

# Nowcasting Covid-19 onset in the UK

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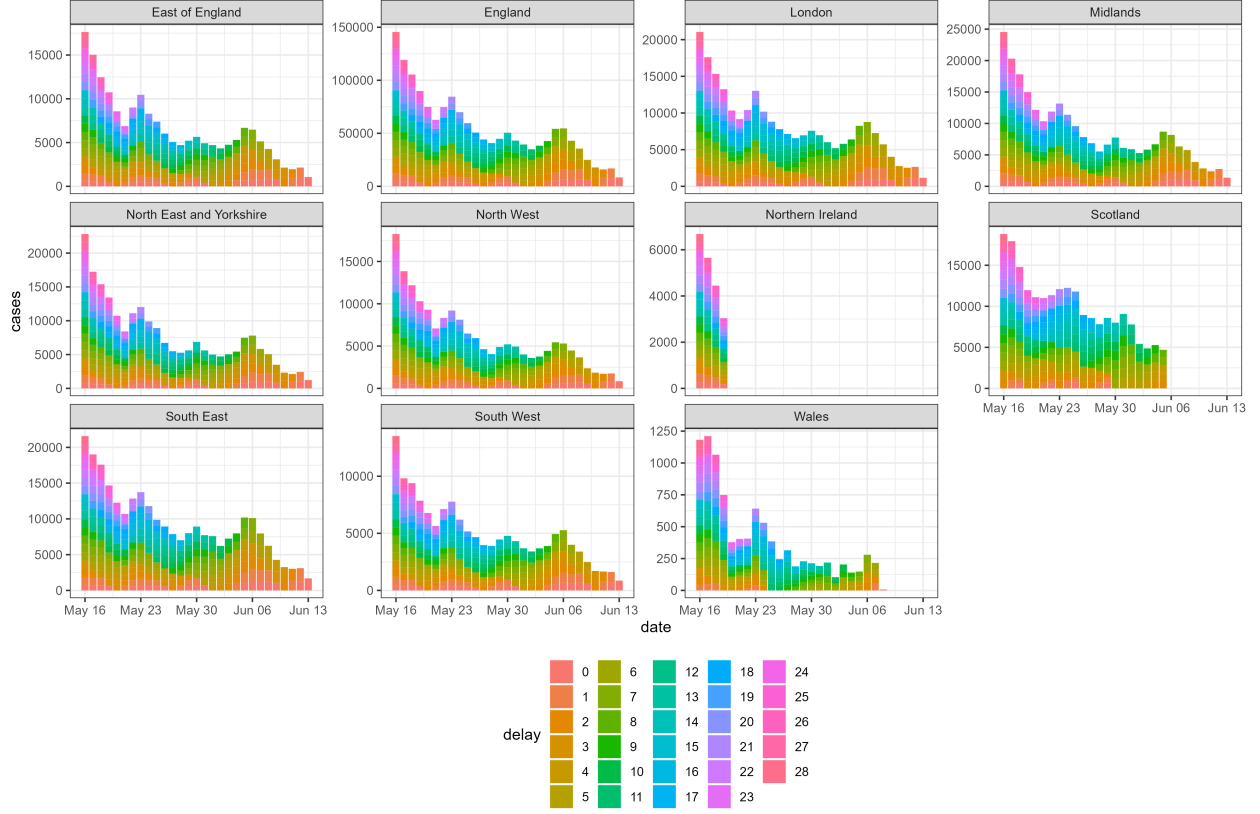
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## Software

- Successfully installed `epinowcast`, managed to run the nowcast for hospitalisations in Germany
- Set up C++ tool chain for `CmdStan` v2.29.2

## Data

- Downloaded `newCasesBySpecimenDate` from the UK Covid-19 dashboard
  - Aggregated by LTLAs then NHS England regions
  - Age stratification also available `newCasesBySpecimenDateAgeDemographics`
  - What is a suitable time frame?
- Plotted cases over time by reporting delay
  - Scotland, Wales, and Northern Ireland seem to be working on different reporting schedules
  - What other visualisations are helpful?
  - Would be interesting to see if the distribution of delays changed over time maybe due to testing advice



## Theory

The nowcasting problem is described as follows (Höhle and Heiden 2014).

- $n_{t,d}$ : number of cases with
  - onset on day  $t, t = 0, \dots, T$
  - reported with a delay of  $d, d = 0, \dots, D$  days, i.e. reports arrive on day  $t + d$ .
    - \* We may assume that delays can occur only up to a maximum of  $D$  days
- We observe  $\bowtie = (n_{t,d} : (t, d) \in A_T^m)$  where

$$\{(t, d) : \max(T - m, 0) \leq t \leq T, 0 \leq d \leq \min(D, T - t)\}$$

is the right-angle trapezoidal observation region at time  $T$  when using a moving window of size  $m$ .

- $N(t, T) = \sum_{d=0}^{\min(T-t, D)} n_{t,d}$ : number of cases which occurred on  $t$  which are reported until time  $T$ 
  - The aim of nowcasting is to predict the total number of cases which occurred on  $t, t = T - D, \dots, T$  given the information at time  $T$
- Reporting delay (in days) of a case occurring at time  $t$  follows a distribution with probability mass function  $f_t(d) = p_{t,d}, d = 0, \dots, D$ , where  $\sum_{d=0}^D p_{t,d} = 1$ .
  - Considering time  $t$  and conditioning on number of cases  $N(t, T)$  observed by time  $T$ , the observed  $n'_{t,d}$ s in the reporting triangle due to the right-truncation have a multinomial distribution with size  $N(t, \infty)$  and cell probabilities  $p'_{t,d} = p_{t,d} / \sum_{i=0}^{\min(D, T-t)} p_{t,i}$  for  $d = 0, \dots, \min(D, T - t)$ .
- If we assume that occurrence of new cases follow an underlying inhomogeneous Poisson process,  $N(t, \infty) \sim \text{Poisson}(\lambda_t), t = \max(0, T - m), \dots, T$ , the above conditional formulation of the  $n'_{t,d}$ s originating from multinomial sampling can be re-formulated as an unconditional likelihood corresponding to an incomplete contingency table

$$n_{t,d} \sim \text{Poi}(\mu_{t,d}), \quad (t, d) \in A_T^m, \text{ where } \log(\mu_{t,d}) = \log(\lambda_t) + \log(p_{t,d})$$

- Nowcasting can thus be divided into

1. Determining the  $\lambda'_t$ s
2. Determining the sequence of  $p'_{t,d}$ s
3. Predicting the unobserved  $n'_{t,d}$ s to compute the total  $N(t, \infty)$

Höhle, Michael, and Matthias an der Heiden. 2014. “Bayesian Nowcasting During the STEC O104:H4 Outbreak in Germany, 2011.” *Biometrics* 70 (4): 993–1002. <https://doi.org/10.1111/biom.12194>.