Multivariate Normal Distribution

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For this presentation, I made references to "Introduction to Probability" by Joseph K. Blitzstein and Jessica Hwang. I would also like to thank Professor Everson for his instruction on this problem.

Multivariate Normal Distribution Definition

4. Multivariate Normal Variables (Blitzstein 7.5)

Definition: A random vector (X_1, \ldots, X_n) follows a *Multivariate Normal* (MVN) distribution if $t_1X_1 + \ldots + t_nX_n$ follows a Normal distribution for any choice of constants t_1, \ldots, t_n .

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a) Explain why this implies that any sample of *n independent* Normal random variables follows and dimensional MVN distribution.

MGF Proof (PSU)

b) Explain why, if X_i is part of a MVN vector, then X_i is Normal. Also give an example of variables X and Y that have marginal Normal distributions, but that are not part of a multivariate Normal vector.

Goal: If
$$X_i$$
 part of a MUN vector, then X_i is Normal X_i , is part of a MUN vector: $t_1X_1 + t_2X_2 + ... + t_1X_1 + ... + t_n \times n$ We know that $t_1,...,t_n$ is any choice of constants (given by MUN distribution definition) Chaose $t_i = 1$, $t_i = 0$

Then, $O(x_i) + O(x_2) + ... + I(x_i) + ... + O(x_n) = X_i$

Therefore X_i is Normal (becomes a univariate Normal)

b) Explain why, if X_i is part of a MVN vector, then X_i is Normal. Also give an example of variables X and Y that have marginal Normal distributions, but that are not part of a multivariate Normal vector.

References Blitzstein textbook (Page 310)

c) The *n*-dimensional multivariate Normal density with $n \times 1$ mean vector $\boldsymbol{\mu}$ and $n \times n$ covariance matrix \mathbf{V} has a proper joint pdf if and only if \mathbf{V} is positive definite, and hence invertible. We write $\mathbf{X} \sim N_n(\boldsymbol{\mu}, \mathbf{V})$, and the joint pdf is

$$f(\mathbf{x}) = (2\pi)^{-n/2} |\mathbf{V}|^{-1/2} \exp \left[-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^{\mathrm{T}} \mathbf{V}^{-1} (\mathbf{x} - \boldsymbol{\mu}) \right]$$

Show that this is a generalization of the univariate Normal density (n = 1) and that the vector of $X_1, \ldots, X_n \stackrel{\text{i.i.d.}}{\sim} N(\mu, \sigma^2)$ has a joint pdf of this form (what are μ and \mathbf{V} ?).

Linear Algebra Refresher:

Transpose:

$$X = \begin{bmatrix} 1 & 2 & 3 \\ 4 & 5 & 6 \end{bmatrix} \quad X^{T} = \begin{bmatrix} 1 & 4 \\ 2 & 5 \\ 3 & 6 \end{bmatrix} \quad |X| = \begin{bmatrix} 1 & 2 \\ 3 & 4 \end{bmatrix} = \frac{1}{(1)(4)-(2)(3)} \quad |X| = \frac{1}{(1)(4)-(2)(3)} \quad |X| = \frac{1}{(1)(4)-(2)(3)} \quad |X| = \frac{1}{(1)(4)-(2)(3)} \quad |X| = \frac{1}{(2)(4)-(2)(3)} \quad |X| = \frac{1}{(2)(4)-(2)(4)} \quad |X| = \frac{1}{($$

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Show that this is a generalization of the univariate Normal density (n = 1) and that the vector of $X_1, \ldots, X_n \stackrel{\text{i.i.d.}}{\sim} N(\mu, \sigma^2)$ has a joint pdf of this form (what are μ and \mathbf{V} ?).

Goal: Starting with univariate Normal Distribution:
$$\frac{1}{\sqrt{2\pi}} = \frac{1}{2} \left(\frac{(x-u)^2}{\sqrt{\sigma}} \right)^2$$
to $\left(\frac{2\pi}{2\pi} \right)^{-\frac{1}{2}} \left[\frac{1}{\sqrt{2\pi}} \left(\frac{1}{\sqrt{2\pi}} \right)^{-\frac{1}{2}} \left(\frac{(x-u)^2}{\sqrt{2\pi}} \right)^{-\frac{1}{2}} \left(\frac{(x-u)^2}{\sqrt{2\pi}} \right)^2$
and find u and V.

Start:
$$\frac{1}{\sqrt{2\pi}} = \frac{1}{\sqrt{2\pi}} \left(\frac{(x-u)^2}{\sqrt{2\pi}} \right)^2$$
We know: The multivariate Normal density is the product of n univariate Normal densities

Start: $\frac{1}{\sqrt{2\pi} \sigma} e^{-\frac{1}{2}\left(\frac{x_{i}}{\sigma}\right)}$ We know: The multivariate Normal density is the product of n Univariate Normal densities $(x_{i},...,x_{n}) \stackrel{\text{iid}}{\sim} N(u,\sigma^{2})$ $(x_{i},...,x_{n}) \stackrel{\text{iid}}{\sim} N(u,\sigma^{2})$ $(x_{i},...,x_{n}) \stackrel{\text{iid}}{\sim} N(u,\sigma^{2})$ $(x_{i},...,x_{n}) \stackrel{\text{iid}}{\sim} N(u,\sigma^{2})$

$$= (2\pi)^{-n/2} (0^2)^{-n/2} \exp \left[-\frac{1}{2} \left(\frac{1}{2}(x_i - u)^2 / o^2\right)\right]$$

$$|V|^{-1/2}: |V| = |O|^{2} O$$
 from $(O^{2})^{-n/2}$ Since Covariance matrix is O^{2} I and $\det = (O^{2})^{n}$

$$Helpful Facts:$$

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$$If \quad X = \begin{bmatrix} x \end{bmatrix}, \text{ then } X^{T} = \begin{bmatrix} x \end{bmatrix} \quad \text{and} \quad (x-u)^{2} = (x-u)^{T}(x-u)$$

$$|x| \quad |x| \quad |x|$$

Then we get:
$$\left(2\pi\right)^{-1/2} \left|V\right| \exp^{-1/2\left(x-u\right)T}V^{-1}\left(x-u\right)\right]$$

We have a nx1 mean vector u = u (!)

From the problem we know:

We have a nxn covariance matrix
$$V = 0^a I = 0^a$$

d) Let $X_1, \ldots, X_n, Y_1, \ldots, Y_m$ represent a multivariate Normal vector of dimension m + n, and suppose $Cov(X_i, Y_j) = 0$ for all $i = 1, \ldots, n$ and $j = 1, \ldots, m$. Show that the joint density of the X's and Y's factors into a joint pdf for X_i 's and a joint pdf for the Y_j 's, meaning these are independent random vectors. State the general result about correlation and independence for elements of a multivariate Normal vector.

independent, random vector while stating the general result of correlation and independence.

Start:
$$X_1, ..., X_n, Y_1, ..., Y_m$$
 is a multivariate normal vector; dimension $m+n$

$$\begin{pmatrix} X_1 \\ X_n \\ Y_1 \\ Y_m \end{pmatrix}$$

$$\begin{pmatrix} X_1 \\ X_n \\ Y_1 \\ Y_m \end{pmatrix}$$

$$\begin{pmatrix} X_1 \\ X_n \\ Y_1 \\ Y_m \end{pmatrix}$$

$$\begin{pmatrix} X_1 \\ X_n \\ Y_1 \\ Y_m \end{pmatrix}$$

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$$\begin{pmatrix} X_1 \\ X_1 \\ X_1 \\ Y_1 \\ Y_1 \\ Y_1 \\ Y_2 \\ Y_3 \end{pmatrix}$$

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$$\begin{pmatrix} X_1 \\ Y_3 \\ Y_4 \\ Y_3 \end{pmatrix}$$

$$\begin{pmatrix} X_1 \\ Y_2 \\ Y_3 \\ Y_4 \\ Y_4 \end{pmatrix}$$

$$\begin{pmatrix} X_1 \\ Y_2 \\ Y_3 \\ Y_4 \\ Y_4 \\ Y_5 \end{pmatrix}$$

$$\begin{pmatrix} X_1 \\ Y_2 \\ Y_3 \\ Y_4 \\ Y_4 \\ Y_5 \\ Y_5 \\ Y_5 \\ Y_5 \\ Y_6 \\ Y_6 \\ Y_7 \\ Y_8 \\$$

Goal: We want to show that f(x, y) = f(x) f(y) to mean that there are

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$$V = \frac{V_{1}}{O} \frac{O}{V_{2}}$$

$$|V| = |V_{1}||V_{2}|$$

$$V^{-1} = \begin{pmatrix} v_{1}^{-1} & 0 \\ 0 & v_{2}^{-1} \end{pmatrix}$$

$$f_{xy} = \begin{pmatrix} a_{1} \end{pmatrix}^{-\frac{(n+m)}{2}} |V|^{-\frac{1}{2}} \exp \left[-\frac{1}{2} \left((x-u_{1})^{T} (y-u_{1})^{T} \right) V^{-1} \left(\frac{x-u_{2}}{y-u_{2}} \right) \right] + \int_{-1}^{1} \left[(x-u_{1})^{T} V_{1}^{-1} (x-u_{2}) + -\frac{1}{2} (y-u_{2})^{T} V_{2}^{-1} (y-u_{2}) \right]$$

$$f_{xy} = \frac{V_{1}}{O} \frac{O}{V_{2}} \left[V_{1} \right]^{-\frac{1}{2}} \exp \left[-\frac{1}{2} \left((x-u_{2})^{T} V_{1}^{-1} (x-u_{2}) + -\frac{1}{2} \left((y-u_{2})^{T} V_{2}^{-1} (y-u_{2}) \right) \right]$$

$$f_{xy} = \frac{V_{1}}{O} \frac{O}{V_{2}} \left[V_{1} \right]^{-\frac{1}{2}} \exp \left[-\frac{1}{2} \left((x-u_{2})^{T} V_{1}^{-1} (x-u_{2}) + -\frac{1}{2} \left((y-u_{2})^{T} V_{2}^{-1} (y-u_{2}) \right) \right]$$

 $f_{xy}(x,y) = f(x)f_y(y)$ if and only if X, y are independent and correlation =0

Example

Example:
$$N = m = I$$

$$\begin{pmatrix} x \\ y \end{pmatrix} \sim \mathcal{N}_{2} \begin{pmatrix} u_{1} \\ u_{2} \end{pmatrix}, \begin{pmatrix} v_{1} & 0 \\ 0 & v_{2} \end{pmatrix}$$

$$f_{(x,y)}(x,y) = \frac{1}{\sqrt{2\pi} V_{1}^{\prime / 2}} e^{-\frac{(x-u_{x})^{2}}{2 V_{1}}} \frac{1}{\sqrt{2\pi} V_{2}^{\prime / 2}} e^{-\frac{(y-u_{y})^{2}}{2 V_{2}}}$$

$$\begin{vmatrix} v_{1} & 0 \\ 0 & v_{2} \end{vmatrix} = v_{1} v_{2} \qquad \begin{bmatrix} v_{1} & 0 \\ 0 & v_{2} \end{bmatrix}^{-1} = \begin{bmatrix} v_{1}^{-1} & 0 \\ 0 & v_{2}^{-1} \end{bmatrix}$$

Application of Multivariate Normal Distributions: Stock Return

Financial Economics: Stock market and how to estimate stock returns

"A Test for Multivariate Normality in Stock Returns" by Richardson and Smith in 1992 aimed to see if stock returns could be a multivariate normal distribution (previously stock return was not able to be described as univariate normal distributions)

Discussed dependence, cross moments, multivariate normal, multivariate time series, and generalized method of moments estimator to run tests on stock/asset returns and market model residuals.

Determined "non normality in both the marginal and joint distributions of these variables." (Richardson and Smith, 1992)

Richardson and Smith Paper