

Helping Children Catch Up: Early Life Shocks and the PROGRESA Experiment*

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Abstract

Children who face significant disadvantage early in life are often found to be worse off years or even decades later. Can conditional cash transfer programs mitigate the negative consequences and help these children catch up with their peers? We answer this question using data from rural Mexico, where rainfall shocks can have substantial effects on household income. We find that adverse rainfall in a child's year of birth decreases grade attainment, post-secondary enrollment, and employment outcomes. But declines were much smaller for children whose families were randomized to receive the conditional cash transfer program, PROGRESA: each additional year of PROGRESA exposure during childhood mitigated almost 20 percent of the early disadvantage in grade attainment.

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Poor circumstance in early life – even when it is temporary – often has long-lasting negative impacts (Almond and Currie, 2011; Currie and Vogl, 2012; Heckman, 2006, 2007). What role can public policy play in lessening the burden of adverse events in a young child’s life? This question is of core relevance to many areas of academic inquiry, and is critical in providing guidance on the allocation of scarce public resources. Much of the related work in economics focuses on evaluating the impacts of safety net policies that provide support to low-income children and families (Aizer et al., 2016; Chetty et al., 2016; Gertler et al., 2014; Hjort et al., 2017; Hoynes et al., 2016). This body of evidence shows that providing material and financial support during childhood can have positive impacts that last well into adulthood, often generating very large social returns (Bailey et al., 2020; Hendren and Sprung-Keyser, 2020).

We study a related but distinct question, for which the evidence thus far is quite limited (Almond et al., 2018). For children who have faced significant disadvantage or trauma early in life, are social protection programs capable of helping them catch up to their more fortunate peers? We focus specifically on conditional cash transfers (CCTs), a popular type of anti-poverty policy used widely across the globe. Our question, therefore, is essentially about heterogeneous returns to CCTs *within* the lower-income populations that are typically targeted by these policies. Do these programs have higher returns among children who have experienced early-life disadvantage compared to children with less exposure to early shocks? The answer to this question is important because it determines whether additional policies are needed above and beyond existing programs, targeting children who have experienced extreme disadvantage or trauma, in order to generate adequate catch-up.

Answering this question poses a substantial empirical challenge. First, we need a causal estimate of the effect of an early-life shock on later-life outcomes. This requires isolating variation in exposure to early life disadvantage that is orthogonal to other determinants of long-run outcomes. Second, in order to measure the extent to which a policy mitigates or exacerbates the effects of early-life disadvantage, we need to isolate exogenous variation in this policy. Because exposure to public programs is determined by parents’ preferences and local access to resources, which could also determine long run outcomes, comparing the outcomes of two people who faced the same shock but were differentially exposed to public policies will likely produce a biased estimate of the remediation value of these programs.¹

¹As Almond and Mazumder (2013) put it in their review of the literature, resolving this identification problem

Our study attempts to overcome this challenge. We leverage the combination of a natural experiment that induced variation in the extent of early disadvantage and a large-scale cluster randomized controlled trial of CCTs in Mexico, documented to have increased educational attainment. In our study's agrarian setting, where weather plays a key role in determining household income (and thus the availability of nutrition and other inputs for children), we verify that adverse rainfall lowers the agricultural wage and affects physical health. We then show that Mexican youth born during periods of adverse rainfall – years in which rainfall levels were more than one standard deviation above or below the locality-specific mean – have worse educational attainment and employment outcomes later in life than those born in normal rainfall periods. Exposure to adverse rainfall in the year of one's birth – a crucial period for the determination of long-term health and human capital – decreased years of completed schooling by more than half a year.

However, for children whose households were randomized to receive CCTs through PROGRESA, Mexico's landmark experiment in anti-poverty policy, each additional year of exposure mitigated the long-term impact of rainfall shocks on educational attainment by 0.1 years, almost 20%. PROGRESA could in theory have had both income and substitution effects, but we argue that the impacts we detect are driven primarily by the latter. By reducing the effective cost of schooling, PROGRESA enabled all children to stay in school longer than they would have otherwise, but had the largest effects on those impacted by negative rainfall shocks at birth. The negative effects of adverse rainfall become discernible after primary school, with the largest impacts measured for completion of grades 7 through 9. The mitigative impact of PROGRESA, as well as the main effect of the program, is also largest precisely in these years.

Finally, for the oldest individuals in our sample, we find a similar pattern of coefficients in regressions on post-high-school education and employment outcomes. Adverse rainfall in the year of birth leads to a reduction of 17 percentage points in the probability of working, but each additional year of PROGRESA exposure offsets nearly 8 percentage points of this impact.

This set of facts constitutes our main contribution: with respect to schooling and early employment outcomes, children born in times of hardship are the ones most responsive to CCTs provided in their school-aged years. This implies that public investment can indeed help children who faced

"may be asking for 'lightning to strike' twice: two identification strategies affecting the same cohort but at adjacent developmental stages."

adversity in early life catch up to their peers. Given that children were exposed to PROGRESA during school-aged years, its success at generating catch-up for disadvantage from the year of birth is striking. Several influential studies argue that there is little scope for catch-up when it comes to nutritional deficiencies that occur before a child's second birthday (Martorell et al., 1994; Victora et al., 2008), or test score gaps that appear by early elementary school (Heckman, 2006). However, there is other work, consistent with our findings, documenting that catch-up on both physical and cognitive dimensions is still possible after age 2 (Crookston et al., 2010, 2013; Lundein et al., 2014; Prentice et al., 2013).

A second important implication is that safety net policies geared toward low-income families in general may in effect target the neediest children within the targeted groups. Our results are similar to heterogeneous impacts found in recent evaluations of preschool policies in Germany (Cornelissen et al., 2018) and Denmark (Rossin-Slater and Wüst, 2020); from the Head Start program in the United States (Bitler et al., 2014); and from micronutrient supplementation in Bangladesh (Gunnsteinsson et al., 2021). Other studies find the opposite result or no evidence of significant interactions (Aguilar and Vicarelli, 2011; Duque et al., 2018; Johnson and Jackson, 2019; Malamud et al., 2016). This emphasizes that differences across settings – types of policies, access to resources, socioeconomic environments, and intervention timing – may determine whether catch-up is possible.

Our empirical context is particularly appealing because of the relatively high potential for external validity. Adverse rainfall is one of the most common type of shocks experienced by poor households in much of the developing world (Dinkelman, 2017), and has large short- and long-term consequences (Maccini and Yang, 2009; Paxson, 1992; Shah and Steinberg, 2017; Wolpin, 1982). Given the rising importance of wide-scale CCT programs around the world – including those modeled closely after PROGRESA itself (see, e.g., Das et al. (2005); Lagarde et al. (2007)) – it is important to learn here that these programs, if administered as successfully as PROGRESA was in Mexico, could potentially mitigate a sizable portion of the adverse impacts of poor rainfall at the time of birth.

1 Program Background

1.1 Description of Program

In 1997, the Mexican government began a CCT program called the Programa de Educación, Salud y Alimentación (PROGRESA). The program provided cash transfers to poor families (mothers, specifically), conditional on certain education and health-related requirements. Since then, the program has been expanded to urban areas and renamed, first to *Oportunidades* in 2002 and to *Prospera* in 2014.

In this paper, we focus on the education component of PROGRESA, which consisted of bimonthly cash payments to mothers during the school year, contingent on their children attending at least 85% of school days. Appendix Table D1 summarizes the monthly grant amounts by gender and grade level for the second semester of 1997, 1998 and 2003. At the program's onset, grants were provided only for children between third and ninth grade. In 2001, the grants were extended to high school.

For evaluation purposes, the program was implemented experimentally in 506 rural localities (in 191 municipalities) in the states of Guerrero, Hidalgo, Michoacan, Puebla, Queretaro, San Luis de Potosi and Veracruz. 320 “treatment” localities were randomly assigned to start receiving benefits in the spring of 1998. 186 localities were kept as a control group and started receiving PROGRESA benefits at the end of 1999. This randomized variation has allowed for rigorous evaluations of the program’s effects on a wide range of outcomes, which we discuss below. For more detail on PROGRESA’s health component, program targeting, and eligibility, see Appendix section A.

1.2 Previous Literature on PROGRESA Effects

An enormous body of research has explored the effects of PROGRESA on a wide array of outcomes (Parker et al., 2017). In Appendix Table D2, we attempt to summarize the key findings of this literature, categorizing studies based on the age of the analysis sample (specifically, how old they were during the years of PROGRESA being used to identify its effects) and by the main outcomes examined: education, health, cognitive or behavioral, or consumption-related.

It is clear that PROGRESA was successful at improving outcomes across all of these dimensions. For school-aged children, however, the main effect of PROGRESA was educational, which is not

surprising: this age group was the only one directly affected by the schooling subsidies and was too old to benefit from the main health benefits targeted toward much younger children. Consistent with this, Table D2 shows that most of the effects that PROGRESA had on health were concentrated among much younger (or much older) samples.

The main question we seek to answer is whether a government policy like PROGRESA can help remediate for disadvantage generated very early in life. We are therefore interested in studying the outcomes of children who were school-aged when the program was rolled out, for whom there is experimental variation in exposure to the schooling grant. When interpreting our results, therefore, we view the education subsidy channel as the main driver of the results we document, not the health component or the actual cash received.² This is consistent with what has been documented in the literature – large education effects for school-aged children but virtually no evidence of health effects for this age group – and with the design of the program.

2 Data

2.1 PROGRESA Data

The data collected for the evaluation of the PROGRESA program include a baseline survey of all households in PROGRESA villages and several follow-ups in 1998, 1999, 2003, and 2007. As we summarize in Table 1, we use the 2003 survey to obtain the outcome variables for our main analysis, and the baseline survey to construct control variables. For supplementary analysis, we also draw on both the 2003 and 2007 waves.

For our primary analysis, we focus on individuals aged 12 to 18 in 2003 in households who were eligible for the program (“poor” households). Following Behrman et al. (2011), we drop individuals who have non-matching genders across the 1997 and 2003 waves (1.9% of the sample), as well as those who report birth years that differ by more than 2 years (1.8% of the sample). For those with non-matching birth years with smaller than 2 year differences, we use the birth year reported in the 1997 wave. We restrict to the 12-18 age range because 12 year-olds are the youngest cohort for which there is differential exposure to PROGRESA in treatment and control villages (see Table 2),

²Though Appendix Table D2 shows that significant consumption effects have been documented, these are on the whole relatively small in magnitude (Parker et al., 2017).

Table 1: Variables and Survey Waves

| Variables | Survey Year | Ages |
|----------------------------------|-------------|--|
| A. Primary Outcomes | | |
| Education | 2003 | 12-18 |
| Employment | 2003 | 18 |
| B. Control Variables | | |
| Household demographics | 1997 | N/A |
| Locality characteristics | 2003 | N/A |
| C. Supplementary Outcomes | | |
| Weight, Height | 2003 | 2-6, 15-21 |
| Weight, Height | 2007 | 0-2, 8-10, adults 30+, mothers of young children |
| Behavioral | 2007 | 8-10 |
| Cognitive Tests | 2003 | 2-6, 15-18 |

while individuals over 18 are more likely to have moved out of the household by the 2003 survey and are therefore not surveyed (though the main respondents are still asked some questions about non-residents).³

For our supporting analysis, we use other physical, cognitive, and behavioral outcomes collected during the 2003 survey for specific age groups (in most cases, different from our main sample of interest). We also use child development measures collected for younger children in 2007.⁴

2.1.1 Outcome Variables

Our main education variables include educational attainment (in grades attained), a dummy for grade progression, and a dummy for having completed the appropriate number of grades for one's age. "Grade progression" is a binary variable equal to 1 if an individual progressed at least five complete grades between 1997 and 2003. We also define an indicator for age-appropriate grade completion. This is equal to 1 if an individual completed the appropriate number of grades for their age (e.g., one grade for a 7 year old). We also generate 10 dummy variables, each indicating whether the individual completed at least a certain number of grades (from 3 to 12 grades) of school.

³As Figure D1 shows, the proportion of 19-year-olds not living in the household is over 40%, and this proportion continues to grow with age.

⁴Unfortunately, high attrition rates prevent us from using the 2007 outcomes of our sample individuals. We lose over half of our 2003 sample, partially due to household-level attrition, but primarily due to individual migration (no proxy information is collected for those no longer living in the originally surveyed household) – likely to be endogenous.

For individuals who are 18 years old in 2003, we also look at continued enrollment and employment outcomes. Specifically, we create indicators for whether an individual is still enrolled in school after having received a high school degree. We are also interested in whether an individual was employed in the past week, employed in the past year, and employed in a non-laborer job in the past year. This last variable attempts to separate the jobs with the lowest earning growth potential from the rest of the employment categories by grouping those working as spot laborers with the unemployed. We verify using the Mexican Family Life Survey that youths who are 18 to 20 years old and working in a laborer job during the 2002 survey have among the lowest hourly wages during the 2009 survey (amounting to about two-thirds of the average of the rest of the sample).⁵

2.1.2 PROGRESA Exposure Variable

One of our main independent variables of interest is years of PROGRESA exposure. The length of exposure to the education component of the PROGRESA program depends on a child's locality and birth year. Table 2 shows, for each birth cohort, the number of years of exposure to PROGRESA by treatment status. This variable takes advantage of the variation in exposure lengths across different age cohorts within the treatment and control groups, in addition to the exogenous variation generated by the randomization of the PROGRESA program.

2.2 Rainfall Data

In addition to PROGRESA data, we use rainfall data from local weather stations collected by Mexico's National Meteorological Service (CONAGUA). We match those rainfall stations to program localities using their geocodes. For each locality, we use data from all stations within a 20 kilometer radius and take an inverse-distance weighted average of rainfall from these nearby stations. Using this procedure, 69 of the 506 localities are still missing rainfall measurements for our study period. Thus, our final sample, after excluding individuals missing rainfall for their particular year of birth, restricting to those from poor households in our desired age group meeting the data quality requirements, consists of individuals from 420 localities.

⁵Job categories differ across the two datasets, but the laborer category is similarly defined.

Table 2: Exposure to PROGRESA

| Cohort Age in 1998 | School Grade in 1998 | Age in 2003 | Age (year) when first exposed to PROGRESA | | Number of years exposed to PROGRESA by 2003 | | |
|-----------------------|----------------------|-------------|--|---------------------|--|---------------------|---------------------------|
| | | | Treatment Villages | Control Villages | Treatment Villages | Control Villages | Difference in Exposure |
| 5 | - | 10 | 8 (2001) | 8 (2001) | 3 | 3 | 0 |
| 6 | 1st year primary | 11 | 8 (2000) | 8 (2000) | 4 | 4 | 0 |
| 7 | 2nd year primary | 12 | 8 (1999) | 9 (2000) | 5 | 4 | 1 |
| 8 | 3rd year primary | 13 | 8 (1998) | 10 (2000) | 6 | 4 | 2 |
| 9 | 4th year primary | 14 | 9 (1998) | 11 (2000) | 6 | 4 | 2 |
| 10 | 5th year primary | 15 | 10 (1998) | 12 (2000) | 6 | 4 | 2 |
| 11 | 6th year primary | 16 | 11 (1998) | 13 (2000) | 6 | 4 | 2 |
| 12 | 1st year junior high | 17 | 12 (1998) | 14 (2000) | 6 | 4 | 2 |
| 13 | 2nd year junior high | 18 | 13 (1998) | 16 (2001) | 4 | 2 | 2 |
| 14 | 3rd year junior high | 19 | 14 (1998) | 17 (2001) | 2 | 1 | 1 |
| 15 | 1st year high school | 20 | - | - | 0 | 0 | 0 |
| 16 | 2nd year high school | 21 | - | - | 0 | 0 | 0 |

Notes:

- Initially, PROGRESA only applied to primary and junior high school. In 2001, the program was extended to all three years of high school. The control cohort aged 14 in 1998 aged out of the program at the end of 1999 and started receiving benefits again in 2001.
- Years of exposure is obtained by calculating the number of months, dividing by 12, and rounding to the nearest year (because there is some ambiguity about the precise month in which treatment households began receiving benefits (Skoufias, 2005; Hoddinott and Skoufias, 2004)).

2.2.1 Rainfall Shock Variable

We use rainfall as an exogenous shock to income during a child's first year of life. Specifically, we define a shock as a level of annual rainfall that is one standard deviation above or below the locality-specific mean (calculated over the 10 years prior to the birth year). We use this relative measure in order to capture the fact that the same amount of rainfall may have different consequences for different regions with different average rainfall levels. As we discuss in detail in section 3, both previous literature as well as our own data show that defining the shock variable in this way captures the contemporaneous relationship between rainfall and agricultural wages: normal rainfall is associated with better outcomes than extreme rainfall.

We use rainfall in an individual's calendar year of birth in their locality of residence in 1997.⁶ To calculate rainfall levels, we simply sum all monthly rainfall during an individual's calendar year of birth. We do not use month of birth to define this annual shock because approximately 30% of our sample reports different birth months in the 1997 and 2003 surveys. Appendix Figure

⁶The data do not include locality of birth, which would be the ideal geographic identifier in this context. We therefore use locality of residence (as of 1997), which should be accurate for most of the individuals in our sample, as long as migration among these young age groups is uncommon.

D2 maps all PROGRESA localities by their rainfall status, separately for treatment and control. Black dots represent localities that experienced a rainfall shock in 1987 (chosen, for illustrative purposes, because this is the modal birth year in our sample), while gray crosses represent those that experienced normal rainfall in that same year. For both treatment and control villages, we see a great deal of variation in rainfall shock status within states, and even within clusters of neighboring localities.

2.3 Summary Statistics

Table 3 reports summary statistics for individual-level variables from the 2003 survey for our sample of interest: individuals aged 12 to 18 (and for employment outcomes, only those aged 18) who live in households eligible for PROGRESA and satisfy the data quality requirements described in section 2.1. Panel A shows all education measures are significantly higher for treatment than control villages, while panel B shows employment outcomes for 18 year olds do not differ by treatment status on average.

Panel C reports summary statistics for the two independent variables of interest: PROGRESA exposure and birth year rainfall, which vary at the locality by birth year level. By experimental design, treatment villages were exposed to PROGRESA for longer than control villages. Mean rainfall, both in raw levels and in normalized terms, is not significantly different across treatment and control villages.

However, there is a small but statistically significant difference in the prevalence of a one-standard deviation shock. Since PROGRESA treatment was randomly allocated and rainfall shocks are arguably exogenous, this difference in the prevalence of a shock does not necessarily indicate an identification issue. However, this imbalance could be problematic if it resulted from a lack of common support across the treatment and control rainfall distributions. Accordingly, we verify in Appendix Figure D3 that the rainfall distributions for treatment and control localities indeed share a common support and are similar overall. Moreover, in Appendix Figure D2, though there are more shocks in control villages than treatment villages, the spatial distribution of rainfall shocks is similar across the two groups (and both quite disperse).

Nevertheless, in order to alleviate concerns that this imbalance is driving our results, we also trim the sample by excluding localities that could be considered outliers. That is, we drop any

Table 3: Summary Statistics

| | <i>Full Sample</i> | <i>Treatment Villages</i> | <i>Control Villages</i> | <i>Treatment - Control Differences</i> |
|---|------------------------|-------------------------------|-----------------------------|--|
| A. Individual Outcomes (12-18 year olds) | | | | |
| Educational Attainment | 6.79 (2.11) | 6.85 (2.09) | 6.69 (2.13) | 0.15*** (0.040) |
| Grade Progression | 0.58 (0.49) | 0.59 (0.49) | 0.56 (0.50) | 0.030*** (0.0096) |
| Appropriate Grade Completion | 0.46 (0.50) | 0.48 (0.50) | 0.44 (0.50) | 0.037*** (0.0094) |
| <i>Number of individuals</i> | 11829 | 7193 | 4636 | |
| <i>Number of localities</i> | 420 | 257 | 163 | |
| B. Individual Outcomes (18 year olds) | | | | |
| Currently Enrolled w/ HS Degree | 0.061 (0.24) | 0.058 (0.23) | 0.064 (0.25) | -0.0057 (0.012) |
| Worked this Week | 0.50 (0.50) | 0.51 (0.50) | 0.48 (0.50) | 0.029 (0.030) |
| Worked this Year | 0.53 (0.50) | 0.54 (0.50) | 0.52 (0.50) | 0.028 (0.030) |
| Worked in Non-Laborer Job | 0.35 (0.48) | 0.36 (0.48) | 0.35 (0.48) | 0.0051 (0.029) |
| <i>Number of individuals</i> | 1597 | 942 | 655 | |
| <i>Number of localities</i> | 368 | 218 | 150 | |
| C. Shock Variables (Full Sample) | | | | |
| Years of PROGRESA exposure | 4.84 (1.17) | 5.57 (0.73) | 3.69 (0.72) | 1.88*** (0.030) |
| Annual rainfall | 1182.4 (644.3) | 1180.6 (654.8) | 1185.3 (628.0) | -4.75 (26.3) |
| Normalized rainfall | -0.070 (0.81) | -0.054 (0.79) | -0.096 (0.84) | 0.042 (0.033) |
| Rainfall Shock | 0.24 (0.43) | 0.22 (0.42) | 0.27 (0.45) | -0.048*** (0.017) |
| <i>Number of locality x birth-year observations</i> | 2519 | 1536 | 983 | |
| <i>Number of localities</i> | 420 | 257 | 163 | |
| D. Shock Variables (Trimmed Sample) | | | | |
| Years of PROGRESA exposure | 4.81 (1.17) | 5.58 (0.72) | 3.71 (0.71) | 1.87*** (0.031) |
| Annual rainfall | 1181.1 (644.0) | 1171.1 (654.8) | 1195.5 (628.0) | -24.4 (28.1) |
| Normalized rainfall | -0.067 (0.84) | -0.051 (0.83) | -0.089 (0.86) | 0.038 (0.037) |
| Rainfall Shock | 0.28 (0.45) | 0.27 (0.44) | 0.29 (0.46) | -0.028 (0.020) |
| <i>Number of locality x birth-year observations</i> | 2170 | 1282 | 888 | |
| <i>Number of localities</i> | 344 | 203 | 141 | |

Notes:

Standard deviations (in the first 3 columns) and standard errors (in the last column) in parentheses (** p<0.01, ** p<0.05, * p<0.1). We do not cluster standard errors in these summary statistics but cluster at the municipality-level in all main results.

localities that either experienced no rainfall shocks throughout the sample period or experienced rainfall shocks in every year throughout the period. As shown in Panel D of Table 3, this trimming results in a sample of balanced rainfall shocks across treatment and control. Appendix Figure D4, which maps the geographic distribution of shocks for this trimmed sample, is not noticeably different from Appendix Figure D2, emphasizing that trimming did not substantially change the distribution of rainfall shocks (by removing localities only from a particular area, for example). In the appendix, we repeat our main empirical analysis using the trimmed sample and show that our results remain nearly unchanged (Table C5).

3 Empirical Strategy

To investigate whether PROGRESA can help generate catch-up for children who experienced early-life disadvantage, we need exogenous variation in early-life disadvantage as well as exogenous variation in exposure to PROGRESA. The randomized rollout of PROGRESA provides the latter. For the former, we turn to variation generated by rainfall shocks.

3.1 Early-Life Rainfall Shocks

In rural settings, good rainfall in early childhood means higher income, which may translate into increased nutritional availability during a crucial stage of development. Children exposed to negative rainfall shocks early in life often remain disadvantaged many years later, in terms of their health, human capital, and labor market outcomes (Dinkelman, 2017; Maccini and Yang, 2009; Shah and Steinberg, 2017).

Drawing on previous literature, as well as new analyses using our data, we argue that negative rainfall shocks do indeed generate substantial disadvantage in this setting. First, studying the same PROGRESA villages that we study in this paper, Bobonis (2009) finds that rainfall shocks, defined as monthly rainfall one standard deviation above or below the historical mean, reduce household expenditures by 16.7%. Next, using locality-level wages reported by village leaders in the PROGRESA data, we find evidence consistent with this. Appendix Figure D5 depicts the lowess-smoothed relationship between average male wages from the 2003 survey and rainfall in that same year, normalized using the locality-specific 10-year historical mean and standard deviation.

The inverted U-shape, which peaks at around zero, shows that wages are highest around the locality mean but fall at the tails of the rainfall distribution.

We also provide evidence that rainfall shocks affect nutrition, by examining effects on BMI. As we show in Table B1 and discuss in more detail in Appendix section B, contemporaneous rainfall shocks reduce BMI. In the same table, we show that these contemporaneous nutrition effects have longer-term implications for child health. Shifting attention to rainfall shocks in the year of birth (instead of the survey year), we find that adverse rainfall increases stunting for children aged 2 and older – by 4.2 percentage points (about 20% of the mean) for those aged 2-6, and 3.7 percentage points (about 40% of the mean) for those aged 8-10 at the time of survey.⁷

Finally, we also examine whether other dimensions of human capital are affected by birth-year rainfall, focusing on cognitive test scores and behavioral measures collected in 2003. We find that adverse birth-year rainfall had no significant effects on cognitive or behavioral measures for 2 to 6 year-olds, but did increase the likelihood of behavioral problems (externalizing problems, in particular) later in childhood. That income shocks in the year of birth can affect non-cognitive development is consistent with the child development literature, which documents that socioeconomic disadvantage is associated with altered maternal responses to infant emotions (Kim et al., 2017) and with other reasons for negative mother-infant interactions that could lead to behavioral problems later in childhood (Goyal et al., 2010).⁸

In sum, exposure to adverse rainfall early in life has substantial effects on household resources, nutrition, and health in our setting. In our analysis, we use adverse rainfall as a proxy for early-life disadvantage, noting that household income at the time of birth is not available in our data (and would be generally difficult to obtain in most settings). Even if this variable were available, however, the exogeneity of the rainfall shock provides an important advantage because it enables us to obtain a causal estimate of the effect of early-life disadvantage (and therefore a valid estimate of the amount

⁷To put these magnitudes into perspective, these increases in stunting correspond to average reductions in height-for-age z-scores of about 0.09 and 0.03 standard deviations for 2-6 year-olds and 8-10 year olds, respectively. For comparison, an additional month of exposure to civil war in Burundi led to a 0.05 standard deviation decrease in height-for-age z-scores (Bundervoet et al., 2009); in Colombia, a one standard deviation increase in early-life exposure to violence reduced height-for-age z-scores by 0.16 standard deviations (Duque, 2017); survivors who were infants during the 1984 Ethiopian famine were 5 centimeters (almost half of the sample standard deviation) shorter than unaffected individuals by young adulthood (Dercon and Porter, 2014).

⁸Because the samples used in Tables B1 and B2 were all exposed to the PROGRESA program by the time of survey (2003 or 2007), and the young cohorts in particular (both treatment and control) were exposed to the health component of the program, the estimated effects could be underestimating the main effect of adverse rainfall if PROGRESA had any remediating effect on these health outcomes.

of catch-up generated by PROGRESA). Because household income at the time of birth could be strongly correlated with household conditions later in life (during exposure to the PROGRESA program), using an exogenous rainfall shock also helps ensure that we are isolating catch-up based on early-life disadvantage rather than current circumstances. Programs like PROGRESA already target recipients based on current income levels – the goal of this paper is to investigate whether these programs help those who experienced additional disadvantage (early in life) catch up to other program recipients.

3.2 Specification

Letting z_{islt} denote education or employment outcomes for individual i , born in year t and living in state s and locality l in 1997, we estimate the regression specification below. See Appendix section E for details on how this estimating equation relates to the structural parameters of a life-cycle utility model of schooling choices, endowments, and conditional transfers.

$$z_{islt} = \beta_1 R_{slt} + \beta_2 P_{slt} + \beta_3 R_{slt}P_{slt} + \alpha' X_{islt} + \mu_s \times \delta_t + \epsilon_{islt}. \quad (1)$$

R_{slt} represents the rainfall shock dummy, indicating that rainfall during the individual's year of birth was more than one standard deviation away from the ten-year locality-specific mean. P_{slt} represents the number of years of PROGRESA exposure, which varies across treatment and control villages as well as across different birth cohorts within villages.

Our basic specification includes state \times birth year fixed effects ($\mu_s \times \delta_t$). In some specifications we add municipality (an administrative region larger than locality but smaller than state) fixed effects. Given that R_{slt} and P_{slt} both vary at the locality and birth year level, we could technically also include locality fixed effects, though these would absorb all of the variation generated by the PROGRESA randomization, the primary source of exogenous variation in this design. Therefore, municipality fixed effects are the smallest set of geographic fixed effects that we use.

β_1 represents the causal effect of a negative early-life income shock, and β_2 provides the causal effect of PROGRESA for individuals who did not experience this negative shock. β_3 provides the differential effect of PROGRESA for disadvantaged individuals (who experienced the negative shock).

A positive β_3 would indicate catch-up: larger effects of PROGRESA for the more disadvantaged individuals; a negative β_3 would suggest that PROGRESA widens the gap between disadvantaged and non-disadvantaged children.

We cluster our standard errors at the municipality level, which is a larger administrative unit than the locality. In addition to this, we also show standard errors that adjust for spatial correlation (unrelated to administrative boundaries) using the method described in Conley (1999). We report standard errors that allow for dependence up to 100km and 500km.

In keeping with previous work on PROGRESA (Behrman et al., 2011; Schultz, 2004; Skoufias and Parker, 2001), we include a rich set of controls in order to obtain more precise estimates of the treatment effects and account for some significant differences across treatment and control villages that exist despite the randomization (see Table D3). All of our specifications include controls for individual gender, household size, household head age, household head gender, a set of household composition variables (specified in Table D3), as well as locality controls for water source type, garbage disposal methods, the existence of a public phone, hospital or health center, and a DICONSA store (nutritional supplement distributor) in the locality.

3.3 Exogeneity of Rainfall and PROGRESA

PROGRESA exposure, the rainfall shock variable, and their interaction form the basis of our empirical specification. To provide support for the exogeneity of these variables, we check whether individuals are observably different across PROGRESA treatment and control villages, as well as rainfall shock versus normal rainfall groups. In Appendix Table D4, we regress each of the individual, household, and village-level characteristics that we use as control variables on a PROGRESA treatment village dummy, the rainfall shock, and their interaction. The vast majority of coefficients are statistically insignificant and/or small in magnitude relative to the means.⁹

Finally, we note some important considerations with respect to the interpretation of the rainfall shock coefficient (β_1). This coefficient provides the reduced-form effect of an early-life income shock on child outcomes in 2003. This includes any direct, biological effect the shock may have on a child's health and human capital, in addition to any changes resulting from compensating or reinforcing

⁹One exception is age, but as we discuss in section C.5 and show in Table C7, these age imbalances do not appear to be driving our main results.

investments that parents may make in response to the shock. Similarly, the coefficient on the interaction term (β_3) indicates whether there is any heterogeneity in the effect of PROGRESA with respect to this reduced-form shock – that is, how the effect of PROGRESA differed for children who experienced an income shock early in life along with any behavioral responses that resulted from this shock. Because children who experience income shocks do not experience these shocks in isolation, we argue this is a policy-relevant parameter of interest.

That said, in order to inform the generalizability of our findings, it would be useful to know whether parents are indeed responding to these early-life income shocks in ways that could in turn influence the effectiveness of the PROGRESA program. For example, parents could adjust their labor supply if they have a child who is less healthy due to an early-life income shock. They could also reallocate resources across siblings. In Table D5, we find no evidence of this. We regress indicators for parental employment, days worked by each parent, and hours worked by each parent in the baseline survey on the child-specific rainfall shock variable of interest (R_{slt}). We also examine average educational attainment among siblings, as well as average grade completion among siblings (which better adjusts for age) from the baseline survey, and find no significant differences across children who experienced and did not experience a rainfall shock at birth. In sum, parental responses to at-birth rainfall shocks are not large, at least in these dimensions we are able to observe.

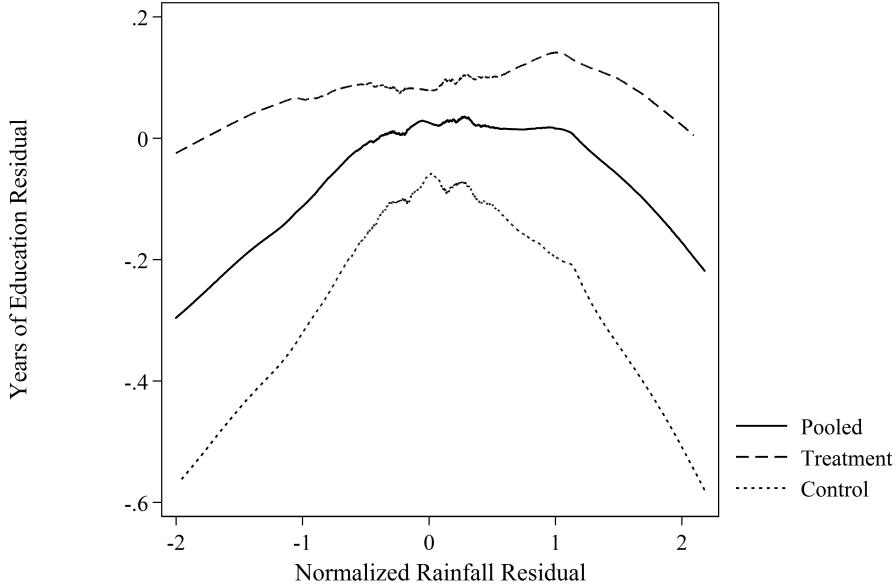
4 Results

4.1 Education Results

Figure 1 illustrates the intuition underlying our identification strategy, using lowess smoothing to depict the non-monotonic relationship between rainfall at birth and educational attainment. We first regress educational attainment and normalized rainfall on our full set of controls (state-by-birth year fixed effects, and all household and locality-level controls described in Section 3). We then plot non-parametrically the relationship between the educational attainment residuals on the y axis and the normalized rainfall residuals on the x axis, separately for treatment and control villages.

The dotted control group line has an inverted U- shape, which reinforces the idea that extreme deviations from mean rainfall are harmful for children. In addition, the treatment line is above the control line across the entire range of rainfall deviations, which is consistent with our summary

Figure 1: Years of Educational Attainment by Birth-Year Rainfall



Notes:

All three lines represent the lowess-smoothed educational attainment residuals for the relevant group. Educational attainment and normalized rainfall residuals are calculated after regressing each variable on state by birth-year fixed effects and the control variables described in section 3. Normalized rainfall residuals are trimmed at the 5th and 95th percentiles.

statistics and previous work on PROGRESA. Finally, the distance between the treatment and control lines is smallest around a normalized rainfall deviation of zero and grows larger in the tails, indicating that PROGRESA exposure mitigates the impacts of extreme rainfall at birth on educational attainment.

Table 4 reports parametric regression estimates analogous to the graphical analysis above. The first three columns show the regression results from our base specification (1), which includes state-by-year fixed effects and household and locality controls. For each coefficient of interest, we report three standard errors: first, clustered at the municipality level; second, allowing for spatial correlation using a 100km cutoff; and third, allowing for spatial correlation using a 500km cutoff.

For educational attainment in column 1, the main effect of PROGRESA is positive, the main effect of a rainfall shock is negative, and the interaction is positive; all are statistically significant at varying conventional levels. Since our sample includes children who may not have completed their schooling yet, we also look at the two other variables that adjust for age: similar patterns hold for grade progression and appropriate grade completion. In columns 2 and 3, the rainfall shock coefficients are negative and statistically significant; the main effects of PROGRESA are small and

Table 4: Effects of PROGRESA and Birth-Year Rainfall on Education Outcomes

| | (1) | (2) | (3) | (4) | (5) | (6) |
|-----------------------------------|--|--|--|---|---|--|
| | Educational Attainment | Grade Progression | Appropriate Grade Completion | Educational Attainment | Grade Progression | Appropriate Grade Completion |
| Years of PROGRESA Exposure | 0.10 (0.038)*** [0.030]*** {0.022}*** | 0.010 (0.0097) [0.0069] {0.0058}* * | 0.013 (0.0076) [0.0070]* {0.0068}* * | 0.015 (0.047) [0.035] {0.034} | -0.013 (0.012) [0.0087] {0.0077}* * | -0.011 (0.011) [0.0082] {0.0092} |
| Rainfall Shock | -0.65 (0.28)** [0.27]** {0.34}* * | -0.11 (0.056)** [0.058]* {0.065}* * | -0.12 (0.051)** [0.049]** {0.047}** * | -0.70 (0.27)*** [0.23]*** {0.25}*** * | -0.12 (0.057)** [0.048]** {0.046}** * | -0.14 (0.054)*** [0.048]*** {0.043}*** * |
| Rainfall Shock x Exposure | 0.11 (0.053)** [0.053]** {0.062}* * | 0.020 (0.011)* [0.012]* {0.013} * | 0.019 (0.010)* [0.010]* {0.0091}** * | 0.13 (0.051)** [0.044]*** {0.045}*** * | 0.024 (0.011)** [0.0095]** {0.0086}*** * | 0.025 (0.011)** [0.0096]*** {0.0081}*** * |
| Observations | 11824 | 11216 | 11824 | 11824 | 11216 | 11824 |
| Mean of Dependent Variable | 6.79 | 0.58 | 0.46 | 6.79 | 0.58 | 0.46 |
| Sample Ages (in 2003) | | | | 12 to 18 | | |
| Fixed Effects | | Birth year x state | | | Birth year x state, Municipality | |

Notes:

- Standard errors clustered at the municipality are reported in parentheses, Conley standard errors using a 100km cutoff are reported in square brackets, and Conley standard errors using a 500km cutoff are reported in curly brackets. (** p<0.01, ** p<0.05, * p<0.1).

- "Rainfall Shock" = 1 for individuals whose birth-year rainfall was more than one standard deviation from the 10-year historical locality-specific mean.

-All specifications include gender, household head gender and age, household size, household composition variables, parental education, parental language, and locality characteristics. Controls for parental language/education and locality distance include dummies for missing values.

only marginally significant, which means that PROGRESA had little effect on those who were not exposed to adverse rainfall. Importantly, however, the sums of the PROGRESA coefficient and the interaction term are positive and statistically significant, indicating large PROGRESA effects for those exposed to early-life disadvantage.

These coefficient estimates imply PROGRESA had larger effects on those disadvantaged at birth. For those who were not exposed to adverse rainfall, PROGRESA increased educational attainment by 0.11 years, but it increased educational attainment for those exposed to adverse rainfall by 0.22 years. Put differently, PROGRESA was able to generate substantial catch-up for individuals exposed to adverse rainfall. A negative rainfall shock decreased educational attainment by 0.65 years. However, one year of PROGRESA exposure mitigated this reduction by 0.11 years. At 2 years of exposure – the average difference between treatment and control exposure – the program mitigated 35% of the disadvantage caused by the rainfall shock at birth in years of completed schooling. For grade progression and appropriate grade completion, the figures are similarly high: 37% and 32%, respectively (all percentages calculated using the results in columns 1 to 3).

In the specification with municipality fixed effects (columns 4 to 6), the pattern of the results is the same: PROGRESA reduces the disadvantage generated by early-life rainfall. The effects of PROGRESA for both groups (i.e., the main PROGRESA coefficient and the sum of the coefficient and interaction) are close to zero, likely due to lack of variation in treatment and control status within municipalities. Although municipality fixed effects help control for location-specific unobservables on a finer level than state, the fact that over half of the municipalities consisted of either all treatment or all control villages reduces the amount of variation we can exploit. For this reason, we focus on the baseline specification for the remainder of the paper.

Table 5 examines schooling completion by grade. We create separate dummy variables for the completion of 3 to 12 grades of school and estimate specification 1 using these dummies as the dependent variables. In each column, we restrict the sample to individuals old enough to have completed the number of grades used in the dependent variable.

The impact of PROGRESA on completing grades 3 to 10 is positive and significant. The main effect of the rainfall shock is negative and significant starting in 7th grade. For grades below this, early life disadvantage does not seem to drive grade completion, possibly because the vast majority of our 12-18 year old sample have completed grades 3 (97%) to 6 (78%). Also starting in 7th grade

Table 5: Effects of PROGRESA and Birth-Year Rainfall on Schooling Completion by Grade

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) |
|-----------------------------------|----------------------------|------------------------------|-----------------------------|-----------------------------|----------------------------|---------------------------------------|----------------------------|---------------------------|-----------------------------|---------------------------------|
| | Primary School | | | | Junior High School | | | | High School | |
| | 3 grades | 4 grades | 5 grades | 6 grades | 7 grades | 8 grades | 9 grades | 10 grades | 11 grades | 12 grades |
| Years of PROGRESA Exposure | 0.0043 (0.0027) | 0.0095 (0.0035)*** | 0.013 (0.0048)*** | 0.018 (0.0058)*** | 0.023 (0.011)** | 0.014 (0.011) | 0.019 (0.011)* | 0.0090 (0.0065) | -0.00032 (0.0060) | -0.0038 (0.0076) |
| | [0.0019]** {0.0015}*** | [0.0026]*** {0.0017}*** | [0.0034]*** {0.0027}*** | [0.0040]*** {0.0034}*** | [0.0080]*** {0.0063}*** | [0.0087]* {0.0083}* [0.0071]*** | [0.010]* [0.0065] | [0.0065] | [0.0068] | [0.0098] |
| Rainfall Shock | 0.012 (0.020) | -0.0090 (0.028) | -0.031 (0.038) | -0.036 (0.047) | -0.20 (0.070)*** | -0.23 (0.072)*** | -0.25 (0.083)*** | -0.065 (0.052) | -0.072 (0.037)* | -0.10 (0.054)* |
| | [0.019] {0.021} | [0.028] {0.034} | [0.034] {0.035} | [0.042] {0.048} | [0.064]*** {0.064}*** | [0.069]*** {0.067}*** | [0.076]*** {0.089}*** | [0.044] [0.042] | [0.035]** [0.029]** | [0.056]* {0.053}* [0.017] |
| Rainfall Shock x Exposure | -0.0020 (0.0040) | 0.0025 (0.0054) | 0.0052 (0.0072) | 0.0047 (0.0090) | 0.032 (0.014)* | 0.040 (0.013)*** | 0.046 (0.016)*** | 0.010 (0.011) | 0.0059 (0.0071) | 0.018 (0.017) |
| | [0.0037] {0.0040} | [0.0053] {0.0061} | [0.0066] {0.0067} | [0.0080] {0.0087} | [0.013]** {0.012}*** | [0.014]*** {0.013}*** | [0.015]*** {0.017}*** | [0.0096] [0.0086] | [0.0078] [0.0055] | [0.018] {0.017} |
| Observations | 11824 | 11824 | 11824 | 11824 | 10068 | 8285 | 6618 | 5002 | 3231 | 1592 |
| Mean of Dependent Variable | 0.97 | 0.93 | 0.88 | 0.78 | 0.56 | 0.52 | 0.45 | 0.14 | 0.097 | 0.058 |
| Sample Ages (in 2003) | 12 to 18 | 12 to 18 | 12 to 18 | 12 to 18 | 13 to 18 | 14 to 18 | 15 to 18 | 16 to 18 | 17 to 18 | 18 |
| Fixed Effects | Birth year x state | | | | | | | | | |

Notes:

- Standard errors clustered at the municipality are reported in parentheses, Conley standard errors using a 100km cutoff are reported in square brackets, and Conley standard errors using a 500km cutoff are reported in curly brackets. (** p<0.01, ** p<0.05, * p<0.1).

- "Rainfall Shock" = 1 for individuals whose birth-year rainfall was more than one standard deviation from the 10-year historical locality-specific mean.

- All specifications include gender, household head gender and age, household size, household composition variables, parental education, parental language, and locality characteristics. Controls for parental language/education and locality distance include dummies for missing values.

(and until 9th grade), there are significant positive interaction coefficients. As in Table 4, these interaction terms are at least as large as the main effects of PROGRESA, implying PROGRESA effects that are double the size for those who experienced adverse early life rainfall compared to those who did not.

Another important related outcome is cognitive ability, but as we explain in Appendix section C.1, the evidence we have on this is somewhat inconclusive. Our regressions reveal no significant effects of PROGRESA, rainfall, or their interaction on Woodcock-Johnson language and math scores for the sub-sample that was tested.

4.2 Employment Outcomes

Table 6: Effects of PROGRESA and Birth-Year Rainfall on Longer-Term Outcomes

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) |
|-----------------------------------|---|---|--|--|--|---|--|
| | Currently Enrolled w/ HS Degree | Worked this Week | Worked this Year | Worked in Non-Laborer Job | Enrolled or Currently Working | Enrolled or Worked this Year | Enrolled or Worked in Non- Laborer Job |
| Years of PROGRESA Exposure | -0.0049 (0.0078) [0.0096] {0.0054} | 0.0010 (0.013) [0.013] {0.0064} | 0.0031 (0.012) [0.012] {0.0066} | -0.014 (0.016) [0.012] {0.0088} | 0.0019 (0.013) [0.014] {0.0066} | 0.0034 (0.012) [0.012] {0.0068} | -0.014 (0.015) [0.014] {0.0095} |
| Rainfall Shock | -0.10 (0.053)* [0.053]* {0.047}** | -0.21 (0.15) [0.15] {0.17} | -0.17 (0.13) [0.10]* {0.059}*** | -0.22 (0.13)* [0.096]** {0.072}*** | -0.21 (0.15) [0.16] {0.19} | -0.21 (0.14) [0.11]* {0.085}** | -0.26 (0.13)* [0.10]** {0.063}*** |
| Rainfall Shock x Exposure | 0.017 (0.017) [0.017] {0.016} | 0.087 (0.044)** [0.043]** {0.046}* 18 | 0.077 (0.039)* [0.030]** {0.018}*** | 0.099 (0.040)** [0.031]*** {0.023}*** | 0.079 (0.046)* [0.049] {0.055} | 0.078 (0.044)* [0.037]** {0.031}** | 0.100 (0.042)** [0.037]*** {0.025}*** |
| Observations | 1597 | 1147 | 1143 | 1143 | 1145 | 1139 | 1138 |
| Mean of Dependent Variable | 0.061 | 0.50 | 0.53 | 0.35 | 0.56 | 0.59 | 0.41 |
| Sample Ages (in 2003) | | | | | Birth year x state | | |
| Fixed Effects | | | | | | | |

Notes:

- Standard errors clustered at the municipality are reported in parentheses, Conley standard errors using a 100km cutoff are reported in square brackets, and Conley standard errors using a 500km cutoff are reported in curly brackets. (** p<0.01, ** p<0.05, * p<0.1).

"Rainfall shock" = 1 for individuals whose birth-year rainfall was more than one standard deviation from the 10-year historical locality-specific mean.

All specifications include gender, household head gender and age, household size, household composition variables, parental education, parental language, and locality characteristics. Controls for parental language/education and locality distance include dummies for missing values.

We are also interested in whether rainfall shocks and PROGRESA exposure have similar effects

on longer-run labor market outcomes that are not directly tied to the PROGRESA cash incentive. Unfortunately, much of our sample is too young for us to study impacts on their employment outcomes, but the oldest cohort – who were 18 at the time of the 2003 survey – were just old enough to be graduating from high school and pursuing either further education or formal employment. About 30% of the 18-year-olds in the 2003 survey were no longer living at home (see Appendix Figure D1) and therefore missing detailed employment information, but as we show in column 4 of Appendix Table C3, the likelihood of a missing employment variable in this sample is not driven by PROGRESA, rainfall, or their interaction.¹⁰ In Table 6, we report the results of regressions on variables related to continuing education and employment after high school for this 18-year-old sample.

Our first dependent variable of interest is the continuation of education after high school: this is an indicator equal to 1 if an individual is enrolled in school (including college or vocational training) and has already completed 12 grades of school. In columns 2 and 3, we create dummies for employment in the week of survey and in the past year. Column 4 attempts to separate those employed in lower-skilled, intermittent jobs from the pool of employed individuals by using an indicator equal to 1 if an individual was employed and worked in a non-laborer job (and 0 if unemployed or a spot laborer). In the last 3 columns, we combine previous outcomes to create indicators for individuals either still in school or working. For instance, the dependent variable in column 5 is an indicator equal to 1 if individuals report either being currently enrolled or having worked that week.

An important takeaway from this table is the consistent pattern of coefficients across all columns: PROGRESA effects are either close to zero or positive, adverse rainfall effects are negative, and interaction terms are all positive. While none of the main effects of PROGRESA are statistically significant, the sum of this coefficient and the interaction term is positive and significant in columns 2 through 7. This indicates, as with the education outcomes, PROGRESA has statistically significant employment effects on those exposed to adverse rainfall at birth. Taken in sum, these findings illustrate the ability of school-aged CCT programs to offset the impacts of insults in early life, in dimensions not limited to school-aged outcomes directly incentivized by the program.

¹⁰The fraction living outside of the household grows even higher after age 18, which is why we do not examine those older than 18 in 2003.

4.3 Robustness Checks

We run a number of checks to address concerns about selective fertility, attrition, migration, and imbalances across treatment and control villages. We discuss these in detail in Appendix section C. In short, we find no evidence that PROGRESA or birth-year rainfall shocks affected fertility (Table C2) or attrition (Table C3). Rainfall shocks do not appear to be correlated with various migration-related outcomes (Table C4). Our results are robust to the use of a trimmed sample that addresses the imbalance in the rainfall shock variable across treatment and control (Tables C5), as well as specifications that address the imbalances in other characteristics, including age (Tables C6 and C7).

We also investigate a number of alternate definitions of our main variables of interest. We find that our conclusions remain the same when we use a simple treatment indicator for PROGRESA (Table C8) . In Table C9, we use a more flexible specification for our rainfall shock variable, allowing for floods and droughts to have different sized effects. We find that floods and droughts both have negative effects that are mitigated by PROGRESA exposure: the coefficients on droughts tend to be slightly larger in magnitude (though not significantly different from) the coefficients on floods. Finally, when we include controls for rainfall shocks in the year before birth, the second year of life, and the third year of life (as well as their interactions with PROGRESA), we only see consistently significant main effects and interaction effects on rainfall in the year of birth. All of these tables are discussed in more detail in Appendix section C.

4.4 Mechanisms

Having documented that negative rainfall shocks at birth affect educational attainment and employment outcomes, and that this effect is reduced by PROGRESA exposure, we now discuss why this might be the case.

One possibility is that parents have inequality averse preferences and when one child is disadvantaged (due to an income shock in their year of birth), they reallocate resources from other children. When schooling becomes more affordable due to a program like PROGRESA, this could result in parents choosing to increase the educational attainment for the disadvantaged child by more than for their other children.

It is difficult to identify the extent to which these kind of preferences exist. However, the evidence we do have does not provide strong support for this possibility. First, Appendix Table D5 finds no evidence that a child's exposure to a rainfall shock at birth affects the educational attainment of their siblings (in the absence of the PROGRESA program). Adding to this, the results in Appendix Table D6 show that a child's educational outcomes do not appear to be affected by their siblings' exposure to rainfall shocks at birth, nor the interaction between sibling rainfall and PROGRESA exposure. Interestingly, for the 18-year-old sample, we find that sibling rainfall shocks increase the probability of work and reduce the effect of PROGRESA exposure, which could indicate that the early-life experiences of children do affect the employment decisions of their siblings – particularly older ones. However, the coefficients on the main variables of interest (own rainfall, own PROGRESA exposure, and their interaction) are very similar to those estimated in the original specifications, suggesting this does not explain our main empirical findings. In these same regressions, we also explore how child outcomes are affected by sibling exposure to PROGRESA. Across all outcomes, there is little evidence that sibling exposure to PROGRESA (or its interaction with the child rainfall shock variable) affects education or employment outcomes.

While these results do not necessarily rule out resource reallocation across siblings as a mechanism for our findings, we are also interested in what could be driving our results in the absence of inequality averse preferences. That is, in a model that abstracts away from parental preferences regarding comparisons between their children, is there anything that could explain why PROGRESA improves educational attainment more for children born in years of adverse rainfall?

To answer this question, we extend the canonical schooling choice model in Card (2001) by allowing individuals to have heterogeneous initial endowments that affect future earnings. That is, the earnings function at period t is given by $y(\omega, S, t)$, which depends not only on years of schooling S but also the initial endowment ω . We describe the model in detail in Appendix section E and summarize the main implications in this section.

4.4.1 Rainfall as a Shock to Endowments

Rainfall shocks are incorporated into the model as a shock to the initial endowment. This is based on the evidence in Appendix Tables B1 and B2, which show that negative rainfall shocks at birth increase stunting and behavioural problems. We acknowledge that rainfall shocks in one year could

affect the income-generating abilities of households in subsequent years, but we argue that the primary effects of birth-year rainfall shocks are concentrated in the first few years of life.¹¹

We assume that rainfall shocks at birth do not affect current household income. This is supported by the evidence in Appendix Table D7, which shows that individuals who experience rainfall shocks at birth are not more or less likely to be classified as poor in 1997. In column 2, we also show that rainfall at birth is not significantly related to current household income.

In Appendix Table D8, we provide further support for the argument that early-life rainfall shocks capture a phenomenon that is distinct from contemporaneous household disadvantage. We estimate a regression that adds to our main specification a measure of household income and its interaction with the PROGRESA exposure variable. Specifically, we use the “poverty score” (which is increasing in household income) that is used to determine program eligibility measure in 1997. The results of this exercise show that the coefficients on our variables of interest (PROGRESA exposure, early life rainfall, and their interaction) are almost identical to the baseline results. In other words, the ability of PROGRESA to remediate early-life disadvantage is separate from any heterogeneous effects based on current household income. This is because current household income appears to be orthogonal to rainfall at birth (as indicated by the similarity between the rainfall-related coefficients in Appendix Table D8 and the corresponding ones in Tables 4 to 6). While there does appear to be some heterogeneity in the effect of PROGRESA by current household income for a subset of outcomes, this is separate from the heterogeneity based on early-life rainfall and does not affect our estimates of PROGRESA’s ability to remediate for early-life disadvantage.

4.4.2 Theoretical Mechanisms for Remediation

Returning to the model, we assume that individuals have an infinite time horizon, attend school during the first S periods of life, and work full-time for the rest of it. While in school, the utility in period t depends on the level of consumption, $u(c(t))$, and the effort cost for the t -th year of schooling, $\phi(\omega, t)$. As we show in Appendix section E, this model allows us to predict how the optimal level of schooling should vary with the initial endowment and with a program like

¹¹Bobonis (2009), for example, finds that household expenditures are affected by rainfall shocks in the previous year. Serial correlation would also imply that rainfall shocks in one year could lead to income effects in subsequent years, but – like other papers that test for serial correlation in rainfall shocks (Kaur, 2014; Shah and Steinberg, 2017) – we do not find that our rainfall shocks are serially correlated over time.

PROGRESA that offsets the cost of schooling. Importantly, the model also provides a mathematical expression describing how the effect of PROGRESA on optimal schooling will vary with the initial endowment. An inspection of this expression helps shed light on the primary mechanisms that could drive remediation.

First, the value of the PROGRESA transfer represents a larger proportion of foregone wages for low endowment individuals compared to high endowment individuals, as low endowment individuals have lower income potential, leading to a larger schooling response to the PROGRESA incentive among low endowment individuals. Second, because high endowment individuals obtain more schooling than do their low endowment counterparts in the absence of the PROGRESA incentive, it would be more difficult for a program like PROGRESA to increase the schooling of high-endowment individuals (vis-a-vis low-endowment individuals) if effort costs are convex in schooling levels. Finally, the shape of the earnings function also plays a role. If the initial endowment and schooling are substitutes in the production function and the marginal returns to schooling increase faster with the initial level of endowment, this would also contribute to remediation. This is because these two conditions imply that the rate at which the benefit of studying an extra year decreases is faster for high endowment individuals.

5 Conclusion

In this paper, we leverage the combination of two sources of exogenous variation – in early life circumstance and costs of schooling during childhood – to study whether (and the extent to which) it is possible to mitigate the impact of early life shocks. We find that a negative shock to early-life circumstance (adverse rainfall) lowers educational attainment and employment probabilities by young adulthood. However, exposure to the PROGRESA program helps mitigate these negative effects, indicating that remediation of early-life shocks is possible through government programs later in life. The magnitude of the interaction term is telling: in most cases, it ranges between 15% to 40% of the size of the main effect of rainfall, suggesting that cash transfer programs like PROGRESA have the potential to offset almost entirely the inequality generated by early life circumstances.

This study contributes to the large literature evaluating PROGRESA, and more specifically,

to our knowledge about the program's ability to mitigate shocks. Two studies investigate the ability of PROGRESA to mitigate for contemporaneous weather shocks and find mixed results. De Janvry et al. (2006), who also focus on the education component of the program, finds that PROGRESA protects school enrollment from falling in response to contemporaneous weather-related income shocks. Aguilar and Vicarelli (2011), on the other hand, find no evidence that PROGRESA mitigated the negative health effects of El Nino flooding on young children, for whom the health component of the program was most relevant.

Our results also speak to the literature on cash transfer programs more generally (Behrman et al., 2011; Blattman et al., 2013; Haushofer and Shapiro, 2013; Schultz, 2004). While most evaluations of such programs tend to focus on average effects, we compare impacts across individuals with different early life experiences and find PROGRESA had a larger impact on those who experienced negative shocks early in life.

References

- Aguilar, A. and Vicarelli, M. (2011). El nino and mexican children: medium-term effects of early-life weather shocks on cognitive and health outcomes. *Cambridge, United States: Harvard University, Department of Economics. Manuscript.*
- Aizer, A., Eli, S., Ferrie, J., and Lleras-Muney, A. (2016). The long-run impact of cash transfers to poor families. *American Economic Review*, 106(4):935–71.
- Almond, D. and Currie, J. (2011). Killing me softly: The fetal origins hypothesis. *The Journal of Economic Perspectives*, 25(3):153–172.
- Almond, D., Currie, J., and Duque, V. (2018). Childhood circumstances and adult outcomes: Act ii. *Journal of Economic Literature*, 56(4):1360–1446.
- Almond, D. and Mazumder, B. (2013). Fetal origins and parental responses. *Annu. Rev. Econ.*, 5(1):37–56.
- Bailey, M. J., Hoynes, H. W., Rossin-Slater, M., and Walker, R. (2020). Is the social safety net a long-term investment? large-scale evidence from the food stamps program. Technical report, National Bureau of Economic Research.
- Behrman, J. R., Parker, S. W., and Todd, P. E. (2011). Do conditional cash transfers for schooling generate lasting benefits? a five-year followup of progresita/oportunidades. *Journal of Human Resources*, 46(1):93–122.
- Bitler, M. P., Hoynes, H. W., and Domina, T. (2014). Experimental evidence on distributional effects of head start. Technical Report 20434, National Bureau of Economic Research.
- Blattman, C., Fiala, N., and Martinez, S. (2013). The economic and social returns to cash transfers: Evidence from a ugandan aid program. Technical report, CEGA Working Paper.
- Bobonis, G. J. (2009). Is the allocation of resources within the household efficient? new evidence from a randomized experiment. *Journal of Political Economy*, 117(3):453–503.
- Bundervoet, T., Verwimp, P., and Akresh, R. (2009). Health and civil war in rural burundi. *Journal of human Resources*, 44(2):536–563.
- Card, D. (2001). Estimating the return to schooling: Progress on some persistent econometric problems. *Econometrica*, 69(5):1127–1160.
- Chetty, R., Hendren, N., and Katz, L. F. (2016). The effects of exposure to better neighborhoods on children: New evidence from the moving to opportunity experiment. *American Economic Review*, 106(4):855–902.
- Conley, T. (1999). Gmm estimation with cross sectional dependence. *Journal of Econometrics*, 92(1):1–45.
- Cornelissen, T., Dustmann, C., Raute, A., and Schönberg, U. (2018). Who benefits from universal child care? estimating marginal returns to early child care attendance. *Journal of Political Economy*, 126(6):2356–2409.

- Crookston, B. T., Penny, M. E., Alder, S. C., Dickerson, T. T., Merrill, R. M., Stanford, J. B., Porucznik, C. A., and Dearden, K. A. (2010). Children who recover from early stunting and children who are not stunted demonstrate similar levels of cognition. *The Journal of nutrition*, 140(11):1996–2001.
- Crookston, B. T., Schott, W., Cueto, S., Dearden, K. A., Engle, P., Georgiadis, A., Lundein, E. A., Penny, M. E., Stein, A. D., and Behrman, J. R. (2013). Postinfancy growth, schooling, and cognitive achievement: Young lives. *The American journal of clinical nutrition*, 98(6):1555–1563.
- Currie, J. and Vogl, T. (2012). Early-life health and adult circumstance in developing countries. *Annual Review of Economics*, 5:1–36.
- Das, J., Do, Q.-T., and Özler, B. (2005). Reassessing conditional cash transfer programs. *The World Bank Research Observer*, 20(1):57–80.
- De Janvry, A., Finan, F., Sadoulet, E., and Vakis, R. (2006). Can conditional cash transfer programs serve as safety nets in keeping children at school and from working when exposed to shocks? *Journal of Development Economics*, 79(2):349–373.
- Dercon, S. and Porter, C. (2014). Live aid revisited: long-term impacts of the 1984 ethiopian famine on children. *Journal of the European Economic Association*, 12(4):927–948.
- Dinkelman, T. (2017). Long-run health repercussions of drought shocks: evidence from south african homelands. *The Economic Journal*, 127(604):1906–1939.
- Duque, V. (2017). Early-life conditions and child development: Evidence from a violent conflict. *SSM-population health*, 3:121–131.
- Duque, V., Rosales-Rueda, M., Sanchez, F., et al. (2018). How do early-life shocks interact with subsequent human-capital investments? evidence from administrative data. In *IZA World of Labor Conference*.
- Gertler, P., Heckman, J., Pinto, R., Zanolini, A., Vermeersch, C., Walker, S., Chang, S. M., and Grantham-McGregor, S. (2014). Labor market returns to an early childhood stimulation intervention in jamaica. *Science*, 344(6187):998–1001.
- Goyal, D., Gay, C., and Lee, K. A. (2010). How much does low socioeconomic status increase the risk of prenatal and postpartum depressive symptoms in first-time mothers? *Women's Health Issues*, 20(2):96–104.
- Gunnsteinsson, S., Molina, T., Adhvaryu, A., Christian, P., Labrique, A., Sugimoto, J., Shamim, A. A., and West Jr, K. P. (2021). Protecting infants from natural disasters: The case of vitamin a supplementation and a tornado in bangladesh. Technical report.
- Haushofer, J. and Shapiro, J. (2013). Household response to income changes: Evidence from an unconditional cash transfer program in kenya. Technical report.
- Heckman, J. J. (2006). Skill formation and the economics of investing in disadvantaged children. *Science*, 312(5782):1900–1902.
- Heckman, J. J. (2007). The economics, technology, and neuroscience of human capability formation. *Proceedings of the national Academy of Sciences*, 104(33):13250–13255.

- Hendren, N. and Sprung-Keyser, B. (2020). A unified welfare analysis of government policies. *The Quarterly Journal of Economics*, 135(3):1209–1318.
- Hjort, J. et al. (2017). Universal investment in infants and long-run health: evidence from denmark's 1937 home visiting program. *American Economic Journal: Applied Economics*, 9(4):78–104.
- Hoddinott, J. and Skoufias, E. (2004). The impact of progesa on food consumption. *Economic development and cultural change*, 53(1):37–61.
- Hoynes, H., Schanzenbach, D. W., and Almond, D. (2016). Long-run impacts of childhood access to the safety net. *American Economic Review*, 106(4):903–34.
- Johnson, R. C. and Jackson, C. K. (2019). Reducing inequality through dynamic complementarity: Evidence from head start and public school spending. *American Economic Journal: Economic Policy*, 11(4):310–49.
- Kaur, S. (2014). Nominal wage rigidity in village labor markets. Technical report, National Bureau of Economic Research.
- Kim, P., Capistrano, C. G., Erhart, A., Gray-Schiff, R., and Xu, N. (2017). Socioeconomic disadvantage, neural responses to infant emotions, and emotional availability among first-time new mothers. *Behavioural brain research*, 325:188–196.
- Lagarde, M., Haines, A., and Palmer, N. (2007). Conditional cash transfers for improving uptake of health interventions in low-and middle-income countries: a systematic review. *Jama*, 298(16):1900–1910.
- Lundein, E. A., Behrman, J. R., Crookston, B. T., Dearden, K. A., Engle, P., Georgiadis, A., Penny, M. E., and Stein, A. D. (2014). Growth faltering and recovery in children aged 1–8 years in four low-and middle-income countries: Young lives. *Public health nutrition*, 17(9):2131–2137.
- Maccini, S. and Yang, D. (2009). Under the weather: Health, schooling, and economic consequences of early-life rainfall. *American Economic Review*, 99(3):1006–1026.
- Malamud, O., Pop-Eleches, C., and Urquiola, M. S. (2016). Interactions between family and school environments: Evidence on dynamic complementarities? *NBER Working Paper*, (w22112).
- Martorell, R., Khan, L. K., and Schroeder, D. G. (1994). Reversibility of stunting: epidemiological findings in children from developing countries. *European journal of clinical nutrition*, 48:S45–57.
- Parker, S. W., Todd, P. E., et al. (2017). Conditional cash transfers: The case of progesa/oportunidades. *Journal of Economic Literature*, 55(3):866–915.
- Paxson, C. H. (1992). Using weather variability to estimate the response of savings to transitory income in thailand. *American Economic Review*, 82(1):15–33.
- Prentice, A. M., Ward, K. A., Goldberg, G. R., Jarjour, L. M., Moore, S. E., Fulford, A. J., and Prentice, A. (2013). Critical windows for nutritional interventions against stunting. *The American journal of clinical nutrition*, 97(5):911–918.
- Rossin-Slater, M. and Wüst, M. (2020). What is the added value of preschool for poor children? long-term and intergenerational impacts and interactions with an infant health intervention. *American Economic Journal: Applied Economics*, 12(3):255–86.

- Schultz, T. P. (2004). School subsidies for the poor: evaluating the mexican progresas poverty program. *Journal of Development Economics*, 74(1):199–250.
- Shah, M. and Steinberg, B. M. (2017). Drought of opportunities: Contemporaneous and long-term impacts of rainfall shocks on human capital. *Journal of Political Economy*, 125(2):527–561.
- Skoufias, E. (2005). Progresa and its impacts on the welfare of rural households in mexico. Technical Report 139, INTERNATIONAL FOOD POLICY RESEARCH INSTITUTE.
- Skoufias, E. and Parker, S. W. (2001). Conditional cash transfers and their impact on child work and schooling: Evidence from the progresas program in mexico. *Economia*, 2(1):45–96.
- Victora, C. G., Adair, L., Fall, C., Hallal, P. C., Martorell, R., Richter, L., Sachdev, H. S., Maternal, Group, C. U. S., et al. (2008). Maternal and child undernutrition: consequences for adult health and human capital. *The lancet*, 371(9609):340–357.
- Wolpin, K. I. (1982). A new test of the permanent income hypothesis: the impact of weather on the income and consumption of farm households in india. *International Economic Review*, pages 583–594.

ONLINE APPENDIX (Not for Publication)

A Additional Program Details

A.1 Health Component

The health and nutrition component of the program involved conditional cash transfers intended to incentivize healthy behaviors. For instance, in order for a household to receive a cash grant for food, all members were required to visit the health facility a specified number of times per year and to attend nutrition and health education lectures. The required number of visits varied by age and gender, with pregnant women and infants required to go every 1-3 months, while anyone aged 5 and older only required to attend 1 or 2 times per year. The program also provided nutrition supplements and other preventative care for pregnant and lactating mothers and young children, and supported the improved provision of primary health care services in PROGRESA localities.

A.2 Program Targeting

PROGRESA was targeted toward poor households in poor localities. To determine which localities would receive the program, a set of marginalized localities was identified using data from the 1990 and 1995 censuses. Within these selected localities, household-level eligibility for PROGRESA was determined based on the results of an income survey administered to all households in each locality. First, household per capita income (excluding child income) was calculated, and households were categorized as above or below a poverty line. Then, separately for each region, the program identified the household characteristics that were the best predictors of poverty status, which were then used to construct the index that ultimately classified households as poor (eligible for PROGRESA) or nonpoor.¹

¹Skoufias et al. (2001) contains more details about the selection of localities and households.

B Rainfall Shocks and Health

In this section, we discuss evidence on the effects of rainfall shocks on BMI, stunting, cognitive test scores, and behavioral measures, using various age cohorts from the 2003 and 2007 surveys.

To study effects on BMI, we pool all individuals for whom BMI was measured across the 2003 and 2007 surveys.² For each individual, we calculate gender- and age-specific BMI z-scores using WHO tables,³ and regress this variable on a rainfall shock dummy (more than one standard deviation from the locality-specific historical mean) in the individual's locality of residence in the relevant survey year (controlling for state-by-survey-year fixed effects and a host of other individual and household-level controls, described in the table notes). In column 1 of Table B1, we see that adverse rainfall in the survey year has negative effects on BMI for the entire sample. This supports the idea that the lower wages and expenditures that result from bad rainfall also translate into lower nutritional intake. In column 2, we show that this result persists (and is much larger) for children under two years old, for whom these measurements are a closer proxy to their initial health endowment. Because BMI is not a conventional measure for young children, we also explore weight-for-length and weight-for-age z-scores in columns 3 and 4. We see that rainfall shocks have a significant negative effect on weight-for-length. In sum, this provides us with evidence that rainfall around the time of birth affects the nutritional intake and therefore BMI of infants.

We next ask whether these contemporaneous nutrition effects have longer-term implications for child health. To answer this question, we use height data, collected for children aged 0-2 in 2007, aged 2-6 in 2003, and aged 8-10 in 2007. We calculate age- and gender-specific height z-scores (once again using WHO tables) and create an indicator for stunted children, with heights falling more than 2 standard deviations below their group-specific mean. We then regress this indicator on the rainfall shock variable that we use in our main analysis. In columns 6 and 7 of Table B1, we see that there is a significant positive relationship between bad rainfall and stunting for children aged 2 and older. In other words, year of birth rainfall

²Specifically, height and weight were measured for sub-samples of children aged 2 to 6 in 2003, adolescents aged 15 to 21 in 2003, infants aged 0 to 2 in 2007, children aged 8 to 10 in 2007, and adults aged 30 and older in 2007.

³We use the means and standard deviations for 20-year-olds, the oldest available age category, for all older adults.

shocks have physical health effects that persist into early childhood.

Table B1: Effect of Rainfall on Weight and Height Outcomes

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) |
|------------------------------------|--|-------------------|---------------------------|------------------------|---------------------|--------------------|--------------------|
| | BMI z-score | BMI z-score | Weight-for-length z-score | Weight-for-age z-score | Stunted | Stunted | Stunted |
| Rainfall Shock (in survey year) | -0.058 (0.033)* | -0.16 (0.082)* | -0.19 (0.098)* | 0.067 (0.099) | | | |
| Rainfall Shock (in birth year) | | | | | -0.00056 (0.026) | 0.042 (0.019)** | 0.037 (0.019)** |
| Observations | 9596 | 1184 | 1184 | 1187 | 1243 | 1978 | 1426 |
| Mean of Dependent Variable | 0.56 | 0.55 | -0.38 | 0.26 | 0.19 | 0.22 | 0.087 |
| Sample Ages (in Survey Year) | All | 0-1 | 0-1 | 0-1 | 0-2 | 2-6 | 8-10 |
| Survey year(s) | 2003; 2007 | 2007 | 2007 | 2007 | 2007 | 2003 | 2007 |
| Fixed Effects | Birth year, state, survey year x state | | | | Birth year x state | | |

Notes:

- Standard errors clustered at the municipality level are in parentheses (** p<0.01, ** p<0.05, * p<0.1).

- "Rainfall shock" = 1 for individuals whose survey year or birth-year rainfall was more than one standard deviation from the 10-year historical locality-specific mean.

-All specifications include gender, household head gender and age, household size, household composition variables, parental education, parental language, and locality characteristics. Controls for parental language/education and locality distance include dummies for missing values. For adults in column 1, all parental variables are missing.

-Column 1 includes all individuals whose height and weight were measured in either 2003 or 2007: 0-2 year-olds in 2007, 2-6 year-olds in 2003, 8-10 year-olds in 2007, 15-21 year-olds in 2003, adults 30 and older and mothers of young children in 2007.

Taking advantage of other measures of child development collected in 2003 (for 2-6 year-olds) and 2007 (for 8-10 year olds), we also explore whether other dimensions of human capital – cognitive and non-cognitive skills – are affected by birth-year rainfall. During the 2003 surveys, a number of cognitive development tests (Woodcock Johnson tests, Peabody Picture Vocabulary tests, and MacArthur communication tests) were administered to a sample of 2-6 year-olds. In addition, mothers were asked to rate their children's behaviors using the Achenbach Child Behavior Checklist. For cognitive measures, we calculate z-scores for each of the cognitive tests and take the mean across all cognitive z-scores. For the Achenbach checklist, we create a z-score after summing the responses to all checklist questions. In 2007, mothers of children aged 8-10 answered the Strengths and Difficulties Questionnaire (SDQ), a list of questions about the behaviors of their children. Using existing recommended methods for scoring and grouping questions, we create z-scores for externalizing problems, internalizing problems, and anti-social problems (and an overall z-score that averages all three).

Results are reported in Table B2, where we use the cognitive z-score (from 2003), the

behavioral z-score (from 2003), and multiple behavioral z-scores (from 2007) as our dependent variables, and run regressions identical to the ones in Table B1. We find that birth-year rainfall had no significant effects on cognitive or behavioral measures for 2 to 6 year-olds, but did increase the likelihood of behavioral problems (externalizing problems, in particular) later in childhood. That income shocks in the year of birth can affect non-cognitive development is consistent with the child development literature, which documents that socioeconomic disadvantage is associated with altered maternal responses to infant emotions (Kim et al., 2017) and with other reasons for negative mother-infant interactions that could lead to behavioral problems later in childhood (Goyal et al., 2010).

Table B2: Effect of Birth-Year Rainfall on Cognitive and Behavioral Outcomes in Childhood

| | (1) | (2) | (3) | (4) | (5) | (6) |
|-----------------------------------|---------------------------|------------------------------|-------------------|------------------------------|------------------------|-----------------|
| | Cognitive measure z-score | Personality measures z-score | Overall | Behavioral problems z-scores | | |
| | | | | Externalizing problems | Internalizing problems | Social problems |
| Rainfall Shock (in birth year) | 0.015 (0.039) | 0.0020 (0.023) | 0.083 (0.050)* | 0.12 (0.065)* | 0.017 (0.067) | 0.11 (0.071) |
| Observations | 2032 | 2014 | 1488 | 1488 | 1488 | 1488 |
| Mean of Dependent Variable | -0.052 | 0.034 | -0.013 | 0.014 | -0.0084 | -0.044 |
| Sample Ages (in Survey Year) | 2-6 | 2-6 | 8-10 | 8-10 | 8-10 | 8-10 |
| Survey year | 2003 | 2003 | 2007 | 2007 | 2007 | 2007 |
| Fixed Effects | | | | Birth year x state | | |

Notes:

- Standard errors clustered at the municipality level are in parentheses (** p<0.01, ** p<0.05, * p<0.1).

- "Rainfall shock" = 1 for individuals whose birth-year rainfall was more than one standard deviation from the 10-year historical locality-specific mean.

-All specifications include gender, household head gender and age, household size, household composition variables, parental education, parental language, and locality characteristics. Controls for parental language/education and locality distance include dummies for missing values

C Robustness Checks and Additional Outcomes

C.1 Cognitive Tests

Although we are also interested in whether cognitive ability, not just educational attainment, was impacted by PROGRESA and birth-year rainfall, the evidence on this is somewhat inconclusive. Woodcock-Johnson dictation, word identification, and applied problems tests were administered to a sub-sample of individuals aged 15 to 21, as part of the 2003 survey. Unfortunately, treatment status is significantly negatively related to the probability of an individual having a non-missing test score, and our main schooling results restricted to this sub-sample reveal smaller and imprecisely estimated coefficients, casting doubt on whether it is representative of our population of interest. Another issue is that the tests may have been unable to capture sufficient variation in cognitive ability: in the letter-word identification test, for example, almost 30% of the sample answered everything correctly (and over 50% only made 2 mistakes) in a test of 58 questions. With these caveats in mind, Appendix Table C1 reveals no significant effects of PROGRESA, rainfall, or their interaction on these test scores, consistent with previous work documenting a null main effect of PROGRESA (Behrman et al., 2009).

C.2 Selective Fertility

In Table C2 we investigate how PROGRESA and rainfall shocks may have affected fertility, which could lead to potential selection issues. One concern might be that negative rainfall shocks during a year may affect the number of children that are born and/or survive to school-aged years. If this were the case, the composition of individuals in our sample who were born in shock years would be different from those in our sample born in regular years. In order to check this, we collapse to the locality by birth year level and count the total number of children born in a particular year in each locality. We then use this constructed panel to regress the total number of children born that year on our rainfall shock. Column 1 of Table C2 reports results from this regression. We find no evidence of selective fertility or selective child mortality.

Table C1: Effects of PROGRESA and Birth-Year Rainfall on Woodcock-Johnson Test Scores

| | (1) | (2) | (3) | (4) |
|----------------------------|-------------------------------|---------------------|------------------|--------------------|
| | Letter Word Identification | Applied Problems | Dictation | Average Score |
| Years of PROGRESA Exposure | 0.0045 (0.040) | 0.014 (0.035) | 0.027 (0.045) | 0.013 (0.035) |
| Rainfall Shock | 0.14 (0.25) | -0.14 (0.25) | -0.19 (0.29) | -0.062 (0.22) |
| Rainfall Shock x Exposure | -0.060 (0.052) | 0.00039 (0.051) | 0.034 (0.059) | -0.0056 (0.045) |
| Observations | 1593 | 1586 | 1581 | 1571 |
| Sample Ages (in 2003) | | 15 to 18 | | |
| Fixed Effects | | Birth year x state | | |

Notes:

- Standard errors clustered at the municipality level are in parentheses (** p<0.01, ** p<0.05, * p<0.1).
- "Rainfall shock" = 1 for individuals whose birth-year rainfall was more than one standard deviation from the 10-year historical locality-specific mean.
- All specifications include gender, household head gender and age, household size, household composition variables, parental education, parental language, and locality characteristics. Controls for parental language/education and locality distance include dummies for missing values
- Scores are standardized by test type, and the average score in column 4 takes the average across all three z-scores.
- These tests were administered to a sample of individuals aged 15-21 in 2003, but we restrict to those aged 15-18 in order to remain consistent with the main sample.

Table C2: Effects of PROGRESA and Birth-Year Rainfall on Fertility

| | (1) | (2) | (3) | (4) | (5) | (6) |
|----------------------------|-------------------------------|----------------------------|---|-----------------------------------|----------------------------|---------------------------------------|
| | Locality-Level | | | Individual-Level | | |
| | Total number of children | Number of younger siblings | Birth spacing (in days) between younger sibling | Mother's ideal number of children | Mother wants more children | Number of additional children desired |
| Years of PROGRESA Exposure | | 0.0026 (0.0085) | -7.42 (10.1) | -0.14 (0.12) | 0.018 (0.011) | -0.027 (0.026) |
| Rainfall Shock | -0.090 (0.16) | -0.10 (0.066) | 29.1 (80.0) | -0.067 (0.52) | 0.089 (0.074) | 0.18 (0.17) |
| Rainfall Shock x Exposure | | 0.020 (0.013) | -6.37 (15.5) | 0.018 (0.10) | -0.022 (0.015) | -0.046 (0.034) |
| Observations | 2519 | 11686 | 7230 | 2057 | 2027 | 2091 |
| Mean of Dependent Variable | 4.83 | 1.98 | 1107.9 | 4.34 | 0.091 | 0.076 |
| Sample | Cohorts aged 12-18 in 2003 | | | Individuals aged 12-18 in 2003 | | |
| Fixed Effects | Birth year x state | | | Birth year x state | | |

Notes:

- Standard errors clustered at the municipality level are in parentheses (** p<0.01, ** p<0.05, * p<0.1).
- For locality-level analysis, the unit of observation is birth-year-locality.
- All specifications include locality controls and individual/household characteristics (gender, household head gender and age, household size, household composition, parental education and language). For the locality-level variables, these are averaged at the locality-birth-year level.
- "Rainfall shock" = 1 for individuals whose birth-year rainfall was more than one standard deviation from the 10-year historical locality-specific mean.
- The dependent variables in columns 4 to 6 are only available for a random subset of mothers.

Our next test is to check whether PROGRESA, rainfall shocks, and their interaction had any impact on mothers' subsequent fertility decisions. Specifically, we might be concerned that a good rainfall shock would increase the likelihood of having more children (or total fertility), or decrease the birth spacing between children, just as exposure to PROGRESA may do the same (by lowering the opportunity cost of having children). If this were the case, an individual's exposure to PROGRESA or rainfall shocks would also be related to intrahousehold allocation issues that may vary with the total number of siblings and spacing between siblings. To check for this, we estimate equation 1, again at the individual level, using number of younger siblings and birth spacing between next youngest sibling (in days) as dependent variables. The main effects and interaction effects in columns 2 and 3 are all insignificant.

In addition to investigating effects on actual fertility, we also ask whether PROGRESA or rainfall affected planned or expected fertility. To answer this question, we use questions on expected and desired fertility for the mothers of our sample children who were part of a detailed fertility questionnaire sub-sample. We do not find that rainfall shocks (or PROGRESA) affected the total number of desired children or the desire for additional children, as we show in the last three columns of Table C2.

C.3 Attrition

As in any longitudinal study, we must consider the extent to which selective attrition may be confounding our results. In Table C3, we show that although attrition between the baseline and 2003 surveys was sizeable, it appears to be uncorrelated with our regressors of interest. In this table, we simply regress various attrition indicators on years of PROGRESA exposure, the rainfall shock indicator, their interaction, and state by birth year fixed effects. In column 1, we investigate household attrition, including all eligible individuals in the baseline survey who would have been aged 12 to 18 in 2003. We do not find that our investment or endowment shocks influenced the likelihood of a household being dropped from the 2003 sample. In column 2, conditional on the household being found in 2003, we show that our regressors of interest do not significantly predict the likelihood of an individual being included in our

Table C3: Effects of PROGRESA and Birth-Year Rainfall on Attrition

| | (1) | (2) | (3) | (4) |
|----------------------------|----------------------------|---------------------------------------|--------------------------------------|---------------------------------------|
| | Household found in 2003 | Meets Data Quality Restrictions | Non-missing education variable | Non-missing employment variable |
| Years of PROGRESA Exposure | 0.0024 (0.0083) | 0.0015 (0.0032) | 0.0035 (0.0030) | 0.0065 (0.014) |
| Rainfall Shock | 0.034 (0.038) | 0.030 (0.023) | -0.0017 (0.019) | 0.13 (0.14) |
| Rainfall Shock x Exposure | -0.0065 (0.0076) | -0.0049 (0.0045) | -0.000028 (0.0037) | -0.039 (0.040) |
| Observations | 14525 | 12917 | 12159 | 1646 |
| Mean of Dependent Variable | 0.89 | 0.94 | 0.97 | 0.70 |
| Sample Ages (in 2003) | 12 to 18 | 12 to 18 | 12 to 18 | 18 |
| Fixed Effects | | | Birth year x state | |

Notes:

- Standard errors clustered at the municipality level are in parentheses (** p<0.01, ** p<0.05, * p<0.1).
- "Rainfall shock" = 1 for individuals whose birth-year rainfall was more than one standard deviation from the 10-year historical locality-specific mean.
- The sample in column 2 restricts to households found in 2003, while columns 3 and 4 restrict to those that meet data quality restrictions.

sample given the data quality restrictions we impose (matching genders across surveys and birth year differences of less than 2 years). Finally, in columns 3 and 4, we investigate whether the shocks predict the probability that an individual – who is found in 2003 and meets the data quality restrictions – has non-missing education and employment variables (restricting to 18-year-olds in column 4). We do not find any evidence of either.

C.4 Migration

In addition to selective fertility and attrition, selective migration in response to rainfall shocks could also be a concern. In particular, we might worry that permanent household-level migration responds to rainfall shocks, which would mean that year-of-birth rainfall shocks might affect the probability of an individual showing up in our PROGRESA localities in the first place. Unfortunately, we cannot study the migration behavior of households that never made it into our sample. What we can do is check whether rainfall shocks in 1997 affect the probability of a household participating in the 2003 survey: this six-year gap between

Table C4: Effects of Rainfall on Migration-Related Variables

| | (1) | (2) | (3) | (4) |
|-----------------------------------|--|----------------------------|--|--|
| | Household found in 2003 | Household found in 2003 | Father living in household in 1997 | Father living in household in 1997 |
| Rainfall Shock (in 1997) | 0.0047 (0.014) | 0.021 (0.016) | | |
| Rainfall Shock (in birth year) | | | -0.0025 (0.0070) | -0.0059 (0.0074) |
| Observations | 6684 | 6684 | 12156 | 12156 |
| Mean of Dependent Variable | 0.88 | 0.88 | 0.91 | 0.91 |
| Sample | Households with children aged 12-18 in 2003 | | Individuals aged 12-18 in 2003 | |
| Fixed Effects | None | Birth year x state | None | Birth year x state |

Notes:

- Standard errors clustered at the municipality level are in parentheses (** p<0.01, ** p<0.05, * p<0.1).
- "Rainfall shock" = 1 for individuals whose birth-year rainfall was more than one standard deviation from the 10-year historical locality-specific mean.
- Columns 1 and 2 are household-level regressions

the shock and migration outcome would capture more delayed migration responses, similar to what might have occurred for individuals whose household migrated before they started school. Although migration is not the only reason a household could be missing from the sample in 2003, it is likely the main one, and as we show in columns 1 and 2 of Table C4, we find no effects of 1997 rainfall shocks on this migration-related outcome.

Another potential issue is that rainfall shocks during an individual's year of birth might affect the temporary (rather than permanent) migration decisions of their parents. If a child grows up without a father in the household as a result, this could generate effects on their development separate from the mechanisms we have focused on in this paper. Therefore, in Table C4, we check to see whether rainfall shocks at birth affect the likelihood of a child's father being present in the household during the 1997 survey (the survey closest to the time of birth), and find no evidence of this (columns 3 and 4).

C.5 Balance

We investigate further the implications of the small but statistically significant imbalance in rainfall shock prevalence across PROGRESA treatment and control villages in our baseline

sample. To test whether our results are being driven by this imbalance, we repeat our analysis using the trimmed sample described in Section 2, in which rainfall shock prevalence is the same across treatment and control villages. This sample omits localities exhibiting shocks in every year, or no shocks in any year, over the study period. As Table C5 shows, our results are virtually identical to the full sample results.

Table C5: Effects of PROGRESA and Birth-Year Rainfall on Education and Employment Outcomes: Trimmed Sample

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) |
|----------------------------|------------------------|--------------------|------------------------------|---------------------|---------------------|---------------------|---------------------------------|---------------------------|---------------------------------------|
| | Educational Attainment | Grade Progression | Appropriate Grade Completion | Completed 7 grades | Completed 8 grades | Completed 9 grades | Currently Enrolled w/ HS Degree | Worked in Non-Laborer Job | Enrolled or Worked in Non-Laborer Job |
| Years of PROGRESA Exposure | 0.12 (0.043)*** | 0.0099 (0.011) | 0.014 (0.0087) | 0.028 (0.012)** | 0.017 (0.012) | 0.023 (0.013)* | -0.0060 (0.0090) | -0.021 (0.016) | -0.019 (0.015) |
| Rainfall Shock | -0.71 (0.29)** | -0.12 (0.058)** | -0.14 (0.053)** | -0.21 (0.071)*** | -0.25 (0.072)*** | -0.27 (0.085)*** | -0.096 (0.056)* | -0.32 (0.13)** | -0.35 (0.15)** |
| Rainfall Shock x Exposure | 0.12 (0.056)** | 0.022 (0.011)* | 0.021 (0.011)** | 0.033 (0.014)** | 0.043 (0.013)*** | 0.048 (0.016)*** | 0.016 (0.018) | 0.12 (0.041)*** | 0.12 (0.046)*** |
| Observations | 10236 | 9713 | 10236 | 8689 | 7160 | 5684 | 1320 | 966 | 962 |
| Mean of Dependent Variable | 6.78 | 0.59 | 0.47 | 0.56 | 0.51 | 0.45 | 0.065 | 0.35 | 0.41 |
| Sample Ages (in 2003) | 12 to 18 | 12 to 18 | 12 to 18 | 13 to 18 | 14 to 18 | 15 to 18 | 18 | 18 | 18 |
| Fixed Effects | | | | Birth year x state | | | | | |

Notes:

- Standard errors clustered at the municipality are reported in parentheses (** p<0.01, ** p<0.05, * p<0.1).

- "Rainfall shock" = 1 for individuals whose birth-year rainfall was more than one standard deviation from the 10-year historical locality-specific mean.

- All specifications include gender, household head gender and age, household size, household composition variables, parental education, parental language, and locality characteristics. Controls for parental language/education and locality distance include dummies for missing values

We also conduct a robustness exercise regarding the unbalanced demographic characteristics across treatment and control villages in Table D3. Table C6 reports the results of regressions on our main outcomes of interest, additionally controlling for interactions between the rainfall shock variable and each of the control variables that are not balanced across treatment and control groups. The results are once again very similar to the main results reported above.

Finally, we address the imbalance in age across treatment villages and rainfall shock groups, which was revealed by the balance checks conducted in Table D4. Because we include birth year fixed effects in our regressions, we are not concerned about these imbalances affecting the estimation of the PROGRESA and rainfall shock main effects. However, these imbalances could be affecting the estimation of the interaction effect, if there is any heterogeneity by

Table C6: Effects of PROGRESA and Birth-Year Rainfall on Education and Employment Outcomes, Controlling for Rainfall Shock Interactions with Unbalanced Characteristics

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) |
|----------------------------|------------------------|-------------------|------------------------------|--------------------|---------------------|---------------------|---------------------------------|---------------------------|---------------------------------------|
| | Educational Attainment | Grade Progression | Appropriate Grade Completion | Completed 7 grades | Completed 8 grades | Completed 9 grades | Currently Enrolled w/ HS Degree | Worked in Non-Laborer Job | Enrolled or Worked in Non-Laborer Job |
| Years of PROGRESA Exposure | 0.11 (0.038)*** | 0.010 (0.0098) | 0.013 (0.0076)* | 0.023 (0.011)** | 0.014 (0.011) | 0.020 (0.011)* | -0.0045 (0.0078) | -0.015 (0.016) | -0.015 (0.015) |
| Rainfall Shock | -0.51 (0.36) | -0.14 (0.084)* | -0.070 (0.077) | -0.17 (0.096)* | -0.16 (0.10) | -0.24 (0.12)** | -0.20 (0.086)** | -0.56 (0.36) | -0.63 (0.38)* |
| Rainfall Shock x Exposure | 0.11 (0.054)** | 0.019 (0.011)* | 0.019 (0.010)* | 0.032 (0.014)** | 0.039 (0.013)*** | 0.041 (0.015)*** | 0.027 (0.018) | 0.096 (0.052)* | 0.11 (0.051)** |
| Observations | 11824 | 11216 | 11824 | 10068 | 8285 | 6618 | 1597 | 1143 | 1138 |
| Mean of Dependent Variable | 6.79 | 0.58 | 0.46 | 0.56 | 0.52 | 0.45 | 0.061 | 0.35 | 0.41 |
| Sample Ages (in 2003) | 12 to 18 | 12 to 18 | 12 to 18 | 13 to 18 | 14 to 18 | 15 to 18 | 18 | 18 | 18 |
| Fixed Effects | | | | | | | Birth year x state | | |

- Standard errors clustered at the municipality level are in parentheses (** p<0.01, ** p<0.05, * p<0.1).

- "Rainfall shock" = 1 for individuals whose birth-year rainfall was more than one standard deviation from the 10-year historical locality-specific mean.

- All specifications include gender, household head gender and age, household size, household composition variables, parental education, parental language, and locality characteristics. Controls for parental language/education and locality distance include dummies for missing values

- All specifications include interactions between the rainfall shock variable and each of the control variables that are unbalanced across treatment and control villages (see Table D4).

age. To address this problem, we use inverse probability weighting to re-weight our sample so that the age distributions are balanced across treatment and control as well as across individuals born during normal and shock years. The results in Table C7, if anything, are stronger after using this weighting procedure, which suggests that our results were not a spurious consequence of this age imbalance.⁴

C.6 Alternate Variable Definitions

We investigate the robustness of our results to other methods of defining our two main independent variables of interest. First, we show that our results are robust to replacing our PROGRESA exposure variable with a simple treatment village indicator (Table C8). The effects are slightly weaker for some outcomes, which indicates that the additional cohort-level variation linked to the schooling incentive is important; however, the pattern of results is preserved across all regressions.

We have chosen the simple rainfall indicator that we use in the main results because it is parsimonious, captures non-linearities, and easy to interpret. However, we can certainly allow

⁴We only report our educational outcomes here because the age-weighting is irrelevant for the employment outcomes that involve only one age cohort.

Table C7: Effects of PROGRESA and Birth-Year Rainfall on Education and Employment Outcomes, Re-weighted on Age

| | (1) | (2) | (3) | (4) | (5) | (6) |
|----------------------------|------------------------|---------------------|------------------------------|---------------------|---------------------|---------------------|
| | Educational Attainment | Grade Progression | Appropriate Grade Completion | Completed 7 grades | Completed 8 grades | Completed 9 grades |
| Years of PROGRESA Exposure | 0.10 (0.038)*** | 0.011 (0.0097) | 0.015 (0.0077)* | 0.022 (0.011)** | 0.014 (0.011) | 0.019 (0.011) |
| Rainfall Shock | -0.93 (0.31)*** | -0.17 (0.052)*** | -0.13 (0.044)*** | -0.22 (0.071)*** | -0.24 (0.073)*** | -0.28 (0.088)*** |
| Rainfall Shock x Exposure | 0.16 (0.059)*** | 0.030 (0.010)*** | 0.019 (0.0090)** | 0.035 (0.014)** | 0.042 (0.014)*** | 0.050 (0.017)*** |
| Observations | 11824 | 11216 | 11824 | 10068 | 8285 | 6618 |
| Mean of Dependent Variable | 6.76 | 0.57 | 0.46 | 0.56 | 0.51 | 0.44 |
| Sample Ages (in 2003) | 12 to 18 | 12 to 18 | 12 to 18 | 13 to 18 | 14 to 18 | 15 to 18 |
| Fixed Effects | | | | Birth year x state | | |

- Standard errors clustered at the municipality level are in parentheses (** p<0.01, ** p<0.05, * p<0.1).

- "Rainfall shock" = 1 for individuals whose birth-year rainfall was more than one standard deviation from the 10-year historical locality-specific mean.

-All specifications include gender, household head gender and age, household size, household composition variables, parental education, parental language, and locality characteristics. Controls for parental language/education and locality distance include dummies for missing values.

-All specifications are weighted to produce the same age distributions across the four groups defined by treatment status and rainfall type.

Table C8: Effects of PROGRESA and Birth-Year Rainfall, Using Treatment Village Dummy

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) |
|----------------------------|------------------------|-------------------|------------------------------|----------------------|---------------------|--------------------|---------------------------------|---------------------------|---------------------------------------|
| | Educational Attainment | Grade Progression | Appropriate Grade Completion | Completed 7 grades | Completed 8 grades | Completed 9 grades | Currently Enrolled w/ HS Degree | Worked in Non-Laborer Job | Enrolled or Worked in Non-Laborer Job |
| Treatment Village | 0.21 (0.073)*** | 0.020 (0.019) | 0.022 (0.015) | 0.049 (0.022)** | 0.031 (0.022) | 0.045 (0.023)* | -0.0097 (0.016) | -0.029 (0.031) | -0.029 (0.030) |
| Rainfall Shock | -0.15 (0.094) | -0.026 (0.022) | -0.050 (0.019)*** | -0.072 (0.026)*** | -0.068 (0.029)** | -0.063 (0.033)* | -0.068 (0.024)*** | -0.023 (0.061) | -0.059 (0.062) |
| Rainfall Shock x Treatment | 0.080 (0.11) | 0.025 (0.025) | 0.040 (0.023)* | 0.050 (0.033) | 0.069 (0.034)** | 0.062 (0.040) | 0.035 (0.033) | 0.20 (0.079)** | 0.20 (0.083)** |
| Observations | 11824 | 11216 | 11824 | 10068 | 8285 | 6618 | 1597 | 1143 | 1138 |
| Mean of Dependent Variable | 6.79 | 0.58 | 0.46 | 0.56 | 0.52 | 0.45 | 0.061 | 0.35 | 0.41 |
| Sample Ages (in 2003) | 12 to 18 | 12 to 18 | 12 to 18 | 13 to 18 | 14 to 18 | 15 to 18 | 18 | 18 | 18 |
| Fixed Effects | | | | Birth year x state | | | | | |

Notes:

- Standard errors clustered at the municipality level are in parentheses (** p<0.01, ** p<0.05, * p<0.1).

- "Rainfall shock" = 1 for individuals whose birth-year rainfall was more than one standard deviation from the 10-year historical locality-specific mean.

- "Treatment village" = 1 for individuals in villages assigned to PROGRESA in the first wave (1998).

-All specifications include gender, household head gender and age, household size, household composition variables, parental education, parental language, and locality characteristics. Controls for parental language/education and locality distance include dummies for missing values

for greater flexibility in this specification, which we do in Table C9. Here, we replace our simple indicator with four dummy variables for normalized birth-year rainfall below the 20th percentile (“droughts”), between the 20th and 40th percentile (“below normal”), between the 60th and 80th percentile (“above normal”), and above the 80th percentile (“floods”) of the normalized rainfall distribution. Rainfall around the median – 40th to 60th percentile – is the omitted category.⁵ Consistent with our main results, we see that floods and droughts have negative effects on our outcomes of interest (larger in magnitude than the insignificant effects of the “below median” and “above median” dummies). We also find that the interactions between both droughts and floods with PROGRESA exposure are significant and indicative of remediation, across the majority of education outcomes. There is very little precision in any of our employment regressions, due to the very small sample sizes and more demanding empirical specification. In terms of magnitudes, the drought main effects and interactions appear to be slightly larger than the respective flood coefficients, but these differences are not significantly different from zero, which validates our use of a simple indicator that combines these two types of shocks.

We focus on rainfall shocks in an individual’s year of birth, specifically, because we are interested in shocks that affect a child’s endowment very early in life. A shock to the endowment during the year of birth should provide the cleanest and earliest source of exogenous variation, but it is of course possible that shocks during early childhood could also affect later-life outcomes. To investigate whether shocks in other years of life had similar positive effects on later-life outcomes, and similar interactions with PROGRESA, we add additional rainfall shock variables to our regressions and report the results in Table C10. Specifically, we add indicators for rainfall shocks during the year before birth, the second year of life, and the third year of life, along with their interactions with PROGRESA exposure. Consistent with Maccini and Yang (2009), we find that year of birth rainfall is the only one that has consistently large and significant effects across all outcomes. Accordingly, PROGRESA’s ability to remediate is only apparent with respect to birth-year rainfall and

⁵Droughts and floods are roughly (though not exactly) equivalent to using the one-standard-deviation cutoff that we use for our main rainfall shock dummy.

Table C9: Effects of PROGRESA and Birth-Year Rainfall, Using Flexible Definition of Rainfall Shock

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) |
|--|------------------------|---------------------|------------------------------|--------------------|---------------------|--------------------|---------------------------------|---------------------------|---------------------------------------|
| | Educational Attainment | Grade Progression | Appropriate Grade Completion | Completed 7 grades | Completed 8 grades | Completed 9 grades | Currently Enrolled w/ HS Degree | Worked in Non-Laborer Job | Enrolled or Worked in Non-Laborer Job |
| Years of PROGRESA Exposure | 0.068 (0.048) | 0.0038 (0.011) | 0.013 (0.0090) | 0.019 (0.014) | 0.011 (0.015) | 0.015 (0.015) | 0.0016 (0.012) | -0.055 (0.029)* | -0.042 (0.029) |
| Drought | -1.03 (0.40)** | -0.22 (0.072)*** | -0.18 (0.064)*** | -0.24 (0.094)** | -0.27 (0.091)*** | -0.28 (0.11)*** | -0.060 (0.063) | -0.16 (0.19) | -0.16 (0.19) |
| Below normal rainfall | -0.16 (0.30) | 0.031 (0.073) | 0.020 (0.059) | 0.028 (0.067) | 0.021 (0.071) | 0.032 (0.073) | 0.065 (0.076) | -0.21 (0.14) | -0.12 (0.14) |
| Above normal rainfall | -0.30 (0.27) | 0.0029 (0.059) | -0.014 (0.050) | -0.065 (0.070) | -0.096 (0.075) | -0.089 (0.077) | -0.025 (0.052) | -0.18 (0.14) | -0.14 (0.14) |
| Flood | -0.65 (0.28)** | -0.19 (0.075)** | -0.11 (0.064)* | -0.19 (0.10)* | -0.18 (0.11)* | -0.28 (0.11)** | -0.054 (0.079) | -0.20 (0.19) | -0.18 (0.22) |
| Drought x Exposure | 0.17 (0.075)** | 0.037 (0.014)*** | 0.027 (0.013)** | 0.040 (0.018)** | 0.047 (0.017)*** | 0.050 (0.021)** | -0.00068 (0.022) | 0.080 (0.054) | 0.067 (0.055) |
| Below normal rainfall x Exposure | 0.027 (0.058) | -0.0024 (0.015) | -0.0051 (0.012) | -0.0051 (0.013) | -0.0057 (0.014) | -0.0097 (0.014) | -0.028 (0.020) | 0.073 (0.044)* | 0.036 (0.044) |
| Above normal rainfall x Exposure | 0.049 (0.052) | -0.00084 (0.012) | -0.0030 (0.011) | 0.0099 (0.014) | 0.014 (0.015) | 0.015 (0.016) | 0.0028 (0.015) | 0.067 (0.040) | 0.055 (0.041) |
| Flood x Exposure | 0.12 (0.056)** | 0.041 (0.015)*** | 0.014 (0.013) | 0.032 (0.019)* | 0.032 (0.021) | 0.053 (0.023)** | -0.0033 (0.022) | 0.085 (0.061) | 0.063 (0.069) |
| Observations | 11824 | 11216 | 11824 | 10068 | 8285 | 6618 | 1597 | 1143 | 1138 |
| Mean of Dependent Variable | 6.79 | 0.58 | 0.46 | 0.56 | 0.52 | 0.45 | 0.061 | 0.35 | 0.41 |
| Sample Ages (in 2003) | 12 to 18 | 12 to 18 | 12 to 18 | 13 to 18 | 14 to 18 | 15 to 18 | 18 | 18 | 18 |
| <i>Tests for equality of coefficients (p-values)</i> | | | | | | | | | |
| drought = flood | 0.66 | 0.68 | 0.57 | 0.19 | 0.16 | 0.17 | 0.22 | 0.82 | 0.86 |
| below normal = above normal | 0.36 | 0.73 | 0.31 | 0.66 | 0.49 | 0.99 | 0.95 | 0.88 | 0.95 |
| drought x exposure = flood x exposure | 0.72 | 0.91 | 0.86 | 0.29 | 0.20 | 0.15 | 0.09 | 0.88 | 0.62 |
| below normal x exposure = above normal x exposure | 0.51 | 0.84 | 0.34 | 0.73 | 0.53 | 0.90 | 0.93 | 0.95 | 0.96 |
| Fixed Effects | | | | | | Birth year x state | | | |

Notes:

- Standard errors clustered at the municipality level are in parentheses (** p<0.01, ** p<0.05, * p<0.1).

- "Drought", "Below normal rainfall," "Above normal rainfall," and "Flood" are dummy variables indicating individuals whose birth-year rainfall (normalized using the locality-specific historical 10-year mean and standard deviation) fell below the 20th percentile, between the 20th and 40th percentile, between the 60th and 80th percentile, and above the 80th percentile, respectively.

- All specifications include gender, household head gender and age, household size, household composition variables, parental education, parental language, and locality characteristics. Controls for parental language/education and locality distance include dummies for missing values

not rainfall in any other year (with the exception of column 3 for rainfall in the year before birth). Across both employment and education outcomes, our three main coefficients of interest are almost identical in magnitude to the estimates in Tables 4, 5, and 6.

Table C10: Effects of PROGRESA and Rainfall in Year of Birth and Other Early-Life Years

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) |
|--|------------------------|---------------------|------------------------------|---------------------|---------------------|---------------------|---------------------------------|---------------------------|---------------------------------------|
| | Educational Attainment | Grade Progression | Appropriate Grade Completion | Completed 7 grades | Completed 8 grades | Completed 9 grades | Currently Enrolled w/ HS Degree | Worked in Non-Laborer Job | Enrolled or Worked in Non-Laborer Job |
| Years of PROGRESA Exposure | 0.087 (0.050)* | 0.0098 (0.011) | 0.013 (0.0099) | 0.015 (0.014) | 0.010 (0.013) | 0.013 (0.014) | 0.0061 (0.011) | -0.0093 (0.026) | -0.0027 (0.026) |
| Rainfall Shock in year of birth | -0.63 (0.30)** | -0.12 (0.060)** | -0.13 (0.053)** | -0.21 (0.072)*** | -0.23 (0.074)*** | -0.25 (0.086)*** | -0.10 (0.062)* | -0.32 (0.14)** | -0.35 (0.15)** |
| Rainfall Shock in year of birth x Exposure | 0.11 (0.057)* | 0.022 (0.012)* | 0.023 (0.011)** | 0.035 (0.014)** | 0.042 (0.014)*** | 0.048 (0.017)*** | 0.022 (0.019) | 0.12 (0.043)*** | 0.12 (0.047)** |
| Rainfall Shock in year before birth | -0.41 (0.25) | 0.021 (0.052) | -0.082 (0.044)* | -0.071 (0.067) | -0.092 (0.070) | -0.072 (0.073) | -0.014 (0.053) | -0.12 (0.11) | -0.17 (0.10)* |
| Rainfall Shock in year before birth x Exposure | 0.082 (0.050) | -0.0043 (0.011) | 0.016 (0.0091)* | 0.013 (0.014) | 0.018 (0.014) | 0.015 (0.015) | 0.0053 (0.015) | 0.026 (0.034) | 0.049 (0.032) |
| Rainfall Shock in second year | -0.25 (0.29) | -0.087 (0.071) | -0.052 (0.053) | -0.11 (0.065)* | -0.11 (0.069) | -0.12 (0.078) | -0.028 (0.042) | 0.090 (0.11) | 0.064 (0.100) |
| Rainfall Shock in second year x Exposure | 0.038 (0.056) | 0.016 (0.013) | 0.0091 (0.010) | 0.019 (0.013) | 0.014 (0.013) | 0.020 (0.015) | -0.0037 (0.012) | -0.040 (0.029) | -0.042 (0.029) |
| Rainfall Shock in third year | -0.064 (0.28) | 0.038 (0.046) | 0.088 (0.047)* | -0.021 (0.068) | -0.027 (0.068) | -0.061 (0.078) | 0.13 (0.054)** | 0.11 (0.11) | 0.21 (0.097)** |
| Rainfall Shock in third year x Exposure | 0.0095 (0.052) | -0.0055 (0.0097) | -0.017 (0.0097)* | 0.0070 (0.013) | 0.0035 (0.013) | 0.0046 (0.016) | -0.039 (0.015)** | -0.044 (0.032) | -0.073 (0.031)** |
| Observations | 10607 | 10057 | 10607 | 8940 | 7404 | 5860 | 1450 | 1028 | 1025 |
| Mean of Dependent Variable | 6.76 | 0.58 | 0.46 | 0.56 | 0.51 | 0.45 | 0.061 | 0.35 | 0.41 |
| Sample Ages (in 2003) | 12 to 18 | 12 to 18 | 12 to 18 | 13 to 18 | 14 to 18 | 15 to 18 | 18 | 18 | 18 |
| Fixed Effects | | | | | | | Birth year x state | | |

Notes:

- Standard errors clustered at the municipality level are in parentheses (** p<0.01, ** p<0.05, * p<0.1).

- "Rainfall shock" = 1 for individuals whose rainfall (in the relevant year) was more than one standard deviation from the 10-year historical locality-specific mean.

- All specifications include gender, household head gender and age, household size, household composition variables, parental education, parental language, and locality characteristics. Controls for parental language/education and locality distance include dummies for missing values

D Additional Tables and Figures

Table D1: Monthly Amount of Educational Transfers to Beneficiary Households (from Behrman et al. 2011)

| | 1997 | | 1998 | | 2003 | |
|---------------------------|------|-------|------|-------|------|-------|
| | Boys | Girls | Boys | Girls | Boys | Girls |
| Primary School | | | | | | |
| 3rd year | 60 | 60 | 70 | 70 | 105 | 105 |
| 4th year | 70 | 70 | 80 | 80 | 120 | 120 |
| 5th year | 90 | 90 | 100 | 100 | 155 | 155 |
| 6th year | 120 | 120 | 135 | 135 | 210 | 210 |
| Junior High School | | | | | | |
| 1st year | 175 | 185 | 200 | 210 | 305 | 320 |
| 2nd year | 185 | 205 | 210 | 235 | 320 | 355 |
| 3rd year | 195 | 205 | 220 | 625 | 335 | 390 |
| High School | | | | | | |
| 1st year | - | - | - | - | 510 | 585 |
| 2nd year | - | - | - | - | 545 | 625 |
| 3rd year | - | - | - | - | 580 | 660 |

Notes:

1. Amounts (in pesos) are for the second semester of the year
2. Grants extended to high school in 2001.

Figure D1: Proportion of Individuals Not Living in Household, by Age

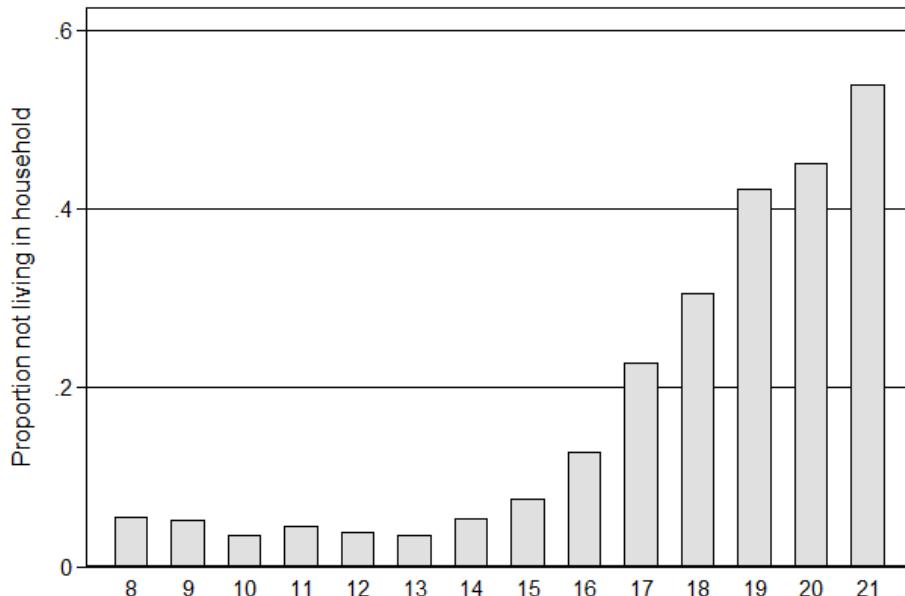


Figure D2: PROGRESA Localities by Treatment Status and Rainfall Shock in 1987

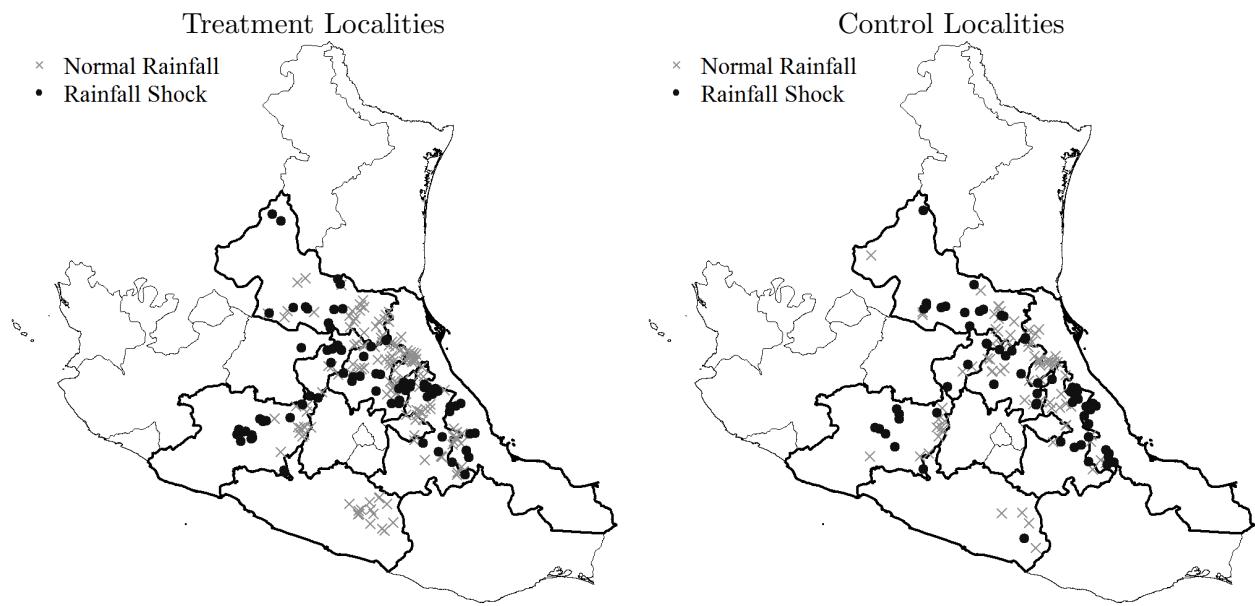
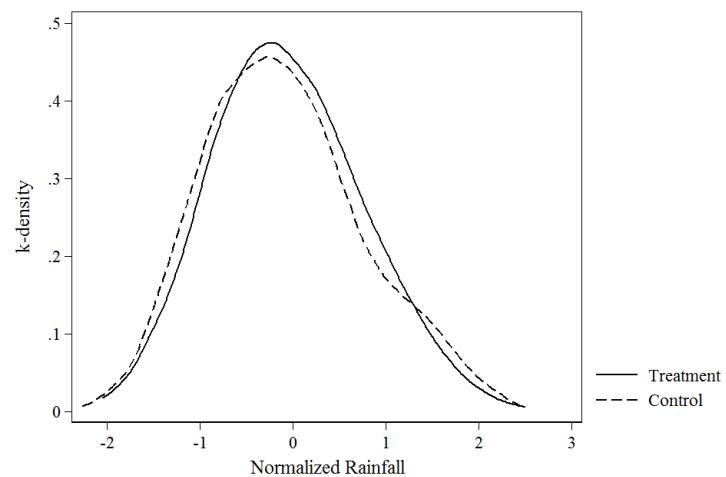


Figure D3: Normalized Rainfall Distributions in Treatment and Control Villages



Notes:

Rainfall levels are normalized using each locality's location-specific 10-year historical mean and standard deviation.

Figure D4: PROGRESA Localities by Treatment Status and Rainfall Shock in 1987, Trimmed Sample

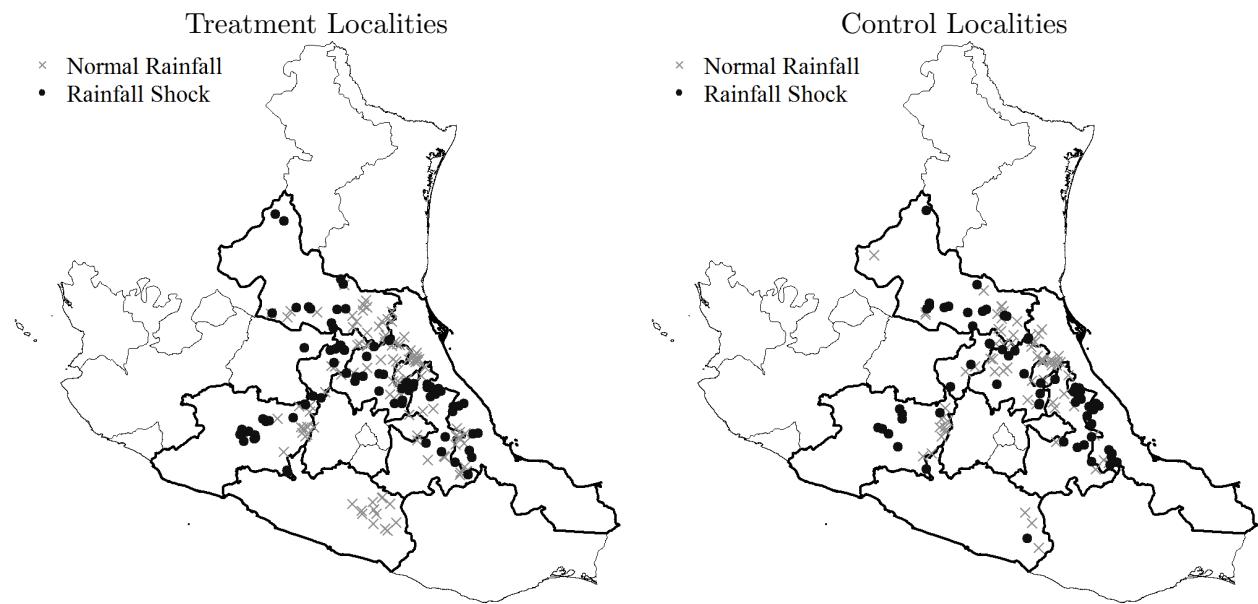
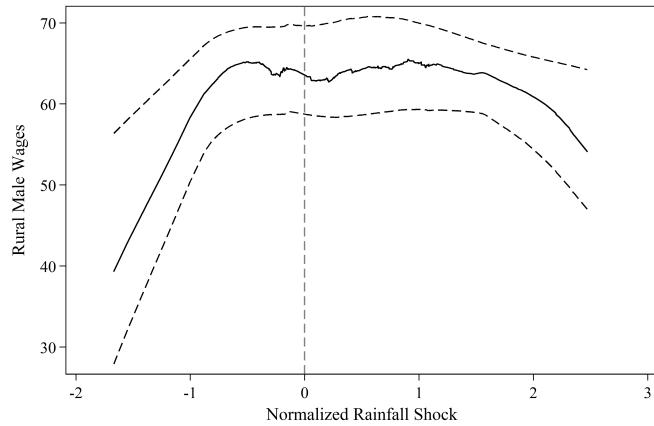


Figure D5: Locality Wages



Notes:

Solid line represents the lowess-smoothed relationship between rural male wages and normalized rainfall. Dashed lines represent 95% confidence intervals, calculated from 1000 bootstrapped samples.

Table D2: Summary of Related PROGRESA Literature

| Age of sample during Progresa exposure | Outcome Category | Outcome | Result | Analysis timeframe | Studies | |
|--|--------------------------|---|---|---|--|--|
| School-aged | Education | Attendance | Increased, particularly for older ages | ST | Skoufias and Parker (2001) | |
| | | Dropouts | Decreased | ST | Behrman, Sengupta, Todd (2005) | |
| | | Educational attainment | Increased by ~0.66 years | ST | Schultz (2004); Behrman, Sengupta, Todd (2005) | |
| | | Educational attainment | Increased by ~0.2 years | MT | Behrman, Parker, Todd (2011); Behrman, Parker, Todd (2009) | |
| | | Enrollment | Increased for younger ages | MT | Behrman, Parker, Todd (2009) | |
| | | Enrollment | Increased for older ages | ST | Behrman, Sengupta, Todd (2000); Schultz (2004) | |
| | | Grade progression | Increased | ST | Behrman, Sengupta, Todd (2005) | |
| | | Grade progression | Increased | MT | Behrman, Parker, Todd (2009) | |
| | | Grade repetition | Decreased | ST | Behrman, Sengupta, Todd (2005) | |
| | | Re-entering school | Increased | ST | Behrman, Sengupta, Todd (2005) | |
| | Health | Schooling gaps | Decreased | ST | Behrman, Sengupta, Todd (2005); Behrman, Sengupta, Todd (2000) | |
| | | Work | Decreased | ST | Skoufias and Parker (2001); Schultz (2004) | |
| | | Work | Decreased for younger boys | MT | Behrman, Parker, Todd (2011); Behrman, Parker, Todd (2009) | |
| Younger than 3rd grade | Education | Overweight | Decreased for girls (in most spec's) | MT | Andalon (2011) | |
| | | Cognitive tests | No significant effect | ST | Behrman, Sengupta, Todd (2000) | |
| | | Grade progression | Increased | MT | Behrman, Parker, Todd (2009b) | |
| Younger than school age | Health | Age of school start | Decreased | MT | Behrman, Parker, Todd (2009b) | |
| | | Educational attainment | Increased | MT | Behrman, Parker, Todd (2009b) | |
| | | Grade progression | Increased | MT | Behrman, Parker, Todd (2009b) | |
| | | Anemia | Decreased | ST | Gertler (2004) | |
| | | BMI | No significant effect | LT | Fernald, Gertler, Neufeld (2009) | |
| | | Height | Increased | ST | Gertler (2004) | |
| | | Height | Increased only for children of mothers with no education | LT | Fernald, Gertler, Neufeld (2009) | |
| | | Salivary cortisol | Decreased for children of mothers with high depressive symptoms | MT | Fernald and Gunnar (2009) | |
| | Cognitive and Behavioral | Self-reported morbidity | Decreased | ST | Gertler (2004) | |
| Younger than school age (including not born) | | Behavioral problems | Decreased | LT | Fernald, Gertler, Neufeld (2009) | |
| | | Cognitive tests | No significant effect | LT | Fernald, Gertler, Neufeld (2009) | |
| | | Language tests | Increased | MT | Fernald, Gertler, Neufeld (2008b) | |
| Health | Anemia | Decreased in ST | ST | Rivera et al (2004) | | |
| | Birthweight (self-rep.) | Increased | ST/MT | Barber and Gertler (2008) | | |
| | Height | Increased | ST | Behrman and Hoddinott (2005); Rivera et al (2004) | | |
| | Infant mortality | Reduced | MT | Barham (2011) | | |
| | Neonatal mortality | No significant effect | MT | Barham (2011) | | |
| Not born | Health | Pre-natal care visits by mother | No significant effect | ST/MT | Barber and Gertler (2008) | |
| | | BMI | Decreased | MT | Fernald, Gertler, Neufeld (2008b) | |
| | | Height | Increased | MT | Fernald, Gertler, Neufeld (2008b) | |
| | Cognitive and Behavioral | Motor development | Increased | MT | Fernald, Gertler, Neufeld (2008b) | |
| Adults | | Cognitive tests | Increased | MT | Fernald, Gertler, Neufeld (2008b) | |
| Health | Blood pressure | Increased (due to cash component) | MT | Fernald, Gertler, Hou (2008a) | | |
| | Blood pressure | Decreased | MT | Fernald, Hou, Gertler (2008c) | | |
| | BMI | Increased (due to cash component) | MT | Fernald, Gertler, Hou (2008a) | | |
| | Elderly mortality | Reduced | MT | Barham and Rowberry (2013) | | |
| | Hypertension | Decreased | MT | Fernald, Hou, Gertler (2008c) | | |
| | Overweight, obesity | Increased (due to cash component) | MT | Fernald, Gertler, Hou (2008a) | | |
| N/A - Households | Income and Consumption | Self-reported health | Increased | MT | Fernald, Hou, Gertler (2008c) | |
| | | Agricultural income, assets, production | Increased | ST | Gertler, Martinez, Rubio-Condina (2012) | |
| | | Consumption | Increased | ST | Djebari and Smith (2008); Angelucci and De Giorgi (2012) | |
| | | Food consumption | Increased | ST | Hoddinott and Skoufias (2004); Angelucci and De Giorgi (2012) | |

- ST (short-term) estimates used outcomes measured before control group received treatment in 2000

- MT (medium-term) estimates used outcomes measured between 2000 and 2003

- LT (long-term) estimates used outcomes measured in 2007

Table D3: Summary Statistics for Control Variables

Panel A: Household-level

| | <i>Full Sample</i> | <i>Treatment Villages</i> | <i>Control Villages</i> | <i>Treatment - Control Differences</i> |
|---|--------------------|---------------------------|-------------------------|--|
| Household size | 7.41 (2.19) | 7.42 (2.22) | 7.40 (2.15) | 0.019 (0.041) |
| Household head age | 41.7 (11.3) | 41.4 (11.1) | 42.2 (11.6) | -0.79*** (0.21) |
| Female household head | 0.057 (0.23) | 0.056 (0.23) | 0.057 (0.23) | -0.00047 (0.0043) |
| Number of children aged 0-2 | 0.073 (0.086) | 0.073 (0.086) | 0.072 (0.087) | 0.0016 (0.0016) |
| Number of children aged 3-5 | 0.10 (0.096) | 0.10 (0.096) | 0.099 (0.096) | 0.0034* (0.0018) |
| Number of boys aged 6-7 | 0.052 (0.077) | 0.051 (0.076) | 0.054 (0.079) | -0.0027* (0.0014) |
| Number of boys aged 8-12 | 0.12 (0.11) | 0.13 (0.11) | 0.12 (0.11) | 0.0049** (0.0021) |
| Number of boys aged 13-18 | 0.070 (0.095) | 0.070 (0.096) | 0.069 (0.093) | 0.00067 (0.0018) |
| Number of girls aged 6-7 | 0.051 (0.076) | 0.052 (0.077) | 0.051 (0.076) | 0.00082 (0.0014) |
| Number of girls aged 8-12 | 0.12 (0.11) | 0.12 (0.11) | 0.12 (0.11) | -0.0023 (0.0021) |
| Number of girls aged 13-18 | 0.066 (0.091) | 0.065 (0.091) | 0.067 (0.091) | -0.0014 (0.0017) |
| Number of women aged 19-54 | 0.16 (0.061) | 0.16 (0.061) | 0.16 (0.062) | -0.00096 (0.0011) |
| Number of men aged 55 and over | 0.019 (0.051) | 0.018 (0.051) | 0.019 (0.051) | -0.00087 (0.00094) |
| Number of women aged 55 and over | 0.017 (0.050) | 0.017 (0.050) | 0.018 (0.051) | -0.0018* (0.00093) |
| Mother's educational attainment (categorical) | 3.93 (2.07) | 3.92 (2.07) | 3.93 (2.07) | -0.0040 (0.048) |
| Mother's educational attainment missing | 0.34 (0.47) | 0.33 (0.47) | 0.36 (0.48) | -0.024*** (0.0088) |
| Father's educational attainment (categorical) | 3.98 (2.25) | 4.03 (2.31) | 3.89 (2.14) | 0.14*** (0.050) |
| Father's educational attainment missing | 0.31 (0.46) | 0.30 (0.46) | 0.31 (0.46) | -0.0080 (0.0086) |
| Mother speaks indigenous language | 0.38 (0.48) | 0.37 (0.48) | 0.39 (0.49) | -0.012 (0.0092) |
| Mother's language missing | 0.041 (0.20) | 0.039 (0.19) | 0.044 (0.20) | -0.0046 (0.0037) |
| Father speaks indigenous language | 0.39 (0.49) | 0.38 (0.49) | 0.40 (0.49) | -0.021** (0.0095) |
| Father's language missing | 0.096 (0.29) | 0.097 (0.30) | 0.094 (0.29) | 0.0030 (0.0055) |
| Number of households | 6233 | 3795 | 2438 | |

Panel B: Locality-level

| | <i>Full Sample</i> | <i>Treatment Villages</i> | <i>Control Villages</i> | <i>Treatment - Control Differences</i> |
|--------------------------------------|--------------------|---------------------------|-------------------------|--|
| Community Well | 0.38 (0.49) | 0.37 (0.48) | 0.39 (0.49) | -0.027 (0.049) |
| Well Spring | 0.48 (0.50) | 0.51 (0.50) | 0.44 (0.50) | 0.074 (0.050) |
| Public Water Network | 0.15 (0.36) | 0.12 (0.33) | 0.19 (0.39) | -0.070* (0.035) |
| Bury Garbage | 0.18 (0.39) | 0.21 (0.41) | 0.14 (0.35) | 0.065* (0.039) |
| Public Dumpster | 0.017 (0.13) | 0.0078 (0.088) | 0.031 (0.17) | -0.023* (0.013) |
| Public Drainage | 0.038 (0.19) | 0.035 (0.18) | 0.043 (0.20) | -0.0079 (0.019) |
| Public Phone | 0.52 (0.50) | 0.52 (0.50) | 0.52 (0.50) | -0.0040 (0.050) |
| Hospital or health center | 0.15 (0.36) | 0.13 (0.34) | 0.18 (0.38) | -0.046 (0.036) |
| Distance to health center | 13.5 (24.4) | 13.7 (24.3) | 13.2 (24.7) | 0.57 (2.45) |
| DICONSA store | 0.24 (0.43) | 0.26 (0.44) | 0.20 (0.40) | 0.058 (0.043) |
| Distance to Bank | 38.7 (51.8) | 40.5 (59.3) | 36.0 (37.6) | 4.48 (5.50) |
| Distance to Secondary School | 11.8 (15.9) | 12.2 (16.0) | 11.3 (15.9) | 0.84 (2.44) |
| Distance to Secondary School Missing | 0.58 (0.49) | 0.60 (0.49) | 0.55 (0.50) | 0.047 (0.049) |
| Number of localities | 420 | 257 | 163 | |

Notes:

Standard deviations (in the first 3 columns) and standard errors (in the last column) in parentheses (** p<0.01, ** p<0.05, * p<0.1). Missing indicators for parental education and language are binary variables equal to 1 for individuals missing the relevant information. Community well, well spring, public water network, public dumpster, public drainage, public phone, hospital or health center, and DICONSA store are all indicators equal to 1 for localities that have the relevant public good or facility. Bury garbage is an indicator equal to 1 for localities that report burying garbage as their main form of garbage disposal. Distances reported in kilometers. Missing distance variables are indicators for localities that did not report a distance to the nearest secondary school or bank.

Table D4: Rainfall Shocks, PROGRESA, and Baseline Characteristics

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) | (11) | (12) | (13) | (14) | (15) | (16) | (17) | (18) | (19) | (20) |
|----------------------------|----------------------------------|-----------------------------------|-----------------------------------|----------------------------------|---------------------------|-----------------------------|---------------------------|-----------------------------|---------------------------|-----------------------------|---------------------------|-----------------------------|---------------------------|-----------------------------|---------------------------|-----------------------------|---------------------------|-----------------------------|--------------------------------|------------------------------|
| Mother's education | Mother's Father's education | Mother's Father's education | Mother's Father's primary only | Mother's Father's secondary only | Mother's Father's primary | Mother's Father's secondary | Number of men aged 55 and over | |
| Treatment Village | -0.0020 (0.12) | 0.16 (0.15) | 0.0077 (0.028) | -0.0069 (0.027) | 0.027 (0.030) | 0.011 (0.085) | -0.047 (0.026)* | -0.025 (0.021) | -0.74 (0.40)* | -0.0036 (0.071) | -0.0033 (0.029) | 0.0012 (0.0033) | -0.0026 (0.026) | 0.0057 (0.038) | 0.0014 (0.0026) | 0.0049 (0.0023) | 0.00011 (0.0040) | -0.0026 (0.0029) | -0.0019 (0.0020) | -0.0065 (0.0016) |
| Rainfall Shock | -0.0043 (0.093) | -0.044 (0.12) | -0.00014 (0.020) | -0.0071 (0.0091) | -0.0067 (0.023) | -0.0080 (0.0077) | -0.011 (0.023) | 0.044 (0.019)** | -0.21 (0.34) | 0.020 (0.089) | -0.0025 (0.0032) | -0.0015 (0.0033) | 0.0014 (0.0045) | -0.0013 (0.0028) | 0.0097 (0.0029) | -0.00076 (0.0027) | 0.0096 (0.0040)** | -0.0041 (0.0030) | -0.0024 (0.0023) | 0.0012 (0.0015) |
| Rainfall Shock x Treatment | -0.044 (0.12) | -0.013 (0.14) | -0.028 (0.028) | 0.0058 (0.012) | 0.022 (0.029) | -0.00064 (0.011) | 0.026 (0.026) | -0.016 (0.023) | -0.041 (0.45) | -0.0051 (0.0094) | 0.00063 (0.0037) | 0.0018 (0.0044) | -0.00041 (0.0029) | 0.0022 (0.0049) | -0.0036 (0.0037) | 0.0017 (0.0032) | -0.0088 (0.0050)* | 0.0047 (0.0039) | 0.0056 (0.0026)** | -0.0071 (0.0021) |
| Observations | 7998 | 8424 | 7998 | 7998 | 8424 | 12159 | 12159 | 12159 | 12159 | 12159 | 12159 | 12159 | 12159 | 12159 | 12159 | 12159 | 12159 | 12159 | 12159 | |
| Mean of Dep. Var. | 3.95 | 3.98 | 0.29 | 0.037 | 0.30 | 0.050 | 0.34 | 0.31 | 41.7 | 0.057 | 0.073 | 0.10 | 0.052 | 0.12 | 0.070 | 0.051 | 0.12 | 0.066 | 0.16 | |
| Sample Ages (in 2003) | | | | | | | | | | | | | | | | | | | 0.019 | |
| Fixed Effects | (21) | (22) | (23) | (24) | (25) | (26) | (27) | (28) | (29) | (30) | (31) | (32) | (33) | (34) | (35) | (36) | (37) | (38) | (39) | |
| | Number of women aged 55 and over | Mother speaks indigenous language | Mother speaks indigenous language | Mother's language missing | Mother's language missing | Community well | Well spring | Public water network | Bury | Public garbage | Public dumpster | Public drainage | Public phone | Hospital or health center | Distance to health center | DICNSA store | Distance to bank | Distance to bank | Distance to secondary school | Distance to secondary school |
| Treatment Village | -0.00052 (0.0014) | -0.0018 (0.065) | -0.011 (0.06) | -0.00065 (0.0057) | 0.0077 (0.0079) | -0.037 (0.062) | 0.063 (0.064) | -0.074 (0.045) | 0.074 (0.044)* | -0.020 (0.018) | -0.011 (0.018) | -0.028 (0.019) | -0.011 (0.020) | -0.028 (0.024) | 2.89 (2.44) | 0.074 (0.055) | -0.57 (5.96) | 0.0028 (0.039) | -0.14 (1.26) | 0.056 (0.063) |
| Rainfall Shock | 0.00025 (0.0018) | -0.088 (0.046)* | -0.086 (0.045)* | 0.0025 (0.0078) | 0.016 (0.012) | 0.00076 (0.0044) | -0.031 (0.031) | 0.036 (0.044) | -0.024 (0.027) | 0.030 (0.024) | 0.00930 (0.016) | -0.077 (0.052) | 0.0073 (0.038) | 0.58 (1.59) | -0.0046 (0.043) | -1.95 (3.14) | 0.019 (0.028) | 1.27 (0.99) | -0.32 (0.04) | -0.69 (0.21)** |
| Rainfall Shock x Treatment | -0.0025 (0.0025) | 0.036 (0.049) | 0.031 (0.049) | -0.087 (0.0087) | -0.015 (0.014) | 0.044 (0.055) | -0.059 (0.052) | -0.045 (0.033) | 0.020 (0.037) | -0.032 (0.022) | -0.0100 (0.016) | 0.086 (0.037) | 0.028 (0.049) | -2.54 (2.30) | 0.042 (0.052) | 3.83 (3.43) | -0.0091 (0.037) | -1.95 (1.10)* | 0.067 (0.051) | 0.30 (0.32) |
| Observations | 12159 | 11663 | 10995 | 12159 | 12159 | 12159 | 12159 | 12159 | 12159 | 12159 | 12159 | 12159 | 12159 | 12159 | 12159 | 12159 | 12159 | 12159 | 12159 | |
| Mean of Dep. Var. | 0.017 | 0.38 | 0.39 | 0.041 | 0.096 | 0.36 | 0.51 | 0.14 | 0.16 | 0.015 | 0.037 | 0.57 | 0.20 | 12.6 | 0.30 | 33.6 | 0.11 | 5.55 | 0.57 | 15.0 |
| Sample Ages (in 2003) | | | | | | | | | | | | | | | | | | | 0.019 | |
| Fixed Effects | | | | | | | | | | | | | | | | | | | | |
| | | | | | | | | | | | | | | | | | | | State | |

Notes:
- Standard errors clustered at the municipality level are in parentheses (** p<0.01, * p<0.05, . p>0.1).

"Rainfall shock" = 1 for individual whose birth-year rainfall was more than one standard deviation from the 10-year historical locality-specific mean.

"Treatment village" = 1 for individuals in villages assigned to PROGRESA in the first wave (1998).

Table D5: Effects of Birth-Year Rainfall on Parent and Sibling Characteristics in Baseline Survey

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
|----------------------------|------------------|---------------------|----------------------|----------------------|-----------------------|-----------------------|--|----------------------------------|
| | Mother Employed | Father Employed | Mother's Days Worked | Father's Days Worked | Mother's Hours Worked | Father's Hours Worked | Average Sibling Educational Attainment | Average Sibling Grade Completion |
| Rainfall Shock | 0.013 (0.011) | -0.0072 (0.0057) | 0.063 (0.060) | -0.0058 (0.047) | 0.11 (0.081) | 0.023 (0.064) | -0.021 (0.038) | -0.0066 (0.0079) |
| Observations | 11677 | 11011 | 11659 | 10950 | 11665 | 10944 | 10577 | 10360 |
| Mean of Dependent Variable | 0.12 | 0.95 | 0.66 | 5.20 | 0.91 | 7.91 | 3.60 | 0.64 |
| Sample Ages (in 2003) | 12 to 18 | | | | | | | |
| Fixed Effects | | | | | Birth year x state | | | |

Notes:

- Standard errors clustered at the municipality level are in parentheses (** p<0.01, ** p<0.05, * p<0.1).
- "Rainfall shock" = 1 for individuals whose birth-year rainfall was more than one standard deviation from the 10-year historical locality-specific mean.
- Outcome variables taken from the baseline survey

Table D6: Effects of Own PROGRESA Exposure, Own Birth-Year Rainfall, Sibling PROGRESA Exposure, and Sibling Birth-Year Rainfall on Education and Employment Outcomes

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) |
|---------------------------------------|------------------------|---------------------|------------------------------|---------------------|---------------------|---------------------|---------------------------------|---------------------------|---------------------------------------|
| | Educational Attainment | Grade Progression | Appropriate Grade Completion | Completed 7 grades | Completed 8 grades | Completed 9 grades | Currently Enrolled w/ HS Degree | Worked in Non-Laborer Job | Enrolled or Worked in Non-Laborer Job |
| Years of PROGRESA Exposure | 0.10 (0.049)** | 0.0095 (0.011) | 0.012 (0.0092) | 0.028 (0.013)** | 0.010 (0.013) | 0.027 (0.015)* | -0.013 (0.013) | 0.023 (0.028) | 0.011 (0.029) |
| Own Rainfall Shock | -0.60 (0.29)** | -0.11 (0.056)* | -0.11 (0.054)** | -0.19 (0.074)** | -0.22 (0.074)*** | -0.26 (0.086)*** | -0.086 (0.054) | -0.25 (0.14)* | -0.28 (0.15)* |
| Own Rainfall Shock x Exposure | 0.12 (0.053)** | 0.021 (0.011)* | 0.019 (0.010)* | 0.030 (0.014)** | 0.039 (0.014)*** | 0.041 (0.016)** | 0.022 (0.018) | 0.14 (0.036)*** | 0.15 (0.039)*** |
| Sibling Rainfall Shock | -0.14 (0.31) | -0.013 (0.069) | -0.028 (0.059) | 0.012 (0.068) | -0.055 (0.077) | 0.023 (0.092) | -0.065 (0.077) | 0.38 (0.17)** | 0.33 (0.17)* |
| Sibling Rainfall x Exposure | 0.017 (0.062) | 0.0011 (0.014) | 0.0018 (0.012) | -0.0083 (0.013) | 0.0060 (0.015) | -0.011 (0.020) | 0.0084 (0.021) | -0.12 (0.051)** | -0.11 (0.052)** |
| Average Sibling PROGRESA Exposure | -0.0091 (0.028) | -0.0038 (0.0075) | 0.00059 (0.0075) | -0.0042 (0.0079) | 0.0012 (0.0093) | -0.0057 (0.010) | 0.011 (0.012) | 0.023 (0.024) | 0.042 (0.026) |
| Sibling Exposure x Own Rainfall Shock | -0.032 (0.036) | -0.0019 (0.0088) | -0.0027 (0.0081) | 0.00020 (0.011) | 0.00086 (0.013) | 0.013 (0.011) | -0.012 (0.011) | -0.047 (0.037) | -0.053 (0.035) |
| Observations | 11379 | 10797 | 11379 | 9679 | 7948 | 6314 | 1487 | 1065 | 1063 |
| Mean of Dependent Variable | 6.77 | 0.58 | 0.47 | 0.56 | 0.51 | 0.45 | 0.060 | 0.36 | 0.42 |
| Sample Ages (in 2003) | 12 to 18 | 12 to 18 | 12 to 18 | 13 to 18 | 14 to 18 | 15 to 18 | 18 | 18 | 18 |
| Fixed Effects | | | | | Birth year x state | | | | |

Notes:

- Standard errors clustered at the municipality are reported in parentheses (** p<0.01, ** p<0.05, * p<0.1).
- "Rainfall shock" = 1 for individuals whose birth-year rainfall was more than one standard deviation from the 10-year historical locality-specific mean.
- "Sibling rainfall shock" is the average of the rainfall shock variable across all of a child's siblings.
- "Average Sibling PROGRESA Exposure" is the average years of PROGRESA exposure across all of a child's siblings.
- These regressions drop all children with no siblings.
- All specifications include gender, household head gender and age, household size, household composition variables, parental education, parental language, and locality characteristics. Controls for parental language/education and locality distance include dummies for missing values.

Table D7: Effects of Birth-Year Rainfall on Future Household Income

| | (1) | (2) |
|----------------------------|-----------------------------------|----------------|
| Eligible for PROGRESA | 1997 Household Income Score | |
| Rainfall Shock | -0.022 (0.014) | 3.16 (5.03) |
| Observations | 16836 | 16818 |
| Mean of Dependent Variable | 0.72 | 677.7 |
| Sample Ages (in 2003) | 12 to 18 | |

Notes:

- Standard errors clustered at the municipality level are in parentheses
- "Rainfall shock" = 1 for individuals whose birth-year rainfall was more than one standard deviation from the 10-year historical locality-specific mean.

Table D8: Effects of PROGRESA, Birth-Year Rainfall, and Household Income on Education and Employment Outcomes

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) |
|--|---------------------------|----------------------|------------------------------------|-----------------------|-----------------------|-----------------------|---------------------------------------|-------------------------------|--|
| | Educational Attainment | Grade Progression | Appropriate Grade Completion | Completed 7 grades | Completed 8 grades | Completed 9 grades | Currently Enrolled w/ HS Degree | Worked in Non- Laborer Job | Enrolled or Worked in Non- Laborer Job |
| Years of PROGRESA Exposure | 0.10 (0.038)*** | 0.0099 (0.0098) | 0.013 (0.0076)* | 0.023 (0.011)** | 0.015 (0.011) | 0.020 (0.011)* | -0.0049 (0.0077) | -0.013 (0.016) | -0.013 (0.015) |
| Rainfall Shock | -0.67 (0.28)** | -0.12 (0.056)** | -0.12 (0.051)** | -0.21 (0.071)*** | -0.23 (0.072)*** | -0.25 (0.083)*** | -0.10 (0.053)* | -0.23 (0.13)* | -0.27 (0.13)** |
| Rainfall Shock x Exposure | 0.12 (0.053)** | 0.021 (0.011)* | 0.020 (0.010)* | 0.033 (0.014)** | 0.041 (0.013)*** | 0.047 (0.016)*** | 0.018 (0.017) | 0.10 (0.039)** | 0.10 (0.042)** |
| Standardized Household Income Score | 0.24 (0.11)** | 0.029 (0.023) | 0.011 (0.020) | 0.094 (0.029)*** | 0.084 (0.026)*** | 0.057 (0.026)** | 0.0038 (0.021) | 0.12 (0.040)*** | 0.12 (0.041)*** |
| Income Score x Exposure | -0.022 (0.021) | -0.0011 (0.0049) | 0.0037 (0.0042) | -0.011 (0.0054)** | -0.0094 (0.0048)* | -0.0040 (0.0047) | 0.000036 (0.0065) | -0.036 (0.013)*** | -0.033 (0.013)*** |
| Observations | 11813 | 11205 | 11813 | 10058 | 8277 | 6612 | 1596 | 1142 | 1137 |
| Mean of Dependent Variable | 6.79 | 0.58 | 0.46 | 0.56 | 0.52 | 0.45 | 0.061 | 0.35 | 0.41 |
| Sample Ages (in 2003) | 12 to 18 | 12 to 18 | 12 to 18 | 13 to 18 | 14 to 18 | 15 to 18 | 18 | 18 | 18 |
| Fixed Effects | | | | | | Birth year x state | | | |

Notes:

- Standard errors clustered at the municipality are reported in parentheses *** p<0.01, ** p<0.05, * p<0.1.

- "Rainfall shock" = 1 for individuals whose birth-year rainfall was more than one standard deviation from the 10-year historical locality-specific mean.

- All specifications include gender, household head gender and age, household size, household composition variables, parental education, parental language, and locality characteristics. Controls for parental language/education and locality distance include dummies for missing values.

E Model

We lay out a simple model of optimal schooling choices that incorporates heterogeneous endowments and a schooling subsidy akin to PROGRESA. We derive our estimating equations from this model and then discuss what could be driving the remediation we find in our empirical analysis.

E.1 Setup

Our model extends the canonical schooling choice model in Card (2001) by allowing individuals to have heterogeneous initial endowments that affect future earnings. We study how the optimal level of schooling changes with the initial endowment and with education policies that attempt to offset the cost of schooling. We assume that individuals have an infinite time horizon: they attend school during the first S periods of life and work full-time for the rest of it. Individuals have an initial level of endowment, ω , that affects the earnings function in each period. While in school, the utility in period t depends on the level of consumption, $u(c(t))$,⁶ and the effort cost for the t -th year of schooling given the initial endowment ω , $\phi(\omega, t)$: specifically, in-school utility is equal to $u(c(t)) - \phi(\omega, t)$. Out of school, the utility is $u(c(t))$. Finally, individuals discount future flows at a rate ρ .⁷

Conditional on schooling S and a consumption profile, life-cycle utility is

$$V(S, c(t)) = \int_0^S [u(c(t)) - \phi(\omega, t)] e^{-\rho t} dt + \int_S^\infty u(c(t)) e^{-\rho t} dt.$$

Let $y(\omega, S, t)$ be the earnings function at period t of an individual with initial endowment ω , S years of schooling, and $t \geq S$ years of post-schooling experience. We assume that while in school, individuals pay tuition costs minus the scholarship provided by PROGRESA at each period of time, $T(t) - x(t)$, and work part-time earning $P(t)$ in period t . Individuals borrow or lend at a fixed interest rate R . Thus the intertemporal budget constraint is

⁶We assume that $u(\cdot)$ is increasing and concave.

⁷Let \mathcal{S} be the set of feasible values of schooling, X the set of feasible values of PROGRESA and Ω the set of feasible values of initial endowment.

$$\int_0^\infty c(t) e^{-Rt} dt \leq \int_0^S [P(t) - T(t) + x(t)] e^{-Rt} dt + \int_S^\infty y(\omega, S, t) e^{-Rt} dt.$$

We introduce the following functional assumptions, which help us characterize the optimal level of schooling and resultant income:

(A1): $y(\omega, S, t) \equiv f(\omega, S) h(t - S)$, where $h(0) = 1$ and $f(\cdot)$ is a CES production function, $f(w, S) = A[\lambda w^k + (1 - \lambda)S^k]^{\frac{1}{k}}$ with $k \in (-\infty, 1]$, $k \neq 0$.⁸

(A2): $\phi(\omega, t) e^{-\rho t}$ is increasing and convex with respect to t , and decreasing with respect to ω .

(A3): $u(c(t)) = \log c(t)$.

(A4): $T(t) - P(t) - x(t) = T - P - x$ for all t .

(A1) assumes that the log of earnings is additively separable in years of post-schooling experience, and a function of an individual's education and initial endowment. (A1) also implies that the earnings are increasing with respect to the initial endowment and the level of schooling, and exhibit decreasing marginal returns to schooling for all levels of the initial endowment.⁹ (A2) is self-evident. (A3) is the standard assumption of log utility. (A4) states that tuition, part-time earnings and PROGRESA subsidies are constant over time.

Proposition 1. There exists a unique S^* , where $S^* \equiv S^*(\omega, x)$ is implicitly defined by:

$$\Gamma(\omega, S^*) - \left(\frac{1}{H(R)} + \frac{T - P - x}{f(\omega, S^*) H(R)} + \rho \phi(\omega, S^*) e^{-\rho S^*} \right) = 0, \quad (2)$$

where $\Gamma(\omega, S) \equiv \frac{f_S(\omega, S)}{f(\omega, S)}$, which is the marginal return to schooling.¹⁰

⁸The results in this section can be generalized by assuming that $f(\cdot)$ is increasing with respect to ω and S , and concave with respect to S , for all ω , and that $\lim_{S \rightarrow 0^+} f_S(\cdot) = \infty$ and $\lim_{S \rightarrow +\infty} f_S(\cdot) = 0$ for all $\omega \in \Omega$.

⁹Note also that $f(\cdot)$ is log-supermodular with respect to (S, ω) if and only if $k < 0$, and log-submodular with respect to (S, ω) if and only if $k > 0$.

¹⁰In order to simplify the discussion we follow Card (2001) and impose three additional assumptions: first, we assume that tuition cost minus part-time earnings minus PROGRESA subsidy are constant over time. Second, we assume that tuition costs minus part-time earnings minus PROGRESA are small relative to lifetime earnings. Finally, we assume that the cost of schooling standardized by the life time earnings is increasing with S , i.e., $d_S = -\frac{(T - P - x)f_S(\omega, S)}{f(\omega, S)^2 H(R)} + \rho \frac{\partial}{\partial S} (\phi(\omega, S) e^{-\rho S}) > 0, \forall S \in \mathbb{S}$.

Proof. The first order condition with respect to S is

$$-\phi(\omega, S) e^{-S\rho} + \lambda \left\{ [P(S) - T(S) + x(S)] e^{-RS} - y(\omega, S, S) e^{-RS} + \int_S^\infty \frac{\partial y(\omega, S, t)}{\partial S} e^{-Rt} dt \right\} = 0, \quad (3)$$

using (A1)

$$f_S(\omega, S) H(R) = f(\omega, S) + [T(s) - P(s) - x(S)] + \frac{\phi(\omega, S) e^{-(\rho-R)S}}{\lambda}, \quad (4)$$

where $H(R)$ is a decreasing function of the interest rate (e.g., following Card (2001), if earnings are fixed after completion of schooling $H(R) = \frac{1}{R}$). From the first order conditions with respect to the consumption, together with the budget constraint

$$\frac{1}{\lambda\rho} = \int_0^S [P(t) - T(t) + x(t)] e^{-Rt} dt + f(\omega, S) H(R) e^{-RS}. \quad (5)$$

Using (4), (5), and A4, S is implicitly defined by

$$\frac{f_S(\omega, S^*)}{f(\omega, S^*)} = \frac{1}{H(R)} + \frac{T - P - x}{f(\omega, S^*) H(R)} + \rho\phi(\omega, S^*) e^{-\rho S^*} \left[1 - \frac{1}{R} \frac{(T - P - x)(e^{RS^*} - 1)}{f(\omega, S^*) H(R)} \right]. \quad (6)$$

Next, given that tuition costs minus part-time earnings minus PROGRESA are small relative to lifetime earnings, then the term in square brackets is close to 1, i.e., $1 - \frac{1}{R} \frac{(T - P - x)(e^{RS^*} - 1)}{f(\omega, S^*) H(R)} \approx 1$. Then, S^* is uniquely defined by

$$\Lambda(x, \omega, S^*) \equiv \Gamma(\omega, S^*) - \underbrace{\left\{ \frac{1}{H(R)} + \frac{T - P - x}{f(\omega, S^*) H(R)} + \rho\phi(\omega, S^*) e^{-\rho S^*} \right\}}_{d(x, \omega, S^*)} = 0, \quad (7)$$

where $\Gamma(\omega, S) \equiv \frac{f_S(\omega, S)}{f(\omega, S)}$, which is the marginal return to schooling. Note that $\Lambda(x, \omega, S)$ is

decreasing with respect to S . Moreover, note that $\Lambda(x, \omega, S) > 0$ as $S \rightarrow 0$, and as $S \rightarrow \infty$, $\Lambda(x, \omega, S) \rightarrow -\infty$. Thus, given $(x, \omega) \in X \times \Omega$, there exists a unique S^* that satisfies (7). By the Implicit Function Theorem, S^* can be expressed as a differentiable function of (ω, S) , $S^* \equiv S^*(\omega, x)$. This completes the proof. \square

Let $d(x, \omega, S) \equiv \frac{1}{H(R)} + \frac{T-P-x}{f(\omega, S)H(R)} + \rho\phi(\omega, S)e^{-\rho S}$, be the marginal cost of schooling standardized by lifetime earnings. From Proposition 1 it follows that the optimal level of schooling is increasing with respect to x , i.e., $\frac{\partial S^*}{\partial x} > 0$. Also, the optimal level of schooling is increasing with respect to the initial endowment, ω , i.e., $\frac{\partial S^*}{\partial \omega} > 0$ if $\Gamma_\omega(\omega, S^*) - d_\omega(x, \omega, S^*) > 0$.^{11,12}

E.2 Estimating Equations

The optimal level of schooling, $S^*(\omega, x)$, and the optimal level of income, $y^*(\omega, x)$, are thus non linear functions of PROGRESA, x , and the initial endowment, ω . To test empirically the predictions of the model we use the fact that $S^*(\omega, x)$ and $y^*(\omega, x)$ are differentiable functions to derive linear approximations for both functions.

Proposition 2. If $S^* \equiv S^*(\omega, x)$ is implicitly defined by (2) and $y(\omega, S, t)$ is a C^2 -function then S^* and $y^* \equiv y^*(\omega, x)$ can be approximated by the polynomials

$$S^* = a_1\omega + a_2x + a_3\omega \cdot x + \varepsilon_s(\omega, x) \quad (8)$$

$$y^* = b_1\omega + b_2x + b_3\omega \cdot x + \varepsilon_y(\omega, x) \quad (9)$$

where $\varepsilon_s(\omega, x) = a_4\omega^2 + a_5x^2 + o(\|\omega, x\|^2)$ and $\varepsilon_y(\omega, x) = b_4\omega^2 + b_5x^2 + o(\|\omega, x\|^2)$.¹³

¹¹For the rest of this appendix we assume that $\Gamma_\omega(\omega, S^*) - d_\omega(x, \omega, S^*) > 0$.

¹²We also explore the effects of PROGRESA and the initial endowment on the equilibrium income, y^* , in the following corollary. From Proposition 1 it follows that the optimal level of income is increasing with respect to x , i.e., $\frac{\partial y^*}{\partial x} > 0$. Also, the optimal level of income is increasing with respect to the initial endowment, ω , i.e., $\frac{\partial y^*}{\partial \omega} > 0$ if $\Gamma_\omega(\omega, S^*) - d_\omega(x, \omega, S^*) > 0$.

¹³If $f(\omega, S)$ is approximated by a CES production function and we assume a cost function of the form $\phi(\omega, S) = [S + (\bar{\omega} - \omega)]e^{\rho S}$, for $\omega \in [0, \bar{\omega}]$, then from the Implicit Function Theorem and equation (2), $\left.\frac{\partial^2 S^*}{\partial x^2}\right|_{(T-p-x)=0} = o(R)$ when $R \rightarrow 0$. Thus, a_5 is close to 0. Similarly, for b_5 (i.e., $\left.\frac{\partial^2 y^*}{\partial x^2}\right|_{(T-p-x)=0} = o(R)$ when $R \rightarrow 0$).

Proof. From equation (7) and the Implicit Function Theorem we know that $S^* \equiv S^*(\omega, x)$ is as smooth as $\Lambda(x, \omega, S)$, and this last function is a C^2 -function since $f(\omega, S)h(t - S)$ is also C^2 -function. Therefore $S^*(\omega, x)$ is a C^2 -function and by the Multivariate Taylor Theorem it can be approximated by a polynomial plus an error:

$$S^*(\omega, x) = a_1\omega + a_2x + a_3\omega \cdot x + \varepsilon(\omega, x) \quad (10)$$

where $\varepsilon_s(\omega, x) = a_4\omega^2 + a_5x^2 + o(||(\omega, x)||^2)$.

Similarly, from the Implicit Function Theorem there exist neighborhoods $U \subset \mathcal{S}$ and $W \subset X \times \Omega$ of S^* and (x, ω) , respectively, on which there is a function $\xi : W \rightarrow U$ such that $(x, \omega, \xi(x, \omega))$ satisfy (2) for all $(x, \omega) \in W$. Then, $y^*(\omega, x) = y(\omega, \xi(x, \omega), t)$ is a C^2 -function and by the Multivariate Taylor Theorem it can be approximated by a polynomial plus an error:

$$y^* = b_1\omega + b_2x + b_3\omega \cdot x + \varepsilon_y(\omega, x) \quad (11)$$

where $\varepsilon_y(\omega, x) = b_4\omega^2 + b_5x^2 + o(||(\omega, x)||^2)$. \square

E.3 Remediation

Our empirical strategy uses adverse rainfall as a (negative) proxy for the initial endowment, i.e., $a_1\omega \approx a_1(-R)$, where R is an adverse rainfall shock. If outcomes are increasing in the endowment, as predicted above, we expect: $a_1 \approx -\beta_1 > 0$ and $b_1 \approx -\tilde{\beta}_1 > 0$, where β_1 and $\tilde{\beta}_1$ are the coefficients for the rainfall shock in our main estimating equation (1) for education and employment outcomes, respectively. Similarly, more years of PROGRESA should increase the optimal level of schooling and income, i.e., $a_2 \approx \beta_2 > 0$ and $b_2 \approx \tilde{\beta}_2 > 0$, where β_2 and $\tilde{\beta}_2$ are the coefficients for PROGRESA in our main estimating equation (1) for education and employment outcomes, respectively. Finally, note that $a_3\omega \cdot x \approx a_3(-R) \cdot x$. Thus, if $a_3 \approx -\beta_3 < 0$ and $b_3 \approx -\tilde{\beta}_3 < 0$, PROGRESA generates remediation in both schooling and income, where β_3 and $\tilde{\beta}_3$ are the coefficients for the interaction of PROGRESA and the

rainfall shock in our main estimating equation (1) for education and employment outcomes, respectively. If the opposite is true (that is, if $a_3 \approx -\beta_3 > 0$ and $b_3 \approx -\tilde{\beta}_3 > 0$) PROGRESA generates reinforcement.

Our empirical results yield negative main effects for the adverse negative shock ($a_1 \approx -\hat{\beta}_1 > 0$, $b_1 \approx -\hat{\beta}_1 > 0$), positive main effects for PROGRESA ($a_2 > 0$, $b_2 > 0$), and positive cross-partial ($a_3 \approx -\hat{\beta}_3 < 0$ and $b_3 \approx -\hat{\beta}_3 < 0$), which means that PROGRESA generates remediation in both schooling and income. To explore what drives the remediation we find, we use this model to study how PROGRESA affects the optimal level of schooling and income, S^* and y^* , respectively, differently for individuals with different levels of the initial endowment.

From the Implicit Function Theorem it follows that $\frac{\partial^2 S^*}{\partial \omega \partial x}$ is given by:

$$\frac{\partial^2 S^*}{\partial \omega \partial x} = \alpha_1 [\Gamma_{S\omega}(\omega, S^*) + \alpha_2 \Gamma_{SS}(\omega, S^*)] + \Theta(x, \omega, S^*), \quad (12)$$

where $\Gamma_{S\omega}(\omega, S) \equiv \frac{\partial^2}{\partial \omega \partial S} \Gamma(S, x, \omega)$ and $\Theta(x, \omega, S) \equiv -\frac{\Lambda_\omega}{\Lambda_S^2} d_{xS} + \frac{1}{\Lambda_S} d_{x\omega} + \rho \frac{\Lambda_x \Lambda_\omega}{\Lambda_S^3} \frac{\partial^2}{\partial^2 S} (\phi(\omega, S) e^{-\rho S})$. Moreover, $\alpha_1 > 0$, $\alpha_2 > 0$ and $\Theta(x, \omega, S^*) < 0$.¹⁴

Note that if the expression in (12) is negative, PROGRESA generates remediation – larger PROGRESA effects (on the optimal level of schooling, S^*) for lower-endowment individuals. If this expression is positive, PROGRESA generates reinforcement – larger effects for higher-endowment individuals. Similarly, if (13) is negative (positive) PROGRESA generates remediation (reinforcement) on the optimal level of income, y^* , for lower (higher)-endowment individuals.

Given that our empirical results imply that the expression in (12) is negative, we now examine why this is the case. This expression depends on features of the production function (i.e., $\alpha_1 [\Gamma_{S\omega}(\omega, S^*) + \alpha_2 \Gamma_{SS}(\omega, S^*)]$) and features of the marginal cost function (i.e.,

¹⁴Also, let $\xi : W \rightarrow U$, a function such that $(x, \omega, \xi(x, \omega))$ satisfy (2) for all $(x, \omega) \in W$.¹⁵ Then, $\frac{\partial^2 y^*}{\partial \omega \partial x}$ is given by the following expression:

$$\left[\underbrace{f_{SS}(\omega, \xi(x, \omega))}_{<0} \underbrace{\frac{\partial \xi(x, \omega)}{\partial \omega}}_{>0} + f_{S\omega}(\omega, \xi(x, \omega)) \right] \underbrace{\frac{\partial \xi(x, \omega)}{\partial x}}_{>0} + \underbrace{f_S(\omega, \xi(x, \omega))}_{>0} \underbrace{\frac{\partial^2 \xi(x, \omega)}{\partial \omega \partial x}}_{>0}. \quad (13)$$

$\Theta(x, \omega, S^*)$). We begin with the latter. The first two terms of $\Theta(x, \omega, S^*)$, $-\frac{\Lambda_\omega}{\Lambda_S^2} d_{xS} + \frac{1}{\Lambda_S} d_{x\omega}$, capture how PROGRESA changes the cost of schooling (relative to foregone earnings) differently for high-endowment and low-endowment individuals. These two terms are negative, which means that PROGRESA decreases the relative cost of schooling more for lower-endowment individuals, and therefore generates larger increases in optimal schooling for these lower-endowment individuals compared to higher-endowment individuals. The third term of $\Theta(x, \omega, S^*)$, $\rho \frac{\Lambda_x \Lambda_\omega}{\Lambda_S^3} \frac{\partial^2}{\partial^2 S} (\phi(\omega, S^*) e^{-\rho S^*})$, is negative due to the convexity of the cost function and the fact that $\frac{\Lambda_x \Lambda_\omega}{\Lambda_S^3} < 0$. In other words, convex costs will make it more difficult to change optimal schooling levels for a high-endowment individual (who obtains more schooling) compared to a low-endowment individual.

With respect to the first term of expression (12), $\alpha_1 [\Gamma_{S\omega}(\omega, S^*) + \alpha_2 \Gamma_{SS}(\omega, S^*)]$: Γ_{SS} depends on the third derivative of $\log f(\cdot)$, which is the rate at which studying an extra year becomes less appealing; and $\Gamma_{S\omega}(\omega, S)$ measures how the curvature of the log earnings function with respect to schooling changes with the initial endowment.¹⁶

In sum, PROGRESA could be generating remediation for one or more of the following three reasons. First, the transfer represents a larger share of foregone earnings for low endowment compared to high endowment children. Second, convex schooling costs would make it more difficult to shift a high-endowment's (higher) level of optimal schooling. Finally, if the second derivative of the earnings function with respect to schooling is decreasing in the initial endowment, this would also contribute to remediation.

¹⁶Whether we see remediation or reinforcement on the optimal level of income (equation 13), depends also on two terms. The second term, $f_S(\omega, \xi(x, \omega)) \frac{\partial^2 \xi(x, \omega)}{\partial \omega \partial x}$, depends on the size of the reinforcement or remediation effect on the optimal level of schooling. The first term (the bracket) depends also on two terms: the first term, $f_{SS}(\omega, \xi(x, \omega)) \frac{\partial \xi(x, \omega)}{\partial \omega}$, is negative, due to the concavity of the earnings function. The second bracketed term, $f_{S\omega}(\omega, \xi(x, \omega))$, is positive from A2. Thus, remediation on the optimal level of income depends on the size of the remediation effect on the optimal level of schooling, and on the total derivative of the marginal return of education on earnings, with respect the initial endowment.

References

- Andalón, M. (2011). Oportunidades to reduce overweight and obesity in mexico? *Health economics*, 20(S1):1–18.
- Angelucci, M. and De Giorgi, G. (2009). Indirect effects of an aid program: how do cash transfers affect ineligibles' consumption? *The American Economic Review*, 99(1):486–508.
- Barber, S. L. and Gertler, P. J. (2008). The impact of mexico's conditional cash transfer programme, oportunidades, on birthweight. *Tropical Medicine & International Health*, 13(11):1405–1414.
- Barham, T. (2011). A healthier start: the effect of conditional cash transfers on neonatal and infant mortality in rural mexico. *Journal of Development Economics*, 94(1):74–85.
- Barham, T. and Rowberry, J. (2013). Living longer: The effect of the mexican conditional cash transfer program on elderly mortality. *Journal of Development Economics*, 105:226–236.
- Behrman, J. R. and Hoddinott, J. (2005). Programme evaluation with unobserved heterogeneity and selective implementation: The mexican progresas impact on child nutrition. *Oxford bulletin of economics and statistics*, 67(4):547–569.
- Behrman, J. R., Parker, S. W., and Todd, P. E. (2009). Medium-term impacts of the oportunidades conditional cash transfer program on rural youth in mexico. *Poverty, Inequality and Policy in Latin America*, pages 219–70.
- Behrman, J. R., Parker, S. W., and Todd, P. E. (2011). Do conditional cash transfers for schooling generate lasting benefits? a five-year followup of progresas/oportunidades. *Journal of Human Resources*, 46(1):93–122.
- Card, D. (2001). Estimating the return to schooling: Progress on some persistent econometric problems. *Econometrica*, 69(5):1127–1160.
- Djebbari, H. and Smith, J. (2008). Heterogeneous impacts in progresas. *Journal of Econometrics*, 145(1):64–80.
- Fernald, L. C., Gertler, P. J., and Hou, X. (2008a). Cash component of conditional cash transfer program is associated with higher body mass index and blood pressure in adults. *The Journal of nutrition*, 138(11):2250–2257.
- Fernald, L. C., Gertler, P. J., and Neufeld, L. M. (2008b). Role of cash in conditional cash transfer programmes for child health, growth, and development: an analysis of mexico's oportunidades. *The Lancet*, 371(9615):828–837.
- Fernald, L. C. and Gunnar, M. R. (2009). Poverty-alleviation program participation and salivary cortisol in very low-income children. *Social science & medicine*, 68(12):2180–2189.
- Fernald, L. C., Hou, X., and Gertler, P. J. (2008c). Oportunidades program participation and body mass index, blood pressure, and self-reported health in mexican adults. *Prev Chronic Dis*, 5(3):A81.

- Gertler, P. (2004). Do conditional cash transfers improve child health? evidence from progesa's control randomized experiment. *American Economic Review*, 94(2):336–341.
- Gertler, P. J., Martinez, S. W., and Rubio-Codina, M. (2012). Investing cash transfers to raise long-term living standards. *American Economic Journal: Applied Economics*, pages 164–192.
- Goyal, D., Gay, C., and Lee, K. A. (2010). How much does low socioeconomic status increase the risk of prenatal and postpartum depressive symptoms in first-time mothers? *Women's Health Issues*, 20(2):96–104.
- Hoddinott, J. and Skoufias, E. (2004). The impact of progesa on food consumption. *Economic development and cultural change*, 53(1):37–61.
- Kim, P., Capistrano, C. G., Erhart, A., Gray-Schiff, R., and Xu, N. (2017). Socioeconomic disadvantage, neural responses to infant emotions, and emotional availability among first-time new mothers. *Behavioural brain research*, 325:188–196.
- Maccini, S. and Yang, D. (2009). Under the weather: Health, schooling, and economic consequences of early-life rainfall. *American Economic Review*, 99(3):1006–1026.
- Rivera, J. A., Sotres-Alvarez, D., Habicht, J.-P., Shamah, T., and Villalpando, S. (2004). Impact of the mexican program for education, health, and nutrition (progesa) on rates of growth and anemia in infants and young children: a randomized effectiveness study. *Jama*, 291(21):2563–2570.
- Schultz, T. P. (2004). School subsidies for the poor: evaluating the mexican progesa poverty program. *Journal of Development Economics*, 74(1):199–250.
- Skoufias, E., Davis, B., and De La Vega, S. (2001). Targeting the poor in mexico: an evaluation of the selection of households into progesa. *World development*, 29(10):1769–1784.
- Skoufias, E. and Parker, S. W. (2001). Conditional cash transfers and their impact on child work and schooling: Evidence from the progesa program in mexico. *Economia*, 2(1):45–96.