Before we wrote that our experimental observation is a sum of sim, discrepancy, and measurement error  $Y_{obs}(\vec{x}) = Y_{sim}(\vec{x}, \Theta) + S(\vec{x}) + \varepsilon$ 

Isim might be expensive to evaluate, a problem for calibrating O's

We can build an emulator of the simulation by running Ysim at Several x's and O's and use a regression model to give output between the points Ism was evaluated

Then we also build a regression model for S The most common way to do this for predictive models through Gaussian Process Regression

This is a non-parametric regression model, meaning we don't specify a functional form for the regression

To demonstrate the GPR methodology we begin with a parametric example before involving the non-parametric case

## GP for Standard Linear Model

We have a set D of data, n observations  $D=\xi(\bar{X}_i, Y_i)|_{i=1..n}$ 

This is called training data We can write this as a vector of length n if and dxn matrix  $\underline{X}$ ,  $D = (\underline{X}, \overline{9})$ 

write a standard linear model with gaussian

E is independent, identically distributed

Therefore, the likelihood of ann observations if given X and is

$$P(\overline{Y}|X,\overline{w}) = \prod_{i=1}^{n} P(Y_{i}|X_{i},\overline{w}) = \prod_{i=1}^{n} \frac{1}{\sigma_{n}\sqrt{2\pi}} \exp\left[-\frac{(Y_{i}-\overline{X_{i}},\overline{w}})^{2}}{2\sigma_{n}^{2}}\right]$$

$$=\frac{1}{(2\pi\sigma_n^2)^{n/2}}\exp\left(-\frac{1}{2\sigma_n^2}|\vec{\gamma}-\vec{X}\vec{w}|^2\right)=N(\vec{X}\vec{w},\sigma_n^2\vec{I})$$

each observation is independent

Now we can use Bayesian formalism to get the weights

Give a prior for  $\vec{w} \sim N(\vec{O}, \Sigma_o) = \pi(\vec{w})$ 

The posterior for w given D is

$$P(\vec{w}|\vec{X},\vec{q}) = \frac{P(\vec{q}|\vec{X},\vec{w}) \pi(\vec{w})}{(\pi(\vec{w}))P(\vec{q}|\vec{X},\vec{w}) d\vec{w}}$$

$$\propto \exp\left(-\frac{1}{2\sigma^2}\left(\vec{\gamma} - \vec{\chi}^T\vec{\alpha}\right)^T\left(\vec{\gamma} - \vec{\chi}^T\vec{\alpha}\right)\right) \exp\left(-\frac{1}{2}\vec{\alpha}^T \vec{\Sigma}_p^{-1}\vec{\alpha}\right)$$

$$\propto \exp\left(-\frac{1}{3}(\vec{w}-\vec{w}^*)^T\left(\frac{1}{3}\sum_{i}\sum_{j}\sum_{i}^{T}+\sum_{p}^{-1}\right)(\vec{w}-\vec{w}^*)\right)$$

$$\overrightarrow{M} = O_{n}^{2} \left( O_{n}^{2} \times \overrightarrow{X} + \Sigma_{p}^{2} \right)^{-1} \times \overrightarrow{A}$$

$$= \frac{1}{2} A^{-1} \times \overrightarrow{A}$$

the posterior is the gassian with mean \$7\* and covariance A1 P(W/X,Y)~N(W\*, A')

Now we can make predictions at test points \* by averaging M over all possible linear models w.r.t to Gaussian posterior

 $p(\vec{x}\vec{n}) = \int p(\vec{x}\vec{n},\vec{X},\vec{n}) = \int p(\vec{x}\vec{n},\vec{x},\vec{n}) p(\vec{n},\vec{X},\vec{n}) d\vec{n}$ 

 $= N\left(\frac{1}{\sigma_{n}^{2}} \overrightarrow{X}^{T} \overrightarrow{A}^{T} \overrightarrow{X}^{T}, \overrightarrow{X}^{*T} \overrightarrow{A}^{T} \overrightarrow{X}^{*}\right)$ 

. The pred. Variance is quadratic in test input so oncertainties grow with X\*

Example

y= W, +W2X, +E

nn(Ö, I) for posterior, E~N(O,1)

We begin with design points

 $X_{1} = -5, 1, 5$ 

y=-5.1,0,25, 4.9

(Show figs)

Now odd two points

and then change on from I to O.1

Generalization to other bases (feature space)

Instead of a linear model, we use a general function space,  $\phi(\vec{x})$ 

Example if  $\vec{x} = x_1$ ,  $\phi(x)$  could be  $\phi(x_1) = \{1, x_1, x_1^2, x_1^3, \dots \}$ 

Later we'll see that we don't have to pick  $\phi(\vec{x})$ 

 $\phi(\vec{x}): \vec{x} \rightarrow f$ length d length N

Define matrix of  $\Phi(\vec{x})$  as  $\underline{P}(X)$ :  $n \times N$  matrix

number of lest points Now y has the form  $y = \phi(\hat{x})^T \vec{w} + \varepsilon$ 

And the analysis is analogous to the linear case and we get predictive dist as  $P(\forall x)^T \vec{w} \mid \vec{x}^*, \vec{X}, \vec{Y}) \sim N(\frac{1}{\sigma_n^2} \phi(x^*)^T \vec{A}' \Phi(\vec{x})^T \vec{A}' \phi(\vec{x}^*))$ 

 $A = \mathcal{L}(X) \Phi(X)^{T} + \Sigma_{P}^{T}$ 

. This could be difficult to compute if A is large

· To get around this we use the Kernel trick

· The feature space pliways enters in the general

torn

if we rewrite

 $P(\phi(\hat{x}^*)^T \hat{\alpha} | \hat{x}^*, \underline{X}, \hat{y}) \sim N(\phi(\hat{x}^*) \Sigma_{\rho} \underline{\Phi}(\underline{x}) (K_{+\sigma_n^2} \underline{T})^T y,$ 

$$K = \overline{\Phi}(\overline{X})^T \Sigma_p \overline{\Phi}(\overline{X})$$

$$\phi_{x}^{T}Z_{p}\phi(\hat{x}^{*})-\phi(\hat{x}^{*})^{T}Z_{p}\bar{\Phi}(\underline{x})(K_{1}\sigma_{x}^{*}\underline{T})^{T}\bar{\Phi}(\underline{x})$$

· Now using k(x,x') we can make the predictions defined only using weighted inner products.

· We can define different kernels to get different feature spaces

Defining a Kernel is easier than defining a feature space

Now we will define Gaussian Process Regression as a space of functions with certain properties

Gaussian Process - A collection of random variables, any finite subset of which have a joint Gaussian distribution

A Gaussian process is completely determined by its mean and covariance function.

Given a real process  $f(\vec{x})$  the mean function is

 $m(\vec{x}) = E[f(\vec{x})]$ the covariance function is

 $k(\vec{x}, \vec{x}') = E[(f(\vec{x}) - m(\vec{x}))(f(\vec{x}') - m(\vec{x}))]$ 

For GPR the random variables are the value of  $f(\vec{x})$  at  $\vec{x}$  they have a known mean and covariance. The way a GP is defined, if  $(y_1, y_2) \sim N(\vec{\mu}, \underline{\Sigma})$  then  $y_1 \sim N(\mu_1, \Sigma_1)$ 

We will write our Gaussian process as  $f(\vec{x}) \sim GP(m(\vec{x}), k(\vec{x}, \vec{x}'))$ 

A common covariance function is the squared exponential function  $R(\vec{X}, \vec{X}') = \sup_{x \to \infty} \exp\left(\frac{-|\vec{X} - \vec{X}'|^2}{20^2}\right)$ 

This cov function is equivalent to Bayesian regression model with an infinite number of basis functions

This is an example of how at is easier to pick kernel than specify  $\phi(\vec{x})$ 

once we have a covariance function we have a distribution over functions.

To demonstrate this lets specify  $f(X^*) \sim N(\vec{0}, K(X^*, X^*))$ 

here  $\underline{X}$  is a Matrix of input points  $\overline{X}$ , K, is a covariance matrix given by  $K_{ij} = R(\overline{X}_i, \overline{X}_j)$ 

Hets specify  $X^*=[-10,-9.9,...,9.9,10]$  and use SE cov function, l=1 Then by sampling from  $N(\vec{0}, K(X^*, Y^*))$  we get our ways random functions evaluated at  $X^*$ 

· Prediction from noise-free observations

We can now generate functions from an distributions of we consider  $f(X^*) \sim N(\vec{O}, K(X^*, X^*))$  a prior and we have a set of training points  $\{(x_i, f_i) | i=1...n3\}$ 

We have a prior distribution

 $\begin{bmatrix} \frac{1}{4} \end{bmatrix} \sim N(0, \begin{bmatrix} K(X,X) & K(X,X^*) \end{bmatrix}$ 

Where for n\* prediction points,  $K(X^*, X^*)$  is  $n^* \times n^*$  and  $K(X, X^*)$  is  $n \times n^*$ 

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What we want to do is restrict this distribution so that functions that don't agree with observations  $\hat{f}$  are rejected. Because  $\hat{f}$ ,  $\hat{f}^*$  are jointly Gaussian (by def. of GP)

we can use the property that for jointly Gaussian radom verts  $\begin{bmatrix} \vec{X} \end{bmatrix} \sim N(\begin{bmatrix} \vec{A} x \\ Y \end{bmatrix}, \begin{bmatrix} A & C \\ C^T & B \end{bmatrix})$ 

P(X19)~N(xx + CB"(7-x0), A-CB"CT)

to write

 $b(\underline{t}_{\star}/\underline{t}, \overline{X}_{\star}, \overline{X}) \sim N(K(\overline{X}^{\star}, \overline{X}) K(\overline{X}, \overline{X}), \underline{t}'$ 

 $|\langle (\bar{X}_{*}, \bar{X}_{*}) - |\langle (\bar{X}_{*}, \bar{X}) | \langle (\bar{X}, \bar{X})_{-1} | \langle (\bar{X}, \bar{X}_{*}) \rangle$ 

Now we simpley need to compute  $K(X^*, X)$ ,  $K(X^*, Y^*)$ ,  $K(X, Y)^T$ 

Example: Take 5, points and constrain previous example.

The data interpolates points, away from training points uncertainty goes up.

Prediction with noisy observations

In this case we don't know f(x), rather  $y = f(x) + \varepsilon$ 

now write

 $Cov(\vec{g}) = K(\underline{X}, \underline{X}) + \sigma_n^2 \underline{T}$ 

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$$\begin{bmatrix} t_* \\ \underline{\zeta} \end{bmatrix} \underset{\sim}{\sim} N \left( \underbrace{Q}, \begin{bmatrix} K(\bar{X}_*, \bar{X}) & K(\bar{X}_*, \bar{X}_*) \\ K(\bar{X}, \bar{X}) + \alpha_s \underline{T} & K(\bar{X}, \bar{X}_*) \end{bmatrix} \right)$$

The posterior distribution is

where 
$$\hat{f}^* = K(\underline{X}^*, \underline{X}) [K(\underline{X}, \underline{X}) + on^* \underline{T}]^{-1} \hat{\eta}$$

$$\operatorname{Con}(\underline{\xi_{\star}}) = K(\bar{X}_{\star}, \bar{X}_{\star}) - K(\bar{X}_{\star}, \bar{X}) \left[ K(\bar{X}, \bar{X}) + o_{s} I \right]_{l} K(\bar{X}, \bar{X}_{\star})$$

What do the covariance params mean

$$R(\vec{X}, \vec{X}_b) = \sigma_b^2 \exp\left(-\frac{1}{2l^2}|\vec{X}_a - \vec{X}_b|^2\right) + \sigma_a^2 \delta_b$$

I = length scale, Of = signal variance, on = noise variance

I controls how oscillatory the GP is

if l is too small function can interpolate data better, but

less accurate between

· if lis to large function might not bend enough to hit the abs.

· low of means that data is not less significant

high of means that all data (both exp and sim) is sig.

Empirical Bayes Approach

What if we have no idea how to pick 1, of, on??

· We can use Bayesian interence to pick them.
To more closely match the notation in the literature we'll write our covariance function

$$\Re(\vec{x}, \vec{x}') = \frac{1}{\lambda_{n}} \exp\left[-\sum_{k=1}^{n} \beta_{k} | x_{k} - x_{k}'|^{\alpha}\right]$$

$$\lambda_n = \frac{1}{\sigma_{\ell}^2}$$
  $\beta_R = \frac{1}{\partial l_R}$ 

Now we have different l's for different elements of  $\vec{X}$ . Given  $\vec{z} = \begin{bmatrix} \vec{Y} \\ \vec{f}^* \end{bmatrix}$  what is the likelihood of getting

the output 2 with our GP with cov. params 7, we also now allow general mean, M

The likelihood is

$$P(\overline{z}|\mu,\lambda_2,\overline{\beta}) \propto |\Sigma_z|^{\frac{1}{2}} \exp[-\frac{1}{2}(\overline{z}-\mu 1)^T \Sigma_z^T (\overline{z}-\mu 1)]$$

$$\sum_{z} = \left[ K(\underline{X}',\underline{X}) + o_{x_{z}}^{z} \underline{I}, K(\underline{X}',\underline{X}') \right]$$

Therefore, the posterior distribution of the parameters is

$$P(\mu, \lambda_n, \hat{B}|\hat{z}) \propto p(\hat{z}|_{\mu, \lambda_n} \hat{B}) \pi(\hat{B}) \pi(\mu) \pi(\lambda_n)$$

Common choices of priors

$$\pi(\mu) \propto \exp\left(-\frac{\mu^2}{2\nu}\right) \quad (\nu=0\Rightarrow \mu=0)$$

$$\pi(\lambda_n) \propto \lambda_n^{\alpha-1} \exp(-b\lambda_n)$$
,  $\lambda_n > 0$ ,  $(a, b | large encourages)$ 

$$T(\vec{\beta}) \propto T(1-e^{-\beta_R})^{-\frac{1}{2}} e^{-\beta_R}$$
 $B_k > 0$  (encourages large  $B_{R_k}$ , Smalle)

Now we can use MCMC to find params This will be inexpensive because we don't need more code runs, the runs are already contained in T

Adding Discrepancy Term

Real discrepancy term:

We use a Gaussian process to model 8, with mean zero Covariance

$$k(\tilde{x},\tilde{x}') = \frac{1}{\lambda_s} \exp\left(-\frac{\hat{x}}{R-1} B_R^{s}(X_R - X_R')^{\alpha_s}\right)$$

The likelihood has the same form

$$p(\hat{z}|_{\mu,\lambda_{z},\hat{\beta},\lambda_{s}},\hat{\beta}^{s}) \propto |\Sigma_{z}|^{\frac{1}{2}} \exp\left[-\frac{1}{a}(\hat{z}-\mu\hat{1})^{T}\Sigma_{z}^{T}(\hat{z}-\mu\hat{1})\right]$$

where now  $\Sigma_z^2 = \left[K(\underline{X}, \underline{X}) + \sigma_n^2 \underline{I} + K_s(\underline{X}, \underline{X}), K(\underline{X}, \underline{X}') + \sigma_n^2 \underline{I} + K_s(\underline{X}, \underline{X}'), K(\underline{X}, \underline{X}')\right]$ 

$$K(\hat{\lambda}, \hat{\lambda})$$