

# Automated and Early Detection of Disease Outbreaks

## AEDDO

Master Thesis





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Master Thesis  
August, 2023

By  
Kasper Schou Telkamp

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## **Approval**

This thesis has been prepared over six months at the Section for Dynamical Systems, Department of Applied Mathematics and Computer Science, at the Technical University of Denmark, DTU, in collaboration with Epidemiologisk Forskning / Modelgruppen at Statens Serum Institut, SSI, in partial fulfilment for the degree Master of Science in Engineering, MSc Eng., Quantitative Biology and Disease Modelling.

It is assumed that the reader has a basic knowledge in the areas of statistics.

Kasper Schou Telkamp - s170397

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*Signature*

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## **Abstract**

Hello, here is some text without a meaning. This text should show what a printed text will look like at this place. If you read this text, you will get no information. Really? Is there no information? Is there a difference between this text and some nonsense like "Huardest gefburn"? Kjift – not at all! A blind text like this gives you information about the selected font, how the letters are written and an impression of the look. This text should contain all letters of the alphabet and it should be written in of the original language. There is no need for special content, but the length of words should match the language.

## Acknowledgements

**Lasse Engbo Christiansen**, Senior Researcher, Statens Serum Institut

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**Jan Kloppenborg Møller**, Associate Professor, Technical University of Denmark

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## Notation

The mathematical notation in this master's thesis is adapted from Madsen and Thyregod 2011. All vectors are column vectors. Vectors and matrices are emphasized using a bold font. Lowercase letters are used for vectors and uppercase letters are used for matrices. Transposing is denoted with the upper index  $T$ . Random variables are always written using uppercase letters. Thus, it is not possible to distinguish between a multivariate random variable and a matrix. However, variables and random variables are assigned to letters from the last part of the alphabet ( $X, Y, Z, U, V, \dots$ ), while constants are assigned to letters from the first part of the alphabet ( $A, B, C, D, \dots$ ). From the context it should be possible to distinguish between a matrix and a random vector.



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# 1 Introduction

This master's thesis will primarily focus on diseases where timely detection of outbreaks results in interventions that are feasible to impose. However, the detection algorithms discussed in this master's thesis are not restricted to these diseases and can be used to detect aberrant counts in any type of disease.

Today, outbreaks of foodborne illnesses can be detected in several ways:

1. Reports from doctors: Physicians may report cases of foodborne illnesses they encounter in their practice to the relevant authorities.
2. Citizen reports: Individuals may directly contact food or health authorities to report suspected cases of foodborne illnesses.
3. Cluster identification through laboratory surveillance: Clusters of cases can be identified through routine laboratory testing and surveillance of samples from patients with suspected foodborne illnesses.
4. Identification of identical "fingerprints": When bacteria or viruses are type-tested, the presence of identical fingerprints among multi cases can strongly indicate a common source of infection.

These various methods help in the early detection and investigation of foodborne disease outbreaks, enabling timely intervention and prevention measures. However, ...

While generalized linear mixed effects models and hierarchical models have earned a reputation within ..., they are fairly unproven in the field of automatic detection of disease outbreaks.

The novel outbreak detection algorithms discussed in this master's thesis are open-source and can be found at <https://github.com/telkamp7/AEDDO>

## 2 Setting the scene

Outbreak investigations have a long history, dating back to John Snow's iconic removal of the handle of London's Broad Street pump during the cholera epidemic in 1865 (Tulchin-sky 2018). Indeed, while John Snow's work was groundbreaking for his time, modern disease outbreak investigations require more advanced and sophisticated techniques. Today, epidemiologists and public health professionals utilize a range of tools and methodologies to effectively tackle disease outbreaks. These may include advanced statistical analysis, mathematical modeling, molecular epidemiology (Honardoost, Rajabpour, and Vakil 2018; Struelens and Brisson 2013), and whole-genome sequencing (WGS) (Koeser et al. 2012; Baldry 2010).

Meanwhile, the focus of this master's thesis is primarily on the field of advanced statistical analysis methods.

In recent years, there has been a notable increase in the interest surrounding statistical methods for automated and early detection of infectious disease outbreaks. These methodologies encompass a wide range of statistical techniques, including regression analysis, time series methodology, methods inspired by statistical process control, approaches incorporating spatial information, and multivariate outbreak detection. A comprehensive review of these methods can be found in Buckeridge (2007) and Unkel et al. (2012).

In addition to the aforementioned studies, state-of-the-art methods for aberration detection is presented in Salmon, Schumacher, and Höhle (2016) and implemented in the R package called **surveillance**. The **surveillance** package provides various techniques for detecting aberrations in disease surveillance data, including the Farrington method initially introduced by Farrington et al. (1996) and subsequent improvements proposed by Noufaily et al. (2013). These methods offer advanced statistical tools for detecting and monitoring disease outbreaks and are currently *the* method choice at European public health institutes (Hulth et al. 2010).

Therefore, in this master's thesis, these established methods will be compared to a novel outbreak detection algorithm based on generalized mixed effects models and hierarchical generalized linear models respectively. The thesis introduces this new algorithm as an innovative approach to outbreak detection and aims to assess its performance in comparison to existing methods.



# 3 Surveillance data in Denmark

This chapter delves into the data collection methods and quality assurance procedures within the Danish surveillance system. Moreover, it introduces the case studies selected for this master's thesis, which include *Listeriosis*, *Shigellosis*, Shiga toxin (verotoxin)-producing *Escherichia coli* (STEC), and *Salmonellosis*. These diseases will be referred to using either their full name or their related case definition:

- LIST: *Listeriosis*
- SHIG: *Shigellosis*
- STEC: Shiga toxin (verotoxin)-producing *Escherichia coli*
- SALM: *Salmonellosis*

## 3.1 Data collection and data quality

In Denmark, the surveillance of infectious diseases is conducted by Statens Serum Institut (SSI). This surveillance system plays a pivotal role in national and international disease preparedness. It encompasses more than just the collection and registration of disease data; it also involves the prompt and ongoing dissemination of knowledge to the relevant authorities responsible for treatment, prevention, and control. This comprehensive approach ensures efficient communication and facilitates appropriate measures to address infectious diseases.

The quality of the Danish surveillance registers is maintained at a high standard, thanks to The National Board of Health Statutory Order on Physicians' Notification of Infectious Diseases (<https://www.retsinformation.dk/eli/ita/2000/277>). This order specifies that several diseases<sup>1</sup> are individually notifiable by physicians and general practitioners. Notifications consist of essential patient information and are submitted in paper form to both the Ministry of Health and to SSI. This rigorous notification process ensures accurate and comprehensive data collection for disease surveillance purposes in Denmark.

In addition to the individually notifiable diseases, SSI has implemented a laboratory notification system for numerous microorganisms. Clinical-microbiological laboratories are obligated to report the identification of specific microorganisms, along with relevant patient information. These data are then stored in the Danish Microbiology Database (MiBa), which was established by SSI in 2010. MiBa is a nationwide and automatically updated database specifically designed to collect and store microbiological test results. In order to utilize the data from MiBa, the information in the test results needs to have a standardized structure with common codes and terminology. MiBa employs national standards to harmonize the data, which initially may be structured in diverse formats. The standards currently used are XRPT05, which is widely employed for the exchange of microbiological test results in the healthcare system, and a specific standard called XRPT06. These standards are regularly revised and exist in various versions<sup>2</sup>. The national surveillance system focuses on diseases of a severe nature, those that are highly contagious, and the majority of vaccine-preventable diseases.

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<sup>1</sup>For a full list of diseases see [https://www.ssi.dk/sygdomme-beredskab-og-forskning/anmeldelse-af-sygdomme/lovpligtige-meldesystemer/individ\\_anmeldelses\\_sygdomme](https://www.ssi.dk/sygdomme-beredskab-og-forskning/anmeldelse-af-sygdomme/lovpligtige-meldesystemer/individ_anmeldelses_sygdomme)

<sup>2</sup>Information on national standards and codes within the healthcare domain can be found on MedCom's website (<https://medcom.dk/>).

## 3.2 Introducing the case studies

For the scope of this master's thesis, only a specific subset of diseases from the mandatory notification system will be considered. This subset consists of *Listeriosis*, *Shigellosis*, Shiga toxin (verotoxin)-producing *Escherichia coli* (STEC), and *Salmonellosis*. These diseases have been chosen for analysis and investigation based on various factors, such as seasonality, incidence, and severity. Additionally, these diseases have been associated with documented outbreaks observed by SSI (<https://www.ssi.dk/sygdome-beredskabog-forskning/sygdomsudbrud/arkiv>) in the past decade, which adds to their relevance for the study.

The count observations are observed in the period from January 2008 to December 2022. An epidemic curve graph for an excerpt of the data for each of the diseases considered in this master's thesis is shown in Figure 3.1.

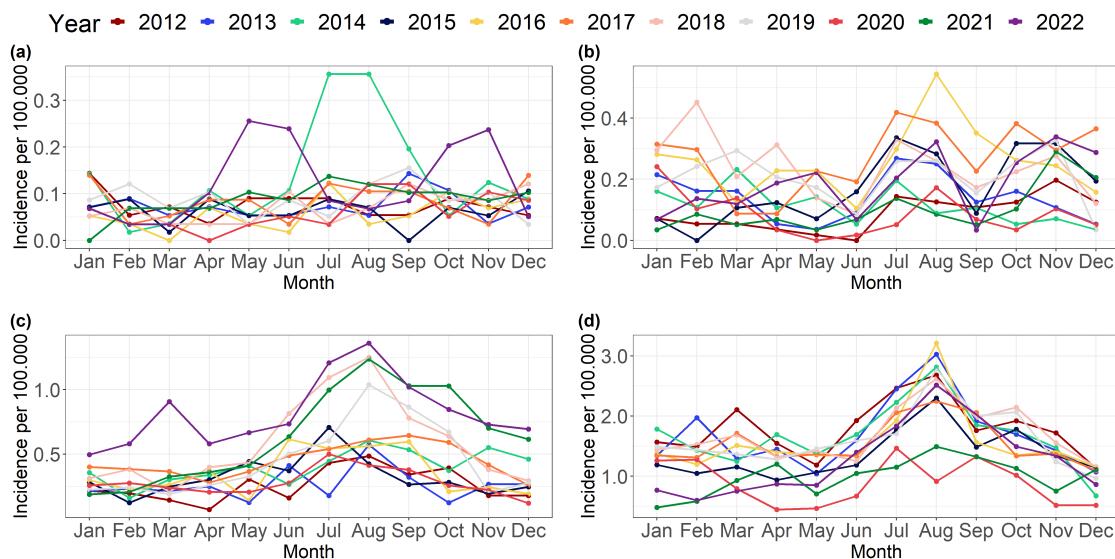


Figure 3.1: Epidemic curve showing the incidence per 100,000 in Denmark, 2012-2022, for the subset of diseases considered in this master thesis. **(a)** Listeriosis, **(b)** Shigellosis, **(c)** STEC, and **(d)** Salmonellosis.

Evidently, the subset of diseases in Figure 3.1 exhibits varying incidences and different levels of seasonal patterns on an annual basis. It is interesting to note that the incidences peak in August, which can be attributed to several factors, including:

- Increased travel activity: Especially when individuals travel to countries with unsafe drinking water, poor sanitation, and insufficient hygiene practices.
- Large social gatherings: Events such as weddings, large festivals, or other gatherings where a significant number of people consume potentially contaminated food or drinking water.

Moreover, one could hypothesize that the warmer climate during the summer could potentially directly influence the proliferation of bacteria, which in turn may have an indirect impact on the transmission and spread of these diseases.

In general, there is a noticeable decrease in the number of observed cases starting from March 2020 and continuing until January 2021. This decline can be attributed to the strict lockdown measures implemented in Denmark in response to the Covid-19 pandemic.

These measures, which involved restrictions on movement and social interactions, likely played a significant role in reducing the transmission of infectious diseases, including the ones being investigated.

In Figure 3.1a...

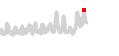
In Figure 3.1b...

Furthermore, in Figure 3.1c, a significant increase in the amplitude of the seasonal variation can be observed starting from 2018, with incidences doubling compared to the preceding years. At a first glance, this increase in the incidences might be recognized as a serious, reoccurring outbreak of the disease, but a more reasonable explanation can be found. Up to 2018, most departments of clinical microbiology used culture-based methods as a diagnostic test for bacterial pathogens and the process of changing the test method to polymerase chain reaction (PCR) methods was ongoing (Svendsen et al. 2023). In general, PCR resulted in higher incidences compared to other test methods, which is no surprise as higher sensitivity is well documented for PCR (Buss et al. 2015; Knabl, Grutsch, and Orth-Höller 2016).

Figure 3.1d ...

Some summary statistics for each of the disease considered in this master's thesis is gathered in Table 3.1.

Table 3.1: Summary statistics of the monthly count observations for the subset of diseases considered in this master's thesis. Boxplot: median (red line), IQR (grey box), whiskers (1.5 IQR), outliers (points). Time series: normalized observations (0-1), first time points minimum and maximum count (red)

Case definition	Min	Max	Mean	Median	Std. Deviation	Boxplot	Time series
LIST	0	20	4.446	4.0	2.934	•	
SHIG	0	28	9.078	8.0	5.652	•	
STEC	2	80	21.933	16.5	14.902	•	
SALM	0	583	108.012	85.5	82.725	—	

In Table 3.1, it is readily seen that the diseases exhibit different properties. WRITE SOME MORE!

All the diseases within the selected subset pose a significant risk of infection and can vary in terms of severity for affected individuals. Therefore, early identification of disease outbreaks is of utmost importance in order to promptly implement necessary interventions. Timely detection allows for swift and targeted actions to control the spread of these diseases and mitigate their impact on public health.

### 3.2.1 *Listeriosis*

*Listeriosis* is a foodborne illness that is caused by consuming food contaminated with *Listeria monocytogenes*. This disease primarily affects pregnant women, unborn or newborn babies, the elderly, and individuals with weakened immune systems. *Listeriosis* is associated with high mortality (Goulet et al. 2012) and manifests in three ways: sepsis, meningitis, and mother-to-child transmission. Pregnancy-associated *Listeriosis* can have severe consequences for the fetus or newborn, including miscarriage, stillbirth, neonatal sepsis, and meningitis (Awofisayo et al. 2015). *Listeriosis* is uncommon among

individuals in other demographic groups. The bacteria is ubiquitous in the environment, found in moist environments, soil, water, decaying vegetation, and animals. Furthermore, it can survive and even grow under refrigeration and other food preservation measures.

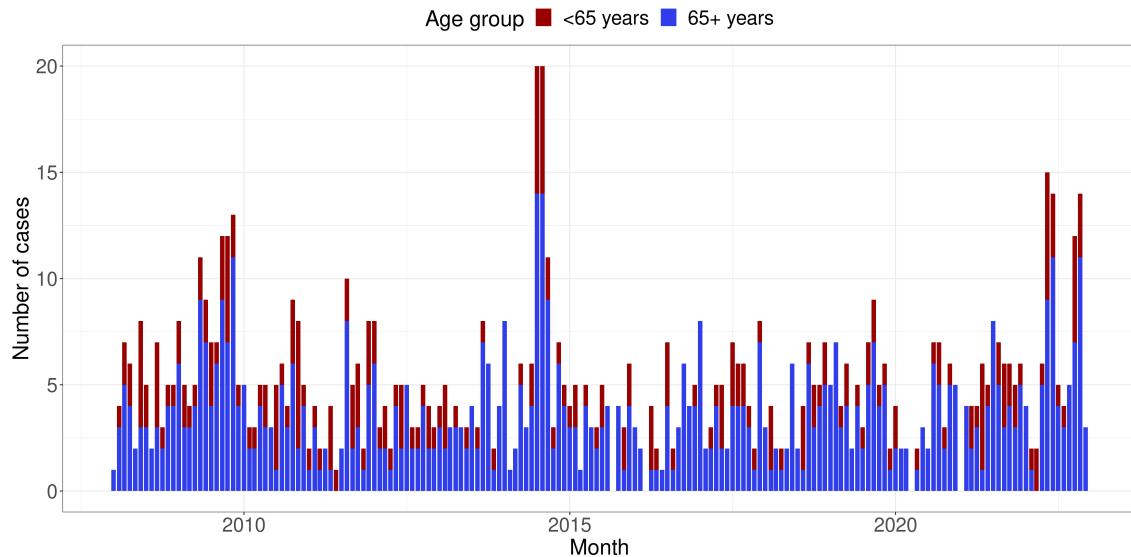


Figure 3.2: Placeholder caption

In general, SSI employs whole-genome sequencing (WGS) as the state-of-the-art method to detect disease outbreaks caused by *Listeria monocytogenes*. This method involves mapping the entire DNA of the bacteria and enables SSI to identify cases where patients are infected with identical Listeria bacteria. However, it is important to note that for this master's thesis, the DNA typing data is unavailable for use.

One notable outbreak investigated by SSI occurred between September 2013 and October 2014. This *Listeriosis* outbreak involved a total of 41 cases, resulting in 17 deaths. Deli meat products from a specific company were identified as the source of the outbreak. The high mortality rate may be attributed to the consumption of these products in nursing homes and hospitals, where patients are more vulnerable. Following the discovery of Listeria at the facility, the Danish Veterinary and Food Administration recalled all products from the company.

In another *Listeriosis* outbreak investigated by SSI, the source was traced back to cold-smoked and cured salmon products. A total of 5 related cases were identified, with 4 of them occurring in August 2017, and the fifth case in May 2017.

In some cases, despite extensive investigations, the source of contamination in an outbreak cannot always be identified. Such was the case in an unresolved outbreak that took place between the 13th of May and the 6th of June, 2022. During this period, a total of nine cases were infected with the same type of Listeria, with the majority of affected patients located in the Capital Region of Denmark. Despite thorough efforts, the specific source of contamination remained unknown.

Early identification of outbreaks caused by *Listeria monocytogenes* is crucial to implement timely interventions and mitigate the impact of the disease. Otherwise, these outbreaks can persist over an extended period. SSI has successfully resolved several long-spanned outbreaks in the last decade. For example, one investigation revealed that a single out-

break was actually two simultaneous outbreaks caused by the consumption of smoked fish. Each outbreak consisted of ten cases and spanned from May 2013 to July 2015 (Gillesberg Lassen et al. 2016).

Other documented long-spanned outbreaks investigated by SSI include:

- A cold-smoked fish outbreak with 9 cases spanning from December 2016 to February 2019.
- A prolonged outbreak with 6 cases from 2016 to 2019, traced back to a local green-grocer.
- An outbreak with 8 cases from October 2021 to June 2022, caused by a deli meat product.
- Two unresolved outbreaks with 9 cases and 12 cases from the end of 2018 to November 2021 and October 2020 to May 2022, respectively.

The use of WGS in these outbreaks provided the ability to link cases that occurred over a period of years and revealed that they were, in fact, continuous-source outbreaks.

In a recent outbreak investigated by SSI, the Danish Veterinary and Food Administration, and the National Food Institute at the Technical University of Denmark, fish patties were identified as the source of contamination. This outbreak occurred from August 2022 to December 2022 and affected a total of 11 cases.

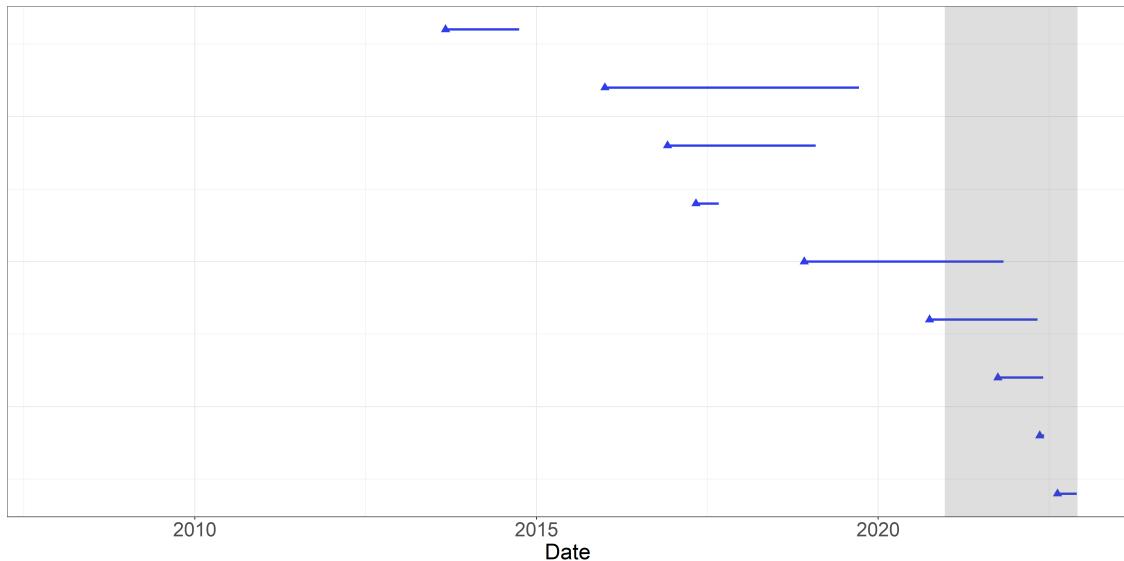


Figure 3.3: Timeline indicating the start (triangle) and duration (line) of documented outbreaks of *Listeriosis* by SSI and possible collaborators.

### 3.2.2 *Shigellosis*

*Shigellosis* is a diarrheal illness that is caused by a group of bacteria called *Shigella*. The bacteria are highly contagious and can be transmitted through direct person-to-person contact, consumption of contaminated food, or ingestion of water contaminated with human feces. *Shigella* infections are most commonly observed in children under the age of 5, individuals traveling to regions with poor sanitation and unsafe water and food practices, as well as gay, bisexual, and other men who have sex with men.

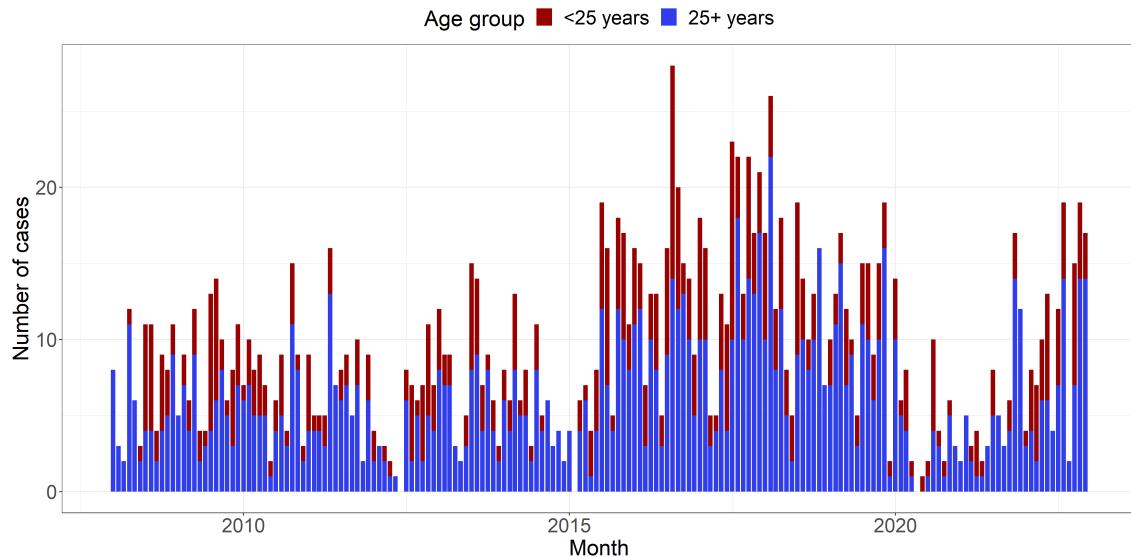


Figure 3.4: Placeholder caption

In Denmark, another significant cause of *Shigellosis* outbreaks is the importation of contaminated vegetables. This was evident in several incidents, including a 2007 outbreak where 215 individuals fell ill after consuming imported contaminated baby corn, a smaller outbreak in 2009 linked to sugar snap peas from Kenya, and a 2020 outbreak associated with fresh mint as the source of infection.

The 2020 outbreak is indeed a significant focus of this study, as it serves as a benchmark for evaluating the effectiveness of outbreak detection algorithms. It took place from the 22th of August to the 9th of September and was investigated by SSI in collaboration with the Danish Veterinary and Food Administration and the National Food Institute at the Technical University of Denmark. The outbreak affected 44 patients, mainly concentrated in the Capital Region of Denmark. During the investigation, at least five events were identified where individuals subsequently developed *Shigellosis*.

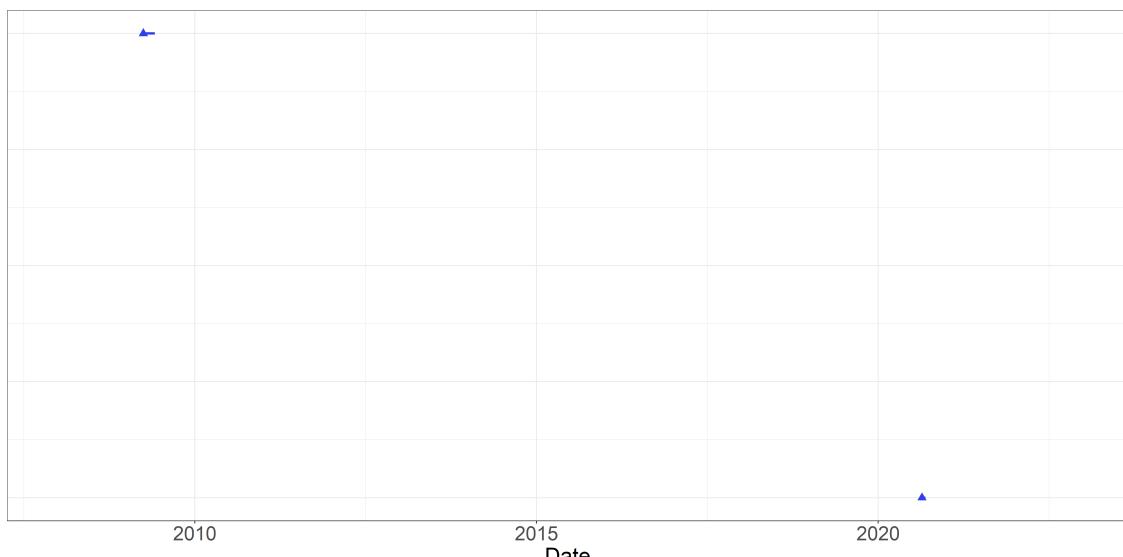


Figure 3.5: Placeholder caption

### 3.2.3 Shiga toxin (verotoxin)-producing *Escherichia coli*

STEC primarily spreads through contaminated food. Less common sources of infection include contaminated drinking and bathing water, as well as direct or indirect contact with infected animals. Cattle and other ruminants are primary reservoirs for STEC serotypes that are frequently associated with human disease (Menge 2020). Therefore, in Denmark, the source of infection is often products derived from beef, non-heat-treated dairy products, or other foods such as ready-to-eat vegetables, leafy greens, vegetable sprouts, and berries contaminated with feces from cows. *Hemolytic uremic syndrome* (HUS) is a severe complication that, in some cases, particularly in children, can develop following an infection with STEC.

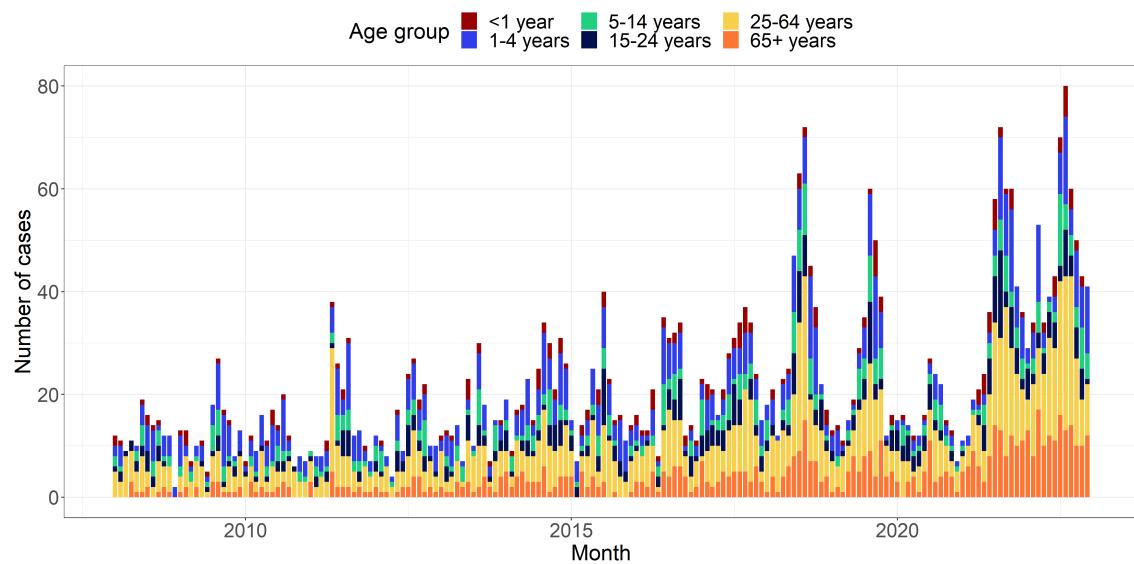


Figure 3.6: Placeholder caption

In general, stool samples are commonly used for diagnostic purposes in cases of STEC infections. Until 2018, most clinical microbiology departments relied on culture-based methods to detect and identify STEC bacteria in stool samples. However, in recent years, PCR methods have been increasingly adopted as a replacement for culture-based methods in the diagnosis of STEC infections (Svendsen et al. 2023). PCR methods offer advantages such as increased sensitivity and faster turnaround time, contributing to their growing popularity in clinical laboratories.

It is important to note that not all patients are routinely tested for STEC, and therefore, physicians need to specifically request STEC testing when submitting stool samples.

One of the earliest documented STEC outbreaks occurred in 2007, involving 18 laboratory-confirmed cases over a six-week period. The outbreak primarily affected children in day-care settings, and most patients experienced mild symptoms without bloody diarrhea. Investigations indicated a specific brand of organic beef sausage as the likely source of infection.

In September to October 2012, a STEC outbreak with a high risk of HUS was observed. Thirteen cases were diagnosed, with eight individuals developing HUS. Epidemiological investigations suggested that ground beef was the vehicle of the outbreak (Soborg et al. 2013).

More recent outbreaks include a 38-case outbreak from September to November 2018,

with a suspected association with beef sausage as the source of infection. Additionally, there were two unresolved outbreaks with 11 and 14 cases occurring from May to July 2019 and from December 2021 to January 2022, respectively. The latter outbreak included three cases of HUS.

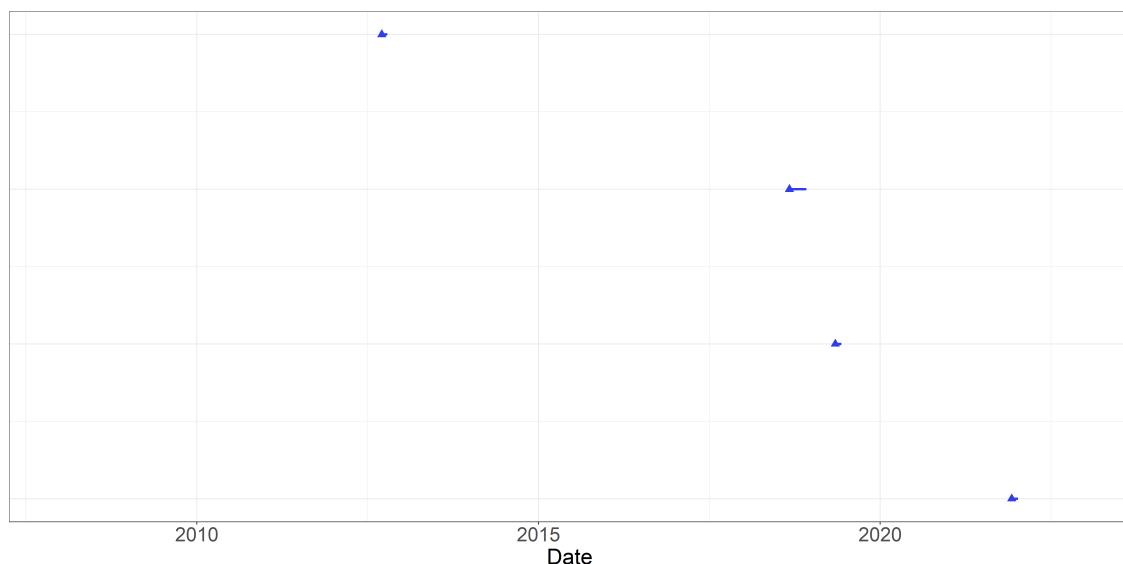


Figure 3.7: Placeholder caption

### 3.2.4 *Salmonellosis*

*Salmonellosis* is a bacterial disease that primarily affects the intestinal tracts of humans. The *Salmonella* bacteria are commonly found in the intestines of animals and humans and are excreted in feces. Human infection typically occurs through the consumption of contaminated food or water. *Salmonella* infections are often associated with the consumption of raw or undercooked meat, poultry, eggs or egg products, as well as unpasteurized milk.

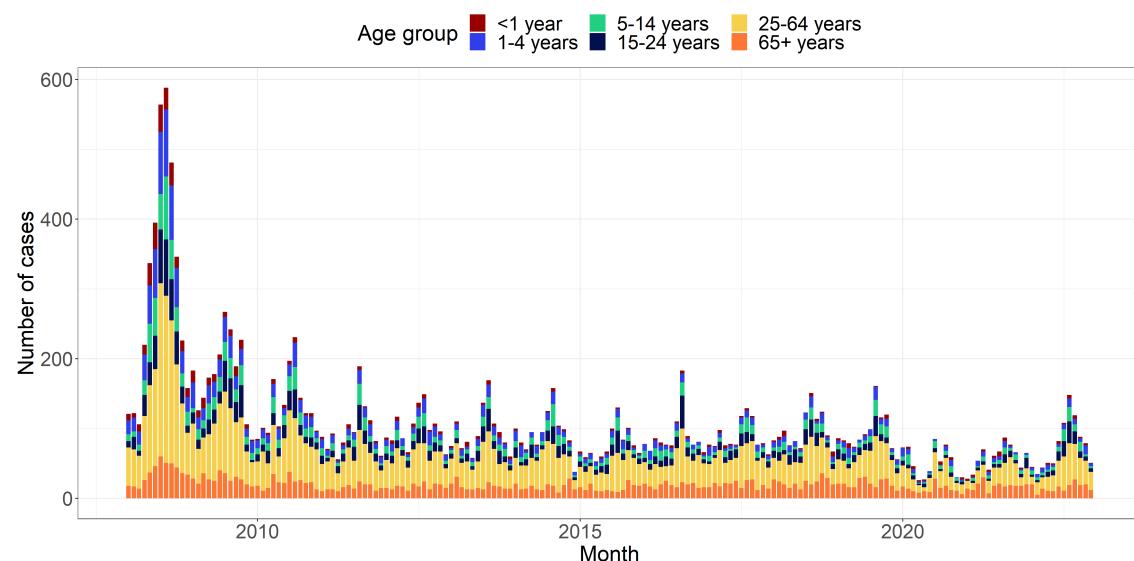


Figure 3.8: Placeholder caption

It is worth noting that an increasing proportion of infections in Denmark are now observed

in connection with international travel, particularly since *Salmonella* has been eliminated from commercial chicken flocks in Denmark, making Danish eggs and poultry meat free from the bacteria. However, imported meat products can still pose a risk of contamination.

In 2015, three outbreaks of *Salmonellosis* with patients in two or more regions were investigated (Helwig, Christensen, and Müller 2016):

- An outbreak caused by *S. Newport* affecting 6 people in the period from March to April 2015.
- A long-lasting outbreak caused by *S. Oranienburg* with 14 genetically linked cases from July 2015 to January 2016.
- An outbreak with 6 patients from November 2015 to January 2016.

Other resolved and well-documented outbreaks include:

- An outbreak with 49 cases from October 2018 to January 2019, where Mediterranean sausage was identified as a possible source of contamination.
- An outbreak with 45 cases from November 2020 to April 2021, linked to the consumption of the natural remedy HUSK Psyllium.
- An international outbreak from March to July 2021, with more than 300 cases in Europe, including 39 cases in Denmark. Imported melons were suspected as the source of infection.
- An outbreak caused by eggs from a Danish producer, resulting in 24 cases registered from September to November 2021.
- An international outbreak with a total of 392 cases across 12 countries in the EU-/EEA and the UK, including 4 cases in Denmark. Kinder chocolate products were identified as the source of infection.

There were also outbreaks where it was not possible to identify the source of contamination, including:

- An outbreak with 26 cases in the period from May to August 2019.
- An outbreak with 11 cases in the period from June to July 2020.
- An outbreak with 24 cases in the period from March to September 2022.
- An outbreak with 15 cases in the period from August to September 2022.

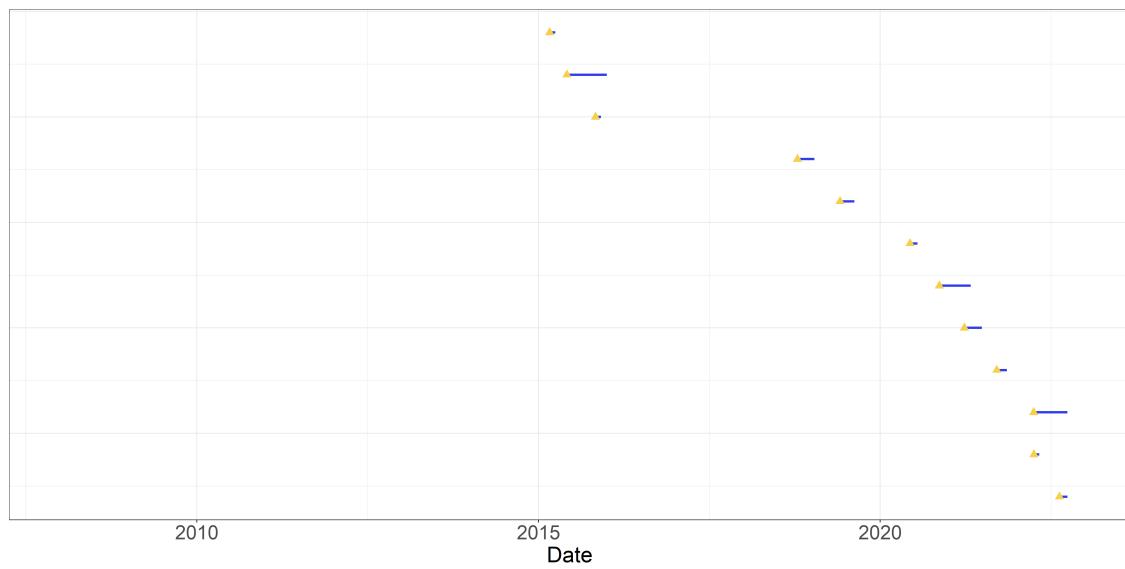


Figure 3.9: Placeholder caption

# 4 Methods

In this chapter, the current state-of-the-art methods for disease outbreak detection will be outlined. Furthermore, the novel outbreak detection algorithm will be introduced, along with the theory related to generalized mixed effects models and hierarchical generalized linear models. These models are utilized in this master's thesis to analyze the count observations denoted as  $y$ , but more importantly they play a crucial role in assessing the unobserved random variables or random effects represented by  $u$ , which are directly employed in the detection algorithm for characterizing outbreaks.

Due to the complexity of generalized mixed effects models, obtaining closed-form solutions is generally not feasible. Therefore, an overview of the Laplace approximation technique will be provided in this chapter, which allows for approximating the likelihood function in these models. Additionally, the implementation of these models in the R programming language will be presented. The presentation of this chapter is mostly inspired by Madsen and Thyregod (2011).

## 4.1 State-of-the-art outbreak detection algorithm

In this section, the Farrington method introduced by Farrington et al. (1996) and the subsequent improvements proposed by Noufaily et al. (2013) will be outlined. These methods are recognized as the current state-of-the-art for disease outbreak detection and will be used as benchmarks to evaluate the performance of the novel outbreak detection algorithm proposed in this master's thesis. Both of the aforementioned methods have been implemented in the R package called **surveillance** by Salmon, Schumacher, and Höhle (2016), which can be accessed from the Comprehensive R Archive Network (CRAN) at <https://cran.r-project.org/web/packages/surveillance/index.html>. The presentation of these methods are strongly inspired by Salmon, Schumacher, and Höhle (2016).

Both methods follow the same steps in the algorithm. The first step involves fitting an over-dispersed Poisson generalized linear model (GLM) with a log link to the reference data  $y_{t_0} \subseteq \{y_t; t \leq t_0\}$ . In this model, the baseline count  $y_t$  corresponding to the baseline time point  $t$  is assumed to have an expected value  $\lambda_t$  and a variance  $\phi\lambda_t$ , where  $\phi \geq 1$  is ensured to account for over-dispersion. The systematic component of the model includes only a linear time trend in the frequency of reports. Therefore, the systematic component can be expressed as

$$\log(\lambda_t) = \alpha + \beta t \quad (4.1)$$

The original method incorporated seasonal effects by considering counts from comparable periods in past years for threshold calculation. This approach is similar to the one used by Stroup et al. (1989). The baseline weeks, which are used as reference, are determined by two integers:  $b$  represents the number of years back, and  $w$  represents the window half-width. For a given current week  $x$  of year  $y$ , only data from weeks  $x - w$  to  $x + w$  of years  $y - b$  to  $y - 1$  are considered, resulting in a total of  $n = b(2w + 1)$  baseline weeks. The default values are  $b = 5$  and  $w = 3$ , resulting in a total of  $n = 35$  baseline values.

However, Noufaily et al. (2013) demonstrated that the algorithm performs better when utilizing more historical data without disregarding seasonality. To achieve this, the author introduced a 10-level factor with a 7-week reference period and nine additional 5-week

periods in each year. As a result, the systematic component of the model is modified as follows

$$\log(\lambda_t) = \alpha + \beta t + \delta_{j(t)} \quad (4.2)$$

In this equation,  $j(t)$  represents the seasonal factor level corresponding to time point  $t$ . The reference week  $t_0$  is always associated with the reference seasonal level, denoted by  $j(t_0) = 0$  and  $\delta_0 = 0$ .

The idea of incorporating more data while preserving seasonality has been further expanded in the implementation of the method in the **surveillance** R package. The package allows the user to choose an arbitrary number of periods in each year. Consequently, the systematic component is adjusted as follows

$$\log(\lambda_t) = \alpha + \beta t + \delta_{c(t)} \quad (4.3)$$

In this equation,  $c(t)$  represents the coefficients of a zero-order spline with `noPeriods` + 1 knots. It can be conveniently represented as a `noPeriods`-level factor that captures seasonality. The function  $c(t)$  indicates which season or period of the year  $t$  belongs to.

Furthermore, Noufaily et al. (2013) demonstrated that it is beneficial to exclude the last 26 weeks before  $t_0$  from the baseline calculation. This exclusion helps prevent a reduction in sensitivity when an outbreak has recently started before  $t_0$ .

In the second step, the algorithm predicts the expected number of counts  $\lambda_{t_0}$  for the current time point  $t_0$  using the fitted generalized linear model. Both methods differ in their assumptions for calculating the upper bound  $U_{t_0}$ .

The original method assumes that a transformation of the prediction error, denoted as  $g(y_{t_0} - \hat{\lambda}_{t_0})$ , follows a normal distribution. For example, when using the identity transformation  $g(x) = x$ , the assumption becomes

$$y_{t_0} - \hat{\lambda}_{t_0} \sim N(0, V(y_{t_0} - \hat{\lambda}_{t_0})) \quad (4.4)$$

The upper bound of the prediction interval is then calculated based on this distribution. The variance of the prediction error is given by

$$V(y_{t_0} - \hat{\lambda}_{t_0}) = V(\hat{y}_{t_0}) + V(\hat{\lambda}_{t_0}) = \phi \lambda_{t_0} \quad (4.5)$$

Here,  $V(\hat{y}_{t_0})$  represents the variance of an observation, and  $V(\hat{\lambda}_{t_0})$  represents the variance of the estimate. The threshold, defined as the upper bound of a one-sided  $(1 - \alpha) \cdot 100\%$  prediction interval, is calculated as

$$U_{t_0} = \hat{\lambda}_{t_0} + z_{1-\alpha} \hat{V}(y_{t_0} - \hat{\lambda}_{t_0}) \quad (4.6)$$

However, this method's weakness lies in the assumption of normality itself. Therefore, an alternative assumption was presented in Noufaily et al. (2013). This approach assumes that  $y_{t_0}$  follows a negative binomial distribution, denoted as  $NB(\lambda_{t_0}, \nu)$ , where  $\lambda_{t_0}$  represents the mean and  $\nu = \frac{\lambda_{t_0}}{\phi - 1}$  represents the over-dispersion parameter. In this parameterization, the expected value of  $y_t$  remains  $\lambda_t$ , and the variance of  $y_t$  is  $\phi \lambda_t$ , with

$\phi > 1$ . If  $\phi \leq 1$ , a Poisson distribution is assumed for the observed count. The threshold is defined as a quantile of the negative binomial distribution using the plug-in estimates  $\hat{\lambda}_{t_0}$  and  $\hat{\phi}$ .

In the final step, the observed count  $y_{t_0}$  is compared to the upper bound  $U_{t_0}$ , and an alarm is raised if  $y_{t_0} > U_{t_0}$ . The fitting of the GLM in both methods involves three important steps.

First, the algorithm optionally performs a power transformation to correct for skewness and stabilize the variance of the data.

Next, the significance of the time trend is checked. The time trend is included in the model only if it is statistically significant at a chosen significance level, there are more than three years of reference data, and there is no over-extrapolation due to the time trend.

Finally, past outbreaks are reweighted based on their Anscombe residuals. If the Anscombe residual of a count exceeds a certain weight threshold, it is reweighted in a second fitting of the GLM. In the original method by Farrington et al. (1996), a reweighting threshold of 1 was used. However, Noufaily et al. (2013) suggests using a value of 2.56 for the weight threshold to make the reweighting procedure less drastic, as it also reduces the variance of the observations.

## 4.2 Novel outbreak detection algorithm

In this section, the novel algorithm for the prospective detection of disease outbreaks proposed in this master's thesis is outlined. The algorithm utilizes a generalized mixed effects model or a hierarchical generalized linear model to model the count observations  $y$  and assess the unobserved random effects  $u$ . These random effects are used directly in the detection algorithm to characterize an outbreak. The theoretical foundations of these models will be further discussed in Section 4.3 and Section 4.4.

The first step involves fitting either a hierarchical Poisson Normal model or a hierarchical Poisson Gamma model with a log link to the reference data  $y_{t_0} \subseteq \{y_t; t \leq t_0\}$ . Here, it is possible for the user to include an arbitrary number of covariates by supplying a model formula. In order to account for structural changes in the time series, e.g. an improved and more sensitive diagnostic method or a new screening strategy at hospitals, a rolling window with width  $k$  is used to estimate the model parameters. Also, it is assumed that the count is proportional to the population size  $n$ . Hence, in terms of the canonical link the model for the fixed effects is

$$\log(\lambda_{it}) = x_{it}\beta + \log(n_{it}), \quad i = 1, \dots, m, \quad t = T, \dots, 0 \quad (4.7)$$

Here  $x_{it}$  and  $\beta$  are  $p$ -dimensional vectors of covariates and fixed effects parameters respectively, where  $p$  denotes the number of covariates or fixed effects parameters,  $m$  denotes the number of groups, and  $T$  denotes the length of the period, i.e.  $T = t_0 - k$ .

In the second step, the algorithm infers the one-step ahead random effect  $u_{t_1}$  using the fitted model. The threshold for detecting outbreaks is defined as a quantile of the distribution of the random effects in the second stage model. This can be either a Gaussian distribution using the plug-in estimate  $\hat{\sigma}$  or a Gamma distribution using the plug-in estimate  $\hat{\phi}$ .

In the final step, the inferred random effect  $u_{t_1}$  is compared to the upper bound  $U_{t_0}$ , and an alarm is raised if  $u_{t_1} > U_{t_0}$ . If an outbreak is detected, the related observation is omitted

from the parameter estimation in the future. Thus, resulting in a smaller sample size for the rolling window until that specific observation is thrown away.

### 4.3 General mixed effects models

In this section selected theory related to generalized mixed effects models is presented. The general mixed effects model can be represented by its likelihood function

$$L_M(\theta; \mathbf{y}) = \int_{\mathbb{R}^q} L(\theta; \mathbf{u}, \mathbf{y}) d\mathbf{u} \quad (4.8)$$

where  $\mathbf{y}$  is the observed random variable,  $\theta$  is the model parameters to be estimated and  $\mathbf{U}$  is the  $q$  unobserved random variables. The likelihood function  $L$  is the joint likelihood of both the observed and the unobserved random variables. The likelihood function for estimating  $\theta$  is the marginal likelihood  $L_M$  obtained by integrating out the unobserved random variables. In general it is difficult to solve the integral in (4.8) if the number of unobserved random variables is more than a few and hence numerical methods must be used. Thus, an outline of the Laplace approximation is included in this section.

#### 4.3.1 Hierarchical models

It is useful to formulate the model as a hierarchical model containing a *first stage model*

$$f_{Y|u}(\mathbf{y}; \mathbf{u}, \beta) \quad (4.9)$$

which is a model for the observed random variables given the unobserved random variables, and a *second stage model*

$$f_U(\mathbf{u}; \Psi) \quad (4.10)$$

which is a model for the unobserved random variables. Here  $\beta$  represent the fixed effects parameters and  $\Psi$  is a model parameter. The total set of parameters is  $\theta = (\beta, \Psi)$ . Hence the joint likelihood is given as

$$L(\beta, \Psi; \mathbf{u}, \mathbf{y}) = f_{Y|u}(\mathbf{y}; \mathbf{u}, \beta) f_U(\mathbf{u}; \Psi) \quad (4.11)$$

To obtain the likelihood for the model parameters  $(\beta, \Psi)$  the unobserved random variables are integrated out. The likelihood function for estimating  $(\beta, \Psi)$  is as in (4.8) the marginal likelihood

$$L_M(\beta, \Psi; \mathbf{y}) = \int_{\mathbb{R}^q} L(\beta, \Psi; \mathbf{u}, \mathbf{y}) d\mathbf{u} \quad (4.12)$$

where  $q$  is the number of unobserved random variables, and  $\beta$  and  $\Psi$  are the parameters to be estimated.

### 4.3.2 Laplace Approximation

The Laplace approximation will be outlined in the following. A thorough description of the Laplace approximation in nonlinear mixed effects models is found in Wolfinger and Lin (1997).

For a given set of model parameters  $\theta$  the joint log-likelihood  $\ell(\theta, u, y) = \log(L(\theta, u, y))$  is approximated using a second order Taylor approximation around the optimum  $\tilde{u} = \hat{u}_\theta$  of the log-likelihood function w.r.t. the unobserved random variables  $u$ , i.e.,

$$\ell(\theta, u, y) \approx \ell(\theta, \tilde{u}, y) - \frac{1}{2}(u - \tilde{u})^T \mathbf{H}(\tilde{u})(u - \tilde{u}) \quad (4.13)$$

where the first-order term of the Taylor expansion disappears since the expansion is done around the optimum  $\tilde{u}$  and  $\mathbf{H}(\tilde{u}) = -\ell''_{uu}(\theta, u, y)|_{u=\tilde{u}}$  is the negative Hessian of the joint log-likelihood evaluated at  $\tilde{u}$ .

It is readily seen that the joint log-likelihood for the hierarchical model specified in Section 4.3.1 is

$$\ell(\theta, u, y) = \ell(\beta, \Psi, u, y) = \log f_{Y|u}(y; u, \beta) + \log f_U(u; \Psi) \quad (4.14)$$

which implies that the Laplace approximation becomes

$$\ell_{M,LA}(\theta, y) = \log f_{Y|u}(y; \tilde{u}, \beta) + \log f_U(\tilde{u}; \Psi) - \frac{1}{2} \log \left| \frac{\mathbf{H}(\tilde{u})}{2\pi} \right| \quad (4.15)$$

### 4.3.3 Formulation of the hierarchical model

**Theorem 4.1** (Hierarchical Poisson Normal model). *In order to simplify the notation, the probability density functions are presented for a specific observation and hence the subscripts indicating the group and time are omitted. The conditional distribution of the count observations is assumed to be a Poisson distribution with intensities  $\lambda$*

$$f_{Y|u}(y; u, \beta) = \frac{\lambda \exp(u)^y}{y!} \exp(-\lambda \exp(u)) \quad (4.16)$$

*Also, it is assumed that the count is proportional to the population size  $x$ . Hence, in terms of the canonical link for the Poisson distribution the model for the fixed effects is*

$$\log(\lambda_{it}) = \mathbf{X}_t^T \beta + \log(x_{it}), \quad i = 1, \dots, m, \quad t = 1, \dots, T \quad (4.17)$$

*The probability density function for the distribution of the random effects is assumed to follow a Gaussian distribution,  $u \sim N(0, \sigma^2)$ , i.e.*

$$f_U(u; \sigma) = \frac{1}{\sigma \sqrt{2\pi}} \exp\left(-\frac{u^2}{2\sigma^2}\right) \quad (4.18)$$

*where  $\sigma$  is a model parameter.*

*Henceforth, the total set of parameters are  $\theta = (\beta, \sigma)$  and the model can be formulated as a two-level hierarchical model*

$$Y|u \sim \text{Pois}(\lambda \exp(u)) \quad (4.19a)$$

$$u \sim N(\mathbf{0}, I\sigma^2) \quad (4.19b)$$

The joint likelihood for the count observations  $y$  and the random effects  $u$  becomes

$$L(\beta, \sigma; u_{it}, y_{it}) = \prod_{t=1}^T \prod_{i=1}^m \frac{(\lambda_{it} \exp(u_{it}))^{y_{it}}}{y_{it}!} \exp(-\lambda_{it} \exp(u_{it})) \prod_{t=1}^T \prod_{i=1}^m \frac{1}{\sigma \sqrt{2\pi}} \exp\left(-\frac{u_{it}^2}{2\sigma^2}\right) \quad (4.20)$$

## 4.4 Hierarchical generalized linear models

In this section selected theory related to hierarchical generalized linear models is presented. The model class was initially formulated by Lee and Nelder (1996) as a natural generalization of the generalized linear models to also incorporate random effects. A starting point in hierarchical modelling is an assumption that the distribution of random effects may be modeled by an exponential dispersion family. This family of models were first introduced by Fisher and Russell (1922), and has proven to play an important role in mathematical statistics because of their simple inferential properties. The exponential dispersion family considers a family of distributions, which can be written on the form

$$f_Y(y; \theta) = c(y, \lambda) \exp(\lambda\{\theta y - \kappa(\theta)\}) \quad (4.21)$$

Here the parameter  $\lambda > 0$  is called the *precision parameter*, which in some cases represents a known shape parameter as for the Gamma distribution. In other cases the precision parameter represents an over-dispersion that is not related to the mean. These distributions combine with the so-called *standard conjugate distributions* in a simple way, and lead to marginal distributions that may be expressed in a closed form suited for likelihood calculations.

### 4.4.1 Standard conjugate distribution

Now the general notion of a *standard conjugate distribution* for an exponential dispersion family is introduced.

Consider an exponential dispersion family  $ED(\mu, V(\mu)/\lambda)$  with density (4.21) for  $\theta \in \Omega$ . Let  $\mathcal{M} = \tau(\Omega)$  denote the mean value space for this family. Let  $m \in \mathcal{M}$  and consider

$$g_\theta(\theta; m, \gamma) = \frac{1}{C(m, \gamma)} \exp\left(\frac{\theta m - \kappa(\theta)}{\gamma}\right) \quad (4.22)$$

with

$$C(m, \gamma) = \int_{\Omega} \exp\left(\frac{\theta m - \kappa(\theta)}{\gamma}\right) d\theta \quad (4.23)$$

for  $\gamma \in \mathbb{R}_+$  for which the integral converges. Then (4.22) defines the density function of a probability distribution for  $\theta$ . This distribution is called the *standard conjugate distribution* for  $\theta$  corresponding to (4.21).

#### 4.4.2 Definition of the hierarchical generalized linear model

Consider a set of observations  $\mathbf{Y} = (Y_1, Y_2, \dots, Y_k)^T$  such that for a given value of a parameter  $\theta$  the distribution of  $Y_i$  is given by an exponential dispersion model with density (4.21) and with canonical parameter space  $\Omega$  (for  $\theta$ ), mean value  $\mu = \kappa'(\theta)$ , mean value space  $\mathcal{M}$  (for  $\mu$ ) and canonical link  $\theta = g(\mu)$ .

Let the *conjugate distribution* of  $\theta$  be given by (4.22), and the corresponding conjugate distribution of  $\mu$ , (e.g.,  $f_Y(y)$  Poisson distribution;  $g_\mu(\mu)$  Gamma distribution; link  $g(\mu) = \log(\mu)$ ).

The variables in a hierarchical generalized linear model are

- i) the observed responses  $y_1, y_2, \dots, y_k$  ( $\in \mathcal{M}$ )
- ii) the unobserved random effects  $u_1, u_2, \dots, u_q$  ( $\in \mathcal{M}$ )
- iii) and the corresponding unobserved canonical variables  $v_i = g(u_i)$  ( $\in \Omega$ )

The *linear predictor* is of the form

$$\boldsymbol{\theta} = g(\boldsymbol{\mu}|\boldsymbol{v}) = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\boldsymbol{v} \quad (4.24)$$

The distribution of  $\mathbf{V} \in \Omega$  is a conjugated distribution to the canonical parameter  $\theta$ . The derived distribution of  $\mathbf{U} \in \mathcal{M}$  is the corresponding conjugated distribution to the mean value parameter  $\mu$  such that  $E[\mathbf{U}] = \boldsymbol{\psi}$ . When the conditional distribution of  $Y|\mu$  is a Poisson distribution, and the distribution of  $V$  is constructed in such a way that the distribution of  $U = \log(V)$  is a Gamma distribution with mean value  $E[U] = \psi = 1$ , then it follows that the distribution of  $Y$  is a negative binomial distribution with parameters determined by  $\mathbf{X}\boldsymbol{\beta}$  and  $\mathbf{Z}\boldsymbol{v}$ .

#### 4.4.3 Formulation of the hierarchical Poisson Gamma model

To motivate the choice of distribution for unobserved random variables  $\mathbf{U}$ , an illustrative example is presented.

##### INCLUDE AN EXAMPLE

```
## # A tibble: 2 x 3
##   ageGroup `mean(y)` `var(y)`
##   <fct>     <dbl>    <dbl>
## 1 <65 years      1.21     1.57
## 2 65+ years      3.23     5.58
```

**Theorem 4.2** (Compound Poisson Gamma model). *In the compound Poisson Gamma model the conditional distribution of the count observations are assumed to be a Poisson distribution with intensities  $\lambda$*

$$f_{Y|u}(y; u, \boldsymbol{\beta}) = \frac{(\lambda u)^y}{y!} \exp(-\lambda u) \quad (4.25)$$

*The probability density function for the random effects  $u$  are assumed to follow a reparametrized Gamma distribution with mean 1,  $u \sim G(1/\phi, \phi)$  that is*

$$f_u(u; \phi) = \frac{1}{\phi \Gamma(1/\phi)} \left(\frac{u}{\phi}\right)^{1/\phi-1} \exp(-u/\phi) \quad (4.26)$$

Subsequently, the model can be formulated as a two-level hierarchical model

$$Y|u \sim \text{Pois}(\lambda u) \quad (4.27a)$$

$$u \sim G(1/\phi, \phi) \quad (4.27b)$$

Given (4.25) and (4.26), the probability function for the marginal distribution of  $Y$  is determined from

$$\begin{aligned} g_Y(y; \beta, \phi) &= \int_{u=0}^{\infty} f_{Y|u}(y; u, \beta) f_u(u; \phi) du \\ &= \int_{u=0}^{\infty} \frac{(\lambda u)^y}{y!} \exp(-\lambda u) \frac{1}{\phi \Gamma(1/\phi)} \left(\frac{u}{\phi}\right)^{1/\phi-1} \exp(-u/\phi) du \\ &= \frac{\lambda^y}{y! \Gamma(1/\phi) \phi^{1/\phi}} \int_{u=0}^{\infty} u^{y+1/\phi-1} \exp(-u(\lambda\phi + 1)/\phi) du \end{aligned} \quad (4.28)$$

In (4.28) it is noted that the integrand is the kernel in the probability density function for a Gamma distribution,  $G(y + 1/\phi, \phi/(\lambda\phi + 1))$ . As the integral of the density shall equal one, we find by adjusting the norming constant that

$$\int_{u=0}^{\infty} u^{y+1/\phi-1} \exp\left(-u/(\phi/(\lambda\phi + 1))\right) du = \frac{\phi^{y+1/\phi} \Gamma(y + 1/\phi)}{(\lambda\phi + 1)^{y+1/\phi}} \quad (4.29)$$

Therefore, it can be shown that the marginal distribution of  $Y$  is a negative binomial distribution,  $Y \sim NB(1/\phi, 1/(\lambda\phi + 1))$ . The probability function for  $Y$  is

$$\begin{aligned} P[Y = y] &= g_Y(y; \beta, \phi) \\ &= \frac{\lambda^y}{y! \Gamma(1/\phi) \phi^{1/\phi}} \frac{\phi^{y+1/\phi} \Gamma(y + 1/\phi)}{(\lambda\phi + 1)^{y+1/\phi}} \\ &= \frac{\Gamma(y + 1/\phi)}{\Gamma(1/\phi) y!} \frac{1}{(\lambda\phi + 1)^{1/\phi}} \left(\frac{\lambda\phi}{\lambda\phi + 1}\right)^y \\ &= \binom{y + 1/\phi - 1}{y} \frac{1}{(\lambda\phi + 1)^{1/\phi}} \left(\frac{\lambda\phi}{\lambda\phi + 1}\right)^y, \quad \text{for } y = 0, 1, 2, \dots \end{aligned} \quad (4.30)$$

where we have used the convention

$$\binom{z}{y} = \frac{\Gamma(z + 1)}{\Gamma(z + 1 - y) y!} \quad (4.31)$$

for  $z$  real and  $y$  integer values. Consequently, the mean and variance of  $Y$  are given by

$$\mathbb{E}[Y] = \lambda \quad \mathbb{V}[Y] = \lambda(\lambda\phi + 1) \quad (4.32)$$

The likelihood function for estimating  $(\beta, \phi)$  is

$$L(\beta, \phi; y_{it}) = \prod_{t=1}^T \prod_{i=1}^m \binom{y_{it} + 1/\phi - 1}{y_{it}} \frac{1}{(\lambda_{it}\phi + 1)^{1/\phi}} \left(\frac{\lambda_{it}\phi}{\lambda_{it}\phi + 1}\right)^{y_{it}} \quad (4.33)$$

### Inference on individual groups

In order to simplify the notation, the subscript indicating the group and time are omitted. Consider the compound Poisson Gamma model in (4.27), and assume that a value  $Y = y$  has been observed.

Then the conditional distribution of  $u$  for given  $Y = y$  is found using Bayes Theorem

$$\begin{aligned}
 g_u(u|Y = y) &= \frac{f_{y,u}(y, u)}{g_Y(y; \lambda, \phi)} \\
 &= \frac{f_{y|u}(y; u)g_u(u)}{g_Y(y; \lambda, \phi)} \\
 &= \frac{1}{g_Y(y; \lambda, \phi)} \left( \frac{(\lambda u)^y}{y!} \exp(-\lambda u) \frac{1}{\phi \Gamma(1/\phi)} \left( \frac{u}{\phi} \right)^{1/\phi-1} \exp(-u/\phi) \right) \\
 &\propto u^{y+1/\phi-1} \exp(-u(\lambda\phi + 1)/\phi)
 \end{aligned} \tag{4.34}$$

We identify the *kernel* of the probability density function

$$u^{y+1/\phi-1} \exp(-u(\lambda\phi + 1)/\phi) \tag{4.35}$$

as the kernel of a Gamma distribution,  $G(y + 1/\phi, \phi/(\lambda\phi + 1))$ , i.e. the conditional distribution of  $u$  for given  $Y = y$  can be written as

$$u|Y = y \sim G(y + 1/\phi, \phi/(\lambda\phi + 1)) \tag{4.36}$$

The mean of the conditional distribution is given by:

$$\mathbb{E}[u|Y = y] = \frac{y\phi + 1}{\lambda\phi + 1} \tag{4.37}$$

And the variance of the conditional distribution is:

$$V[u|Y = y] = \frac{(\phi^2 + \phi)}{(\lambda\phi + 1)^2} \tag{4.38}$$

These formulas provide the mean and variance of the conditional distribution of  $u$  given the observed value  $Y = y$ .

### Why do we choose the Gamma distribution to represent the variation between days?

The choice of the Gamma distribution for modeling the random effects has been motivated by several reasons. Firstly, the support of the Gamma distribution, which ranges from 0 to infinity, aligns with the mean-value space, denoted as  $\mathcal{M}$ , for the Poisson distribution. This ensures that the random effects are constrained within a meaningful range for the underlying Poisson process.

Secondly, the two-parameter family of Gamma distributions offers considerable flexibility, encompassing a wide range of shapes and distributions that can span from exponential-like distributions to fairly symmetrical distributions on the positive real line. This flexibility allows the model to capture various patterns and characteristics observed in the data.

Additionally, the choice of the Gamma distribution has benefits in terms of the derivation of the marginal distribution of the response variable  $Y$ . The kernel  $u^{\alpha-1} \exp(-u/\beta)$  of the Gamma distribution used for modeling the random effects exhibits a similar structure to the kernel  $u^y \exp(-u)$  of the likelihood function corresponding to the sampling distribution of  $Y$ . This similarity facilitates the analytical computation of the integral involved in deriving the marginal distribution, as it can be expressed in terms of known functions.

Overall, the Gamma distribution is selected due to its alignment with the mean-value space of the Poisson distribution, its flexibility in capturing diverse distributions, and its analytical convenience in computing the marginal distribution of the response variable.

## 4.5 Parameter estimation

In this section, the parameter estimation and implementation of the novel outbreak detection algorithm in R using the **TMB** (Template model Builder) package is presented. **TMB** is an open-source R package developed by Kristensen et al. (2016). This package facilitates efficient maximum likelihood estimation and uncertainty calculations for the parameter set  $\theta = (\beta, \Psi)$  and random effects  $u$ . The presentation of the parameter estimation conducted in **TMB** is strongly inspired by Chapter 2 in Kristensen et al. (2016) and Section 5.10 in Madsen and Thyregod (2011).

The **TMB** package maximizes a user-provided objective function in the form of a C++ template to estimate the maximum likelihood for the parameter set  $\theta = (\beta, \Psi)$ . See Appendix B to access the C++ template files used in this master's thesis. The objective function maximizes the marginal log-likelihood function, which integrates out the random effects  $u$

$$\ell_M(\theta; y) = \int_{\mathbb{R}^q} \ell(\theta; u, y) du \quad (4.39)$$

where  $\ell(\theta, u, y)$  is the joint log-likelihood function of the data given the parameters and random effects. We use  $\hat{u}_\theta$  to denote the maximizer of the joint log-likelihood  $\ell(\theta; u, y)$  w.r.t.  $u$ ; i.e.,

$$\hat{u}_\theta = \arg \max_u \ell(\theta; u, y) \quad (4.40)$$

Using  $H(\hat{u}_\theta)$  to denote the negative Hessian of the joint log-likelihood evaluated at  $\hat{u}_\theta$ ; i.e.,

$$H(\hat{u}_\theta) = -\ell''_{uu}(\theta, u, y)|_{u=\hat{u}_\theta} \quad (4.41)$$

The Laplace approximation for the marginal log-likelihood  $\ell_M(\theta)$  is

$$\ell_{M,LA}(\theta, y) = \ell(\theta, u, y) - \frac{1}{2} \log \left| \frac{H(\hat{u}_\theta)}{2\pi} \right| \quad (4.42)$$

Our estimate of  $\theta$  minimizes the negative log of the Laplace approximation, i.e.,

$$-\ell_{M,LA}(\theta, y) = -\ell(\theta, u, y) + \frac{1}{2} \log \left| \frac{H(\hat{u}_\theta)}{2\pi} \right| \quad (4.43)$$

The maximization of the Laplace approximation for the marginal likelihood is then performed using conventional R optimization routines (e.g., BFGS) to optimize the objective and obtain our estimate  $\hat{\theta}$ . Uncertainty of the estimate  $\hat{\theta}$ , or any differentiable function of the estimate  $\phi(\hat{\theta})$ , is obtained by the  $\delta$ -method:

$$V(\phi(\hat{\theta})) = -\phi'_{\theta}(\hat{\theta}) \left( \Delta^2 \ell_{M,LA}(\hat{\theta}, \mathbf{y}) \right)^{-1} \phi'_{\theta}(\hat{\theta})^T \quad (4.44)$$

WRITE SOMETHING ABOUT THE LIKELIHOOD ESTIMATION IN THE HIERARCHICAL POISSON GAMMA MODEL!

Additionally, **TMB** utilized Automatic Differentiation (AD) techniques (Griewank and Walther 2008) to evaluate first, second, and potentially third-order derivatives. For a comprehensive introduction to the concept of AD, it is recommended to read Section 2.1 and Section 2.2 of Fournier et al. (2012). This approach enhances the computational efficiency and accuracy of the parameter estimation process in the implemented models.

## 4.6 Scoring rule

In this section, the scoring rule used to evaluate the overall score of the models is outlined. The approach is inspired by Bjerregård, Møller, and Madsen (2021). For a given time series  $y_t = y_1, \dots, y_N$ , each forecast and its corresponding realized observation pair  $(G_t, y_t)$  is evaluated. The overall score of the model is then reported as the average score:

$$\bar{S}(G, y) = \frac{1}{N} \sum_{t=1}^N S(G_t, y_t) \quad (4.45)$$

One commonly used scoring rule is the *logarithmic score* derived from likelihood theory, which is defined as  $S(G, y) = -\log(f(y))$  (Good 1992). This scoring rule is based on the probability density function and is equivalent to the log-likelihood of the forecast model. It has desirable properties as it captures all possible information about the observed data in relation to the model. However, it has a potential drawback in that it heavily penalizes unlikely observations. Consequently, even small changes in the tails of a density forecast can lead to significant changes in the *logarithmic score*, even when the overall shape of the density remains unchanged.

The calculation of the logarithmic score is shown in Example 4.1, which is adapted from Bjerregård, Møller, and Madsen (2021).

**Example 4.1** (Calculation of the logarithmic score). The Gamma distribution is used to represent the probabilistic forecast in this example. The Gamma distribution is parametrized by two parameters, shape ( $\alpha$ ) and rate ( $\beta$ ), and its probability density function (PDF) is given by:

$$f(y) = \frac{1}{\Gamma(\alpha)\beta} \left( \frac{y}{\beta} \right)^{\alpha-1} \exp\left(-\frac{y}{\beta}\right) \quad (4.46)$$

In this example, the parameters of the true model, denoted as  $f$ , is chosen to be  $(\alpha, \beta) = (3, 3)$ . We simulate 10 observations, denoted as  $y_1, y_2, \dots, y_{10}$ , which are shown in Table 4.1. The true model  $f$  is compared to a competing model, denoted as  $g$ , which is a Gamma distribution with parameters  $(\alpha, \beta) = (3, 8)$ . The true model  $f$ , the competing model  $g$ , and the observations  $y$  are illustrated in Figure 4.1.

Table 4.1: 10 simulated observations following a G(3,3)-distribution.

$i$	1	2	3	4	5	6	7	8	9	10
$y_i$	1.278	2.233	0.657	0.831	1.281	0.287	1.363	1.612	0.790	0.161

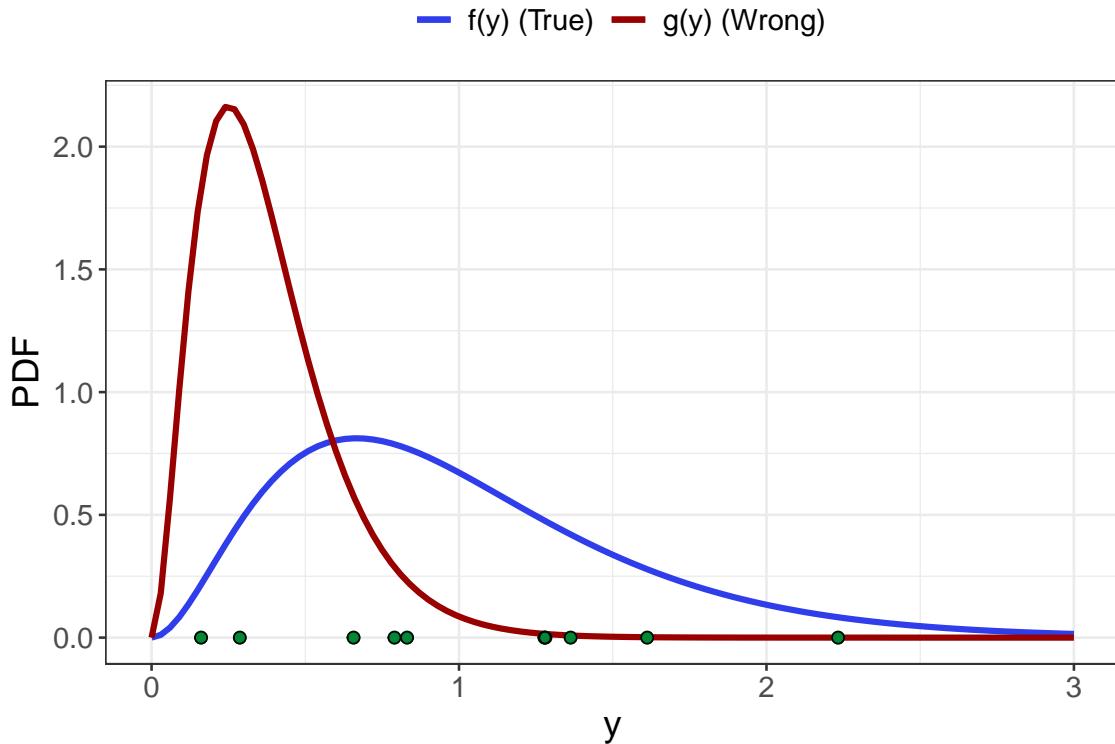


Figure 4.1: Observations (green dots) were simulated from a G(3,3)-distribution. The true model  $f$  (blue) is shown, along with the PDF for the competing model  $g$ , which follows a  $G(3,8)$ -distribution.

The logarithmic score of the true model  $f$  for the first observation,  $y_1 = 1.278$ , is calculated as follows

$$\begin{aligned}
 -\log(f(y_1)) &= -\log(f(1.278)) \\
 &= -\log\left[\frac{1}{\Gamma(3)4}\left(\frac{1.278}{4}\right)^{3-1} \exp(-1.278/4)\right] \\
 &= -1.156
 \end{aligned} \tag{4.47}$$

Similarly, the logarithmic scores for the other observations can be calculated using the same formula. The individual logarithmic scores for all 10 observations are presented in Table 4.2. Among the 10 observations, 8 of them are more likely to occur under the true model  $f$  compared to the competing model  $g$ . The final key quantity, the average logarithmic score,  $\bar{S}(G, y)$ , can be calculated using Equation (4.45).

$$\begin{aligned}
 \bar{S}(f, y) &= 1.1 \\
 \bar{S}(g, y) &= 3.22
 \end{aligned} \tag{4.48}$$

Table 4.2: Logarithmic scores of the two different Gamma models w.r.t. the 10 individual observations.

$i$	1	2	3	4	5	6	7	8	9	10
$\bar{S}(f, y_i)$	1.16	3.86	0.00	0.23	1.16	0.18	1.37	2.03	0.17	0.83
$\bar{S}(g, y_i)$	4.19	10.71	0.55	1.47	4.21	-0.75	4.74	6.40	1.25	-0.60



# 5 Case studies

This chapter presents the findings obtained by applying both the state-of-the-art outbreak detection algorithm and the novel outbreak detection algorithm to the subset of diseases examined in this master's thesis. The performance of these algorithms in identifying outbreaks within the selected diseases will be extensively discussed and analyzed. Additionally, a comprehensive comparative analysis will be conducted to offer valuable insights into the strengths and limitations of both the state-of-the-art algorithms and, more importantly, the novel algorithms. For a detailed presentation of the data used to generate these results, please refer to Chapter 3.

To demonstrate the practical application of these algorithms, a comprehensive case study centered on STEC will be presented as the primary focus. Following that, the results of the remaining case studies will be presented in a more concise manner to keep the reader engaged. However, for the complete collection of related figures and tables, please refer to Appendix D.

## 5.1 Shiga toxin (verotoxin)-producing *Escherichia coli*

Initially, the ... is investigated.

ALL RESULTS REGARDING STEC IS GATHERED HERE

### 5.1.1 Applying the state-of-the-art outbreak detection algorithm to Shiga toxin (verotoxin)-producing *Escherichia coli*

Initially, outbreak detection using the Farrington method and the Noufaily method, described in Section 4.1, are investigated. The methods are implemented in the `farringtonFlexible` function, which can be accessed through the R package known as `surveillance`. The control arguments for each of the methods can be found in Appendix C.1. The reference values for the analysis are based on data from January 2008 to February 2011. Subsequently, surveillance is conducted using the data from March 2011 to December 2021. The resulting series is visualized in Figure 5.1.

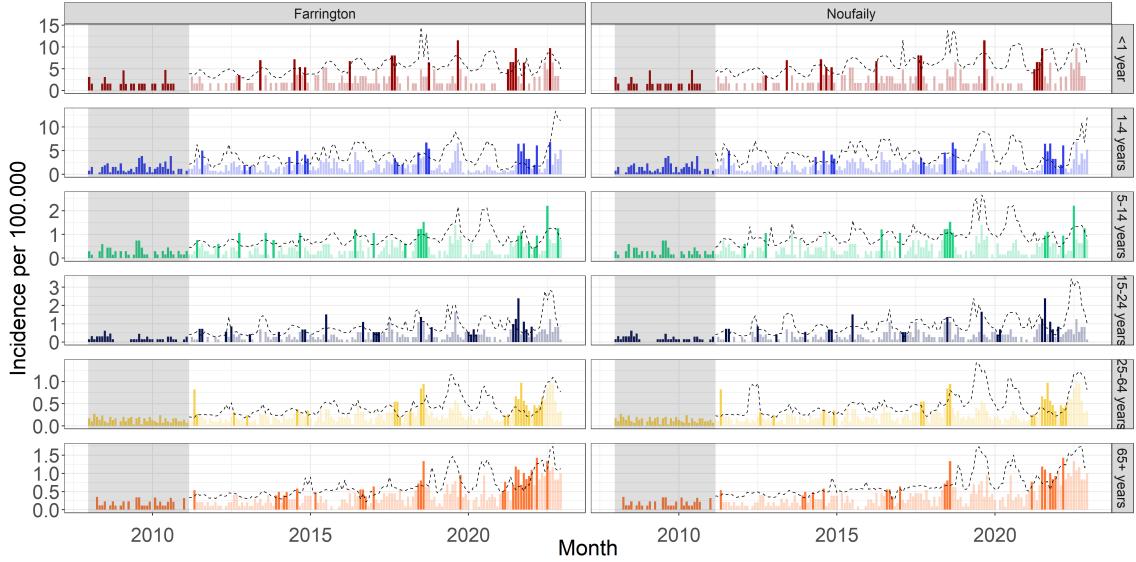


Figure 5.1: Monthly Shiga toxin (verotoxin)-producing *Escherichia coli* incidence per 100,000 in Denmark, 2008-2022. Monitored by (left) Farrington and (right) Noufaily method. Reference data for the estimation of model parameters from January 2008 to March 2011 (grey area). Threshold (dashed line) is computed for observations timepoints outside reference data. Alarm triggered (full opacity) if observations exceeds threshold.

In Figure 5.1, multiple alarms are observed for both the Farrington method and the Noufaily method. The Farrington method triggers a total of 133 alarms, while the Noufaily method produces slightly fewer, specifically 102 alarms.

### 5.1.2 Applying the novel outbreak detection algorithm to Shiga toxin (verotoxin)-producing *Escherichia coli*

At first, a suite of different models for the fixed effects is proposed

$$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \log(n_{it}) \quad (5.1)$$

$$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \beta_{\text{trend}} t + \log(n_{it}) \quad (5.2)$$

$$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \sin\left(\frac{\pi \cdot \tau_t}{6}\right)\beta_{\text{sin}} + \cos\left(\frac{\pi \cdot \tau_t}{6}\right)\beta_{\text{cos}} + \log(n_{it}) \quad (5.3)$$

$$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \beta_{\text{trend}} t + \sin\left(\frac{\pi \cdot \tau_t}{6}\right)\beta_{\text{sin}} + \cos\left(\frac{\pi \cdot \tau_t}{6}\right)\beta_{\text{cos}} + \log(n_{it}) \quad (5.4)$$

The average logarithmic score,  $\bar{S}(G, y)$ , and parameter estimate for the models assuming either the hierarchical Poisson normal model or the hierarchical Poisson Gamma model is gathered in Table 5.1.

**Table 5.1:** The average logarithmic score,  $\bar{S}(G, y)$ , along with the parameter estimates at  $t_0$  for a suite of models modelling Shiga toxin (verotoxin)-producing *Escherichia coli*, assuming either the hierarchical Poisson Normal model or the hierarchical Poisson Gamma model. The confidence intervals for the estimates are calculated using profile likelihood confidence intervals.

$\bar{S}(G, y)$	Parameter	Estimate (95% CI)
<b>Poisson Normal</b>		
$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \log(n_{it})$		
14.07	$\beta_{<1\text{year}}$	11.23 (10.88, 11.55)
	$\beta_{1-4\text{years}}$	13.45 (13.18, 13.72)
	$\beta_{5-14\text{years}}$	14.3 (14.04, 14.55)
	$\beta_{15-24\text{years}}$	14.71 (14.46, 14.94)
	$\beta_{25-64\text{years}}$	16.95 (16.73, 17.16)
	$\beta_{65+\text{years}}$	15.87 (15.66, 16.08)
	$\log(\sigma)$	-0.79 (-1.04, -0.57)
$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \beta_{\text{trend}} t + \log(n_{it})$		
13.76	$\beta_{\text{trend}}$	0.05 (0.04, 0.06)
	$\beta_{<1\text{year}}$	12.12 (11.79, 12.43)
	$\beta_{1-4\text{years}}$	14.61 (14.36, 14.85)
	$\beta_{5-14\text{years}}$	15.12 (14.85, 15.39)
	$\beta_{15-24\text{years}}$	15.54 (15.28, 15.79)
	$\beta_{25-64\text{years}}$	17.95 (17.73, 18.17)
	$\beta_{65+\text{years}}$	16.71 (16.48, 16.94)
	$\log(\sigma)$	-0.88 (-1.11, -0.68)
$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \beta_{\sin} \sin\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \beta_{\cos} \cos\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \log(n_{it})$		
13.19	$\beta_{<1\text{year}}$	11.29 (10.95, 11.6)
	$\beta_{1-4\text{years}}$	13.77 (13.52, 14)
	$\beta_{5-14\text{years}}$	14.3 (14.03, 14.56)
	$\beta_{15-24\text{years}}$	14.71 (14.46, 14.95)
	$\beta_{25-64\text{years}}$	17.14 (16.93, 17.34)
	$\beta_{65+\text{years}}$	15.91 (15.7, 16.12)
	$\beta_{\sin}$	-0.47 (-0.61, -0.33)
	$\beta_{\cos}$	-0.24 (-0.38, -0.1)
	$\log(\sigma)$	-0.65 (-0.84, -0.47)
$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \beta_{\text{trend}} t + \beta_{\sin} \sin\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \beta_{\cos} \cos\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \log(n_{it})$		
13.12	$\beta_{\text{trend}}$	0.04 (0.03, 0.05)
	$\beta_{<1\text{year}}$	11.95 (11.68, 12.31)
	$\beta_{1-4\text{years}}$	14.51 (14.28, 14.73)
	$\beta_{5-14\text{years}}$	15 (14.75, 15.26)
	$\beta_{15-24\text{years}}$	15.42 (15.19, 15.66)
	$\beta_{25-64\text{years}}$	17.82 (17.64, 18.03)
	$\beta_{65+\text{years}}$	16.59 (16.4, 16.81)
	$\beta_{\sin}$	-0.33 (-0.43, -0.21)
	$\beta_{\cos}$	-0.22 (-0.33, -0.1)
	$\log(\sigma)$	-1.1 (-1.45, -0.87)

Table 5.1: (continued)

Model	Formula	$\bar{S}(G, y)$	Parameter	Estimate (95% CI)
<b>Poisson Gamma</b>				
	$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \log(n_{it})$			
14.16			$\beta_{<1year}$	11.51 (11.2, 11.8)
			$\beta_{1-4years}$	13.56 (13.29, 13.82)
			$\beta_{5-14years}$	14.4 (14.15, 14.65)
			$\beta_{15-24years}$	14.8 (14.57, 15.03)
			$\beta_{25-64years}$	16.97 (16.76, 17.2)
			$\beta_{65+years}$	15.95 (15.74, 16.16)
			$\log(\phi)$	-1.52 (-2.02, -1.09)
	$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \beta_{trend}t + \log(n_{it})$			
13.82			$\beta_{trend}$	0.05 (0.04, 0.05)
			$\beta_{<1year}$	12.2 (11.88, 12.52)
			$\beta_{1-4years}$	14.71 (14.47, 14.95)
			$\beta_{5-14years}$	15.2 (14.94, 15.47)
			$\beta_{15-24years}$	15.64 (15.38, 15.89)
			$\beta_{25-64years}$	18.02 (17.81, 18.24)
			$\beta_{65+years}$	16.77 (16.54, 17)
			$\log(\phi)$	-1.77 (-2.24, -1.35)
	$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \beta_{\sin} \sin\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \beta_{\cos} \cos\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \log(n_{it})$			
13.24			$\beta_{<1year}$	11.42 (11.1, 11.73)
			$\beta_{1-4years}$	13.96 (13.73, 14.19)
			$\beta_{5-14years}$	14.43 (14.17, 14.68)
			$\beta_{15-24years}$	14.83 (14.59, 15.07)
			$\beta_{25-64years}$	17.26 (17.07, 17.47)
			$\beta_{65+years}$	16.05 (15.84, 16.26)
			$\beta_{\sin}$	-0.46 (-0.59, -0.32)
			$\beta_{\cos}$	-0.24 (-0.38, -0.1)
			$\log(\phi)$	-1.31 (-1.68, -0.98)
	$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \beta_{trend}t + \beta_{\sin} \sin\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \beta_{\cos} \cos\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \log(n_{it})$			
13.13			$\beta_{trend}$	0.04 (0.03, 0.05)
			$\beta_{<1year}$	12.06 (11.74, 12.35)
			$\beta_{1-4years}$	14.58 (14.36, 14.8)
			$\beta_{5-14years}$	15.06 (14.81, 15.31)
			$\beta_{15-24years}$	15.48 (15.25, 15.71)
			$\beta_{25-64years}$	17.88 (17.69, 18.07)
			$\beta_{65+years}$	16.65 (16.45, 16.85)
			$\beta_{\sin}$	-0.32 (-0.43, -0.21)
			$\beta_{\cos}$	-0.21 (-0.32, -0.1)
			$\log(\phi)$	-2.29 (-2.95, -1.77)

As expected, the models with including both trend and seasonality minimizes the logarithmic score assuming the hierarchical Poisson Normal model and the hierarchical Poisson Gamma model, respectively. Thus, these models are chosen for further consideration.

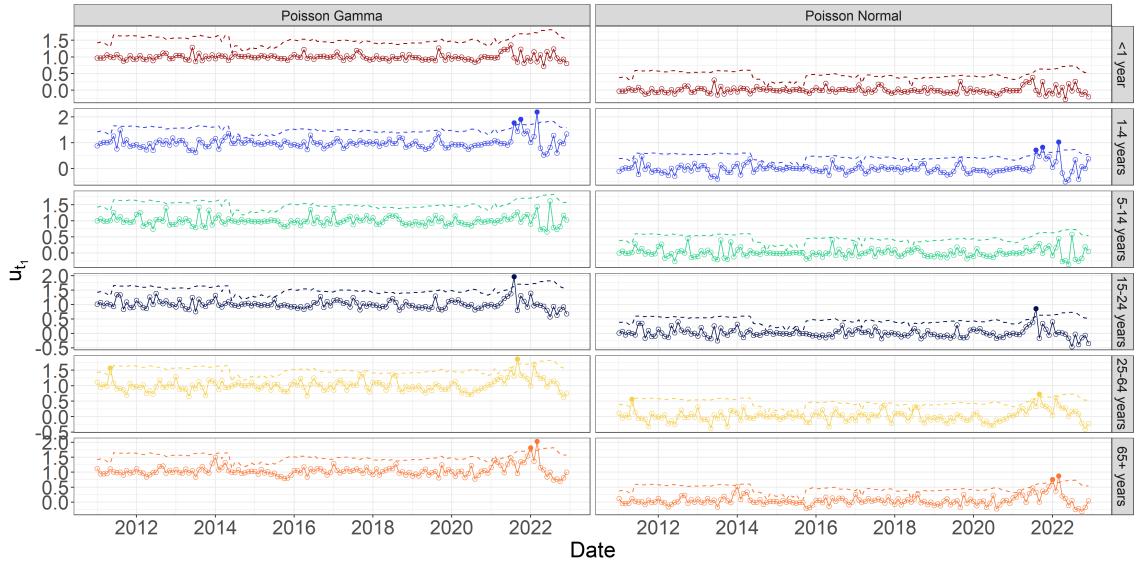


Figure 5.2: One-step ahead random effects  $u_{t1}$  of Shiga toxin (verotoxin)-producing *Escherichia coli* in Denmark, 2011-2022, as circles. Poisson Normal model (top) and Poisson Gamma model (bottom) monitor the disease. Alarm raised (solid circle) if  $u_{t1}$  exceeds the threshold (dashed line).

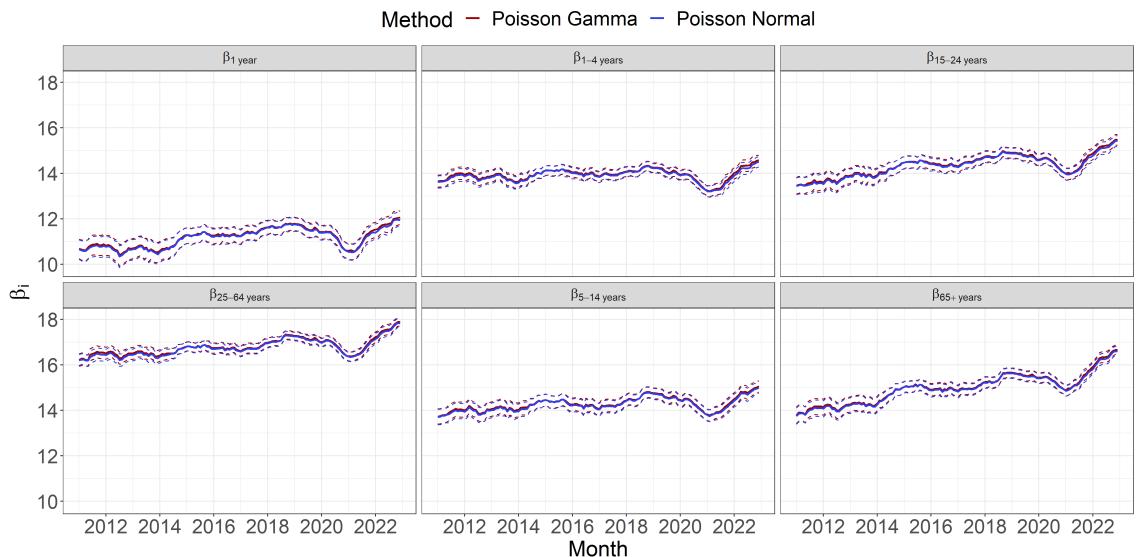


Figure 5.3: Placeholder caption

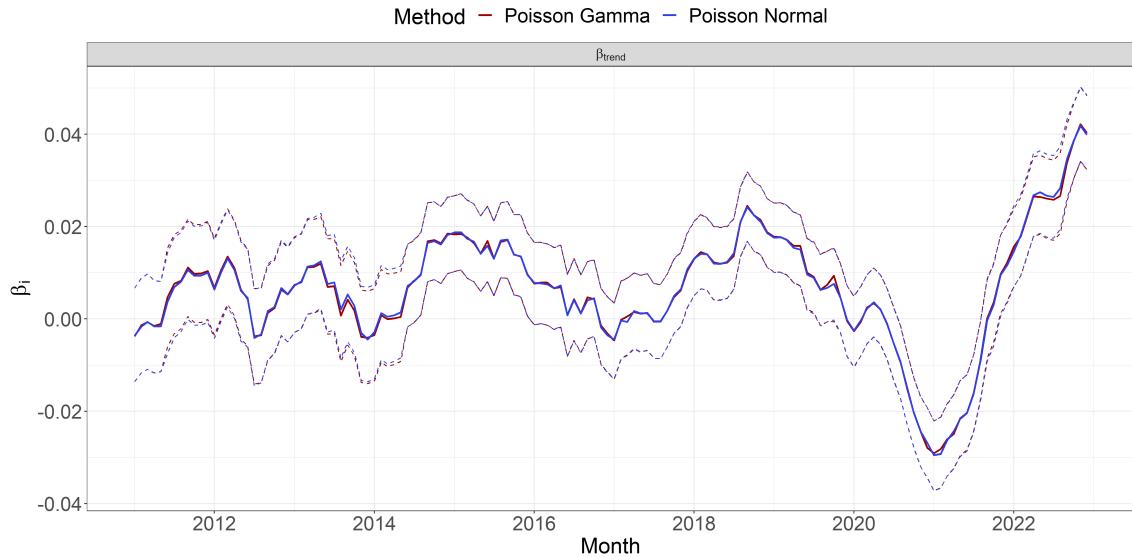


Figure 5.4: Placeholder caption

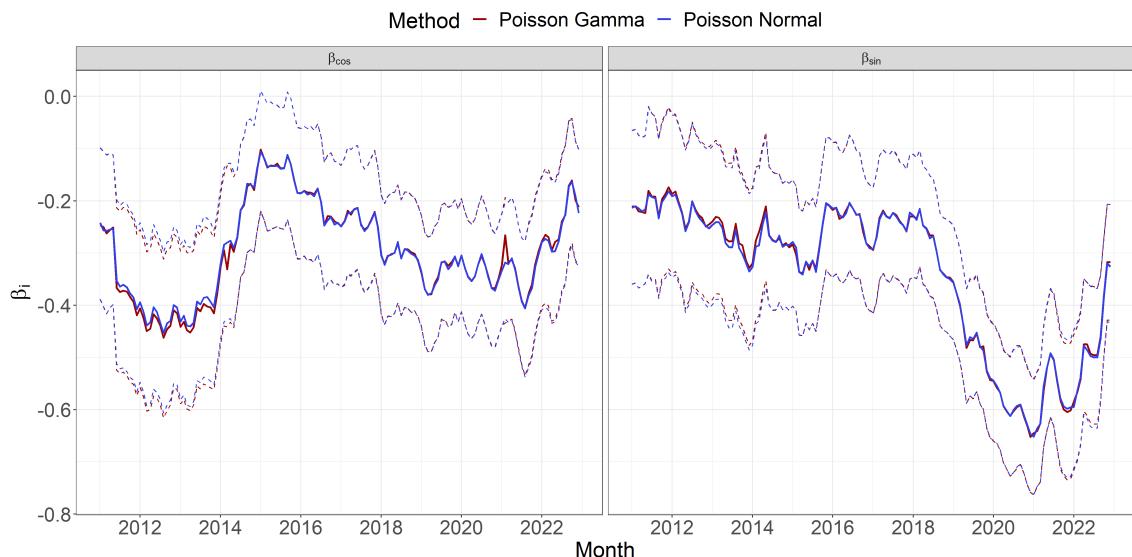


Figure 5.5: Placeholder caption

Apart from the Poisson Gamma model being seemingly more robust, the estimated fixed effects parameters for the models are on a comparable scale. This can be seen in Figure 5.3

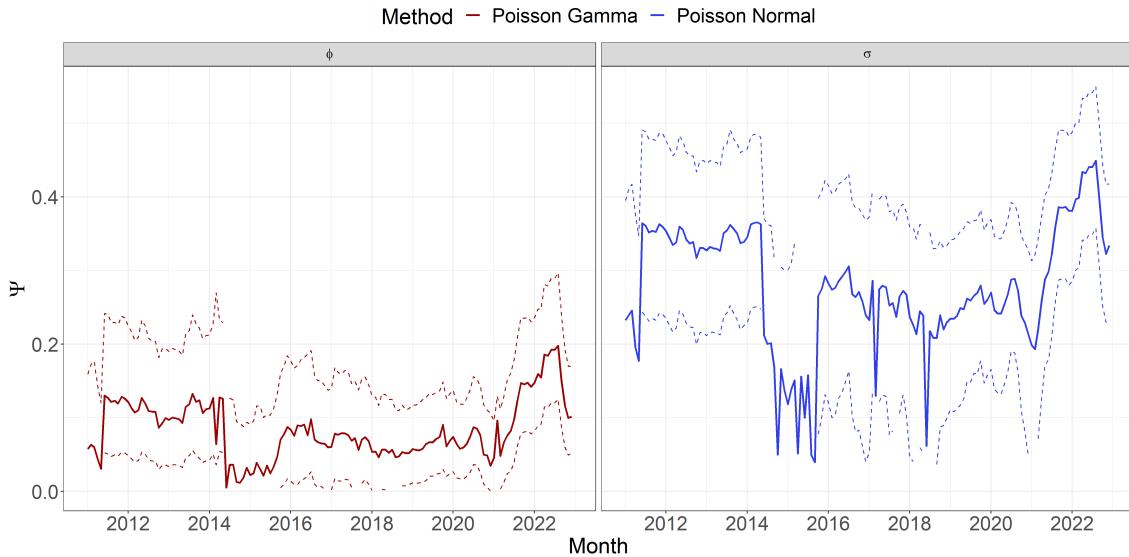


Figure 5.6: Placeholder caption

## 5.2 *Listeriosis*

ALL THE RESULTS REGARDING *LISTERIOSIS* IS GATHERED HERE!

### 5.2.1 Applying the state-of-the-art outbreak detection algorithm to *Listeriosis*

Initially, outbreak detection using the Farrington method and the Noufaily method, described in Section 4.1, are investigated. The methods are implemented in the `farringtonFlexible` function, which can be accessed through a R package known as **surveillance**. The control arguments for each of the functions can be found in Appendix C.1. The reference values for the analysis are based on the data from January 2008 to February 2011. Subsequently, surveillance is conducted using the data from March 2011 to December 2021. The resulting series is visualized in Figure 5.7.

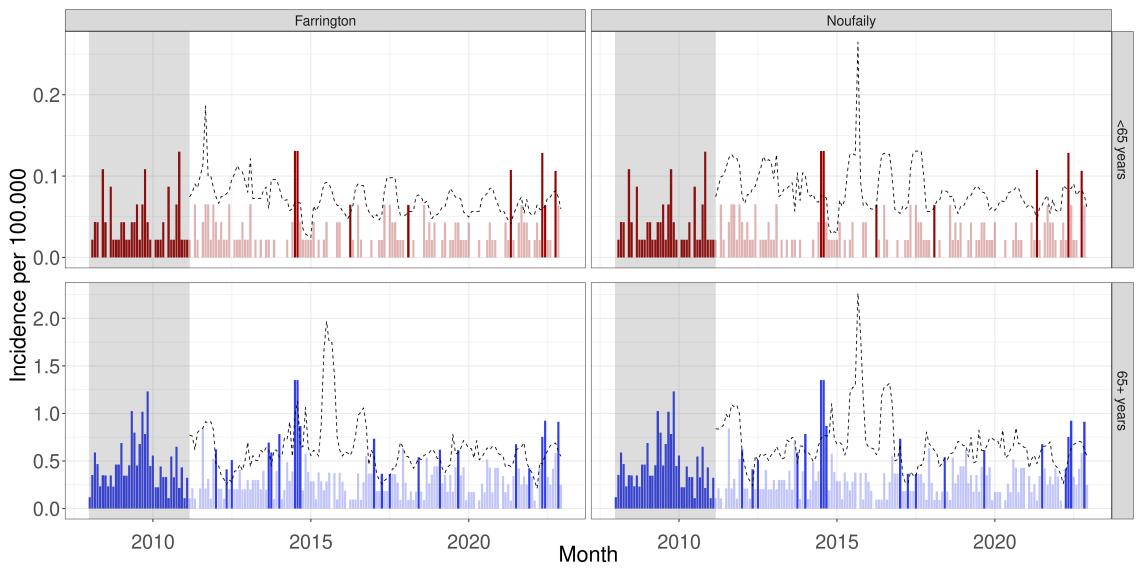


Figure 5.7: Monthly *Listeriosis* incidence per 100,000 in Denmark, 2008-2022. Monitored by (left) Farrington and (right) Noufaily method. Reference data for the estimation of model parameters from January 2008 to March 2011 (grey area). Threshold (dashed line) is computed for observations timepoints outside reference data. Alarm triggered (full opacity) if observations exceeds threshold.

In Figure 5.7, multiple alarms can be observed for both the Farrington method and the Noufaily method. The Farrington method triggers a total of 28 alarms, while the Noufaily method produces slightly fewer, specifically 25 alarms.

Interestingly, both the Farrington method and the Noufaily method demonstrate the ability to correctly raise alarms during ongoing outbreak investigations by SSI. One notable example is the outbreak investigated by SSI, which involved 41 cases and resulted in 17 deaths. The Farrington method successfully flags this event in September 2013, with alarms occurring sporadically in the subsequent period. The Noufaily method also identifies this outbreak, although it does so one month later in October 2013.

It is worth noting that the other observations flagged by the methods may be related to some of the numerous other long-spanned outbreaks that occurred concurrently in the period from 2016 to 2021. For a full list over the observations flagged by the state-of-the-art outbreak detection algorithms see Table C.1.

### 5.2.2 Applying the novel outbreak detection algorithm to *Listeriosis*

At this stage, the novel outbreak detection algorithm presented in Section 4.2 is evaluated on its ability to detect outbreaks of *Listeriosis*.

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avgLogS	$\theta$	Estimate (CI)
<b>Poisson Normal</b>		
ageGroup		
3.597	$\beta_{<65years}$	15.43 (15.09, 15.73)
	$\beta_{65+years}$	15.19 (14.98, 15.36)
	$\log(\sigma)$	-3.09 (NA, -0.85)
ageGrouptrend		
3.639	$\beta_{trend}$	0 (NA, NA)
	$\beta_{<65years}$	15.46 (NA, NA)
	$\beta_{65+years}$	15.22 (NA, NA)
	$\log(\sigma)$	-2.87 (NA, NA)
ageGroupseasonality		
3.577	$\beta_{<65years}$	15.4 (NA, NA)
	$\beta_{65+years}$	15.16 (NA, NA)
	$\beta_{\sin}$	-0.34 (NA, NA)
	$\beta_{\cos}$	0.02 (NA, NA)
	$\log(\sigma)$	-6 (NA, NA)
ageGrouptrendseasonality		
3.663	$\beta_{trend}$	0 (NA, NA)
	$\beta_{<65years}$	15.46 (NA, NA)
	$\beta_{65+years}$	15.22 (NA, NA)
	$\beta_{\sin}$	0 (NA, NA)
	$\beta_{\cos}$	0.02 (NA, NA)
	$\log(\sigma)$	-2.83 (NA, NA)
<b>Poisson Gamma</b>		
ageGroup		
3.609	$\beta_{<65years}$	15.43 (15.1, 15.73)
	$\beta_{65+years}$	15.19 (15, 15.37)
	$\log(\phi)$	-6 (-36.77, -1.69)

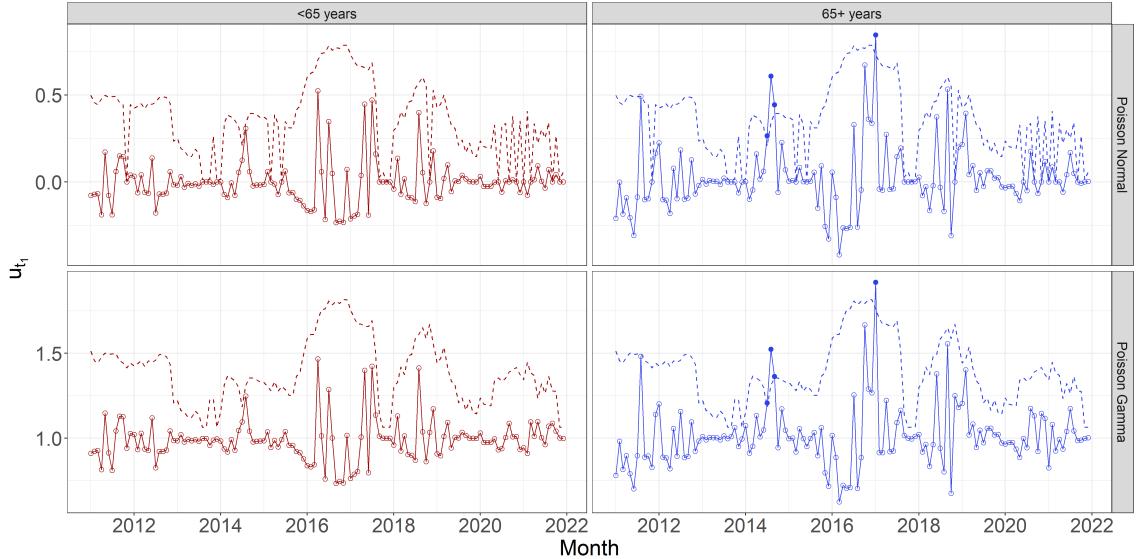


Figure 5.8: One-step ahead random effects  $u_{t_1}$  of *Listeriosis* in Denmark, 2011-2022, as circles. Poisson Normal model (top) and Poisson Gamma model (bottom) monitor the disease. Alarm raised (solid circle) if  $u_{t_1}$  exceeds the threshold (dashed line).

The one-step ahead random effects  $u_{t_1}$  for the disease are visualized in Figure 5.8. It is clear that the models struggle to adequately describe the underlying process of the disease. Only a single alarm is raised in July 2014, August 2014, September 2014, January 2017, which is in the midst of several concurrent outbreaks investigated by SSI, with several observations being borderline significant. When compared to the numerous outbreaks investigated by SSI in the same time period, this is not an impressive result. However, it is important to note that the current state-of-the-art detection method for *Listeriosis* is laboratory-based and requires the use of WGS to link cases that occurred over an extended time period. Moreover, the start of outbreaks with *Listeriosis* are usually dated back to the first identified case, which can be up to several months prior to the peak of the disease.

Mostly, it is noted that the Poisson Gamma model seems more robust, than the Poisson Normal model. Specifically, this is observed in periods with low case counts, i.e. in November 2011, from August 2013 to January 2014, in February 2019, and again in the period from June 2020 to November 2021.

Apart from the Poisson Gamma model being seemingly more robust, the estimated fixed effects parameters for the models are on a comparable scale. This can be seen in Figure 5.9

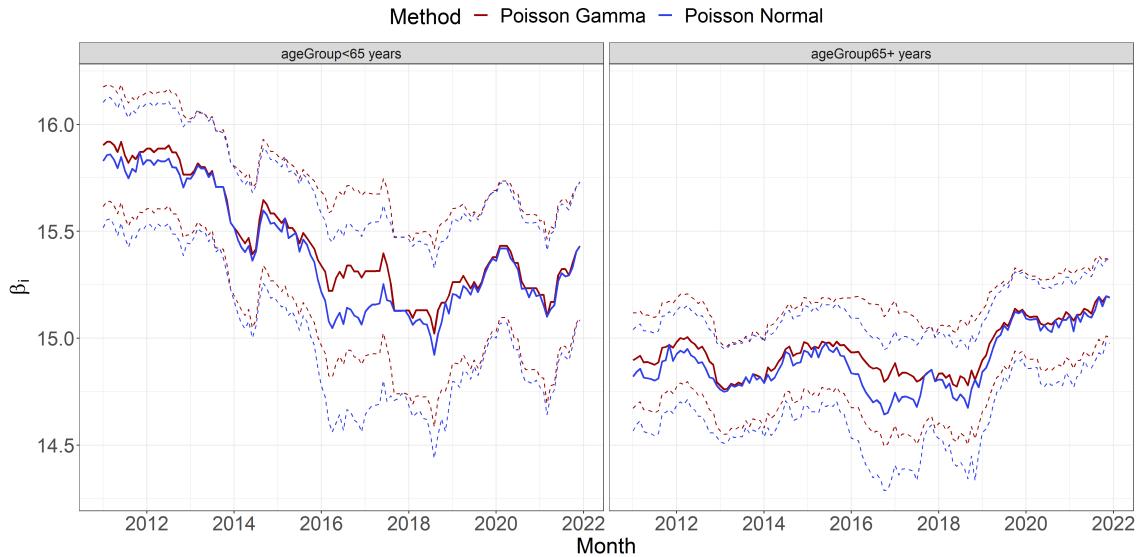


Figure 5.9: Placeholder caption

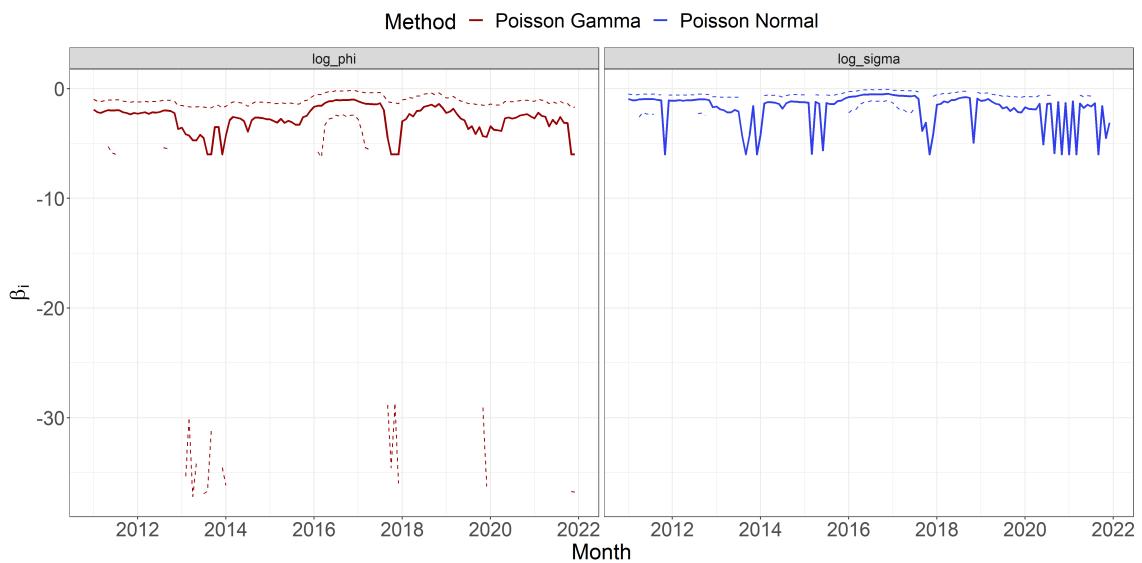


Figure 5.10: Placeholder caption

Consider the models given in (4.19) and (4.27)

### 5.3 *Shigello*sis

ALL THE RESULTS REGARDING *LISTERIOSIS* IS GATHERED HERE!

### 5.3.1 Applying the state-of-the-art outbreak detection algorithm to *Shigellosis*

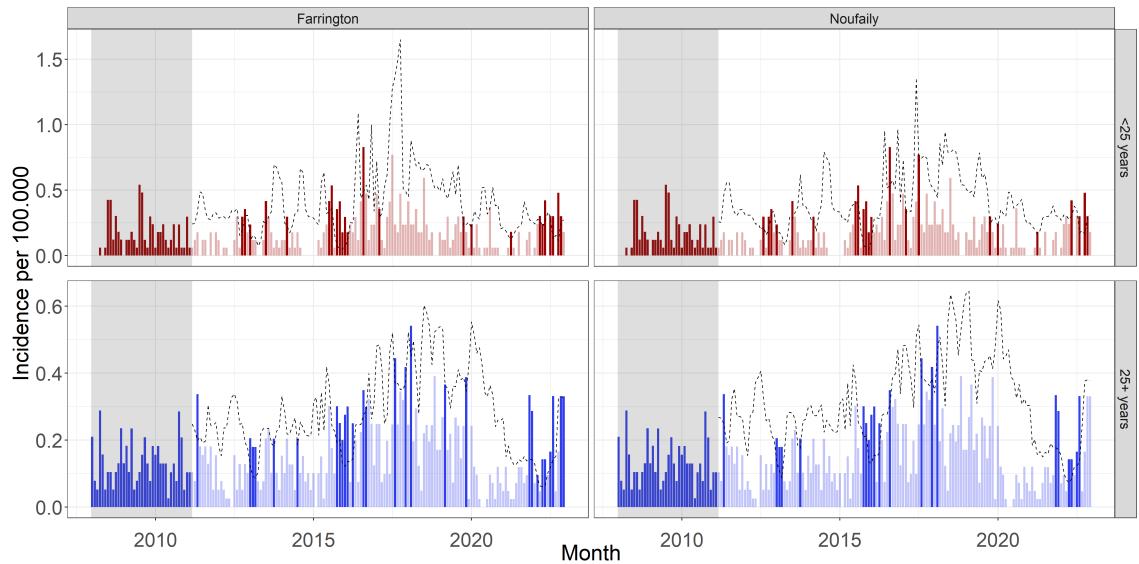


Figure 5.11: Monthly *Shigellosis* incidence per 100,000 in Denmark, 2008-2022. Monitored by (left) Farrington and (right) Noufaily method. Reference data for the estimation of model parameters from January 2008 to March 2011 (grey area). Threshold (dashed line) is computed for observations timepoints outside reference data. Alarm triggered (full opacity) if observations exceeds threshold.

### 5.3.2 Applying the novel outbreak detection algorithm to *Shigellosis*

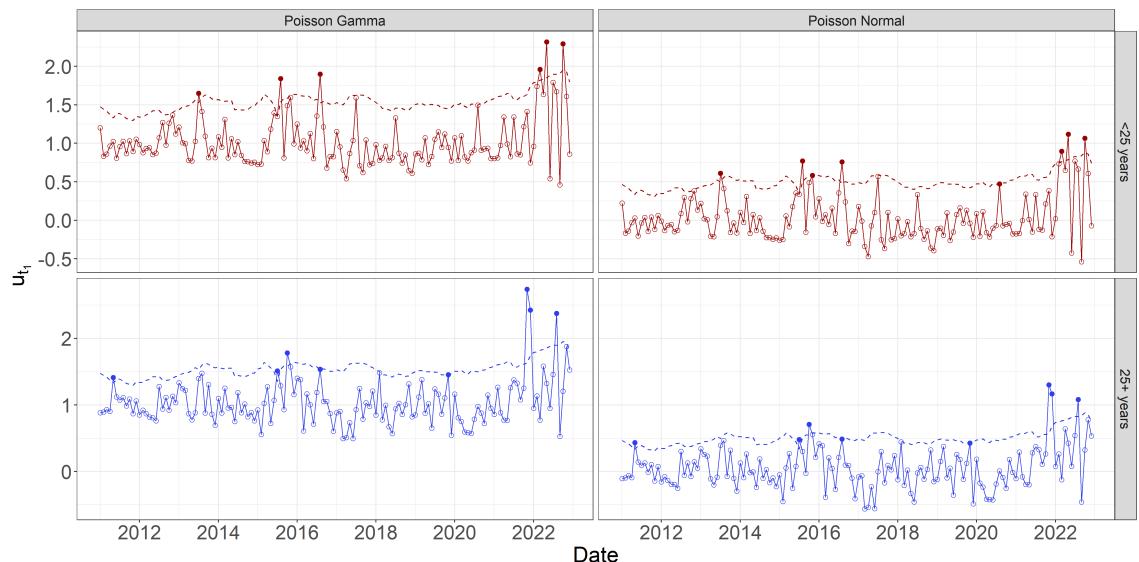


Figure 5.12: One-step ahead random effects  $u_{t_1}$  of *Shigellosis* in Denmark, 2011-2022, as circles. Poisson Normal model (top) and Poisson Gamma model (bottom) monitor the disease. Alarm raised (solid circle) if  $u_{t_1}$  exceeds the threshold (dashed line).

## 5.4 *Salmonellosis*

### 5.4.1 Applying the state-of-the-art outbreak detection algorithm to *Salmonellosis*

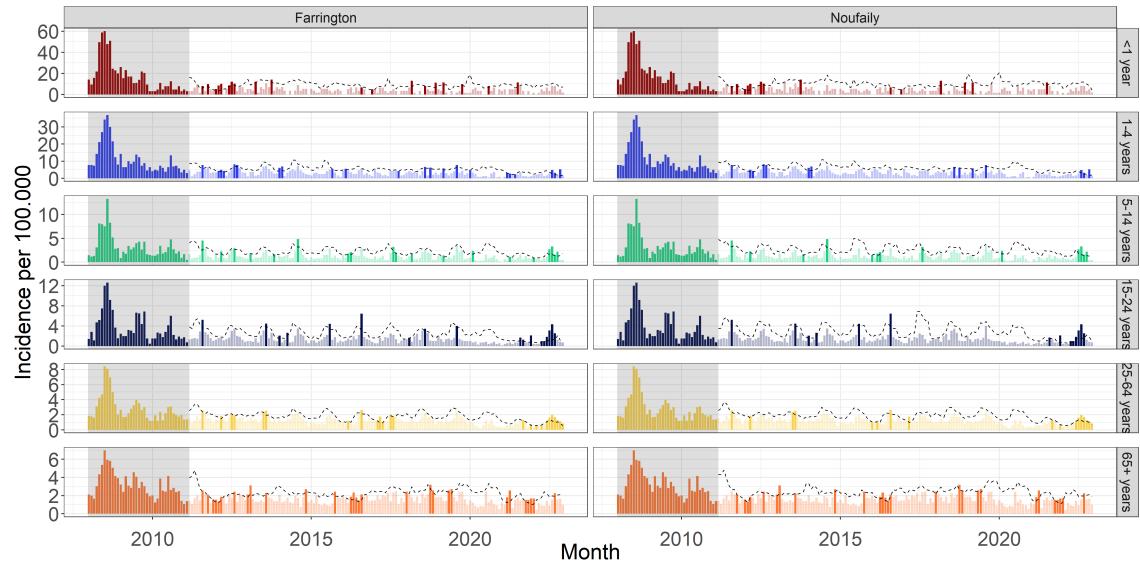


Figure 5.13: Monthly *Salmonellosis* incidence per 100.000 in Denmark, 2008-2022. Monitored by (left) Farrington and (right) Noufaily method. Reference data for the estimation of model parameters from January 2008 to March 2011 (grey area). Threshold (dashed line) is computed for observations timepoints outside reference data. Alarm triggered (full opacity) if observations exceeds threshold.

### 5.4.2 Applying the novel outbreak detection algorithm to *Salmonellosis*

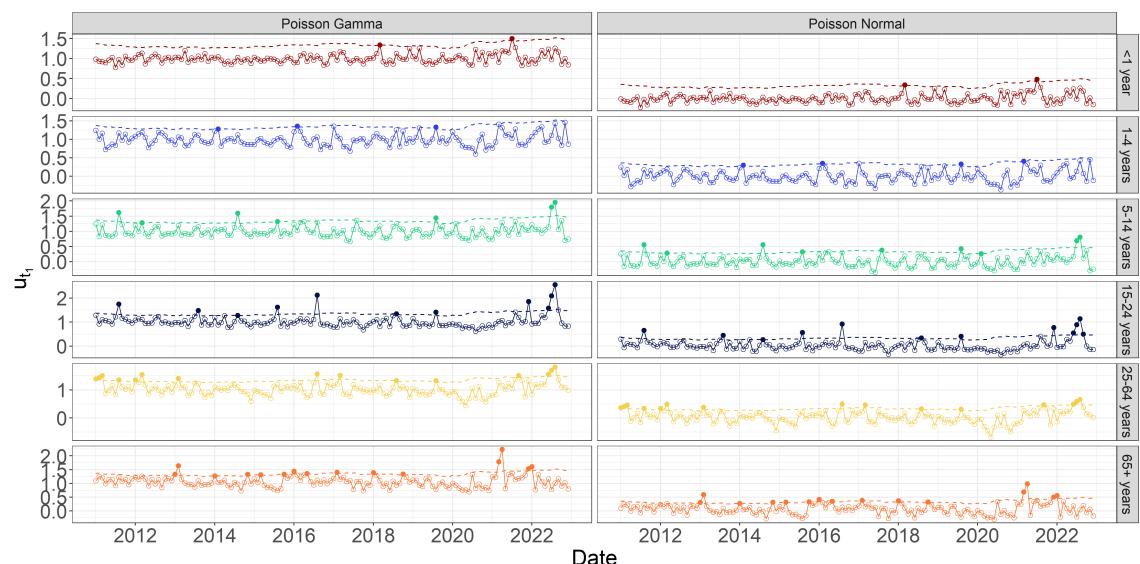


Figure 5.14: (ref:ref:CompareNovelSALM)

(#fig:ref:CompareNovelSALM)

## 5.5 Performance comparison of novel methods

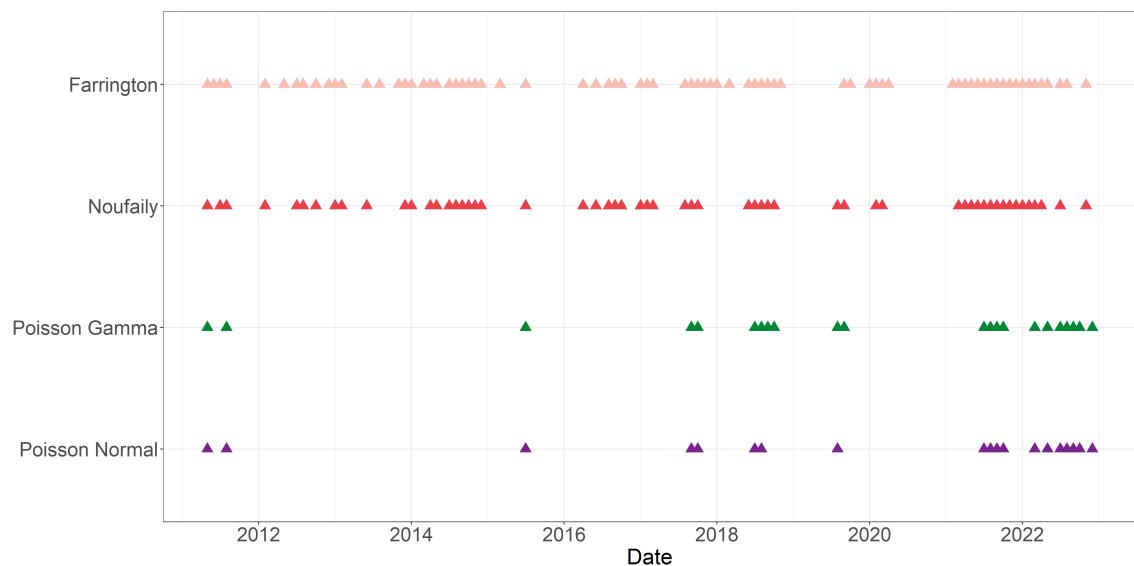


Figure 5.15: Placeholder caption

## 6 Simulation study

Hello, here is some text without a meaning. This text should show what a printed text will look like at this place. If you read this text, you will get no information. Really? Is there no information? Is there a difference between this text and some nonsense like “Huardest gefburn”? Kjift – not at all! A blind text like this gives you information about the selected font, how the letters are written and an impression of the look. This text should contain all letters of the alphabet and it should be written in of the original language. There is no need for special content, but the length of words should match the language.

Hello, here is some text without a meaning. This text should show what a printed text will look like at this place. If you read this text, you will get no information. Really? Is there no information? Is there a difference between this text and some nonsense like “Huardest gefburn”? Kjift – not at all! A blind text like this gives you information about the selected font, how the letters are written and an impression of the look. This text should contain all letters of the alphabet and it should be written in of the original language. There is no need for special content, but the length of words should match the language.



## 7 Discussion

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In the future, the utilization of MiBa-based surveillance has immense potential for disease surveillance. It has already demonstrated its value in various surveillance systems, such as the Healthcare-Associated Infections Database (HAIBA) for monitoring hospital-acquired infections and the COVID-19 surveillance system.

HAIBA, launched in 2015, was the first fully automated surveillance system built on MiBa data. It provides monitoring capabilities for hospital-acquired infections, enabling healthcare professionals to track and manage these infections more effectively. Similarly, the COVID-19 surveillance system, developed during 2020 and 2021, utilizes MiBa data to monitor and respond to the COVID-19 pandemic.

In addition to these systems, MiBa-based surveillance includes monitoring respiratory infections (such as influenza, pertussis, Mycoplasma pneumonia, and respiratory syncytial virus) and sexually transmitted diseases like chlamydia. While these surveillance systems currently have partial automation in data processing, there are plans to fully automate them in the near future.

Expanding on the field of automated disease outbreak detection is crucial to fully harness the potential of MiBa. By developing advanced algorithms and methodologies, it becomes possible to automatically analyze MiBa data and detect disease outbreaks in a timely manner. This can lead to early identification of outbreaks, allowing for prompt interventions and preventive measures.

Further research and development in automated disease outbreak detection, specifically tailored to leverage MiBa data, can significantly enhance our ability to detect and respond to infectious disease outbreaks more proactively and efficiently. By maximizing the potential of MiBa-based surveillance and continuously improving automated detection methods, we can strengthen our overall disease surveillance efforts and better protect public health.



## 8 Conclusion

Hello, here is some text without a meaning. This text should show what a printed text will look like at this place. If you read this text, you will get no information. Really? Is there no information? Is there a difference between this text and some nonsense like “Huardest gefburn”? Kjift – not at all! A blind text like this gives you information about the selected font, how the letters are written and an impression of the look. This text should contain all letters of the alphabet and it should be written in of the original language. There is no need for special content, but the length of words should match the language.



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# A Some probability functions

This chapter serves as a reference, specifying notation, properties, and moments related to the various distributions used in this master thesis.

Name	Support	Density	$E[Y]$	$V[Y]$
Poisson	$0, 1, 2, \dots$			
$\text{Pois}(\lambda)$	$\lambda \in \mathbb{R}_+$	$\frac{\lambda^y}{y!} \exp(-\lambda)$	$\lambda$	$\lambda$
Gamma	$\mathbb{R}_+$			
$G(\alpha, \beta)$	$\alpha \in \mathbb{R}_+, \beta \in \mathbb{R}_+$	$\frac{1}{\Gamma(\alpha)\beta} \left(\frac{y}{\beta}\right)^{\alpha-1} \exp(-y/\beta)$	$\alpha\beta$	$\alpha\beta^2$
Neg. Bin.	$0, 1, 2, \dots$			
$\text{NB}(r, p)$	$r \in \mathbb{R}_+, p \in ]0, 1]$	$\binom{r+y-1}{y} p^r (1-p)^y$	$\frac{r(1+p)}{p}$	$\frac{r(1-p)}{p^2}$
Normal	$\mathbb{R}$			
$N(\mu, \sigma^2)$	$\mu \in \mathbb{R}, \sigma^2 \in \mathbb{R}_+$	$\frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(y-\mu)^2}{2\sigma^2}\right)$	$\mu$	$\sigma^2$

Table A.1: Density, support, mean value, and variance for a number of distributions used in this master thesis.

## B C++ templates for the negative joint log-likelihood

This chapter presents the user templates for the hierarchical Poisson Normal model in (4.19) and the hierarachical Poisson Gamma model in (4.27).

The user template for the hierarchical Poisson Normal model specified in (4.19) is

```
#include <TMB.hpp>
template<class Type>
Type objective_function<Type>::operator() ()
{
    // R input data
    DATA_VECTOR(y);                                // Count data
    DATA_VECTOR(x);                                // Population size
    DATA_MATRIX(X);                               // Design matrix
    PARAMETER_VECTOR(u);                           // Random effects

    // Parameters
    PARAMETER_VECTOR(beta);                      // Fixed effects parameters
    PARAMETER(log_sigma_u);                     // Model parameter
    vector<Type> lambda = exp(X*beta-log(x)+u); // Construct 'lambda'
    Type sigma_u = exp(log_sigma_u);             // And the model parameters
    Type mean_ran = Type(0);

    // Objective function
    Type f = 0;                                    // Declare the objective
    f -= sum(dnorm(u,mean_ran,sigma_u,true));     // Calculate the objective
    f -= sum(dpois(y,lambda,true));                // Calculate the objective
    return f;
}
```

The user template for the hierarchical Poisson Gamma model specified in (4.27) is

```
#include <TMB.hpp>
template<class Type>
Type objective_function<Type>::operator() ()
{
    // Data
    DATA_VECTOR(y);                                // Count data
    DATA_VECTOR(x);                                // Population size
    DATA_MATRIX(X);                               // Design matrix

    // Parameters
    PARAMETER_VECTOR(beta);                      // Fixed effects parameters
    PARAMETER(log_phi_u);                     // Model parameter
    vector<Type> lambda = exp(X*beta-log(x));   // Construct 'lambda'
    Type phi_u = exp(log_phi_u);                 // And the model parameters
    Type r = 1/phi_u;                            // Construct the size
    vector<Type> p = 1/(lambda*phi_u+1);        // And the prob. parameter

    // Objective function
    Type f = -sum(dnbinom(y, r, p,true));       // Calculate the objective
```

```
    return f;  
}
```

# C State-of-the-art detection algorithm

## C.1 Controls

In the function `farringtonFlexible`, users can select either the original Farrington method or the improved method by Noufaily by specifying the appropriate `control` arguments. The choice of algorithm variant is determined by the contents of the `control` slot. In the example provided, `con.farrington` indicates the use of the original method, while `con.noufaily` represents the options for the improved method.

```
con.farrington <- list(
  range = NULL, b = 3, w = 2,
  reweight = TRUE, weightsThreshold = 1,
  verbose = TRUE, glmWarnings = TRUE,
  alpha = 0.05, trend = TRUE, pThresholdTrend = 0.05,
  limit54 = c(0,4), powertrans = "2/3",
  fitFun = "algo.farrington.fitGLM.flexible",
  populationOffset = TRUE,
  noPeriods = 1, pastWooksNotIncluded = NULL,
  thershodMethod = "delta"
)

con.noufaily <- list(
  range = NULL, b = 3, w = 2,
  reweight = TRUE, weightsThreshold = 2.58,
  verbose = TRUE, glmWarnings = TRUE,
  alpha = 0.05, trend = TRUE, pThresholdTrend = 0.05,
  limit54 = c(0,4), powertrans = "2/3",
  fitFun = "algo.farrington.fitGLM.flexible",
  populationOffset = TRUE,
  noPeriods = 1, pastWooksNotIncluded = NULL,
  thershodMethod = "Noufaily"
)
```

In this chapter an excerpt of the results

Table C.1: Longtable

Date	Age group	$y_t$	Farrington		Noufaily	
			Threshold	Alarm	Threshold	Alarm
2012-01-01	<65 years	2	3.407	FALSE	3.407	FALSE
	65+ years	6	4.345	TRUE	5.646	TRUE
2012-05-01	<65 years	1	3.594	FALSE	3.906	FALSE
	65+ years	4	2.444	TRUE	2.327	TRUE
2012-07-01	<65 years	0	4.424	FALSE	4.542	FALSE
	65+ years	5	2.715	TRUE	4.976	TRUE

Table C.1: (continued)

Date	Age group	$y_t$	Threshold	Alarm	Threshold	Alarm
<b>2013-09-01</b>	<65 years	1	2.749	FALSE	2.604	FALSE
	65+ years	7	6.321	TRUE	7.397	FALSE
<b>2013-10-01</b>	<65 years	0	4.384	FALSE	4.773	FALSE
	65+ years	6	4.551	TRUE	4.828	TRUE
<b>2014-01-01</b>	<65 years	0	3.920	FALSE	4.621	FALSE
	65+ years	8	5.530	TRUE	5.952	TRUE
<b>2014-07-01</b>	<65 years	6	3.093	TRUE	3.462	TRUE
	65+ years	14	10.223	TRUE	7.628	TRUE
<b>2014-08-01</b>	<65 years	6	3.093	TRUE	3.462	TRUE
	65+ years	14	11.690	TRUE	7.978	TRUE
<b>2014-09-01</b>	<65 years	2	3.061	FALSE	3.338	FALSE
	65+ years	9	6.705	TRUE	7.916	TRUE
<b>2016-04-01</b>	<65 years	3	2.474	TRUE	2.881	TRUE
	65+ years	1	5.990	FALSE	6.082	FALSE
<b>2017-01-01</b>	<65 years	0	2.228	FALSE	2.505	FALSE
	65+ years	8	6.005	TRUE	7.332	TRUE
<b>2017-04-01</b>	<65 years	1	2.440	FALSE	2.913	FALSE
	65+ years	4	2.789	TRUE	2.856	TRUE
<b>2017-07-01</b>	<65 years	3	4.541	FALSE	6.060	FALSE
	65+ years	4	3.736	TRUE	3.688	TRUE
<b>2018-02-01</b>	<65 years	3	2.618	TRUE	2.881	TRUE
	65+ years	1	6.040	FALSE	6.784	FALSE
<b>2018-06-01</b>	<65 years	0	3.569	FALSE	3.824	FALSE
	65+ years	6	5.935	TRUE	5.935	TRUE
<b>2019-02-01</b>	<65 years	0	2.622	FALSE	3.215	FALSE
	65+ years	7	5.962	TRUE	7.482	FALSE
<b>2019-09-01</b>	<65 years	2	3.446	FALSE	3.696	FALSE
	65+ years	7	6.686	TRUE	6.949	TRUE
<b>2021-05-01</b>	<65 years	5	1.905	TRUE	2.072	TRUE
	65+ years	1	5.209	FALSE	5.531	FALSE
<b>2021-07-01</b>	<65 years	0	2.802	FALSE	2.992	FALSE
	65+ years	8	7.258	TRUE	7.723	TRUE
<b>2021-12-01</b>	<65 years	1	2.595	FALSE	2.727	FALSE
	65+ years	5	4.979	TRUE	5.359	FALSE

Table C.1: (*continued*)

Date	Age group	$y_t$	Threshold	Alarm	Threshold	Alarm
<b>2022-04-01</b>	<65 years	1	2.759	FALSE	3.976	FALSE
	65+ years	5	5.776	FALSE	4.461	TRUE
<b>2022-05-01</b>	<65 years	6	2.714	TRUE	3.923	TRUE
	65+ years	9	5.562	TRUE	6.473	TRUE
<b>2022-06-01</b>	<65 years	3	2.957	TRUE	3.955	FALSE
	65+ years	11	6.700	TRUE	7.544	TRUE
<b>2022-10-01</b>	<65 years	5	3.864	TRUE	3.908	TRUE
	65+ years	7	8.298	FALSE	8.401	FALSE
<b>2022-11-01</b>	<65 years	3	3.481	FALSE	3.580	FALSE
	65+ years	11	7.213	TRUE	7.438	TRUE

## D Figures and Tables related to the case studies

### D.1 Listeriosis

Table D.1: The average logarithmic score,  $\bar{S}(G, y)$ , along with the parameter estimates at  $t_0$  for a suite of models modelling *Listeriosis*, assuming either the hierarchical Poisson Normal model or the hierarchical Poisson Gamma model. The confidence intervals for the estimates are calculated using profile likelihood confidence intervals.

$\bar{S}(G, y)$	Parameter	Estimate (95% CI)
<b>Poisson Normal</b>		
$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \log(n_{it})$		
3.597	$\beta_{<65\text{years}}$	15.43 (15.09, 15.73)
	$\beta_{65+\text{years}}$	15.19 (14.98, 15.36)
	$\log(\sigma)$	-3.09 (NA, -0.85)
$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \beta_{trend}t + \log(n_{it})$		
3.639	$\beta_{trend}$	0 (NA, NA)
	$\beta_{<65\text{years}}$	15.46 (NA, NA)
	$\beta_{65+\text{years}}$	15.22 (NA, NA)
	$\log(\sigma)$	-2.87 (NA, NA)
$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \beta_{\sin} \sin\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \beta_{\cos} \cos\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \log(n_{it})$		
3.577	$\beta_{<65\text{years}}$	15.4 (NA, NA)
	$\beta_{65+\text{years}}$	15.16 (NA, NA)
	$\beta_{\sin}$	-0.34 (NA, NA)
	$\beta_{\cos}$	0.02 (NA, NA)
	$\log(\sigma)$	-6 (NA, NA)
$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \beta_{trend}t + \beta_{\sin} \sin\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \beta_{\cos} \cos\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \log(n_{it})$		
3.663	$\beta_{trend}$	0 (NA, NA)
	$\beta_{<65\text{years}}$	15.46 (NA, NA)
	$\beta_{65+\text{years}}$	15.22 (NA, NA)
	$\beta_{\sin}$	0 (NA, NA)
	$\beta_{\cos}$	0.02 (NA, NA)
	$\log(\sigma)$	-2.83 (NA, NA)
<b>Poisson Gamma</b>		
$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \log(n_{it})$		
3.609	$\beta_{<65\text{years}}$	15.43 (15.1, 15.73)
	$\beta_{65+\text{years}}$	15.19 (15, 15.37)
	$\log(\phi)$	-6 (-36.77, -1.69)

### D.2 Shigellosis

Table D.2: The average logarithmic score,  $\bar{S}(G, y)$ , along with the parameter estimates at  $t_0$  for a suite of models modelling *Shigellosis*, assuming either the hierarchical Poisson Normal model or the hierarchical Poisson Gamma model. The confidence intervals for the estimates are calculated using profile likelihood confidence intervals.

$\bar{S}(G, y)$	Parameter	Estimate (95% CI)
<b>Poisson Normal</b>		
$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \log(n_{it})$		
5.291	NA	14.76 (14.39, 15.08)
	NA	16.39 (16.09, 16.65)
	$\log(\sigma)$	-0.67 (-1.33, -0.25)
$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \beta_{\text{trend}} t + \log(n_{it})$		
5.027	$\beta_{\text{trend}}$	0.04 (0.02, 0.06)
	NA	15.53 (15.07, 15.94)
	NA	17.21 (16.79, 17.6)
	$\log(\sigma)$	-0.55 (-0.95, -0.21)
$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \beta_{\sin} \sin\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \beta_{\cos} \cos\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \log(n_{it})$		
5.422	NA	14.74 (14.36, 15.07)
	NA	16.43 (16.13, 16.7)
	$\beta_{\sin}$	-0.01 (-0.3, 0.29)
	$\beta_{\cos}$	0.08 (-0.23, 0.38)
	$\log(\sigma)$	-0.59 (-1.1, -0.2)
$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \beta_{\text{trend}} t + \beta_{\sin} \sin\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \beta_{\cos} \cos\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \log(n_{it})$		
5.084	$\beta_{\text{trend}}$	0.04 (0.02, 0.06)
	NA	15.55 (15.1, 15.96)
	NA	17.22 (16.8, 17.61)
	$\beta_{\sin}$	-0.02 (-0.3, 0.27)
	$\beta_{\cos}$	0.21 (-0.06, 0.49)
	$\log(\sigma)$	-0.6 (-1.04, -0.24)
<b>Poisson Gamma</b>		
$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \log(n_{it})$		
5.257	NA	14.9 (14.58, 15.22)
	NA	16.5 (16.24, 16.78)
	$\log(\phi)$	-1.25 (-2.53, -0.46)
$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \beta_{\text{trend}} t + \log(n_{it})$		
5.019	$\beta_{\text{trend}}$	0.04 (0.02, 0.05)
	NA	15.66 (15.26, 16.07)
	NA	17.33 (16.94, 17.72)
	$\log(\phi)$	-1.04 (-1.83, -0.4)
$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \beta_{\sin} \sin\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \beta_{\cos} \cos\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \log(n_{it})$		
5.408	NA	14.97 (14.65, 15.29)
	NA	16.57 (16.3, 16.85)
	$\beta_{\sin}$	-0.03 (-0.33, 0.26)
	$\beta_{\cos}$	0.13 (-0.17, 0.42)

Table D.2: (continued)

Model	Formula	$\bar{S}(G, y)$	Parameter	Estimate (95% CI)
			$\log(\phi)$	-1.06 (-2.02, -0.36)
5.078	$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \beta_{trend} + \beta_{\sin} \sin\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \beta_{\cos} \cos\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \log(n_{it})$		$\beta_{trend}$	0.04 (0.02, 0.06)
			NA	15.68 (15.28, 16.1)
			NA	17.32 (16.94, 17.72)
			$\beta_{\sin}$	-0.02 (-0.3, 0.26)
			$\beta_{\cos}$	0.22 (-0.05, 0.5)
			$\log(\phi)$	-1.14 (-2.01, -0.47)

### D.3 Shiga toxin (verotoxin)-producing *Escherichia coli*

Table D.3: The average logarithmic score,  $\bar{S}(G, y)$ , along with the parameter estimates at  $t_0$  for a suite of models modelling Shiga toxin (verotoxin)-producing *Escherichia coli*, assuming either the hierarchical Poisson Normal model or the hierarchical Poisson Gamma model. The confidence intervals for the estimates are calculated using profile likelihood confidence intervals.

		$\bar{S}(G, y)$	Parameter	Estimate (95% CI)
<b>Poisson Normal</b>				
14.07	$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \log(n_{it})$		$\beta_{<1year}$	11.23 (10.88, 11.55)
			$\beta_{1-4years}$	13.45 (13.18, 13.72)
			$\beta_{5-14years}$	14.3 (14.04, 14.55)
			$\beta_{15-24years}$	14.71 (14.46, 14.94)
			$\beta_{25-64years}$	16.95 (16.73, 17.16)
			$\beta_{65+years}$	15.87 (15.66, 16.08)
			$\log(\sigma)$	-0.79 (-1.04, -0.57)
13.76	$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \beta_{trend} + \log(n_{it})$		$\beta_{trend}$	0.05 (0.04, 0.06)
			$\beta_{<1year}$	12.12 (11.79, 12.43)
			$\beta_{1-4years}$	14.61 (14.36, 14.85)
			$\beta_{5-14years}$	15.12 (14.85, 15.39)
			$\beta_{15-24years}$	15.54 (15.28, 15.79)
			$\beta_{25-64years}$	17.95 (17.73, 18.17)
			$\beta_{65+years}$	16.71 (16.48, 16.94)
13.19	$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \beta_{\sin} \sin\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \beta_{\cos} \cos\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \log(n_{it})$		$\beta_{<1year}$	11.29 (10.95, 11.6)
			$\beta_{1-4years}$	13.77 (13.52, 14)
			$\beta_{5-14years}$	14.3 (14.03, 14.56)
			$\beta_{15-24years}$	14.71 (14.46, 14.95)
			$\beta_{25-64years}$	17.14 (16.93, 17.34)

Table D.3: (continued)

Model	Formula	$\bar{S}(G, y)$	Parameter	Estimate (95% CI)
			$\beta_{65+years}$	15.91 (15.7, 16.12)
			$\beta_{\sin}$	-0.47 (-0.61, -0.33)
			$\beta_{\cos}$	-0.24 (-0.38, -0.1)
			$\log(\sigma)$	-0.65 (-0.84, -0.47)
13.12	$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \beta_{trend}t + \beta_{\sin} \sin\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \beta_{\cos} \cos\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \log(n_{it})$		$\beta_{trend}$	0.04 (0.03, 0.05)
			$\beta_{<1year}$	11.95 (11.68, 12.31)
			$\beta_{1-4years}$	14.51 (14.28, 14.73)
			$\beta_{5-14years}$	15 (14.75, 15.26)
			$\beta_{15-24years}$	15.42 (15.19, 15.66)
			$\beta_{25-64years}$	17.82 (17.64, 18.03)
			$\beta_{65+years}$	16.59 (16.4, 16.81)
			$\beta_{\sin}$	-0.33 (-0.43, -0.21)
			$\beta_{\cos}$	-0.22 (-0.33, -0.1)
			$\log(\sigma)$	-1.1 (-1.45, -0.87)
<b>Poisson Gamma</b>				
	$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \log(n_{it})$			
14.16			$\beta_{<1year}$	11.51 (11.2, 11.8)
			$\beta_{1-4years}$	13.56 (13.29, 13.82)
			$\beta_{5-14years}$	14.4 (14.15, 14.65)
			$\beta_{15-24years}$	14.8 (14.57, 15.03)
			$\beta_{25-64years}$	16.97 (16.76, 17.2)
			$\beta_{65+years}$	15.95 (15.74, 16.16)
			$\log(\phi)$	-1.52 (-2.02, -1.09)
13.82	$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \beta_{trend}t + \log(n_{it})$		$\beta_{trend}$	0.05 (0.04, 0.05)
			$\beta_{<1year}$	12.2 (11.88, 12.52)
			$\beta_{1-4years}$	14.71 (14.47, 14.95)
			$\beta_{5-14years}$	15.2 (14.94, 15.47)
			$\beta_{15-24years}$	15.64 (15.38, 15.89)
			$\beta_{25-64years}$	18.02 (17.81, 18.24)
			$\beta_{65+years}$	16.77 (16.54, 17)
			$\log(\phi)$	-1.77 (-2.24, -1.35)
13.24	$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \beta_{\sin} \sin\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \beta_{\cos} \cos\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \log(n_{it})$		$\beta_{<1year}$	11.42 (11.1, 11.73)
			$\beta_{1-4years}$	13.96 (13.73, 14.19)
			$\beta_{5-14years}$	14.43 (14.17, 14.68)
			$\beta_{15-24years}$	14.83 (14.59, 15.07)
			$\beta_{25-64years}$	17.26 (17.07, 17.47)
			$\beta_{65+years}$	16.05 (15.84, 16.26)
			$\beta_{\sin}$	-0.46 (-0.59, -0.32)
			$\beta_{\cos}$	-0.24 (-0.38, -0.1)
			$\log(\phi)$	-1.31 (-1.68, -0.98)

Table D.3: (continued)

Model	Formula	$\bar{S}(G, y)$	Parameter	Estimate (95% CI)
	$\log(\lambda_{it}) = \beta(\text{ageGroup}_i) + \beta_{trend}t + \beta_{\sin} \sin\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \beta_{\cos} \cos\left(\frac{\pi \cdot \text{monthInYear}_t}{6}\right) + \log(n_{it})$			
13.13	$\beta_{trend}$			0.04 (0.03, 0.05)
	$\beta_{<1year}$			12.06 (11.74, 12.35)
	$\beta_{1-4years}$			14.58 (14.36, 14.8)
	$\beta_{5-14years}$			15.06 (14.81, 15.31)
	$\beta_{15-24years}$			15.48 (15.25, 15.71)
	$\beta_{25-64years}$			17.88 (17.69, 18.07)
	$\beta_{65+years}$			16.65 (16.45, 16.85)
	$\beta_{\sin}$			-0.32 (-0.43, -0.21)
	$\beta_{\cos}$			-0.21 (-0.32, -0.1)
	$\log(\phi)$			-2.29 (-2.95, -1.77)

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Denmark

Richard Petersens Plads, Building 324  
2800 Kgs. Lyngby  
Tlf. 4525 1700

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