

LMMSE

1 LMMSE

1.1 MMSE \rightarrow LMMSE

1.1.1 Mean

The zero forcing estimation is

$$\hat{x} = R_x H^H (H^H R_x H + \sigma^2 I)^{-1} H^H y = R_x (H^H H R_x + \sigma^2 I)^{-1} y \quad (1)$$

The proof is given. Suppose $\hat{h} = Cy$, we can minimize

$$\begin{aligned} J(C) &= E\{\|x - \hat{x}\|^2\} \\ &= E\{(x - Cy)^H (x - Cy)\} \\ &= E\{x^H x - x^H Cy - y^H C^H x + y^H C^H Cy\} \\ &= E\{tr(h^H h) - tr(h^H Cy) - tr(y^H C^H h) + tr(y^H C^H Cy)\} \end{aligned} \quad (2)$$

Therefore, we calculate the derivative,

$$\begin{aligned} \frac{\partial J(C)}{\partial C} &= \frac{\partial tr(x^H x)}{\partial C} - \frac{\partial tr(x^H Cy)}{\partial C} - \frac{\partial tr(y^H C^H x)}{\partial C} + \frac{\partial tr(y^H C^H Cy)}{\partial C} = 0 \\ &= 0 - x^* y^T - 0 + \frac{\partial tr(Cy y^H C^H)}{\partial C} \\ 0 &= -x^* y^T + C^* (y y^H)^T \\ x^* y^T &= C^* (y y^H)^T \\ (x y^H)^* &= C^* (y y^H)^T \\ R_{xy}^* &= C^* R_y^T \end{aligned} \quad (3)$$

Here, R_y is a Hermitian matrix, so $R_y^T = R_y^*$. (3) can be written as,

$$\begin{aligned} R_{xy}^* &= C^* R_y^* \\ R_{xy} &= C R_y \\ C &= R_{xy} R_y^{-1} \end{aligned} \quad (4)$$

Here,

$$\begin{aligned} R_{xy} &= E\{x y^H\} = E\{x(Hx + z)^H\} \\ &= E\{x x^H H^H\} + E\{x z^H\} \\ &= E\{x x^H\} H^H \\ &= R_x H^H \end{aligned} \quad (5)$$

and,

$$\begin{aligned}
R_y &= E\{yy^H\} = E\{(Hx + z)(Hx + z)^H\} \\
&= E\{Hxx^H H^H\} + E\{zz^H\} \\
&= HR_x H^H + \sigma^2 I
\end{aligned} \tag{6}$$

We load (5) and (6) into (4) to estimate symbols,

$$\begin{aligned}
\hat{x} &= Cy = R_{xy} R_y^{-1} y \\
&= R_x H^H (H R_x H^H + \sigma^2 I)^{-1} y \\
&= R_x (H^H H R_x + \sigma^2 I)^{-1} H^H y
\end{aligned} \tag{7}$$

As with the result in (7),

$$\begin{aligned}
(H^H H R_x + \sigma^2 I) H^H &= H^H H R_x H^H + \sigma^2 I H^H \\
&= H^H H R_x H^H + H^H \sigma^2 I \\
&= H^H (H R_x H^H + \sigma^2 I) \\
(H^H H R_x + \sigma^2 I)^{-1} (H^H H R_x + \sigma^2 I) H^H &= (H^H H R_x + \sigma^2 I)^{-1} H^H (H R_x H^H + \sigma^2 I) \\
H^H &= (H^H H R_x + \sigma^2 I)^{-1} H^H (H R_x H^H + \sigma^2 I) \\
H^H (H R_x H^H + \sigma^2 I)^{-1} &= (H^H H R_x + \sigma^2 I)^{-1} H^H (H R_x H^H + \sigma^2 I)^{-1} \\
H^H (H R_x H^H + \sigma^2 I)^{-1} &= (H^H H R_x + \sigma^2 I)^{-1} H^H \\
R_x H^H (H R_x H^H + \sigma^2 I)^{-1} &= R_x (H^H H R_x + \sigma^2 I)^{-1} H^H
\end{aligned} \tag{8}$$

If we estimate R_x as a power of x , R_x is a diagonal matrix with the power of x . (7) can be written as

$$\hat{x} = \text{diag}(P_x) \cdot (H^H H \cdot \text{diag}(P_x) + \sigma^2 I)^{-1} H^H y, \tag{9}$$

where $P_x = [P_{x_0}, P_{x_1}, \dots, P_{x_n}]$ is a vector of the signal power for each x .

1.1.2 Variance

The error of LMMSE is

$$\begin{aligned}
D_e &= E\{(\hat{x} - x)(\hat{x} - x)^H\} \\
&= E\{\hat{x}\hat{x}^H\} - E\{x\hat{x}^H\} - E\{\hat{x}x^H\} + E\{xx^H\}
\end{aligned} \tag{10}$$

Please note that the error is orthogonal to the estimation (that is, the correlation matrix is all zero), i.e.,

$$\begin{aligned}
E\{e\hat{x}^H\} &= 0 \\
E\{(\hat{x} - x)\hat{x}^H\} &= 0 \\
E\{(\hat{x}\hat{x}^H - E\{x\hat{x}^H\})\} &= 0 \\
E\{\hat{x}\hat{x}^H\} &= E\{x\hat{x}^H\} \\
E\{(\hat{x}\hat{x}^H)^H\} &= E\{(x\hat{x}^H)^H\} \\
E\{\hat{x}\hat{x}^H\} &= E\{\hat{x}x^H\}
\end{aligned} \tag{11}$$

We load (11) into (10),

$$\begin{aligned}
cov(\hat{x}) &= E\{\hat{x}\hat{x}^H\} - E\{\hat{x}\hat{x}^H\} - E\{\hat{x}\hat{x}^H\} + E\{xx^H\} \\
&= E\{xx^H\} - E\{\hat{x}\hat{x}^H\} \\
&= R_x - R_x H^H R_y^{-1} y y^H R_y^{-1} H R_x \\
&= R_x - R_x H^H R_y^{-1} R_y R_y^{-1} H R_x \\
&= R_x - R_x H^H R_y^{-1} H R_x \\
&= R_x - R_x H^H (H R_x H^H + \sigma^2 I)^{-1} H R_x \\
&= R_x - R_x (H^H H R_x + \sigma^2 I)^{-1} H^H H R_x \\
&= R_x (I - (H^H H R_x + \sigma^2 I)^{-1} H^H H R_x) \\
&= R_x ((H^H H R_x + \sigma^2 I)^{-1} (H^H H R_x + \sigma^2 I) - (H^H H R_x + \sigma^2 I)^{-1} H^H H R_x) \\
&= R_x (H^H H R_x + \sigma^2 I)^{-1} (H^H H R_x + \sigma^2 I - H^H H R_x) \\
&= R_x (H^H H R_x + \sigma^2 I)^{-1} \sigma^2 I \\
&= \sigma^2 R_x (H^H H R_x + \sigma^2 I)^{-1}
\end{aligned} \tag{12}$$

Therefore, the variance is

$$var(\hat{x}) = \sigma^2 \cdot diag(R_x (H^H H R_x + \sigma^2 I)^{-1}) \tag{13}$$

1.2 Using the Estimated Channel Matrix

If we use the estimated channel matrix \hat{H} , the system model can be updated as

$$\begin{aligned}
y &= (\hat{H} + H - \hat{H})x + z \\
&= \hat{H}x + (H - \hat{H})x + z \\
&= \hat{H}x + \Delta H x + z \\
&= \hat{H}x + \tilde{z}
\end{aligned} \tag{14}$$

As $(H - \hat{H})x$ and z are independent, the covariance of \tilde{z} is

$$\begin{aligned}
cov(\tilde{z}) &= E\{\Delta H x (\Delta H x)^H\} + E\{z z^H\} \\
&= E\{\Delta H x x^H \Delta H^H\} + \sigma^2 I
\end{aligned} \tag{15}$$

the error is

$$\begin{aligned}
e &= \hat{x} - x = C y - x \\
&= C(\hat{H}x + z + Hx - \hat{H}x) - x \\
&= C(\hat{H}x + z) + C(H - \hat{H})x \\
&= \hat{x} + C(H - \hat{H})x
\end{aligned} \tag{16}$$