

# Introduction to Regression with R

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# The Regression Model

This is a linear regression model

$$Y = \beta_0 + \beta_1 X + U \quad (1)$$

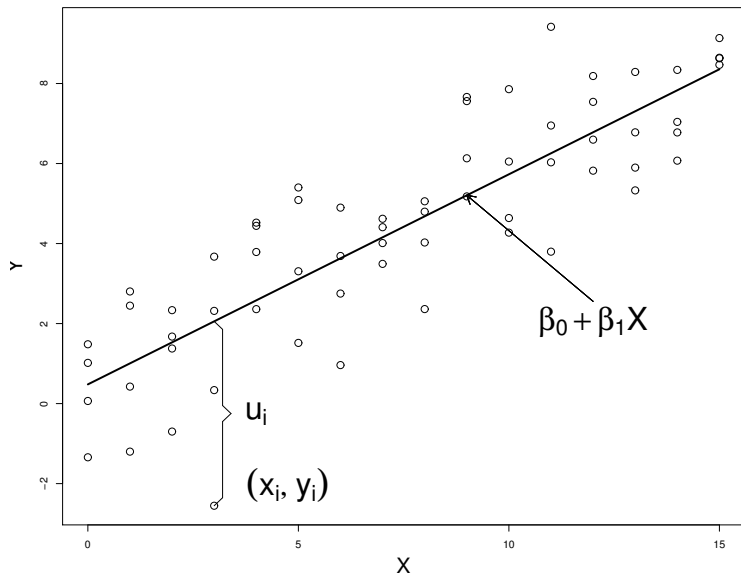
where

$Y$  : Dependent variable (also called respond, regressand and explained variable).

$X$  : Independent variable (also called covariate, regressor and explanatory variable).

$U$  : Error term,  $E(U) = 0$ . The error term includes all factors that affects  $Y$  except for  $X$ .

# The Regression model



# The regression line as a conditional average

Let us take the conditional mean of  $Y$  w.r.t.  $X$

$$E(Y|X) = E(\beta_0 + \beta_1 X + U|X) = \beta_0 + \beta_1 X, \quad (2)$$

$\beta_0 + \beta_1 X$  is the regression line.

- ▶ For the regression line to actually be a **conditional average**, we have to assume that  $E(U|X) = E(U) = 0$ .
- ▶ If this assumption is wrong, the regression line is not a conditional average but merely a **linear projection**
  - ▶ A linear approximation to the conditional average

# Least squares (LS)

First, we collect a (random) sample of data  $(X_i, Y_i)$ ,  $i = 1, \dots, n$

- ▶ Carl Friedrich Gauss realized, in the late 18<sup>th</sup> century, that we need a measure between the observation points and the line.
- ▶ His choice was

$$\sum_{i=1}^n U_i^2 = \sum_{i=1}^n (Y_i - \beta_0 - \beta_1 X_i)^2$$

and the LS estimators are based on the minimization problem:

$$\min_{\beta_0, \beta_1} \sum_{i=1}^n (Y_i - \beta_0 - \beta_1 X_i)^2$$

# The LS estimators

$$\hat{\beta}_1 = \frac{\sum(Y_i - \bar{Y})(X_i - \bar{X})}{\sum(X_i - \bar{X})^2},$$

$$\hat{\beta}_0 = \bar{Y} - \hat{\beta}_1\bar{X},$$

the fitted (predicted) values ( $\hat{E}(Y_i|X_i)$ )

$$\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i,$$

and the residuals

$$\hat{U}_i = Y_i - \hat{Y}_i.$$

## Exercise

Use the data *Prestige* in the package *car*. Consider the following model:

$$\text{prestige} = \beta_0 + \beta_1 \text{income} + U$$

1. We compute  $\hat{\beta}_1$  using the formula from the previous slide and then you compute:
2.  $\hat{\beta}_0$
3. and  $\hat{Y}_i$
4. and  $\hat{U}_i$

## Distributions of $\hat{\beta}_0$ and $\hat{\beta}_1$

- ▶ The  $\hat{\beta}_0$  and  $\hat{\beta}_1$  are estimators and random variables.
- ▶ They will vary between different samples and have distributions with probabilities of each possible outcome
- ▶ In small samples we have to assume  $U \sim N(0, \sigma^2)$ , then

$$T = \frac{\hat{\beta}_j - \beta_j}{S_{\hat{\beta}_j}} \sim t_{n-2}$$

- ▶ However if  $n$  is large we don't need the normality assumption, then

$$T = \frac{\hat{\beta}_j - \beta_j}{S_{\hat{\beta}_j}} \overset{a.}{\approx} N(0, 1)$$

- ▶ We will come back to the details in coming lectures



# T-testing of hypotheses about the parameters in a regression model (P-values)

1.  $H_0 : \beta_j = \beta_{j0}$

2.

$$H_a : \begin{cases} \beta_j > \beta_{j0} & \text{(upper-tail alternative)} \\ \beta_j < \beta_{j0} & \text{(lower-tail alternative)} \\ \beta_j \neq \beta_{j0} & \text{(two-tailed alternative)} \end{cases}$$

3. Test statistic:  $T = T_0 = \frac{\hat{\beta}_j - \beta_{j0}}{s_{\hat{\beta}_j}}$  under  $H_0$

4.

$$\text{P-value} = \begin{cases} P(T > t_0) & \text{(upper-tail alternative)} \\ P(T < t_0) & \text{(lower-tail alternative)} \\ 2P(T > |t_0|) & \text{(two-tailed alternative)} \end{cases}$$

# T-tests given by default by Statistical packages

1.  $H_0 : \beta_j = 0$
2.  $H_a : \beta_j \neq 0$
3. Test statistic:  $T = T_0 = \frac{\hat{\beta}_j}{S_{\hat{\beta}_j}}$  under  $H_0$
4. P-value =  $2P(T > |t_0|)$

i.e. for the coefficient on  $X$  the  $H_0$  is:  $X$  does not have an effect on  $Y$ .

If one rejects  $H_0$  it is common to say that  $X$  *has a significant effect on  $Y$* .

## Example using the `lm()` function

Now we will use the `lm()` function in R for computing the estimates in the previous exercise. Check the help file for `lm()` (`?lm()`).

Use the data *Prestige* in the package *car*. Consider the following model:

$$\text{prestige} = \beta_0 + \beta_1 \text{income} + U$$

Does *income* have a significant effect on *prestige*?

# Why LS?

## 1. Practical

- ▶ Most researchers know how LS works and how to interpret the results.
- ▶ All computing softwares have preprogrammed functions for LS estimation (Excel, Minitab, Matlab, R, gretl ...).

## 2. Theoretically attractive

- ▶ LS is consistent, asymptotic normal and also BLUE ("Best Linear Unbiased Estimator") under a special assumption (more about this latter).

# Why not LS?

- ▶ A major concern is that we assume a linear functional form.
  - ▶ Why? Because it is simple, mathematically
- ▶ Because more general models requires more data to give significant results.
  - ▶ There is often a choice between a general specification and no significant results and a linear model and significant results

## $R^2$

- How well does the linear relationship explain the data? One measure of fit is the so called  $R^2$  value.

$$R^2 = \frac{\sum(\hat{Y}_i - \bar{Y})^2}{\sum(Y_i - \bar{Y})^2} = \frac{ESS}{TSS} = 1 - \frac{\sum(Y_i - \hat{Y}_i)^2}{\sum(Y_i - \bar{Y})^2} = 1 - \frac{RSS}{TSS}.$$

where

ESS: Explained (variation) sum of squares

RSS: Residual (variation) sum of squares

TSS: Total (variation) sum of squares, and

$$TSS = ESS + RSS$$

$R^2$  is a consistent estimator of the following measure:

$$1 - \frac{\sigma_U^2}{\sigma_Y^2}$$

# $R^2$

- ▶  $R^2$  is interpreted as the relative fraction of 'variation' of  $Y$  explained by  $X$
- ▶ The range of  $R^2$  is limited to

$$0 \leq R^2 \leq 1$$

- ▶ If  $X$  do not explain any variation in  $Y \Leftrightarrow \hat{\beta}_1 = 0 \Rightarrow$

$$\begin{aligned} \sum (\hat{Y}_i - \bar{Y})^2 &= \sum (\hat{\beta}_0 + \hat{\beta}_1 X_i - \bar{Y})^2 = \sum (\hat{\beta}_0 - \bar{Y})^2 = \\ &\underbrace{=}_{[\hat{\beta}_0 = \bar{Y} - \hat{\beta}_1 \bar{X}]} \sum (\bar{Y} - \bar{Y})^2 = 0 \Rightarrow \underline{\underline{R^2 = 0}} \end{aligned}$$

- ▶ If  $X$  explains all variation in  $Y \Rightarrow \hat{Y}_i = Y_i \Rightarrow \underline{\underline{R^2 = 1}}$

# The Standard Error of the Regression

- ▶ The Standard Error of the Regression is defined as

$$s_U = \sqrt{s_U^2} = \sqrt{\frac{\sum \hat{U}_i^2}{n-2}} = \sqrt{\frac{\sum (Y_i - \hat{Y}_i)^2}{n-2}}$$

This is an estimator of the standard deviation of  $U$ , and thus measures the spread around the  $E(Y|X) - line$ .



## Exercise

Compute  $R^2$  for the Prestige regression in the previous example.  
Compare your  $r^2$  to the ones reported by `lm()`.

# Remarks

- ▶ A relatively small  $R^2$  value (e.g 0.1) and a large  $S_U$  do not imply that  $X$  is not important; only that there is other factors accumulated in  $U$  that also explain variation in  $Y$
- ▶  $X$  could still be the single variable that explains most variation in  $Y$ , but of course the probability for this decreases with  $R^2$ .
- ▶ How large  $R^2$  is supposed to be is arbitrary and differs between different research disciplines and applications.

# Regression when the explanatory variables is binary

Binary variables, also called Dummy or indicator variables, e.g.

$$D = \begin{cases} 1 & \text{if Female} \\ 0 & \text{if Male} \end{cases} \quad \text{or} \quad D = \begin{cases} 1 & \text{if Medicine} \\ 0 & \text{if Placebo} \end{cases}$$

$$Y = \beta_0 + \beta_1 D + U$$

$$E(Y|D = 0) = \beta_0$$

$$E(Y|D = 1) = \beta_0 + \beta_1$$

$$E(Y|D = 1) - E(Y|D = 0) = \beta_1$$

- Thus in the simple regression framework, with one binary regressor, the interpretation of  $\beta_1$  is the difference in average of  $Y$  given group belonging,  $D = 1$  or  $D = 0$ .

# Next

- ▶ Next time we start with the second exercise set. You will mainly work with the `lm()` function in R. Estimate regressions and interpret results based on today's lecture.
- ▶ Next lecture we will examine some matrix algebra related to least-squares regression (when the number of X-variables are more than one).