

Math 170E: Fall 2020

Lecture 5, Fri 20th Jan

Conditional probability and Bayes' theorem

Last time:

Last time, we were interested in taking r samples from n objects

- we can do this **with** or **without** replacement
- we can seek **ordered** or **unordered** samples
- the Binomial theorem

Today:

We'll discuss today:

- the notion of *conditional probability*
- Bayes' theorem

Suppose $A, B \subseteq \Omega$ are events.

If we know that event B has occurred, how does this affect the probability of event A occurring?

Example 13: You roll a fair six-sided die twice. You know one of the die rolls is a 6. What is the probability that the sum of the two rolls is a 7?

$$\Omega = \{(1,1), (1,2), \dots, (6,6)\},$$

$$A = \{\text{sum is a 7}\} = \{(1,6), (6,1), (2,5), (5,2), (3,4), (4,3)\}.$$

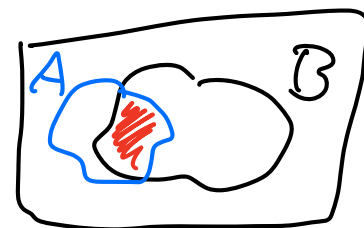
$$\hookrightarrow P(A) = \frac{|A|}{|\Omega|} = \frac{6}{36} = 1/6, \rightsquigarrow P(A \text{ given that } B \text{ happened})$$

$$B = \{\text{one of the rolls is a 6}\} = \{(1,6), (2,6), \dots, (6,6), (6,1), \dots, (6,5)\}$$

$$\hookrightarrow |B| = 11.$$

$$\hookrightarrow A \cap B = \{\text{elements in } A \text{ also in } B\} = \{(1,6), (6,1)\} \rightarrow |A \cap B| = 2. \Omega.$$

$$P(A \text{ given that } B \text{ happened}) = \frac{|A \cap B|}{|B|} = \frac{2}{11}$$



$$\frac{\frac{|A \cap B|}{|\Omega|}}{|B|/|\Omega|} = \frac{P(A \cap B)}{P(B)}$$

Definition 1.13: (Conditional probability)

Let $B \subseteq \Omega$ be an event so that $\mathbb{P}(B) \neq 0$. The probability of an event $A \subseteq \Omega$ conditioned on the event B is given by

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)}.$$

Given B

↪ Because we know that B has occurred
our state space (all possible outcomes) has shrunk!

"probability that A occurs given that
 B has occurred."

Theorem 1.14: (Properties of the conditional probability) $(\Omega, \mathcal{F}, \mathbb{P})$
 $(B, \mathbb{P}(B) \neq 0)$
 If $B \subseteq \Omega$ such that $\mathbb{P}(B) \neq 0$, the $\mathbb{P}(\cdot|B)$ is a probability measure.
 i.e. $\mathbb{P}(\cdot|B) : \mathcal{F} \mapsto [0, 1]$ satisfies: $(B, \mathcal{F}, \mathbb{P}(\cdot|B))$
is a prob.-space

1. $\mathbb{P}(\Omega|B) = 1$
2. (Countable additivity) If $\{A_j\}_{j=1}^k$ are mutually exclusive events, then

$$\mathbb{P}\left(\bigcup_{j=1}^k A_j \middle| B\right) = \sum_{j=1}^k \mathbb{P}(A_j|B),$$

and (when " $k = +\infty$ "),

$$\mathbb{P}\left(\bigcup_{j=1}^{\infty} A_j \middle| B\right) = \sum_{j=1}^{\infty} \mathbb{P}(A_j|B),$$

\implies all properties of prob. measures we proved in Lectures 1 & 2 hold for $\mathbb{P}(\cdot|B)$:
 e.g. $\mathbb{P}(A'|B) = 1 - \mathbb{P}(A|B)$

Proof: 1) $P(\Omega | B) = \frac{P(\Omega \cap B)}{P(B)} = \frac{P(B)}{P(B)} = 1.$

2) Suppose $\{A_j\}_{j=1}^K$ are mutually exclusive.

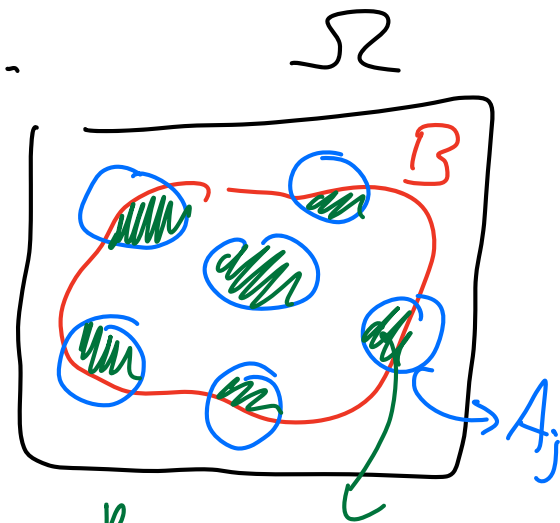
$$P\left(\bigcup_{j=1}^K A_j | B\right) = \frac{P\left(\left(\bigcup_{j=1}^K A_j\right) \cap B\right)}{P(B)}$$

by Countable additivity

$$= \frac{P\left(\bigcup_{j=1}^K (A_j \cap B)\right)}{P(B)}$$

$$= \frac{\sum_{j=1}^K P(A_j \cap B)}{P(B)}$$

$$= \sum_{j=1}^K \frac{P(A_j \cap B)}{P(B)} = \sum_{j=1}^K P(A_j | B).$$



$$\left(\bigcup_{j=1}^K A_j\right) \cap B = \bigcup_{j=1}^K (A_j \cap B)$$

$\{A_j\}$ are m.e. \Downarrow $\{A_j \cap B\}$ are m.e.

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Example 14:

- You have a regular pack of 52 cards
- You draw 3 cards at random
- Given that the first card you draw is not an ace, what is the probability that you draw at least one ace?

→ 4 aces in 52 card deck.

$\Omega = \{\text{all possible 3 card hands}\}$

$A = \{\text{drew at least 1 ace}\}$

$B = \{\text{first is not an ace}\}$

Want:

$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$

A) $\frac{4}{52} \times \frac{3}{51} \times \frac{2}{50}$

B) $1 - \frac{48}{52} \times \frac{47}{51} \times \frac{46}{50}$

C) $\frac{1}{52} \times \frac{48}{51}$

D) $1 - \frac{47}{51} \times \frac{46}{50}$

$P(B) = \frac{48}{52}$

$P(A \cap B) = ??$

$A \cap B = \{ \text{1st not an ace, 2nd is ace, 3rd not ace} \}$
 $\cup \{ \text{1st not an ace, 2nd not ace, 3rd is ace} \}$
 $\cup \{ \text{1st not an ace, 2nd is ace, 3rd is ace} \}$

$\hookrightarrow A' = \{\text{drew no aces}\}$

$\hookrightarrow A' \subseteq B \hookrightarrow A' \cap B = A'$

$$P(A|B) = 1 - P(A'|B).$$

$$= 1 - \frac{P(A' \cap B)}{P(B)}$$

$$P(A' \cap B) = P(A')$$

$$= \frac{48}{52} \times \frac{47}{51} \times \frac{46}{50}$$

$$= 1 - \frac{\cancel{\frac{48}{52}} \times \frac{47}{51} \times \frac{46}{50}}{\cancel{48/52}}$$

$$= 1 - \frac{47}{51} \times \frac{46}{50} //$$

Proposition 1.15: If $A, B \subseteq \Omega$ are independent and $\mathbb{P}(B) \neq 0$, then

$$\mathbb{P}(A|B) = \mathbb{P}(A)$$

Proof: If A & B are indep, knowing B tells me nothing about A .

$$\hookrightarrow \mathbb{P}(A|B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)} \stackrel{\text{indep}}{=} \frac{\mathbb{P}(A) \cancel{\mathbb{P}(B)}}{\cancel{\mathbb{P}(B)}} = \mathbb{P}(A). //$$

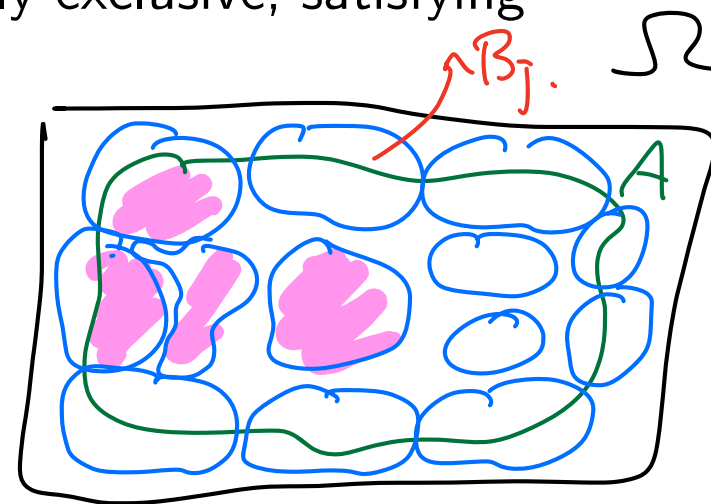
Theorem 1.16: (The law of total probability)

Let $A \subseteq \Omega$ be an event and $\{B_j\}_{j=1}^k \subseteq \Omega$ be mutually exclusive, satisfying $\mathbb{P}(B_j) \neq 0$, for every $j \in \{1, \dots, k\}$, and

$$A \subseteq \bigcup_{j=1}^k B_j.$$

Then

$\{B_j\}$ cover A .



$$\mathbb{P}(A) = \mathbb{P}(A|B_1)\mathbb{P}(B_1) + \dots + \mathbb{P}(A|B_k)\mathbb{P}(B_k) = \sum_{j=1}^k \mathbb{P}(A|B_j)\mathbb{P}(B_j).$$

$$\mathbb{P}(A|B_j) = \frac{\mathbb{P}(A \cap B_j)}{\mathbb{P}(B_j)}$$

Proof: $A \subseteq \bigcup_{j=1}^{\infty} B_j \hookrightarrow A = A \cap \left(\bigcup_{j=1}^{\infty} B_j \right) = \bigcup_{j=1}^{\infty} (A \cap B_j)$

$\{B_j\}$ m.e. $\Rightarrow \{A \cap B_j\}$ are m.e.

By countable additivity,

$$\begin{aligned} P(A) &= P\left(\bigcup_{j=1}^{\infty} (A \cap B_j)\right) \\ &= \sum_{j=1}^{\infty} P(A \cap B_j) = \sum_{j=1}^{\infty} P(A|B_j) P(B_j). \end{aligned}$$

Example 15: A bin has three types of disposable flashlights in it.

- The probability a flashlight of type 1 lasts more than 100 hours of use is 70%
- for types 2 and 3, the probability is 40% and 30% respectively
- suppose 20% in the bin are type 1, 30% are type 2 and 50% are type 3

What is the probability that a randomly selected flashlight has more than 100 hours of use?

$\Omega = \{\text{all flashlights in the bin}\}$

$A = \{\text{chosen has } \geq 100 \text{ hrs}\}$

WANT: $P(A)$.

$B_j = \{\text{chosen is of type } j\}, j=1, 2, 3.$ \rightarrow mutually exclusive

$P(B_1) = 0.2, P(B_2) = 0.3, P(B_3) = 0.5.$

$P(A|B_j): P(A|B_1) = 0.7, P(A|B_2) = 0.4,$
 $P(A|B_3) = 0.3.$

By the law of total. Prob.,

$$P(A) = \sum_{j=1}^3 P(A|B_j) P(B_j)$$

$$= (0.7) \times (0.2) + (0.4) \times (0.3) + (0.3) \times (0.5)$$

$$= 0.41.$$

↳ 41% chance that a randomly selected flashlight lasts for ≥ 100 hrs.

$$\frac{7}{10} \times \frac{1}{5} + \frac{2}{5} \times \frac{3}{10} + \frac{3}{10} \cdot \frac{1}{2}.$$