



# Outline

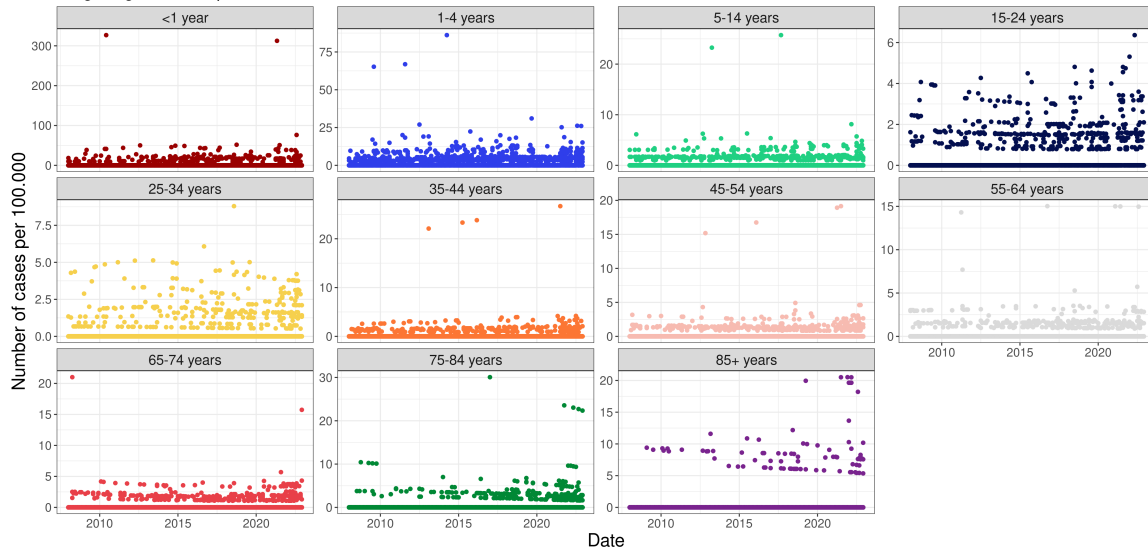
- Data exploration
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  - Implementation
- Hierarchical Poisson-Gamma model
  - Formulation
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Date	ageGroup	y	n
2008-01-01	<1 year	2	10120
2009-01-01	<1 year	1	10288
2010-01-01	<1 year	1	10654
2011-01-01	<1 year	0	11199
...	...	...	...
2019-12-01	85+ years	1	14153
2020-12-01	85+ years	0	14613
2021-12-01	85+ years	2	14976
2022-12-01	85+ years	3	15203

# Data exploration

## VTEC / STEC

Shiga- og verotoxin producerende E. coli.



$$Y_t^a | u_t^a \sim \text{Pois} (w_t^a \lambda_a \exp(u_t^a)) \quad (1a)$$

$$u_t^a \sim N(0, \sigma^2) \quad (1b)$$

## Implementation - Objective function in C++

```

#include <TMB.hpp>           // Links in the TMB libraries
template<class Type>
Type objective_function<Type>::operator() ()
{
    DATA_VECTOR(y);           // Data vector transmitted from R
    DATA_VECTOR(w);           // Data vector transmitted from R
    DATA_FACTOR(ageGroup);     // Data factor transmitted from R

    PARAMETER_VECTOR(u);       // Random effects

    // Parameters
    PARAMETER_VECTOR(lambda);   // Parameter value transmitted from R
    PARAMETER(log_sigma_u);     // Parameter value transmitted from R

    Type sigma_u = exp(log_sigma_u);

    int nobs = y.size();
    Type mean_ran = Type(0);

    int j;

    Type f = 0;                // Declare the "objective function" (neg. log. likelihood)
    for(int i=0; i < nobs; i++){
        f -= dnorm(u[i],mean_ran,sigma_u,true);
        j = ageGroup[i];
        f -= dpois(y[i],exp(log(lambda[j])-log(w[j]))*exp(u[i]),true);
    }

    return f;
}

```

## Hierarchical Poisson-Normal model

# Implementation - Call from R

```
# Import libraries
library(readr)
library(dplyr)
library(TMB)

# Import the data
dat <- read_rds(file = "../data/processed/dat.rds")

# Only consider some of the data
y <- dat %>%
  filter(caseDef == "Shiga- og veratotoxin producerende E. coli.") %>%
  group_by(Date, ageGroup) %>%
  mutate(y = sum(cases)) %>%
  select(Date, ageGroup, y, n)

compile(file = "PoissonLognormal.cpp") # Compile the C++ file
dyn.load(dynlib("PoissonLognormal")) # Dynamically link the C++ code

# Function and derivative
PoisLN <- MakeADFun(
  data = list(y = y$, ageGroup = y$ageGroup, w = y$n),
  parameters = list(u = rep(1, length(y$y)),
                    lambda = rep(1, nlevels(y$ageGroup)),
                    log_sigma_u = log(1)),
  random = "u",
  DLL = "PoissonLognormal"
)

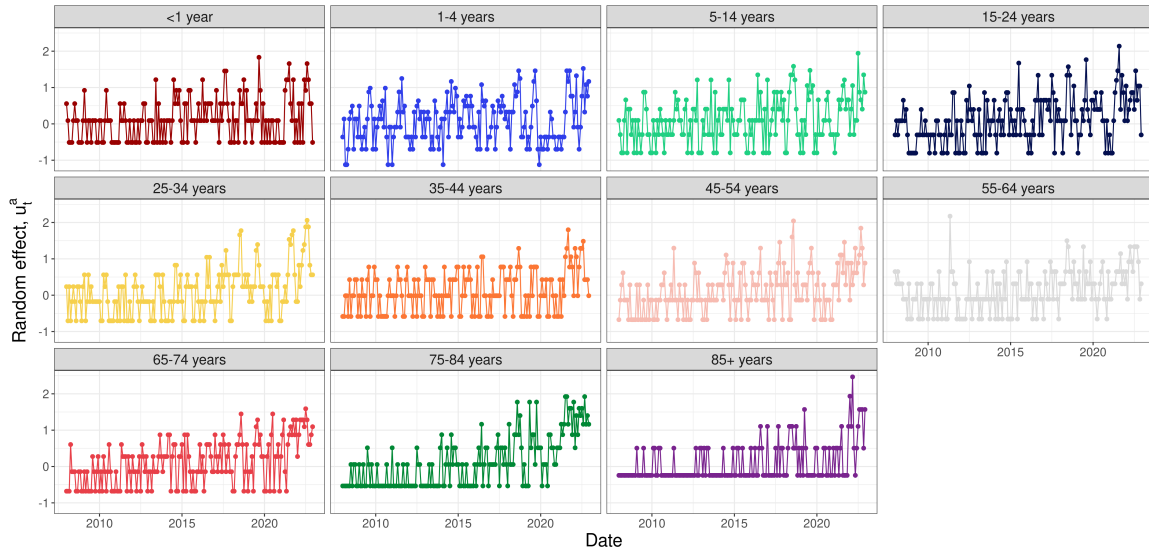
opt <- nlminb(start = PoisLN$par, PoisLN$fn, PoisLN$gr, lower = c(0.01, 0.01))
```

Parameter	Estimate	Std. Error
$\log(\lambda_{<1year})$	410814.56	...
$\log(\lambda_{1-4years})$	305337.95	28427.34
$\log(\lambda_{5-14years})$	578647.21	64256.49
$\log(\lambda_{15-24years})$	782223.95	96931.77
$\log(\lambda_{25-34years})$	148651.67	13670.29
$\log(\lambda_{35-44years})$	738191.24	113778.60
$\log(\lambda_{45-54years})$	305727.09	31939.10
$\log(\lambda_{55-64years})$	457519.57	53292.63
$\log(\lambda_{65-74years})$	166840.39	16968.32
$\log(\lambda_{75-84years})$	280202.19	54641.59
$\log(\lambda_{85+years})$	917515.87	...
$\log(\sigma_u)$	0.45	0.02



## Hierarchical Poisson-Normal model

## Results

Shiga- og verotoxin producerende *E. coli*.

$$Y_i | u_i \sim \text{Pois}(\lambda_i u_i) \quad (2a)$$

$$u_i \sim G(1/\beta, \beta) \quad (2b)$$

$$\begin{aligned} P[Y = y] &= g_Y(y; \lambda, \beta) \\ &= \frac{\lambda^y}{y! \Gamma(1/\beta) \beta^{1/\beta}} \frac{\beta^{y+1/\beta} \Gamma(y + 1/\beta)}{(\lambda\beta + 1)^{y+1/\beta}} \\ &= \frac{\Gamma(y + 1/\beta)}{\Gamma(1/\beta) y!} \frac{1}{(\lambda\beta + 1)^{1/\beta}} \left( \frac{\lambda\beta}{\lambda\beta + 1} \right)^y \\ &= \binom{y + 1/\beta - 1}{y} \frac{1}{(\lambda\beta + 1)^{1/\beta}} \left( \frac{\lambda\beta}{\lambda\beta + 1} \right)^y, \text{ for } y = 0, 1, 2, \dots \end{aligned} \quad (3)$$

where we have used the convention

$$\binom{z}{y} = \frac{\Gamma(z + 1)}{\Gamma(z + 1 - y) y!} \quad (4)$$

The marginal distribution of  $Y$  is a negative binomial distribution,  $Y \sim \text{NB}(1/\beta, 1/(\lambda\beta + 1))$

The probability function for the conditional distribution of  $Y$  for given  $u$

$$f_{Y|u}(y; \lambda, u) = \frac{(\lambda u)^y}{y!} \exp(-\lambda u) \quad (5)$$

and the probability density function for the distribution of  $u$  is

$$f_u(u; \beta) = \frac{1}{\beta \Gamma(1/\beta)} \left( \frac{u}{\beta} \right)^{1/\beta-1} \exp(-u/\beta) \quad (6)$$

Given (5) and (6), the probability function for the marginal distribution of  $Y$  is determined from

$$\begin{aligned} g_Y(y; \lambda, \beta) &= \int_{u=0}^{\infty} f_{Y|u}(y; \lambda, u) f_u(u; \beta) du \\ &= \int_{u=0}^{\infty} \frac{(\lambda u)^y}{y!} \exp(-\lambda u) \frac{1}{\beta \Gamma(1/\beta)} \left(\frac{u}{\beta}\right)^{1/\beta-1} \exp(-u/\beta) du \\ &= \frac{\lambda^y}{y! \Gamma(1/\beta) \beta^{1/\beta}} \int_{u=0}^{\infty} u^{y+1/\beta} \exp(-u(\lambda\beta + 1)/\beta) du \end{aligned} \quad (7)$$

In (7) it is noted that the integrand is the *kernel* in the probability density function for a Gamma distribution,  $G(y + 1/\beta, \beta/(\lambda\beta + 1))$ . As the integral of the density shall equal one, we find by adjusting the norming constant that

$$\int_{u=0}^{\infty} u^{y+1/\beta} \exp\left(-u/(\beta/(\lambda\beta + 1))\right) du = \frac{\beta^{y+1/\beta} \Gamma(y + 1/\beta)}{(\lambda\beta + 1)^{y+1/\beta}} \quad (8)$$

and then (3) follows

**Inference on individual group means**

Consider the hierarchical Poisson-Gamma model in (2), and assume that a value  $Y = y$  has been observed. Then the conditional distribution of  $u$  for given  $Y = y$  is a Gamma distribution,

$$u|Y = y \sim G(y + 1/\beta, \beta/(\lambda\beta + 1)) \quad (9)$$

with mean

$$E[u|Y = y] = \frac{y\beta + 1}{\lambda\beta + 1} \quad (10)$$

The conditional distribution is found using Bayes Theorem

$$\begin{aligned}
 g_u(u|Y = y) &= \frac{f_{y,u}(y, u)}{g_Y(y; \lambda, \phi)} \\
 &= \frac{f_{y|u}(y; u)g_u(u)}{g_Y(y; \lambda, \beta)} \\
 &= \frac{1}{g_Y(y; \lambda, \beta)} \left( \frac{(\lambda u)^y}{y!} \exp(-\lambda u) \frac{1}{\beta \Gamma(1/\beta)} \left( \frac{u}{\beta} \right)^{1/\beta-1} \exp(-u/\beta) \right) \\
 &\propto u^{y+1/\beta-1} \exp(-u(\lambda\beta + 1)/\beta)
 \end{aligned} \tag{11}$$

We identify the *kernel* of the probability density function

$$u^y \exp(-u(\lambda + 1)/\phi) \tag{12}$$

as the kernel of a Gamma distribution,  $G(y + 1/\beta, \beta/(\lambda\beta + 1))$