

Introduction

- Scientists express hypotheses through complex and expensive simulators.
- Posterior distributions give insight into models for both understanding underlying phenomena and improving hypotheses.
- Approximate Bayesian Computation provides a Bayesian framework for posterior analysis, but is very inefficient.
- This work uses a surrogate of the simulator to speed-up ABC.
- Based on the *Metropolis-Hastings Error*, our algorithms determine when a simulation is necessary and provides the user with a “knob” ξ to control it.

Approximate Bayesian Computation

- ABC is a *likelihood-free* method because $\pi(\mathbf{y}|\boldsymbol{\theta})$ is either not computable or very expensive:

$$\pi_{\epsilon}(\boldsymbol{\theta}|\mathbf{y}) = \frac{\pi(\boldsymbol{\theta})}{\pi(\mathbf{y})} \int \pi_{\epsilon}(\mathbf{y}|\mathbf{x}) \pi(\mathbf{x}|\boldsymbol{\theta}) d\mathbf{x}$$

- Kernel functions $\pi_{\epsilon}(\mathbf{y}|\mathbf{x})$ are proxies for the likelihood, based on draws $\mathbf{x} \stackrel{\text{sim}}{\sim} \pi(\mathbf{x}|\boldsymbol{\theta})$ from simulator.
- Rejection sampling with ϵ -tube kernel is very inefficient.
- We use ABC MCMC, which approximates the likelihood by Monte Carlo approximation:

$$\pi_{\epsilon}(\mathbf{y}|\boldsymbol{\theta}') \approx \frac{1}{S} \sum_{s=1}^S \pi_{\epsilon}(\mathbf{y}|\mathbf{x}^{(s)}, \boldsymbol{\theta}')$$

We accept the proposed parameter $\boldsymbol{\theta}'$ with probability equal to $\alpha(\boldsymbol{\theta}'|\boldsymbol{\theta}) =$

$$\min \left(1, \frac{\pi(\boldsymbol{\theta}') \sum_s \pi_{\epsilon}(\mathbf{y}|\mathbf{x}^{(s)}, \boldsymbol{\theta}') q(\boldsymbol{\theta}|\boldsymbol{\theta}')}{\pi(\boldsymbol{\theta}) \sum_s \pi_{\epsilon}(\mathbf{y}|\mathbf{x}^{(s)}, \boldsymbol{\theta}) q(\boldsymbol{\theta}'|\boldsymbol{\theta})} \right)$$

The Synthetic Likelihood

- Introduced by Wood (2010), replace the Monte Carlo approximation with a Gaussian with estimators based on the pseudo-data $\{\mathbf{x}_1, \dots, \mathbf{x}_S\}$ simulated at $\boldsymbol{\theta}$:

$$\hat{\boldsymbol{\mu}}_{\boldsymbol{\theta}} = \frac{1}{S} \sum_s \mathbf{x}^{(s)} \quad \hat{\boldsymbol{\Sigma}}_{\boldsymbol{\theta}} = \frac{1}{S-1} \sum_s \left(\mathbf{x}^{(s)} - \hat{\boldsymbol{\mu}}_{\boldsymbol{\theta}} \right) \left(\mathbf{x}^{(s)} - \hat{\boldsymbol{\mu}}_{\boldsymbol{\theta}} \right)^T$$

- With a Gaussian kernel:

$$\pi_{\epsilon}(\mathbf{y}|\mathbf{x}) = K_{\epsilon}(\mathbf{y}, \mathbf{x}) = \frac{1}{(2\pi\epsilon)^{J/2}} e^{-\frac{1}{2\epsilon}(\mathbf{x}-\mathbf{y})^T(\mathbf{x}-\mathbf{y})}$$

- The synthetic likelihood can then be computed analytically:

$$\pi(\mathbf{y}|\boldsymbol{\theta}) = \int K_{\epsilon}(\mathbf{y}, \mathbf{x}) \mathcal{N}(\hat{\boldsymbol{\mu}}_{\boldsymbol{\theta}}, \hat{\boldsymbol{\Sigma}}_{\boldsymbol{\theta}}) d\mathbf{x} = \mathcal{N}(\hat{\boldsymbol{\mu}}_{\boldsymbol{\theta}}, \hat{\boldsymbol{\Sigma}}_{\boldsymbol{\theta}} + \epsilon^2 \mathbf{I})$$

MCMC with a Random Acceptance Probability

- The randomness of $\hat{\boldsymbol{\mu}}_{\boldsymbol{\theta}}$ (by re-running S simulations), induces a distribution over the acceptance probabilities.
- Approximate randomness in $\boldsymbol{\mu}_{\boldsymbol{\theta}}$ by drawing M times $\boldsymbol{\mu}_{\boldsymbol{\theta}}^{(m)} \sim \mathcal{N}(\hat{\boldsymbol{\mu}}_{\boldsymbol{\theta}}, \hat{\boldsymbol{\Sigma}}_{\boldsymbol{\theta}}/S)$
- Estimate $p(\alpha)$ using Monte Carlo approximation based on these M samples:

$$\alpha^{(m)} = \min \left(1, \frac{\pi(\boldsymbol{\theta}') \mathcal{N}(\mathbf{y}|\boldsymbol{\mu}_{\boldsymbol{\theta}'}^{(m)}, \hat{\boldsymbol{\Sigma}}_{\boldsymbol{\theta}'} + \epsilon^2 \mathbf{I}) q(\boldsymbol{\theta}|\boldsymbol{\theta}')}{\pi(\boldsymbol{\theta}) \mathcal{N}(\mathbf{y}|\boldsymbol{\mu}_{\boldsymbol{\theta}}^{(m)}, \hat{\boldsymbol{\Sigma}}_{\boldsymbol{\theta}} + \epsilon^2 \mathbf{I}) q(\boldsymbol{\theta}'|\boldsymbol{\theta})} \right)$$

- Based on $p(\alpha)$, compute the probability of making a M-H Error. This is the area under the folded CDF and is minimized at the median of $p(\alpha)$, denoted τ .

Probability of acceptance error Probability of rejection error

$$P(\alpha < u) = \frac{1}{M} \sum_m [\alpha^{(m)} < u] \quad P(\alpha > u) = \frac{1}{M} \sum_m [\alpha^{(m)} \geq u]$$

- Total *conditional* error: $\mathcal{E}_u(\alpha) = [u \leq \tau] P(\alpha < u) + [u > \tau] P(\alpha \geq u)$
- MHE: $\mathcal{E}(\alpha) = \int \mathcal{E}_u(\alpha) \mathcal{U}(0, 1) du$
- The crux of the MH step is to run simulations while $\mathcal{E}(\alpha) > \xi$.

Gaussian Process Surrogate ABC

- Adaptive SL is wasteful for expensive simulations: all results are discarded.
- GPS-ABC follow directly from synthetic likelihood ABC with randomized acceptance.
- Gaussian processes provides uncertainty estimates of the marginal likelihood informing us of the need to conduct additional experiments in order to make confident accept/reject decisions.
- For each statistic j , the surrogate provides the following conditional predictive distribution of the expected value of statistic j based on N training points $\boldsymbol{\mu}_{\boldsymbol{\theta}_j} \sim \mathcal{N}(\bar{\boldsymbol{\mu}}_{\boldsymbol{\theta}_j}, \sigma_{\boldsymbol{\theta}_j}^2)$

$$\bar{\boldsymbol{\mu}}_{\boldsymbol{\theta}_j} = \mathbf{k}_{\boldsymbol{\theta}\boldsymbol{\theta}_j} [\mathbf{K}_{\boldsymbol{\theta}\boldsymbol{\theta}_j} + \sigma_j^2 \mathbf{I}]^{-1} \mathbf{X}[:, j]$$

$$\sigma_{\boldsymbol{\theta}_j}^2 = k_{\boldsymbol{\theta}\boldsymbol{\theta}_j} - \mathbf{k}_{\boldsymbol{\theta}\boldsymbol{\theta}_j} [\mathbf{K}_{\boldsymbol{\theta}\boldsymbol{\theta}_j} + \sigma_j^2 \mathbf{I}]^{-1} \mathbf{k}_{\boldsymbol{\theta}\boldsymbol{\theta}_j}$$

- Adjusting ξ affects precision and computation:

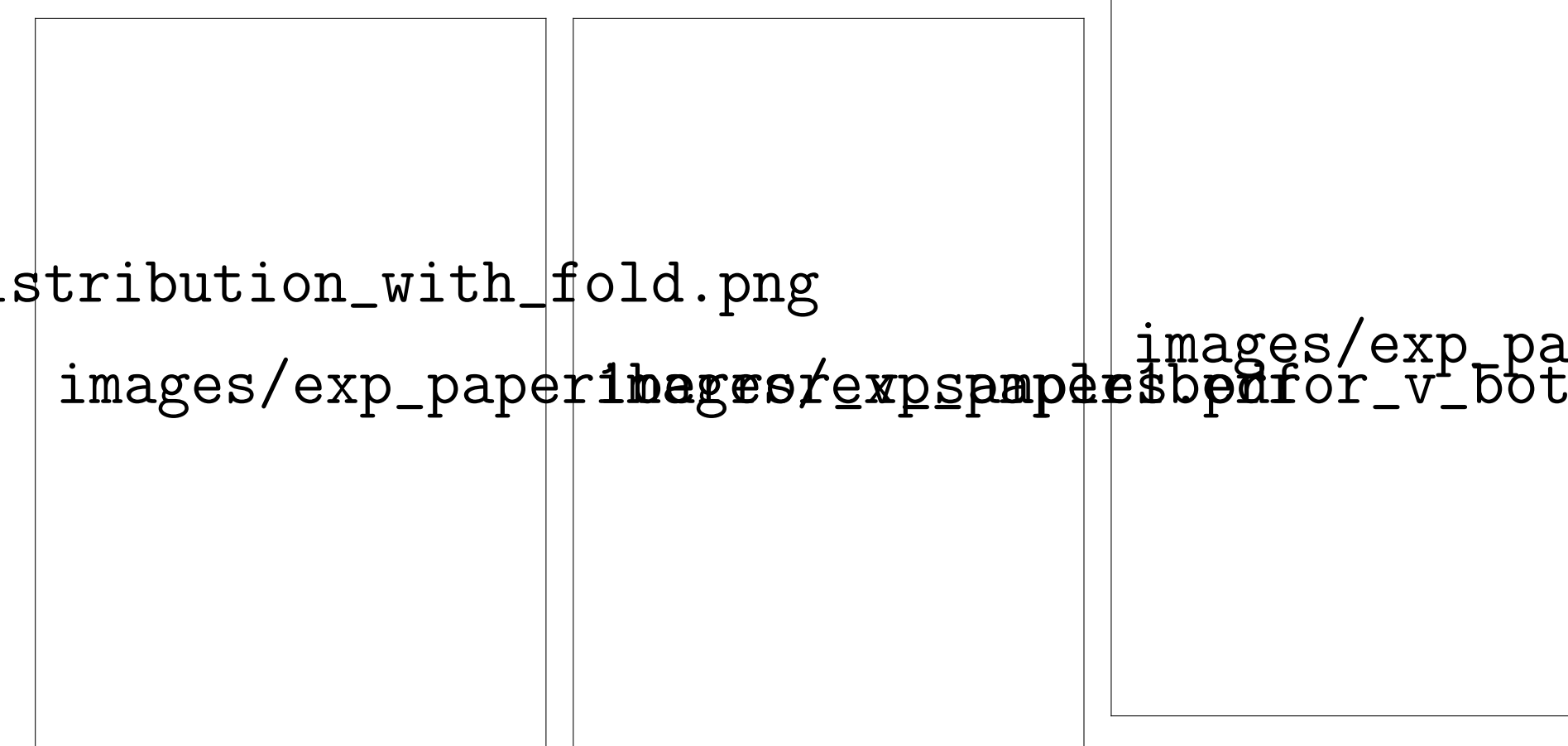
$$\xi = 0.4, N = 9 \quad \xi = 0.2, N = 184 \quad \xi = 0.05, N = 1297$$

high model uncertainty. model uncertainty drops with ξ high precision, many simulations

- The key advantage of GPS-ABC is that once we have trained the GP surrogate, we will not have to do any expensive simulations whatsoever during a MH step because the GP surrogate is sufficiently confident about the statistics' surface in that region of parameter space.

Toy problem: Inferring parameters of an exponential distribution

- GPS-ABC learns the surface, eventually eliminating any new simulations:



Chaotic Ecological Systems

- Adult blowfly populations exhibit dynamic, sometimes chaotic, behavior for which several competing population models exist.
- We use observational data and a simulation model from Wood (2010).
- Population dynamics are modeled using (discretized) differential equations that can produce chaotic behavior for some parameter settings.
- The population dynamics equation generates N_1, \dots, N_T : $N_{t+1} = PN_{t-\tau} \exp(-N_{t-\tau}/N_0) + N_t \exp(-\delta\epsilon_t)$
- Injected noise: $e_t \sim \mathcal{G}(1/\sigma_p^2, 1/\sigma_p^2)$, $\epsilon_t \sim \mathcal{G}(1/\sigma_d^2, 1/\sigma_d^2)$
- $\boldsymbol{\theta}$: $\{\log P, \log \delta, \log N_0, \log \sigma_d, \log \sigma_p, \tau\}$
- Statistics \mathbf{y} ($J = 10$): the log of the mean of all 25% quantiles of $N/1000$, the mean of the 25% quantiles of the first-order differences of $N/1000$, and the maximal peaks of smoothed N , with 2
- Models: rejection sampling ($\epsilon = 0.5$), SL ($S = 10$), GPS-ABC ($\xi = 0.3$)
- Results: generated observations (top), posterior samples $\boldsymbol{\theta}$ (middle), convergence to \mathbf{y}^* using posterior predictive $p(\mathbf{y}|\mathbf{y}^*)$ (bottom).

Conclusions

- Adaptive ABC algorithms using Metropolis-Hastings Error controls the computational complexity of the inference
- GP surrogate models incorporate both model and pseudo-uncertainty into MCMC
- Improvements: Hamiltonian dynamics on the GP surface dependent samples; covariance on the outputs; alternative gate models; acquisition functions.