STANDARDIZING TYPE Ia SUPERNOVA ABSOLUTE MAGNITUDES USING GAUSSIAN PROCESS DATA REGRESSION

A. G. Kim¹, R. C. Thomas², G. Aldering¹, P. Antilogus³, C. Aragon¹, S. Bailey¹, C. Baltay⁴, S. Bongard³, C. Buton⁵, A. Canto³, F. Cellier-Holzem³, M. Childress^{1,6}, N. Chotard^{7,8}, Y. Copin^{9,10,11}, H. K. Fakhouri^{1,6}, E. Gangler^{9,10,11}, J. Guy³, M. Kerschhaggl⁵, M. Kowalski⁵, J. Nordin¹, P. Nugent², K. Paech⁵, R. Pain³, E. Pecontal¹², R. Pereira^{9,10,11}, S. Perlmutter^{1,6}, D. Rabinowitz⁴, M. Rigault^{9,10,11}, K. Runge¹, C. Saunders^{1,6}, R. Scalzo¹³, G. Smadja^{9,10,11}, C. Tao^{7,14}, B. A. Weaver¹⁵, and C. Wu^{3,8} ¹ Physics Division, Lawrence Berkeley National Laboratory, 1 Cyclotron Road, Berkeley, CA 94720, USA ² Computational Cosmology Center, Computational Research Division, Lawrence Berkeley National Laboratory, 1 Cyclotron Road MS 50B-4206, Berkeley, CA 94720, USA ³ Laboratoire de Physique Nucléaire et des Hautes Énergies, Université Pierre et Marie Curie Paris 6, Université Denis Diderot Paris 7, CNRS-IN2P3, 4 place Jussieu, F-75252 Paris Cedex 05, France Department of Physics, Yale University, New Haven, CT 06250-8121, USA ⁵ Physikalisches Institut, Universität Bonn, Nußallee 12, D-53115 Bonn, Germany ⁶ Department of Physics, University of California Berkeley, 366 LeConte Hall MC 7300, Berkeley, CA 94720-7300, USA ⁷ Tsinghua Center for Astrophysics, Tsinghua University, Beijing 100084, China ⁸ National Astronomical Observatories, Chinese Academy of Sciences, Beijing 100012, China ⁹ Université de Lyon, F-69622 Lyon, France ¹⁰ Université de Lyon 1, Villeurbanne, France ¹¹ CNRS/IN2P3, Institut de Physique Nucléaire de Lyon, Lyon, France 12 Centre de Recherche Astronomique de Lyon, Université Lyon 1, 9 Avenue Charles André, F-69561 Saint Genis Laval Cedex, France ¹³ Research School of Astronomy and Astrophysics, The Australian National University, Mount Stromlo Observatory, Cotter Road, Weston Creek, ACT 2611, Australia ¹⁴ Centre de Physique des Particules de Marseille, 163 Avenue de Luminy, Case 902, F-13288 Marseille Cedex 09, France ¹⁵ Center for Cosmology and Particle Physics, New York University, 4 Washington Place, New York, NY 10003, USA

ABSTRACT

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We present a novel class of models for Type Ia supernova time-evolving spectral energy distributions (SEDs) and absolute magnitudes: they are each modeled as stochastic functions described by Gaussian processes. The values of the SED and absolute magnitudes are defined through well-defined regression prescriptions, so that data directly inform the models. As a proof of concept, we implement a model for synthetic photometry built from the spectrophotometric time series from the Nearby Supernova Factory. Absolute magnitudes at peak B brightness are calibrated to 0.13 mag in the g band and to as low as 0.09 mag in the g bulleting to spectrophotometric time series of supernovae that span a range of redshifts to simultaneously standardize supernovae together with fitting cosmological parameters.

Key words: distance scale – methods: data analysis – supernovae: general *Online-only material:* color figures

1. INTRODUCTION

Type I supernovae were found to be good standard candles with a low dispersion in visible-wavelength peak luminosities (Kowal 1968). As supernovae were further subclassified, Type Ia supernovae (SNe Ia) were found to be even better calibrated candles by accounting for temporal (Phillips 1993) and color (Tripp 1998) information transmitted in their light curves. Modern distance determinators fit supernova measurements to a time-evolving spectral energy distribution (SED) model with free parameters that adjust the light-curve shapes and colors. An individual supernova's peak absolute magnitude is derived directly from the fit (e.g., Jha et al. 2007; Mandel et al. 2011, MLCS2k2, BayeSN) or from a secondary relation based on the fit parameters (e.g., Guy et al. 2007; Conley et al. 2008; Burns et al. 2011, SALT2, SIFTO, SNooPy).

There are challenges in empirically modeling the family of SN Ia SEDs and relating supernova observables with absolute magnitude. Temporal coverage and sampling, and SN-frame photometric coverage vary from supernova to super-

nova; seldom are there data of different supernovae taken at common phase and band that can be directly compared. Data are fit to SED models to transform these heterogeneous observations into parameters that have common meaning across all supernovae. Constructing these models requires physical insight, diligent scrutiny of data, and/or extreme flexibility in order to ensure capturing the parameters that describe SN Ia heterogeneity. Once those parameters are identified, selecting the functional form to model their relationship with absolute magnitude is non-trivial: the behavior is known to be nonlinear and there is the difficulty of choosing a function that leaves no residual bias over the full parameter space while not overfitting the realized data set. We therefore seek an alternative to choosing the explicit functions for the time-evolving SED and the parameter-absolute-magnitude relationship.

This article introduces an alternative approach: the evolving SED for a supernova plus foreground dimming are modeled as a Gaussian process: a function whose value at each input is known to lie within a normal distribution but whose precise value is

unknown and whose values at different inputs may be correlated (Rasmussen & Williams 2006). The simple form of the normal distribution leads to analytic expressions for the likelihood of the data for the Gaussian process model, and for the predicted mean and covariance of the fluxes at any wavelength and epoch. Modeled as a Gaussian process, the description of the underlying SED of a single supernova is influenced by each datum and thus has more degrees of freedom than either of the SEDs underlying MLCS2k2 and SALT2. Here supernova absolute magnitudes are modeled as a function of the flux values at different epochs and wavelengths regressed using the first Gaussian process. In one case a linear model is used. In a second, Gaussian processes are again used to model supernova absolute magnitudes as a nonlinear function. Previous applications of Gaussian processes in the analysis of supernova data are given in Holsclaw et al. (2010) and Shafieloo et al. (2013).

Our approach contrasts with the usage of Gaussian processes in Mandel et al. (2011). We model the covariance between single-band absolute magnitudes of different supernovae with a parameterized function that depends on differences between light-curve shapes and colors. The absolute magnitude of a supernova is regressed from the absolute magnitudes of other supernovae. The distance is then inferred from the observed and absolute magnitude in that band; independent information potentially available in the distances from other bands is not included. Mandel et al. (2011) focus on supernova distance and its covariance with a suite of light-curve parameters. There, independent elements of the covariance matrix are fit in training. The distance of a supernova is then inferred directly from its light-curve parameters and the trained covariance matrix.

The approach presented in this article can be used in the analysis of existing supernova data and can inform the design of future experiments. To date, the Nearby Supernova Factory (SNfactory; Aldering et al. 2002) has obtained and processed spectrophotometric time series for a sample of 132 supernovae within the linear Hubble flow. These data can serve as a template sample to search for indicators of diversity in luminosity, and contribute to the low-redshift anchor when combined with available high-redshift SN Ia data and future photometric surveys such as the Dark Energy Survey (DES), 16 LSST, 17 or high-redshift spectrophotometric surveys as espoused by the Interim Science Working Group of the Joint Dark Energy Mission (Baltay & Moos 2010).

In Section 2, we present in detail how Gaussian processes can determine SN Ia distances from broadband photometry based on training with a low-redshift spectrophotometric sample. The framework for analyzing spectrophotometric data to calibrate absolute magnitudes and/or simultaneously fit for cosmological parameters is sketched in Section 4. We finish with conclusions in Section 5.

2. CASE ILLUSTRATION: INFERRING LUMINOSITY FROM BROADBAND PHOTOMETRY

The methodology to determine absolute magnitudes from broadband light curves is applied to synthetic photometry generated from low-redshift SNfactory flux-calibrated spectroscopy (spectrophotometry). Two steps are involved. First, the photometry in all bands of all supernovae is modeled as a Gaussian process. Based on that model, interpolated multi-band light curves are constructed for each supernova. Then, the light-curve shapes

and colors derived from the interpolations are treated as parameters that are related to absolute magnitudes using a second Gaussian process. Based on that model, the absolute magnitude of a supernova with given light-curve shapes and colors is predicted.

The motivation is to calibrate absolute magnitudes for supernovae at redshift z observed in the DES griz photometric system. We work in the supernova rest frame, treating observations as if taken in the appropriate blueshifted DES filter set. Focusing the analysis for specific redshifts is made possible for the first time by our extensive set of spectrophotometric data and eliminates K-correction uncertainties (Kim et al. 1996). In this article, we consider four redshifts and thus four filter sets corresponding to the rest-frame coverage of z=0.00, 0.25, 0.50, 0.75 supernovae. The absolute magnitudes in each of the blueshifted griz bands at the epoch of B-band peak brightness are calibrated.

The analysis procedure for each redshift of interest is summarized as follows.

- 1. Synthetic photometry of SNfactory data is generated (Section 2.2).
- 2. The synthetic light curves from all supernovae are modeled as a Gaussian process. Best-fit parameters for the model are determined (Section 2.3.1).
- 3. SN light curves for a grid of epochs and bands and their associated covariance matrix are constructed by interpolating between a supernova's photometry using the regression feature of Gaussian processes. Subtracting the interpolated absolute magnitude at *B* peak for a fiducial band gives light-curve shapes and colors. For computational convenience the shapes and colors are re-expressed in terms of their principal components (Section 2.3.2).
- 4. The absolute magnitude at *B* peak as a function of the light-curve shapes and colors is modeled either as a second Gaussian process or a linear function and trained (Section 2.3.3).
- 5. The trained procedure is applied to predict absolute magnitudes for an independent validation set. This is done four times, running with completely independent validation sets and corresponding complementary training sets. We report the average weighted rms (wrms) and standard deviation of the four-fold validation runs (Section 2.4).

2.1. Computing

Certain steps of Gaussian process regression involve the inversion or decomposition of matrices having dimension equal to the number of data points n. To handle large data sets (thousands of points and up), we implemented a framework of Gaussian process algorithms especially for use with high-performance parallel architectures. However, we note that there are alternative strategies for mitigating the $\mathcal{O}(n^3)$ matrix inversion cost, such as selective pruning of data and sparse linear algebra methods, but exploring them is beyond the scope of the present paper. Our implementation uses the well-tested ScaLAPACK parallel distributed linear algebra library (Blackford et al. 1997).

The code base has the flexibility to operate on multiple platforms. Development was performed using one or two processors of a 2.53 GHz Intel Core 2 Duo MacBook Pro. The full analysis was run using 144 processors of Hopper, NERSC's Cray XE6. A single training/validation analysis took 5 wall-clock minutes to complete.

High-performance computing facilitated code development and training: we were able to quickly cycle through tests using

¹⁶ http://www.darkenergysurvey.org/

¹⁷ http://www.lsst.org/lsst/

different Gaussian process models and optimize parameters in likelihood fits. Once training of the model is completed, determining the absolute magnitudes of the validation set is not computationally intensive.

2.2. Data

The input to our analysis is the portion of the spectrophotometric data set obtained by the SNfactory between 2004 and 2009 with the SuperNova Integral Field Spectrograph (SNIFS; Lantz et al. 2004) that had been fully processed as of early 2011. SNIFS is a fully integrated instrument optimized for automated observation of point sources on a structured background over the full ground-based optical window at moderate spectral resolution ($R \sim 500$). It consists of a high-throughput wide-band pure-lenslet integral field spectrograph (IFS; "à la TIGER"; Bacon et al. 1995, 2000, 2001), a multi-filter photometric channel to image the field in the vicinity of the IFS for atmospheric transmission monitoring simultaneous with spectroscopy, and an acquisition/guiding channel. The IFS possesses a fully filled $6''.4 \times 6''.4$ spectroscopic field of view subdivided into a grid of 15 × 15 spatial elements, a dual-channel spectrograph covering 3200–5200 Å and 5100–10000 Å simultaneously, and an internal calibration unit (continuum and arc lamps). SNIFS is mounted on the south bent Cassegrain port of the University of Hawaii 2.2 m telescope on Mauna Kea, and is operated remotely. Observations are reduced using the SNfactory's dedicated data reduction pipeline, similar to that presented in Section 4 of Bacon et al. (2001). A discussion of the software pipeline is presented in Aldering et al. (2006) and is updated in Scalzo et al. (2010). A detailed description of host-galaxy subtraction is given in Bongard et al. (2011).

The instrument characteristics and observing strategy enable the construction of spectrophotometric time series of supernovae sampled every 2–3 days up until ~25 days after maximum, and then 5–7 days out to ~45 days after maximum. The median signal-to-noise ratio (S/N) within 2.5 days of peak brightness is 10.2 per 2.4 Å bin. In our sample, 83% of the objects have z>0.03 where peculiar velocities of order $300~{\rm km~s^{-1}}$ are expected to contribute less than 0.07 mag to the absolute magnitude dispersion (in fact 98% have z>0.01). At the other extreme, 90% have z<0.08 and so are insensitive to present uncertainties in Ω_M or the dark energy equation of state, w.

The spectral time series used are corrected for Milky Way dust extinction (Cardelli et al. 1989; Schlegel et al. 1998) but no attempt is made to correct for the effects of circumstellar or host-galaxy dust extinction. Reference observations when the supernova has faded away facilitate host-galaxy subtraction (Bongard et al. 2011). Synthetic photometry light curves are fit using SALT2 (Guy et al. 2007) to obtain the date of *B*-band maximum in order to assign a phase to each spectrum. For our analysis, each spectral time series is blueshifted to rest frame and the fluxes normalized to represent that observed at a common distance for all supernovae assuming a flat universe with $\Omega_M = 0.28$, h = 0.7—thus, ratios in flux between supernovae can be interpreted as ratios in luminosity or equivalently differences in absolute magnitude.

Synthetic photometry and uncertainties are generated for each of the blueshifted DES filter sets when there is sufficient wavelength coverage, including the few cases when data from only one arm of the SNIFS spectrograph are available. Spectra from the two arms generally provide coverage of DES gri for z=0.00, griz for z=0.25, and riz for z=0.50, 0.75; these

redshifts provide different rest-frame wavelength ranges with which to calibrate supernova absolute magnitudes. Photon fluxes are calculated and are converted to magnitudes, as we find that the dominant sources of Gaussian dispersion are multiplicative (e.g., flux calibration) rather than additive (e.g., detector dark current). As the magnitudes refer to supernovae normalized to be at a common distance we refer to them as absolute magnitudes; the zero point of the magnitude system is irrelevant for the purposes of this article.

Application of SALT2 on Johnson–Cousins synthetic photometry of the full input data set calibrates absolute magnitudes to a wrms of 0.147 mag, similar to the 0.16 mag found in earlier data productions (Bailey et al. 2009; Chotard et al. 2011).

2.3. Modeling Supernova Absolute Magnitudes

2.3.1. Modeling Supernova Light Curves

We begin by training a parameterized model for the multi-band light curves that make up our data set. For a single blueshifted filter set, the multi-band light curves underlying supernovae are modeled as arising from a Gaussian process,

$$m_{(t,\lambda)} \sim \text{GP}(\bar{m}(t,\lambda;\bar{\mathbf{m}}_0), k_m(t,\lambda,t',\lambda'; l_{k_m}, \boldsymbol{\sigma}_{k_m})),$$
 (1)

where $m_{(t,\lambda)}$ is the photometric magnitude at epoch t relative to the B-band peak in the filter indexed by λ . (The definition of a Gaussian process and associated equations used in the analysis of this article are given in the Appendix.) Subscripted (t,λ) represent points whereas those in unsubscripted parentheses are function variables. This notation means that for a set of coordinates $\{(t,\lambda)\}$, the magnitude values are drawn from a normal distribution

$$m_{\{(t,\lambda)\}} \sim \mathcal{N}\left(\{\bar{m}(t,\lambda;\bar{\mathbf{m}}_0)\},K\right),$$
 (2)

where the elements of the model covariance matrix are

$$K_{(t,\lambda),(t',\lambda')} = k_m(t,\lambda,t',\lambda';l_{k_m},\boldsymbol{\sigma}_{k_m}). \tag{3}$$

Measurements of a realized supernova are then also described by a Gaussian process,

$$m_{(t,\lambda)}^{\star} \sim \text{GP}(\bar{m}(t,\lambda;\bar{\mathbf{m}}_0), k_m(t,\lambda,t',\lambda'; l_{k_m}, \boldsymbol{\sigma}_{k_m}) + n_m(t,\lambda,t',\lambda'; \boldsymbol{\sigma}_{n_m})). \tag{4}$$

Given a set of measurements $m_{(t,\lambda)}^{\star}$, the likelihood that the data arise from a Gaussian process can be expressed analytically, as well as the mean and covariance of the probability distribution function of a finite set of underlying points $m_{(t,\lambda)}$ (see the Appendix).

For our model, we choose as mean function $\bar{m}(t,\lambda;\bar{\mathbf{m}}_0)$ the synthetic photometry of band λ from the updated templates of Hsiao et al. (2007) offset by zero points $\bar{\mathbf{m}}_0$ for each band. The mean function could include a parameter for the date of peak B magnitude but the spectral time series we work with has been preprocessed and is already expressed in terms of days relative to peak determined by SALT2. A zero mean function would lead to expectations of zero flux in temporal regions lacking data; the Hsiao mean is used to avoid abrupt deviations toward zero where there are temporal gaps in the light curves, whose locations differ from supernova to supernova.

The kernel represents the covariance of the underlying flux: we use the square exponential kernel

$$k_m(t, \lambda, t', \lambda'; l_{k_m}, \boldsymbol{\sigma}_{k_m}) = \sigma_{k_m}^2(\lambda) \delta_{\lambda \lambda'} \exp \left[-\left(\frac{t - t'}{l_{k_m}}\right)^2 \right],$$
(5)

 Table 1

 Best-fit Hyperparameters of the Gaussian Process that Models the Supernova Multi-band Light Curves, Given in Equation (1)

System z	l_{k_m}	$\sigma_{k_m}(g)$	$\sigma_{k_m}(r)$	$\sigma_{k_m}(i)$	$\sigma_{k_m}(z)$	$\sigma_{n_m}(g)$	$\sigma_{n_m}(r)$	$\sigma_{n_m}(i)$	$\sigma_{n_m}(z)$
0.00	5.63	0.09	0.09	0.12		0.05	0.06	0.07	
0.25	5.87	0.13	0.08	0.09	0.13	0.14	0.05	0.06	0.06
0.50	6.18		0.11	0.08	0.10		0.06	0.05	0.05
0.75	6.56		0.10	0.11	0.11		0.14	0.06	0.05

Notes. Given are the results of one of the four training sets for each of the blueshifted DES photometric systems. All values have magnitude units except l_{k_m} , given in supernova-frame days.

where each band has its own normalization factor $\sigma_{k_m}(\lambda)$ but shares a common timescale l_{k_m} . The l_{k_m} does not represent the timescale of the light curves themselves, but rather those of deviations from the mean model. The kernel does not correlate the light curves in different bands. Although supernova flux across bands is correlated, for this stage of the analysis we preferred imposing no prior assumptions on color. We opt for the square exponential kernel as it is frequently used in the Gaussian process literature; model testing comparing the likelihoods of other kernels is deferred to future study.

Measurements of m include a nugget term that contains all other sources of variance in the photometry. Our nugget model is

$$n_m(t, \lambda, t', \lambda'; \boldsymbol{\sigma}_{n_m}) = \left(\sigma_{n_m}^2(\lambda) + \operatorname{var}(m_{(t,\lambda)})\right) \delta_{t,t'} \delta_{\lambda,\lambda'}, \quad (6)$$

where within the scope of our analysis there are no two measurements with the same (t, λ) so the nugget applies per measurement. Each band has its own intrinsic dispersion and the SNfactory photon noise has variance $\text{var}(m_{(t,\lambda)})$. The measurement covariance is very small given the small overlap of the griz bands, and so is ignored.

The kernel includes an additional term that accounts for radial peculiar velocities (Bonvin et al. 2006):

$$\sigma_{\alpha\beta}^2 = \left[\sigma_z^{\text{pec}} \frac{5}{\ln{(10)}} \frac{(1+z_\alpha)}{\chi H} \right]^2 \delta_{\alpha\beta}, \tag{7}$$

where the peculiar velocity dispersion, $\sigma_z^{\rm pec}$, can be either a fixed or fit parameter, and z_α is the actual redshift of the supernova; in this paper, we fix $\sigma_z^{\rm pec}=300~{\rm km~s^{-1}}$. Here α and β are the supernova indices, so the $\delta_{\alpha\beta}$ reflects the common realization of peculiar velocity for a single supernova. At low redshifts, the conformal distance can be conveniently approximated as $\chi \approx z - 0.75 \Omega_M z^2$ for a flat $\Lambda \rm CDM$ cosmology.

For each of four training/validation runs, 75% of the SNfactory supernovae are designated as the training set. Photometry with S/N < 50 are culled as often being associated with poor extractions and a supernova must have at least 16 light-curve points total over all bands to be included. Cuts are applied after the division into training and validation sets, so each of the analyses do not have exactly the same sample sizes: after the S/N and number-of-point cuts there remains 80–90 training supernovae per run.

All training-set supernovae are used to fit the free parameters (and hyperparameters) that give the maximum likelihood of the Gaussian process model. The hyperparameters l_{k_m} , σ_{k_m} , and σ_{n_m} are common to all SNe Ia and are fit simultaneously to the full ensemble of data, whereas a distinct $\bar{\mathbf{m}}_0$ is fit for each supernova. The best-fit hyperparameters for l_{k_m} , σ_{k_m} , and σ_{n_m} for one of the four-fold training runs are shown in Table 1; these values do

not differ significantly between the different realizations. The strength of the model covariance between points in a single light-curve range from 0.08 to 0.14 mag with correlation-length scales of around 6 days. The intrinsic dispersion is 0.05–0.07 mag except in the rest-frame UV where the dispersion is 0.14 mag.

The regressed g-band date of maximum has a distribution with mean and standard deviation of 0.058 ± 0.76 days. The temporal alignment of our model is consistent with that of SALT2.

To summarize this first step in the analysis, supernova light curves are described by a parameterized model for the covariance in their residuals about a parameterized mean function. Light-curve photometry of a supernova training set is used to optimize the parameters that best describe their scatter.

2.3.2. Supernova Parameterization

We now turn to the supernova parameters used to describe the light curve of a single event, in analogy to the light-curve shape and color parameters used in classical light-curve fitters. The parameters adopted in this article are the (linearly transformed) values of interpolated light curves.

Since we are ultimately interested in determining the magnitude at peak brightness, both training and validation sets are pared down to include only those supernovae having a first measurement at least two days before B peak to avoid extrapolating the value of $M_{(\lambda_0, t_0)}$. This leaves 49–57 training supernovae.

Observed supernovae have disparate temporal sampling; those data must be transformed into a set of parameters that can be directly compared between supernovae. The Gaussian process model of the light curve is used to predict the fluxes on a common set of epochs, daily from -10 to 35 rest-frame days relative to *B* peak, in the blueshifted DES bands that are spanned by SNfactory data. (The choice of daily sampling of the light curves was arbitrary and probably finer than necessary.) The predictions (Equation (A5)) have covariance (Equation (A6)) so 50 light-curve grids are realized per supernova to include uncertainties in the training and propagate uncertainties in the inferences. Predictions and realizations of regressed light curves from sample supernova photometric measurements are shown in Figures 1–3.

We want to calibrate the supernova absolute magnitude in some band λ_0 at the phase of peak B brightness t_0 : this is referred to as the "true" absolute magnitude and is estimated by the $M_{(\lambda_0,t_0)}$ predicted by our model. This information is removed from the predicted light curves by subtracting $M_{(\lambda_0,t_0)}$ from all other magnitudes. What remains is a light-curve shape in the selected band λ_0 and colors for the other bands, giving 137 (46 × 3 – 1 for the case of three bands) or 183 (46 × 4 – 1 for four bands) potential parameters to describe a supernova realization.

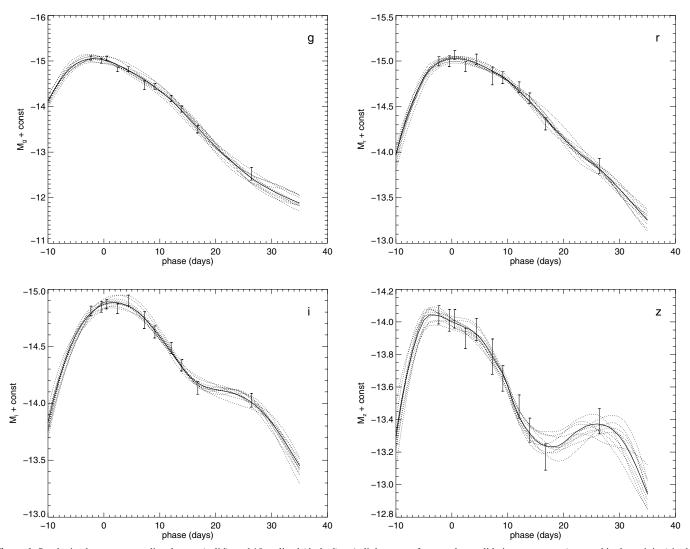


Figure 1. Synthetic photometry, predicted mean (solid), and 10 realized (dashed) griz light curves for a random validation supernova (not used in the training) in the run with the z = 0.25 filter set.

For practical reasons, we re-express the light-curve shapes and colors in terms of their principal components ordered by the dimensions that carry the leading population variance by applying a principal component analysis (PCA) on all 50 realizations of all training supernovae. The coordinates of the leading eigenvectors that account for 95% of the variance are used as the supernova parameters, and are denoted by x. This reduces the dimensionality of the problem down to ~ 15 in anticipation that the truncated components carry little real signal that correlates with absolute magnitude. As a result of this transformation each parameter no longer represents a specific color at a specific phase, but rather the contribution to the temporal evolution of all colors from a single PCA eigenvector. Note that due to the correlated noise in the regressed light curves, PCA is not encapsulating pure supernova diversity efficiently in a minimal number of components.

As a representative example of one of the training sets, we show results from one of the runs calibrating the i band of the z=0.25 filter set. The first several PCA eigenvectors describing light-curve shapes and colors (after subtraction of the mean) are shown in Figure 4. These eigenvectors capture 54%, 15%, 9%, and 5% of the population variance.

The distribution of PCA coefficient values for the training set is shown in Figure 5. For each component *i*, the coefficients have

standard deviation σ_{x_i} . In solid (red) are the coefficients for the mean light curves and colors for each training-set supernova. Surrounding each of those points is a cloud of gray points of the coefficients from the 50 realizations; the size of the cloud is reflective of the dispersion in the Gaussian process model as well as the data quality. The \sim 50 training-set supernovae do not fill a continuous distribution in this space; there are clear outliers due to realizations of a single supernova. The distinct blobs containing multiple supernovae seen in the x(0)–x(2) plot are not significant as they are not as prominent in the three other training runs.

Mean and perturbed light curves from $1\sigma_{x_j}$ and $2\sigma_{x_j}$ changes in the x(j) parameter are shown for j=0,1,2 in Figures 6–9. The j=0 eigenvector (Figure 6) behaves as if it were due to dust: the correction is (mostly) phase independent and monotonically decreases from blue to redder bands. The j=1 eigenvector (Figure 7) generates light curves that vary in width, with supernovae with relatively skinny blueshifted-g (observer-frame central wavelength of 380 nm) light curves having relatively broader light curves in the redder bands. The g-X colors vary but the other colors do not vary. The j=2 eigenvector (Figure 8) has a particularly strong influence in the shape and colors in the near-UV band. It has the property that the curves intersect at later phases; the effect is pronounced

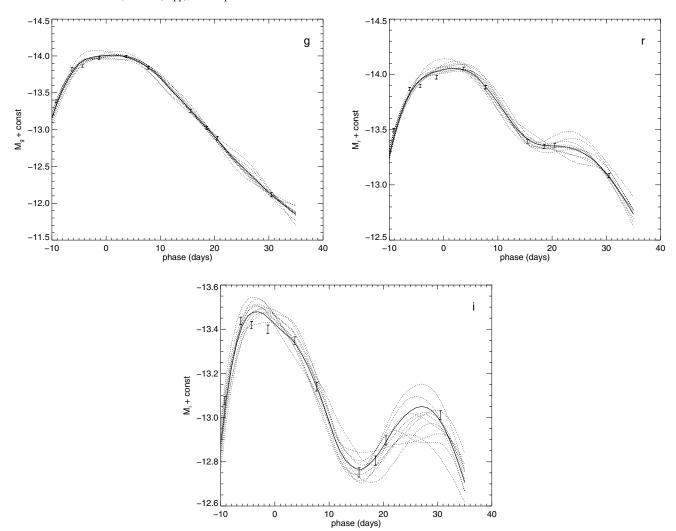


Figure 2. Light curves for SN 19 in the z=0 filter set with an arbitrary normalization offset. The predicted peak magnitudes for this object are off by \sim 0.6 mag.

in the blueshifted i band where the phase and height of the secondary maximum vary. Light-curve shapes and all color combinations also are influenced by this component. The j=3 eigenvector (Figure 9) is similar to j=2 and has a particularly exaggerated influence on the shape of the near-UV band.

The PCA coefficients are correlated with the SALT2 color and x_1 parameters. Figure 10 contains scatter plots of the first four principal components with the SALT2 color and x_1 parameters for one of the runs for the z=0 filter set: the first PCA component is strongly correlated with color and the second and fourth are correlated with x_1 .

To summarize this second step of the analysis, the mean multi-band flux and covariance of the underlying light curve of an individual supernova are calculated for a set of epochs and filters using the sparsely measured photometry combined with the trained light-curve Gaussian process model. Each realization of the light curves is decomposed into one "true" absolute magnitude and light-curve shapes and color curves. Anticipating that absolute-magnitude variations are encoded in SN Ia heterogeneity, for the light-curve shapes and colors we transform to a new coordinate system that is defined with a PCA on the training set.

2.3.3. Parameterized Model for Supernova Absolute Magnitudes

The light-curve and shape coordinates in the PCA basis are correlated with true absolute magnitude (modulo a constant

offset for the arbitrary distance at which the SNe are placed; see Section 2.2), as seen in Figure 11.

In this section, we relate light-curve and color parameters to absolute magnitudes with a linear and a Gaussian process model. The linear model makes an explicit choice for the relationship between light-curve parameters and absolute magnitude that can be well constrained by our small data set. The Gaussian process methodology lets the data drive the model. While having the advantage of not requiring an explicit model, results can be sensitive to statistical fluctuations in the training set or large interpolations in sparsely populated regions. An analogous contrast can be made between SALT (Guy et al. 2005) which used parameterized functions, and MLCS (Riess et al. 1996) which used data averages to construct light-curve models.

Linear Model. The absolute magnitude at phase t_0 and band λ_0 is a linear function of the light-curve parameters $M_{(t_0,\lambda_0)}(\mathbf{x}; p_{\bar{M}}) = \bar{M}_0 + \sum_{j=0}^{n_{\max}-1} p_j x(j)$, where \bar{M}_0 and p_j are fit parameters. Most of the over 100 supernova parameters capture little of the population variance: as is justified in Section 2.4, $n_{\max} = 4$ is adopted in this article.

Figure 12 shows the linear-model solution for the z = 0.00 filter set calibrating the g band, which closely corresponds to the bands for which SALT2 is directly trained. As in SALT2, there are clear correlations between the first two parameters, which in turn are correlated with color and x_1 , respectively

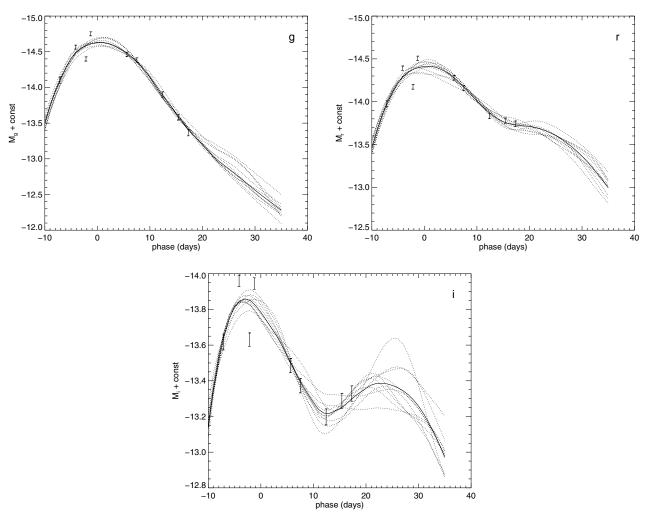


Figure 3. Light curves in the z = 0 filter set for SN 13, whose predicted peak magnitudes are off by ~ 0.3 mag.

(as shown in Figure 10). There is no correlation for the third parameter, but a correlation with the fourth parameter is apparent.

Gaussian Process Model. While previous analyses have modeled linear and quadratic correlations between light-curve parameters and absolute magnitude, it is well known that such simple relationships are not representative of the full range of light-curve shapes (Hamuy et al. 1996). Here we infer the absolute magnitude through regression in the high-dimensional PCA parameter space.

The fiducial magnitude at phase t_0 and band λ_0 , $M_{(t_0,\lambda_0)}$, is modeled as a Gaussian process dependent on light-curve shapes and colors expressed by the parameters \mathbf{x} ,

$$M_{(t_0,\lambda_0)} \sim \operatorname{GP}(\bar{M}(\mathbf{x}; \bar{M}_0), k_M(\mathbf{x}, \mathbf{x}'; \mathbf{l}_{k_M}, \sigma_{k_M}) + n_M(\mathbf{x}, \mathbf{x}'; \sigma_{n_M})).$$
(8)

The mean function $\bar{M}(\mathbf{x}; p_{\bar{M}}) = \bar{M}_0 + \sum_{j=0}^{n_{\text{max}}-1} p_j x(j)$ is linear and has free parameters \bar{M}_0 and p_j . The kernel function is again a square exponential

$$k_M(\mathbf{x}, \mathbf{x}'; \mathbf{l}_{k_M}, \sigma_{k_M}) = \sigma_{k_M}^2 \exp\left[-\sum_{j=1}^n \left(\frac{x(j) - x'(j)}{l_{k_M}(j)}\right)^2\right],$$
(9)

where n is the number of elements in \mathbf{x} , i.e., the number of eigenvectors capturing 95% of the population variance. There is an independent per-point nugget, $n_M(\mathbf{x}, \mathbf{x}'; \sigma_{n_M}) = \sigma_{n_M}^2 \delta_{\mathbf{x}, \mathbf{x}'}$.

Training-set supernovae are used to find the best-fit values for \bar{M}_0 , \mathbf{l}_{k_M} , p_j , σ_{k_M} , and σ_{n_M} . The values of σ_{k_M} , the strength of magnitude correlations, are found to be around 0.18 mag. The residual absolute magnitude dispersion after training is given by the σ_{n_M} s, which range from 0.04 to 0.07 mag except for the z=0.75 bands that have a range from 0.07 to 0.11 mag.

The hyperparameters for the length scale of the correlation \mathbf{l}_{k_M} alone do not provide much physical insight. A more interesting statistic is the ratio of the correlation scale and a typical value of the parameter in the training set; the metric represents a typical scale for each parameter in the exponent in Equation (9), which in turn gives the relative contribution to the correlation strength. Parameters with low values in Figure 13 contribute more to the correlation. Figure 13 shows the best-fit $l_{k_M}(j)/\sigma_{x_j}$, where σ_{x_j} is the standard deviation of the *j*th coefficient of all realizations in the training set of a representative run.

The relationship between each parameter and absolute magnitude is illustrated for a slice in parameter space for the same run shown for the linear model; Figure 12 shows the most-likely absolute magnitude as a function of parameter value while holding the coordinates for all other dimensions at zero. In this example, the coordinate range spanned by the supernova set decreases with the dimension order: -6 < x(0) < 4, -2 < x(1) < 2, -1.8 < x(2) < 1.8, and -1 < x(3) < 1.

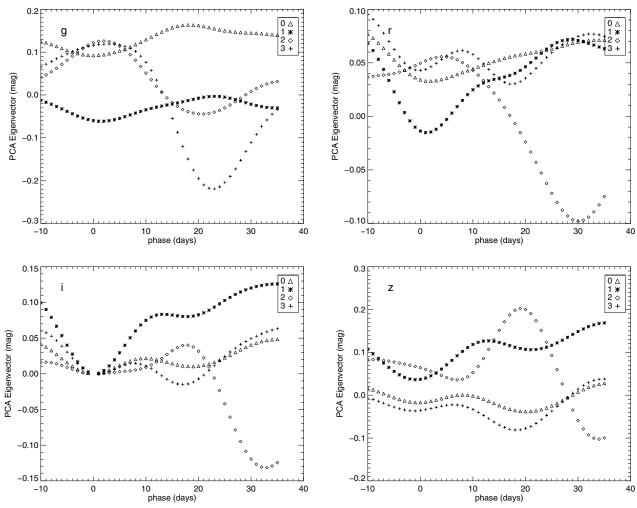


Figure 4. First four PCA eigenvectors of the z=0.25 filter set plotted as a function of corresponding phase relative to the B peak split into g (top left), r (top right), i (bottom left), and z (bottom right), from a run calibrating the i flux (hence the gap in the bottom left panel). These eigenvectors capture 54%, 15%, 9%, and 5% of the population variance.

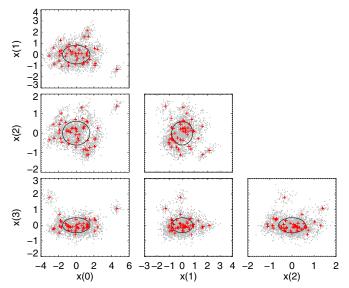


Figure 5. Distribution of the first four PCA coefficients for one of the runs calibrating the i band of the z=0.25 filter set. The predicted mean of each training-set supernova is shown in red + symbols with the 50 associated realizations shown in gray. The ellipses have semimajor and semiminor axes with the standard deviations of the distribution σ_{x_i} .

(A color version of this figure is available in the online journal.)

The correlation between parameters and absolute magnitude is nonlinear.

The Gaussian process and linear solutions for the first two parameters are in general agreement, consistent with the SALT2 model. X(2) does not have a monotonic trend in the region and extends a small range of influence on luminosity. The fourth parameter has agreement in X(3) within the range where there are training-set data.

To summarize this step of the analysis, supernova absolute magnitudes are taken to be a function of the light-curve shapes and colors. In one analysis a linear fit is performed. In the other the functional form is not explicit; instead its stochastic behavior is modeled. Training-set supernovae provide several measurements of this function and we use these to train a Gaussian process description of their correlation. The trained Gaussian process provides a prescription for interpolating an "inferred" absolute magnitude from the light-curve shape and color. It should be emphasized that the training procedure is not optimized to minimize the dispersion in the residuals of absolute magnitudes but rather to maximize the likelihood of the data for the Gaussian process model.

2.4. Validation

The validation-set supernovae are used to test how well true absolute magnitudes can be predicted from the inferred

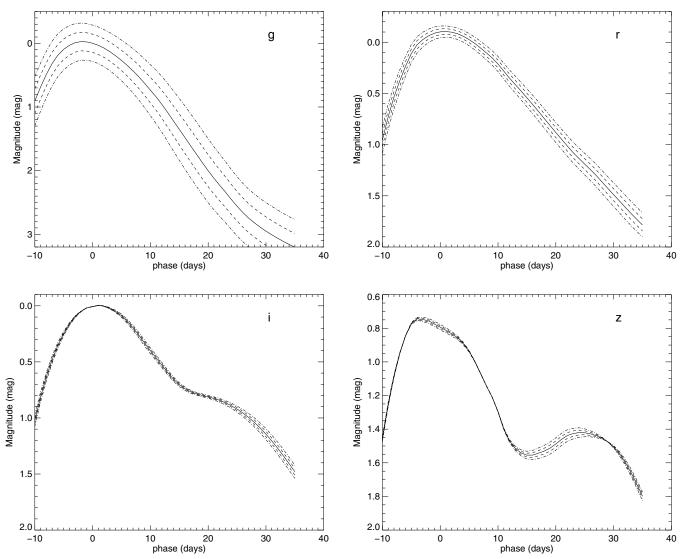


Figure 6. Mean and perturbed light curves from $1\sigma_{x_0}$ and $2\sigma_{x_0}$ changes in the x(0) parameter for the z=0.25 filter set from a run calibrating the i flux, plotted as a function of corresponding phase relative to the B peak. The four bands are plotted: g (top left), r (top right), i (bottom left), and z (bottom right).

absolute magnitudes derived from the trained model in the procedure given in Section 2.3.1. We perform cross-validation by dividing the supernovae into distinct training and validation sets with 25% assigned to the latter: the results are not expected to be as good as if all supernovae are used for training, but in this approach there is no question of the independence of the magnitude dispersion of the validation set from the data used in training. We measure the stability of our results by repeating the same analysis on validation sets containing different supernovae; as 25% of the objects enter the validation set the analysis is run four times with independent validation but partially intersecting training sets. Over four runs each supernova enters a validation set once and only once. Comparing the four almostindependent cross-validations diagnoses whether the number and fraction of training supernovae are a fair representation of the full population.

2.4.1. Pruned Supernova Sample

The SNfactory data do not have perfect temporal sampling or S/N. As a selection criterion, we study those validation supernovae with synthetic photometry in at least eight distinct phases in each of the bands. This leaves 13-23 validation supernovae per run except in the z=0.75 filter set, which

has 7–11 SNe per run. Measurement uncertainties contribute to uncertainties in the interpolated absolute magnitude and light curves. To illustrate, the distribution of light curves realized from the Gaussian-process-predicted mean and covariance shown in Figures 1–3 indicates the size of the uncertainties. Figure 15 contains a histogram of the standard deviations of the measured true absolute magnitude at B peak for the i band of the z=0.25 filter set; here the absolute-magnitude uncertainties per supernova range from 0.02 to 0.06 mag.

As a representative example of the light-curve regression, we present results for one random supernova from one of the runs with the z=0.25 filter set. Figure 1 shows the synthetic photometry, the predicted mean of the underlying light curves in solid, and 10 random realizations in dashed lines.

We identify two odd supernovae, SN 13 and SN 19, that are culled from the sample. SN 19 data and predicted light curves for the z=0 filter set are shown in Figure 2. This supernova does not have the wavelength coverage to make the z=0.25 sample and its PCA coefficients fail the distance cut for z=0.75. It is an extremely slow decliner with a g band $\Delta m_{15}=0.7$ derived from the Gaussian process prediction. The PCA coefficients from the z=0 case are shown among those of the training set in Figure 14; it is clear that SN 19 does not have a corresponding

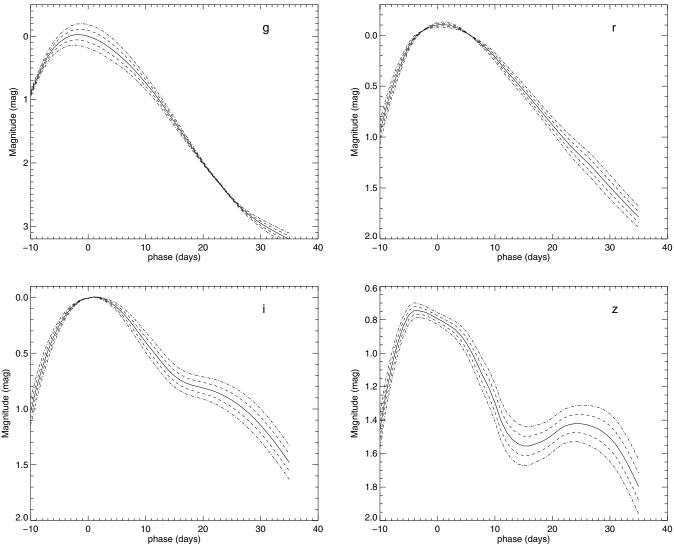


Figure 7. As in Figure 6 except for perturbations in the x(1) parameter.

analog in the training set. With z=0.0055, this supernova has the lowest cosmic microwave background-frame redshift of the full sample of which >96% has z>0.015; a peculiar velocity of 300 km s⁻¹ would contribute a 0.40 mag residual. SN 13 data and predicted light curves for the z=0.25 filter set are shown in Figure 3. The discrepancy is accounted for by one spectrum near the peak that is fainter than expected; indeed this discrepancy disappears in subsequent versions of the spectral reduction.

Validation supernovae may not have an analog in the training set. We approximate the PCA coefficient distribution as Gaussian, described by the standard deviations σ_{x_i} of each dimension. A validation supernova is included in the analysis if it lies within the core of the distribution

$$\sum_{j} \left(\frac{x_{j}}{\sigma_{x_{j}}}\right)^{2} < \chi^{2}(p, n_{\text{PCA}}), \tag{10}$$

where $\chi^2(p,n)$ is the value where the cumulative χ^2 distribution has p-value 0.05 and n_{PCA} is the number of principal components that capture 95% of the variance. This cut criterion was made after the distributions were constructed and is therefore not blind. However, 0.05 is selected as a nice round number that should capture 95% of the population and ensuing results are

insensitive to departures from this value. This cut typically removes 0–2 objects from the sample.

2.4.2. True and Inferred Magnitudes and their Differences

The covariance in the regression is accounted for in our analysis by using 50 light-curve realizations of each supernovae. Each of the 50 multi-band light-curve realizations of each supernova is temporally sampled from -10 to 35 days using the best-fit Gaussian process model of the training set. The light curves are normalized by the true $M_{(t_0,\lambda_0)}$ and expressed in the coordinate system defined by the PCA eigenvectors. The PCA coefficients are used to infer $M_{(t_0,\lambda_0)}$. The standard deviation of its 50 realizations provides the measurement uncertainty of the $M_{(t_0,\lambda_0)}$ of a supernova; this value depends on the quality of the measured light curve and ranges. The means and standard deviations of the 50 realizations of true and inferred absolute magnitudes from the validation procedure (modulo a constant offset) are given in Tables 2–5 for the Gaussian process model. (In the tables, omitted supernovae were used for lightcurve training but did not pass the selection criteria for the peak-magnitude analysis. In italics are supernovae that have data before two-day pre-maximum with light curves that have at least eight points. In bold are supernovae used for validation

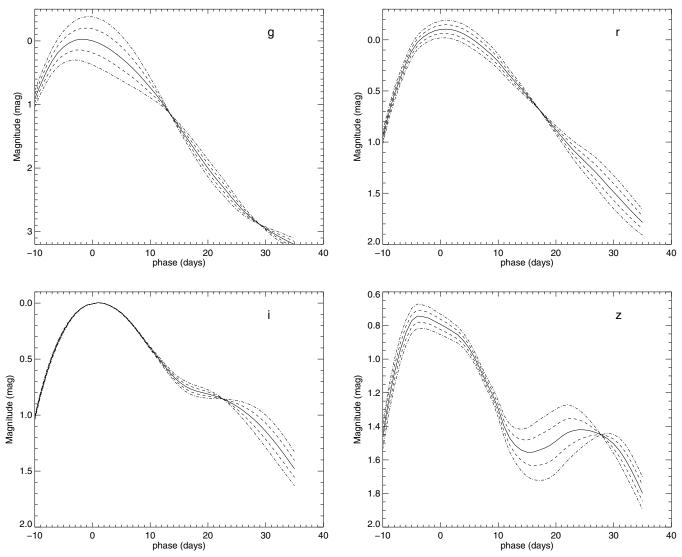


Figure 8. As in Figure 6 except for perturbations in the x(2) parameter.

that have data before the two-day pre-maximum and at least eight points in all light curves.) Figure 15 shows the histogram of the standard deviations of the inferred absolute magnitude at the B peak for the i band of the z=0.25 filter set: the inferred absolute magnitude measurement uncertainties per supernova mostly fall within the range 0.04-0.10 mag.

The accuracy to which true absolute magnitudes are predicted by absolute magnitudes inferred from SNfactory synthetic light curves is the root mean square of the difference between the true and inferred $M_{(t_0,\lambda_0)}$ over all realizations of all supernovae. The wrms of the differences of each validation set, and then the mean and standard deviation of those wrms' over the four training/validations ($\overline{\text{wrms}}$ and $\sigma(\text{wrms})$, respectively) are calculated. The wrms weights the squared residual by the inverse measurement variance visualized in Figure 15 (the expression for the wrms is given in Blondin et al. 2011). The mean shows how well the algorithm standardizes the distance indicator over all supernovae. The standard deviation of the wrms of the four runs indicates the sensitivity to the sample selection. The wrms' have contributions from measurement errors, peculiar velocities, and the intrinsic ability that the methodology has to calibrate supernovae, σ_{int} . Assuming a peculiar velocity dispersion of 300 km s⁻¹, for each run we report the mean and standard deviations of the fit value the intrinsic dispersion σ_{int} .

The wrms calculated above represents an estimate of the absolute-magnitude uncertainty when using 75% of the data training, such that none of the validation dispersions are calculated with supernovae used in the training. Our analysis lends to easy calculation of the wrms from K-fold cross analysis (as done by Blondin et al. 2011) for K = 4, where the dispersion is calculated for all supernovae together despite the fact that they are based on different training sets; these results are given in Tables 6 and 7. We provide two numbers for the K-fold analysis; the first using the subset after applying the sample cut of Equation (10) and the second using the full sample of objects. The former statistic is calculated anticipating that the supernova parameter space is not densely sampled and considering that our absolute-magnitude model may not extrapolate well to objects outside the training regime.

For the linear model, the number of light-curve parameters related to absolute magnitude is set as $n_{\rm max}=4$. This choice is made as follows: we calculate $\overline{\rm wrms}$ for the linear model, varying $n_{\rm max}$ from 1 to 8 for the *i*-band calibration of the z=0.25 filter set, obtaining 0.097, 0.099, 0.104, 0.091, 0.092, 0.094, 0.093, and 0.095 mag. There is a step in the $\overline{\rm wrms}$ when four parameters are used. In the Gaussian process model, $n_{\rm max}$ is set to the number of principal components that capture 95% of the variance.

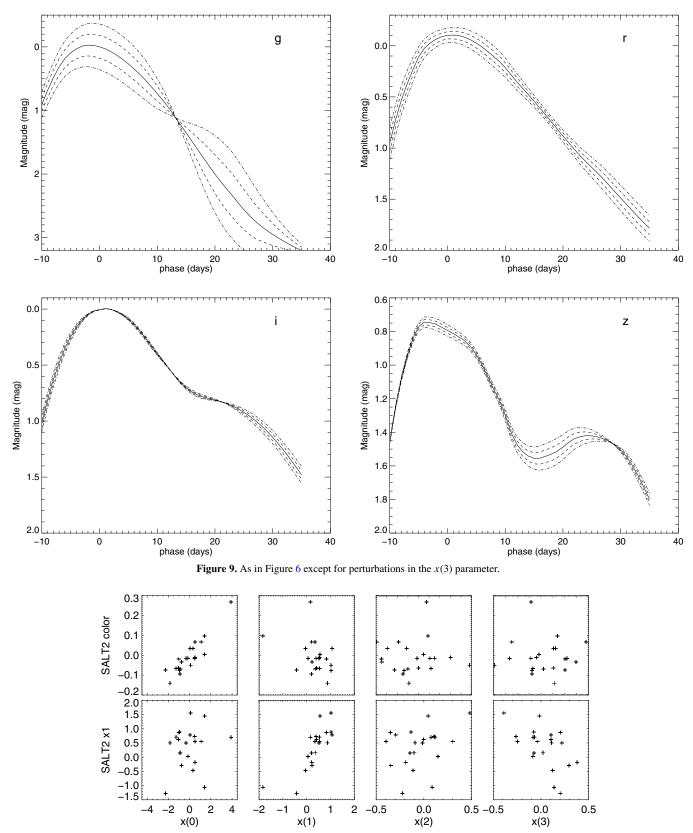


Figure 10. Scatter plots of the first four PCA coefficients, x(0), x(1), x(2), and x(3) with the SALT2 color and x_1 (light-curve shape) parameters for one of the runs for the z = 0 filter set. The correlation coefficients of color with the PCA coefficients are $\{-0.91, -0.33, 0.02, -0.03\}$ and those of x_1 with the PCA coefficients are $\{-0.18.0.80, 0.27, -0.47\}$.

Calculated dispersions from the validation are shown in Table 6 for the linear fit, and in Table 7 for the Gaussian process model. The weighted mean of each validation set is $\lesssim 0.01$ mag.

The statistics in Efron (1983) are calculated for the case of the *i* band from the z = 0.25 filter set: the *apparent error rate* is determined from the residuals of the training-set supernovae run through the same processing as the validation supernovae. The

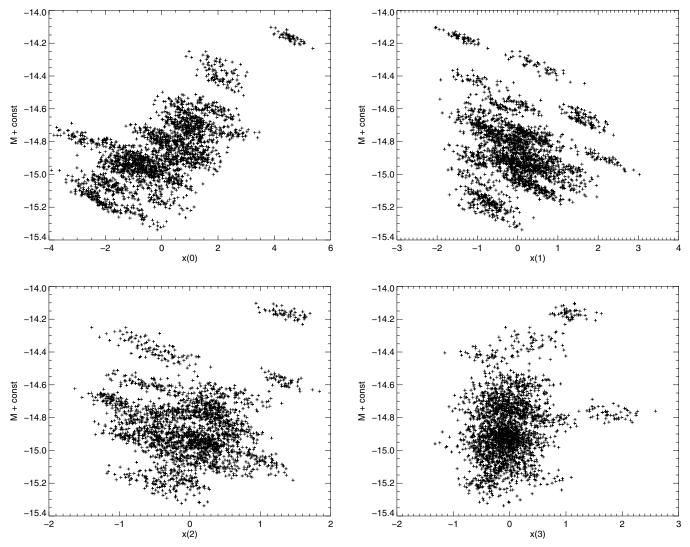


Figure 11. Training-set absolute magnitudes (with constant offset) as a function of the first four light-curve coordinates in the PCA basis for one run of the z = 0.25 filters calibrating the i band. Each point represents one supernova realization.

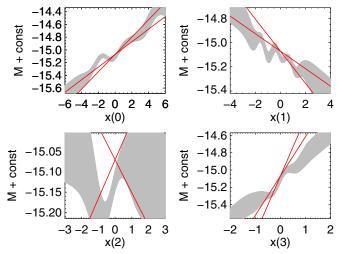


Figure 12. 1σ range of the inferred absolute magnitude found for the linear fit is shown by the red solid lines and in shaded gray are the 1σ bands from the Gaussian process model as a function of parameter value while holding the coordinates for all other dimensions at zero, for one run of the z=0.00 filters calibrating the g band. Note the different ordinate ranges of the plots.

(A color version of this figure is available in the online journal.)

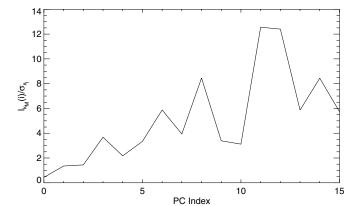


Figure 13. Characteristic correlation scale in the kernel for each parameter, $l_{k_M}(i)/\sigma_{x_i}$, where σ_{x_i} is the standard deviation of the *i*th parameter over the training set.

linear model gives 0.089 mag. The resulting 0.025 mag from the Gaussian process model is an over-optimistic error estimate: the best-fit $\sigma_{n_M} = 0.060$ mag nugget parameter (introduced after Equation (9)) gives an estimate of the absolute magnitude dispersion of the training supernovae. The *bootstrap* estimate

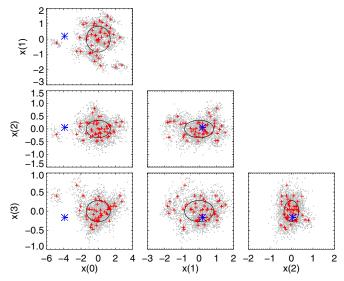


Figure 14. Distribution of the first four PCA coefficients for one of the z=0.00 filter-set runs used in the training of the analysis of SN 19. The symbols are the same as in Figure 5. In addition, blue stars show the PCA coefficients of the predicted mean of SN 19.

(A color version of this figure is available in the online journal.)

is calculated from 50 realizations of training sets generated using random draws with replacement from the full set and the \sim 37% unselected supernovae forming the validation set; this error estimate is 0.107 (0.117) mag for the linear (Gaussian process) model. The .632 estimator is not rigorously motivated but is empirically exhibiting small biases in the estimator of the true prediction error in certain cases; it is a function of the apparent and bootstrap estimators for which we get 0.100 (0.083) mag for the linear (Gaussian process) model.

3. DISCUSSION

Training and validation sets have different variances in their predicted magnitudes. The discrepancy between the two is not surprising given the sparsity of supernovae; it indicates that the current training set does not fairly represent the population of the current validation set. We anticipate that these values should converge as the number of training supernovae increases. The validation wrms may decrease with interpolation from more local values, while the training standard deviation may increase as singletons disappear.

The subsample of \sim 50 training-set supernovae used to calibrate absolute magnitudes is too small to cover the full range of intrinsic and extrinsic supernova variability. This issue is highlighted in SN 19, where highly discrepant absolute magnitudes are predicted when no analogue exists in the training set. We remark that SN 19 (IAU name SN 2007le) does have unique properties noted elsewhere (Simon et al. 2009): it is one of only three supernovae that exhibit time-varying Na I D absorption features and the only one of those that is not highly reddened, suggesting circumstellar extinction. Nevertheless, with the loose requirement of Equation (10), many other supernovae manage to be reasonably calibrated. With a larger sample size, it will be interesting to see if large outliers persist despite having training-set analogues in regions of parameter space where SNe Ia are not well standardized. The trade between criteria for deciding if a supernova has training-set analogues versus the numbers of supernovae

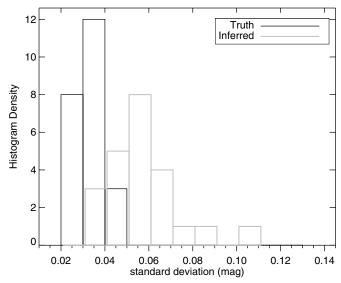


Figure 15. Histograms of uncertainties in the Gaussian process prediction variance including measurement error in the true (solid) and inferred (gray) absolute magnitudes per supernova, for a run calibrating the i band of the z=0.25 filter set.

that meet that criteria should be studied in actual cosmological analyses.

Whether the linear or Gaussian process model results in lower dispersion depends on the filter set being calibrated and the specific statistic of interest. Figure 12 shows examples of how either model can go wrong. The wiggles in the X(1) dependence of the absolute magnitude suggest that the Gaussian process regression is sensitive to fluctuations in the small training set. The Gaussian process regression for X(2) is clearly not well described by the linear model. While both linear and Gaussian process solutions for X(3) agree in the region covered by the training sample, beyond that they extrapolate out with different slopes. For the data used, neither model for absolute magnitude universally outperforms the other.

We get dispersions of 0.09-0.14 mag (linear fit) and 0.09–0.11 mag (Gaussian process model) using the four bands of the z = 0.25 filter set. Filter sets that use only three bands exhibit higher dispersion, although only modestly so. Although the input magnitude dispersion strongly depends on the band of the magnitude being calibrated, this dependency mostly disappears after calibration in the Gaussian process fit. Some of the dispersion is due to limitations in the data themselves. There are light curves with suspected misestimated fluxes or uncertainties, SN 13 being a prime example (Figure 3). Sometimes there are significant temporal interpolations being made due to coarse light-curve sampling, particularly over the secondary maximum in the redder bands as seen in Figures 1–3. A further decrease in the dispersion may be possible as more supernova spectrophotometry with higher S/N and denser temporal sampling becomes available.

In an attempt to avoid an explicit description of the supernova light-curve model, we use the relatively simple normalized Hsiao templates for the mean function. Certainly, Gaussian processes can be used to describe residuals about the SALT2 model whose shape and color parameters (including date of maximum) are included as variables in the absolute-magnitude model. Inclusion of a dust-absorption model in the mean function should reduce the number of training-set supernovae needed to densely sample the full range of dust absorption.

SN	8	3			i		
	True	Inferred	True	Inferred	True	Inferred	
1	-15.23 ± 0.04	-15.22 ± 0.06	-14.91 ± 0.05	-14.90 ± 0.05	-14.19 ± 0.05	-14.19 ± 0.0	
	-15.12 ± 0.03	-15.16 ± 0.06	-14.90 ± 0.03	-14.93 ± 0.05	-14.18 ± 0.04	-14.18 ± 0.0	
	-15.25 ± 0.03	-15.23 ± 0.07	-14.94 ± 0.03	-14.92 ± 0.06	-14.25 ± 0.04	-14.24 ± 0.0	
	-15.08 ± 0.03	-14.76 ± 0.07	-14.83 ± 0.04	-14.53 ± 0.08	-14.17 ± 0.04	-13.86 ± 0.0	
	-15.29 ± 0.03	-15.20 ± 0.13	-14.99 ± 0.03	-14.88 ± 0.10	-14.32 ± 0.05	-14.22 ± 0.1	
	-15.18 ± 0.03 -15.22 ± 0.04	-15.27 ± 0.08 -15.06 ± 0.05	-14.90 ± 0.03 -14.90 ± 0.04	-14.97 ± 0.07 -14.70 ± 0.02	-14.14 ± 0.03 -14.25 ± 0.06	-14.17 ± 0.0 -14.11 ± 0.0	
	-15.22 ± 0.04 -15.16 ± 0.02	-15.00 ± 0.03 -15.23 ± 0.08	-14.84 ± 0.03	-14.70 ± 0.02 -14.93 ± 0.07	-14.23 ± 0.00 -14.13 ± 0.04	-14.11 ± 0.0 -14.19 ± 0.0	
0	-15.29 ± 0.03	-15.17 ± 0.08	-15.00 ± 0.03	-14.90 ± 0.05	-14.39 ± 0.04	-14.25 ± 0.0	
2	-15.22 ± 0.03	-15.09 ± 0.07	-14.94 ± 0.03	-14.86 ± 0.08	-14.29 ± 0.04	-14.17 ± 0.0	
3	-14.63 ± 0.04	-14.97 ± 0.08	-14.41 ± 0.04	-14.76 ± 0.05	-13.78 ± 0.05	-14.12 ± 0.0	
1	-14.91 ± 0.02	-14.96 ± 0.11	-14.74 ± 0.02	-14.75 ± 0.08	-14.14 ± 0.03	-14.18 ± 0.0	
5	-15.15 ± 0.03	-15.11 ± 0.09	-14.88 ± 0.03	-14.82 ± 0.07	-14.23 ± 0.04	-14.19 ± 0.0	
7	-15.33 ± 0.04	-15.12 ± 0.08	-15.00 ± 0.04	-14.87 ± 0.08	-14.39 ± 0.04	-14.20 ± 0.0	
9	-14.01 ± 0.04	-14.55 ± 0.09	-14.05 ± 0.04	-14.60 ± 0.05	-13.42 ± 0.05	-13.98 ± 0.0	
1 2	-14.61 ± 0.02 -15.43 ± 0.03	-14.70 ± 0.05 -15.21 ± 0.09	-14.51 ± 0.03 -15.15 ± 0.03	-14.63 ± 0.04 -14.93 ± 0.08	-14.00 ± 0.04 -14.45 ± 0.04	-14.13 ± 0.0 -14.21 ± 0.1	
1	-13.43 ± 0.03 -14.79 ± 0.03	-13.21 ± 0.09 -14.86 ± 0.07	-13.13 ± 0.03 -14.66 ± 0.03	-14.73 ± 0.06 -14.73 ± 0.06	-14.43 ± 0.04 -14.02 ± 0.03	-14.21 ± 0.1 -14.10 ± 0.0	
7	-15.42 ± 0.05	-15.47 ± 0.05	-15.05 ± 0.05	-15.09 ± 0.03	-14.34 ± 0.06	-14.38 ± 0.0	
)	-15.10 ± 0.02	-15.14 ± 0.07	-14.86 ± 0.03	-14.95 ± 0.05	-14.30 ± 0.04	-14.38 ± 0.0	
3	-15.14 ± 0.03	-15.14 ± 0.11	-14.87 ± 0.03	-14.84 ± 0.09	-14.21 ± 0.04	-14.18 ± 0.1	
ó	-15.30 ± 0.03	-14.68 ± 0.06	-15.16 ± 0.04	-14.58 ± 0.04	-14.94 ± 0.05	-14.30 ± 0.0	
3	-15.41 ± 0.03	-15.36 ± 0.06	-15.04 ± 0.03	-14.98 ± 0.06	-14.37 ± 0.05	-14.32 ± 0.0	
)	-15.34 ± 0.03	-15.34 ± 0.05	-14.94 ± 0.03	-14.97 ± 0.04	-14.30 ± 0.04	-14.30 ± 0.0	
3	-15.34 ± 0.04	-15.31 ± 0.07	-15.03 ± 0.05	-14.99 ± 0.05	-14.45 ± 0.06	-14.40 ± 0.0	
,	-15.47 ± 0.06	-15.34 ± 0.09	-15.10 ± 0.07	-14.97 ± 0.07	-14.37 ± 0.09	$-14.22 \pm 0.$	
,	-15.02 ± 0.03 -15.07 ± 0.03	-15.09 ± 0.06 -15.22 ± 0.09	-14.85 ± 0.03 -14.77 ± 0.03	$-14.89 \pm 0.06 \ -14.91 \pm 0.07$	$-14.21 \pm 0.04 \\ -14.14 \pm 0.04$	-14.29 ± 0.0 -14.26 ± 0.0	
	-15.07 ± 0.03 -15.37 ± 0.05	-15.22 ± 0.09 -15.21 ± 0.08	-14.77 ± 0.03 -15.07 ± 0.04	-14.91 ± 0.07 -14.92 ± 0.06	-14.14 ± 0.04 -14.22 ± 0.05	-14.20 ± 0.0 -14.06 ± 0.0	
	-15.05 ± 0.03	-15.21 ± 0.00 -15.21 ± 0.07	-14.76 ± 0.03	-14.94 ± 0.05	-14.13 ± 0.05	-14.31 ± 0.0	
,	-15.41 ± 0.03	-15.20 ± 0.05	-15.10 ± 0.03	-14.95 ± 0.04	-14.46 ± 0.04	-14.31 ± 0.0	
)	-15.10 ± 0.03	-15.11 ± 0.07	-14.88 ± 0.03	-14.87 ± 0.07	-14.24 ± 0.04	-14.24 ± 0.0	
!	-15.25 ± 0.03	-15.29 ± 0.07	-14.92 ± 0.03	-14.96 ± 0.05	-14.38 ± 0.05	-14.41 ± 0.0	
,	-14.95 ± 0.04	-14.92 ± 0.08	-14.76 ± 0.03	-14.76 ± 0.07	-14.16 ± 0.04	-14.14 ± 0.0	
	-14.87 ± 0.03	-14.93 ± 0.08	-14.69 ± 0.03	-14.74 ± 0.05	-13.96 ± 0.04	-14.05 ± 0.0	
)	-15.52 ± 0.04	-15.30 ± 0.09	-15.23 ± 0.05	-15.04 ± 0.08	-14.64 ± 0.05	-14.45 ± 0.1	
)	-15.09 ± 0.04 -14.85 ± 0.03	-15.20 ± 0.08 -15.12 ± 0.09	-14.84 ± 0.05 -14.61 ± 0.04	-14.96 ± 0.10 -14.77 ± 0.08	-14.16 ± 0.05 -13.85 ± 0.04	-14.27 ± 0.0 -14.11 ± 0.0	
	-15.02 ± 0.03	-13.12 ± 0.09 -14.91 ± 0.10	-14.84 ± 0.03	-14.77 ± 0.06 -14.73 ± 0.06	-13.03 ± 0.04 -14.12 ± 0.04	-14.02 ± 0.0	
	-15.24 ± 0.03	-15.09 ± 0.09	-14.99 ± 0.04	-14.86 ± 0.07	-14.38 ± 0.04	-14.21 ± 0.1	
	-14.74 ± 0.03	-14.79 ± 0.06	-14.56 ± 0.03	-14.61 ± 0.04	-13.96 ± 0.04	-14.02 ± 0.0	
)	-15.13 ± 0.04	-15.19 ± 0.06	-14.90 ± 0.06	-14.94 ± 0.07	-14.18 ± 0.08	-14.22 ± 0.0	
)	-14.82 ± 0.04	-14.99 ± 0.11	-14.64 ± 0.04	-14.82 ± 0.07	-13.94 ± 0.05	$-14.14 \pm 0.$	
	-15.15 ± 0.02	-15.07 ± 0.09	-14.96 ± 0.03	-14.89 ± 0.07	-14.32 ± 0.03	-14.23 ± 0.0	
	-15.20 ± 0.03	-15.20 ± 0.09	-14.90 ± 0.02	-14.90 ± 0.09	-14.23 ± 0.04	-14.20 ± 0.1	
ļ ;	-15.03 ± 0.03 -14.97 ± 0.02	-14.85 ± 0.06 -15.06 ± 0.04	-14.80 ± 0.03 -14.66 ± 0.03	-14.63 ± 0.04 -14.74 ± 0.04	-14.21 ± 0.03 -14.03 ± 0.03	-14.04 ± 0.0 -14.12 ± 0.0	
;	-14.97 ± 0.02 -15.44 ± 0.03	-15.00 ± 0.04 -15.23 ± 0.06	-14.00 ± 0.03 -15.16 ± 0.03	-14.74 ± 0.04 -14.96 ± 0.07	-14.03 ± 0.03 -14.56 ± 0.04	-14.12 ± 0.0 -14.38 ± 0.0	
,	-15.40 ± 0.03	-15.31 ± 0.06	-15.07 ± 0.03	-15.01 ± 0.04	-14.37 ± 0.04	-14.26 ± 0.0	
	-15.00 ± 0.03	-15.21 ± 0.08	-14.74 ± 0.03	-14.95 ± 0.07	-14.04 ± 0.03	-14.21 ± 0.0	
	-15.01 ± 0.04	-15.18 ± 0.07	-14.75 ± 0.04	-14.90 ± 0.05	-14.08 ± 0.05	$-14.22 \pm 0.$	
	-15.02 ± 0.03	-15.20 ± 0.05	-14.71 ± 0.03	-14.94 ± 0.05	-14.01 ± 0.04	$-14.23 \pm 0.$	
	-15.32 ± 0.03	-15.34 ± 0.04	-15.03 ± 0.03	-15.05 ± 0.03	-14.30 ± 0.03	$-14.31 \pm 0.$	
	-14.91 ± 0.03	-15.13 ± 0.08	-14.67 ± 0.04	-14.88 ± 0.06	-13.87 ± 0.04	$-14.08 \pm 0.$	
	-14.75 ± 0.03	-14.85 ± 0.07	-14.63 ± 0.02	-14.71 ± 0.05	-13.85 ± 0.04	$-13.97 \pm 0.$	
	-14.59 ± 0.03	-14.94 ± 0.07	-14.55 ± 0.03	-14.85 ± 0.04	-13.99 ± 0.03	$-14.38 \pm 0.$	
	-13.96 ± 0.03 -15.01 ± 0.03	-14.21 ± 0.03 -15.09 ± 0.05	-14.16 ± 0.03 -14.77 ± 0.03	-14.34 ± 0.03 -14.84 ± 0.04	-13.63 ± 0.03 -14.07 ± 0.03	$-13.89 \pm 0.$ -14.12 \pm 0.	
0	-15.01 ± 0.03 -14.79 ± 0.03	-15.09 ± 0.05 -14.76 ± 0.09	-14.77 ± 0.03 -14.69 ± 0.03	-14.84 ± 0.04 -14.64 ± 0.07	-14.07 ± 0.03 -14.08 ± 0.04	-14.12 ± 0.0 -14.07 ± 0.0	
3	-14.79 ± 0.03 -14.92 ± 0.03	-14.70 ± 0.09 -14.93 ± 0.09	-14.69 ± 0.03 -14.69 ± 0.04	-14.71 ± 0.07	-14.03 ± 0.04 -14.15 ± 0.04	$-14.07 \pm 0.$ $-14.13 \pm 0.$	
5	-15.12 ± 0.03	-15.06 ± 0.10	-14.91 ± 0.03	-14.84 ± 0.10	-14.23 ± 0.04	$-14.18 \pm 0.$	
)7	-15.15 ± 0.03	-14.88 ± 0.05	-14.89 ± 0.03	-14.66 ± 0.06	-14.27 ± 0.04	-14.00 ± 0.0	
8	-15.32 ± 0.05	-15.26 ± 0.08	-15.00 ± 0.06	-14.93 ± 0.06	-14.31 ± 0.06	-14.26 ± 0.0	
)9	-15.18 ± 0.03	-15.18 ± 0.09	-14.89 ± 0.03	-14.93 ± 0.08	-14.29 ± 0.04	-14.33 ± 0.1	
11	-15.08 ± 0.02	-15.13 ± 0.11	-14.76 ± 0.03	-14.85 ± 0.08	-14.04 ± 0.04	-14.08 ± 0.0	

Table 2 (Continued)

SN	8	g		r		i	
	True	Inferred	True	Inferred	True	Inferred	
112	-14.71 ± 0.04	-14.66 ± 0.08	-14.61 ± 0.04	-14.59 ± 0.05	-14.10 ± 0.06	-14.09 ± 0.05	
113	-14.77 ± 0.03	-14.87 ± 0.07	-14.65 ± 0.03	-14.73 ± 0.06	-13.97 ± 0.04	-14.07 ± 0.07	
114	-14.86 ± 0.03	-14.95 ± 0.09	-14.59 ± 0.04	-14.70 ± 0.07	-13.92 ± 0.05	-14.10 ± 0.12	
115	-15.52 ± 0.03	-15.30 ± 0.11	-15.21 ± 0.04	-15.02 ± 0.09	-14.65 ± 0.04	-14.38 ± 0.11	
116	-14.53 ± 0.05	-14.93 ± 0.07	-14.33 ± 0.05	-14.71 ± 0.04	-13.81 ± 0.06	-14.20 ± 0.09	
117	-15.26 ± 0.04	-15.21 ± 0.07	-14.98 ± 0.04	-14.93 ± 0.07	-14.21 ± 0.06	-14.12 ± 0.11	
118	-15.32 ± 0.04	-15.19 ± 0.09	-15.04 ± 0.05	-14.93 ± 0.07	-14.38 ± 0.05	-14.25 ± 0.12	

Notes. In italics are supernovae that have data before two-day pre-maximum with light curves that have at least eight points. In bold are supernovae that have data before two-day pre-maximum and at least eight points in all light curves.

The new expanded set of light-curve parameters can standardize SNe Ia to a level that appears competitive with other fitters. A direct comparison with the quoted dispersions of 0.12 mag for SALT2 and 0.14 mag for MLCS2k2 is difficult because there is no sharp delineation between training and validation supernovae reported for these fitters. A SALT2 analysis on synthetic photometry from this same data set yields a 0.147 mag residual on the Hubble diagram. An alternative reference is the 0.18–0.22 mag scatter found using these fitters on the CfA3 sample (Hicken et al. 2009). Assuming the dust parameter $R_V = 3.1$, Mandel et al. (2011) calculate an apparent error rate of 0.10 mag, a bootstrap error of 0.15 mag, and a .632 estimator of 0.13 mag based on optical-NIR light curves; for a more flexible model where the R_V distribution differs in different A_V intervals they obtain a .632 estimator of 0.11 mag. Folatelli et al. (2010) calibrate absolute magnitudes in the redder optical bands r and i with training-set errors of 0.13 and 0.15 mag rms scatter and inferred intrinsic dispersions of 0.10 and 0.12 mag, respectively.

We compare the results of our methodology using the *g*-band calibration of the z=0, gri filter set, with SALT2 in wavelength regions that correspond roughly to Johnson-Cousins B, V, and R. The combined subset dispersions and for the same set of supernovae are calculated to allow direct comparison with the preceding analysis. The results are $\overline{\text{wrms}} = 0.132 \pm 0.008$ mag and four-fold cross-validation dispersion of 0.132 mag; the SALT2 uncertainties in the distance are large and result in an intrinsic dispersion consistent with zero. The linear fit performs somewhat worse in reducing the total dispersions relative to this second-generation light-curve analyzer, although the Gaussian process method compares favorably.

Most of the leverage in absolute magnitude comes from the first four parameters, as shown in Figure 13. This contrasts with MLCS2k2 and SALT2, which describe supernovae with one free light-curve shape and one free color parameter. Unlike those light-curve fitters, many parameters are readily available in our analysis to account for color variance due to physical mechanisms within the supernova explosion and different distributions and properties of absorbing dust.

The intrinsic dispersion is dependent on the blueshift of the filter set, a reasonable result considering that fluxes from different rest-frame wavelength regions are measured. Note that Jha et al. (2007, MLCS2k2) and Guy et al. (2007, SALT2) each provide one redshift-independent intrinsic dispersion. As an example, the z band of the z=0.5 filter system has a dispersion of 0.114 mag and the i band of the z=0.75 filter system is 0.131 mag. Given that the griz filters are

not logarithmically distributed it is difficult to make direct inferences as to the exact cause of the different dispersions: whether local wavelength regions or the full wavelength range is playing the more important role in determining absolute magnitudes. It would be of interest to find correlations between the PCA coefficients of the z=0.5 analysis with the excess residuals of the z=0.75 analysis; this would mean that systematic errors that depend on redshift are incurred due to the limited resolution and wavelength coverage of a fixed filter system. No such correlations are apparent with the available statistics.

In the Gaussian process analysis, we restrict ourselves to using as parameters the coefficients of the leading principal components that account for 95% of the dispersion. The reduction of the dimensionality of the parameter space was motivated by computational tractability and anticipating that it is overkill to use >100 parameters to deduce absolute magnitudes. Indeed, our results do not change appreciably by using the ~30 parameters that describe 99% of the dispersion. Figure 13 hints that even our restricted parameter set is more than needed for this data set. This validates our non-trivial expectation that the largest variances in light-curve shape and color be correlated with absolute magnitude. Further compression of information may be possible; in our methodology the PCA also accounts for the covariance in the light-curve prediction. This may be avoided by accounting for measurement noise while performing the PCA, which is possible using expectation maximization techniques (Bailey 2012). Other projections such as those from independent component analysis and kernel PCA may prove more efficient in encapsulating the information, which could then be used in a Gaussian process.

For a given supernova, we calculate the inferred peak absolute magnitude for each of the observer bands. These magnitudes and their residuals from the truth are highly correlated and so their resulting inferred distances are not independent. The same would be true if we had calibrated absolute magnitude at phases other than the *B* maximum. The methodology presented optimizes the determination of the absolute magnitude in a single band and phase.

In actual cosmology analysis, a customized light-curve model can be made for each redshift or on a grid sufficiently fine to make interpolation errors unimportant. For all redshifts, the distance moduli (difference between observed peak magnitude and inferred peak absolute magnitude) are based on a common set of supernovae and can be directly compared, at least to the same extent being done with current light-curve fitters.

SN		g		,		i		Ζ
511	True	Inferred	True	Inferred	True	Inferred	True	Inferred
0	-15.07 ± 0.07	-15.07 ± 0.09	-15.08 ± 0.04	-15.11 ± 0.07	-14.94 ± 0.04	-14.95 ± 0.05	-14.08 ± 0.05	-14.09 ± 0.06
1	-14.88 ± 0.05	-14.85 ± 0.09	-15.01 ± 0.02	-14.94 ± 0.06	-14.92 ± 0.03	-14.90 ± 0.05	-14.05 ± 0.03	-14.05 ± 0.05
3	-15.11 ± 0.05	-15.16 ± 0.08	-15.12 ± 0.02	-15.17 ± 0.03	-14.98 ± 0.02	-15.05 ± 0.05	-14.11 ± 0.03	-14.18 ± 0.07
4	-14.82 ± 0.07	-14.67 ± 0.13	-14.96 ± 0.03	-14.84 ± 0.09	-14.85 ± 0.03	-14.70 ± 0.09	-14.07 ± 0.04	-13.96 ± 0.09
7	-15.05 ± 0.07	-15.13 ± 0.11	-15.04 ± 0.03	-15.19 ± 0.07	-14.93 ± 0.02	-15.05 ± 0.06	-14.02 ± 0.03	-14.11 ± 0.07
8 9	-15.02 ± 0.08 -15.02 ± 0.05	-14.80 ± 0.15 -15.10 ± 0.08	-15.05 ± 0.04 -15.02 ± 0.02	-14.88 ± 0.06 -15.10 ± 0.07	-14.91 ± 0.04 -14.87 ± 0.03	-14.69 ± 0.04 -14.96 ± 0.05	-14.14 ± 0.04 -14.00 ± 0.03	-13.95 ± 0.05 -14.07 ± 0.04
10	-15.02 ± 0.03 -15.18 ± 0.07	-15.10 ± 0.08 -15.11 ± 0.10	-15.02 ± 0.02 -15.19 ± 0.03	-15.10 ± 0.07 -15.13 ± 0.06	-14.87 ± 0.03 -15.03 ± 0.03	-14.96 ± 0.05 -14.96 ± 0.05	-14.00 ± 0.03 -14.21 ± 0.04	-14.07 ± 0.04 -14.14 ± 0.06
11	-15.05 ± 0.10	-14.97 ± 0.15	-15.12 ± 0.04	-15.03 ± 0.07	-14.99 ± 0.03	-14.92 ± 0.05	-14.12 ± 0.05	-14.01 ± 0.07
12	-15.05 ± 0.07	-15.02 ± 0.13	-15.10 ± 0.03	-15.09 ± 0.15	-14.97 ± 0.02	-14.90 ± 0.11	-14.13 ± 0.03	-14.09 ± 0.13
13	-14.29 ± 0.06	-14.55 ± 0.09	-14.53 ± 0.04	-14.77 ± 0.06	-14.43 ± 0.04	-14.70 ± 0.05	-13.70 ± 0.04	-13.95 ± 0.05
14	-14.62 ± 0.03	-14.60 ± 0.10	-14.80 ± 0.02	-14.76 ± 0.08	-14.75 ± 0.02	-14.69 ± 0.07	-14.01 ± 0.02	-14.00 ± 0.08
16	-14.97 ± 0.07	-14.86 ± 0.12	-15.01 ± 0.04	-14.84 ± 0.09	-14.92 ± 0.03	-14.78 ± 0.08	-14.10 ± 0.04	-14.00 ± 0.09
17	-15.06 ± 0.07	-14.86 ± 0.09	-15.23 ± 0.04	-14.96 ± 0.05	-15.01 ± 0.03	-14.90 ± 0.04	-14.24 ± 0.05	-14.10 ± 0.05
21 22	-14.24 ± 0.06	-14.24 ± 0.08	-14.55 ± 0.02	-14.56 ± 0.04	-14.53 ± 0.03	-14.53 ± 0.04	-13.85 ± 0.04	-13.84 ± 0.03
24	-15.38 ± 0.06 -14.50 ± 0.07	-15.20 ± 0.09 -14.56 ± 0.11	-15.30 ± 0.03 -14.71 ± 0.02	-15.10 ± 0.06 -14.77 ± 0.06	-15.20 ± 0.03 -14.68 ± 0.03	-14.97 ± 0.06 -14.74 ± 0.06	-14.34 ± 0.04 -13.90 ± 0.04	-14.15 ± 0.07 -13.95 ± 0.07
25	-15.05 ± 0.10	-15.03 ± 0.13	-15.12 ± 0.04	-15.10 ± 0.08	-14.99 ± 0.05	-14.99 ± 0.07	-14.11 ± 0.07	-14.09 ± 0.08
27	-15.37 ± 0.07	-15.45 ± 0.10	-15.28 ± 0.04	-15.34 ± 0.05	-15.09 ± 0.04	-15.16 ± 0.03	-14.15 ± 0.06	-14.23 ± 0.05
30	-14.90 ± 0.06	-14.91 ± 0.11	-14.99 ± 0.02	-14.97 ± 0.07	-14.88 ± 0.03	-14.88 ± 0.05	-14.11 ± 0.04	-14.05 ± 0.06
33	-15.00 ± 0.08	-15.13 ± 0.12	-15.03 ± 0.03	-15.13 ± 0.08	-14.91 ± 0.03	-15.05 ± 0.06	-14.07 ± 0.03	-14.19 ± 0.07
36	-15.21 ± 0.07	-14.80 ± 0.12	-15.19 ± 0.03	-14.75 ± 0.05	-15.15 ± 0.04	-14.75 ± 0.04	-14.62 ± 0.04	-14.06 ± 0.04
38	-15.29 ± 0.06	-15.26 ± 0.08	-15.25 ± 0.03	-15.22 ± 0.05	-15.07 ± 0.03	-15.05 ± 0.05	-14.22 ± 0.04	-14.19 ± 0.06
39	-15.21 ± 0.06	-15.26 ± 0.10	-15.18 ± 0.03	-15.23 ± 0.05	-14.97 ± 0.02	-15.09 ± 0.04	-14.17 ± 0.04	-14.25 ± 0.05
43 45	-15.25 ± 0.08 -15.37 ± 0.11	-15.26 ± 0.13 -15.30 ± 0.16	-15.20 ± 0.04 -15.32 ± 0.06	-15.25 ± 0.06 -15.24 ± 0.07	-15.07 ± 0.05 -15.17 ± 0.05	-15.10 ± 0.06 -15.07 ± 0.04	-14.26 ± 0.06 -14.25 ± 0.07	-14.30 ± 0.06 -14.17 ± 0.07
43 47	-13.57 ± 0.11 -14.70 ± 0.07	-13.50 ± 0.10 -14.66 ± 0.08	-13.32 ± 0.00 -14.94 ± 0.03	-13.24 ± 0.07 -14.91 ± 0.05	-13.17 ± 0.03 -14.87 ± 0.03	-13.07 ± 0.04 -14.84 ± 0.05	-14.23 ± 0.07 -14.07 ± 0.03	-14.07 ± 0.07 -14.04 ± 0.04
49	-14.82 ± 0.05	-14.87 ± 0.13	-14.94 ± 0.03	-14.99 ± 0.07	-14.78 ± 0.04	-14.87 ± 0.05	-14.02 ± 0.04	-14.09 ± 0.08
54	-15.20 ± 0.07	-15.09 ± 0.11	-15.24 ± 0.03	-15.15 ± 0.06	-15.12 ± 0.04	-14.99 ± 0.06	-14.14 ± 0.06	-14.02 ± 0.07
55	-14.83 ± 0.06	-14.96 ± 0.10	-14.91 ± 0.03	-15.02 ± 0.10	-14.79 ± 0.02	-14.92 ± 0.08	-13.99 ± 0.04	-14.11 ± 0.08
57	-15.26 ± 0.04	-15.05 ± 0.06	-15.29 ± 0.02	-15.07 ± 0.03	-15.12 ± 0.02	-14.95 ± 0.03	-14.30 ± 0.03	-14.06 ± 0.06
59	-14.96 ± 0.06	-14.95 ± 0.12	-14.98 ± 0.03	-14.98 ± 0.07	-14.90 ± 0.03	-14.90 ± 0.07	-14.11 ± 0.03	-14.09 ± 0.07
62	-15.12 ± 0.06	-15.08 ± 0.08	-15.10 ± 0.03	-15.09 ± 0.05	-14.94 ± 0.03	-14.96 ± 0.04	-14.19 ± 0.04	-14.17 ± 0.05
63 66	-14.69 ± 0.05 -14.60 ± 0.07	-14.63 ± 0.09 -14.60 ± 0.10	-14.84 ± 0.03 -14.77 ± 0.03	$-14.77 \pm 0.06 \ -14.77 \pm 0.06$	-14.77 ± 0.03 -14.72 ± 0.03	-14.73 ± 0.06 -14.71 ± 0.05	-14.06 ± 0.04 -13.90 ± 0.04	-13.97 ± 0.06 -13.91 ± 0.04
67	-15.40 ± 0.07 -15.40 ± 0.08	-15.25 ± 0.10	-15.42 ± 0.03	-15.19 ± 0.00	-15.26 ± 0.03	-15.06 ± 0.05	-13.90 ± 0.04 -14.43 ± 0.06	-13.91 ± 0.04 -14.21 ± 0.09
69	-14.86 ± 0.09	-14.86 ± 0.13	-14.98 ± 0.03	-14.93 ± 0.07	-14.88 ± 0.05	-14.86 ± 0.07	-14.04 ± 0.04	-14.05 ± 0.06
70	-14.63 ± 0.07	-14.76 ± 0.08	-14.73 ± 0.04	-14.90 ± 0.07	-14.63 ± 0.03	-14.75 ± 0.07	-13.76 ± 0.05	-13.95 ± 0.05
71	-14.70 ± 0.06	-14.62 ± 0.07	-14.92 ± 0.03	-14.86 ± 0.05	-14.87 ± 0.03	-14.80 ± 0.04	-14.05 ± 0.04	-13.98 ± 0.06
73	-15.03 ± 0.06	-14.90 ± 0.14	-15.13 ± 0.03	-14.93 ± 0.08	-15.01 ± 0.03	-14.86 ± 0.05	-14.18 ± 0.04	-14.00 ± 0.09
74	-14.48 ± 0.07	-14.58 ± 0.12	-14.63 ± 0.02	-14.75 ± 0.05	-14.59 ± 0.03	-14.69 ± 0.05	-13.85 ± 0.03	-13.98 ± 0.05
76 79	-14.94 ± 0.09	-14.99 ± 0.10 -14.59 ± 0.09	-15.01 ± 0.06 -14.71 ± 0.04	-15.07 ± 0.08 -14.81 ± 0.06	-14.92 ± 0.05 -14.68 ± 0.04	-14.98 ± 0.06 -14.76 ± 0.05	-14.09 ± 0.08 -13.88 ± 0.04	-14.13 ± 0.08 -13.97 ± 0.07
81	-14.49 ± 0.07 -14.88 ± 0.04	-14.39 ± 0.09 -14.71 ± 0.08	-14.71 ± 0.04 -15.06 ± 0.02	-14.81 ± 0.00 -14.88 ± 0.06	-14.03 ± 0.04 -15.00 ± 0.02	-14.78 ± 0.06 -14.78 ± 0.06	-13.88 ± 0.04 -14.16 ± 0.03	-13.97 ± 0.07 -14.00 ± 0.07
82	-15.00 ± 0.04	-15.04 ± 0.08	-15.07 ± 0.02	-15.13 ± 0.05	-14.92 ± 0.04	-15.00 ± 0.04	-14.10 ± 0.03	-14.15 ± 0.06
84	-14.87 ± 0.06	-14.83 ± 0.10	-14.91 ± 0.03	-14.89 ± 0.05	-14.82 ± 0.03	-14.71 ± 0.06	-14.07 ± 0.03	-14.06 ± 0.04
85	-14.80 ± 0.05	-14.78 ± 0.10	-14.83 ± 0.02	-14.87 ± 0.04	-14.68 ± 0.02	-14.65 ± 0.03	-13.90 ± 0.04	-13.91 ± 0.03
88	-15.37 ± 0.05	-15.30 ± 0.09	-15.34 ± 0.03	-15.18 ± 0.07	-15.19 ± 0.03	-15.06 ± 0.06	-14.35 ± 0.04	-14.26 ± 0.09
89	-15.25 ± 0.07	-15.18 ± 0.10	-15.25 ± 0.03	-15.19 ± 0.06	-15.11 ± 0.03	-15.04 ± 0.04	-14.25 ± 0.04	-14.18 ± 0.05
90	-14.76 ± 0.04	-14.87 ± 0.07	-14.87 ± 0.03	-14.99 ± 0.04	-14.76 ± 0.03	-14.90 ± 0.05	-13.93 ± 0.03	-14.09 ± 0.05
91 94	-14.79 ± 0.06 -14.83 ± 0.05	-15.00 ± 0.14 -15.05 ± 0.09	-14.87 ± 0.03 -14.87 ± 0.03	-15.03 ± 0.07 -15.08 ± 0.07	$-14.76 \pm 0.04 \ -14.75 \pm 0.03$	$-14.94 \pm 0.07 \ -14.97 \pm 0.07$	-13.98 ± 0.05 -13.92 ± 0.03	-14.17 ± 0.10 -14.15 ± 0.08
95	-15.12 ± 0.06	-15.03 ± 0.09 -15.13 ± 0.08	-15.20 ± 0.03	-15.03 ± 0.07 -15.22 ± 0.04	-15.06 ± 0.03	-15.07 ± 0.07 -15.07 ± 0.04	-13.92 ± 0.03 -14.18 ± 0.03	-14.13 ± 0.03 -14.20 ± 0.06
96	-14.61 ± 0.07	-14.78 ± 0.14	-14.79 ± 0.02	-15.03 ± 0.09	-14.70 ± 0.03	-14.87 ± 0.01	-13.84 ± 0.03	-14.03 ± 0.12
97	-14.48 ± 0.07	-14.49 ± 0.11	-14.65 ± 0.02	-14.68 ± 0.06	-14.65 ± 0.02	-14.67 ± 0.07	-13.85 ± 0.04	-13.85 ± 0.09
98	-14.21 ± 0.05	-14.65 ± 0.07	-14.56 ± 0.02	-14.98 ± 0.03	-14.58 ± 0.03	-14.93 ± 0.03	-13.80 ± 0.04	-14.27 ± 0.04
99	-13.36 ± 0.06	-13.49 ± 0.08	-13.99 ± 0.02	-14.17 ± 0.03	-14.18 ± 0.02	-14.35 ± 0.03	-13.46 ± 0.03	-13.61 ± 0.02
100	-14.82 ± 0.05	-15.04 ± 0.07	-14.91 ± 0.02	-15.11 ± 0.04	-14.80 ± 0.03	-15.02 ± 0.04	-13.93 ± 0.03	-14.16 ± 0.05
102	-14.55 ± 0.06	-14.61 ± 0.07	-14.71 ± 0.03	-14.75 ± 0.05	-14.71 ± 0.03	-14.75 ± 0.06	-13.90 ± 0.04	-13.94 ± 0.05
103 105	-14.72 ± 0.05 -14.93 ± 0.06	$-14.75 \pm 0.08 \ -14.88 \pm 0.11$	-14.80 ± 0.02 -15.03 ± 0.03	-14.89 ± 0.06 -14.99 ± 0.07	-14.70 ± 0.03 -14.94 ± 0.03	$-14.76 \pm 0.05 \ -14.82 \pm 0.05$	-13.99 ± 0.03 -14.10 ± 0.04	-14.03 ± 0.05 -14.03 ± 0.08
103	-14.93 ± 0.06 -14.98 ± 0.06	-14.83 ± 0.11 -14.94 ± 0.08	-15.03 ± 0.03 -15.03 ± 0.02	-14.99 ± 0.07 -14.87 ± 0.06	-14.94 ± 0.03 -14.92 ± 0.03	-14.82 ± 0.03 -14.84 ± 0.06	-14.10 ± 0.04 -14.15 ± 0.03	-14.03 ± 0.08 -14.11 ± 0.05
108	-15.19 ± 0.09	-15.11 ± 0.16	-15.03 ± 0.02 -15.18 ± 0.05	-15.11 ± 0.09	-15.03 ± 0.05	-14.98 ± 0.09	-14.13 ± 0.03 -14.20 ± 0.07	-14.11 ± 0.03 -14.12 ± 0.08
109	-15.00 ± 0.06	-15.05 ± 0.11	-15.05 ± 0.02	-15.11 ± 0.08	-14.92 ± 0.02	-15.01 ± 0.07	-14.12 ± 0.04	-14.20 ± 0.08
111	-14.93 ± 0.06	-15.25 ± 0.09	-14.94 ± 0.03	-15.26 ± 0.05	-14.77 ± 0.03	-15.13 ± 0.04	-14.00 ± 0.04	-14.36 ± 0.05

Table 3 (Continued)

SN	8	<i>g</i>		r		i		z	
	True	Inferred	True	Inferred	True	Inferred	True	Inferred	
112	-14.36 ± 0.08	-14.29 ± 0.14	-14.63 ± 0.04	-14.54 ± 0.08	-14.63 ± 0.03	-14.54 ± 0.08	-13.93 ± 0.04	-13.87 ± 0.07	
113	-14.50 ± 0.07	-14.68 ± 0.11	-14.72 ± 0.03	-14.89 ± 0.07	-14.68 ± 0.03	-14.86 ± 0.05	-13.84 ± 0.04	-14.02 ± 0.05	
114	-14.63 ± 0.07	-14.77 ± 0.13	-14.73 ± 0.04	-14.91 ± 0.08	-14.62 ± 0.04	-14.78 ± 0.06	-13.82 ± 0.05	-13.98 ± 0.10	
115	-15.38 ± 0.08	-15.10 ± 0.14	-15.41 ± 0.03	-15.11 ± 0.09	-15.24 ± 0.04	-14.97 ± 0.09	-14.49 ± 0.04	-14.18 ± 0.07	
116	-14.18 ± 0.08	-14.50 ± 0.13	-14.42 ± 0.04	-14.77 ± 0.07	-14.35 ± 0.05	-14.67 ± 0.05	-13.69 ± 0.07	-14.00 ± 0.07	
117	-15.06 ± 0.09	-15.06 ± 0.13	-15.12 ± 0.04	-15.14 ± 0.07	-15.00 ± 0.05	-14.99 ± 0.04	-14.13 ± 0.06	-14.13 ± 0.06	
118	-15.09 ± 0.07	-14.87 ± 0.12	-15.18 ± 0.04	-14.96 ± 0.05	-15.05 ± 0.05	-14.83 ± 0.07	-14.27 ± 0.05	-14.02 ± 0.05	

4. FUTURE OUTLOOK: SPECTROPHOTOMETRIC DATA

The SNfactory generates spectrophotometry for objects at low redshift while the JDEM/ISWG design for a space telescope dark energy experiment called for spectrophotometric time series of high-redshift supernovae. In combination, such data would bring the potential of providing improved distance measurements with lower statistical and systematic uncertainty by preserving information lost in broadband photometry.

Specific spectral features have been correlated with absolute magnitude (Bongard et al. 2006; Bailey et al. 2009; Wang et al. 2009; Foley & Kasen 2011; Chotard et al. 2011; Silverman et al. 2012) and may provide additional information not contained in broadband light curves that further standardize absolute magnitudes. Using the approach of Section 2 with each dispersion element serving as a filter, the variance in resolved features can be correlated with absolute magnitude.

Supernovae have varying magnitude dispersion depending on wavelength and phase (James et al. 2006) that can be modeled as a Gaussian process. Measurements $m_{(1+z)t,(1+z)\lambda}$ for the observer phase and wavelength (1+z)t and $(1+z)\lambda$ are drawn as

$$m_{((1+z)t,(1+z)\lambda)} \sim \text{GP}(\bar{M}(t,\lambda; p_{\bar{M}}) + 2.5 \log (\mu(z, \mathbf{p}_{\Omega}), \times k_m(t,\lambda,t',\lambda'; \mathbf{p}_{k_m}) + n_m(t,\lambda,t',\lambda'; \mathbf{p}_{n_m})).$$

$$(11)$$

With spectrophotometric data, the cosmological parameters $\mathbf{p}_{\Omega} = \{\Omega_M, \Omega_{DE}, w_0, w_a\}$ that determine the distance modulus μ predicted by theory can be fit simultaneously with the supernova parameters.

Such an analysis using SNfactory data alone is of interest. Although insensitive to the cosmological parameters, the hyperparameters \mathbf{p}_{k_m} in particular provide insight into the interplay between wavelength and phase with the standard candle nature of SNe Ia. This information aids in the planning of future missions and hardware and quantifies the systematic uncertainties potentially induced when using degraded data (e.g., broadband photometry) only.

5. CONCLUSIONS

The regressed values from supernova light curves provide a large number of parameters that may correlate with absolute magnitude. Gaussian processes provide a convenient and parameterizable method of performing interpolations between measurements of light-curve shape and color and for using this information to predict absolute magnitudes.

For a proof of concept, we present an example application showing how spectrophotometry of low-redshift supernovae can be used to determine distances from broadband photometry of high-redshift supernovae, such as will be obtained by DES and LSST. Using SNfactory data, we find dispersions in distance determinations as high as 0.44 mag corrected down to 0.04 mag within the training set and to as low as 0.09 mag when applied to a validation set. Some of this dispersion is attributable to the data themselves. Our procedure is highly competitive with SALT2 and MLCS2k2, which give training-set dispersions of 0.12 and 0.14, respectively.

The spectrophotometry of SNfactory provides an excellent data set for exploring supernova studies. Traditionally, broadband light curves of $z \sim 0$ supernovae are used to train light-curve fitters. The spectrophotometric templates allow the calibration of the absolute magnitude to be optimized for any redshift. With spectroscopy, a light-curve model can be tailor made for every object from an imaging survey. For planning of future supernova-specific imaging surveys, our methodology and SNfactory data can be used to determine optimal filter sets that minimize intrinsic dispersion over the targeted redshift range.

There are directions for future development of the method and exploration of the data.

- There are further improvements possible in the implementation of the methodology. A kernel other than the square exponential may better describe supernova light curves and their relationship with the absolute magnitude. We did not perform an exhaustive search for the absolute maximum likelihoods in the parameter space since the local maxima found in this analysis were sufficient to establish our proof of concept.
- 2. When constraining the problem to broadband photometry, the full wavelength range and resolution available is not exploited. Analysis of intrinsic supernova properties within SNfactory and simultaneous analysis of low- and high-redshift samples are planned in future work.
- 3. There is a correlation between supernova absolute magnitudes and host-galaxy properties after light-curve shape corrections (Kelly et al. 2010; Sullivan et al. 2010; Lampeitl et al. 2010; Gupta et al. 2011; Childress et al. 2013). The search for such correlations with the absolute-magnitude residuals from our method is ongoing.
- 4. The observed variance of SNfactory measurements, and indeed of all supernova data, cannot be entirely attributed to photon noise. Gaussian processes provide a natural way to model and fit for external sources of photometric uncertainty.

We thank Bruce Bassett for fruitful discussions and Dan Birchall for observing assistance. We thank the technical and scientific staffs of the Palomar Observatory, the High

SN	i	r	i	i		Ζ
	True	Inferred	True	Inferred	True	Inferred
0	-15.27 ± 0.04	-15.21 ± 0.08	-14.98 ± 0.04	-14.90 ± 0.06	-14.75 ± 0.03	-14.69 ± 0.05
1	-15.14 ± 0.03	-15.22 ± 0.07	-14.91 ± 0.02	-14.83 ± 0.06	-14.73 ± 0.03	-14.69 ± 0.06
3	-15.27 ± 0.03	-15.27 ± 0.07	-15.01 ± 0.03	-15.01 ± 0.06	-14.79 ± 0.03	-14.79 ± 0.06
4 6	-15.11 ± 0.04	-14.95 ± 0.12	-14.85 ± 0.03	-14.72 ± 0.09	-14.63 ± 0.03	-14.50 ± 0.09
7	-15.32 ± 0.04 -15.22 ± 0.04	-15.23 ± 0.09 -15.29 ± 0.08	-15.03 ± 0.04 -14.96 ± 0.03	-14.91 ± 0.08 -15.05 ± 0.06	-14.82 ± 0.03 -14.75 ± 0.02	-14.70 ± 0.07 -14.81 ± 0.07
8	-15.23 ± 0.05	-15.11 ± 0.08	-14.94 ± 0.04	-14.81 ± 0.06	-14.72 ± 0.02	-14.58 ± 0.04
9	-15.21 ± 0.02	-15.35 ± 0.08	-14.92 ± 0.03	-15.05 ± 0.08	-14.68 ± 0.02	-14.82 ± 0.07
10	-15.27 ± 0.03	-15.10 ± 0.09	-15.10 ± 0.03	-14.96 ± 0.08	-14.84 ± 0.03	-14.76 ± 0.06
12	-15.20 ± 0.04	-15.15 ± 0.11	-15.01 ± 0.02	-14.96 ± 0.10	-14.80 ± 0.02	-14.72 ± 0.12
13	-14.60 ± 0.03	-14.77 ± 0.11	-14.42 ± 0.03	-14.61 ± 0.07	-14.21 ± 0.03	-14.46 ± 0.08
14 16	-14.89 ± 0.03 -15.17 \pm 0.04	-14.84 ± 0.11 -15.14 \pm 0.13	-14.70 ± 0.02 -14.92 ± 0.03	-14.64 ± 0.11 -14.87 \pm 0.09	-14.53 ± 0.03 -14.70 ± 0.03	-14.41 ± 0.09 -14.69 ± 0.09
17	-15.17 ± 0.04 -15.31 ± 0.04	-13.14 ± 0.13 -14.99 ± 0.07	-14.92 ± 0.03 -15.14 ± 0.04	-14.76 ± 0.06	-14.70 ± 0.03 -14.84 ± 0.05	-14.55 ± 0.05 -14.55 ± 0.05
19	-13.91 ± 0.04	-14.36 ± 0.07	-13.91 ± 0.03	-14.37 ± 0.06	-13.84 ± 0.04	-14.32 ± 0.05
21	-14.55 ± 0.03	-14.41 ± 0.03	-14.44 ± 0.02	-14.35 ± 0.03	-14.31 ± 0.03	-14.23 ± 0.03
22	-15.49 ± 0.03	-15.29 ± 0.09	-15.19 ± 0.02	-15.01 ± 0.09	-15.00 ± 0.03	-14.79 ± 0.11
24	-14.73 ± 0.03	-14.79 ± 0.06	-14.63 ± 0.03	-14.66 ± 0.05	-14.48 ± 0.03	-14.50 ± 0.04
27	-15.42 ± 0.05	-15.35 ± 0.07	-15.19 ± 0.04	-15.13 ± 0.06	-14.91 ± 0.04	-14.84 ± 0.04
30 33	-15.07 ± 0.03 -15.15 ± 0.03	-15.27 ± 0.09 -15.29 ± 0.08	-14.90 ± 0.03 -14.94 ± 0.03	-15.08 ± 0.10 -15.07 ± 0.08	-14.68 ± 0.03 -14.71 ± 0.03	-14.87 ± 0.08 -14.83 ± 0.07
36	-15.13 ± 0.03 -15.30 ± 0.04	-15.29 ± 0.08 -15.02 ± 0.08	-15.10 ± 0.03	-13.07 ± 0.08 -14.82 ± 0.08	-14.95 ± 0.04	-14.67 ± 0.06
38	-15.44 ± 0.03	-15.40 ± 0.06	-15.15 ± 0.03	-15.11 ± 0.05	-14.88 ± 0.03	-14.84 ± 0.06
39	-15.36 ± 0.03	-15.50 ± 0.07	-15.06 ± 0.03	-15.18 ± 0.06	-14.77 ± 0.03	-14.87 ± 0.05
45	-15.49 ± 0.07	-15.37 ± 0.11	-15.21 ± 0.05	-15.09 ± 0.06	-14.98 ± 0.06	-14.87 ± 0.07
47	-14.94 ± 0.04	-14.93 ± 0.09	-14.85 ± 0.03	-14.86 ± 0.09	-14.67 ± 0.03	-14.66 ± 0.06
49	-15.04 ± 0.03	-15.15 ± 0.09	-14.84 ± 0.03	-14.96 ± 0.07	-14.59 ± 0.03	-14.73 ± 0.05
54 55	-15.38 ± 0.05 -15.05 ± 0.03	-15.24 ± 0.08 -15.20 ± 0.08	-15.15 ± 0.04 -14.81 ± 0.03	-15.01 ± 0.09 -14.91 ± 0.08	-14.93 ± 0.04 -14.60 ± 0.03	-14.78 ± 0.06 -14.72 ± 0.07
57	-15.40 ± 0.03	-15.23 ± 0.04	-15.20 ± 0.03	-15.05 ± 0.03	-14.94 ± 0.03	-14.80 ± 0.03
59	-15.14 ± 0.04	-15.20 ± 0.08	-14.87 ± 0.03	-14.96 ± 0.08	-14.71 ± 0.03	-14.73 ± 0.07
62	-15.26 ± 0.04	-15.31 ± 0.06	-15.00 ± 0.04	-15.06 ± 0.05	-14.75 ± 0.03	-14.81 ± 0.05
63	-14.99 ± 0.03	-15.00 ± 0.10	-14.74 ± 0.02	-14.75 ± 0.09	-14.56 ± 0.03	-14.58 ± 0.10
66	-14.88 ± 0.04	-14.94 ± 0.08	-14.67 ± 0.02	-14.74 ± 0.08	-14.51 ± 0.03	-14.55 ± 0.07
67 69	$-15.50 \pm 0.05 -15.09 \pm 0.05$	$-15.39 \pm 0.10 \\ -15.13 \pm 0.08$	$-15.32 \pm 0.04 \ -14.89 \pm 0.04$	-15.19 ± 0.11 -14.93 ± 0.07	$-15.08 \pm 0.04 \ -14.66 \pm 0.04$	-14.97 ± 0.11 -14.72 ± 0.07
70	-13.05 ± 0.03 -14.85 ± 0.04	-13.13 ± 0.08 -14.95 ± 0.08	-14.62 ± 0.04 -14.62 ± 0.03	-14.74 ± 0.07	-14.44 ± 0.04	-14.72 ± 0.07 -14.51 ± 0.08
71	-15.01 ± 0.03	-14.93 ± 0.08	-14.82 ± 0.03	-14.78 ± 0.06	-14.67 ± 0.03	-14.62 ± 0.06
73	-15.23 ± 0.04	-14.98 ± 0.10	-15.04 ± 0.03	-14.79 ± 0.09	-14.82 ± 0.03	-14.61 ± 0.08
74	-14.75 ± 0.03	-14.81 ± 0.10	-14.51 ± 0.03	-14.62 ± 0.06	-14.36 ± 0.03	-14.43 ± 0.08
76	-15.15 ± 0.05	-15.24 ± 0.08	-14.95 ± 0.05	-15.05 ± 0.07	-14.72 ± 0.05	-14.84 ± 0.06
79 81	$-14.81 \pm 0.04 \\ -15.10 \pm 0.02$	$-14.90 \pm 0.09 \ -14.90 \pm 0.07$	$-14.62 \pm 0.03 \\ -14.98 \pm 0.02$	$-14.71 \pm 0.08 \ -14.82 \pm 0.08$	$-14.47 \pm 0.04 \\ -14.81 \pm 0.03$	-14.54 ± 0.07 -14.64 ± 0.06
82	-15.10 ± 0.02 -15.20 ± 0.03	-14.30 ± 0.07 -15.27 ± 0.08	-14.93 ± 0.02 -14.97 ± 0.03	-14.82 ± 0.03 -15.01 ± 0.07	-14.31 ± 0.03 -14.72 ± 0.02	-14.04 ± 0.00 -14.78 ± 0.07
84	-15.04 ± 0.03	-14.97 ± 0.09	-14.81 ± 0.02	-14.75 ± 0.09	-14.61 ± 0.03	-14.48 ± 0.06
85	-14.99 ± 0.03	-15.12 ± 0.05	-14.72 ± 0.03	-14.85 ± 0.03	-14.48 ± 0.03	-14.60 ± 0.03
88	-15.42 ± 0.03	-15.30 ± 0.07	-15.25 ± 0.03	-15.15 ± 0.08	-15.01 ± 0.03	-14.91 ± 0.08
89	-15.44 ± 0.03	-15.33 ± 0.08	-15.15 ± 0.02	-15.02 ± 0.07	-14.93 ± 0.03	-14.79 ± 0.05
90	-15.01 ± 0.03	-15.12 ± 0.09	-14.76 ± 0.02	-14.86 ± 0.07	-14.56 ± 0.03	-14.69 ± 0.06
91 94	$-15.01 \pm 0.04 -15.06 \pm 0.04$	$-15.14 \pm 0.06 \ -15.22 \pm 0.08$	$-14.78 \pm 0.03 \ -14.77 \pm 0.04$	$-14.92 \pm 0.06 \ -14.92 \pm 0.08$	-14.57 ± 0.03 -14.56 ± 0.03	-14.70 ± 0.06 -14.71 ± 0.06
95	-15.00 ± 0.04 -15.31 ± 0.04	-15.22 ± 0.03 -15.28 ± 0.05	-14.77 ± 0.04 -15.10 ± 0.03	-14.92 ± 0.03 -15.07 ± 0.04	-14.88 ± 0.03 -14.88 ± 0.03	-14.71 ± 0.00 -14.81 ± 0.04
96	-14.92 ± 0.05	-15.16 ± 0.09	-14.69 ± 0.03	-14.92 ± 0.07	-14.50 ± 0.03	-14.71 ± 0.06
97	-14.79 ± 0.03	-14.86 ± 0.09	-14.53 ± 0.02	-14.63 ± 0.07	-14.43 ± 0.02	-14.49 ± 0.07
98	-14.50 ± 0.03	-14.76 ± 0.08	-14.46 ± 0.03	-14.71 ± 0.04	-14.38 ± 0.03	-14.54 ± 0.04
99	-13.78 ± 0.04	-13.80 ± 0.06	-13.92 ± 0.02	-13.98 ± 0.04	-13.95 ± 0.03	-13.98 ± 0.03
100	-15.01 ± 0.03	-15.17 ± 0.05	-14.81 ± 0.02	-15.00 ± 0.04	-14.61 ± 0.03	-14.78 ± 0.03
102 103	-14.77 ± 0.03 -14.93 ± 0.04	$-14.79 \pm 0.06 \ -14.93 \pm 0.09$	-14.62 ± 0.03 -14.68 ± 0.03	$-14.59 \pm 0.05 \ -14.73 \pm 0.07$	-14.51 ± 0.03 -14.49 ± 0.03	-14.48 ± 0.06 -14.54 ± 0.08
105	-14.93 ± 0.04 -15.11 ± 0.03	-14.93 ± 0.09 -15.07 ± 0.09	-14.08 ± 0.03 -14.93 ± 0.03	-14.73 ± 0.07 -14.91 ± 0.07	-14.49 ± 0.03 -14.76 ± 0.03	-14.54 ± 0.08 -14.73 ± 0.08
107	-15.11 ± 0.03 -15.19 ± 0.04	-15.07 ± 0.09 -15.02 ± 0.08	-14.91 ± 0.03	-14.76 ± 0.06	-14.70 ± 0.03 -14.71 ± 0.03	-14.75 ± 0.03 -14.55 ± 0.07
108	-15.37 ± 0.07	-15.29 ± 0.11	-15.06 ± 0.05	-14.99 ± 0.07	-14.85 ± 0.06	-14.76 ± 0.09
109	-15.18 ± 0.03	-15.21 ± 0.08	-14.96 ± 0.02	-14.99 ± 0.09	-14.73 ± 0.03	-14.74 ± 0.08
111	-15.07 ± 0.04	-15.19 ± 0.10	-14.85 ± 0.03	-14.93 ± 0.07	-14.56 ± 0.03	-14.74 ± 0.07

Table 4 (Continued)

SN		r		i	z	
	True	Inferred	True	Inferred	True	Inferred
112	-14.63 ± 0.05	-14.53 ± 0.10	-14.56 ± 0.04	-14.47 ± 0.06	-14.41 ± 0.05	-14.36 ± 0.05
113	-14.75 ± 0.05	-14.79 ± 0.14	-14.63 ± 0.03	-14.73 ± 0.09	-14.49 ± 0.03	-14.57 ± 0.09
114	-14.86 ± 0.04	-15.01 ± 0.11	-14.62 ± 0.03	-14.74 ± 0.10	-14.42 ± 0.04	-14.55 ± 0.08
115	-15.53 ± 0.04	-15.26 ± 0.11	-15.31 ± 0.03	-15.02 ± 0.07	-15.03 ± 0.04	-14.79 ± 0.07
116	-14.48 ± 0.05	-14.76 ± 0.08	-14.32 ± 0.05	-14.61 ± 0.08	-14.13 ± 0.05	-14.42 ± 0.06
117	-15.28 ± 0.06	-15.20 ± 0.08	-15.03 ± 0.04	-14.96 ± 0.05	-14.82 ± 0.04	-14.76 ± 0.06
118	-15.33 ± 0.05	-15.25 ± 0.12	-15.07 ± 0.04	-14.99 ± 0.09	-14.86 ± 0.05	-14.76 ± 0.09

SN		r		i		ζ
	True	Inferred	True	Inferred	True	Inferred
0	-14.83 ± 0.07	-14.83 ± 0.10	-15.18 ± 0.04	-15.17 ± 0.06	-14.73 ± 0.04	-14.74 ± 0.05
1	-14.67 ± 0.05	-14.58 ± 0.09	-15.04 ± 0.02	-14.96 ± 0.09	-14.68 ± 0.02	-14.64 ± 0.06
3	-14.89 ± 0.05	-14.89 ± 0.10	-15.18 ± 0.03	-15.20 ± 0.08	-14.78 ± 0.03	-14.78 ± 0.05
4	-14.56 ± 0.07	-14.34 ± 0.13	-15.03 ± 0.03	-14.81 ± 0.11	-14.60 ± 0.03	-14.40 ± 0.10
6	-14.81 ± 0.07	-14.58 ± 0.14	-15.22 ± 0.03	-14.99 ± 0.07	-14.79 ± 0.03	-14.55 ± 0.09
7	-14.89 ± 0.06	-15.01 ± 0.12	-15.12 ± 0.03	-15.22 ± 0.09	-14.70 ± 0.03	-14.82 ± 0.07
8	-14.82 ± 0.05	-14.65 ± 0.07	-15.16 ± 0.05	-15.01 ± 0.10	-14.73 ± 0.05	-14.55 ± 0.05
9	-14.82 ± 0.05	-14.90 ± 0.09	-15.09 ± 0.02	-15.15 ± 0.07	-14.69 ± 0.02	-14.76 ± 0.00
10	-15.02 ± 0.06	-14.94 ± 0.11	-15.19 ± 0.04	-15.07 ± 0.06	-14.87 ± 0.03	-14.76 ± 0.06
12	-14.86 ± 0.06	-14.95 ± 0.12	-15.12 ± 0.04	-15.18 ± 0.10	-14.77 ± 0.03	-14.87 ± 0.10
13	-14.02 ± 0.06	-14.19 ± 0.12	-14.54 ± 0.04	-14.73 ± 0.09	-14.18 ± 0.04	-14.41 ± 0.09
14	-14.33 ± 0.05	-14.28 ± 0.15	-14.82 ± 0.03	-14.80 ± 0.16	-14.47 ± 0.02	-14.38 ± 0.12
17	-14.86 ± 0.07	-14.58 ± 0.14	-15.24 ± 0.04	-14.95 ± 0.11	-14.88 ± 0.03	-14.58 ± 0.12
19	-13.28 ± 0.06	-13.89 ± 0.11	-13.87 ± 0.03	-14.49 ± 0.05	-13.68 ± 0.03	-14.28 ± 0.03
21	-14.02 ± 0.04	-14.07 ± 0.09	-14.52 ± 0.04	-14.58 ± 0.07	-14.21 ± 0.03	-14.31 ± 0.06
22	-15.24 ± 0.07	-15.15 ± 0.12	-15.38 ± 0.03	-15.28 ± 0.09	-14.94 ± 0.03	-14.86 ± 0.07
24	-14.23 ± 0.06	-14.33 ± 0.10	-14.66 ± 0.04	-14.77 ± 0.09	-14.40 ± 0.03	-14.45 ± 0.07
27	-15.18 ± 0.06	-15.18 ± 0.09	-15.32 ± 0.04	-15.29 ± 0.05	-14.93 ± 0.04	-14.89 ± 0.04
30	-14.74 ± 0.06	-14.84 ± 0.10	-15.00 ± 0.03	-15.15 ± 0.07	-14.66 ± 0.03	-14.82 ± 0.09
33	-14.77 ± 0.05	-14.91 ± 0.08	-15.05 ± 0.03	-15.18 ± 0.05	-14.70 ± 0.03	-14.81 ± 0.07
36	-15.06 ± 0.06	-15.10 ± 0.13	-15.23 ± 0.04	-15.26 ± 0.08	-14.86 ± 0.03	-14.88 ± 0.08
38	-15.10 ± 0.06	-15.03 ± 0.11	-15.35 ± 0.03	-15.30 ± 0.07	-14.91 ± 0.03	-14.90 ± 0.08
39	-14.99 ± 0.05	-14.98 ± 0.10	-15.27 ± 0.03	-15.31 ± 0.07	-14.83 ± 0.03	-14.83 ± 0.03
45	-15.18 ± 0.08	-15.10 ± 0.11	-15.41 ± 0.07	-15.34 ± 0.08	-14.97 ± 0.06	-14.90 ± 0.04
47	-14.44 ± 0.07	-14.39 ± 0.16	-14.89 ± 0.03	-14.87 ± 0.10	-14.63 ± 0.03	-14.55 ± 0.09
49	-14.57 ± 0.05	-14.47 ± 0.14	-15.00 ± 0.04	-14.90 ± 0.14	-14.61 ± 0.03	-14.54 ± 0.12
54	-14.96 ± 0.07	-14.81 ± 0.14	-15.28 ± 0.05	-15.13 ± 0.11	-14.92 ± 0.04	-14.75 ± 0.07
55	-14.60 ± 0.04	-14.63 ± 0.11	-14.98 ± 0.03	-14.99 ± 0.07	-14.57 ± 0.03	-14.58 ± 0.08
57	-15.14 ± 0.04	-14.96 ± 0.11	-15.31 ± 0.03	-15.21 ± 0.07	-14.96 ± 0.03	-14.83 ± 0.06
59	-14.73 ± 0.06	-14.76 ± 0.09	-15.04 ± 0.03	-15.06 ± 0.06	-14.63 ± 0.03	-14.70 ± 0.05
62	-14.90 ± 0.05	-14.88 ± 0.09	-15.20 ± 0.03	-15.21 ± 0.08	-14.77 ± 0.03	-14.76 ± 0.07
63	-14.36 ± 0.05	-14.24 ± 0.10	-14.88 ± 0.04	-14.73 ± 0.11	-14.51 ± 0.03	-14.38 ± 0.09
66	-14.33 ± 0.06	-14.35 ± 0.14	-14.78 ± 0.04	-14.82 ± 0.09	-14.45 ± 0.03	-14.42 ± 0.10
67	-15.25 ± 0.06	-15.14 ± 0.10	-15.42 ± 0.05	-15.30 ± 0.09	-15.06 ± 0.04	-14.94 ± 0.06
69	-14.58 ± 0.05	-14.41 ± 0.13	-15.03 ± 0.05	-14.90 ± 0.11	-14.65 ± 0.04	-14.53 ± 0.12
70	-14.30 ± 0.07	-14.43 ± 0.14	-14.80 ± 0.04	-14.99 ± 0.12	-14.39 ± 0.03	-14.55 ± 0.11
71	-14.40 ± 0.05	-14.23 ± 0.08	-14.93 ± 0.03	-14.74 ± 0.11	-14.59 ± 0.03	-14.40 ± 0.09
73	-14.84 ± 0.05	-14.56 ± 0.11	-15.13 ± 0.04	-14.83 ± 0.09	-14.80 ± 0.03	-14.54 ± 0.07
74	-14.23 ± 0.05	-14.58 ± 0.13	-14.68 ± 0.03	-14.99 ± 0.12	-14.28 ± 0.03	-14.64 ± 0.08
76	-14.70 ± 0.06	-14.68 ± 0.08	-15.04 ± 0.04	-15.02 ± 0.05	-14.70 ± 0.05	-14.66 ± 0.05
79	-14.12 ± 0.07	-14.10 ± 0.11	-14.74 ± 0.03	-14.71 ± 0.06	-14.39 ± 0.04	-14.35 ± 0.06
81	-14.67 ± 0.04	-14.49 ± 0.09	-15.03 ± 0.03	-14.78 ± 0.09	-14.75 ± 0.02	-14.53 ± 0.06
82	-14.77 ± 0.05	-14.79 ± 0.08	-15.13 ± 0.03	-15.13 ± 0.09	-14.73 ± 0.03	-14.73 ± 0.06
84	-14.66 ± 0.04	-14.72 ± 0.13	-14.97 ± 0.03	-15.06 ± 0.11	-14.58 ± 0.03	-14.60 ± 0.08
85	-14.58 ± 0.05	-14.78 ± 0.05	-14.92 ± 0.03	-15.14 ± 0.03	-14.47 ± 0.03	-14.68 ± 0.04
88	-15.27 ± 0.06	-15.24 ± 0.11	-15.36 ± 0.03	-15.33 ± 0.09	-15.00 ± 0.03	-15.00 ± 0.04
89	-15.07 ± 0.06	-14.98 ± 0.10	-15.34 ± 0.03	-15.25 ± 0.08	-14.91 ± 0.03	-14.87 ± 0.07
90	-14.51 ± 0.07	-14.59 ± 0.14	-14.93 ± 0.03	-14.99 ± 0.10	-14.53 ± 0.04	-14.59 ± 0.09
	-14.51 ± 0.07	-14.58 ± 0.09	-14.94 ± 0.04	-15.02 ± 0.06	-14.54 ± 0.04	-14.61 ± 0.04

Table 5 (Continued)

SN	i	r		i		z
	True	Inferred	True	Inferred	True	Inferred
94	-14.53 ± 0.05	-14.74 ± 0.13	-14.95 ± 0.03	-15.20 ± 0.10	-14.54 ± 0.03	-14.74 ± 0.09
95	-14.91 ± 0.05	-14.87 ± 0.10	-15.22 ± 0.03	-15.18 ± 0.07	-14.86 ± 0.03	-14.82 ± 0.05
96	-14.33 ± 0.06	-14.54 ± 0.13	-14.83 ± 0.03	-15.04 ± 0.13	-14.47 ± 0.04	-14.65 ± 0.11
97	-14.10 ± 0.05	-14.16 ± 0.13	-14.68 ± 0.03	-14.72 ± 0.13	-14.32 ± 0.02	-14.41 ± 0.10
98	-14.17 ± 0.04	-14.79 ± 0.09	-14.44 ± 0.03	-15.13 ± 0.06	-14.23 ± 0.02	-14.76 ± 0.05
99	-13.00 ± 0.06	-13.43 ± 0.10	-13.75 ± 0.03	-14.20 ± 0.05	-13.69 ± 0.03	-14.04 ± 0.05
100	-14.52 ± 0.07	-14.60 ± 0.11	-14.91 ± 0.03	-15.01 ± 0.06	-14.57 ± 0.03	-14.68 ± 0.06
102	-14.34 ± 0.04	-14.62 ± 0.11	-14.69 ± 0.03	-14.90 ± 0.07	-14.38 ± 0.03	-14.64 ± 0.09
103	-14.50 ± 0.05	-14.55 ± 0.12	-14.86 ± 0.03	-14.91 ± 0.07	-14.44 ± 0.03	-14.48 ± 0.09
105	-14.72 ± 0.05	-14.80 ± 0.11	-15.03 ± 0.03	-15.12 ± 0.07	-14.68 ± 0.03	-14.74 ± 0.11
107	-14.75 ± 0.05	-14.67 ± 0.10	-15.10 ± 0.03	-15.01 ± 0.07	-14.66 ± 0.03	-14.59 ± 0.07
108	-14.92 ± 0.08	-14.83 ± 0.14	-15.29 ± 0.04	-15.21 ± 0.10	-14.83 ± 0.06	-14.75 ± 0.09
109	-14.85 ± 0.06	-14.97 ± 0.13	-15.10 ± 0.04	-15.20 ± 0.10	-14.72 ± 0.03	-14.82 ± 0.09
111	-14.54 ± 0.05	-14.51 ± 0.10	-14.98 ± 0.04	-14.94 ± 0.07	-14.62 ± 0.03	-14.60 ± 0.06
112	-14.10 ± 0.07	-14.17 ± 0.15	-14.59 ± 0.05	-14.72 ± 0.10	-14.33 ± 0.04	-14.46 ± 0.08
113	-14.27 ± 0.06	-14.33 ± 0.11	-14.64 ± 0.04	-14.70 ± 0.09	-14.41 ± 0.03	-14.50 ± 0.06
114	-14.34 ± 0.07	-14.53 ± 0.16	-14.81 ± 0.04	-15.06 ± 0.12	-14.39 ± 0.04	-14.61 ± 0.09
115	-15.18 ± 0.07	-14.88 ± 0.12	-15.44 ± 0.04	-15.15 ± 0.09	-15.06 ± 0.03	-14.82 ± 0.10
116	-13.91 ± 0.06	-14.25 ± 0.12	-14.44 ± 0.05	-14.78 ± 0.08	-14.10 ± 0.05	-14.45 ± 0.06
117	-14.88 ± 0.08	-14.84 ± 0.14	-15.17 ± 0.05	-15.12 ± 0.09	-14.79 ± 0.03	-14.76 ± 0.08
118	-14.87 ± 0.06	-14.57 ± 0.15	-15.24 ± 0.05	-14.96 ± 0.09	-14.83 ± 0.05	-14.53 ± 0.08

Table 6Results from the Linear Model

Filter System z	Band	σ_0	wrms	σ (wrms)	$\overline{\sigma_{ m int}}$	$\sigma(\sigma_{ m int})$	Four-fold Cross-validation	$N_{ m SN}$
0.00	g	0.322	0.146	0.026	0.123	0.049	0.154 (0.158)	50
0.00	r	0.229	0.118	0.014	0.049	0.036	0.120 (0.122)	49
0.00	i	0.265	0.137	0.018	0.099	0.021	0.136 (0.140)	50
0.25	g	0.437	0.141	0.039	0.119	0.052	0.155 (0.169)	44
0.25	r	0.301	0.119	0.039	0.068	0.052	0.122 (0.147)	43
0.25	i	0.236	0.091	0.019	0.025	0.030	0.090 (0.102)	43
0.25	z	0.250	0.136	0.032	0.111	0.032	0.130 (0.132)	44
0.50	r	0.391	0.153	0.033	0.126	0.042	0.156 (0.157)	50
0.50	i	0.301	0.149	0.023	0.123	0.021	0.148 (0.149)	48
0.50	z	0.248	0.135	0.025	0.093	0.027	0.132 (0.131)	49
0.75	r	0.320	0.140	0.036	0.093	0.062	0.148 (0.184)	35
0.75	i	0.237	0.129	0.042	0.088	0.059	0.133 (0.184)	35
0.75	z	0.210	0.131	0.035	0.080	0.058	0.133 (0.160)	35

Notes. For each filter system and calibrated band: σ_0 , the raw standard deviation of $M_{(t_0,\lambda_0)}$ for the input sample; $\overline{\text{wrms}}$ and $\sigma(\text{wrms})$, the mean and standard deviation over the four validation samples of the wrms difference between true and inferred magnitude; the mean intrinsic dispersion and its standard deviation over the four validation samples, the four-fold cross-validation error estimate with (and without) rejection of objects with extreme parameter values; N_{SN} is the number of validated supernovae.

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APPENDIX

REVIEW OF GAUSSIAN PROCESSES

In this appendix, we introduce the Gaussian process and provide expressions for the likelihood and normal probability distribution function used in this paper. A broader description of Gaussian processes can be found in Rasmussen & Williams (2006).

A Gaussian process is denoted as

$$f(\mathbf{x}) \sim \text{GP}(m(\mathbf{x}), k(\mathbf{x}, \mathbf{x}')).$$
 (A1)

Filter System z Four-fold Cross-validation Band σ (wrms) $\overline{\sigma_{\rm int}}$ $\sigma(\sigma_{\rm int})$ $N_{\rm SN}$ σ_0 0.322 0.044 0.00 0.129 0.040 0.099 0.134 (0.145) 50 g0.00 0.229 0.120 0.041 0.087 0.039 0.123 (0.130) 49 r 0.040 0.00 i 0.265 0.131 0.100 0.045 0.136 (0.146) 50 0.25 0.437 0.101 0.017 0.065 0.018 0.102 (0.141) 44 g 0.25 0.301 0.100 0.021 0.068 0.026 0.102 (0.139) 43 0.25 0.236 0.093 0.013 0.064 0.017 0.094 (0.128) 43 0.250 0.025 0.024 0.25 0.113 0.076 0.113 (0.158) 44 0.50 0.391 0.019 0.033 50 0.122 0.0850.121 (0.118) 0.50 0.301 0.118 0.013 0.077 0.021 0.116 (0.121) 48 0.50 0.248 0.114 0.005 0.073 0.006 0.113 (0.112) 49 Z. 35 0.75 0.320 0.129 0.046 0.072 0.061 0.136 (0.189) 0.054 0.050 35 0.75 0.237 0.131 0.068 0.147 (0.212) 0.039 0.75 0.210 0.120 0.032 0.047 0.123 (0.163) 35

Table 7
As in Table 6, but for the Gaussian Process Model

The expectation of the stochastic function $f(\mathbf{x})$ is $m(\mathbf{x})$ with a model covariance between any two points $\operatorname{cov}(f(\mathbf{x}_p), f(\mathbf{x}_q)) = k(\mathbf{x}_p, \mathbf{x}_q)$. Both m and k may be parameterized and the parameters of the latter are referred to as hyperparameters since they describe function scatter rather than the function itself. For a set of input points X_{\star} , the values of the function are drawn from a normal distribution

$$\mathbf{f}_{\star} \sim \mathcal{N}(m(X_{\star}), K(X_{\star}, X_{\star})),$$
 (A2)

where the covariance matrix K has elements filled with all input pairs of $k(\mathbf{x}_{\star,i}, \mathbf{x}_{\star,j})$.

Since the probability distribution function is described by a Gaussian, several statistics are expressed analytically after algebraic manipulation. The likelihood that \mathbf{y} , a set of n measurements of f at inputs X, with measurement covariance V described by a Gaussian process is written as

$$\log p(\mathbf{y}|X) = -\frac{1}{2}(\mathbf{y} - m(X))^{T}(K + V)^{-1}(\mathbf{y} - m(X))$$
$$-\frac{1}{2}\log|K + V| - \frac{n}{2}\log 2\pi. \tag{A3}$$

A set of measurements and function values is drawn from a normal distribution

$$\begin{bmatrix} \mathbf{y} \\ \mathbf{f}_{\star} \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} m(X) \\ m(X_{\star}) \end{bmatrix}, \begin{bmatrix} K(X, X) + V & K(X, X_{\star}) \\ K(X_{\star}, X) & K(X_{\star}, X_{\star}) \end{bmatrix} \right). \tag{A4}$$

The conditional distribution of the function values is Gaussian with expected mean

$$\bar{\mathbf{f}}_{\star} = m(X) + K(X_{\star}, X)[K(X, X) + V]^{-1}(\mathbf{y} - m(X))$$
 (A5)

and covariance

$$\operatorname{cov}(\bar{\mathbf{f}}_{\star}) = K(X_{\star}, X_{\star}) - K(X_{\star}, X)[K(X, X) + V]^{-1}K(X, X_{\star}). \tag{A6}$$

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