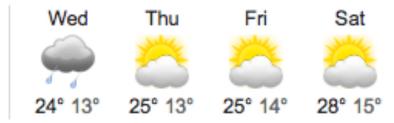
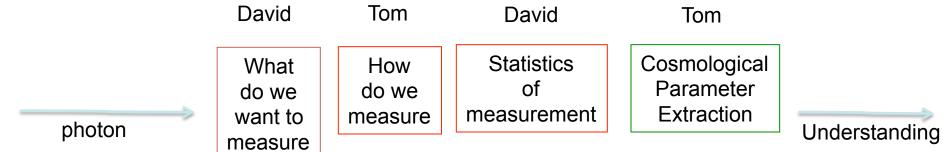
#### Weather for Cargèse





### Statistics and Prediction

Tom Kitching tdk@roe.ac.uk @tom\_kitching



# Recap

- Lensing equation
- Local mapping

$$\beta = \theta - \hat{\alpha} \frac{D_{ds}}{D_s}$$
$$\beta = \theta - \alpha.$$

General Relativity relates this to the gravitational potential

- Distortion matrix implies that distortion is elliptical: shear and convergence
- Simple formalise that relates the shear and convergence (observable) to the underlying gravitational potential

### Recap

- Observed galaxies have instrinsic ellipticity and shear
- Reviewed shape measurement methods
  - Moments KSB
  - Model fitting lensfit
- Still an unsolved problem for largest most ambitious surveys
- Simulations
  - STEP 1, 2
  - GREAT08, GREAT10

# Recap

- Mass mapping
- Correlation functions
- Power Spectra

What will we learn?

- Predictive parameter and model estimation
  - Fisher Matrix
  - Evidence

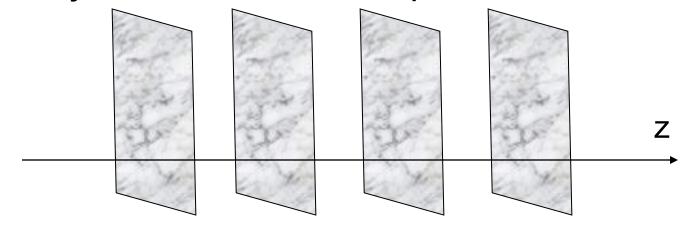
- Real Life parameter and model selection
  - MCMC

$$C_{ij}(\ell) = \int_0^{r_H} \mathrm{d}r \, W_{ij}^{\mathrm{GG}}(r) P_{\delta\delta}igg(rac{\ell}{S_k(r)}; rigg)$$

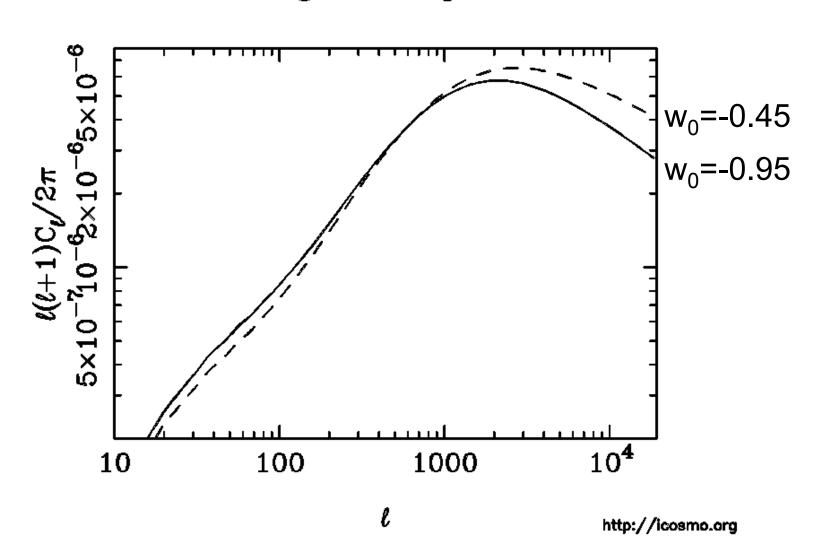
$$W_{ij}^{ ext{GG}}(r) = rac{q_i(r)q_j(r)}{S_k^2(r)}$$

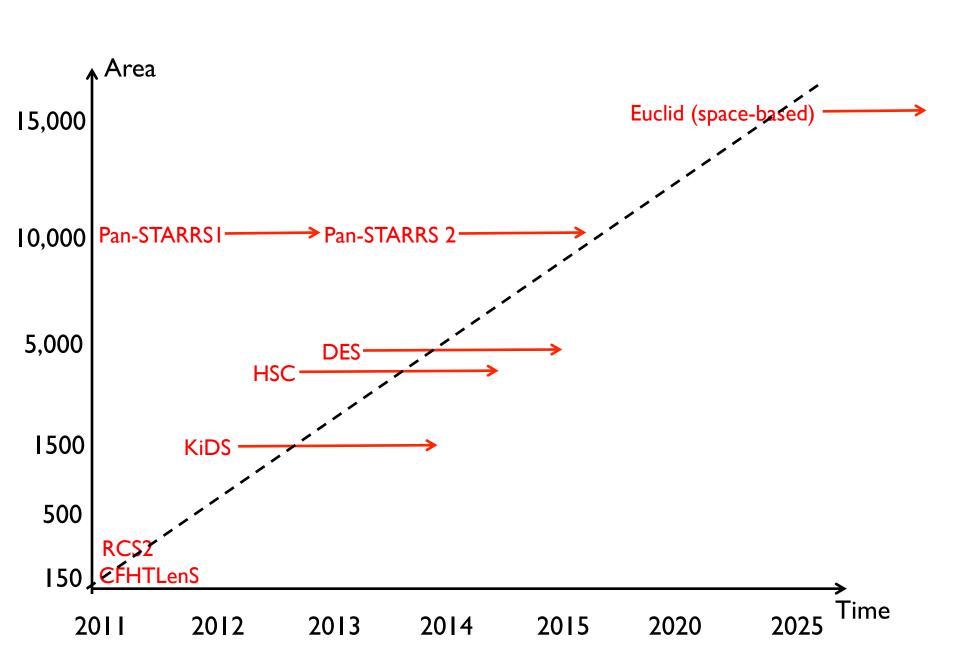
$$W_{ij}^{ ext{GG}}(r) = rac{q_i(r)q_j(r)}{S_k^2(r)} \hspace{0.5cm} q_i(r) = rac{3H_0^2\Omega_m S_k(r)}{2a(r)} \int_r^{r_H} dr' \, p_i(r') rac{S_k(r'-r)}{S_k(r')}$$

- Tomography
  - Generate 2D shear correlation in redshift bins
  - Can "auto" correlate in a bin
  - Or "cross" correlate between bin pairs
  - i and j refer to redshift bin pairs



### Lensing Power Spectrum



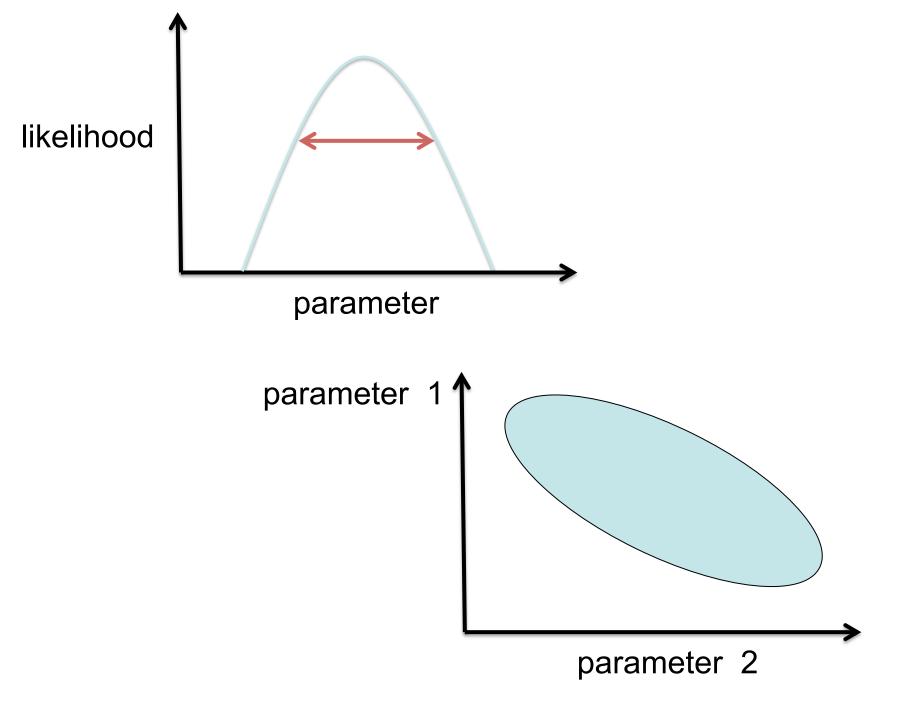


### What do we want?

 How accurately can we estimate a model parameter from a given data set?

- Given a set of N data point x<sub>1</sub>,...,x<sub>N</sub>
  - Want the estimator to be *unbiased*  $\langle \theta \rangle = \theta_0$
  - Give small error bars as possible  $\Delta \theta_{\alpha} \equiv (\langle \theta_{\alpha}^2 \rangle \langle \theta_{\alpha} \rangle^2)^{1/2}$
- The Best Unbiased Estimator
- A key Quantity in this is the Fisher (Information) Matrix

- Fisher 1935
- Tegmark, Taylor, Heavens 1997



# What is the (Fisher) Matrix?

Lets expand a likelihood surface about the maximum likelihood point

$$\ln L(\mathbf{x}; \boldsymbol{\theta}) = \ln L(\mathbf{x}; \boldsymbol{\theta}_0) + \frac{1}{2} (\boldsymbol{\theta}_{\alpha} - \boldsymbol{\theta}_{0\alpha}) \frac{\partial^2 \ln L}{\partial \boldsymbol{\theta}_{\alpha} \partial \boldsymbol{\theta}_{\beta}} (\boldsymbol{\theta}_{\beta} - \boldsymbol{\theta}_{0\beta}) + \dots$$

Can write this as a Gaussian

$$L(\mathbf{x}; \boldsymbol{\theta}) = L(\mathbf{x}; \boldsymbol{\theta}_0) \exp \left[ -\frac{1}{2} (\boldsymbol{\theta}_{\alpha} - \boldsymbol{\theta}_{0\alpha}) H_{\alpha\beta} (\boldsymbol{\theta}_{\beta} - \boldsymbol{\theta}_{0\beta}) \right]$$

Where the Hessian (covariance) is

$$\mathsf{H}_{\alpha\beta} \equiv -\frac{\partial^2 \ln L}{\partial \theta_{\alpha} \partial \theta_{\beta}}$$

### What is the Fisher Matrix?

The Hessian Matrix has some nice properties

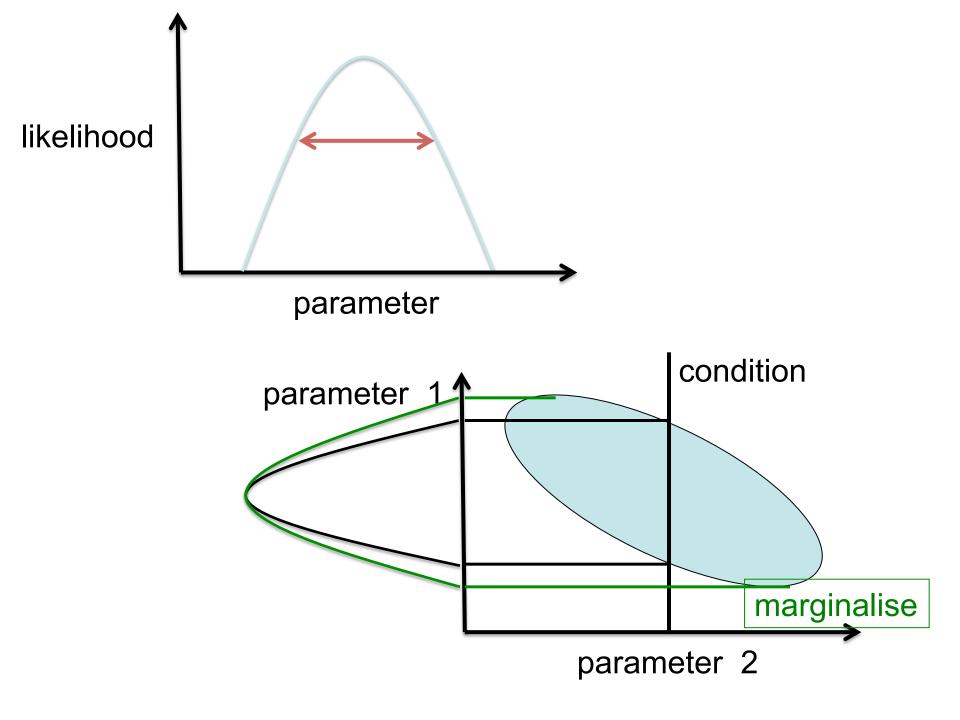
Conditional Error on α

$$\sigma_{\text{conditional},\alpha} = \frac{1}{\sqrt{\mathsf{H}_{\alpha\alpha}}}$$

• Marginal error on  $\alpha$ 

$$\sigma_{\alpha} = \sqrt{(\mathsf{H}^{-1})_{\alpha\alpha}}$$

Matrix inversion performed



### What is the Fisher Matrix?

 The Fisher Matrix defined as the expectation of the Hessian matrix

$$\mathsf{F}_{\alpha\beta} \equiv \left\langle \mathsf{H}_{\alpha\beta} \right\rangle = \left\langle -\frac{\partial^2 \ln L}{\partial \theta_\alpha \partial \theta_\beta} \right\rangle$$

- This allows us to make <u>predictions</u> about the performance of an experiment!
- The <u>expected</u> conditional error on  $\alpha$ :  $\sigma_{\alpha} = \sqrt{1/F_{\alpha\alpha}}$
- The expected marginal error on  $\alpha$

$$\sigma_{\alpha} = \sqrt{(\mathsf{F}^{-1})_{\alpha\alpha}}.$$
 Matrix inversion performed

### The Gaussian Case

How do we calculate Fisher Matrices in practice?

Assume that the likelihood is Gaussian

$$2\mathcal{L} = \ln \det \mathsf{C} + (\mathbf{x} - \boldsymbol{\mu})\mathsf{C}^{-1}(\mathbf{x} - \boldsymbol{\mu})^T$$

### The Gaussian Case

$$D \equiv (\mathbf{x} - \boldsymbol{\mu})(\mathbf{x} - \boldsymbol{\mu})^T$$
 
$$2\mathcal{L} = \operatorname{Tr} \left[ \ln \mathsf{C} + \mathsf{C}^{-1} \mathsf{D} \right]$$
 
$$\operatorname{derivative}$$
 
$$\operatorname{derivative}$$
 
$$2\mathcal{L},_{\alpha} = \operatorname{Tr} \left[ \mathsf{C}^{-1} \mathsf{C},_{\alpha} - \mathsf{C}^{-1} \mathsf{C},_{\alpha} \mathsf{C}^{-1} \mathsf{D} + \mathsf{C}^{-1} \mathsf{D},_{\alpha} \right]$$
 
$$\operatorname{derivative}$$
 
$$\left[ \mathcal{L},_{\alpha} \right] = 0.$$

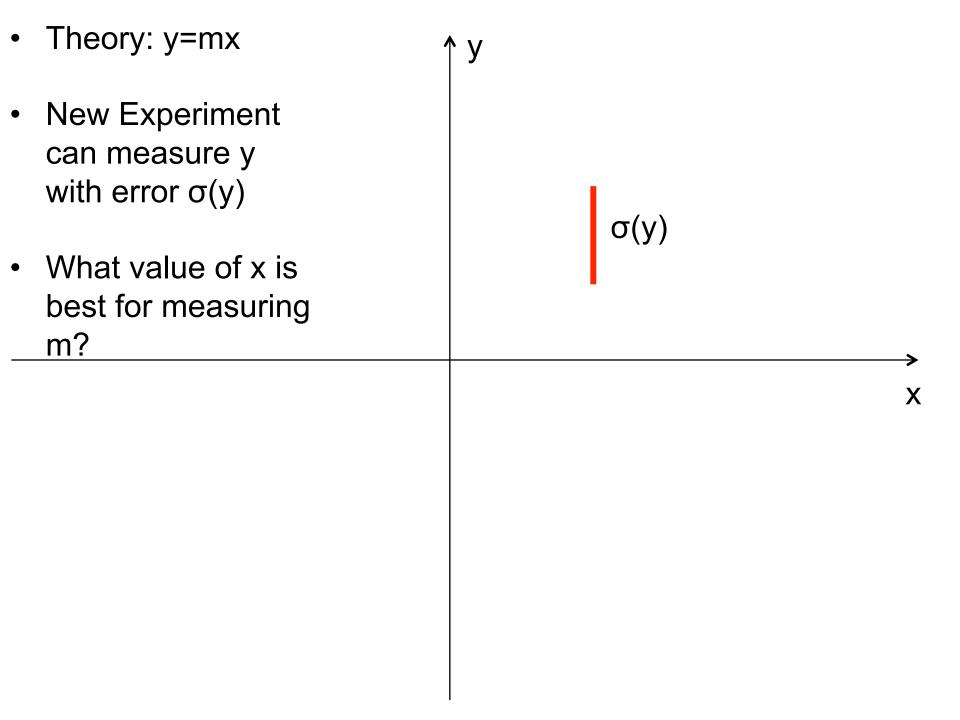
$$\mathsf{F}_{\alpha\beta} = \langle \mathcal{L},_{\alpha\beta} \rangle = \frac{1}{2} \mathrm{Tr} [\mathsf{C}^{-1} \mathsf{C},_{\alpha} \mathsf{C}^{-1} \mathsf{C},_{\beta} + \mathsf{C}^{-1} [ \boldsymbol{\mu},_{\alpha} \boldsymbol{\mu},_{\beta}^T + \boldsymbol{\mu},_{\beta} \boldsymbol{\mu},_{\alpha}^T ] ].$$

### How to Calculate a Fisher Matrix

$$\mathsf{F}_{\alpha\beta} = \left\langle \mathcal{L},_{\alpha\beta} \right. \right\rangle = \frac{1}{2} \mathrm{Tr} [\mathsf{C}^{-1} \mathsf{C},_{\alpha} \mathsf{C}^{-1} \mathsf{C},_{\beta} + \mathsf{C}^{-1} \big[ \, \boldsymbol{\mu},_{\alpha} \, \boldsymbol{\mu},_{\beta}^T + \boldsymbol{\mu},_{\beta} \, \boldsymbol{\mu},_{\alpha}^T \, \, \big] \, \big].$$

- We know the (expected) covariance and mean from theory
- Requires NO DATA!

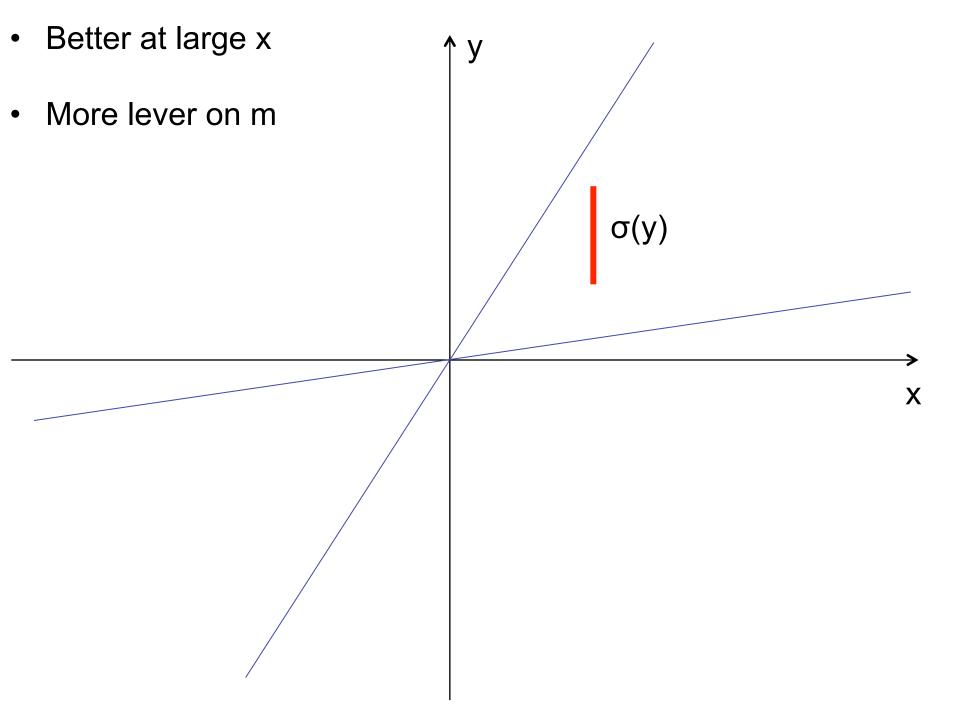
Worked example y=mx+c



$$\mathsf{F}_{\alpha\beta} = \left\langle \mathcal{L},_{\alpha\beta} \right. \right\rangle = \frac{1}{2} \mathrm{Tr} [\mathsf{C}^{-1} \mathsf{C},_{\alpha} \mathsf{C}^{-1} \mathsf{C},_{\beta} + \mathsf{C}^{-1} \big[ \, \boldsymbol{\mu},_{\alpha} \, \boldsymbol{\mu},_{\beta}^T + \boldsymbol{\mu},_{\beta} \, \boldsymbol{\mu},_{\alpha}^T \, \, \big] \, \big].$$

- The theory: y=mx
- Experiment can measure σ(y)
- Question: what x value is best?

- Covariance does not depend on m so first term zero
- dy/dm=x
- $F_{mm} = (1/\sigma^2(y))x^2$
- $\sigma(m) = \sqrt{[F^{-1}_{mm}]} = \sqrt{[\sigma^2(y) \ x^{-2}]} = \sigma(y) \ x^{-1}$
- Better measure of m at large x



Some nomenclature

• 
$$\sigma(m) = \sqrt{[F^{-1}_{mm}]} = \sqrt{[\sigma^2(y) \ x^{-2}]} = \sigma(y) \ x^{-1}$$

- Need a "fiducial" value to make quantitative predictions
- If derivative is analytic then fairly straightforward (not always the case!)

### Some nice properties of Fisher matrices

• 
$$F = \begin{pmatrix} A & B \\ C & D \end{pmatrix}$$

- Matrix Manipulation:
  - Inversion F<sup>-1</sup>
  - Addition F=G+H
  - Rotation G=RFR<sup>T</sup>
  - Schur Complement F<sub>A</sub>=A-BD<sup>-1</sup>C

# Adding Extra Parameters

To add parameters to a Fisher Matrix

Simply extend the matrix

$$F = \begin{pmatrix} F^{\theta\theta} & F^{\theta w(\phi)} \\ F^{w(\phi)\theta} & F^{w(\phi)w(\phi)} \end{pmatrix}$$

# Combining Experiments

If two experiments are independent then the combined error is simply

$$F_{comb} = F_1 + F_2$$

Same for n experiments

 If not independent need to have a single Fisher matrix with a joint covariance

# Re-Parameterising

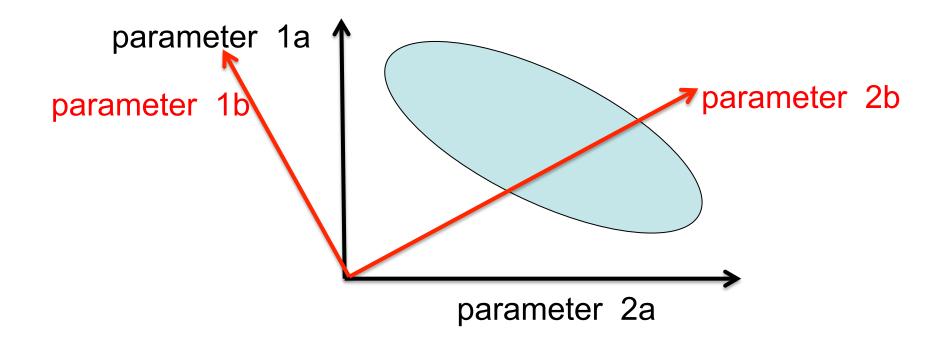
 Can change variables between two parameter sets through a Jacobian transform (rotation)

F(new)=J
$$^{\mathsf{T}}$$
F(old)J  $J_{ij}=rac{\partial b_j}{\partial a_i}$ 

- Where J is a matrix of derivatives
  - NOTE can only do this if the basis sets are mutually complete (Kitching & Amara, 2009)
- Eigendecomposition is a special case of parameter rotation

$$J_{ij}=rac{\partial b_j}{\partial a_i}$$

$$F(new)=J^TF(old)J$$



# Schur Complement

$$F = \begin{pmatrix} A & B \\ C & D \end{pmatrix}$$

Schur Complement F<sub>A</sub>=A-BD<sup>-1</sup>C

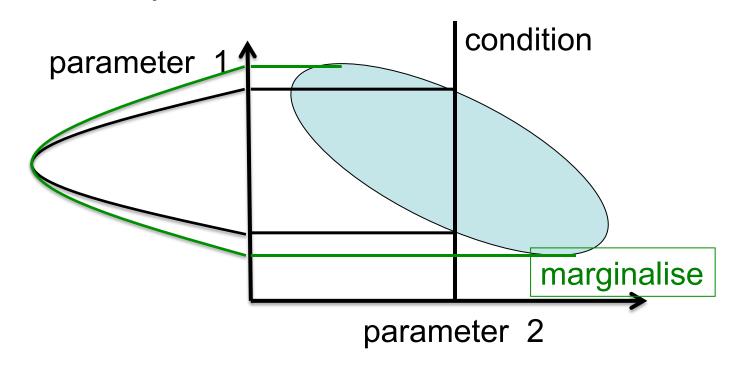
- This is equivalent of the following operation
  - Invert entire matrix F<sup>-1</sup>
  - Select the A-part of the inverse F<sup>-1</sup>(A)
  - Reinvert
  - What you have is a new "sub" Fisher matrix

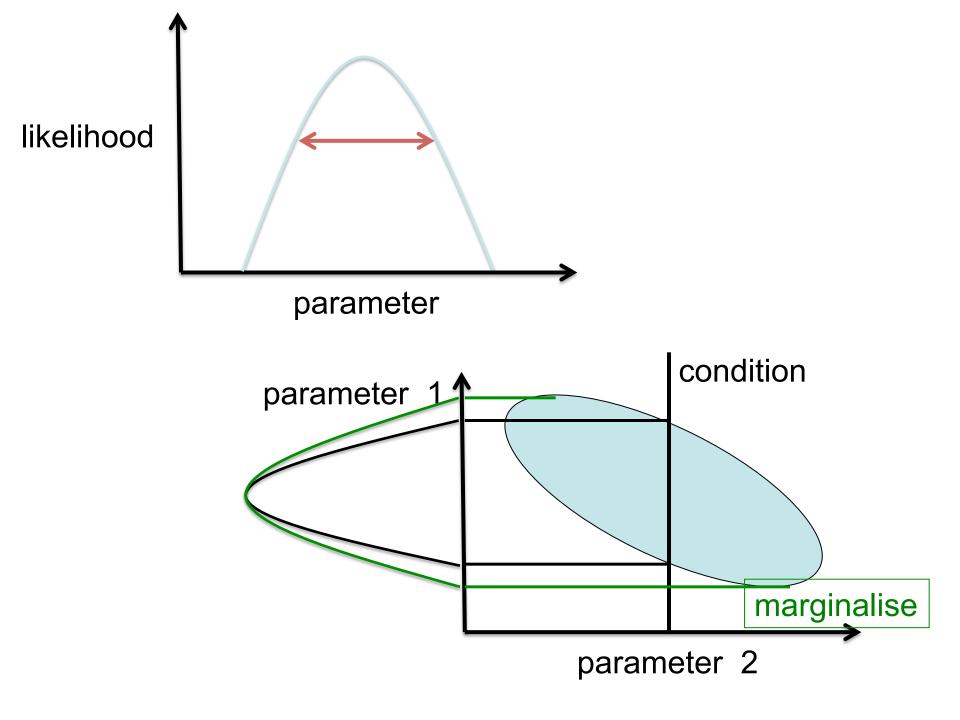
What is does this correspond to ?

Schur Complement is equivalent of marginalisation over parameters

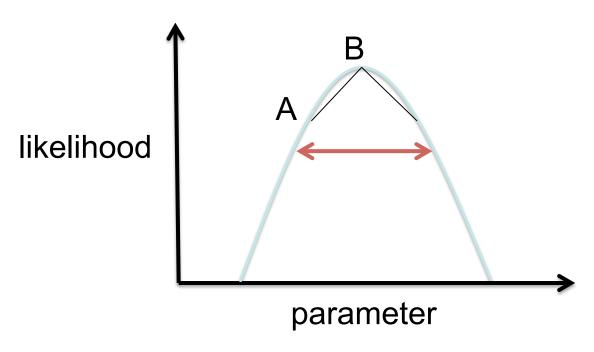
$$F = \begin{pmatrix} A & B \\ C & D \end{pmatrix}$$

- Schur Complement
  - $F_A = A BD^{-1}C < A$
- Marginalises over parameters not-A





### A Warning about derivatives



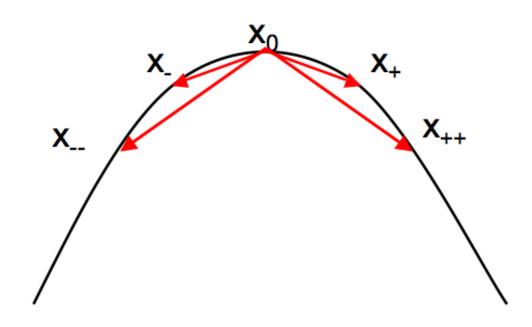
dL/dparameter= (B-A)/dp

Or can approximate using q parabola

Numerically must test if this is stable

Quadractic approximation Exercise Prove This

 $df(x)/dx = [f(x_{+})-f(x_{-}) - \{f(x_{++})-f(x_{--})-(2 f(x_{+})+2 f(x_{-}))/6\}]/[x_{+}-x_{-}]$ 

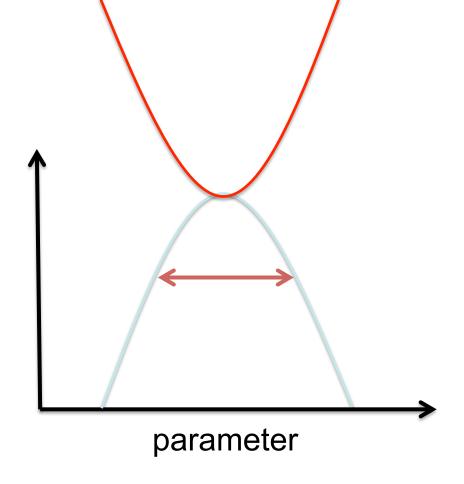


# Warning about the two terms

$$\mathsf{F}_{\alpha\beta} = \left< \mathcal{L},_{\alpha\beta} \right. \right> = \frac{1}{2} \mathrm{Tr} [\mathsf{C}^{-1} \mathsf{C},_{\alpha} \mathsf{C}^{-1} \mathsf{C},_{\beta} + \mathsf{C}^{-1} \big[ \, \boldsymbol{\mu},_{\alpha} \, \boldsymbol{\mu},_{\beta}^T + \boldsymbol{\mu},_{\beta} \, \boldsymbol{\mu},_{\alpha}^T \, \big] \, \big].$$

- Can we use both terms and get twice the information?
- No not usually (almost never in cosmology)

- Fisher matrices must be positive definite
- A positive definite matrix has
  - ONLY POSITIVE EIGENVALUES
  - Because by definition the likelihood surface is assumed to be single peaked



- Can get non-positive definite due to numerical inaccuracies
  - Corresponds to a convex(negatively) curved surface
- Should always check matrices

# Fisher Future Forecasting

 We now have a tool with which we can predict the accuracy of future experiments!

$$\mathsf{F}_{\alpha\beta} = \langle \mathcal{L},_{\alpha\beta} \, \rangle = \frac{1}{2} \mathrm{Tr} [\mathsf{C}^{-1} \mathsf{C},_{\alpha} \mathsf{C}^{-1} \mathsf{C},_{\beta} + \mathsf{C}^{-1} \big[ \, \boldsymbol{\mu},_{\alpha} \, \boldsymbol{\mu},_{\beta}^T + \boldsymbol{\mu},_{\beta} \, \boldsymbol{\mu},_{\alpha}^T \, \big] \, \big].$$

- Can easily
  - Calculate expected parameter errors
  - Combine experiments
  - Change variables
  - Add extra parameters

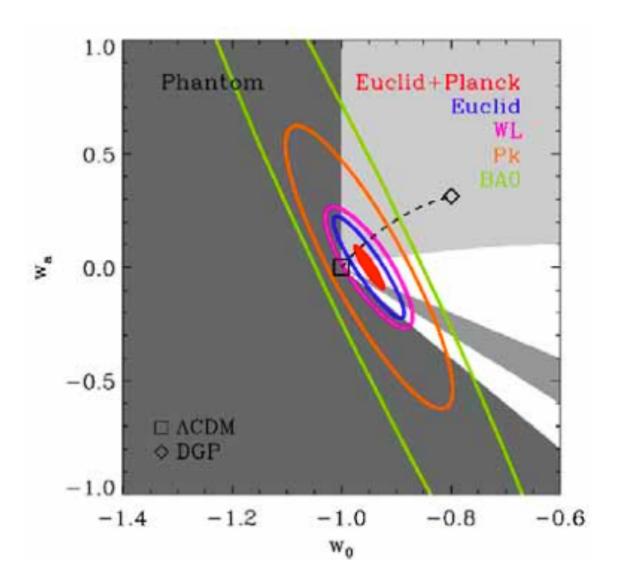
 For shear the mean shear is zero, the information is in the covariance so (Hu, 1999)

$$\mathbf{F}_{lphaeta} = \sum_{\ell=2}^{\ell_{ ext{max}}} (\ell+1/2) f_{ ext{sky}} ext{tr}[\mathbf{C}^{-1}\mathbf{C}_{,lpha}\mathbf{C}^{-1}\mathbf{C}_{,eta}]$$

 This is what is used to make predictions for cosmic shear and dark energy experiments

# Dark Energy

Expect constraints of 1% from Euclid



$$\mathsf{F}_{\alpha\beta} = \left< \mathcal{L},_{\alpha\beta} \right. \right> = \frac{1}{2} \mathrm{Tr} [\mathsf{C}^{-1} \mathsf{C},_{\alpha} \mathsf{C}^{-1} \mathsf{C},_{\beta} + \mathsf{C}^{-1} \big[ \, \boldsymbol{\mu},_{\alpha} \, \boldsymbol{\mu},_{\beta}^T + \boldsymbol{\mu},_{\beta} \, \boldsymbol{\mu},_{\alpha}^T \, \, \big] \, \big].$$

- The theory: y=mx
- Experiment can measure σ(y)
- Question: what x value is best?

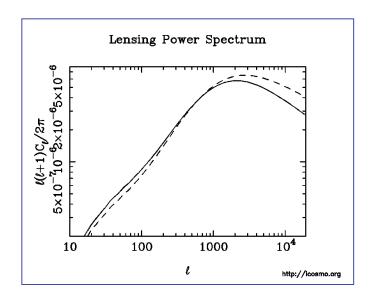
dy/dm=x

- $F_{mm} = (1/\sigma^2(y))x^2$
- $\sigma(m) = \sqrt{[F^{-1}_{mm}]} = \sqrt{[\sigma^2(y) \ x^{-2}]} = \sigma(y) \ x^{-1}$
- Better measure of m at large x

$$\mathsf{F}_{\alpha\beta} = \left\langle \mathcal{L},_{\alpha\beta} \right. \right\rangle = \frac{1}{2} \mathrm{Tr} [\mathsf{C}^{-1} \mathsf{C},_{\alpha} \mathsf{C}^{-1} \mathsf{C},_{\beta} + \mathsf{C}^{-1} \big[ \, \boldsymbol{\mu},_{\alpha} \, \boldsymbol{\mu},_{\beta}^T + \boldsymbol{\mu},_{\beta} \, \boldsymbol{\mu},_{\alpha}^T \, \big] \, \big].$$

- The theory:  $w(z)=w_0+w_a(1-a)$
- Experiment can measure σ(C<sub>I</sub>)
- Question what redshift/area?

$$C_{ij}(\ell) = \int_0^{r_H} \mathrm{d}r \, W_{ij}^{\mathrm{GG}}(r) P_{\delta\delta} igg(rac{\ell}{S_k(r)}; rigg)$$



Know from Theory/Simulations

$$\mathsf{F}_{\alpha\beta} = \langle \mathcal{L},_{\alpha\beta} \rangle = \frac{1}{2} \mathrm{Tr} [\mathsf{C}^{-1} \mathsf{C},_{\alpha} \mathsf{C}^{-1} \mathsf{C},_{\beta} + \mathsf{C}^{-1} [\underline{\mu},_{\alpha} \underline{\mu},_{\beta}^T + \underline{\mu},_{\beta} \underline{\mu},_{\alpha}^T]].$$
 Mean is shear is zero

## Hu 1999

#### POWER SPECTRUM TOMOGRAPHY WITH WEAK LENSING

WAYNE HU

Institute for Advanced Study, Princeton, NJ 08540 Submitted June 16, 2012

Generalizing the results of Hu & Tegmark (1998) to multiple correlated power spectra, we obtain<sup>1</sup>

$$\mathbf{F}_{\alpha\beta} = \sum_{\ell=2}^{\ell_{\text{max}}} (\ell + 1/2) f_{\text{sky}} \text{tr}[\mathbf{C}^{-1} \mathbf{C}_{,\alpha} \mathbf{C}^{-1} \mathbf{C}_{,\beta}], \qquad (8)$$

under the assumption of Gaussian signal and noise, where  $f_{\text{sky}}$  is fraction of sky covered by the survey, the covari-

I+1/2 sums over m-modes f<sub>sky</sub> scales the covariance with the survey area

Note also need to include noise on covariance C=C(I)+N $N=\sigma(e)/n_{galaxy}$  Question we have address is:

 Given an experiment how accurate can I measure parameter values?

Alternative/additional question is

 How accurately can I determine a model (set of parameters) • Bayes' Theorem  $p_{i}(e|\boldsymbol{y}_{i}) = \frac{\mathcal{P}\left(\boldsymbol{e}\right)\mathcal{L}\left(\boldsymbol{y}_{i}|\boldsymbol{e}\right)}{\int \mathcal{P}\left(\boldsymbol{e}\right)\mathcal{L}\left(\boldsymbol{y}_{i}|\boldsymbol{e}\right)d\boldsymbol{e}}$  Evidence

#### **Posterior**

- Measure likelihood of data given parameters
- Assume prior on parameters
- Evidence for a Model (set of parameters)

#### How to compute expected evidence?

• Evidence  $p(D|M) = \int d\theta \, p(D|\theta, M) p(\theta|M)$ 

Bayes Factor=Ratio of Evidences

$$\frac{p(M'|D)}{p(M|D)} = \frac{p(M')}{p(M)} \frac{\int d\theta' \ p(D|\theta', M') p(\theta'|M')}{\int d\theta \ p(D|\theta, M) p(\theta|M)}.$$

$$B \equiv \frac{\int d\theta' \, p(D|\theta', M') p(\theta'|M')}{\int d\theta \, p(D|\theta, M) p(\theta|M)}$$

What does the Bayes factor mean?

$$B \equiv \frac{\int d\theta' \, p(D|\theta', M') p(\theta'|M')}{\int d\theta \, p(D|\theta, M) p(\theta|M)}$$

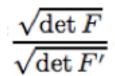
- Odds: how much would you gamble?
- Jeffereys Scale (take with a pinch of salt)
  - LnB < 1 "inconclusive"</li>
  - 1< LnB < 2.5 "significant" odds ~1:12</li>
  - 2.5 < LnB < 5.0 "strong" odds ~1:150</li>
  - LnB > 5.0 "decisive" odds better than 1:150

- Can assume Gaussian likelihoods and perform the integration (Heavens, Kitching, Verde, 2008)
- Can compute from the Fisher matrix the expected evidence for nested models

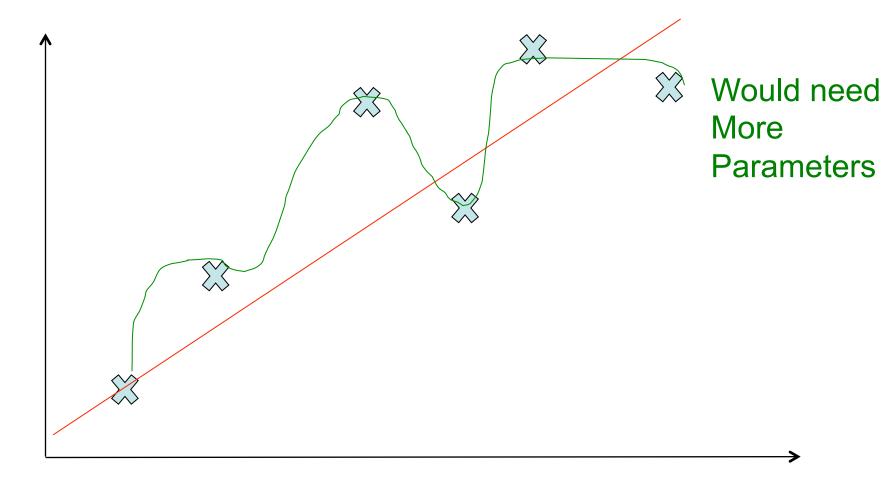
$$\langle B \rangle = (2\pi)^{-p/2} \frac{\sqrt{\det F}}{\sqrt{\det F'}} \exp\left(-\frac{1}{2}\delta\theta_{\alpha}F_{\alpha\beta}\delta\theta_{\beta}\right) \prod_{q=1}^{p} \Delta\theta_{n'+q'}$$

- Other similar approaches in Trotta (2008)
- "Occam Factor" can be seen

## Occam Factor



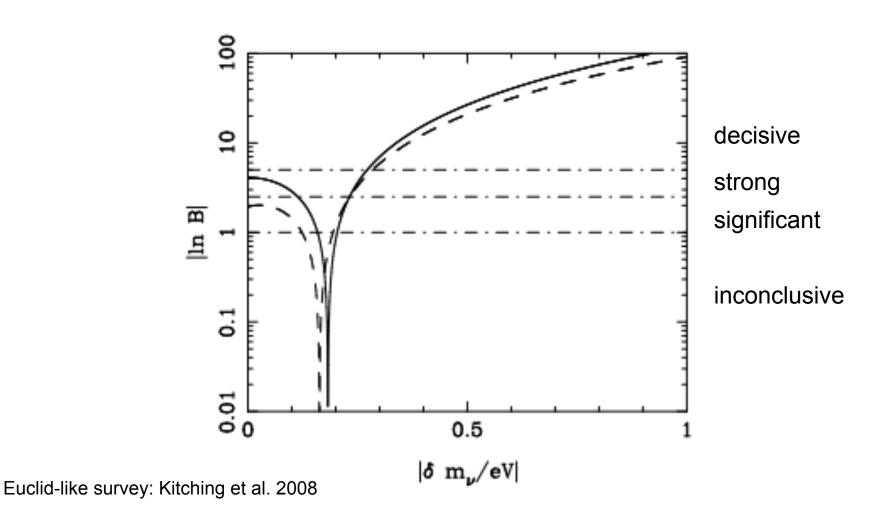
- Occam's Razor: Simpler models are prefered
- Stops you over fitting your data



Example from neutrino mass from weak lensing

- Neutrinos have mass = model A
- Neutrinos do not have mass = model B

# Example from neutrino mass from weak lensing

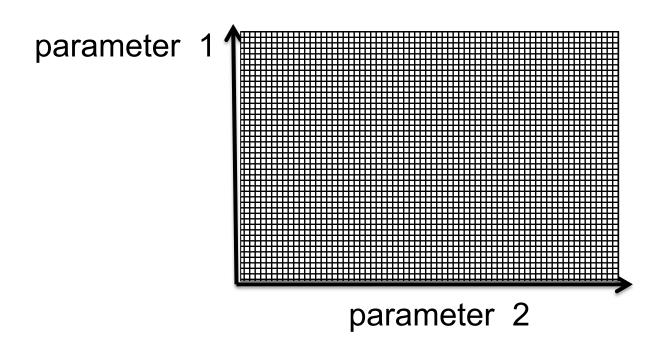


- What will we learn?
  - Fisher matrices
  - Likelihood sampling

# Likelihood Sampling

- We have at least 10 cosmological parameters
- Others may be
  - non-zero
  - functions of scale and/ or redshift
  - w(z)+1
  - b(z,k)-1

Parameter	Symbol
Hubble parameter	h
Total matter density	$\Omega_{ m m}$
Baryon density	$\Omega_{ m b}$
Cosmological constant	$\Omega_{\Lambda}$
Radiation density	$\Omega_{ m r}$
Neutrino density	$\Omega_{ u}$
Density perturbation amplitude	$\Delta_R^2(k_*)$
Density perturbation spectral index	n
Tensor to scalar ratio	r
Ionization optical depth	au



Grid. Evaluate likelihood function at a grid of points in parameter space

What is wrong with this approach? ND

#### MCMC Methods

Monte Carlo Markov Chain

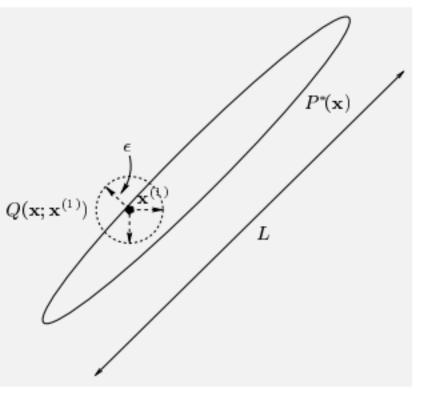
Randomly Sample Likelihood Space

A chain of likelihood (or other) evaluations that is "memoryless"

## Metropolis-Hastings

Sample likelihood space with a random walk

- Pick a point in parameter space x<sub>i</sub>
- Evaluate likelihood L(x<sub>i</sub>)
- 3) Pick a new random point x<sub>i+1</sub> from a *proposal distribution*
- 4) Evaluate likelihood  $L(x_{i+1})$
- 5) If  $a=L(x_{i+1})/L(x_i)\geq 1$  ACCEPT (and goto 3)
- 6) Accept with probability a
  - Draw a uniform random number b and if b<a ACCEPT else reject</li>



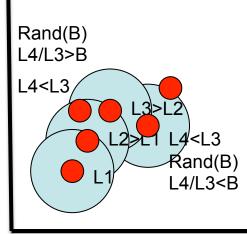
What choice of proposal?

Common to choose multivariate Gaussian (inefficient for degenerate parameters)

Could also choose the Fisher matrix

- Pick a point in parameter space x<sub>i</sub>
- Evaluate likelihood L(x<sub>i</sub>)
- 3) Pick a new random point x<sub>i+1</sub> from a proposal distribution
- 4) Evaluate likelihoodL(x<sub>i+1</sub>)
- 5) If  $a=L(x_{i+1})/L(x_i)\geq 1$ ACCEPT (and goto 3)
- 6) Accept with probability a
  - Draw a uniform random number b and if b<a ACCEPT else reject

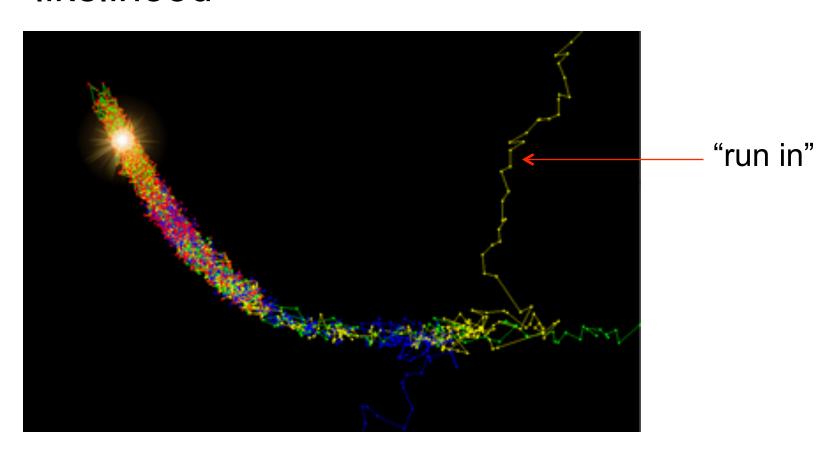
parameter 1



parameter

parameter 2

 The density of points is proportional to the likelihood



# Convergence

- When to stop (never!)
- But chain will "converge"
- Gelman-Rubin
  - "variance between chains consistent with variance within a chain"
  - Run multiple chains; m group of n chains n chains

$$\hat{\sigma}^2 = \frac{n-1}{n} W + \frac{1}{n} B$$

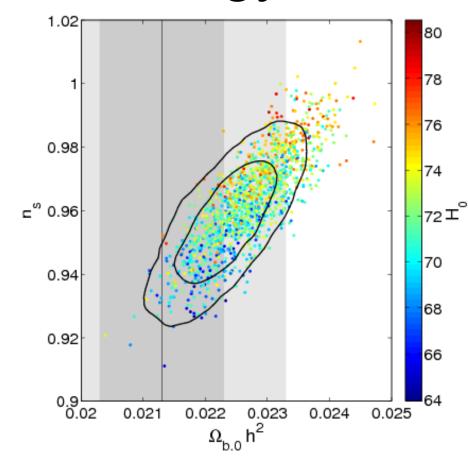
- W=mean variance over chains
- B=variance between chains
- σ unbiased estimate of target
- m groups of n chains

$$\sqrt{\hat{V}} = \sqrt{\hat{\sigma}^2 + B/mn}$$

V=1 means convergence

## MCMC for Cosmology

- Most commonly used package is
  - cosmomc
     http://cosmologist.info/ cosmomc/
  - Coupled with CAMB
  - Used for WMAP analysis (and many others)
  - Can download WMAP MCMC chains to play around with
  - http://lambda.gsfc.nasa.gov/ toolbox/



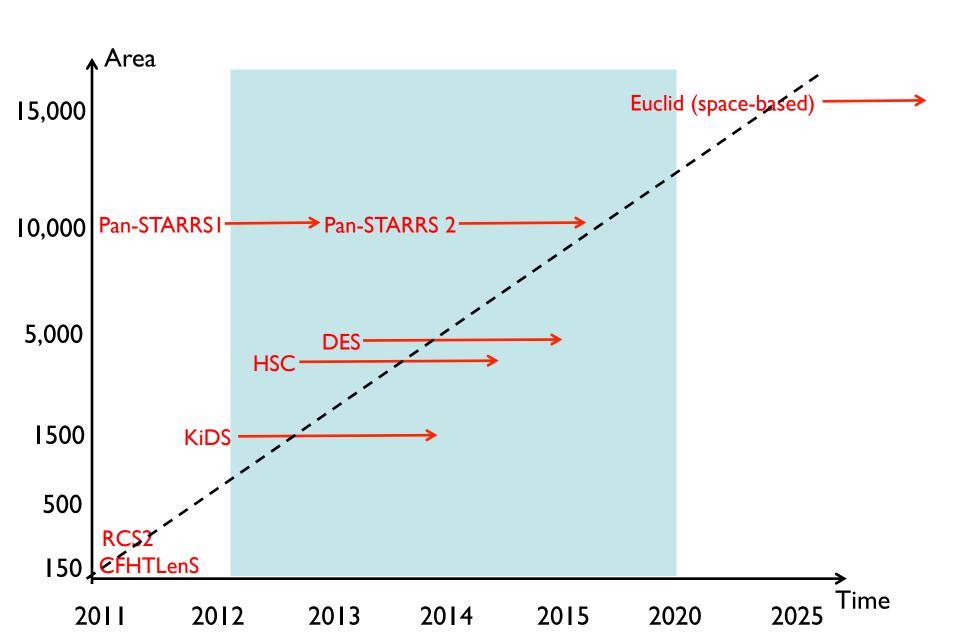
#### Other MCMC methods

- What are the problems with MCMC?
  - Multiple peaks
  - Evidence?
- Have some different types/flavours
  - Gibbs sampling
  - Simulated annealing
- What alternatives?
  - Nested Sampling
    - Feroz & Hobson (2008), refined by Feroz, Hobson & Bridges (2008)
  - Population Monte Carlo
    - Kilbinger et al. 2009

## Recap

- Predictive parameter and model estimation
  - Fisher Matrix
  - Evidence

- Real Life parameter and model selection
  - MCMC



#### Conclusion

- Lensing is a simple cosmological probe
  - Directly related to General Relativity
  - Simple linear image distortions

- Measurement from data is challenging
  - Need lots of galaxies and very sophisticated experiments

 Lensing is a powerful probe of dark energy and dark matter