# The Effect of Expected Income on Individual Migration Decisions

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## Previous models of migration choice

- Blanchard & Katz (1992) Income differentials play a large role in inter-state migration decisions
- Gallin (2004) Migration as a function of expected wages, but individual decision problem un-modeled
- Holt (2009) Dynamic discrete choice of migration, but move / stay not location-specific
- Dahl (2002) Many alternatives but only a single life-time migration choice

#### Motivation

#### Interstate Migration, NLSY 1979-1994

Movers (%)	32.3
Moves per mover	1.98
Repeat moves (% of moves)	49.5
Movers returning home (%)	50.2

- Migration decisions aren't one-off or isolated
- Need a dynamic model to study migration decisions

## Overview of the paper

- Tractable dynamic search model of migration decisions based on wage differentials
- Estimated using NLSY79

- Expected income change, distance, home / previous location, population size, age, climate
- Empirical estimate of moving costs (\$312k+) and home premium (\$23k)

#### Model – environment

- Finitely-lived individuals i
- ullet Get linear utility from wages and amenities in one of J locations
- Pay a moving cost to try a new location
- Make migration decisions to maximize lifetime utility
- Wage earned in I is best offer available only chance of permanent wage change is I'

#### Model – recursive structure

$$V(x,\zeta) = \max_{j} \left( u(x,j) + \zeta_{j} + \beta \sum_{x'} p(x'|x,j) \mathbb{E}_{\zeta'} V(x',\zeta') \right)$$

- Finite horizon Bellman equation
- State vector is x current and previous locations ( $I^0, I^1, \ldots$ ), age, home location
- j is the location choice
- ullet  $\zeta_j$  is preference / moving-cost shock  $\sim$  type I extreme value

## Model - specification

• Individual wakes up aged a in location  $l^0$ , gets paid and enjoys amenities, decides where to move j in the evening

$$u(x,j) = \alpha_0 w(x) + \sum_k \alpha_k A_k(x) + \alpha^H \mathbb{1}_{\{I^0 = h\}} - \Delta_{\tau}(x,j) + \xi_{I^0}$$

- $\xi_{I^0}$  is an individual-location utility fixed effect
- Moving costs  $\Delta_{\tau}(x,j)$  depend on distance, adjacency, recent occupation, etc.

$$w_i(x) = \mu_{i0} + \eta_i + \nu_{ii0} + G(X_i, a, t) + \varepsilon_i(x)$$

•  $G(X_i, a, t)$  is time-trend + effects of observables

$$w_i(x) = \mu_{i0} + \eta_i + \nu_{ii0} + G(X_i, a, t) + \varepsilon_i(x)$$

- Wage effect on location decisions  $\mu_{l^0} + \nu_{il^0}$
- Probability of moving depends only on  $\nu$ :  $ho(\mathsf{move}|
  u)$
- Shape of  $\rho(\mathsf{move}|\nu)$  determines income effect on migration
- Try to identify the choice probability by Bayes theorem

$$f_{\nu}(\nu|\mathsf{move}) = \frac{\rho(\mathsf{move}|\nu)f_{\nu}(\nu)}{\mathsf{Pr}(\mathsf{move})}$$

- Just need to identify distributions of  $\nu$  for movers and stayers
- Use Kotlarski's lemma for contaminated observations

Let  $X_1$ ,  $X_2$ , and  $\theta$  be three independent real-valued random variables and define  $Y_1 = X_1 + \theta$  and  $Y_2 = X_2 + \theta$ . Then the joint distribution of  $(Y_1, Y_2)$  determines the distributions of  $X_1$ ,  $X_2$ , and  $\theta$ .

• Define the wage residual for i at t in location j as

$$y(i,j,t) = w(i,j,t) - \mu(j) - G(X_i, a, t) = \eta(i) + \nu(i,j) + \varepsilon(i,j,t)$$

$$y(i,j,t) = \eta(i) + \nu(i,j) + \varepsilon(i,j,t)$$

• For stayer i in location j

$$y(i,j,t) = \underline{\eta(i) + \nu(i,j) + \varepsilon(i,j,t)}$$
$$y(i,j,t') = \underline{\eta(i) + \nu(i,j) + \varepsilon(i,j,t')}$$

Kotlarski means that distributions of  $\eta + \nu$  and  $\varepsilon$  are identified!

$$y(i,j,t) = \eta(i) + \nu(i,j) + \varepsilon(i,j,t)$$

• For mover i from j to j'

$$y(i,j,t) = \underline{\eta(i)} + \nu(i,j) + \varepsilon(i,j,t)$$
$$y(i,j',t') = \underline{\eta(i)} + \nu(i,j') + \varepsilon(i,j',t')$$

Kotlarski means that distributions of  $\eta$  and  $\nu+\varepsilon$  are identified!

$$f_{\nu}(\nu|\mathsf{move}) = \frac{\rho(\mathsf{move}|\nu)f_{\nu}(\nu)}{\mathsf{Pr}(\mathsf{move})}$$

• By de-convolution, can identify  $f_{\nu}(\nu|\text{move})$  and  $f_{\nu}(\nu) \to \rho(\text{move}|\nu)$ 

## Empirical implementation

- Fact that wage information in all past locations is known makes state space grow
- With J locations and n points of support for each wage distribution  $J(n+1)^J$  states for each person
- ullet Assume only recently-observed (last M < J) locations are known
- Individuals with identical recent histories have same state, irrespective of older histories
- *M*-vector *I* of locations in state

#### Data

- NLSY79 white non-Hispanic, HS grads, no post-secondary education
- n = 432 people, 4,274 person-years, 123 interstate moves
- $\mu_i$  estimated with PUMS data from 1990 Census

## Estimation - distribution assumptions

- Wage residual  $arepsilon_{it} \sim \mathcal{N}(0, \sigma_arepsilon^2)$
- Use constant-probability discrete distributions of  $\nu,~\eta,~\xi,\sigma_{arepsilon}$
- e.g. for  $\nu_{ij}$  the distribution with  $n_{\nu}=3$  is

$$\Pr(
u_{ij}) = egin{cases} 0 & \text{w.p. } 1/3 \ au_{
u} & \text{w.p. } 1/3 \ - au_{
u} & \text{w.p. } 1/3 \end{cases}$$

where  $au_{
u}$  is a parameter to be estimated

#### Estimation – likelihood function

- If individual i visits  $N_i$  locations in the data there are  $n_{\eta}n_{\sigma}(n_{\xi}n_{\nu})^{N_i} = 7\cdot 4\cdot (3\cdot 3)^{N_i}$  values  $\omega_i$ , all equally likely
- A draw is a wage-location pair get information from both

$$arepsilon_{it}(\omega^i| heta_ au) = w_{it} - \mu_{I^0(it)} - G(X_i, a_{it}| heta_ au) - 
u(\omega^i) - \eta(\omega^i)$$
 $ho(I(i,t), \omega^i| heta_ au)$  is CCP of  $I(i,t)$ 

• MLE by integrating over all  $\omega_i$ 

#### Estimation – likelihood function

• Given parameters  $\theta$ , likelihood of an individual history for a person of type  $\tau$  is

$$L_i(\theta_{\tau}) = \frac{1}{7 \cdot 4 \cdot (3 \cdot 3)^{N_i}} \times \sum_{\omega_i} \left( \prod_{t=1}^{\tau} \Pr(\varepsilon_{it}(\omega^i | \theta_{\tau})) \rho(I(i, t), \omega^i | \theta_{\tau}) \right)$$

- Types  $\tau$  for 'movers / stayers' (Johnston, 1971)
- If fraction  $\pi_{\tau}$  is of type  $\tau$  then the sample log likelihood is

$$\Lambda( heta) = \sum_{i=1}^N \log \left( \sum_{ au=1}^K \pi_ au L_i( heta_ au) 
ight)$$

## Results – goodness of fit

GOODNESS OF FIT

Moves	Binomial		NLSY		Model	
None	325.1	75.3%	361	83.6%	36,257	83.9%
One	91.5	21.2%	31	7.2%	2,466	5.7%
More	15.4	3.6%	40	9.3%	4,478	10.4%
Movers with more than one move	14.4%		56.3%		64.5%	
Total observations	432		432		43,201	

## Results – goodness of fit

#### RETURN MIGRATION STATISTICS

Movers	NLSY	Model
Proportion who		
Return home	34.7%	37.5%
Return elsewhere	4.8%	6.2%
Move on	60.5%	61.9%
Proportion who ever		
Leave home	14.4%	13.7%
Move from not-home	40.0%	42.5%
Return from not-home	25.7%	32.3%

#### Results

- Home premium is \$23,000 per year
- Average move away from a bad match increases income \$8,400.
   Move to better state increases by \$9,500
- Need high moving costs to fit lack of movement (NPV of gain is \$312,000)

Young (20)	\$384,743
'Old' (24.4)	\$312,146

#### Contribution

- First tractable model of lifetime migration decisions
  - Improves on previous literature by explicitly modeling choice procedure and increasing time- and location-dimensions
- First structural estimation of moving costs, but implausibly(?) high (\$312K+)
- First structural estimation of home premium