

FACTORIZATION THEOREM

$X \sim \text{DISCRETE}$

$V = V(X) \iff \exists h, g, \nu$ ST $\forall x, \nu$
 SUFFICIENT $P_{\theta}(x) = g_{\theta}(V(x))h(x)$

$(\Rightarrow) V$ IS SUFFICIENT
 $P_{\theta}(x) = P_{\theta}(x=x)$
 $= P_{\theta}(x=x, V=\nu)$
 $= P(x=x|V=\nu)P_{\theta}(V=\nu)$ $\left| \begin{array}{l} P(x=x|V=\nu) \text{ DOES NOT DEPEND ON } \theta \\ \text{BECAUSE } V \text{ IS SUFF} \end{array} \right.$
 $= h(x) \cdot g_{\theta}(V(x))$

$(\Leftarrow) \exists h(x), g_{\theta}(V(x))$ ST $P_{\theta}(x) = g_{\theta}(V(x))h(x)$
 WTS: $P(X=x|V=\nu)$ DOES NOT DEP. ON θ
 TAKE x_0 ST. $V(x_0)=\nu$ $P_{\theta}(X=x_0|V=\nu) = \frac{P_{\theta}(X=x_0, V=\nu)}{P_{\theta}(V=\nu)}$
 $\left| \begin{array}{l} P_{\theta}(V=\nu) \leftarrow V \text{ IS } f(X) \\ P_{\theta}(X=x_0) \end{array} \right.$

$$= \frac{\sum_{x: V(x)=\nu} P_{\theta}(x=x)}{P_{\theta}(V=\nu)} = \frac{\sum_{x: V(x)=\nu} g_{\theta}(V(x)) \cdot h(x)}{\sum_{x: V(x)=\nu} g_{\theta}(V(x)) \cdot h(x)} = \frac{h(x_0)}{\sum_{x: V(x)=\nu} h(x)} \quad \left. \begin{array}{l} \text{NO } \theta \text{ DEPEND.} \\ \text{VALUE IS THE SAME} \end{array} \right\}$$

RAO-BLACKWELL

$V = V(X)$ SUFF.
 TESTINATOR FOR $g(\theta)$
 $\exists T^* = T^*(V)$ ST.
 $E_{\theta}(T) = E_{\theta}(T^*)$
 $V_{\theta}(T^*) \leq V_{\theta}(T)$
 $(\Rightarrow \text{NSE WILL ALSO BE SMALLER})$

$X \sim \text{DISCRETE}$
 $T^* = E(T|V)$

1) $E_{\theta}(T) = E_{\theta}(T^*)$

$$E_{\theta} T^* = E_{\theta} [E_{\theta}(T|V)] = E_{\theta}(T) \quad (\text{TOWER RULE})$$

2) $V_{\theta} T^* \leq V_{\theta} T$

$$\begin{aligned} E_{\theta} T T^* &= \sum_{\nu} E[T T^* | V=\nu] P_{\theta}(V=\nu) \quad (\text{CONDITION BY } V) \\ &= \sum_{\nu} T^*(\nu) E[T | V=\nu] P_{\theta}(V=\nu) \quad (\text{EXTRACT } T^* \text{ FROM INDEP.}) \\ &= \sum_{\nu} T^*(\nu)^2 P_{\theta}(V=\nu) \quad (\text{AGAIN}) \\ &= E_{\theta} [(T^*)^2] \end{aligned}$$

$$\begin{aligned} E_{\theta} T^2 &= E_{\theta} (T - T^*)^2 + 2E_{\theta} (T - T^*) T^* + E_{\theta} (T^*)^2 \\ &= E_{\theta} (T - T^*)^2 + 0 + E_{\theta} (T^*)^2 \\ &\geq E_{\theta} [(T^*)^2] \end{aligned}$$

USED TO SHOW $= 0$

T, T^* HAVE SAME EXPECTATION,
 $V(T) = E(T^2) - E(T)^2$ IMPLIES

$$V_{\theta} T^* \leq V_{\theta} T$$

LEHMAN-SCHIFFE

$V = V(X)$ SUFFICIENT, COMPLETE $\Rightarrow T$ IS UNIVU
 $T = T(V)$ UNBIASED ESTIMATOR FOR $g(\theta)$

1) RAO-BLACKWELL: $\forall S$ ESTIMATOR $S(V)$ OF $g(\theta)$
 FOR $g(\theta)$ ST. $E_{\theta}(S) = g(\theta)$.

USE SUFFICIENCY OF V

BY $V = V(X)$ SUFFICIENT,
 $\exists S^* = S^*(V)$ FOR $g(\theta)$, USING ONLY V ,
 $E_{\theta}(S^*) = E(S) = g(\theta)$ AND $V(S^*) \leq V(S)$

FOR ALL ESTIMATORS THERE IS A BEST ONE

2) $E_{\theta}(S^* - T) = E_{\theta}(S^*) - E_{\theta}(T) = 0 \rightarrow$ THE BEST ESTIMATOR IS UNIQUE

BY V BEING : $P_{\theta}(S^* = T) = 1$ VS EST. FOR $g(\theta)$
 $\Rightarrow S^* = T$ A.S. AND $V_{\theta}(T) \leq V_{\theta}(S)$

UNIQUENESS

FISHER INFORMATION

$X, Y \perp$

$$i_{(X,Y)} = i_X + i_Y$$

$\theta \in \mathbb{R}$

$$P_{\theta}(x, y) = P_{\theta}(x) P_{\theta}(y)$$

$$i_{\theta} = \text{Var} \left(\frac{\partial}{\partial \theta} \ln(P_{\theta}(x, y)) \right)$$

$$= \text{Var} \left(\frac{\partial}{\partial \theta} [\ln(P_{\theta}(x)) + \ln(P_{\theta}(y))] \right) \quad \left. \begin{array}{l} X, Y \text{ iid SO} \\ \text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y) \end{array} \right\}$$

NEYMAN-PEARSON LEMMA

$\exists C_{\alpha_0}$ ST. $P(L(\vartheta_0, \vartheta_1, X) \geq C_{\alpha_0}) = \alpha_0$
 THEN THE TEST
 WITH $K = \{X: L(\vartheta_0, \vartheta_1, X) \geq C_{\alpha_0}\}$
 IS THE MOST POWERFUL AT α_0
 FOR $H_0: \vartheta \in \Theta_0$
 $H_1: \vartheta \in \Theta_1$

FIX $\alpha = \alpha_0$
 BY ASSUMPTION, $\exists K'$ ST. $P_{\Theta_0}(X \in K') \leq \alpha_0$
 ANY OTHER CONFIDENCE SET
 MEANS THAT K IS MOST POWERFUL
 MORE POWERFUL TEST REGION

$$X \in K, \frac{P_{\vartheta_1}(X)}{P_{\vartheta_0}(X)} \geq C_{\alpha_0}$$

$$P_{\vartheta_1}(X) \geq C_{\alpha_0} P_{\vartheta_0}(X)$$

$$1) \forall x: (I_{K'}(x) - I_K(x)) (P_{\vartheta_1}(x) - C_{\alpha_0} P_{\vartheta_0}(x)) \leq 0$$

DEPENDING ON THE
 PLACEMENT OF X , THE
 TEST BEHAVES IN
 ONE OF TWO WAYS

SAME TEST,
 DIFFERENT
 CRITICAL
 REGION

$$2) P_{\vartheta_1}(X \in K') - P_{\vartheta_1}(X \in K) = \int I_{K'}(x) P_{\vartheta_1}(x) dx - \int I_K(x) P_{\vartheta_1}(x) dx$$

$$= \int (I_{K'}(x) - I_K(x)) P_{\vartheta_0}(x) dx \quad (\text{SEE INEQUALITY})$$

$$\leq C_{\alpha_0} \int (I_{K'}(x) - I_K(x)) P_{\vartheta_0}(x) dx$$

$$= C_{\alpha_0} (P_{\vartheta_0}(X \in K') - P_{\vartheta_0}(X \in K))$$

$$\leq C_{\alpha_0} (\alpha - \alpha_0) = 0$$

REASON BY
 CONTRADICTION,
 EXPECT VALUE
 TO BE ≥ 0

BREAKING ASSUMPTION $\Rightarrow K$ IS MOST
 POWERFUL
 TEST

CRAMER-RAO LOWER BOUND

$\vartheta \mapsto p_{\vartheta}(x)$ DIFFERENTIABLE

IN ANY UNBIASED
 ESTIMATOR FOR ϑ : $V_{\vartheta}(T) \geq \frac{g'(\vartheta)^2}{I_{\vartheta}}$

$$g'(\vartheta) = E[T]$$

$$= \frac{\partial}{\partial \vartheta} \int T(x) p_{\vartheta}(x) dx$$

$$= \int T(x) \dot{p}_{\vartheta}(x) dx$$

$$= \int T(x) p_{\vartheta}(x) \dot{\ell}(x) dx$$

$$= E[T(X) \dot{\ell}(X)]$$

(IF $X \sim$ DISCRETE,
 USE \sum INSTEAD)

$$E_{\vartheta}[\dot{\ell}(X)] = 0 \quad \text{"PHANTOM TERM"}$$

$$g'(\vartheta) = E_{\vartheta}[T \dot{\ell}(X)] - E_{\vartheta}(T) E_{\vartheta}[\dot{\ell}(X)]$$

$$= \text{COV}(T, \dot{\ell}(X))$$

$$g'(\vartheta)^2 = \text{COV}(T, \dot{\ell}(X))^2$$

$$\leq V_{\vartheta}(T) \cdot V_{\vartheta}(\dot{\ell}(X)) \quad \text{CAUCHY SCHWARTZ}$$

$$= V_{\vartheta}(T) \cdot I_{\vartheta}$$

$$V_{\vartheta}(T) \geq \frac{g'(\vartheta)^2}{I_{\vartheta}}$$