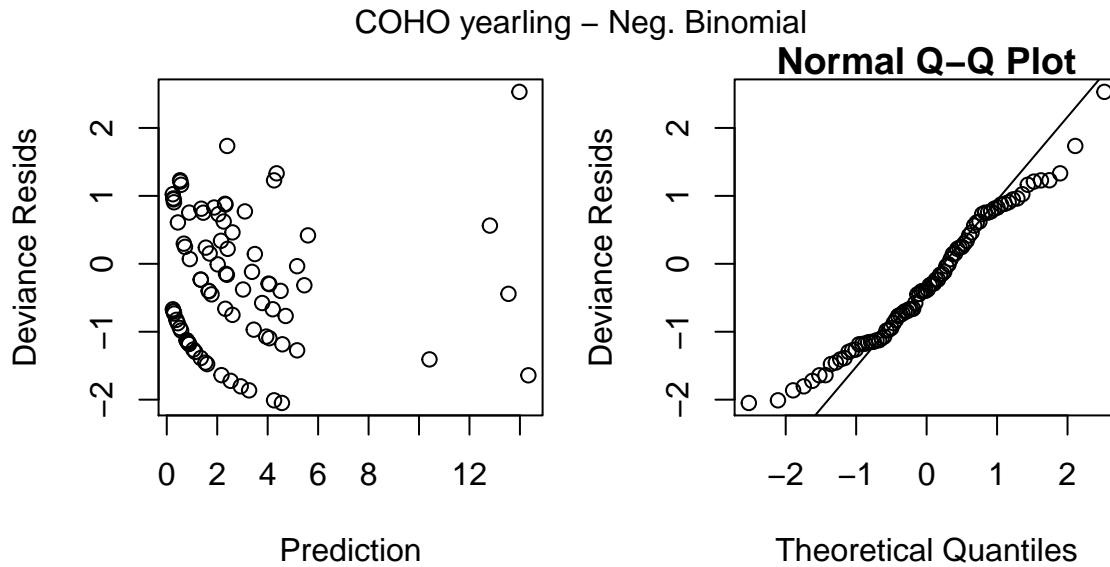


Figure 12:

```
## Gear    2      2833      72      6670 < 2.2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##           Up      Down
## Mean 0.3499610 0.3202018
## 0.05 0.3099832 0.3073303
## 0.25 0.3326078 0.3148456
## 0.5  0.3490732 0.3201062
## 0.75 0.3663537 0.3254546
## 0.95 0.3930925 0.3334131
##
## *** METHOD 2: GLM ANODEV, negative binomial ***
## Initial Theta: 0.7905173
##
## Initial fit to estimate theta
## Estimated theta: 0.9008459 , SE: 0.1607128
##
## Final fit with theta = 0.9008459
##
## FIT STATISTICS:Analysis of Deviance Table
##
## Model: Negative Binomial(0.9008), link: log
##
## Response: Count
##
## Terms added sequentially (first to last)
##
##
##      Df Deviance Resid. Df Resid. Dev Pr(>Chi)
## NULL      85      433.55
## Block 11    329.48      74    104.08 <2e-16 ***
## Gear    2      4.46      72     99.61  0.1631
```

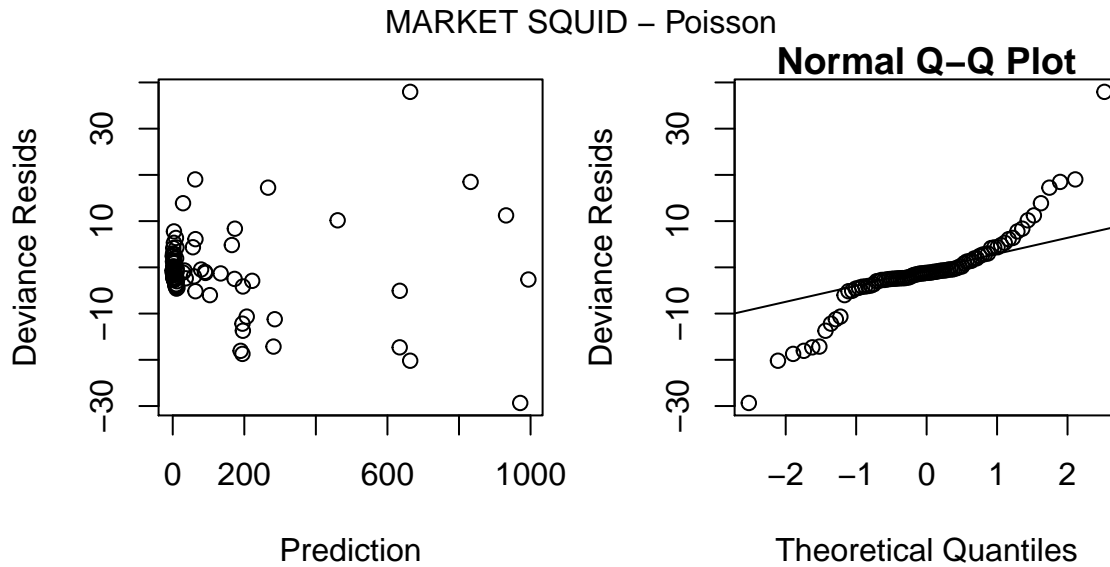
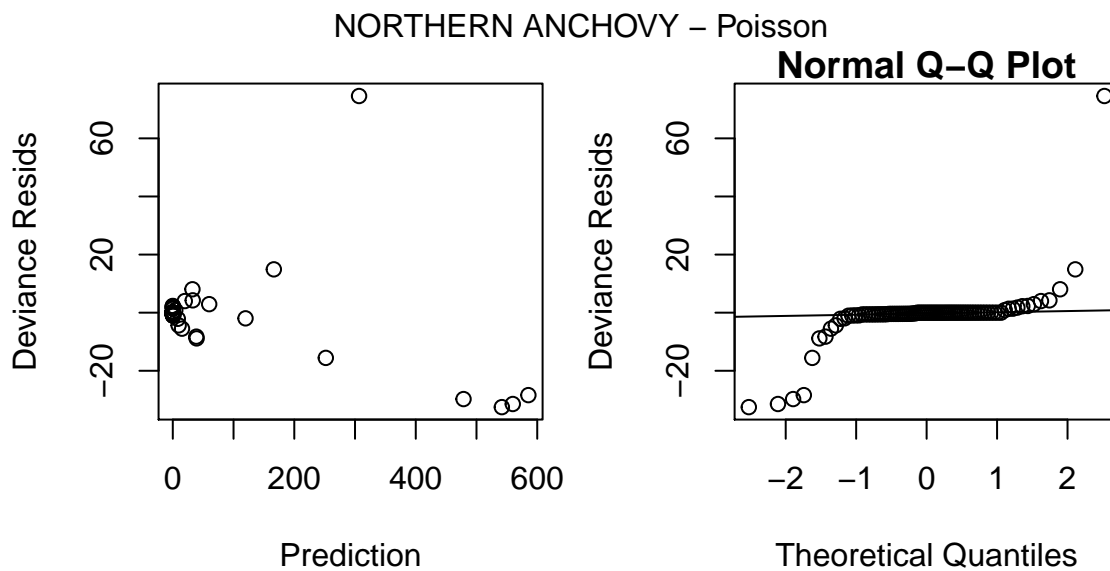
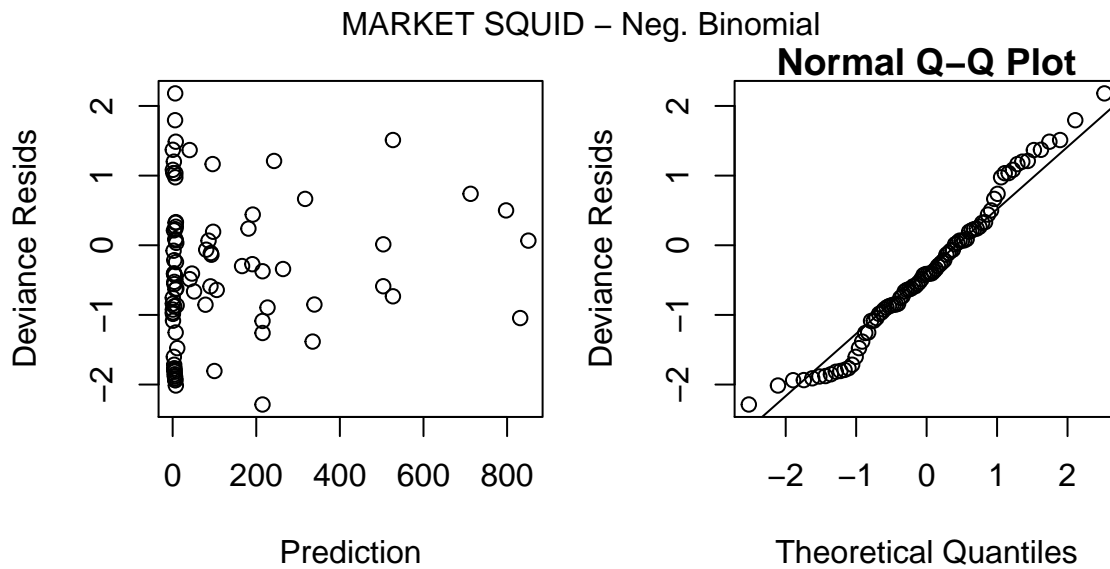


Figure 13:

```
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##           Up      Down
## Mean 0.8765406 0.4845325
## 0.05 0.4519912 0.2228432
## 0.25 0.6444103 0.3356991
## 0.5  0.8218839 0.4446347
## 0.75 1.0482347 0.5889203
## 0.95 1.4944830 0.8871711
##
## ***** NORTHERN ANCHOVY *****
##
## *** METHOD 1: GLM ANODEV, Poisson ***
##
## FIT STATISTICS:Analysis of Deviance Table
##
## Model: poisson, link: log
##
## Response: Count
##
## Terms added sequentially (first to last)
##
##      Df Deviance Resid. Df Resid. Dev  Pr(>Chi)
## NULL                85      22390
## Block 11      9361.5      74  13029 < 2.2e-16 ***
## Gear   2      2929.5      72   10100 < 2.2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##           Up      Down
## Mean 0.06762749 4.6141497
## 0.05 0.05972477 0.2828815
```



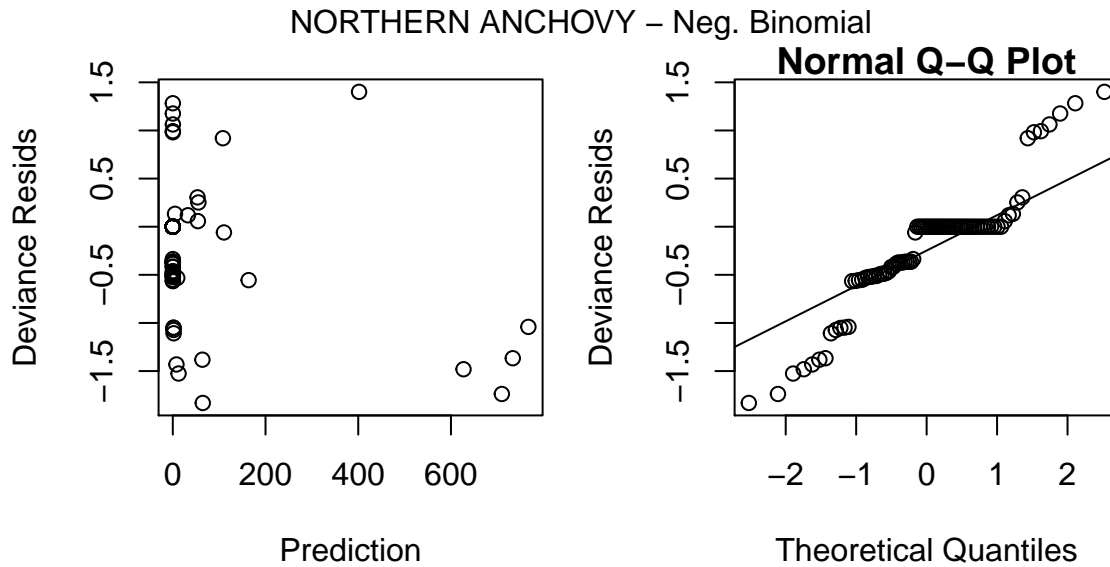


Figure 16:

```
## 0.25 0.06419196 0.9494745
## 0.5 0.06744764 2.1785820
## 0.75 0.07086843 4.9987855
## 0.95 0.07616913 16.7781196
##
## *** METHOD 2: GLM ANODEV, negative binomial ***
## Initial Theta: 0.3260892
##
## Initial fit to estimate theta
## Estimated theta: 0.314627 , SE: 0.09187918
##
## Final fit with theta = 0.314627
##
## FIT STATISTICS:Analysis of Deviance Table
##
## Model: Negative Binomial(0.3146), link: log
##
## Response: Count
##
## Terms added sequentially (first to last)
##
##      Df Deviance Resid. Df Resid. Dev Pr(>Chi)
## NULL                85    253.813
## Block 11    210.880      74    42.933 < 2.2e-16 ***
## Gear 2      6.113      72    36.821 0.002488 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##           Up      Down
## Mean 0.09891328 3.8901959
## 0.05 0.03385024 0.4174618
## 0.25 0.05841659 1.1492074
```

```

## 0.5 0.08493225 2.3016355
## 0.75 0.12348353 4.6097213
## 0.95 0.21310000 12.6898443
##
## ***** PACIFIC HERRING *****
##
## *** METHOD 1: GLM ANODEV, Poisson ***
##
## FIT STATISTICS:Analysis of Deviance Table
##
## Model: poisson, link: log
##
## Response: Count
##
## Terms added sequentially (first to last)
##
##
##      Df Deviance Resid. Df Resid. Dev Pr(>Chi)
## NULL                        85      19043.0
## Block 11  10470.5          74      8572.5 < 2.2e-16 ***
## Gear  2    2986.8          72      5585.7 < 2.2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##           Up      Down
## Mean 0.11263741 233.0833
## 0.05 0.08864923 126.3052
## 0.25 0.10160500 175.7400
## 0.5  0.11156983 220.4245
## 0.75 0.12251196 276.4707
## 0.95 0.14041665 384.6789
##
## *** METHOD 2: GLM ANODEV, negative binomial ***
## Initial Theta: 0.09446641
##
## Initial fit to estimate theta
## [1] "Error in while ((it <- it + 1) < limit && abs(del) > eps) { : \n missing value where TRUE/FALSE"
## attr(,"class")
## [1] "try-error"
## attr(,"condition")
## <simpleError in while ((it <- it + 1) < limit && abs(del) > eps) { t0 <- abs(t0) del <- score(
## Warning in GLMAnoDevEst(cbind(MMEDcnt[[sp]], MMEDhauls), .qprobs, nb =
## TRUE, : glm.nb failed to estimate theta; using default value
##
## Final fit with theta = 0.09446641
##
## FIT STATISTICS:Analysis of Deviance Table
##
## Model: Negative Binomial(0.0945), link: log
##
## Response: Count
##
## Terms added sequentially (first to last)

```

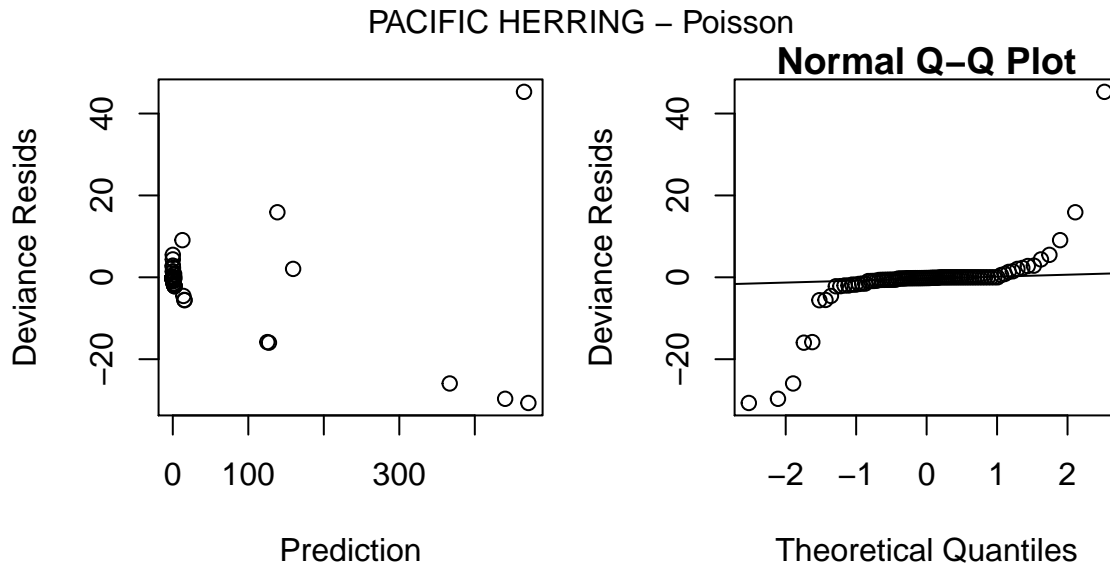


Figure 17:

```
##
##
##      Df Deviance Resid. Df Resid. Dev Pr(>Chi)
## NULL                        85      98.474
## Block 11      70.323        74      28.151 < 2.2e-16 ***
## Gear   2       3.531        72      24.620 0.0006172 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##              Up      Down
## Mean 0.07043164 2.9876844
## 0.05 0.01206889 0.7724318
## 0.25 0.02781045 1.5053250
## 0.5  0.04930240 2.3789116
## 0.75 0.08740338 3.7594672
## 0.95 0.20140445 7.3264980
##
## ***** SEA NETTLE *****
##
## *** METHOD 1: GLM ANODEV, Poisson ***
##
## FIT STATISTICS:Analysis of Deviance Table
##
## Model: poisson, link: log
##
## Response: Count
##
## Terms added sequentially (first to last)
##
##
##      Df Deviance Resid. Df Resid. Dev Pr(>Chi)
## NULL                        85      5737.1
## Block 11      4848.0        74      889.0 <2e-16 ***
```

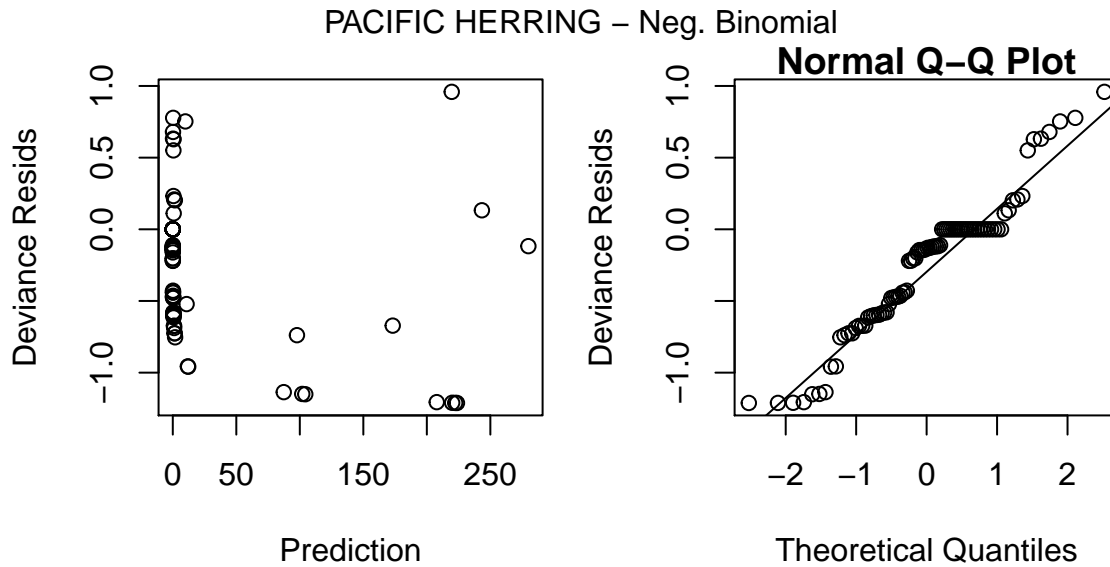


Figure 18:

```
## Gear    2      0.7      72      888.4    0.7119
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##           Up      Down
## Mean 0.9945518 0.8239794
## 0.05 0.9317755 0.4837972
## 0.25 0.9680875 0.6467325
## 0.5  0.9938081 0.7892015
## 0.75 1.0202120 0.9630551
## 0.95 1.0599704 1.2873971
##
## *** METHOD 2: GLM ANODEV, negative binomial ***
## Initial Theta:  2.154861
##
## Initial fit to estimate theta
## Estimated theta:  2.16212 , SE:  0.5290027
##
## Final fit with theta =  2.16212
##
## FIT STATISTICS:Analysis of Deviance Table
##
## Model: Negative Binomial(2.1621), link: log
##
## Response: Count
##
## Terms added sequentially (first to last)
##
##
##      Df Deviance Resid. Df Resid. Dev Pr(>Chi)
## NULL                85      632.74
## Block 11      559.76      74      72.98 <2e-16 ***
## Gear    2       1.65      72      71.33  0.4406
```

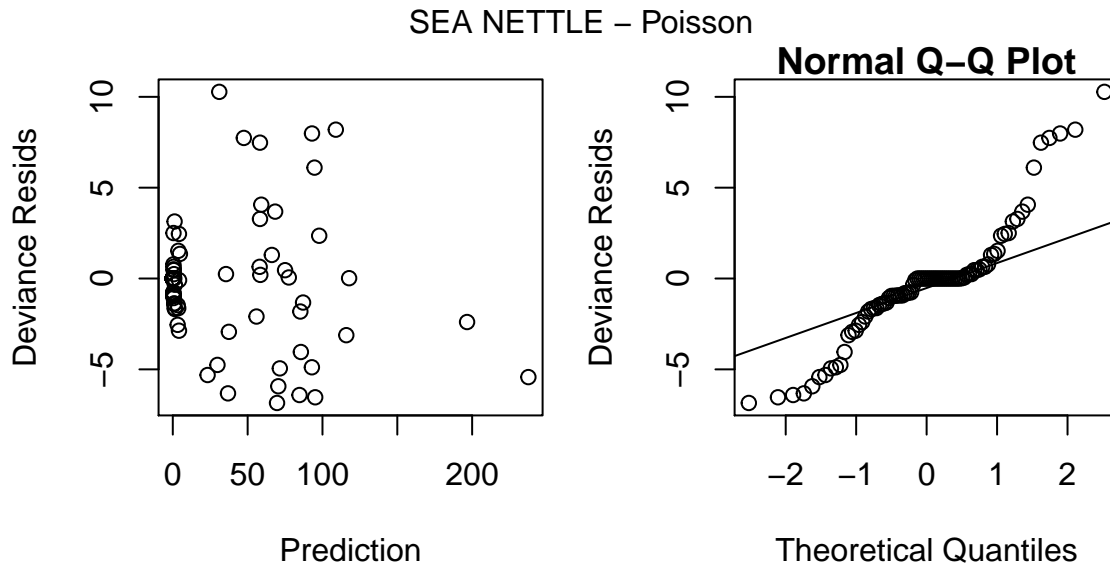


Figure 19:

```
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##           Up      Down
## Mean 0.9357340 0.6465416
## 0.05 0.6115211 0.2874541
## 0.25 0.7738936 0.4400243
## 0.5  0.9095491 0.5892529
## 0.75 1.0689838 0.7890904
## 0.95 1.3528226 1.2079109
##
## ***** WATER JELLY *****
##
## *** METHOD 1: GLM ANODEV, Poisson ***
##
## FIT STATISTICS:Analysis of Deviance Table
##
## Model: poisson, link: log
##
## Response: Count
##
## Terms added sequentially (first to last)
##
##      Df Deviance Resid. Df Resid. Dev Pr(>Chi)
## NULL                85      547720
## Block 11      468382      74      79338 < 2.2e-16 ***
## Gear   2       3313      72      76025 < 2.2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##           Up      Down
## Mean 1.634405 0.7221444
## 0.05 1.454432 0.7152668
```



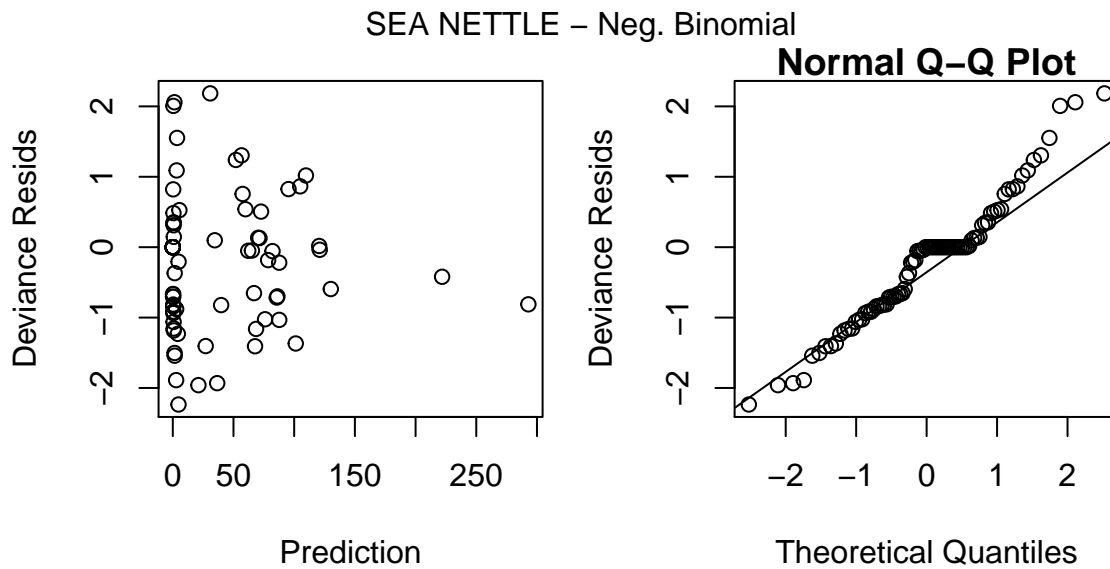


Figure 20:

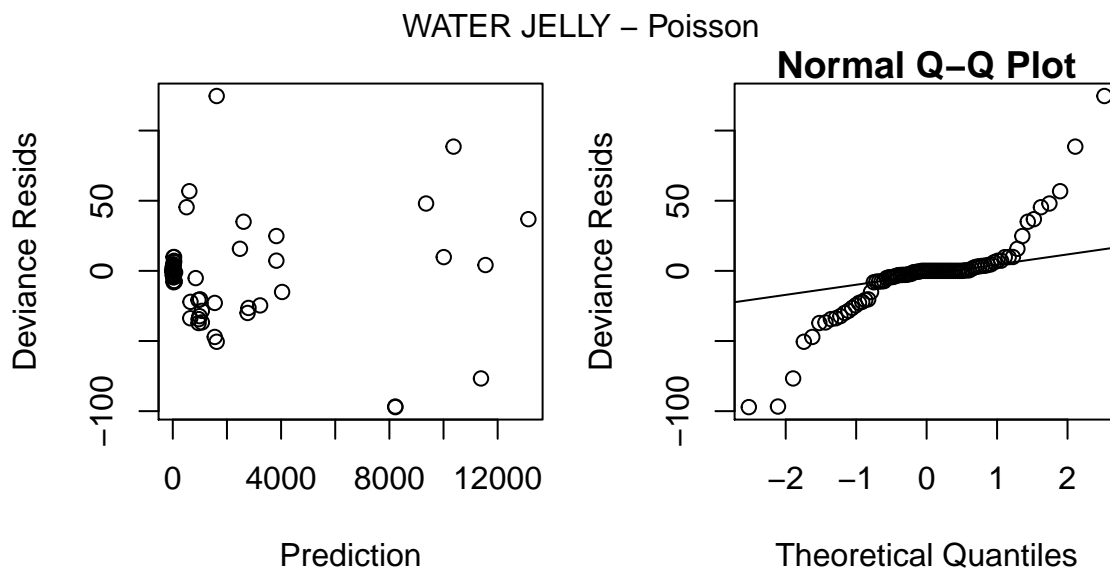


Figure 21:

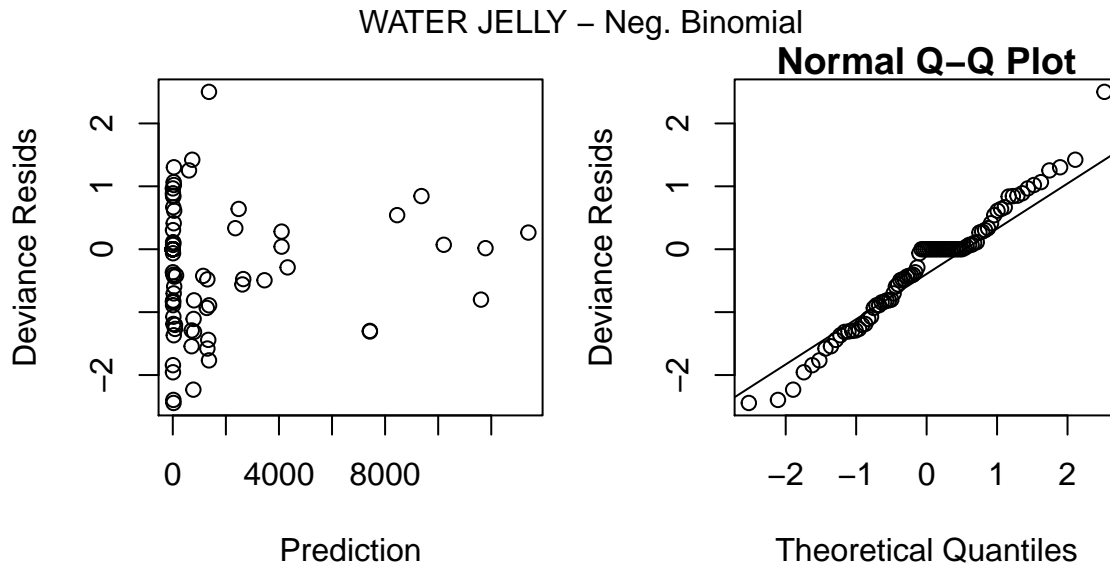


Figure 22:

```
## 0.25 1.556469 0.7193314
## 0.5 1.630563 0.7221325
## 0.75 1.708185 0.7249445
## 0.95 1.828025 0.7290641
##
## *** METHOD 2: GLM ANODEV, negative binomial ***
## Initial Theta: 0.85263
##
## Initial fit to estimate theta
## Estimated theta: 0.8668664 , SE: 0.1430373
##
## Final fit with theta = 0.8668664
##
## FIT STATISTICS:Analysis of Deviance Table
##
## Model: Negative Binomial(0.8669), link: log
##
## Response: Count
##
## Terms added sequentially (first to last)
##
##      Df Deviance Resid. Df Resid. Dev Pr(>Chi)
## NULL                85      653.02
## Block 11      572.25      74      80.77 <2e-16 ***
## Gear    2        2.98      72      77.79 0.217
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##           Up      Down
## Mean 1.9209114 0.6862222
## 0.05 0.9583039 0.3424639
## 0.25 1.3884082 0.4960827
```

```
## 0.5 1.7904148 0.6396456
## 0.75 2.3088203 0.8247545
## 0.95 3.3450611 1.1947141
```

### 3.2 Summary of results

Method abbreviations for the summary are:

- GLM.Po - GLM blocked AnoDev with Poisson
- GLM.nb - GLM blocked AnoDev with negative binomial

```
print(stat.sum)
```

```
## $`CHINOOK subadult`
## , , Up
##
##      Mean q0.05 q0.25 Median q0.75 q0.95
## GLM.Po 1.030 0.681 0.856 1.003 1.174 1.478
## GLM.nb 1.074 0.647 0.854 1.033 1.249 1.649
##
## , , Down
##
##      Mean q0.05 q0.25 Median q0.75 q0.95
## GLM.Po 1.425 0.844 1.123 1.366 1.662 2.211
## GLM.nb 1.452 0.769 1.082 1.368 1.729 2.433
##
##
## $`CHINOOK subyearling`
## , , Up
##
##      Mean q0.05 q0.25 Median q0.75 q0.95
## GLM.Po 0.985 0.903 0.95 0.983 1.018 1.072
## GLM.nb 0.894 0.538 0.71 0.860 1.040 1.373
##
## , , Down
##
##      Mean q0.05 q0.25 Median q0.75 q0.95
## GLM.Po 0.649 0.554 0.607 0.647 0.689 0.755
## GLM.nb 0.510 0.296 0.398 0.488 0.598 0.803
##
##
## $`CHINOOK yearling`
## , , Up
##
##      Mean q0.05 q0.25 Median q0.75 q0.95
## GLM.Po 0.806 0.692 0.756 0.803 0.853 0.932
## GLM.nb 0.787 0.474 0.625 0.756 0.915 1.207
##
## , , Down
##
##      Mean q0.05 q0.25 Median q0.75 q0.95
## GLM.Po 1.128 0.863 1.004 1.114 1.236 1.439
## GLM.nb 1.239 0.705 0.957 1.181 1.457 1.978
##
```

```

##
## $`CHUM SALMON`
## , , Up
##
##      Mean q0.05 q0.25 Median q0.75 q0.95
## GLM.Po    0    NA    NA    NA    NA    NA
## GLM.nb     0    NA    NA    NA    NA    NA
##
## , , Down
##
##      Mean q0.05 q0.25 Median q0.75 q0.95
## GLM.Po 0.391 0.296 0.346 0.386 0.429 0.502
## GLM.nb 0.386 0.236 0.309 0.372 0.448 0.587
##
##
## $`COHO subadult`
## , , Up
##
##      Mean q0.05 q0.25 Median q0.75 q0.95
## GLM.Po 0.619 0.397 0.507 0.601 0.711 0.909
## GLM.nb 0.664 0.382 0.516 0.634 0.779 1.053
##
## , , Down
##
##      Mean q0.05 q0.25 Median q0.75 q0.95
## GLM.Po 3.043 1.758 2.370 2.908 3.568 4.809
## GLM.nb 2.873 1.485 2.114 2.695 3.434 4.891
##
##
## $`COHO yearling`
## , , Up
##
##      Mean q0.05 q0.25 Median q0.75 q0.95
## GLM.Po 0.186 0.126 0.156 0.182 0.211 0.263
## GLM.nb 0.340 0.172 0.247 0.317 0.407 0.586
##
## , , Down
##
##      Mean q0.05 q0.25 Median q0.75 q0.95
## GLM.Po 0.705 0.487 0.599 0.690 0.794 0.977
## GLM.nb 0.719 0.351 0.514 0.667 0.867 1.269
##
##
## $`MARKET SQUID`
## , , Up
##
##      Mean q0.05 q0.25 Median q0.75 q0.95
## GLM.Po 0.350 0.310 0.333 0.349 0.366 0.393
## GLM.nb 0.877 0.452 0.644 0.822 1.048 1.494
##
## , , Down
##
##      Mean q0.05 q0.25 Median q0.75 q0.95
## GLM.Po 0.320 0.307 0.315 0.320 0.325 0.333

```

```

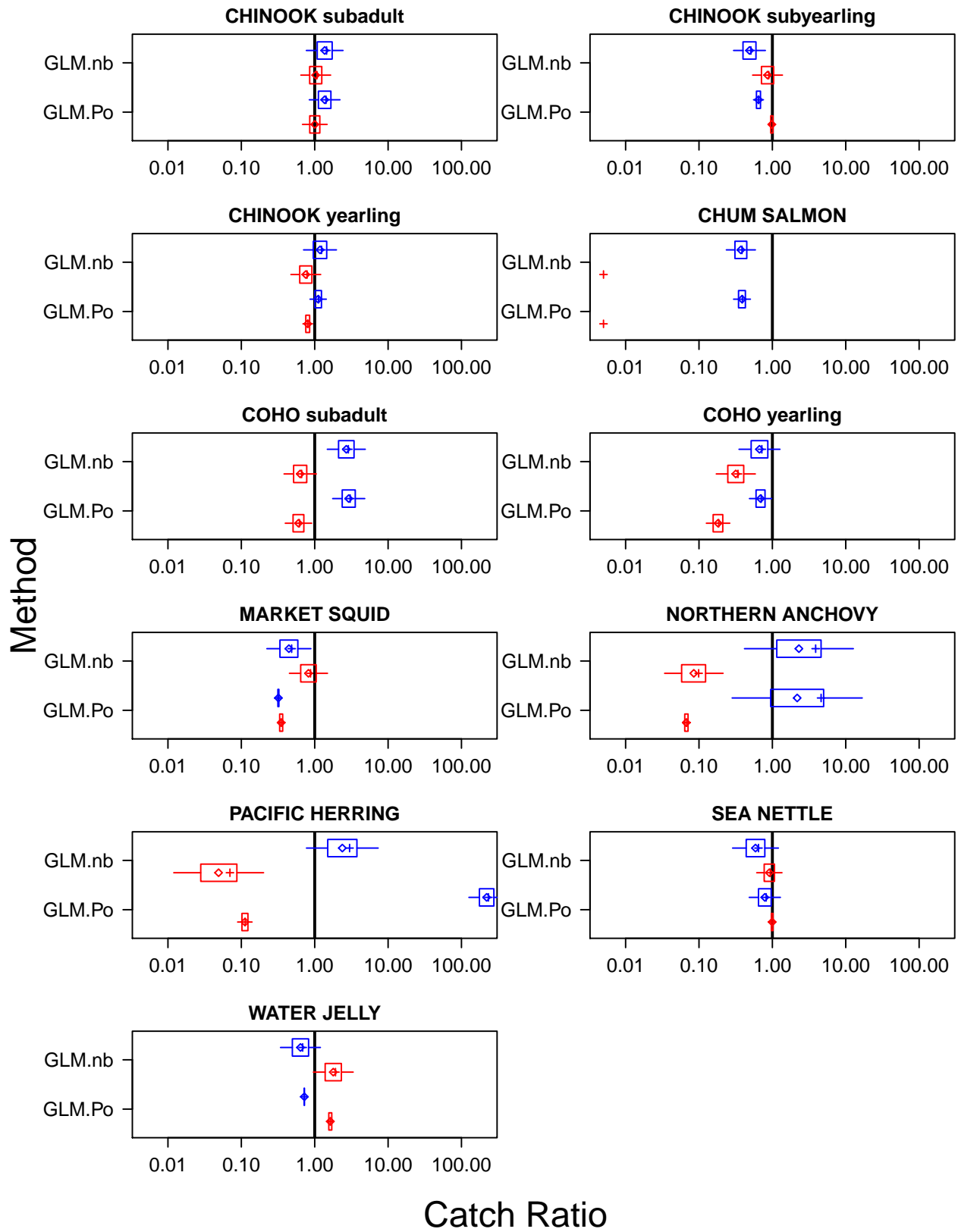
## GLM.nb 0.485 0.223 0.336 0.445 0.589 0.887
##
##
## $`NORTHERN ANCHOVY`
## , , Up
##
##      Mean q0.05 q0.25 Median q0.75 q0.95
## GLM.Po 0.068 0.060 0.064 0.067 0.071 0.076
## GLM.nb 0.099 0.034 0.058 0.085 0.123 0.213
##
## , , Down
##
##      Mean q0.05 q0.25 Median q0.75 q0.95
## GLM.Po 4.614 0.283 0.949 2.179 4.999 16.778
## GLM.nb 3.890 0.417 1.149 2.302 4.610 12.690
##
##
## $`PACIFIC HERRING`
## , , Up
##
##      Mean q0.05 q0.25 Median q0.75 q0.95
## GLM.Po 0.113 0.089 0.102 0.112 0.123 0.140
## GLM.nb 0.070 0.012 0.028 0.049 0.087 0.201
##
## , , Down
##
##      Mean q0.05 q0.25 Median q0.75 q0.95
## GLM.Po 233.083 126.305 175.740 220.424 276.471 384.679
## GLM.nb 2.988 0.772 1.505 2.379 3.759 7.326
##
##
## $`SEA NETTLE`
## , , Up
##
##      Mean q0.05 q0.25 Median q0.75 q0.95
## GLM.Po 0.995 0.932 0.968 0.994 1.020 1.060
## GLM.nb 0.936 0.612 0.774 0.910 1.069 1.353
##
## , , Down
##
##      Mean q0.05 q0.25 Median q0.75 q0.95
## GLM.Po 0.824 0.484 0.647 0.789 0.963 1.287
## GLM.nb 0.647 0.287 0.440 0.589 0.789 1.208
##
##
## $`WATER JELLY`
## , , Up
##
##      Mean q0.05 q0.25 Median q0.75 q0.95
## GLM.Po 1.634 1.454 1.556 1.631 1.708 1.828
## GLM.nb 1.921 0.958 1.388 1.790 2.309 3.345
##
## , , Down
##

```

```
##           Mean q0.05 q0.25 Median q0.75 q0.95
## GLM.Po 0.722 0.715 0.719 0.722 0.725 0.729
## GLM.nb 0.686 0.342 0.496 0.640 0.825 1.195
```

Next, generate summary figures of catch ratio estimates. In the plots, the estimated mean is marked with “+” and the median with a diamond; boxes span the quartiles, and whiskers extend to the 5% and 95% quantiles.

```
par(mfrow=c(ceiling(length(stat.sum)/plcol), plcol), oml=c(0.5,0.5,0,0), mar=c(3,4,2,1))
for (sp in names(stat.sum)) {
  .sumtbl <- stat.sum[[sp]]
  .sumtbl[.sumtbl==Inf] <- 99 # recode infinite values as +99
  .minx <- 5e-3
  .sumtbl[.sumtbl<.minx] <- .minx # recode zeros as small pos. value (for log scale plots)
  .nstats <- dim(.sumtbl)[1]
  .ngears <- dim(.sumtbl)[3]
  .ny <- .ngears*.nstats # number of elements along y axis.
  .gclr <- rep(c('red','blue','green3')[1:.ngears], .nstats) # gear color codes
  .xmax <- ceiling(max(.sumtbl[, 'q0.75', ], na.rm=TRUE)) # make sure the quartiles are covered
  .xmax <- max(.xmax, 2) #make sure upper bound is above 1
  .xmax <- min(.xmax, 10) #truncate high values so plot is readable
  .xlim <- c(.minx, 1/.minx)
  plot(t(.sumtbl[, 'Mean', ]), log='x', 1:.ny, col=.gclr, pch=3, cex=0.75, axes=F,
       xlim=.xlim, ylim=c(0.5, .ny+0.5),
       main=sp, xlab='', ylab='')
  box()
  abline(v=1, lwd=2, col='black')
  axis(side=1, cex.axis=1.2)
  axis(side=2, at=seq(1,.ny,.ngears)+1/.ngears, labels=dimnames(.sumtbl)[[1]], las=2, cex.axis=1.2)
  points(t(.sumtbl[, 'Median', ]), 1:.ny, col=.gclr, pch=5, cex=0.75)
  rect(t(.sumtbl[, 'q0.25', ]), (1:.ny)-0.35, t(.sumtbl[, 'q0.75', ]), (1:.ny)+0.35, border=.gclr)
  segments(t(.sumtbl[, 'q0.05', ]), 1:.ny, t(.sumtbl[, 'q0.25', ]), 1:.ny, col=.gclr, lwd=1)
  segments(t(.sumtbl[, 'q0.75', ]), 1:.ny, t(.sumtbl[, 'q0.95', ]), 1:.ny, col=.gclr, lwd=1)
} # for (sp)
mtext('Catch Ratio', outer=T, side=1, line=1, cex=1.5)
mtext('Method', outer=T, side=2, line=1, cex=1.5)
```



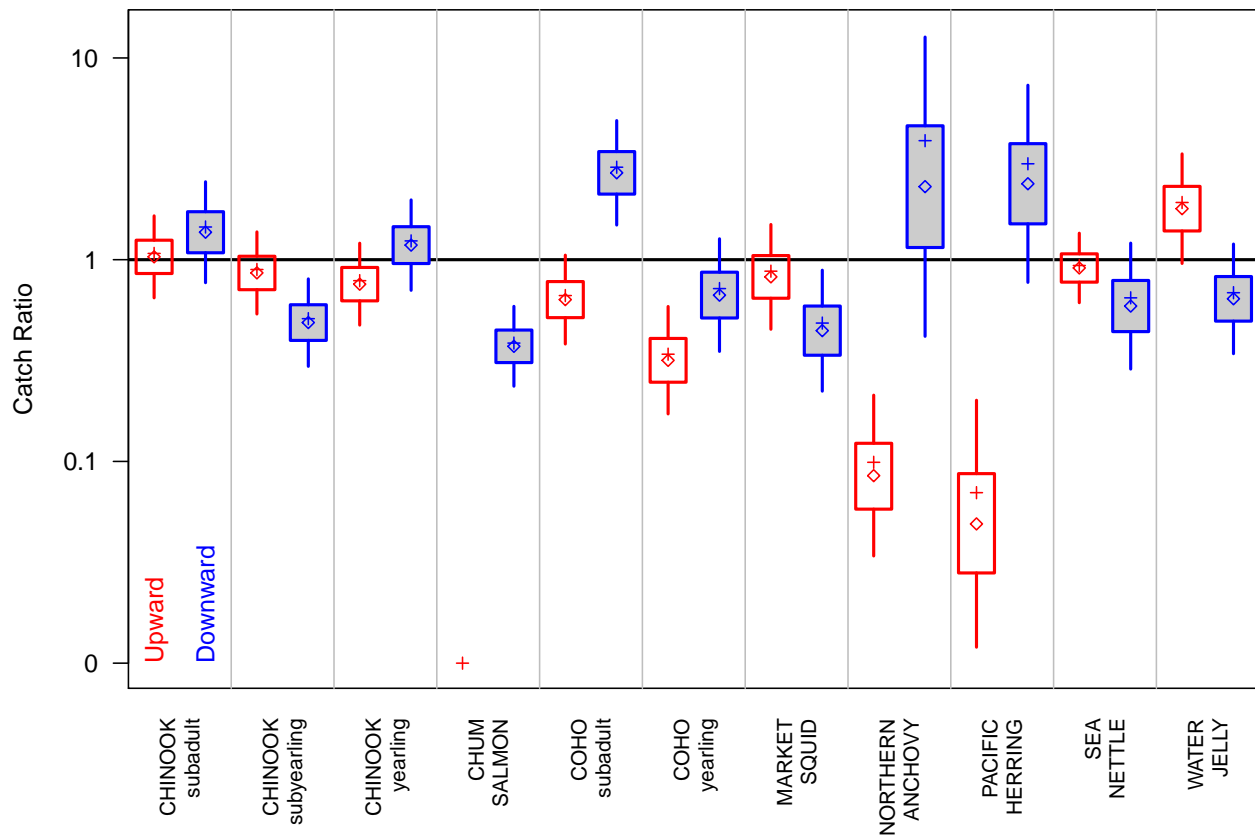
And a single panel summary of just the GLM.nb results for both excluder types.

```

par(mfrow=c(1,1), omi=c(0.5,0.5,0,0), mar=c(4,4,1,1))
.sumbtbl <- simplify2array(stat.sum)[ "GLM.nb", , , ] # array: probs x gear x species
.sumbtbl[.sumtbl==Inf] <- 99 # recode infinite values as +99
.miny <- 1e-2
.sumbtbl[.sumtbl<.minx] <- .miny # recode zeros as small pos. value (for log scale plots)
.ngears <- dim(.sumtbl)[2]
.nspecs <- dim(.sumtbl)[3]
.nx <- .ngears*.nspecs # number of elements along y axis.
.gbox <- rep(c('red','blue','green3')[1:.ngears], .nspecs) # box colors
.gpnt <- rep(c('red','blue','grey80')[1:.ngears], .nspecs) # point colors
.gfill <- rep(c(NA,'grey80','green3')[1:.ngears], .nspecs) # fill colors
.maxy <- ceiling(max(.sumtbl[ 'q0.95', , ], na.rm=TRUE))
.ylim <- c(.miny, .maxy)
plot(1:.nx, .sumtbl['Mean', , ], type='n', log='y', xaxs='i', axes=F,
     ylim=.ylim, xlim=c(0.5, .nx+0.5),
     xlab='', ylab='Catch Ratio')
box()
abline(h=1, lwd=2, col='black')
# label "fake zero" as zero:
axis(side=2, at=c(.miny,0.10,1,10,100), labels=c('0','0.1','1','10','100'), las=2)
.labs <- dimnames(.sumtbl)[[3]]
.labs <- sub(' ', '\n', .labs)
axis(side=1, at=seq(1, .nx, .ngears) + 1/.ngears, tick=FALSE, labels=.labs,
     las=2, cex.axis=0.8)
abline(v=seq(.ngears+1, .nx, .ngears) - 1/.ngears, col="grey")
rect((1:.nx)-0.35, .sumtbl['q0.25', , ], (1:.nx)+0.35, .sumtbl['q0.75', , ],
     border=.gbox, col=.gfill, lwd=2)
segments(1:.nx, .sumtbl['q0.05', , ], 1:.nx, .sumtbl['q0.25', , ], col=.gbox, lwd=2)
segments(1:.nx, .sumtbl['q0.75', , ], 1:.nx, .sumtbl['q0.95', , ], col=.gbox, lwd=2)
points(1:.nx, .sumtbl['Median', , ], col=.gpnt, pch=5, cex=0.75)
points(1:.nx, .sumtbl['Mean', , ], col=.gpnt, pch=3, cex=0.75)
text(1, .miny, "Upward", col="red", adj=c(0,0.5), srt=90)
text(2, .miny, "Downward", col="blue", adj=c(0,0.5), srt=90)

```





## Part 4: Size-selectivity analysis

#### 4.1: Size-frequency Data

We analyze size-frequencies for the same set of species selected above, but do not subdivide salmon species by age group. For an unbiased analysis, we first compute the size subsampling ratio (individuals measured / number caught) for each species in each haul, and compute adjusted counts by size. We then bin sizes into 5 mm intervals, and eliminate a couple size outliers for anchovy and water jelly (appear to be erroneous data that are overly influential in the analysis).

```
# Restrict to species selected above, but no age groups for salmon
len.spec <- c("CHINOOK SALMON", "CHUM SALMON", "COHO SALMON",
             "MARKET SQUID", "NORTHERN ANCHOVY", "PACIFIC HERRING",
             "SEA NETTLE", "WATER JELLY")
lenData <- MMEDdata[ , c("Cruise", "MMED", "Species", "Length",
                        "Number", "Distance", "Haul")]
lenData <- lenData[lenData$Species %in% len.spec, ]
# Add subsample ratio:
# Total number by Haul (rows) and Species (cols)
cnt <- with(lenData, tapply(Number, list(Haul, Species), FUN=sum, simplify=T))
cnt[is.na(cnt)] <- 0
cnt <- as.data.frame(cnt)
# Total measured by Haul & species
meas <- with(lenData, tapply(!is.na(Length), list(Haul, Species), FUN=sum, simplify=T))
meas[is.na(meas)] <- 0
meas <- as.data.frame(meas)
# Subsampling ratio by Haul & Species:
ssr <- meas / cnt
```

```
cat('\nSubsampling Ratios:\n')
```

```
##
```

```
## Subsampling Ratios:
```

```
print(summary(ssr))
```

```
## CHINOOK SALMON      CHUM SALMON      COHO SALMON      MARKET SQUID
## Min.   :0.1328      Min.   :0.6923      Min.   :0.6618      Min.   :0.008602
## 1st Qu.:0.7674      1st Qu.:0.8000      1st Qu.:1.0000      1st Qu.:0.114096
## Median :0.9045      Median :1.0000      Median :1.0000      Median :0.660256
## Mean   :0.8346      Mean   :0.9054      Mean   :0.9851      Mean   :0.542583
## 3rd Qu.:1.0000      3rd Qu.:1.0000      3rd Qu.:1.0000      3rd Qu.:1.000000
## Max.   :1.0000      Max.   :1.0000      Max.   :1.0000      Max.   :1.000000
## NA's   :2          NA's   :53         NA's   :11         NA's   :18
## NORTHERN ANCHOVY    PACIFIC HERRING    SEA NETTLE        WATER JELLY
## Min.   :0.00253     Min.   :0.01261     Min.   :0.06829     Min.   :0.000679
## 1st Qu.:0.33095     1st Qu.:0.46711     1st Qu.:0.21333     1st Qu.:0.010392
## Median :0.92857     Median :1.00000     Median :0.41463     Median :0.140222
## Mean   :0.67990     Mean   :0.75391     Mean   :0.53817     Mean   :0.287572
## 3rd Qu.:1.00000     3rd Qu.:1.00000     3rd Qu.:1.00000     3rd Qu.:0.528846
## Max.   :1.00000     Max.   :1.00000     Max.   :1.00000     Max.   :1.000000
## NA's   :66         NA's   :70         NA's   :37         NA's   :24
```

```
# Adjusted Numbers (expanded by ssr)
```

```
lenData <- lenData[!is.na(lenData$Length), ] #remove non-measured counts
```

```
lenData$AdjNum <- lenData$Number / unlist(apply(lenData[c("Haul", "Species")], 1,
function(x){ssr[x["Haul"], x["Species"]]}))
```

```
# Length bin size (mm)
```

```
binsize <- 5
```

```
lenData$LenBin <- binsize * round(lenData$Length/binsize)
```

```
# Remove size outliers for anchovy & water jelly (likely data errors)
```

```
lenData <- lenData[!(lenData$Species == "NORTHERN ANCHOVY") &
(lenData$Length < 100)], ]
```

```
lenData <- lenData[!(lenData$Species == "WATER JELLY") &
(lenData$Length > 150)], ]
```

```
cat('\nSummary of Length Data:\n')
```

```
##
```

```
## Summary of Length Data:
```

```
print(summary(lenData))
```

```
##      Cruise      MMED      Species      Length
## Min.   :41.00    Length:4385    Length:4385    Min.   : 15.0
## 1st Qu.:43.00    Class :character    Class :character    1st Qu.: 73.0
## Median :50.00    Mode  :character    Mode  :character    Median :115.0
## Mean   :47.99                                Mean   :137.1
## 3rd Qu.:53.00                                3rd Qu.:158.0
## Max.   :53.00                                Max.   :794.0
##      Number      Distance      Haul      AdjNum
## Min.   : 1.000    Min.   :1.129    Length:4385    Min.   : 1.000
## 1st Qu.: 1.000    1st Qu.:1.506    Class :character    1st Qu.: 1.179
## Median : 1.000    Median :1.860    Mode  :character    Median : 3.154
## Mean   : 1.702    Mean   :2.396                                Mean   :51.898
## 3rd Qu.: 2.000    3rd Qu.:3.513                                3rd Qu.:12.056
```

```
## Max.      :33.000    Max.      :4.906                      Max.      :5890.571
##      LenBin
## Min.      : 15.0
## 1st Qu.: 75.0
## Median :115.0
## Mean     :137.1
## 3rd Qu.:160.0
## Max.     :795.0
```

## 4.2: Analysis Methods

Next, run the Size-Frequency (SF) analysis for each species. Because the length data is sparse for most species, we analyze only the bulk data across all haul samples, ignoring the data blocks. First tabulate the distributions across all samples by species and gear type. Then, apply both a Wilcoxon-Mann-Whitney test and a Kolmogorov-Smirnov test for gear differences. **NOTE** the several warnings that p-values are approximate for the K-S test. We'll ignore the warnings, assuming they're close enough as we're concerned with the patterns of size-selectivity, and only note the highly significant results.

Then, conduct full gear Size-Selectivity Analysis, fitting a smooth curve (3rd-order polynomial) to the size-specific catch ratio data using a binomial GLM with logit link function, similar to models in Krag et al. (2014 PLOS One), Herrmann et al. (2017 Fish. Res.), & Kotwicki et al. (2017 Fish. Res.). The model is first fit to the full data set, then the error distribution is approximated via a double bootstrap (Millar 1993 Fish. Bull.) resampling both among hauls and among fish within hauls. We define a function for this analysis, so it can be applied to combinations of gear and species.

```
boot_GLM3P <- function(sdat, nrep=10, binsz=5, L.pr=NULL) {
  fit.model <- function(sdat) {
    NumTotL <- with(sdat, tapply(AdjNum, list(LenBin, MMED), sum,
                                   na.rm=TRUE, default=0))
    EffTotL <- with(sdat, tapply(Distance, list(LenBin, MMED), sum,
                                   na.rm=TRUE, default=0))

    cpue <- NumTotL/EffTotL
    cpue[is.na(cpue)] <- 0
    STD <- match("None", colnames(cpue))
    TST <- match("Up", colnames(cpue))
    if (is.na(TST)) TST <- match("Down", colnames(cpue))
    std <- cpue[ , STD]
    tst <- cpue[ , TST]
    p.L12 <- std / (std + tst)
    # Binomial weights based on number measured in both gears:
    Nmeas <- with(sdat, tapply(Number, list(LenBin, MMED), sum,
                                   na.rm=TRUE, default=0))

    wts <- Nmeas[ , STD] + Nmeas[ , TST]
    L <- as.numeric(names(p.L12))
    old.opt <- options(warn = -1) # suppress warnings about non-integer values
    fit.glm <- glm(p.L12 ~ L + I(L^2) + I(L^3), family=binomial, weights=wts)
    options(old.opt)
    return(fit.glm)
  } # fit.model()

  # Fit the model to the original (full) dataset:
  fit.full <- fit.model(sdat)
  # Predictions of full model, with rough SE's
  if (is.null(L.pr)) L.pr <- seq(min(sdat$LenBin), max(sdat$LenBin), 5)
```

```

pred.full <- predict(fit.full, newdata=data.frame(L=L.pr, wts=1.0),
                    type="response")
names(pred.full) <- L.pr
# Bootstrap predictions:
bs <- matrix(NA, nrow=length(L.pr), ncol=nrep, dimnames=list(L.pr, NULL))
rep <- 0
while (rep < nrep) {
  hauls <- unique(sdat$Haul)
  hauls.samp <- sample(hauls, length(hauls), replace=TRUE)
  .data <- data.frame()
  for (h in hauls.samp) {
    .hdata <- sdat[sdat$Haul == h, ]
    ssr <- with(.hdata, sum(Number)/sum(AdjNum)) # subsample rate
    L.ex <- with(.hdata, rep(LenBin, Number)) #expand Number to indiv. lengths
    if(length(L.ex) > 1) {
      L.smp <- sample(L.ex, length(L.ex), replace=TRUE) #resample lengths
    } else {
      L.smp <- L.ex # sample() doesn't work for length 1 vector
    }
    new.freq <- as.data.frame(table(L.smp))
    .ndata <- data.frame(MMED=unique(.hdata$MMED),
                        Haul=unique(.hdata$Haul),
                        Distance=mean(.hdata$Distance),
                        LenBin=as.numeric(levels(new.freq$L.smp)),
                        Number=new.freq$Freq,
                        AdjNum=new.freq$Freq/ssr)
    .data <- rbind(.data, .ndata)
  } # for (h)
  names(.data) <- names(sdat)
  fit.rep <- fit.model(.data)
  rep <- rep+1
  bs[, rep] <- predict(fit.rep, newdata=data.frame(L=L.pr, wts=1.0),
                      type="response")
} # for (rep)
rownames(bs) <- L.pr
bs.mn <- apply(bs, 1, mean, na.rm=FALSE)
bs.q <- t(apply(bs, 1, quantile, probs=c(0,0.05,0.25,0.50,0.75,0.95,1),
              na.rm=FALSE))
return(list(glm=fit.full, pred=pred.full, boot=bs,
            boot.sum=data.frame(mean=bs.mn, q=bs.q)))
} # boot_GLM3P()

```

### 4.3 Results

Because the two MED orientations were used in different years and locations, we can't compare either MED to all the samples from the standard net, so we run this analysis into two groups: the upward MED tests in years 2011 & 2014, and the downward MED tests in 2015.

```

# Set number of bootstrap replicates:
## nbsr <- 50   ### TESTING ###
nbsr <- 1000 ### PRODUCTION ###

for (excl in c("Up", "Down")) {

```

```

cat('\n***** Excluder: ', excl, ' *****\n')
1D <- if(excl %in% "Up") {
  lenData[lenData$Cruise %in% c(41,43,50), ]
} else {
  lenData[lenData$Cruise %in% 53, ]
}
1D$Species <- factor(as.character(1D$Species))
.tab <- with(1D, tapply(AdjNum, list(Species), sum, na.rm=T))
cat('Total Adjusted Catch:\n')
print(.tab)
lf.sel.spec <- names(.tab)[.tab >= 100]
lenFreq <- with(1D, tapply(AdjNum, list(LenBin, MMED, Species), sum, na.rm=TRUE))
lenFreq[is.na(lenFreq)] <- 0
mfrow <- if(lndscp) c(3,3) else c(3,2)
par(mfrow=mfrow, omi=c(0.5,0.5,0,0.5), mar=c(3,3,2,3))
for (sp in lf.sel.spec) {
  cat('\n*****', sp, '*****\n')
  if (sp %in% dimnames(lenFreq)[[3]]) {
    .dat <- lenFreq[, , sp]
    .maxN <- max(.dat, na.rm=T)
    .len <- 1D[1D$Species %in% sp,
      c("MMED", "Haul", "Distance", "LenBin", "Number", "AdjNum")]
    .maxL <- max(.len$LenBin, na.rm=T)
    .minL <- min(.len$LenBin, na.rm=T)
    .len.std <- .len[.len$MMED=="None", ]
    .len.mmed <- .len[.len$MMED==excl, ]
    .x <- rep(.len.std$LenBin, .len.std$Number)
    .y <- rep(.len.mmed$LenBin, .len.mmed$Number)
    # Run analysis only if > 40 measurements in each gear:
    if ((length(.x) > 40) & (length(.y) > 40)) {
      # Wilcox & KS test for overall difference in size-frequencies
      print(wilcox.test(.x, .y, alt="two.sided"))
      print(ks.test(.x, .y))
      # GLM fit of Catch Ratio to size:
      mod.fit <- boot_GLM3P(sdat=.len, nrep=nbsr, binsz=binsize)
      cat("\tSummary of GLM fit: \n")
      print(summary(mod.fit$glm))
      print(anova(mod.fit$glm, test="Chisq"))
      cat("\n\tSummary of bootstrap fits: \n")
      print(summary(mod.fit$boot.sum))
      p.pred <- mod.fit$pred
      L.pred <- as.numeric(names(p.pred))
      # Convert probability to Catch Ratio:
      CR.obs <- 1/mod.fit$glm$model$p.L12 - 1
      CR.pred <- 1/p.pred - 1
      CR.boot <- 1/mod.fit$boot.sum - 1
      CR.boot[CR.boot > 1000] <- 1000 #recode infinite values
      CR.boot[CR.boot < 1/1000] <- 1/1000 #recode zero values
      plot(as.numeric(rownames(.dat)), -.dat[, 'None'], type='h', col=BLACK, lwd=1,
        xlim=c(.minL, .maxL), ylim=c(-.maxN, .maxN), axes=F,
        xlab='', ylab='')
      box()
      axis(side=1, lwd=0, lwd.ticks=1)
    }
  }
}

```

```

axis(side=2, at=pretty(c(-max(.dat,na.rm=T),max(.dat,na.rm=T))),
      labels=abs(pretty(c(-max(.dat,na.rm=T),max(.dat,na.rm=T)))))
points(as.numeric(rownames(.dat)), .dat[, excl], type='h',
       col={if(excl %in% "Up") RED else BLUE}, lwd=1)
mtext(sp, side=3, cex=0.75, line=0)
abline(h=0, lwd=1)
# add Catch Ratio plot on right axis
par(new=TRUE)
plot(L.pred, CR.pred, log="y", type='l', lwd=3, axes=FALSE, bty="n",
     ylim=c(1/50, 50), xlab="", ylab="")
axis(side=4, at=c(0.02, 0.2, 1, 5, 50),
     labels=c("0.02", "0.2", "1", "5", "50"))
lines(L.pred, CR.boot[ , "q.50."], lty=1, col='gray50', lwd=2) # bs median
lines(L.pred, CR.boot[ , "q.5."], lty=2, col='gray50', lwd=2) # bs lower 5%
lines(L.pred, CR.boot[ , "q.95."], lty=2, col='gray50', lwd=2) # bs upper 5%
} else {
  cat('\n Insufficient data \n')
} # if (length...)
} else {
  cat('\n\tNO LENGTH DATA FOR ', sp, '\n')
} # if (sp %in% ...)
} # for(sp)
mtext('Size (mm)', side=1, outer=TRUE, at=c(0.5,0.5), cex=1.5)
mtext('Observed Catch Per Km', side=2, outer=TRUE, at=c(0.5,0.5), cex=1.5)
mtext('Estimated Catch Ratio', side=4, outer=TRUE, at=c(0.5,0.5), cex=1.5)
} # for (excl)

```

```
## Warning in ks.test(.x, .y): p-value will be approximate in the presence of
## ties
```

```
## Warning in eval(family$initialize): non-integer #successes in a binomial
## glm!
```

```
## Warning in eval(family$initialize): non-integer #successes in a binomial
## glm!
```

```
## Warning in ks.test(.x, .y): cannot compute exact p-value with ties
```

```
## Warning in eval(family$initialize): non-integer #successes in a binomial
## glm!
```

```
## Warning in eval(family$initialize): non-integer #successes in a binomial
## glm!
```

```
## Warning in ks.test(.x, .y): p-value will be approximate in the presence of
## ties
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## Warning in eval(family$initialize): non-integer #successes in a binomial
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```

```
## Warning in eval(family$initialize): non-integer #successes in a binomial
## glm!
```

```
## Warning in ks.test(.x, .y): p-value will be approximate in the presence of
## ties
```

```
## Warning in eval(family$initialize): non-integer #successes in a binomial
```

```

## glm!

## Warning in eval(family$initialize): non-integer #successes in a binomial
## glm!

## Warning in ks.test(.x, .y): p-value will be approximate in the presence of
## ties

## Warning in eval(family$initialize): non-integer #successes in a binomial
## glm!

## Warning in eval(family$initialize): non-integer #successes in a binomial
## glm!

## Warning in ks.test(.x, .y): p-value will be approximate in the presence of
## ties

## Warning in eval(family$initialize): non-integer #successes in a binomial
## glm!

## Warning in eval(family$initialize): non-integer #successes in a binomial
## glm!

## Warning in ks.test(.x, .y): p-value will be approximate in the presence of
## ties

## Warning in eval(family$initialize): non-integer #successes in a binomial
## glm!

## Warning in eval(family$initialize): non-integer #successes in a binomial
## glm!

## Warning in ks.test(.x, .y): cannot compute exact p-value with ties

## Warning in ks.test(.x, .y): non-integer #successes in a binomial glm!

## Warning in ks.test(.x, .y): non-integer #successes in a binomial glm!

## Warning in ks.test(.x, .y): cannot compute exact p-value with ties

## Warning in eval(family$initialize): non-integer #successes in a binomial
## glm!

## Warning in eval(family$initialize): non-integer #successes in a binomial
## glm!

## Warning in ks.test(.x, .y): p-value will be approximate in the presence of
## ties

## Warning in eval(family$initialize): non-integer #successes in a binomial
## glm!

## Warning in eval(family$initialize): non-integer #successes in a binomial
## glm!

## Warning in ks.test(.x, .y): p-value will be approximate in the presence of
## ties

## Warning in eval(family$initialize): non-integer #successes in a binomial
## glm!

```

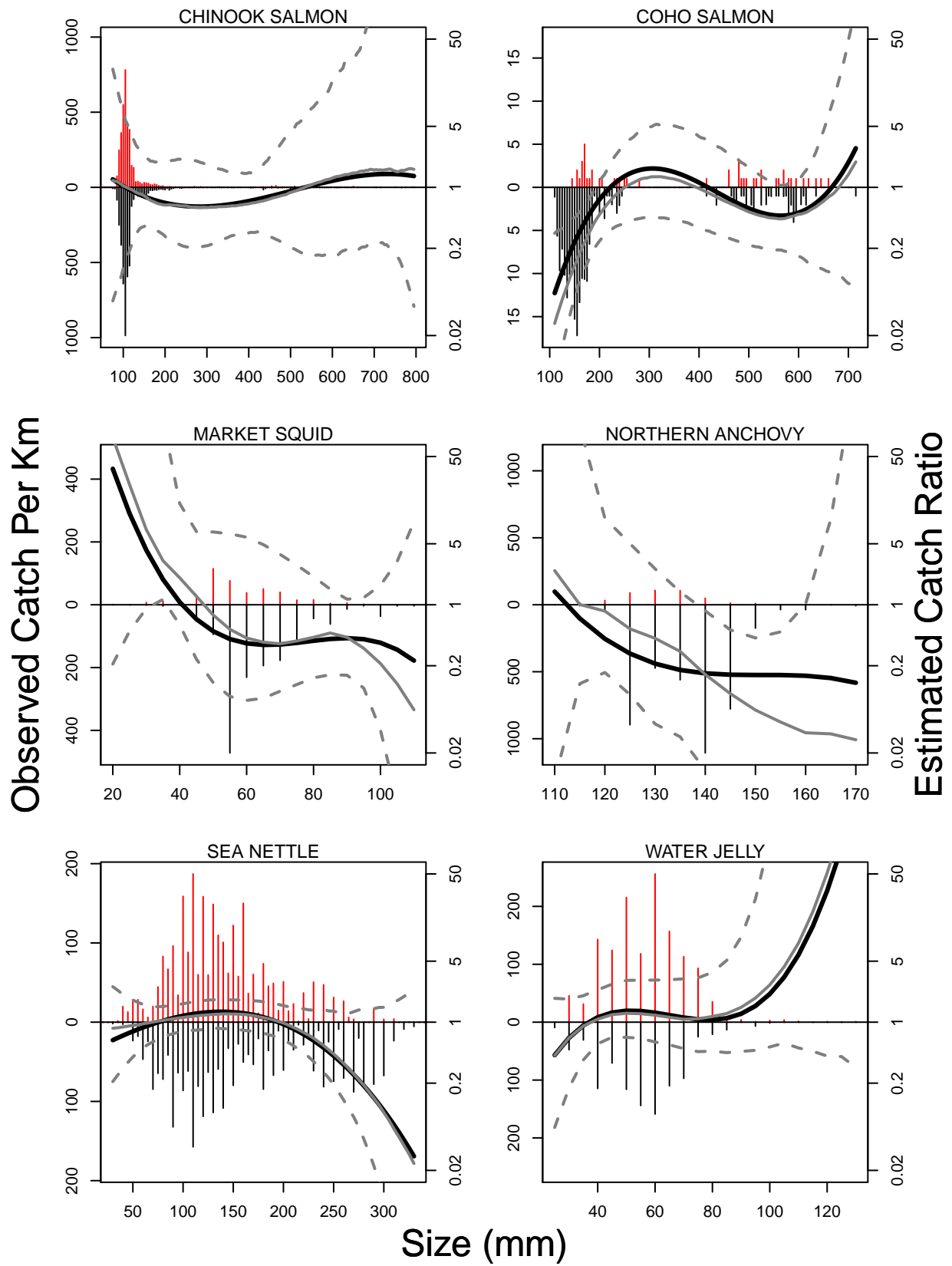


Figure 23:



```

## Warning in eval(family$initialize): non-integer #successes in a binomial
## glm!

##
## ***** Excluder: Up *****
## Total Adjusted Catch:
##   CHINOOK SALMON      CHUM SALMON      COHO SALMON      MARKET SQUID
##      7806.36006        24.34091        271.59921        1930.73203
## NORTHERN ANCHOVY  PACIFIC HERRING      SEA NETTLE      WATER JELLY
##      4545.80019        788.94286        5328.01283        2321.85789
##
## ***** CHINOOK SALMON *****
##
## Wilcoxon rank sum test with continuity correction
##
## data: .x and .y
## W = 557520, p-value = 0.3175
## alternative hypothesis: true location shift is not equal to 0
##
##
## Two-sample Kolmogorov-Smirnov test
##
## data: .x and .y
## D = 0.028936, p-value = 0.7764
## alternative hypothesis: two-sided
##
## Summary of GLM fit:
##
## Call:
## glm(formula = p.L12 ~ L + I(L^2) + I(L^3), family = binomial,
##      weights = wts)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -2.4824  -1.0048   0.0389   1.0590   3.2229
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept) -9.376e-01  3.682e-01  -2.546  0.0109 *
## L           1.175e-02  4.754e-03   2.471  0.0135 *
## I(L^2)      -2.881e-05  1.573e-05  -1.832  0.0670 .
## I(L^3)       1.893e-08  1.423e-08   1.330  0.1835
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 132.65  on 87  degrees of freedom
## Residual deviance: 120.58  on 84  degrees of freedom
## AIC: 275.2
##
## Number of Fisher Scoring iterations: 4
##

```

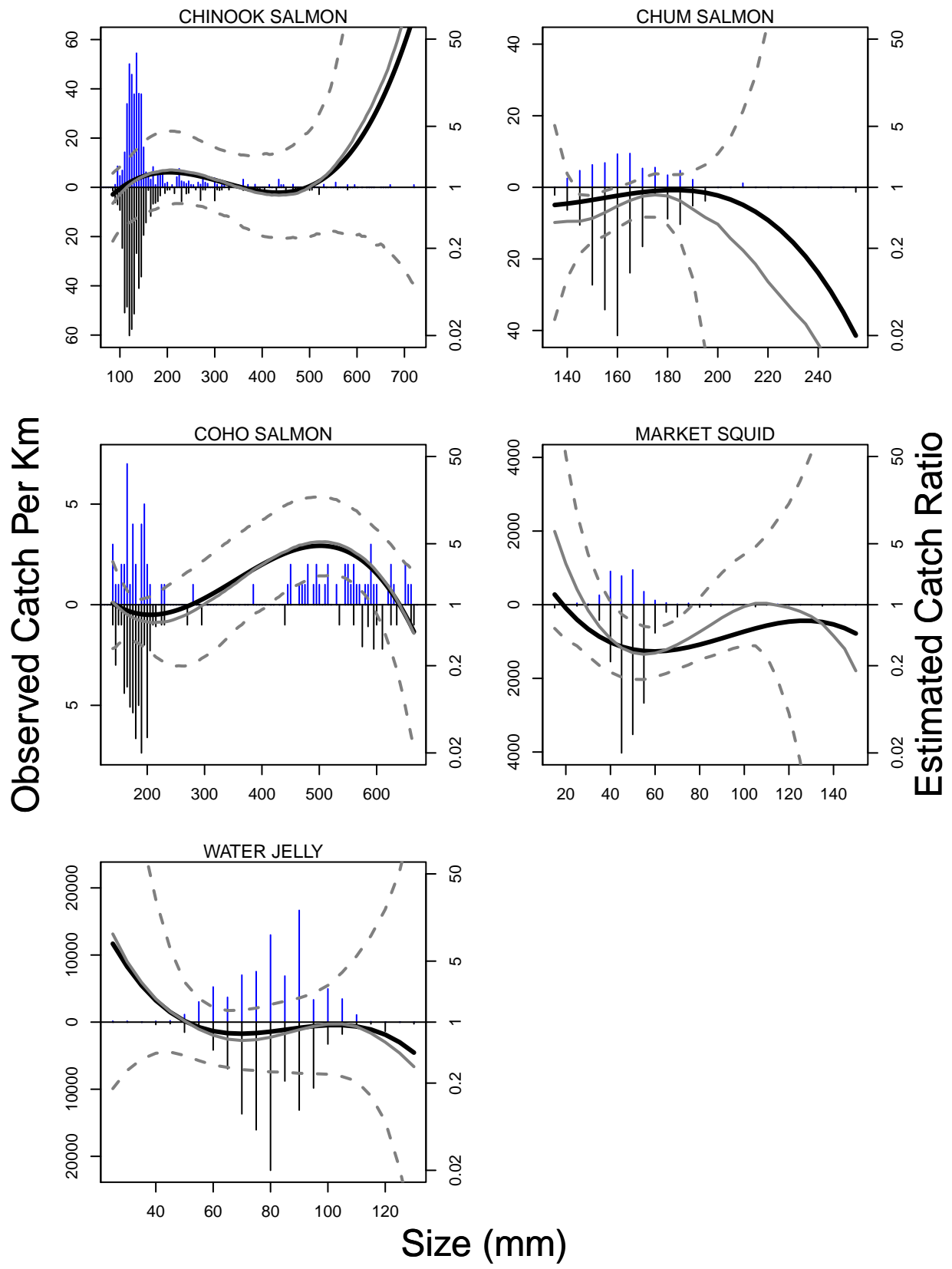


Figure 24:

```

## Analysis of Deviance Table
##
## Model: binomial, link: logit
##
## Response: p.L12
##
## Terms added sequentially (first to last)
##
##           Df Deviance Resid. Df Resid. Dev Pr(>Chi)
## NULL                87      132.65
## L             1    1.8014      86      130.85  0.1795
## I(L^2)        1    8.5266      85      122.32  0.0035 **
## I(L^3)        1    1.7446      84      120.58  0.1866
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Summary of bootstrap fits:
##           mean           q.0.           q.5.           q.25.
## Min.      :0.3934   Min.      :0.00000   Min.      :0.0000753   Min.      :0.1196
## 1st Qu.:0.4412   1st Qu.:0.00000   1st Qu.:0.0438000   1st Qu.:0.2456
## Median :0.5291   Median :0.01454   Median :0.2182649   Median :0.4172
## Mean      :0.5167   Mean      :0.05429   Mean      :0.2035302   Mean      :0.3756
## 3rd Qu.:0.5941   3rd Qu.:0.09490   3rd Qu.:0.3370246   3rd Qu.:0.4924
## Max.      :0.6155   Max.      :0.20724   Max.      :0.4112790   Max.      :0.5377
##           q.50.           q.75.           q.95.           q.100.
## Min.      :0.3786   Min.      :0.5583   Min.      :0.7325   Min.      :0.8186
## 1st Qu.:0.4281   1st Qu.:0.6313   1st Qu.:0.7898   1st Qu.:0.9282
## Median :0.5381   Median :0.6581   Median :0.8231   Median :0.9721
## Mean      :0.5206   Mean      :0.6544   Mean      :0.8227   Mean      :0.9589
## 3rd Qu.:0.6136   3rd Qu.:0.6899   3rd Qu.:0.8483   3rd Qu.:0.9987
## Max.      :0.6295   Max.      :0.7293   Max.      :0.9586   Max.      :1.0000
##
## ***** COHO SALMON *****
##
## Wilcoxon rank sum test with continuity correction
##
## data:  .x and .y
## W = 2478.5, p-value = 0.000001831
## alternative hypothesis: true location shift is not equal to 0
##
##
## Two-sample Kolmogorov-Smirnov test
##
## data:  .x and .y
## D = 0.37991, p-value = 0.00002893
## alternative hypothesis: two-sided
##
## Summary of GLM fit:
##
## Call:
## glm(formula = p.L12 ~ L + I(L^2) + I(L^3), family = binomial,
##      weights = wts)
##

```

```

## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -2.3673  -0.9730   0.1931   1.2725   2.2525
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)  9.004e+00  2.250e+00   4.001 0.000063 ***
## L            -7.567e-02  2.202e-02  -3.437 0.000588 ***
## I(L^2)        1.904e-04  6.037e-05   3.154 0.001610 **
## I(L^3)       -1.457e-07  5.075e-08  -2.872 0.004084 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 106.243  on 63  degrees of freedom
## Residual deviance:  87.423  on 60  degrees of freedom
## AIC: 143.29
##
## Number of Fisher Scoring iterations: 4
##
## Analysis of Deviance Table
##
## Model: binomial, link: logit
##
## Response: p.L12
##
## Terms added sequentially (first to last)
##
##
##      Df Deviance Resid. Df Resid. Dev Pr(>Chi)
## NULL                                63    106.243
## L              1    4.8453             62    101.398 0.02772 *
## I(L^2)         1    5.1492             61     96.249 0.02326 *
## I(L^3)         1    8.8261             60     87.423 0.00297 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Summary of bootstrap fits:
##      mean          q.0.          q.5.          q.25.
## Min.    :0.3854   Min.    :0.0002299   Min.    :0.008153   Min.    :0.06874
## 1st Qu.:0.4684   1st Qu.:0.0481034   1st Qu.:0.181501   1st Qu.:0.33857
## Median :0.5765   Median :0.0860872   Median :0.287034   Median :0.46960
## Mean    :0.5887   Mean    :0.1431427   Mean    :0.325847   Mean    :0.48068
## 3rd Qu.:0.6683   3rd Qu.:0.2050709   3rd Qu.:0.455102   3rd Qu.:0.59969
## Max.    :0.9414   Max.    :0.4960444   Max.    :0.770876   Max.    :0.93796
##      q.50.          q.75.          q.95.          q.100.
## Min.    :0.3382   Min.    :0.5485   Min.    :0.6862   Min.    :0.8389
## 1st Qu.:0.4832   1st Qu.:0.6094   1st Qu.:0.7338   1st Qu.:0.9003
## Median :0.5936   Median :0.7083   Median :0.8199   Median :0.9219
## Mean    :0.6028   Mean    :0.7021   Mean    :0.8147   Mean    :0.9310
## 3rd Qu.:0.6808   3rd Qu.:0.7533   3rd Qu.:0.8868   3rd Qu.:0.9758
## Max.    :0.9732   Max.    :0.9863   Max.    :0.9949   Max.    :0.9999
##

```

```

## ***** MARKET SQUID *****
##
## Wilcoxon rank sum test with continuity correction
##
## data: .x and .y
## W = 23156, p-value = 4.907e-08
## alternative hypothesis: true location shift is not equal to 0
##
## Two-sample Kolmogorov-Smirnov test
##
## data: .x and .y
## D = 0.23432, p-value = 0.00006958
## alternative hypothesis: two-sided
##
## Summary of GLM fit:
##
## Call:
## glm(formula = p.L12 ~ L + I(L^2) + I(L^3), family = binomial,
##      weights = wts)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -2.9502  -0.9026   0.2298   1.1338   2.3714
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept) -10.81427430   3.97277650  -2.722  0.00649 **
## L              0.47533570   0.19060262   2.494  0.01264 *
## I(L^2)        -0.00623093   0.00292930  -2.127  0.03341 *
## I(L^3)         0.00002659   0.00001440   1.846  0.06482 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 57.182  on 17  degrees of freedom
## Residual deviance: 35.140  on 14  degrees of freedom
## AIC: 91.578
##
## Number of Fisher Scoring iterations: 5
##
## Analysis of Deviance Table
##
## Model: binomial, link: logit
##
## Response: p.L12
##
## Terms added sequentially (first to last)
##
##      Df Deviance Resid. Df Resid. Dev Pr(>Chi)
## NULL              17      57.182
## L               1  10.0445        16      47.138 0.001528 **

```

```

## I(L^2)  1   8.2683          15   38.869 0.004034 **
## I(L^3)  1   3.7298          14   35.140 0.053450 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Summary of bootstrap fits:
##      mean          q.0.          q.5.          q.25.
## Min.   :0.1580   Min.   :0.000000   Min.   :0.00000   Min.   :0.0003288
## 1st Qu.:0.3921   1st Qu.:0.000245   1st Qu.:0.08284   1st Qu.:0.2295661
## Median :0.6117   Median :0.028827   Median :0.14399   Median :0.3013745
## Mean   :0.5373   Mean   :0.057190   Mean   :0.18944   Mean   :0.3720510
## 3rd Qu.:0.6721   3rd Qu.:0.084406   3rd Qu.:0.29174   3rd Qu.:0.5775720
## Max.   :0.8371   Max.   :0.208333   Max.   :0.46349   Max.   :0.8156187
##      q.50.          q.75.          q.95.          q.100.
## Min.   :0.01195   Min.   :0.1848   Min.   :0.4649   Min.   :0.8525
## 1st Qu.:0.38712   1st Qu.:0.5279   1st Qu.:0.8176   1st Qu.:0.9659
## Median :0.70216   Median :0.7909   Median :0.8864   Median :0.9736
## Mean   :0.56856   Mean   :0.6814   Mean   :0.8334   Mean   :0.9700
## 3rd Qu.:0.73289   3rd Qu.:0.8380   3rd Qu.:0.9201   3rd Qu.:0.9998
## Max.   :0.94112   Max.   :0.9872   Max.   :0.9994   Max.   :1.0000
##
## ***** NORTHERN ANCHOVY *****
##
## Wilcoxon rank sum test with continuity correction
##
## data:  .x and .y
## W = 24497, p-value = 3.043e-09
## alternative hypothesis: true location shift is not equal to 0
##
##
## Two-sample Kolmogorov-Smirnov test
##
## data:  .x and .y
## D = 0.32536, p-value = 6.01e-09
## alternative hypothesis: two-sided
##
## Summary of GLM fit:
##
## Call:
## glm(formula = p.L12 ~ L + I(L^2) + I(L^3), family = binomial,
##      weights = wts)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -1.58746  -1.04605  -0.00453   1.13477   1.85831
##
## Coefficients:
##              Estimate      Std. Error z value Pr(>|z|)
## (Intercept) -106.87895097  145.70040268  -0.734    0.463
## L             2.15694582   3.19635961   0.675    0.500
## I(L^2)        -0.01426172   0.02328293  -0.613    0.540
## I(L^3)         0.00003143   0.00005632   0.558    0.577
##
## (Dispersion parameter for binomial family taken to be 1)

```

```

##
## Null deviance: 26.876 on 11 degrees of freedom
## Residual deviance: 18.727 on 8 degrees of freedom
## AIC: 54.403
##
## Number of Fisher Scoring iterations: 4
##
## Analysis of Deviance Table
##
## Model: binomial, link: logit
##
## Response: p.L12
##
## Terms added sequentially (first to last)
##
##
## Df Deviance Resid. Df Resid. Dev Pr(>Chi)
## NULL 11 26.876
## L 1 6.1933 10 20.683 0.01282 *
## I(L^2) 1 1.6251 9 19.058 0.20238
## I(L^3) 1 0.3305 8 18.727 0.56537
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Summary of bootstrap fits:
## mean q.0. q.5. q.25.
## Min. :0.4058 Min. :0.0000000 Min. :0.000028 Min. :0.05113
## 1st Qu.:0.6008 1st Qu.:0.0000001 1st Qu.:0.092484 1st Qu.:0.42175
## Median :0.7934 Median :0.0324816 Median :0.279141 Median :0.77341
## Mean :0.7265 Mean :0.0407347 Mean :0.317851 Mean :0.60818
## 3rd Qu.:0.8841 3rd Qu.:0.0691386 3rd Qu.:0.539925 3rd Qu.:0.84873
## Max. :0.9117 Max. :0.1044731 Max. :0.705958 Max. :0.88293
## q.50. q.75. q.95. q.100.
## Min. :0.2888 Min. :0.6826 Min. :0.8567 Min. :1
## 1st Qu.:0.6539 1st Qu.:0.7948 1st Qu.:0.9587 1st Qu.:1
## Median :0.8629 Median :0.9295 Median :0.9902 Median :1
## Mean :0.7728 Mean :0.8904 Mean :0.9666 Mean :1
## 3rd Qu.:0.9567 3rd Qu.:0.9948 3rd Qu.:1.0000 3rd Qu.:1
## Max. :0.9724 Max. :1.0000 Max. :1.0000 Max. :1
##
## ***** PACIFIC HERRING *****
##
## Insufficient data
##
## ***** SEA NETTLE *****
##
## Wilcoxon rank sum test with continuity correction
##
## data: .x and .y
## W = 161020, p-value = 0.9677
## alternative hypothesis: true location shift is not equal to 0
##
##
## Two-sample Kolmogorov-Smirnov test

```

```

##
## data: .x and .y
## D = 0.064328, p-value = 0.1919
## alternative hypothesis: two-sided
##
## Summary of GLM fit:
##
## Call:
## glm(formula = p.L12 ~ L + I(L^2) + I(L^3), family = binomial,
##      weights = wts)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -2.3007  -0.8560   0.1096   0.9071   2.5380
##
## Coefficients:
##              Estimate      Std. Error z value Pr(>|z|)
## (Intercept)  0.8708333569   0.8311327335   1.048   0.295
## L            -0.0140396513   0.0180557455  -0.778   0.437
## I(L^2)        0.0000232098   0.0001210210   0.192   0.848
## I(L^3)        0.0000001328   0.0000002502   0.531   0.596
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 113.782  on 55  degrees of freedom
## Residual deviance:  69.779  on 52  degrees of freedom
## AIC: 232.81
##
## Number of Fisher Scoring iterations: 4
##
## Analysis of Deviance Table
##
## Model: binomial, link: logit
##
## Response: p.L12
##
## Terms added sequentially (first to last)
##
##
##      Df Deviance Resid. Df Resid. Dev  Pr(>Chi)
## NULL                    55    113.782
## L              1  15.2083         54     98.573 9.628e-05 ***
## I(L^2)         1  28.5074         53     70.066 9.334e-08 ***
## I(L^3)         1   0.2869         52     69.779  0.5922
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Summary of bootstrap fits:
##      mean          q.0.          q.5.          q.25.
## Min.      :0.4466   Min.      :0.00009821   Min.      :0.2816   Min.      :0.4048
## 1st Qu.:0.4692   1st Qu.:0.08633950   1st Qu.:0.3586   1st Qu.:0.4263
## Median :0.5208   Median :0.22947680   Median :0.3850   Median :0.4538
## Mean      :0.5887   Mean      :0.18142280   Mean      :0.3813   Mean      :0.5487
## 3rd Qu.:0.7009   3rd Qu.:0.26332753   3rd Qu.:0.4080   3rd Qu.:0.6616

```



```

## Max. :0.8944 Max. :0.29427590 Max. :0.4427 Max. :0.9224
## q.50. q.75. q.95. q.100.
## Min. :0.4456 Min. :0.4848 Min. :0.5399 Min. :0.6495
## 1st Qu.:0.4693 1st Qu.:0.5093 1st Qu.:0.5636 1st Qu.:0.6995
## Median :0.5194 Median :0.5829 Median :0.6652 Median :0.8166
## Mean :0.6033 Mean :0.6523 Mean :0.7193 Mean :0.8318
## 3rd Qu.:0.7231 3rd Qu.:0.7791 3rd Qu.:0.8637 3rd Qu.:0.9769
## Max. :0.9766 Max. :0.9937 Max. :0.9995 Max. :1.0000
##
## ***** WATER JELLY *****
##
## Wilcoxon rank sum test with continuity correction
##
## data: .x and .y
## W = 30119, p-value = 0.0192
## alternative hypothesis: true location shift is not equal to 0
##
##
## Two-sample Kolmogorov-Smirnov test
##
## data: .x and .y
## D = 0.099297, p-value = 0.1566
## alternative hypothesis: two-sided
##
## Summary of GLM fit:
##
## Call:
## glm(formula = p.L12 ~ L + I(L^2) + I(L^3), family = binomial,
## weights = wts)
##
## Deviance Residuals:
## Min 1Q Median 3Q Max
## -2.15870 -0.60214 0.04482 0.72553 1.81500
##
## Coefficients:
## Estimate Std. Error z value Pr(>|z|)
## (Intercept) 5.73520424 2.98626868 1.921 0.0548 .
## L -0.29910317 0.15325488 -1.952 0.0510 .
## I(L^2) 0.00479145 0.00251495 1.905 0.0568 .
## I(L^3) -0.00002448 0.00001324 -1.849 0.0644 .
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
## Null deviance: 29.520 on 19 degrees of freedom
## Residual deviance: 23.455 on 16 degrees of freedom
## AIC: 82.13
##
## Number of Fisher Scoring iterations: 5
##
## Analysis of Deviance Table
##
## Model: binomial, link: logit

```

```

##
## Response: p.L12
##
## Terms added sequentially (first to last)
##
##
##      Df Deviance Resid. Df Resid. Dev Pr(>Chi)
## NULL                19      29.520
## L      1    0.9636      18      28.556 0.32627
## I(L^2) 1    0.0609      17      28.495 0.80513
## I(L^3) 1    5.0398      16      23.455 0.02477 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Summary of bootstrap fits:
##      mean      q.0.      q.5.      q.25.
## Min.    :0.0843  Min.    :0.00000  Min.    :0.00000  Min.    :0.0000198
## 1st Qu.:0.2448  1st Qu.:0.00000  1st Qu.:0.003028  1st Qu.:0.0860607
## Median :0.4369  Median :0.04005  Median :0.224277  Median :0.3505161
## Mean    :0.3729  Mean    :0.05378  Mean    :0.163054  Mean    :0.2670565
## 3rd Qu.:0.4642  3rd Qu.:0.09941  3rd Qu.:0.249009  3rd Qu.:0.3689303
## Max.    :0.6809  Max.    :0.15506  Max.    :0.352928  Max.    :0.5628794
##      q.50.      q.75.      q.95.      q.100.
## Min.    :0.001415  Min.    :0.01977  Min.    :0.5980  Min.    :0.7273
## 1st Qu.:0.209401  1st Qu.:0.35123  1st Qu.:0.6406  1st Qu.:0.8366
## Median :0.443781  Median :0.52233  Median :0.6801  Median :0.9337
## Mean    :0.356115  Mean    :0.44815  Mean    :0.6888  Mean    :0.9077
## 3rd Qu.:0.475020  3rd Qu.:0.56342  3rd Qu.:0.7051  3rd Qu.:0.9999
## Max.    :0.706089  Max.    :0.82215  Max.    :0.9415  Max.    :1.0000
##
## ***** Excluder:  Down *****
## Total Adjusted Catch:
##      CHINOOK SALMON      CHUM SALMON      COHO SALMON      MARKET SQUID
##      1002.0864      255.3848      155.4833      17572.2192
## NORTHERN ANCHOVY  PACIFIC HERRING      SEA NETTLE      WATER JELLY
##      3.0000      2400.1818      47.0000      183119.7267
##
## ***** CHINOOK SALMON *****
##
## Wilcoxon rank sum test with continuity correction
##
## data:  .x and .y
## W = 34425, p-value = 0.000005302
## alternative hypothesis: true location shift is not equal to 0
##
##
## Two-sample Kolmogorov-Smirnov test
##
## data:  .x and .y
## D = 0.17929, p-value = 0.0001577
## alternative hypothesis: two-sided
##
## Summary of GLM fit:
##

```

```

## Call:
## glm(formula = p.L12 ~ L + I(L^2) + I(L^3), family = binomial,
##      weights = wts)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -2.27913  -1.05892   0.02342   0.95216   1.90680
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)  1.645e+00  1.030e+00   1.597   0.1102
## L           -2.376e-02  1.329e-02  -1.787   0.0739 .
## I(L^2)        8.571e-05  4.855e-05   1.766   0.0775 .
## I(L^3)       -8.979e-08  5.197e-08  -1.728   0.0840 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 79.036  on 74  degrees of freedom
## Residual deviance: 74.352  on 71  degrees of freedom
## AIC: 175.45
##
## Number of Fisher Scoring iterations: 5
##
## Analysis of Deviance Table
##
## Model: binomial, link: logit
##
## Response: p.L12
##
## Terms added sequentially (first to last)
##
##
##      Df Deviance Resid. Df Resid. Dev Pr(>Chi)
## NULL                                74      79.036
## L              1    0.5877           73      78.449  0.44333
## I(L^2)         1    0.0303           72      78.418  0.86174
## I(L^3)         1    4.0662           71      74.352  0.04375 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Summary of bootstrap fits:
##      mean          q.0.          q.5.          q.25.
## Min.    :0.1137   Min.    :0.00000   Min.    :0.00000   Min.    :0.0000496
## 1st Qu.:0.3509   1st Qu.:0.00000   1st Qu.:0.02879   1st Qu.:0.1964383
## Median :0.4295   Median :0.01713   Median :0.20650   Median :0.3300689
## Mean    :0.4009   Mean    :0.04579   Mean    :0.17861   Mean    :0.2866721
## 3rd Qu.:0.5074   3rd Qu.:0.07319   3rd Qu.:0.27843   3rd Qu.:0.4121642
## Max.    :0.6056   Max.    :0.19220   Max.    :0.40821   Max.    :0.5294551
##      q.50.          q.75.          q.95.          q.100.
## Min.    :0.00395   Min.    :0.0503   Min.    :0.6053   Min.    :0.7908
## 1st Qu.:0.33047   1st Qu.:0.4766   1st Qu.:0.6688   1st Qu.:0.8369
## Median :0.42611   Median :0.5201   Median :0.7691   Median :0.9654

```

```

## Mean :0.37941 Mean :0.4828 Mean :0.7422 Mean :0.9209
## 3rd Qu.:0.51048 3rd Qu.:0.6084 3rd Qu.:0.7908 3rd Qu.:1.0000
## Max. :0.60440 Max. :0.6838 Max. :0.9284 Max. :1.0000
##
## ***** CHUM SALMON *****
##
## Wilcoxon rank sum test with continuity correction
##
## data: .x and .y
## W = 3579, p-value = 0.1798
## alternative hypothesis: true location shift is not equal to 0
##
##
## Two-sample Kolmogorov-Smirnov test
##
## data: .x and .y
## D = 0.14104, p-value = 0.4074
## alternative hypothesis: two-sided
##
## Summary of GLM fit:
##
## Call:
## glm(formula = p.L12 ~ L + I(L^2) + I(L^3), family = binomial,
## weights = wts)
##
## Deviance Residuals:
## Min 1Q Median 3Q Max
## -1.3788 -0.6636 0.1496 0.5312 1.9372
##
## Coefficients:
## Estimate Std. Error z value Pr(>|z|)
## (Intercept) -15.301418766 67.256511774 -0.228 0.820
## L 0.324906462 1.156819324 0.281 0.779
## I(L^2) -0.002174849 0.006600616 -0.329 0.742
## I(L^3) 0.000004691 0.000012502 0.375 0.707
##
## (Dispersion parameter for binomial family taken to be 1)
##
## Null deviance: 14.167 on 14 degrees of freedom
## Residual deviance: 12.646 on 11 degrees of freedom
## AIC: 55.107
##
## Number of Fisher Scoring iterations: 6
##
## Analysis of Deviance Table
##
## Model: binomial, link: logit
##
## Response: p.L12
##
## Terms added sequentially (first to last)
##
##
## Df Deviance Resid. Df Resid. Dev Pr(>Chi)

```

```

## NULL          14      14.166
## L            1  0.03756      13      14.129      0.8463
## I(L^2)       1  1.29445      12      12.835      0.2552
## I(L^3)       1  0.18892      11      12.646      0.6638
##
## Summary of bootstrap fits:
##      mean          q.0.          q.5.          q.25.
## Min.    :0.5495    Min.    :0.00000    Min.    :0.00000    Min.    :0.4308
## 1st Qu.:0.6457    1st Qu.:0.00000    1st Qu.:0.002317   1st Qu.:0.4945
## Median :0.7052    Median :0.01162    Median :0.292037   Median :0.5199
## Mean    :0.6778    Mean    :0.11755    Mean    :0.251715   Mean    :0.5689
## 3rd Qu.:0.7151    3rd Qu.:0.21143    3rd Qu.:0.417583   3rd Qu.:0.6082
## Max.    :0.7662    Max.    :0.41454    Max.    :0.554902   Max.    :0.8774
##      q.50.          q.75.          q.95.          q.100.
## Min.    :0.5499    Min.    :0.6051    Min.    :0.6868    Min.    :0.8499
## 1st Qu.:0.6400    1st Qu.:0.7066    1st Qu.:0.8044    1st Qu.:0.9311
## Median :0.7166    Median :0.8895    Median :0.9863    Median :1.0000
## Mean    :0.7721    Mean    :0.8531    Mean    :0.9037    Mean    :0.9655
## 3rd Qu.:0.9465    3rd Qu.:1.0000    3rd Qu.:1.0000    3rd Qu.:1.0000
## Max.    :0.9982    Max.    :1.0000    Max.    :1.0000    Max.    :1.0000
##
## ***** COHO SALMON *****
##
## Wilcoxon rank sum test with continuity correction
##
## data:  .x and .y
## W = 2252.5, p-value = 0.01996
## alternative hypothesis: true location shift is not equal to 0
##
##
## Two-sample Kolmogorov-Smirnov test
##
## data:  .x and .y
## D = 0.31455, p-value = 0.001098
## alternative hypothesis: two-sided
##
## Summary of GLM fit:
##
## Call:
## glm(formula = p.L12 ~ L + I(L^2) + I(L^3), family = binomial,
##      weights = wts)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -1.8298  -0.7191  -0.3292   0.5503   2.4104
##
## Coefficients:
##              Estimate      Std. Error z value Pr(>|z|)
## (Intercept) -3.5712824495   3.4261844962  -1.042   0.2972
## L             0.0432693656   0.0322471088   1.342   0.1797
## I(L^2)       -0.0001484696   0.0000864551  -1.717   0.0859 .
## I(L^3)        0.0000001400   0.0000000711   1.969   0.0490 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

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##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 69.991  on 56  degrees of freedom
## Residual deviance: 55.305  on 53  degrees of freedom
## AIC: 104.59
##
## Number of Fisher Scoring iterations: 4
##
## Analysis of Deviance Table
##
## Model: binomial, link: logit
##
## Response: p.L12
##
## Terms added sequentially (first to last)
##
##
##      Df Deviance Resid. Df Resid. Dev Pr(>Chi)
## NULL                                56      69.991
## L              1   7.0902         55      62.901 0.00775 **
## I(L^2)         1   3.6468         54      59.254 0.05618 .
## I(L^3)         1   3.9485         53      55.305 0.04691 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Summary of bootstrap fits:
##      mean          q.0.          q.5.          q.25.
## Min.   :0.1691   Min.   :0.006330   Min.   :0.05510   Min.   :0.1112
## 1st Qu.:0.2212   1st Qu.:0.008683   1st Qu.:0.07558   1st Qu.:0.1429
## Median :0.3855   Median :0.041698   Median :0.14638   Median :0.2606
## Mean   :0.3900   Mean   :0.076715   Mean   :0.19264   Mean   :0.2951
## 3rd Qu.:0.5527   3rd Qu.:0.097809   3rd Qu.:0.28581   3rd Qu.:0.4420
## Max.   :0.6654   Max.   :0.323010   Max.   :0.46416   Max.   :0.5552
##      q.50.          q.75.          q.95.          q.100.
## Min.   :0.1595   Min.   :0.2130   Min.   :0.3167   Min.   :0.5079
## 1st Qu.:0.2071   1st Qu.:0.2748   1st Qu.:0.4174   1st Qu.:0.7609
## Median :0.3670   Median :0.4989   Median :0.6793   Median :0.9190
## Mean   :0.3800   Mean   :0.4757   Mean   :0.6188   Mean   :0.8478
## 3rd Qu.:0.5520   3rd Qu.:0.6526   3rd Qu.:0.7893   3rd Qu.:0.9758
## Max.   :0.6806   Max.   :0.8359   Max.   :0.9777   Max.   :0.9998
##
## ***** MARKET SQUID *****
##
## Wilcoxon rank sum test with continuity correction
##
## data:  .x and .y
## W = 77344, p-value = 0.003271
## alternative hypothesis: true location shift is not equal to 0
##
##
## Two-sample Kolmogorov-Smirnov test
##
## data:  .x and .y

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## D = 0.12, p-value = 0.009663
## alternative hypothesis: two-sided
##
## Summary of GLM fit:
##
## Call:
## glm(formula = p.L12 ~ L + I(L^2) + I(L^3), family = binomial,
##      weights = wts)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -2.71791  -0.94032   0.00664   0.87627   1.82742
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept) -1.730989505  1.529021884  -1.132   0.2576
## L             0.118148391  0.067601893   1.748   0.0805 .
## I(L^2)       -0.001458737  0.000946183  -1.542   0.1231
## I(L^3)        0.000005212  0.000004133   1.261   0.2074
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 38.511  on 24  degrees of freedom
## Residual deviance: 33.295  on 21  degrees of freedom
## AIC: 97.49
##
## Number of Fisher Scoring iterations: 4
##
## Analysis of Deviance Table
##
## Model: binomial, link: logit
##
## Response: p.L12
##
## Terms added sequentially (first to last)
##
##
##      Df Deviance Resid. Df Resid. Dev Pr(>Chi)
## NULL                24      38.511
## L             1  0.45022        23      38.061  0.50223
## I(L^2)       1  3.13409        22      34.927  0.07667 .
## I(L^3)       1  1.63221        21      33.295  0.20140
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Summary of bootstrap fits:
##      mean          q.0.          q.5.          q.25.
## Min.      :0.2063   Min.      :0.00000   Min.      :0.0002774   Min.      :0.03008
## 1st Qu.:0.5043   1st Qu.:0.00000   1st Qu.:0.0324803   1st Qu.:0.31165
## Median :0.6015   Median :0.00336   Median :0.2094750   Median :0.42239
## Mean      :0.5880   Mean      :0.11489   Mean      :0.2675123   Mean      :0.45524
## 3rd Qu.:0.6971   3rd Qu.:0.23866   3rd Qu.:0.4927264   3rd Qu.:0.64031

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## Max. :0.7762 Max. :0.45413 Max. :0.6483986 Max. :0.73924
## q.50. q.75. q.95. q.100.
## Min. :0.1263 Min. :0.3312 Min. :0.6514 Min. :0.9234
## 1st Qu.:0.5176 1st Qu.:0.6460 1st Qu.:0.7908 1st Qu.:0.9344
## Median :0.6267 Median :0.7483 Median :0.8504 Median :0.9610
## Mean :0.6068 Mean :0.7368 Mean :0.8526 Mean :0.9630
## 3rd Qu.:0.7454 3rd Qu.:0.8223 3rd Qu.:0.8962 3rd Qu.:0.9947
## Max. :0.8513 Max. :0.9979 Max. :1.0000 Max. :1.0000
##
## ***** PACIFIC HERRING *****
##
## Insufficient data
##
## ***** WATER JELLY *****
##
## Wilcoxon rank sum test with continuity correction
##
## data: .x and .y
## W = 79448, p-value = 0.8924
## alternative hypothesis: true location shift is not equal to 0
##
##
## Two-sample Kolmogorov-Smirnov test
##
## data: .x and .y
## D = 0.025505, p-value = 0.9995
## alternative hypothesis: two-sided
##
## Summary of GLM fit:
##
## Call:
## glm(formula = p.L12 ~ L + I(L^2) + I(L^3), family = binomial,
## weights = wts)
##
## Deviance Residuals:
## Min 1Q Median 3Q Max
## -3.5114 -0.5073 0.2832 0.9065 2.5335
##
## Coefficients:
## Estimate Std. Error z value Pr(>|z|)
## (Intercept) -7.018755017 3.943695235 -1.780 0.0751 .
## L 0.272045342 0.159724393 1.703 0.0885 .
## I(L^2) -0.003280210 0.002107257 -1.557 0.1196
## I(L^3) 0.000012696 0.000009052 1.403 0.1607
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
## Null deviance: 44.978 on 19 degrees of freedom
## Residual deviance: 40.236 on 16 degrees of freedom
## AIC: 110.61
##
## Number of Fisher Scoring iterations: 3

```



```

##
## Analysis of Deviance Table
##
## Model: binomial, link: logit
##
## Response: p.L12
##
## Terms added sequentially (first to last)
##
##
##          Df Deviance Resid. Df Resid. Dev Pr(>Chi)
## NULL                19      44.978
## L             1  0.22654      18      44.751  0.6341
## I(L^2)        1  2.39043      17      42.361  0.1221
## I(L^3)        1  2.12431      16      40.236  0.1450
##
## Summary of bootstrap fits:
##      mean      q.0.      q.5.
## Min.   :0.2142   Min.   :0.0000000   Min.   :0.0000541
## 1st Qu.:0.4971   1st Qu.:0.0001554   1st Qu.:0.0405347
## Median :0.5431   Median :0.0373356   Median :0.2345563
## Mean   :0.5062   Mean   :0.1080724   Mean   :0.2115019
## 3rd Qu.:0.5930   3rd Qu.:0.2291246   3rd Qu.:0.3614782
## Max.   :0.6254   Max.   :0.2953500   Max.   :0.4242308
##      q.25.      q.50.      q.75.      q.95.
## Min.   :0.01017   Min.   :0.08892   Min.   :0.3148   Min.   :0.6877
## 1st Qu.:0.33944   1st Qu.:0.49707   1st Qu.:0.6011   1st Qu.:0.7593
## Median :0.39869   Median :0.54857   Median :0.6552   Median :0.7888
## Mean   :0.37196   Mean   :0.51063   Mean   :0.6367   Mean   :0.8024
## 3rd Qu.:0.48714   3rd Qu.:0.60762   3rd Qu.:0.6797   3rd Qu.:0.8213
## Max.   :0.54969   Max.   :0.76425   Max.   :0.9400   Max.   :0.9961
##      q.100.
## Min.   :0.8376
## 1st Qu.:0.8947
## Median :0.9267
## Mean   :0.9297
## 3rd Qu.:0.9880
## Max.   :1.0000

```