

# Tutoriat 10 - Exam training - Seria 23

## Contents

<b>1 Exam from 2020-2021</b>	<b>1</b>
<b>2 Exam from 2022-2023</b>	<b>1</b>
2.1 Probabilities . . . . .	1
2.1.1 Exercise 1 . . . . .	1
2.1.2 Exercise 3 . . . . .	3
2.2 Statistics . . . . .	3
2.2.1 Exercise 4 . . . . .	3
2.2.2 Exercise 5: . . . . .	4
2.3 Exercise 6: . . . . .	5
2.4 Exercise 7 (Extra) . . . . .	5

## 1 Exam from 2020-2021

Already solved [here](#)

## 2 Exam from 2022-2023

### 2.1 Probabilities

#### 2.1.1 Exercise 1

First exercise from [here!](#)

## Exercise 2

Let  $X$  and  $Y$  be random variables with joint support

$$(x, y) \in [0, i] \times [0, i],$$

and joint pdf

$$f_{X,Y}(x, y) = cxy^2.$$

#### (a) Determination of the constant $c$

The normalization condition gives

$$\int_0^i \int_0^i cxy^2 dx dy = 1.$$

First,

$$\int_0^i x dx = \frac{i^2}{2}, \quad \int_0^i y^2 dy = \frac{i^3}{3}.$$

Thus,

$$c \cdot \frac{i^2}{2} \cdot \frac{i^3}{3} = 1 \implies \boxed{c = \frac{6}{i^5}}.$$

**(b) Marginal pdf of  $X$**

$$f_X(x) = \int_0^i f_{X,Y}(x,y) dy = \frac{6x}{i^5} \int_0^i y^2 dy = \boxed{\frac{2x}{i^2}}, \quad 0 \leq x \leq i.$$

**(c) Marginal CDF of  $Y$**

First, compute the marginal pdf of  $Y$ :

$$f_Y(y) = \int_0^i f_{X,Y}(x,y) dx = \frac{6y^2}{i^5} \int_0^i x dx = \frac{3y^2}{i^3}, \quad 0 \leq y \leq i.$$

Then, the marginal CDF is

$$F_Y(y) = \int_0^y f_Y(t) dt = \int_0^y \frac{3t^2}{i^3} dt = \boxed{\frac{y^3}{i^3}}, \quad 0 \leq y \leq i.$$

**(d) Probability that  $X + Y < i$  (Extra)**

We compute

$$\mathbb{P}(X + Y < i) = \int_0^i \int_0^{i-x} f_{X,Y}(x,y) dy dx.$$

Substitute the joint pdf:

$$\mathbb{P}(X + Y < i) = \int_0^i \int_0^{i-x} \frac{6}{i^5} xy^2 dy dx.$$

First integrate with respect to  $y$ :

$$\int_0^{i-x} y^2 dy = \frac{(i-x)^3}{3}.$$

Thus,

$$\mathbb{P}(X + Y < i) = \frac{6}{i^5} \int_0^i x \cdot \frac{(i-x)^3}{3} dx = \frac{2}{i^5} \int_0^i x(i-x)^3 dx.$$

Expand the integrand:

$$x(i-x)^3 = x(i^3 - 3i^2x + 3ix^2 - x^3) = i^3x - 3i^2x^2 + 3ix^3 - x^4.$$

Integrate term by term:

$$\begin{aligned} \int_0^i i^3 x dx &= \frac{i^5}{2}, & \int_0^i 3i^2 x^2 dx &= i^5, \\ \int_0^i 3ix^3 dx &= \frac{3i^5}{4}, & \int_0^i x^4 dx &= \frac{i^5}{5}. \end{aligned}$$

Putting everything together:

$$\int_0^i x(i-x)^3 dx = i^5 \left( \frac{1}{2} - 1 + \frac{3}{4} - \frac{1}{5} \right) = \frac{i^5}{20}.$$

Therefore,

$$\boxed{\mathbb{P}(X + Y < i) = \frac{2}{i^5} \cdot \frac{i^5}{20} = \frac{1}{10}.}$$

### (e) Checking independence of $X$ and $Y$ (Extra)

Two random variables are independent if and only if

$$f_{X,Y}(x,y) = f_X(x)f_Y(y).$$

From earlier results,

$$f_X(x) = \frac{2x}{i^2}, \quad f_Y(y) = \frac{3y^2}{i^3}.$$

Their product is

$$f_X(x)f_Y(y) = \frac{2x}{i^2} \cdot \frac{3y^2}{i^3} = \frac{6}{i^5}xy^2.$$

Since

$$f_{X,Y}(x,y) = f_X(x)f_Y(y),$$

we conclude that

$X$  and  $Y$  are independent.

### 2.1.2 Exercise 3

Let  $X$  be a discrete random variable with distribution

$$X \sim \begin{pmatrix} -1 & 0 & i \\ \frac{1}{3} & \frac{1}{3} & \frac{1}{3} \end{pmatrix}.$$

**Expectation**

$$\mathbb{E}[X] = (-1)\frac{1}{3} + 0 + i\frac{1}{3} = \frac{i-1}{3}.$$

**Second Moment**

$$\mathbb{E}[X^2] = \frac{1+i^2}{3}.$$

**Variance**

$$\text{Var}(X) = \mathbb{E}[X^2] - (\mathbb{E}[X])^2 = \boxed{\frac{2(i^2 + i + 1)}{9}}.$$

## 2.2 Statistics

### 2.2.1 Exercise 4

Calculate the odds of choosing a natural number between 0 and 20 that is divisible by 3.

**Solution**

Let

$$\Omega = \{1, 2, \dots, 20\}.$$

Let  $A$  be the event that the chosen number is divisible by 3.

The numbers divisible by 3 in this interval are:

$$3, 6, 9, 12, 15, 18,$$

so

$$|A| = 6, \quad |\Omega| = 20.$$

Hence,

$$P(A) = \frac{6}{20} = \frac{3}{10}, \quad P(A^c) = 1 - P(A) = \frac{7}{10}.$$

The odds in favor of  $A$  are defined as

$$O(A) = \frac{P(A)}{P(A^c)}.$$

Therefore,

$$O(A) = \frac{3/10}{7/10} = \frac{3}{7}.$$

### 2.2.2 Exercise 5:

**Given** We observe a single data point

$$x = i,$$

generated from a normal distribution with unknown mean  $\theta$  and unit variance:

$$x | \theta \sim \mathcal{N}(\theta, 1).$$

The prior distribution for  $\theta$  is

$$\theta \sim \mathcal{N}(50, 1).$$

Estimate the parameter  $\theta$

### Solution

**Likelihood** The likelihood function is

$$f(x | \theta) = c_1 \exp\left(-\frac{1}{2}(x - \theta)^2\right), \quad c_1 = \frac{1}{\sqrt{2\pi}}.$$

Substituting  $x = i$ :

$$f(i | \theta) = c_1 \exp\left(-\frac{1}{2}(i - \theta)^2\right).$$

**Prior** The prior density is

$$f(\theta) = c_2 \exp\left(-\frac{1}{2}(\theta - 50)^2\right), \quad c_2 = \frac{1}{\sqrt{2\pi}}.$$

**Bayes' Theorem** Applying Bayes' theorem:

$$f(\theta | i) = \frac{f(i | \theta)f(\theta)}{f(i)}.$$

Let

$$c_3 = \frac{c_1 c_2}{f(i)}.$$

Then

$$f(\theta | i) = c_3 \exp\left(-\frac{1}{2}(i - \theta)^2 - \frac{1}{2}(\theta - 50)^2\right).$$

**Completing the Square** Expanding and simplifying the exponent gives

$$f(\theta | i) = c_4 \exp\left(-\frac{1}{2}\left(\theta - \frac{i+50}{2}\right)^2\right).$$

**Posterior Distribution** Therefore, the posterior distribution is

$$\theta | i \sim \mathcal{N}\left(\frac{i+50}{2}, \frac{1}{2}\right)$$

The posterior mean is

$$\mathbb{E}[\theta | i] = \frac{i+50}{2}.$$

### 2.3 Exercise 6:

Let

$$X \sim \mathcal{N}(i, 1).$$

#### (a) Symmetric $\frac{1}{2}$ -Probability Interval for $X$

A symmetric probability interval  $(-a, a)$  for a standard normal random variable  $Z \sim \mathcal{N}(0, 1)$  satisfies

$$P(-a \leq Z \leq a) = \frac{1}{2}.$$

From the standard normal table,

$$a \approx 0.674.$$

The standardization of  $X$  is

$$Z = \frac{X - i}{1} = X - i.$$

Thus,

$$P(i - a \leq X \leq i + a) = \frac{1}{2}.$$

Therefore, the symmetric  $\frac{1}{2}$ -probability interval for  $X$  is

$$\boxed{(i - 0.674, i + 0.674)}.$$

#### (b) Symmetric $\frac{1}{2}$ -Probability Interval for $Y = 3X - 5$ (Extra)

Define the random variable

$$Y = 3X - 5.$$

Since  $X \sim \mathcal{N}(i, 1)$ , it follows that

$$Y \sim \mathcal{N}(3i - 5, 9).$$

Using the linear transformation of intervals, the symmetric  $\frac{1}{2}$ -probability interval for  $Y$  is obtained by applying the transformation  $y = 3x - 5$  to the interval for  $X$ .

Thus,

$$Y \in (3(i - 0.674) - 5, 3(i + 0.674) - 5).$$

Therefore, the required interval is

$$\boxed{(3i - 7.022, 3i - 2.978)}.$$

### 2.4 Exercise 7 (Extra)

Let  $X_1$  and  $X_2$  be independent observations from a normal distribution

$$X_i \sim \mathcal{N}(\mu, \sigma^2), \quad i = 1, 2,$$

where both  $\mu$  and  $\sigma^2$  are unknown.

The observed values are

$$x_1 = 4, \quad x_2 = 8.$$

Estimate the parameters  $\mu$  and  $\sigma$  using the method of maximum likelihood.

---

# Solution

## Likelihood Function

The likelihood function is

$$L(\mu, \sigma^2) = \prod_{i=1}^2 \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x_i - \mu)^2}{2\sigma^2}\right).$$

Taking the logarithm:

$$\ell(\mu, \sigma^2) = -\ln(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^2 (x_i - \mu)^2.$$

—

## MLE of the Mean

Differentiate with respect to  $\mu$  and set equal to zero:

$$\frac{\partial \ell}{\partial \mu} = \frac{1}{\sigma^2} \sum_{i=1}^2 (x_i - \mu) = 0.$$

Thus,

$$\hat{\mu}_{\text{MLE}} = \frac{x_1 + x_2}{2} = \frac{4 + 8}{2} = \boxed{6}.$$

—

## MLE of the Variance

Differentiate with respect to  $\sigma^2$  and set equal to zero:

$$\frac{\partial \ell}{\partial \sigma^2} = -\frac{1}{\sigma^2} + \frac{1}{2(\sigma^2)^2} \sum_{i=1}^2 (x_i - \mu)^2 = 0.$$

Substitute  $\hat{\mu}_{\text{MLE}} = 6$ :

$$\sum_{i=1}^2 (x_i - 6)^2 = (4 - 6)^2 + (8 - 6)^2 = 4 + 4 = 8.$$

Thus,

$$\hat{\sigma}_{\text{MLE}}^2 = \frac{1}{2} \cdot 8 = 4.$$

Therefore,

$$\boxed{\hat{\sigma}_{\text{MLE}} = 2.}$$

—

## Final Estimates

$$\boxed{\hat{\mu}_{\text{MLE}} = 6, \quad \hat{\sigma}_{\text{MLE}} = 2.}$$