

Random Numbers

Abstract—This manual provides a simple introduction to the generation of random numbers

1 UNIFORM RANDOM NUMBERS

Let U be a uniform random variable between 0 and 1.

- 1.1 Generate 10^6 samples of U using a C program and save into a file called uni.dat .

Solution:

```
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/exrand.c
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/coeffs.h
```

- 1.2 Load the uni.dat file into python and plot the empirical CDF of U using the samples in uni.dat. The CDF is defined as

$$F_U(x) = \Pr(U \leq x) \quad (1.1)$$

Solution: The following code plots Fig. 1.2

```
https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/cdf_plot.py
```

- 1.3 Find a theoretical expression for $F_U(x)$.

Solution:

$$f_U(x) = \begin{cases} 1, & x \in (0, 1) \\ 0, & \text{otherwise} \end{cases} \quad (1.2)$$

$$F_U(x) = \int_{-\infty}^x f_U(x) dx \quad (1.3)$$

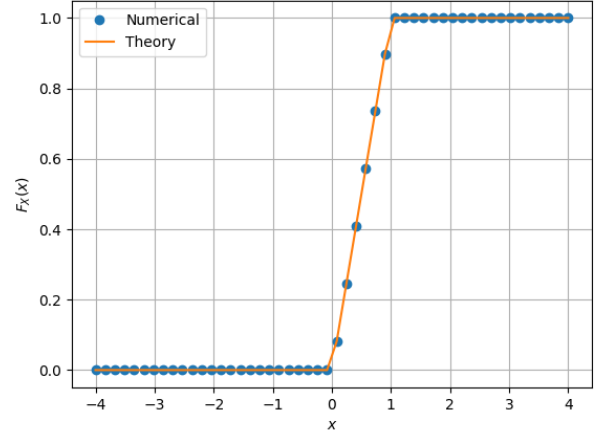


Fig. 1.2: The CDF of U

Hence, If $x \leq 0$,

$$F_U(x) = \int_{-\infty}^x f_U(x) dx \quad (1.4)$$

$$F_U(x) = \int_{-\infty}^x 0 dx \quad (1.5)$$

$$= 0 \quad (1.6)$$

If $0 < x < 1$,

$$F_U(x) = \int_0^x f_U(x) dx \quad (1.7)$$

$$F_U(x) = \int_0^x 1 dx \quad (1.8)$$

$$= x \quad (1.9)$$

If $x \geq 1$,

$$F_U(x) = \int_1^x f_U(x) dx \quad (1.10)$$

$$F_U(x) = \int_1^x 0 dx \quad (1.11)$$

$$= 0 \quad (1.12)$$

Hence,

$$F_U(x) = \begin{cases} x, & x \in (0, 1) \\ 0, & \text{otherwise} \end{cases} \quad (1.13)$$

1.4 The mean of U is defined as

$$E[U] = \frac{1}{N} \sum_{i=1}^N U_i \quad (1.14)$$

and its variance as

$$\text{var}[U] = E[U - E[U]]^2 \quad (1.15)$$

Write a C program to find the mean and variance of U .

Solution:

```
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/mean_variance.c
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/coeffs.h
```

$$\text{Mean obtained} = 0.500007 \quad (1.16)$$

$$\text{Variance obtained} = 0.083301 \quad (1.17)$$

1.5 Verify your result theoretically given that

$$E[U^k] = \int_{-\infty}^{\infty} x^k dF_U(x) \quad (1.18)$$

Solution:

$$E[U^k] = 0 + \int_0^1 x^k dF_U(x) + 0 \quad (1.19)$$

$$= \frac{1}{k+1} \quad (1.20)$$

$$\Rightarrow E[U] = \frac{1}{2} \quad (1.21)$$

$$= 0.50 \quad (1.22)$$

$$E[U^2] = \frac{1}{3} \quad (1.23)$$

Hence,

$$\text{variance} = E[U^2] - (E[U])^2 \quad (1.24)$$

$$= \frac{1}{3} - \left(\frac{1}{2}\right)^2 \quad (1.25)$$

$$= \frac{1}{12} \quad (1.26)$$

$$= 0.0833 \quad (1.27)$$

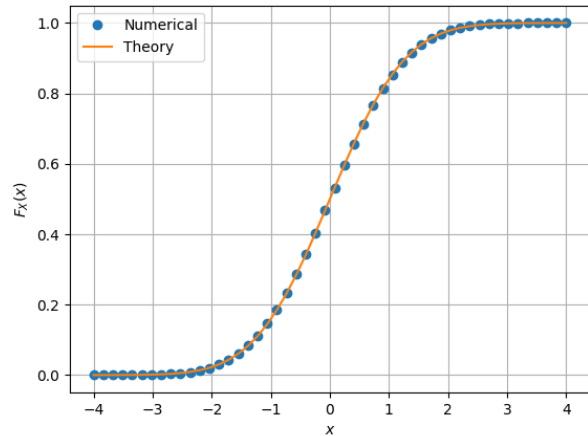


Fig. 2.2: The CDF of X

2 CENTRAL LIMIT THEOREM

2.1 Generate 10^6 samples of the random variable

$$X = \sum_{i=1}^{12} U_i - 6 \quad (2.1)$$

using a C program, where $U_i, i = 1, 2, \dots, 12$ are a set of independent uniform random variables between 0 and 1 and save in a file called gau.dat **Solution:**

```
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/exrand.c
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/coeffs.h
```

2.2 Load gau.dat in python and plot the empirical CDF of X using the samples in gau.dat. What properties does a CDF have?

Solution: The CDF of X is plotted in Fig. 2.2 using code

```
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/cdf_plot.py
```

The CDF of a random variable U has the following properties:

- a) $F_U(x)$ is a non decreasing function of x where $-\infty < x < \infty$

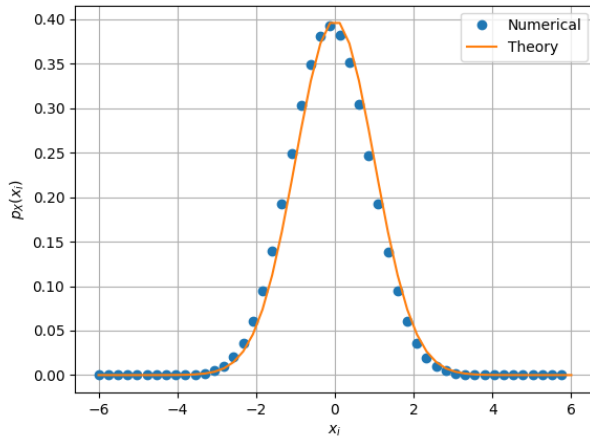


Fig. 2.3: The PDF of X

- b) $F_U(x)$ ranges from 0 to 1
- c) $F_U(x) = 0$ as $x \rightarrow -\infty$
- d) $F_U(x) = 1$ as $x \rightarrow \infty$

2.3 Load gau.dat in python and plot the empirical PDF of X using the samples in gau.dat. The PDF of X is defined as

$$p_X(x) = \frac{d}{dx} F_X(x) \quad (2.2)$$

What properties does the PDF have?

Solution: The PDF of X is plotted in Fig. 2.3 using the code below

```
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/pdf_plot.py
```

The PDF of a random variable X has the following properties:

- a) The probability density function is non-negative for all the possible values.
- b) $\int_{-\infty}^{\infty} f(x) dx = 1$
- c) $f(x) = 0$ as $x \rightarrow -\infty$
- d) $f(x) = 0$ as $x \rightarrow \infty$

2.4 Find the mean and variance of X by writing a C program.

Solution:

```
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/mean_variance.c
```

```
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/coeffs.h
```

$$\text{Mean obtained} = 0.000326 \quad (2.3)$$

$$\text{Variance obtained} = 1.000907 \quad (2.4)$$

2.5 Given that

$$p_X(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right), -\infty < x < \infty, \quad (2.5)$$

repeat the above exercise theoretically.

Solution: By definition,

$$p_X(x) dx = dF_U(x) \quad (2.6)$$

Hence,

$$E[U] = \int_{-\infty}^{\infty} x p_X(x) dx \quad (2.7)$$

$$= \int_{-\infty}^{\infty} x \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) dx \quad (2.8)$$

$$= 0 \quad (2.9)$$

Also,

$$\text{variance} = E[U^2] - (E[U])^2 \quad (2.10)$$

$$E[U^2] = \int_{-\infty}^{\infty} x^2 p_X(x) dx \quad (2.11)$$

$$= \int_{-\infty}^{\infty} x^2 \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) dx \quad (2.12)$$

Using integration by parts, (2.13)

$$= \int_{-\infty}^{\infty} x x \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) dx \quad (2.14)$$

$$= \left[-x \frac{1}{\sqrt{2\pi}} e^{\left(-\frac{x^2}{2}\right)} + \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) dx \right] \quad (2.15)$$

$$= 0 + \frac{1}{\sqrt{2\pi}} \sqrt{2\pi} \quad (2.16)$$

$$= 1 \implies \text{variance} = 1 \quad (2.17)$$

Also,

$$F_U(x) = \int_{-\infty}^x p_X(x) dx \quad (2.18)$$

$$= \int_{-\infty}^x e^{-\frac{x^2}{2}} dx \quad (2.19)$$

3 FROM UNIFORM TO OTHER

3.1 Generate samples of

$$V = -2 \ln(1 - U) \quad (3.1)$$

and plot its CDF.

Solution:

```
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/exrand.c
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/coeffs.h
```

The CDF of X is plotted in Fig. 3.2 using the code below

```
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/cdf_plot.py
```

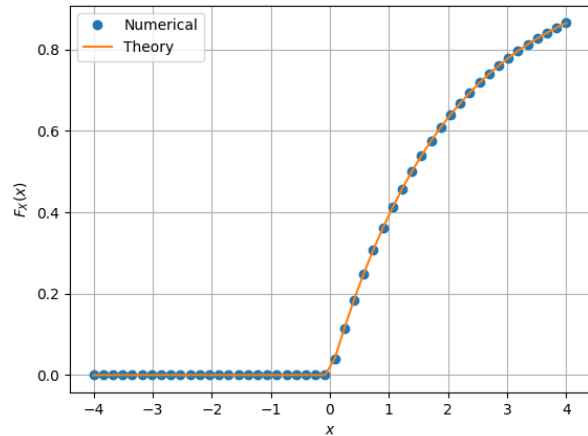


Fig. 3.2: CDF of V

```
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/coeffs.h
```

3.2 Find a theoretical expression for $F_V(x)$.

Solution: Let, $V = g(U)$

$$V = -2 \ln(1 - U) \quad (3.2)$$

$$\Rightarrow U = 1 - e^{-\frac{V}{2}} \quad (3.3)$$

Now,

$$F_V(x) = P(g(U) \leq x) \quad (3.4)$$

$$= P(X < g^{-1}(V)) \quad (3.5)$$

$$= F_U(g^{-1}(V)) \quad (3.6)$$

$$= F_U(1 - e^{-\frac{V}{2}}) \quad (3.7)$$

$$F_U(1 - e^{-\frac{V}{2}}) = \begin{cases} 1 - e^{-\frac{V}{2}}, & V \in (0, \infty) \\ 0, & \text{otherwise} \end{cases} \quad (3.8)$$

4 TRIANGULAR DISTRIBUTION

4.1 Generate

$$T = U_1 + U_2 \quad (4.1)$$

Solution:

```
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/exrand.c
```

4.2 Find the CDF of T .

Solution: The CDF of T is plotted in Fig. 4.2

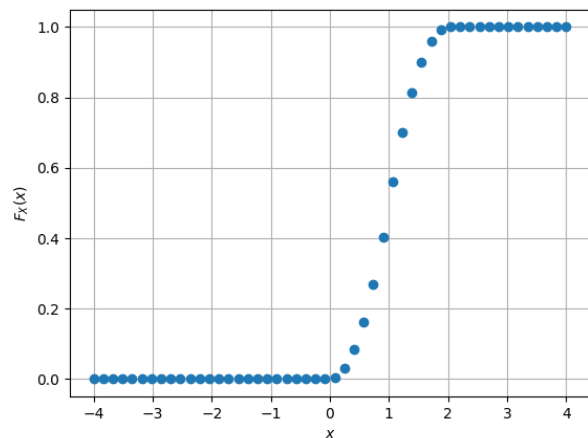


Fig. 4.2: CDF of T

4.3 Find the PDF of T .

Solution: The PDF of T is plotted in Fig. 4.3

4.4 Find the theoretical expressions for the PDF and CDF of T .

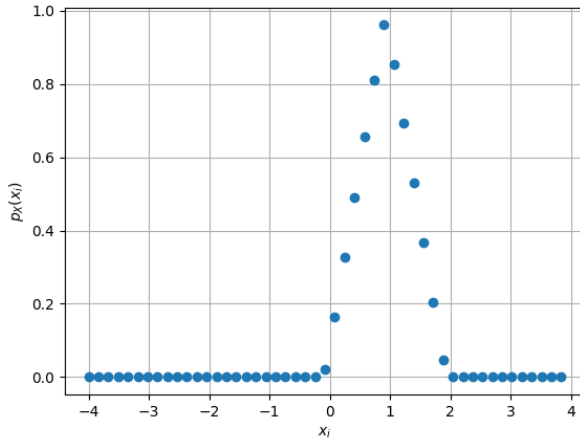


Fig. 4.3: PDF of T

Solution:

$$T = U_1 + U_2 \quad (4.2)$$

$$\Rightarrow f_T(T) = f_{U_1+U_2}(T) \quad (4.3)$$

$$= f_{U_1}(T) * f_{U_2}(T) \quad (4.4)$$

$$= \int_{-\infty}^{\infty} f_{U_1}(T - U_2) f_{U_2}(U_2) dU_2 \quad (4.5)$$

since U_1 and U_2 are independent Also,

$$f_{U_1}(U_1) = f_{U_2}(U_2) = \begin{cases} 1, & x \in (0, 1) \\ 0, & \text{otherwise} \end{cases} \quad (4.6)$$

Hence for $0 \leq U_2 \leq 1$, Eq:(4.5) becomes,

$$f_T(T) = \int_0^1 f_{U_1}(T - U_2) dU_2 \quad (4.7)$$

If $0 \leq T \leq 1$,

$$f_T(T) = \int_0^T dU_2 \quad (4.8)$$

$$= T \quad (4.9)$$

If $1 < T \leq 2$,

$$f_T(T) = \int_{T-1}^1 dU_2 \quad (4.10)$$

$$= 2 - T \quad (4.11)$$

$$f_T(T) = \begin{cases} T, & 0 \leq T \leq 1 \\ 2 - T, & 1 < T \leq 2 \\ 0, & \text{otherwise} \end{cases} \quad (4.12)$$

Now,

$$F_T(T) = \int_{-\infty}^T f_T(T) dT \quad (4.13)$$

Hence, if $T < 0$,

$$F_T(T) = \int_{-\infty}^T f_T(T) dT \quad (4.14)$$

$$= \int_{-\infty}^T 0 dT \quad (4.15)$$

$$= 0 \quad (4.16)$$

If $0 \leq T \leq 1$,

$$F_T(T) = \int_0^T f_T(T) dT \quad (4.17)$$

$$= \int_0^T T dT \quad (4.18)$$

$$= \frac{T^2}{2} \quad (4.19)$$

If $1 < T \leq 2$,

$$F_T(T) = \int_0^1 f_T(T) dT + \int_1^T f_T(T) dT \quad (4.20)$$

$$= \int_0^1 T dT + \int_1^T (2 - T) dT \quad (4.21)$$

$$= \frac{1}{2} + \frac{4T - T^2 - 3}{2} \quad (4.22)$$

$$= 1 - \frac{(2 - T)^2}{2} \quad (4.23)$$

If $T > 2$,

$$F_T(T) = \int_2^T f_T(T) dT \quad (4.24)$$

$$= \int_2^T 0 dT \quad (4.25)$$

$$= 0 \quad (4.26)$$

$$F_T(T) = \begin{cases} \frac{T^2}{2}, & 0 \leq T \leq 1 \\ 1 - \frac{(2-T)^2}{2}, & 1 < T \leq 2 \\ 0, & \text{otherwise} \end{cases} \quad (4.27)$$

4.5 Verify your results through a plot.

Solution: The CDF of T is plotted in Fig. 4.5 below

The PDF of T is plotted in Fig. 4.5 below
Using codes

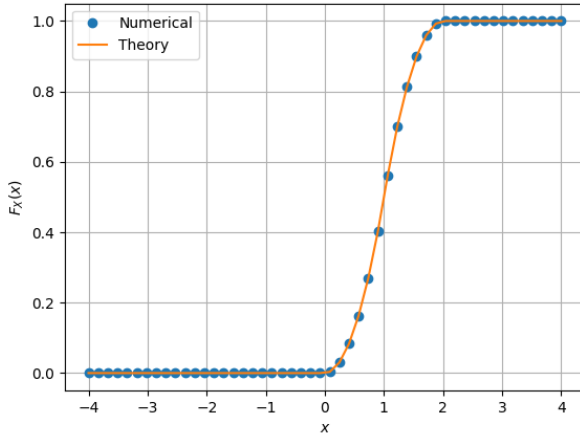


Fig. 4.5: CDF of T

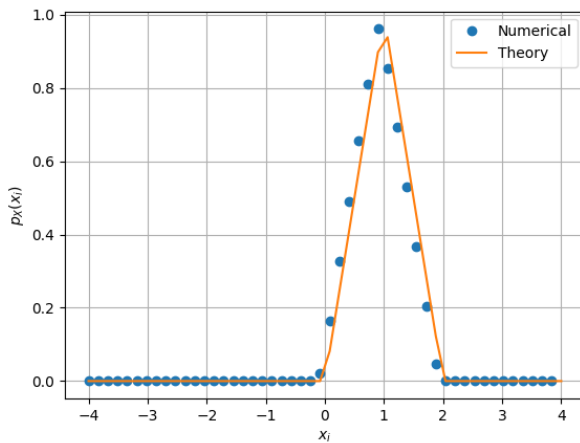


Fig. 4.5: PDF of T

```
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/cdf_plot.py
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/pdf_plot.py
```

5 MAXIMUM LIKELIHOOD

5.1 Generate equiprobable $X \in \{1, -1\}$.

Solution:

```
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/exrand.c
```

```
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/coeffs.h
```

5.2 Generate

$$Y = AX + N, \quad (5.1)$$

where $A = 5$ dB, and $N \sim \mathcal{N}(0, 1)$.

Solution:

```
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/exrand.c
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/coeffs.h
```

5.3 Plot Y using a scatter plot. The scatter plot of Y is plotted in Fig. 5.3 below

Solution:

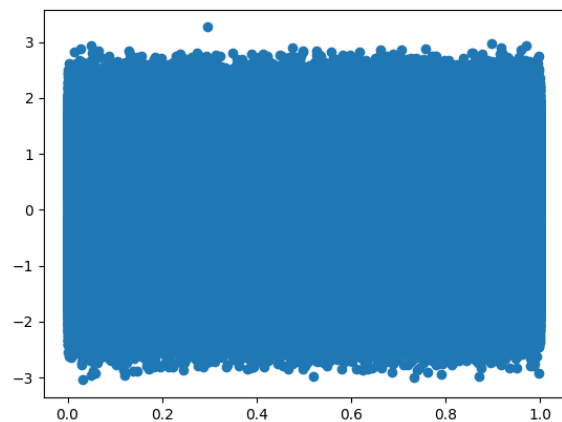


Fig. 5.3: Scatter plot of Y

5.4 Guess how to estimate X from Y .

Solution:

To estimate X from Y we define the following function,

$$g(y) = \begin{cases} 1, & y \in (0, \infty) \\ -1 & y \in (-\infty, 0] \end{cases} \quad (5.2)$$

Hence using this function we can operate on Y to find X

5.5 Find

$$P_{e|0} = \Pr(\hat{X} = -1 | X = 1) \quad (5.3)$$

and

$$P_{e|1} = \Pr(\hat{X} = 1|X = -1) \quad (5.4)$$

Solution:

```
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/exrand.c
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/coeffs.h
```

Values obtained

$$P_{e|0} = 0.311084 \quad (5.5)$$

$$P_{e|1} = 0.311586 \quad (5.6)$$

5.6 Find P_e assuming that X has equiprobable symbols.

Solution: We have,

$$P_{e|0} = \Pr(\hat{X} = -1|X = 1) \quad (5.7)$$

$$= \Pr(AX + N < 0|X = 1) \quad (5.8)$$

$$= \Pr(N < -A) \quad (5.9)$$

$$= \int_{-\infty}^{-A} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx \quad (5.10)$$

$$= \int_A^{\infty} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx \quad (5.11)$$

$$= Q_N(A) \quad (5.12)$$

$$P_{e|1} = \Pr(\hat{X} = 1|X = -1) \quad (5.13)$$

$$= \Pr(AX + N > 0|X = -1) \quad (5.14)$$

$$= \Pr(N > A) \quad (5.15)$$

$$= Q_N(A) \quad (5.16)$$

$$P_e = P_{e|0}P(X = 1) + P_{e|1}P(X = -1) \quad (5.17)$$

$$= P_{e|0}\frac{1}{2} + P_{e|1}\frac{1}{2} \quad (5.18)$$

$$= Q_N(A)\frac{1}{2} + Q_N(A)\frac{1}{2} \quad (5.19)$$

$$= Q_N(A) \quad (5.20)$$

5.7 Verify by plotting the theoretical P_e with respect to A from 0 to 10 dB.

Solution:

5.8 Now, consider a threshold δ while estimating X from Y . Find the value of δ that minimizes

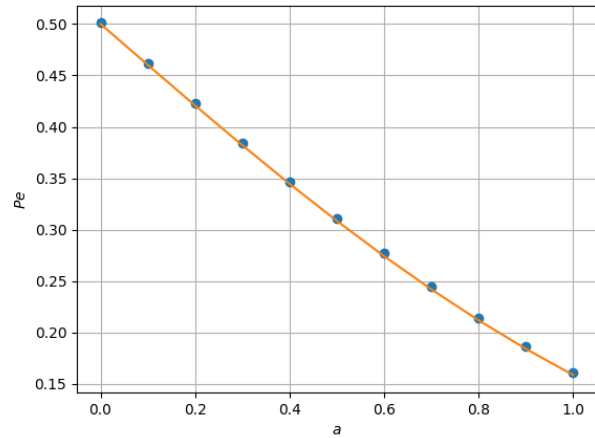


Fig. 5.7: P_e vs A

the theoretical P_e . To estimate X from Y we now define the following function,

$$g(y) = \begin{cases} 1, & y \in (\delta, \infty) \\ -1 & y \in (-\infty, \delta] \end{cases} \quad (5.21)$$

We have,

$$P_{e|0} = \Pr(\hat{X} = -1|X = 1) \quad (5.22)$$

$$= \Pr(AX + N < \delta|X = 1) \quad (5.23)$$

$$= \Pr(N < \delta - A) \quad (5.24)$$

$$= \int_{-\infty}^{\delta-A} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx \quad (5.25)$$

$$= \int_{A-\delta}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx \quad (5.26)$$

$$= Q_N(A - \delta) \quad (5.27)$$

$$P_{e|1} = \Pr(\hat{X} = 1|X = -1) \quad (5.28)$$

$$= \Pr(AX + N > \delta|X = -1) \quad (5.29)$$

$$= \Pr(N > \delta + A) \quad (5.30)$$

$$= Q_N(\delta + A) \quad (5.31)$$

Hence,

$$P_e = P_{e|0}P(X = 1) + P_{e|1}P(X = -1) \quad (5.32)$$

$$= P_{e|0}\frac{1}{2} + P_{e|1}\frac{1}{2} \quad (5.33)$$

$$= Q_N(A - \delta)\frac{1}{2} + Q_N(A + \delta)\frac{1}{2} \quad (5.34)$$

To minimize P_e we differentiate P_e wrt δ ,

$$0 = \frac{d}{d\delta} \left(\frac{1}{2} (Q_N(A - \delta) + Q_N(A + \delta)) \right) \quad (5.35)$$

$$= \frac{1}{2} \left(\frac{1}{\sqrt{2\pi}} e^{-\frac{(\delta-A)^2}{2}} - \frac{1}{\sqrt{2\pi}} e^{-\frac{(A+\delta)^2}{2}} \right) \quad (5.36)$$

$$(5.37)$$

Therefore,

$$(\delta - A)^2 = (A + \delta)^2 \quad (5.38)$$

$$\implies \delta = 0 \quad (5.39)$$

5.9 Repeat the above exercise when

$$p_X(0) = p \quad (5.40)$$

Solution: Let $PX = -1 = p$, hence $PX = 1 = 1 - p$ Hence,

$$P_e = P_{e|0}P(X = 1) + P_{e|1}P(X = -1) \quad (5.41)$$

$$P_e = P_{e|0}(1 - p) + P_{e|1}p \quad (5.42)$$

$$= Q_N(A - \delta)(1 - p) + Q_N(A + \delta)p \quad (5.43)$$

Differentiating like before,

$$0 = p \frac{1}{\sqrt{2\pi}} e^{-\frac{(\delta-A)^2}{2}} - (1 - p) \frac{1}{\sqrt{2\pi}} e^{-\frac{(A+\delta)^2}{2}} \quad (5.44)$$

Taking log from both sides,

$$\ln p - \frac{(\delta - A)^2}{2} = \ln(1 - p) + \frac{(\delta + A)^2}{2} \quad (5.45)$$

$$\implies \delta = \frac{1}{2A} \ln \frac{1 - p}{p} \quad (5.46)$$

5.10 Repeat the above exercise using the MAP criterion.

Solution: Using the MAP criterion we need to maximize $p_{X|Y}(x|y)$. Using Bayes Theorem,

$$p_{X|Y}(x|y) = \frac{p_{Y|X}(y|x) p_X(x)}{p_Y(y)} \quad (5.47)$$

There are 2 cases for $X = 1$ and $X = -1$,

$$p_{Y|X}(y|x) = \begin{cases} p_{Y|X}(y|1) & X = 1 \\ p_{Y|X}(y|-1) & X = -1 \end{cases} \quad (5.48)$$

$$p_{Y|X}(y|x) = \begin{cases} p_N(y - A) & X = 1 \\ p_N(y + A) & X = -1 \end{cases} \quad (5.49)$$

Now,

$$p_Y(y) = p_X(1) p_{Y|X}(y|1) + p_X(-1) p_{Y|X}(y|-1) \quad (5.50)$$

$$p_Y(y) = (p) p_N(y - A) + (1 - p) p_N(y + A) \quad (5.51)$$

Hence,

$$p_{X|Y}(x|y) = \begin{cases} \frac{\frac{p_N(y-A)(p)}{(p)p_N(y-A)+(1-p)p_N(y+A)}}{\frac{p_N(y+A)(1-p)}{(p)p_N(y-A)+(1-p)p_N(y+A)}} & X = 1 \\ \frac{p_N(y+A)(1-p)}{(p)p_N(y-A)+(1-p)p_N(y+A)} & X = -1 \end{cases} \quad (5.52)$$

$$p_{X|Y}(x|y) = \begin{cases} \frac{(p)}{(p)+(1-p)\frac{p_N(y+A)}{p_N(y-A)}} & X = 1 \\ \frac{(1-p)}{(p)\frac{p_N(y-A)}{p_N(y+A)}+(1-p)} & X = -1 \end{cases} \quad (5.53)$$

Simplifying using $p_N(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$,

$$p_{X|Y}(x|y) = \begin{cases} \frac{(p)}{(p)+(1-p)e^{-2yA}} & X = 1 \\ \frac{(1-p)}{(p)e^{2yA}+(1-p)} & X = -1 \end{cases} \quad (5.54)$$

Therefore when $p_{X|Y}(1|y) > p_{X|Y}(-1|y)$ we have,

$$\frac{(p)}{(p) + (1 - p) e^{-2yA}} > \frac{(1 - p)}{(p) e^{2yA} + (1 - p)} \quad (5.55)$$

$$p^2 e^{2yA} > (1 - p)^2 e^{-2yA} \quad (5.56)$$

$$e^{4yA} > \left(\frac{1 - p}{p} \right)^2 \quad (5.57)$$

$$e^{2yA} > \frac{1 - p}{p} \quad (5.58)$$

$$y > \frac{1}{2A} \ln \frac{1 - p}{p} \quad (5.59)$$

Hence if $y > \frac{1}{2A} \ln \frac{1-p}{p}$ we can say that $X = 1$ and if $y \leq \frac{1}{2A} \ln \frac{1-p}{p}$ then $X = -1$ Now if $p = \frac{1}{2}$

$$y > \frac{1}{2A} \ln \frac{1 - p}{p} \quad (5.60)$$

$$y > \frac{1}{2A} \ln 1 \quad (5.61)$$

$$y > 0 \quad (5.62)$$

Hence $\delta = 0$ for $p = \frac{1}{2}$

6 GAUSSIAN TO OTHER

6.1 Let $X_1 \sim 01$ and $X_2 \sim 01$. Plot the CDF and PDF of

$$V = X_1^2 + X_2^2 \quad (6.1)$$

Solution: Data for the plots is generated using

```
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/exrand.c
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/coeffs.h
```

```
Random%20Variables%20Assignment/
codes/cdf_plot.py
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/pdf_plot.py
```

The CDF of V is plotted in Fig. 6.1

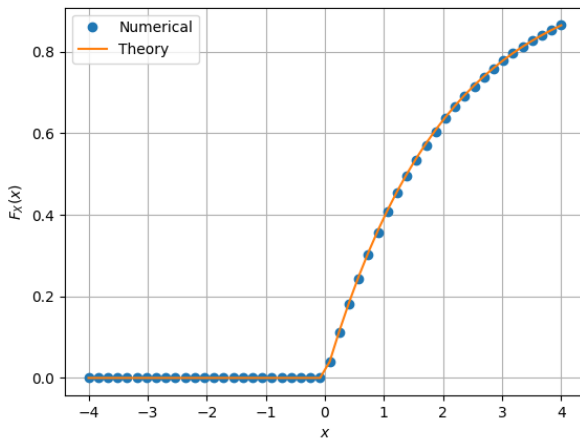


Fig. 6.1: CDF of V

The PDF of V is plotted in Fig. 6.1

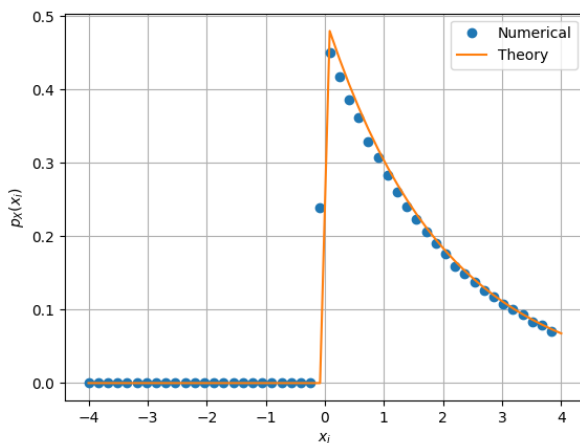


Fig. 6.1: PDF of V

Using codes

```
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
```

6.2 If

$$F_V(x) = \begin{cases} 1 - e^{-\alpha x} & x \geq 0 \\ 0 & x < 0, \end{cases} \quad (6.2)$$

find α .

Solution: X_1 and X_2 are i.i.d . Let,

$$X_1 = R \cos \theta \quad (6.3)$$

$$X_2 = R \sin \theta \quad (6.4)$$

Using Jacobian transform we have,

$$f_{X_1 X_2}(X_1, X_2) = \frac{1}{|J(R, \theta)|} f_{R\theta}(R, \theta) \quad (6.5)$$

Now, Jacobian matrix is given as follows,

$$J = \begin{pmatrix} \frac{\partial x_1}{\partial R} & \frac{\partial x_1}{\partial \theta} \\ \frac{\partial x_2}{\partial R} & \frac{\partial x_2}{\partial \theta} \end{pmatrix} \quad (6.6)$$

$$J = \begin{pmatrix} \cos \theta & -R \sin \theta \\ \sin \theta & R \cos \theta \end{pmatrix} \quad (6.7)$$

$$|J| = R \quad (6.8)$$

$$\Rightarrow f_{X_1 X_2}(X_1, X_2) = \frac{1}{R} f_{R\theta}(R, \theta) \quad (6.9)$$

Now,

$$f_{X_1 X_2}(X_1, X_2) = f_{X_1}(X_1) f_{X_2}(X_2) \quad (6.10)$$

$$= \frac{1}{\sqrt{2\pi}} e^{-\frac{x_1^2}{2}} \frac{1}{\sqrt{2\pi}} e^{-\frac{x_2^2}{2}} \quad (6.11)$$

$$= \frac{1}{2\pi} e^{-\frac{(x_1^2 + x_2^2)}{2}} \quad (6.12)$$

$$= \frac{1}{2\pi} e^{-\frac{R^2}{2}} \quad (6.13)$$

Hence,

$$f_{R\theta}(R, \theta) = \frac{R}{2\pi} e^{-\frac{R^2}{2}} \quad (6.14)$$

$$\Rightarrow f_R(R) = \int_0^{2\pi} \frac{R}{2\pi} e^{-\frac{R^2}{2}} d\theta \quad (6.15)$$

$$= R e^{-\frac{R^2}{2}} \quad (6.16)$$

Hence,

$$f_R(R) = \begin{cases} Re^{-\frac{R^2}{2}} & x \geq 0 \\ 0 & x < 0, \end{cases} \quad (6.17)$$

$$F_R(R) = \begin{cases} 1 - e^{-\frac{R^2}{2}} & x \geq 0 \\ 0 & x < 0, \end{cases} \quad (6.18)$$

Now,

$$F_V(x) = P(V \leq x) \quad (6.19)$$

$$= P(R^2 \leq x) \quad (6.20)$$

$$= P(R \leq \sqrt{x}) \quad (6.21)$$

$$= F_R(\sqrt{x}) \quad (6.22)$$

Hence,

$$F_V(x) = \begin{cases} 1 - e^{-\frac{x}{2}} & x \geq 0 \\ 0 & x < 0, \end{cases} \quad (6.23)$$

$$\Rightarrow \alpha = \frac{1}{2} \quad (6.24)$$

6.3 Plot the CDF and PDF of

$$A = \sqrt{V} \quad (6.25)$$

Solution: Data for the plots is generated using

```
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/exrand.c
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/coeffs.h
```

The CDF of A is plotted in Fig. 6.3 below
The PDF of A is plotted in Fig. 6.3 below
Using codes

```
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/cdf_plot.py
wget https://github.com/tejalkul/AI1110-
Assignments/blob/main/AI1110%20
Random%20Variables%20Assignment/
codes/pdf_plot.py
```

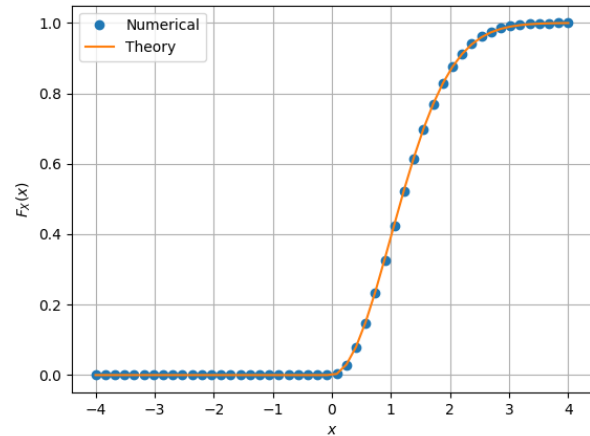


Fig. 6.3: CDF of A

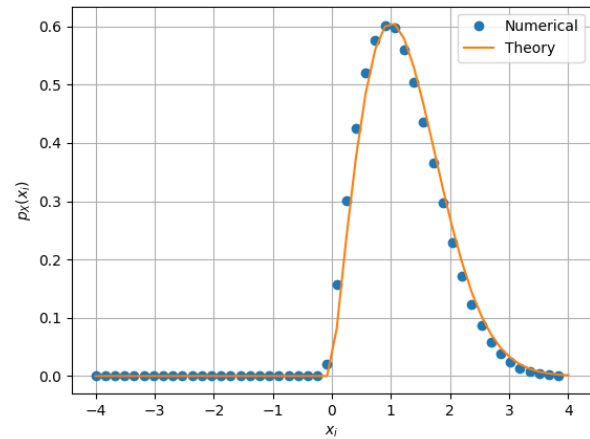


Fig. 6.3: PDF of A

7 CONDITIONAL PROBABILITY

7.1 Plot

$$P_e = \Pr(\hat{X} = -1 | X = 1) \quad (7.1)$$

for

$$Y = AX + N, \quad (7.2)$$

where A is Raleigh with $E[A^2] = \gamma$, $N \sim 01$, $X \in (-1, 1)$ for $0 \leq \gamma \leq 10$ dB.

7.2 Assuming that N is a constant, find an expression for P_e . Call this $P_e(N)$

7.3 For a function g ,

$$E[g(X)] = \int_{-\infty}^{\infty} g(x)p_X(x)dx \quad (7.3)$$

Find $P_e = E[P_e(N)]$.

7.4 Plot P_e in problems 7.1 and 7.3 on the same

graph w.r.t γ . Comment.

8 TWO DIMENSIONS

Let

$$\mathbf{y} = A\mathbf{x} + \mathbf{n}, \quad (8.1)$$

where

$$x \in (\mathbf{s}_0, \mathbf{s}_1), \mathbf{s}_0 = \begin{pmatrix} 1 \\ 0 \end{pmatrix}, \mathbf{s}_1 = \begin{pmatrix} 0 \\ 1 \end{pmatrix} \quad (8.2)$$

$$\mathbf{n} = \begin{pmatrix} n_1 \\ n_2 \end{pmatrix}, n_1, n_2 \sim 01. \quad (8.3)$$

8.1 Plot

$$\mathbf{y}|\mathbf{s}_0 \text{ and } \mathbf{y}|\mathbf{s}_1 \quad (8.4)$$

on the same graph using a scatter plot.

8.2 For the above problem, find a decision rule for detecting the symbols \mathbf{s}_0 and \mathbf{s}_1 .

8.3 Plot

$$P_e = \Pr(\hat{\mathbf{x}} = \mathbf{s}_1 | \mathbf{x} = \mathbf{s}_0) \quad (8.5)$$

with respect to the SNR from 0 to 10 dB.

8.4 Obtain an expression for P_e . Verify this by comparing the theory and simulation plots on the same graph.