

## 0 Introduction

**Example 0.1. Perfect riffle shuffle**  $\sigma \in S_{52}$  s.t.  $\sigma(j) := 2j \mod 51$ . I.e.

$$\sigma(j) := \begin{cases} 2j & \text{if } 0 \leq j \leq 25 \\ 2j - 51 & \text{if } 26 \leq j \leq 51 \end{cases}$$

The perfect riffle is deterministic, i.e. as a distribution  $\mu : S_{52} \rightarrow [0, 1]$ ,  $\mu = \delta_\sigma$ . It has order 8:

*Proof.* We seek minimal  $k$  s.t.  $\sigma^k = \text{id}$ . Since  $\sigma^k(j) = 2^k j \mod 51$  we require  $k$  s.t.  $2^k = 1 \mod 51$ . Checking  $k = 1, 2, \dots, 8$  shows 8 is the correct value.  $\square$

**Example 0.2. Random transposition** Randomly pick two cards and swap, with replacement after first choice (so doing nothing is possible). Has probability distribution  $\mu : S_{52} \rightarrow [0, 1]$ ,

$$\mu(\sigma) = \begin{cases} \sum^{52} \frac{1}{52^2} = \frac{1}{52} & \sigma = \text{id} \\ \frac{2}{52^2} & \sigma = (ij) \\ 0 & \text{otherwise} \end{cases}$$

**Example 0.3. Pass the broccoli (i.e.  $\mathbb{Z}_p$ )**  $\mathbb{Z}_p = \{0, \dots, p-1\}$ , the broccoli starts at 0, and at each step the broccoli is passed either left or right with probability  $\frac{1}{2}$ , generating random walk on  $\mathbb{Z}_p$ .

**Definition 0.4.** • **Sample space** is a set  $\Omega$  (i.e.  $\Omega = \mathbb{Z}_p$ ), and an **event** is element of the power set of the sample space:  $P(\Omega)$ .

- **Probability distribution** is a map  $\mu : \Omega \rightarrow [0, 1]$  such that  $\sum_{\omega \in \Omega} \mu(\omega) = 1$ .
- **Probability measure** is the extension of the distribution to the event space:  $\mu : P(\Omega) \rightarrow [0, 1]$ , such that for  $A \subset \Omega$ :  $\mu(A) := \sum_{\omega \in A} \mu(\omega)$ , with  $\mu(\emptyset) := 0$ . Its properties:
  - $\mu(\Omega) = 1$
  - $A \subset B \subset \mathbb{Z}_p \implies \mu(A) \leq \mu(B)$  “monotonicity”
  - $\{A_i\}$  countable pairwise disjoint subsets, then  $\mu(\cup_{k=1}^{\infty} A_k) = \sum_{k=1}^{\infty} \mu(A_k)$  “ $\sigma$ -additive”
  - $\mu(\mathbb{Z}_p) = 1$
- **Probability space** is triple  $(\Omega, P(\Omega), \mu)$  - Sample space, event space, probability measure.
- For a set  $S$ , an  **$S$ -valued random variable** is a map  $X : \Omega \rightarrow S$ .
- **Probability distribution wrt random variable  $X : \Omega \rightarrow S$  and probability space  $(\Omega, P(\Omega), \mu)$**  is the function  $\tilde{\mu} : X(\Omega) \rightarrow [0, 1]$  s.t.  $\tilde{\mu}(b) := \mu(X^{-1}(\{b\}))$ . This allows us to associate probabilities to “areas” in the image space of the random variable.

## 1 Probability on $\mathbb{Z}_p$

Course notes use  $t \oplus s := t + s \mod p$ ,  $t \ominus s := t - s \mod p$ . We use  $+$ ,  $-$  here.

**Definition 1.1. Probability distribution on  $\mathbb{Z}_p$**  is a map  $\mu : \mathbb{Z}_p \rightarrow [0, 1]$  s.t.  $\sum_{t=0}^{p-1} \mu(t) = 1$ .

**Example 1.2.** • **Uniform distribution** is  $\lambda : \mathbb{Z}_p \rightarrow [0, 1]$  s.t.  $\lambda(t) = 1/p \forall t \in \mathbb{Z}_p$ .

- **Dirac distribution** at  $s \in \mathbb{Z}_p$  is  $\delta_s : \mathbb{Z}_p \rightarrow [0, 1]$  s.t.  $\delta_s(t) := 1$  if  $t = s$ , 0 otherwise.
- **Fair Pass the broccoli distribution**

$$\mu(t) = \begin{cases} \frac{1}{2} & t = 1 \\ \frac{1}{2} & t = -1 \\ 0 & \text{otherwise} \end{cases}$$

**Definition 1.3.** The **expectation** of random var  $f : \mathbb{Z}_p \rightarrow \mathbb{C}$  wrt distribution  $\mu : \mathbb{Z}_p \rightarrow [0, 1]$  is:

$$\mathbb{E}_\mu(f) = \mu(f) := \sum_{t \in \mathbb{Z}_p} f(t)\mu(t) \in \mathbb{C}$$

Note the probability of event  $A \subset \Omega$  is the expectation of the indicator function (rv) over  $A$ :

$$\mathbf{1}_A(t) := \begin{cases} 1 & t \in A \\ 0 & t \notin A \end{cases}$$

**Theorem 1.4.** For distributions  $\mu_1, \dots, \mu_n$  and  $\alpha_1, \dots, \alpha_n \in [0, 1]$  s.t.  $\sum_i \alpha_i = 1$ , the **convex combination**  $\mu(t) := \sum_i \alpha_i \mu_i(t)$  is also a probability distribution.

**Definition 1.5.** **Biased pass the broccoli** is the convex combination of delta distributions: for  $\alpha \in (0, 1)$

$$\mu_\alpha = \alpha \delta_1 + (1 - \alpha) \delta_{-1}$$

**Definition 1.6.** • **Total variation distance** of dists  $\mu, \nu$  is:  $d(\mu, \nu) := \max_{A \subset \mathbb{Z}_p} |\mu(A) - \nu(A)|$

- **$L_1$ -norm** on the set of functions  $\{f : \mathbb{Z}_p \rightarrow \mathbb{R}\}$  is  $\|f\|_1 := \sum_{t \in \mathbb{Z}_p} |f(t)|$
- **$L^\infty$ -norm** on  $f : \mathbb{Z}_p \rightarrow \mathbb{R}$  is  $\|f\|_\infty = \max_{t \in \mathbb{Z}_p} |f(t)|$

**Lemma 1.7.** Total variation distance is a metric, i.e. satisfies triangle inequality:  $d(\mu, \nu) \leq d(\mu, \tau) + d(\tau, \nu)$ , symmetric:  $d(\mu, \nu) = d(\nu, \mu)$ , equality:  $d(\mu, \nu) = 0 \iff \mu = \nu$ .

**Theorem 1.8** (Total variation distance  $\equiv L_1$  norm).

$$d(\mu, \nu) = \frac{1}{2} \|\mu - \nu\|_1 = \frac{1}{2} \sum_{t \in \mathbb{Z}_p} |\mu(t) - \nu(t)|$$

**Theorem 1.9** (Variational formula). The tvd = the max difference in expectations over all rv's with max value ( $L_\infty$ ) at most 1:

$$d(\mu, \nu) = \frac{1}{2} \max\{ |\mu(f) - \nu(f)| \mid f : \mathbb{Z}_p \rightarrow \mathbb{R} \text{ s.t. } \|f\|_\infty \leq 1 \}$$

**Definition 1.10.** • **Information** of distribution  $\mu$  is  $I_\mu : \mathbb{Z}_p \rightarrow [0, \infty)$ ,  $I_\mu(t) := -\ln(\mu(t))$

- **Entropy** of distribution  $\mu$  is  $H(\mu) = -\sum_{t \in \mathbb{Z}_p} \mu(t) \ln(\mu(t))$  (i.e. the expected information).  
Note: the uniform and dirac distributions have entropies:

$$H(\lambda) = \ln(p), \quad H(\delta_s) = 0$$

**Theorem 1.11** (Pinsker's Inequality). For distribution  $\mu : \mathbb{Z}_p \rightarrow [0, 1]$  and uniform dist  $\lambda$

$$\frac{1}{2(H(\lambda) + 1)} |H(\mu) - H(\lambda)| \leq d(\mu, \lambda) \leq \sqrt{2|H(\mu) - H(\lambda)|}$$

## 2 Dynamics

**Definition 2.1.** The **convolution** of  $f, g : \mathbb{Z}_p \rightarrow [0, 1]$  is  $(f * g)(t) := \sum_{s \in \mathbb{Z}_p} f(t \ominus s)g(s)$ .

Iterated convolutions are denoted:  $\mu^{*n} := \mu^{*(n-1)} * \mu$ , with  $\mu^{*0} := \delta_0$

**Theorem 2.2.** • *Commutativity:*  $f * g = g * f$

- *Associative:*  $f * (g * h) = (f * g) * h$
- *Bilinear:*  $f * (\lambda g + \mu h) = \lambda(f * g) + \mu(f * h)$
- *for distributions  $\mu, \nu$ , then  $\mu * \nu$  is also a distribution.*
- *for distribution  $\mu$  and uniform dist  $\lambda$ , we have:  $\mu * \lambda = \lambda$*
- *for distribution  $\mu$  and dirac dist  $\delta_s$ , we have:  $(\delta_s * \mu)(t) = \mu(t - s)$ , hence:  $\delta_0 * \mu = \mu$ .*

**Example 2.3.** Fair pass the broccoli:  $(\mu * \mu)(t) = \frac{\delta_2(t)}{4} + \frac{\delta_0(t)}{2} + \frac{\delta_{-2}(t)}{4}$  (distribution after two steps).

**Theorem 2.4.** For distributions  $\mu, \nu : \mathbb{Z}_p \rightarrow [0, 1]$ , “entropy grows under convolution”:

$$\max\{H(\mu), H(\nu)\} \leq H(\mu * \nu) \leq H(\mu) + H(\nu)$$

**Definition 2.5.** • **Sumset** of  $A, B \subset \mathbb{Z}_p$  is  $A \oplus B := \{t + s : t \in A, s \in B\}$ .

- Iterated sumsets:  $A^{\oplus n} := A^{\oplus(n-1)} \oplus A$ , with  $A^{\oplus 0} := \emptyset$ .
- **Support** of distribution  $\mu$  is:  $\text{spt}(\mu) = \{t \in \mathbb{Z}_p : \mu(t) > 0\} \subset \mathbb{Z}_p$

Properties of sumsets:

- $\max\{|A|, |B|\} \leq |A \oplus B| \leq |A||B|$
- (Cauchy-Davenport inequality) If  $p$  prime, then:  $\min\{|A| + |B| - 1, p\} \leq |A \oplus B|$

**Theorem 2.6.** For distributions  $\mu, \nu$ ,  $\text{spt}(\mu * \nu) = \text{spt}(\mu) \oplus \text{spt}(\nu)$

**Definition 2.7.** A **random walk** on  $\mathbb{Z}_p$  with  $n$  steps is the  $\mathbb{Z}_p$ -valued random variable:

$$X_n := t_1 + \dots + t_n$$

for  $\mathbb{Z}_p$ -valued random variables  $t_1, \dots, t_n$  identically distributed wrt distribution  $\mu$ , so for each  $j$ :  $\mathbb{P}(t_j = t) = \mu(t)$ . By probability, probability distribution for the sum of random variables is the convolution of the distributions:  $\mathbb{P}(X_n = t) = \mu^{*n}(t)$ .

**Definition 2.8.**

$$\mathbb{P}(X_1 = s, X_n = t) := \mathbb{P}(X_1 = s)\mathbb{P}(X_n = t) = \mu(s)\mu^{*n}(t)$$

Note: here  $X_1$  and  $X_n$  are two **distinct** rw’s, so are independent, and therefore the probability of both events is the product of the probabilities of each event. Hence

$$P(X_n = t \mid X_1 = s) := \frac{\mathbb{P}(X_1 = s, X_n = t)}{\mathbb{P}(X_1 = s)} = \mathbb{P}(X_n = t)$$

This is stupid as we should have:  $P(X_n = t \mid X_1 = s) = \mathbb{P}(s + t_2 + \dots + t_n = t)$  (see below).

Note: The probability the  $n$ -th step is  $t \in \mathbb{Z}_p$  given the first step was  $s \in \mathbb{Z}_p$  is:

$$\mathbb{P}(s + t_2 + \cdots + t_n = t) = \delta_s * \mu^{*(n-1)}(t)$$

**Definition 2.9.** The **limit** of the sequence of distributions  $\mu_1, \mu_2, \dots : \mathbb{Z}_p \rightarrow [0, 1]$  is  $\mu_\infty : \mathbb{Z}_p \rightarrow [0, 1]$  such that:  $\lim_{n \rightarrow \infty} \mu_n(t) = \mu_\infty(t) \forall t \in \mathbb{Z}_p$ . In this case:  $\mu_\infty$  is also a distribution.

**Theorem 2.10** (Characterisation of limits).  $\mu_\infty$  is limit of  $\mu_1, \mu_2, \dots \iff \lim_{n \rightarrow \infty} d(\mu_n, \mu_\infty) = 0$

**Definition 2.11.** Distribution  $\mu$  is **ergodic** if  $\lim_{n \rightarrow \infty} \mu^{*n}(t) = \lambda(t)$ , for  $\lambda$  the uniform distribution.

**Lemma 2.12.** If  $A \subset \mathbb{Z}_p$  is not contained within a coset of a proper subgroup, then  $\exists n \in \mathbb{N}$  s.t.

$$A^{\oplus n} = \mathbb{Z}_p$$

**Theorem 2.13.** For distribution  $\mu$ , the support  $\text{spt}(\mu)$  is not contained within a coset of a proper subgroup  $\iff \exists n \in \mathbb{N}$  such that  $\text{spt}(\mu^{*n}) = \mathbb{Z}_p$ .

**Theorem 2.14** (Ergodic theorem).  $\mu$  is ergodic  $\iff \text{spt}(\mu)$  is not contained within a coset of a proper subgroup.

**Definition 2.15.** • The **mixing time**  $n_{\text{mix}}(\epsilon) \in \mathbb{N}$  of a random walk driven by distribution  $\mu$ , with a **threshold**  $\epsilon > 0$ , is such that:  $d(\nu * \mu^{*n}, \lambda) < \epsilon \forall n \geq n_{\text{mix}}(\epsilon)$ ,  $\forall$  starting dists  $\nu$ .

- For  $\phi : \mathbb{N} \rightarrow [0, \infty)$  s.t.  $\lim_{n \rightarrow \infty} \phi(n) = 0$ ,  $\mu$  is **mixing with rate**  $\phi$  if

$$d(\nu * \mu^{*n}, \lambda) \leq \phi(n) \forall n, \nu$$

- $\mu$  is **exponentially mixing** if  $\exists C \in (0, \infty), \theta \in [0, 1)$  such that  $\mu$  is mixing with a rate function  $\phi$  that is exponentially decaying:  $\phi(n) \leq C\theta^n$ .

### 3 Harmonic Analysis

**Definition 3.1.** • The **discrete fourier transform** of  $f : \mathbb{Z}_p \rightarrow \mathbb{C}$  is  $\hat{f} : \mathbb{Z}_p \rightarrow \mathbb{C}$  such that:

$$\hat{f}(k) := \sum_{t=0}^{p-1} f(t) e^{\frac{-2\pi i k t}{p}}$$

- The maps  $\chi_k(t) := e^{\frac{-2\pi i k t}{p}}$  are **characters** of  $\mathbb{Z}_p$  (i.e. group homomorphism  $\mathbb{Z}_p \rightarrow \mathbb{C}$ ).

**Lemma 3.2.** 1.  $\hat{\mu}(0) = 1$  (by definition of  $\mu$  being a distribution).

2.  $\hat{\mu}(k) \leq 1 \forall k$

3. (Exponential sum formula) For  $\theta \neq 0$ ,  $\sum_{t=0}^{p-1} e^{it\theta} = \frac{1 - e^{ip\theta}}{1 - e^{i\theta}}$

*Proof.* Pt 2:  $|\hat{\mu}(k)| = |\sum_t \mu(t) e^{\frac{-2\pi i k t}{p}}| \leq \sum_t |\mu(t) e^{\frac{-2\pi i k t}{p}}| = \sum_t |\mu(t)| \cdot |e^{\frac{-2\pi i k t}{p}}| = \sum_t |\mu(t)| = 1 \quad \square$

**Example 3.3.** Fourier transform of distributions. The moral is the more spread out  $\hat{\mu}$  is, the more confined  $\mu$  is, and vice versa.

- $\hat{\lambda} = \delta_0$

- $\hat{\delta}_0 = \lambda$ . So  $\hat{\delta}_0(k) = 1 \forall k$ . More generally:  $\hat{\delta}_s(k) = e^{\frac{-2\pi i k s}{p}}$
- $\mu = \frac{1}{2}\delta_{-1} + \frac{1}{2}\delta_1$ , then  $\hat{\mu}(k) = \frac{1}{2}e^{\frac{-2\pi i k}{p}} + \frac{1}{2}e^{\frac{2\pi i k}{p}} = \cos(\frac{2\pi k}{p})$

**Theorem 3.4** (Fourier series theorem). *Every  $f : \mathbb{Z}_p \rightarrow \mathbb{C}$  has **fourier expansion/inverse FT**:*

$$f(t) = \frac{1}{p} \sum_{k=0}^{p-1} \hat{f}(k) e^{\frac{2\pi i k t}{p}}$$

**Definition 3.5.** • The **inner product** of  $f, g : \mathbb{Z}_p \rightarrow \mathbb{C}$  is  $\langle f, g \rangle := \sum_t f(t) \overline{g(t)}$

- The  **$L_2$ -norm** is  $\|f\|_2 := \sqrt{\langle f, f \rangle}$
- For  $1 < p < \infty$ , the  **$L_p$ -norm** is  $\|f\|_p = (\sum_t |f(t)|^p)^{\frac{1}{p}}$

**Lemma 3.6.** • Characters  $\chi_k$  are orthonormal, i.e.  $\langle \chi_k, \chi_l \rangle = \begin{cases} 1 & \text{if } k = l, \\ 0 & \text{otherwise.} \end{cases}$

- (Cauchy-Schwartz Inequality)  $\forall f, g: |\langle f, g \rangle| \leq \|f\|_2 \|g\|_2$
- (Holders Inequality) For  $1 < p, q < \infty$  s.t.  $\frac{1}{p} + \frac{1}{q} = 1$ , then  $|\langle f, g \rangle| \leq \|f\|_p \|g\|_q$

**Theorem 3.7** (Plancherels theorem/Parsevals identity).

$$\langle f, g \rangle = \frac{1}{p} \langle \hat{f}, \hat{g} \rangle \quad \|f\|_2 = \frac{1}{\sqrt{p}} \|\hat{f}\|_2$$

**Theorem 3.8** (Convolution theorem).

$$\widehat{f * g} = \hat{f} \hat{g}$$

## 4 Mixing Time

**Theorem 4.1.** • (Upper Bound Lemma) For distribution  $\mu$ , then  $\forall n \in \mathbb{N}$

$$d(\mu^{*n}, \lambda) \leq \frac{1}{2} \sqrt{\sum_{k \in \mathbb{Z}_p \setminus \{0\}} |\hat{\mu}(k)|^{2n}}$$

- Generalisation:

$$d(\mu_1 * \dots * \mu_n, \lambda) \leq \frac{1}{2} \sqrt{\sum_{k \in \mathbb{Z}_p \setminus \{0\}} \prod_{j=1}^n |\hat{\mu}_j(k)|^2}$$

- (Lower Bound Lemma)  $\forall n \in \mathbb{N}$

$$d(\mu^{*n}, \lambda) \geq \frac{1}{2} \sqrt{\frac{1}{p} \sum_{k \in \mathbb{Z}_p \setminus \{0\}} |\hat{\mu}(k)|^{2n}}$$

- (Entropy)

$$H(\mu^{*n}) \geq \ln(p) - (\ln(p) + 1) \sqrt{\sum_{k \in \mathbb{Z}_p \setminus \{0\}} |\hat{\mu}(k)|^{2n}}$$

**Definition 4.2.** Distribution  $\mu$  has a **spectral gap** if  $|\hat{\mu}(k)| < 1 \forall k \in \mathbb{Z}_p \setminus \{0\}$ .

**Theorem 4.3.** • Distribution  $\mu$  has a spectral gap  $\implies$  it is exponentially mixing.

- Distribution  $\mu$  has a spectral gap  $\iff$  it is ergodic.

## 5 Beyond $\mathbb{Z}_p$

For general group  $G$ , probability distributions, total variation distance,  $L_1$ -norms, ergodicity are defined identically. A random walks are denoted  $X_n = a_1 \dots a_n$ , as products rather sums since  $G$  not necessarily abelian.

**Definition 5.1.** For distributions  $f, g$  over  $G$  (the following coincide for abelian  $G$ ):

- The **left convolution**  $(f *_L g)(a) := \sum_{b \in G} f^{-1}(b^{-1}a)g(b)$
- The **right convolution** is  $(f *_R g)(a) := \sum_{b \in G} f^{-1}(ab^{-1})g(b)$

**Theorem 5.2.**  $\mu$  is ergodic  $\iff$  the support of  $\text{spt}(\mu)$  is not contained within a coset of a proper subgroup. (Result as above, but requires new proof).

**Definition 5.3.**  $\mathbb{Z}_p^d = \{(t_1, \dots, t_d) : t_j \in \mathbb{Z}_p\}$  is a vector space, so group over addition.

**Example 5.4.** Ehrenfests Urn model is made up of  $d$  balls in 2 urns, whose states are modelled via a vector  $v \in \mathbb{Z}_2^d$ , one coordinate per ball, with  $v_j = 0$  if  $j$ -th ball is in the left urn,  $v_j = 1$  if in the right urn. Each move selects a ball/coordinate randomly and swaps its urn/value, so is equivalent to adding a standard basis vec  $e_j \in \mathbb{Z}_2^d$ . Hence the distribution of possible moves is:

$$\mu(t') = \begin{cases} \frac{1}{d} & t' = e_j \text{ for some } 1 \leq j \leq n \\ 0 & \text{otherwise} \end{cases} \quad (1)$$

The distribution after one move, starting at configuratio  $t \in \mathbb{Z}_2^d$  is then  $\mu * \delta_t$ , and after  $n$  moves:  $\mu^{*n} * \delta_t$ .

**Definition 5.5.** The **fourier transform** of  $f : \mathbb{Z}_2^d \rightarrow \mathbb{C}$  is  $\hat{f} : \mathbb{Z}_2^d \rightarrow \mathbb{C}$  s.t.

$$\hat{f}(k) := \sum_{t \in \mathbb{Z}_2^d} f(t)(-1)^{k \cdot t}$$

With this definition, all the Harmonic analysis such as the Plancheral/Parseval/Convolution theorems for  $\mathbb{Z}_p$  hold for  $\mathbb{Z}_2^d$  too.

**Theorem 5.6** (Upper Bound Lemma for  $\mathbb{Z}_2^d$ ).  $\forall n \in \mathbb{N}$

$$d(\mu^{*n}, \lambda) \leq \frac{1}{2} \sqrt{\sum_{k \in \mathbb{Z}_2^d \setminus \{0\}} |\hat{\mu}(k)|^{2n}}$$

**Lemma 5.7.** The fourier transform of the distribution of moves in Ehrenfests Urn model (in Equation (1)) is:  $\hat{\mu}(k) = 1 - \frac{2}{d} |\{1 \leq j \leq d : k_j = 1\}|$ .