

# Stat 111

Olivia R. McClammy

February 12, 2025

*For this problem, I made references to "Introduction to Probability" by Joseph K. Blitzstein and Jessica Hwang and "Mathematical Statistics: Third Edition" by John A. Rice. I would also like to thank Professor Everson on his instruction and for discussion with Sarah Cooper on this problem.*

## 1 Question 2

a.) Point out that, for a non-linear function  $g(x)$ , it is unusual for  $E(g(X))$  to equate to  $g(E(X))$ . Give justification for Jensen's inequality and state its implications for expected values. For example, show that  $E(1/X) \geq 1/E(X)$ , and that  $E(\log(X)) \leq \log(E(X)) \leq E(X) - 1$ .

For a non-linear function  $g(X)$ , it is unusual for  $E(g(X)) = g(E(X))$

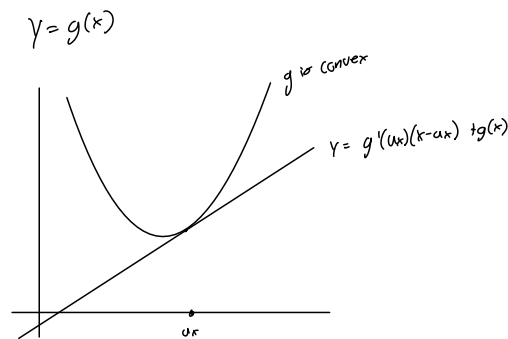
**Convex** = second derivative is positive

**Concave** = second derivative is negative

**Consequences of convex and concave:**

For a convex function, the tangent line is equal or below the convex function.

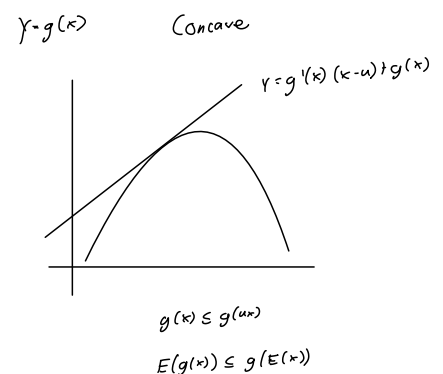
For a concave function, the tangent line is equal or above the concave function.



$$g(x) \geq g(u)$$

$$E(g(x)) \geq g(E(x))$$

1

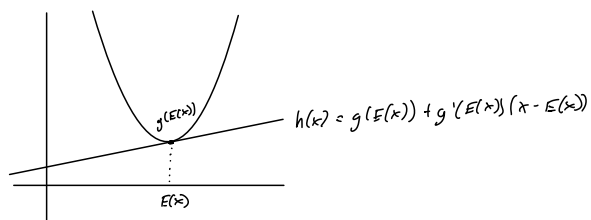


$$g(x) \leq g(u)$$

$$E(g(x)) \leq g(E(x))$$

**Jensen's Inequality** (From Blitzstein - page 425): Let  $X$  be a random variable. If  $g$  is a convex function, then  $E(g(X)) \geq g(E(X))$ . If  $g$  is a concave function, then  $E(g(X)) \leq g(E(X))$ . In both cases, the only way that the equality holds is if there are constants  $a$  and  $b$  such that  $g(X) = a + bX$  with probability 1.

$$\text{Let } g(x) = x^2$$



$$\text{Then } g(x) \geq h(x)$$

$$E(g(x)) \geq E(h(x))$$

$$E(h(x)) = g(E(x)) + 0$$

$$\text{So, } E(g(x)) \geq g(E(x))$$

$$E(1/X) \geq 1/E(X)$$

$$g(x) = 1/x$$

$$g'(x) = -1/x^2$$

$$g''(x) = 2/x^3 \leftarrow g \text{ is convex}$$

$$E(g(x)) \geq g(E(x))$$

$$E(1/x) \geq 1/E(x)$$

$$E(\log(X)) \leq \log(E(X)) \leq E(X) - 1$$

$$g(x) = \log(x)$$

$$g'(x) = 1/x$$

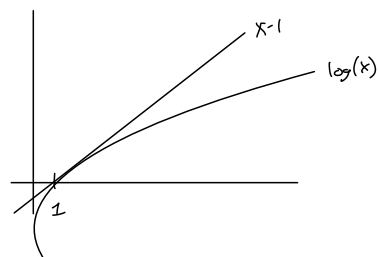
$$g''(x) = -1/x^2$$

$$E(g(x)) \leq g(E(x))$$

$$E(\log(x)) \leq \log(E(x))$$

$$\log(E(x)) \leq E(x) - 1$$

$$\log(x) \leq x - 1$$



$$\log(E(x)) \leq E(x) - 1$$

b.) State Theorem A of 4.1 in Rice and explain why this is called the law of the unconscious statistician (see also Blitzstein(4.5). Explain how LOTUS is used when we find  $Var(X) = E((X - E(X))^2) = E(X^2) - E(X)^2$ . Describe what would be involved to find a variance without using LOTUS (e.g., for a Gamma variable).

**Theorem A:** Rice (page 122)

Suppose that  $Y = g(X)$

a.) If  $X$  is discrete with frequency function  $p(X)$ , then

$$E(Y) = \sum_x g(X)p(X)$$

provided that  $\sum |g(X)|p(X)$

b.) If  $X$  is continuous with density function  $f(X)$ , then

$$E(Y) = \int_{-\infty}^{\infty} (g(X)f(X))dx$$

provided that  $\int |g(X)|f(X)dx$

We call this the law of unconscious statistician's law because we do not think about using it and we use LOTUS all the time.

How is LOTUS used in  $Var(X) = E((X - E(X))^2) = E(X^2) - E(X)^2$ ?

When we use LOTUS:

$$E(X^2) = \int_{-\infty}^{\infty} x^2 f_x(x) dx$$

If we don't use LOTUS

$$Y = X^2$$

$$E(Y) = \int_{-\infty}^{\infty} y f_Y(y) dy, \text{ where we would have to find the pdf of } Y.$$

$$Var(X) = E((X - E(X))^2) = E(X^2) - E(X)^2$$

$$= E(X^2 - 2(E(X)X) + E(X)^2)$$

$$= E(X^2) - E(2(E(X)X)) + E(E(X)^2)$$

$$= E(X^2) - 2E(X)^2 + E(X)^2$$

$$= E(X^2) - E(X)^2$$

Suppose that  $Y = g(X)$ . If we did not have LOTUS, we would need to find the pdf for  $Y$  and evaluate the integral of  $y$  multiplied by  $y$ 's pdf.

c.) State the multivariate version of LOTUS and show this implies  $E(X + Y) = E(X) + E(Y)$ , even if  $X$  and  $Y$  are not independent. Describe what would be involved to find a covariance without using LOTUS.

**Multivariate Version of LOTUS:** Rice (page 123)

Suppose that  $X_1, \dots, X_n$  are jointly distributed random variables and  $Y = g(X_1, \dots, X_n)$

a.) If the  $X_i$  are discrete with frequency function  $p(X_1, \dots, X_n)$ , then

$$E(Y) = \sum_{X_1, \dots, X_n} g(X_1, \dots, X_n) p(X_1, \dots, X_n)$$

provided that  $\sum_{X_1, \dots, X_n} |g(X_1, \dots, X_n)| p(X_1, \dots, X_n) < \infty$

b.) If the  $X_i$  are continuous with joint density function  $f(X_1, \dots, X_n)$ , then

$$E(Y) = \int \dots \int g(X_1, \dots, X_n) f(X_1, \dots, X_n) dX_1, \dots, dX_n$$

provided that the integral with  $|g|$  in place of  $g$  converges.

How does the multivariate version of LOTUS imply  $E(X + Y) = E(X) + E(Y)$ ?

$$E(X+Y) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (x+y) f_{X,Y}(x,y) dx dy$$

$$\text{Let } T = X+Y$$

$$\text{Find the pdf } T, \quad E(X+Y) = \int t f_T(t) dt$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (x f_{X,Y}(x,y) + y f_{X,Y}(x,y)) dx dy$$

$$\underbrace{\int_{-\infty}^{\infty} x \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy dx}_{E(X)} + \underbrace{\int_{-\infty}^{\infty} y \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx dy}_{E(Y)}$$

$$E(X+Y) = E(X) + E(Y)$$

If we did not have LOTUS, to find the covariance between  $X$  and  $Y$ , we would have to use density  $Z = X + Y$  and find the pdf of  $Z$ , and then find  $\int_{-\infty}^{\infty} f(Z) Z dZ = E(X + Y)$ , we would also have to

find the density of  $E(T)$

where  $T = X \cdot Y$

d.) Show that, for uncorrelated random variables,  $Var(X+Y) = Var(X) + Var(Y)$ , but that this is not true in general. Give examples using extremely correlated variables (e.g.,  $X_1 = X_2$  and  $X_1 = -X_2$ ).

$$Var(X+Y) = Var(X) + Var(Y)$$

$$\text{Let } Z = X+Y$$

$$Var(Z) = E(Z^2) - E(Z)^2$$

$$= E((X+Y)^2) - (E(X+Y))^2$$

$$= E(X^2 + 2XY + Y^2) - \underbrace{E(X+Y)^2}$$

$$= \underbrace{E(X^2) - E(X)^2}_{Var(X)} + \underbrace{E(Y^2) - E(Y)^2}_{Var(Y)} + \underbrace{2(E(XY) - E(X)E(Y))}_{Covariance}$$

$$Var(X+Y) = Var(X) + Var(Y) + 2(E(XY) - E(X)E(Y)) \quad \text{If independent} = 2Var(X_1)$$

$$\text{If } X_1 = X_2, \rho = 1$$

$$Var(\underbrace{X_1 + X_2}_{X_1}) = Var(2X_1) = 2^2 Var(X_1) = 4Var(X_1)$$

$$\text{If } X_1 = -X_2, \rho = -1$$

$$Var(\underbrace{X_1 + X_2}_{-X_2})$$

$$Var(-X_2 + X_2) = 0$$

$$Var(0) = 0$$

e.) Show how to find approximations to the mean and variance of a transformation of a random variable using Taylor's approximation. As an example, find approximations to the mean and variance of  $Y = \log(X)$ , for  $X \sim \text{Gamma}(\alpha, \lambda)$ .

**Taylor's approximation:**  $g(x) = g(\mu_x) + \underbrace{(x - \mu_x)g'(\mu_x)}_{\sqrt{x} - \mu_x} + ((x - \mu_x)/2)g''(\mu_x) + \dots$

$$Y = \log(x), \text{ for } X \sim \text{Gamma}(\alpha, \lambda)$$

Find an approximation to the mean:

$$Y = \log(X) = g(X)$$

$$g'(x) = 1/x, \quad g''(x) = -1/x^2$$

$$\text{Gamma} = u_x = \frac{\alpha}{\lambda}, \quad \text{Var}(x) = \frac{\alpha}{\lambda^2}$$

$$= g(u_x) + \underbrace{0} + \frac{\text{Var}(x)}{2} g''(u_x)$$

$$= \log(g(u_x)) + \log\left(\frac{\text{Var}(x)}{2}\right) g''(u_x)$$

$$= \log\left(\frac{\alpha}{\lambda}\right) + -\frac{1}{2} \left(\frac{\alpha}{\lambda^2}\right) \left(\frac{\alpha}{\lambda^2}\right) = \log\left(\frac{\alpha}{\lambda}\right) - \frac{1}{2\alpha} \leq \log(E(X))$$

Find an approximation for variance:

Jensen's inequality

We want to use first order approximation

$$g(x) \approx \underbrace{g(u_x)} + \underbrace{(x-u_x)} \underbrace{g'(u_x)}$$

$$\text{Var}(g(x)) \approx \text{Var}(x) (g'(u_x))^2$$

$$= \left(\frac{\alpha}{\lambda^2}\right) \left(\frac{1}{\alpha^2}\right) = 1/\alpha$$