

Variational Bayes

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2025 年 12 月 15 日

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1 Common Distributions

1.1 Gamma Distribution

Supposing we have a gamma distribution $p(x; a, b)$. a 表示事件次数, b 表示每次发生的概率。The probability density function (PDF) of gamma distribution is

$$p(x) = \frac{x^{(a-1)} b^a e^{-bx}}{\Gamma(a)}$$

$$\ln p(x) \propto (a-1)\ln(x) - bx$$

The mean is

$$\mu_x = \frac{a}{b}$$

The variance is

$$\sigma_x^2 = \frac{a}{b^2} \quad (1)$$

1.2 Complex Gaussian Distribution

- For a complex Gaussian distributed variable x , its PDF is

$$p(x) = \frac{1}{\pi\sigma^2} e^{-\frac{|x-\mu|^2}{\sigma^2}}$$

$$\ln p(x) \propto -\ln(\sigma^2) - \sigma^{-2}|x - \mu|^2 \quad (2)$$

- For a complex Gaussian distributed variable vector $x = [x_0, \dots, x_{N-1}]$, its PDF is

$$p(x) = \frac{1}{\pi^N \det(\Sigma)} e^{-(x-\mu)^H \Sigma^{-1} (x-\mu)}$$

$$\ln p(x) \propto -\ln \det(\Sigma) - (x - \mu)^H \Sigma^{-1} (x - \mu) \quad (3)$$

where Σ is the covariance matrix.

2 Variational Bayes

For a telecommunication system, we have

$$y = Hx + z, \quad (4)$$

where y is the received signal, x is the transmitted signal and $z \in \mathcal{CN}(0, \sigma^2)$. Please note that,

$$x = x_p + x_d, \quad (5)$$

where x_p is the pilot and x_d is data.

2.1 Bayes Interference

In the Rx, given the prior $p(y)$, we compute posterior distribution $p(y|x)$,

$$p(x|y) = \frac{p(x, y)}{p(y)} = \frac{\overbrace{p(y|x)}^{likelihood} \overbrace{p(x)}^{prior}}{\underbrace{p(y)}_{evidence}} = \frac{p(y|x)p(x)}{\int_y p(x, y) dy} \quad (6)$$

Usually, we assume the evidence is 100%, i.e., $p(y) = 1$. Hence,

$$p(x|y) \propto \overbrace{p(y|x)}^{likelihood} \overbrace{p(x)}^{prior} \quad (7)$$

Here, we need to choose the likelihood and the prior .

2.2 Variational Interference

However, the posterior may have no closed form, i.e., computing $p(x|y)$ is not feasible. Instead, we use a distribution Q over the symbols x to approximate $p(x|y)$, i.e.,

$$\begin{aligned} q^*(x) &= \arg \min_{q(x) \in Q} KL(q(x)||p(x|y)) \\ &= \arg \min_{q(x) \in Q} \int_x q(x) \ln \frac{q(x)}{p(x|y)} dx \\ &= \arg \min_{q(x) \in Q} - \int_x q(x) \ln \frac{p(x|y)}{q(x)} dx \end{aligned} \quad (8)$$

Here, $q^*(x)$ is the optimal $q(x)$ but $p(x|y)$ is unknown. Therefore,

$$KL(q(x)||p(x|y)) = - \int_x q(x) \ln \frac{p(x|y)}{q(x)} dx \quad (9)$$

$$= \int_x q(x) \ln q(x) dx - \int_x q(x) \ln p(x|y) dx \quad (10)$$

$$= \int_x q(x) \ln q(x) dx - \int_x q(x) \ln \frac{p(x,y)}{p(y)} dx \quad (11)$$

$$= \int_x q(x) \ln q(x) dx - \int_x q(x) \ln p(x,y) dx + \int_x q(x) \ln p(y) dx \quad (12)$$

$$= \int_x q(x) \ln q(x) dx - \int_x q(x) \ln p(x,y) dx + \ln p(y) \int_x q(x) dx \quad (13)$$

$$= \underbrace{\mathbb{E}_q[\ln q(x)] - \mathbb{E}_q[\ln p(x,y)]}_{-ELBO} + \ln p(y) \quad (14)$$

$$= -ELBO(q) + \ln p(y) \quad (15)$$

2.2.1 ELBO

Here, ELBO is Evidence Lower Bound, i.e.,

$$ELBO(q) = \mathbb{E}_q[\ln p(x,y)] - \mathbb{E}_q[\ln q(x)] \quad (16)$$

$$= \int_x q(x) \ln \overbrace{\frac{p(x,y)}{q(x)}}^{known} dx \quad (17)$$

$$= \int_x q(x) \ln \frac{p(y|x)p(x)}{q(x)} dx \quad (18)$$

(15) can be rewritten as,

$$\begin{aligned} \ln p(y) &= \underbrace{ELBO(q)}_{\text{CONST}} + \underbrace{KL(q(x)||p(x|y))}_{\geq 0} \\ &\geq ELBO(q) \end{aligned} \quad (19)$$

The minimizing KL can be taken as the maximizing ELBO, i.e.,

$$\begin{aligned} q^*(x) &= \arg \min_{q(x) \in Q} KL(q(x)||p(x|y)) \\ &= \arg \max_{q(x) \in Q} ELBO(q) \end{aligned} \quad (20)$$

2.3 Mean Field

Now, we know the problem has been simplified as the maximizing ELBO. Here, we use the mean-field assumption, i.e.,

$$\begin{aligned} q(x) &= \prod_{i=1}^m q_i(x_i) \\ \ln q(x) &= \sum_{i=1}^m \ln q_i(x_i) \\ \mathbb{E}_q[\ln q(x)] &= \sum_{i=1}^m \mathbb{E}_{q_i}[\ln q_i(x_i)] \end{aligned} \quad (21)$$

2.3.1 Coordinate Ascent Optimization

In $q = [q_1, q_2, \dots, q_j, \dots, q_m]$, we fix others to update q_j , i.e.,

$$\begin{aligned} q_j^*(x_j) &= \arg \min_{q_j} \text{ELBO}(q_j) \\ &= \frac{\exp\{\mathbb{E}_{q_{-j}}[\ln p(x, y)]\}}{\int_{x_j} \exp\{\mathbb{E}_{q_{-j}}[\ln p(x, y)]\} dx_j} \end{aligned} \quad (22)$$

To prove (22), we need to load (21) into (16),

$$\text{ELBO}(q) = \mathbb{E}_q[\ln p(x, y)] - \mathbb{E}_q[\ln q(x)] \quad (23)$$

$$= \int_x q(x) \ln p(x, y) dx - \left[\mathbb{E}_{q_j}[\ln q_j(x_j)] + \sum_{i \neq j} \mathbb{E}_{q_i}[\ln q_i(x_i)] \right] \quad (24)$$

Here, $\sum_{i \neq j} \mathbb{E}_{q_i}[\ln q_i(x_i)]$ can be seen as a constant because it is not related to q_j . Therefore, (24) can be simplified as ($*_{-j}$ represents the other elements except j),

$$\text{ELBO}(q) = \int_x q(x) \ln p(x, y) dx - \mathbb{E}_{q_j}[\ln q_j(x_j)] + \text{const} \quad (25)$$

$$= \int_x q(x) \ln p(x, y) dx - \int_{x_j} q(x_j) \ln q_j(x_j) dx_j + \text{const} \quad (26)$$

$$= \int_{x_j} \int_{x_{-j}} q(x_j) q(x_{-j}) \ln p(x, y) dx_j dx_{-j} - \int_{x_j} q(x_j) \ln q_j(x_j) dx_j + \text{const} \quad (27)$$

$$= \int_{x_j} q(x_j) \left[\int_{x_{-j}} q(x_{-j}) \ln p(x, y) dx_{-j} \right] dx_j - \int_{x_j} q(x_j) \ln q_j(x_j) dx_j + \text{const} \quad (28)$$

$$= \int_{x_j} q(x_j) \mathbb{E}_{q_{-j}}[\ln p(x, y)] dx_j - \int_{x_j} q(x_j) \ln q_j(x_j) dx_j + \text{const} \quad (29)$$

Here, we define a new distribution

$$\begin{aligned} \ln \tilde{p}_j(x_j, y) &= \mathbb{E}_{q_{-j}}[\ln p(x, y)] + \text{const} \\ \tilde{p}_j(x_j, y) &\propto \exp\{\mathbb{E}_{q_{-j}}[\ln p(x, y)]\} \end{aligned} \quad (30)$$

Here, we load (30) into (29),

$$\text{ELBO}(q) = \int_{x_j} q(x_j) \ln \tilde{p}_j(x_j, y) dx_j - \int_{x_j} q(x_j) \ln q_j(x_j) dx_j + \text{const} \quad (31)$$

$$= \int_{x_j} q(x_j) \ln \frac{\tilde{p}_j(x_j, y)}{\ln q_j(x_j)} dx_j + \text{const} \quad (32)$$

$$= -KL(q_j(x_j) || \tilde{p}_j(x_j, y)) \quad (33)$$

The KL divergence reaches the minimum when

$$\begin{aligned}
 q_{x_j}^* &= \tilde{p}_j(x_j, y) \\
 &\propto \exp\{\mathbb{E}_{q_{-j}}[\ln p(x, y)]\} \\
 &= \frac{\exp\{\mathbb{E}_{q_{-j}}[\ln p(x, y)]\}}{\int_{x_j} \exp\{\mathbb{E}_{q_{-j}}[\ln p(x, y)]\} dx_j}
 \end{aligned} \tag{34}$$

$$\int_{x_j} \exp\{\mathbb{E}_{q_{-j}}[\ln p(x, y)]\} dx_j \text{是为了让总体概率为1}$$

2.4 Algorithm Structure

The structure is given as below:

```

1: initialize  $q_j(x_j)$  for  $j \in 1, \dots, m$ 
2: while ELBO not converge do
3:   for  $j \in 1, \dots, m$  do
4:      $q_{x_j}^* = \frac{\exp\{\mathbb{E}_{q_{-j}}[\ln p(x, y)]\}}{\int_{x_j} \exp\{\mathbb{E}_{q_{-j}}[\ln p(x, y)]\} dx_j}$ 
5:   end for
6:   ELBO(q)=  $\mathbb{E}_q[\ln p(x, y)] - \mathbb{E}_q[\ln q(x)]$ 
7: end while
8: return  $q(x)$ 

```

3 System Model

We first introduce the OTFS mod/demod and its frame structure. Subsequently, we derive the input-output relation in the delay-Doppler (DD) domain for the two most widely adopted pulse-shaping waveforms.

3.1 OTFS System

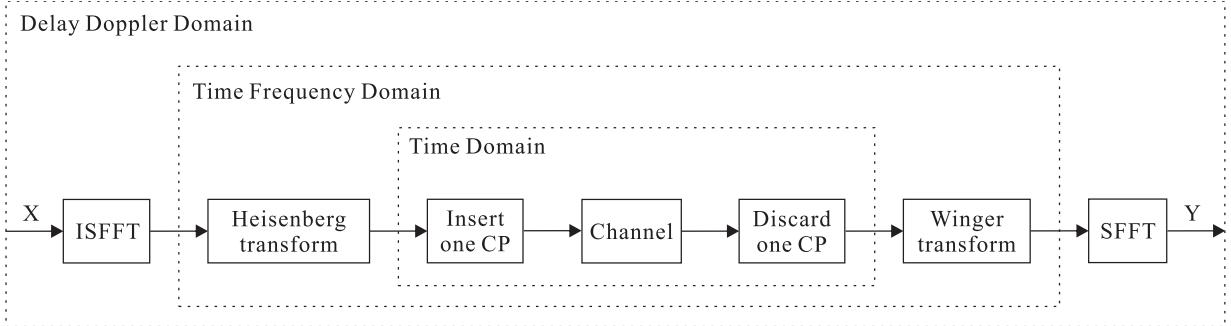


图 1: OTFS mode/demod

We consider a single input single output (SISO) OTFS system as illustrated in Fig. 1. The transmitter operates an OTFS frame (detailed in 3.2), $\mathbf{X}[k, l] \in \mathbb{C}^{K \times L}$, with $k = 0, \dots, K - 1$ and $l = 0, \dots, L - 1$ indexing discretized Doppler and delay shifts, respectively. After transposition, the frame is converted to the time-frequency (TF) domain via the inverse symplectic finite Fourier transform (ISFFT), mapping the data on $L \times K$ grids with uniform intervals Δf (Hz) and $T = 1/\Delta f$ (seconds). The time-domain signal is synthesized using (discrete) Heisenberg transform with a pulse-shaping waveform employing a single initial cyclic prefix spanning the full OTFS frame duration. The time-domain signal is transmitted over

a time-varying wireless channel characterized by the delay-Doppler impulse response $h(\tau, v)$ as [1],

$$h(\tau, v) = \sum_{i=1}^P h_i \delta(\tau - \tau_i) \delta(v - v_i), \quad (35)$$

where $\delta(\cdot)$ denotes the Dirac delta function, $h_i \sim \mathcal{N}(0, \frac{1}{P})$ is the gain of the i -th propagation path, and P represents the total number of paths. Each path is characterized by distinct delay and/or Doppler shifts, modeling the channel response between the receiver and either moving reflectors or the transmitting source. The delay and Doppler shifts are given as,

$$\tau_i = l_i \frac{T}{L}, v_i = k_i \frac{\Delta f}{K}, \quad (36)$$

respectively. Let the integers $l_i \in [0, l_{\max}]$ and $k_i \in [-k_{\max}, k_{\max}]$ represent the delay and Doppler shift indices, respectively, where l_{\max} and k_{\max} denote the maximum delay index and maximum Doppler shift index across all propagation paths. Note that we restrict our consideration to integer-valued indices, as fractional delay and Doppler shifts can be equivalently represented through virtual integer taps in the delay-Doppler domain using the techniques described in [2–4].

3.2 OTFS Frame Structure

As illustrated in Fig. 2, a superimposed OTFS frame structure is considered, where pilot and data symbols are jointly embedded over delay-Doppler grids, i.e.,

$$\mathbf{X} = \mathbf{X}_d + \mathbf{X}_p, \quad (37)$$

where $\mathbf{X}_d[k, l] \in \mathbb{C}^{K \times L}$ denotes the data frame composed of quadrature amplitude modulation (QAM) symbols drawn from a constellation \mathcal{A} with average energy E_d . The pilot frame $\mathbf{X}_p[k, l]$ contains nonzero elements only at designated positions, i.e.,

$$\mathbf{X}_p[k, l] = \begin{cases} x_p, & k = k_p, l = l_p, \\ 0, & \text{otherwise,} \end{cases} \quad (38)$$

where x_p is the pilot symbol with energy E_p , $k_p = \lfloor (K-1)/2 \rfloor$ is the Doppler index of all pilots, and $l_p = i(l_{\max} + 1)$ for $i = 0, \dots, N_p - 1$ are their delay indices. Here, $N_p = \lfloor L/(l_{\max} + 1) \rfloor$ denotes the total number of pilots. Each pilot facilitates channel estimation over a region of size $K \times (l_{\max} + 1)$ in the DD domain.

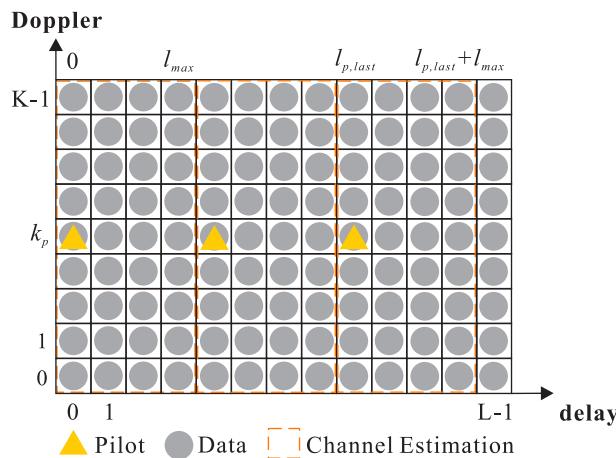


图 2: OTFS frame structure with the last pilot delay index at $l_{p,\text{last}} = (N_p - 1)(l_{\max} + 1)$

3.2.1 Inital Channel Estimation

In this section, we have two methods: the first, we average all channel estimation area together and do absolute square; the second, we do absolute square and average. We will introduce two thresholds individually.

- First

$$Y_{CHE}[k, l] = \frac{1}{N_b} \sum_{i=0}^{N_b-1} Y_{DD}[k, l + (l_{\max} + 1)(i - 1)] \quad (39)$$

where $k \in [0, K]$ and $l \in [0, l_{\max}]$ and the variance of $Y_{CHE}[k, l]$ is $\sigma_{CHE}^2 = \frac{1+\sigma^2}{N_b}$

$$|\eta[k, l]|^2 = |Y_{CHE}[k, l]|^2 \quad (40)$$

$$\frac{|\eta[k, l]|^2}{\sigma_{CHE}^2/2} = \left| \frac{\text{real}(Y_{CHE}[k, l])}{\sigma_{CHE}/\sqrt{2}} \right|^2 + \left| \frac{\text{imag}(Y_{CHE}[k, l])}{\sigma_{CHE}/\sqrt{2}} \right|^2 \quad (41)$$

Therefore, $\frac{|\eta[k, l]|^2}{\sigma_{CHE}^2/2}$ is a chi-square distribution of freedom 2. We set the a threshold ρ_0 and its probability p_0 (assuming it is noise and data), i.e.,

$$p\left(\frac{|\eta[k, l]|^2}{\sigma_{CHE}^2/2} \leq \rho_0\right) = p_0 \quad (42)$$

$$\rho_0 = \text{chi2inv}(p_0, 2) \quad (43)$$

$$\frac{|\eta[k, l]|^2}{\sigma_{CHE}^2/2} \leq \rho_0 \quad (44)$$

$$|\eta[k, l]|^2 \leq \rho_0 \sigma_{CHE}^2 / 2 \quad (45)$$

$$|\eta[k, l]|^2 \leq \rho_0 \frac{1+\sigma^2}{2N_b} \quad (46)$$

If $|\eta[k, l]|^2 \geq \rho_0 \frac{1+\sigma^2}{2N_b}$, we assume that it is not noise and data at a probability of $1 - p_0$. Then, we set another threshold ρ_1 and its probability p_1 (assuming it is a path where $\sigma_{CHE}^2 = \frac{1+\sigma^2}{N_b}$), i.e.,

$$p(|\eta[k, l]|^2 \geq \rho_1) \geq p_1 \quad (47)$$

$$p(|\eta[k, l]|^2 \leq \rho_1) \leq 1 - p_1 \quad (48)$$

$$\rho_1 = \text{chi2inv}(1 - p_1, 2) \quad (49)$$

$$\frac{|\eta[k, l]|^2}{(E_s/p + (1 + \sigma^2)/N_b)/2} \geq \rho_1 \quad (50)$$

$$|\eta[k, l]|^2 \geq \rho_1 \frac{E_s/p + (1 + \sigma^2)/N_b}{2} \quad (51)$$

To make sure that a path exist when it is not a noise

$$\rho_0 \frac{1 + \sigma^2}{2N_b} = \rho_1 \frac{E_s/p + (1 + \sigma^2)/N_b}{2} \quad (52)$$

$$\frac{\rho_0}{\rho_1} \frac{1 + \sigma^2}{N_b} = E_s/p + (1 + \sigma^2)/N_b \quad (53)$$

$$E_s/p = \left(\frac{\rho_0}{\rho_1} - 1 \right) \frac{1 + \sigma^2}{N_b} \quad (54)$$

- Second

$$|\eta[k, l]|^2 = \sum_{i=0}^{N_b-1} |Y_{DD}[k, l + (l_{\max} + 1)(i - 1)]|^2 \quad (55)$$

$$\frac{|\eta[k, l]|^2}{(1 + \sigma^2)/2} = \sum_{i=0}^{N_b-1} \frac{\text{real}(Y_{DD}[k, l + (l_{\max} + 1)(i - 1)])^2}{(1 + \sigma^2)/2} + \frac{\text{imag}(Y_{DD}[k, l + (l_{\max} + 1)(i - 1)])^2}{(1 + \sigma^2)/2} \quad (56)$$

Therefore, $\frac{|\eta[k,l]|^2}{(1+\sigma^2)/2}$ is a chi-square distribution of freedom $2N_b$. We set the a threshold ρ_0 and its probability p_0 (assuming it is noise and data), i.e.,

$$p\left(\frac{|\eta[k,l]|^2}{(1+\sigma^2)/2} \leq \rho_0\right) = p_0 \quad (57)$$

$$\rho_0 = chi2inv(p_0, 2N_b) \quad (58)$$

$$|\eta[k,l]|^2 \leq \rho_0(1 + \sigma^2)/2 \quad (59)$$

Similarly,

$$\rho_1 = chi2inv(1 - p_1, 2N_b) \quad (60)$$

$$|\eta[k,l]|^2 \geq \rho_1 \frac{E_s/p + 1 + \sigma^2}{2} \quad (61)$$

Therefore,

$$E_s/p = \left(\frac{\rho_0}{\rho_1} - 1\right)(1 + \sigma^2)$$

4 Variational Bayes in QJCHESD

4.1 OTFS Channel Estimation using VB

We assume the channel follows the Gaussian distribution, i.e.,

$$p(h|\gamma) = \prod_{i=0}^{P_{\max}-1} p(h_i|\gamma_i) = \prod_{i=0}^{P_{\max}-1} \mathcal{CN}(h_i; 0, \gamma_i^{-1}), \quad (62)$$

where $P_{\max} = (l_{\max} + 1)(2k_{\max} + 1)$, $\gamma = [\gamma_0, \gamma_1, \dots, \gamma_{P_{\max}-1}]$ is the precision vector of h . γ follows the Gamma distribution, i.e.,

$$p(\gamma) = \text{Gamma}(\gamma; a, b) \quad (63)$$

where a is the shape parameter and b is the inverse scale parameter. Also, we assume the noise obeys the Gaussian distribution, i.e.,

$$p(z) = \mathbf{z} \sim \mathcal{CN}(0, \alpha^{-1} \mathbf{I}) \quad (64)$$

where α obeys the Gamma distribution, i.e.,

$$p(\alpha) = \text{Gamma}(\alpha; c, d) \quad (65)$$

where c is the shape parameter and d is the inverse scale parameter. For all parameters, we define $\Theta = \{h, \gamma, \alpha\}$.

Here, we use the mean field assumption to estimate the channel, i.e.,

$$\begin{aligned} p(y, \Theta) &= p(y|h, \alpha)p(h|\gamma)p(\gamma)p(\alpha) \\ \ln p(y, \Theta) &= \ln p(y|h, \alpha) + \ln p(h|\gamma) + \ln p(\gamma) + \ln p(\alpha) \end{aligned} \quad (66)$$

Here, we use a distribution family Q over Θ to approximate $p(\Theta|y)$, i.e.,

$$\begin{aligned} p(\Theta|y) &= \arg \min_{q(\Theta) \in Q} KL(q(\Theta)||p(\Theta|y)) \\ &= \arg \min_{q(\Theta) \in Q} - \int_{\Theta} q(\Theta) \ln \frac{p(\Theta|y)}{q(\Theta)} d\Theta \end{aligned} \quad (67)$$

where $q(\Theta)$ follows the mean field assumption, i.e.,

$$q(\Theta) = q(h)q(\gamma)q(\alpha) \quad (68)$$

Therefore, we can update the probability functions as follows:

$$q^{(t+1)}(\alpha) \propto \exp(\mathbb{E}_{q_{-\alpha}^{(t)}}[\ln p(y, \Theta)]) \quad (69)$$

$$q^{(t+1)}(h) \propto \exp(\mathbb{E}_{q_h^{(t)}}[\ln p(y, \Theta)]) \quad (70)$$

$$q^{(t+1)}(\gamma) \propto \exp(\mathbb{E}_{q_{-\gamma}^{(t)}}[\ln p(y, \Theta)]) \quad (71)$$

The update is computed as follows

1) Update $q(\alpha)$

$$q^{(t+1)}(\alpha) \propto \exp(\mathbb{E}_{q_{-\alpha}^{(t)}}[\ln p(y, \Theta)]) \quad (72)$$

$$\ln q^{(t+1)}(\alpha) \propto \mathbb{E}_{q_{-\alpha}^{(t)}}[\ln p(y, \Theta)] \quad (73)$$

$$\ln q^{(t+1)}(\alpha) \propto \mathbb{E}_{q_h^{(t)}}[\ln p(y|h, \alpha)] + \ln p(\alpha) \quad (74)$$

Here,

$$\begin{aligned} p(y|h, \alpha) &= \mathcal{CN}(\mathbf{y}_p; \Phi_p \mathbf{h}, \alpha^{-1} \mathbf{I}) \\ &= \frac{1}{\pi^Z \det(\alpha^{-1} \mathbf{I})} e^{-(\mathbf{y}_p - \Phi_p \mathbf{h})^H (\alpha^{-1} \mathbf{I})^{-1} (\mathbf{y}_p - \Phi_p \mathbf{h})} \end{aligned} \quad (75)$$

$$= \frac{1}{\pi^Z \det(\alpha^{-1} \mathbf{I})} e^{-\alpha (\mathbf{y}_p - \Phi_p \mathbf{h})^H (\mathbf{y}_p - \Phi_p \mathbf{h})} \quad (76)$$

$$= \frac{1}{\pi^Z \det(\alpha^{-1} \mathbf{I})} e^{-\alpha \|\mathbf{y}_p - \Phi_p \mathbf{h}\|^2} \quad (77)$$

where Z is the dimension of \mathbf{y}_p . Here,

$$\det(\alpha^{-1} \mathbf{I}) = (\alpha^{-1})^Z = \alpha^{-Z} \quad (78)$$

Therefore,

$$\ln p(y|h, \alpha) \propto Z \ln(\alpha) - \alpha \|\mathbf{y}_p - \Phi_p \mathbf{h}\|^2 \quad (79)$$

$$\mathbb{E}_{q_h^{(t)}}\{\ln p(y|h, \alpha)\} \propto \mathbb{E}_{q_h^{(t)}}\{Z \ln(\alpha)\} - \mathbb{E}_{q_h^{(t)}}\{\alpha \|\mathbf{y}_p - \Phi_p \mathbf{h}\|^2\} \quad (80)$$

$$\mathbb{E}_{q_h^{(t)}}\{\ln p(y|h, \alpha)\} \propto Z \ln(\alpha) - \alpha \mathbb{E}_{q_h^{(t)}}\{\|\mathbf{y}_p - \Phi_p \mathbf{h}\|^2\} \quad (81)$$

Now, we need to get $\mathbb{E}_{q_h^{(t)}}\{\|\mathbf{y}_p - \Phi_p \mathbf{h}\|^2\}$, i.e.,

$$\|\mathbf{y}_p - \Phi_p \mathbf{h}\|^2 = \mathbf{y}_p^H \mathbf{y}_p - \mathbf{y}_p^H \Phi_p \mathbf{h} - \mathbf{h}^H \Phi_p^H \mathbf{y}_p + \mathbf{h}^H \Phi_p^H \Phi_p \mathbf{h} \quad (82)$$

$$\mathbb{E}_{q_h^{(t)}}\{\|\mathbf{y}_p - \Phi_p \mathbf{h}\|^2\} = \mathbf{y}_p^H \mathbf{y}_p - \mathbf{y}_p^H \Phi_p \boldsymbol{\mu}_h - \boldsymbol{\mu}_h^H \Phi_p^H \mathbf{y}_p + \mathbb{E}_{q_h^{(t)}}\{\mathbf{h}^H \Phi_p^H \Phi_p \mathbf{h}\} \quad (83)$$

Here, as in [5],

$$\mathbb{E}_{q_h^{(t)}}\{\mathbf{h}^H \Phi_p^H \Phi_p \mathbf{h}\} = \boldsymbol{\mu}_p^H \Phi_p^H \Phi_p \boldsymbol{\mu}_p + \text{tr}(\Phi_p^H \Phi_p \boldsymbol{\Sigma}_h) \quad (84)$$

$$= \boldsymbol{\mu}_p^H \Phi_p^H \Phi_p \boldsymbol{\mu}_p + \text{tr}(\Phi_p \boldsymbol{\Sigma}_h \Phi_p^H) \quad (85)$$

Therefore,

$$\mathbb{E}_{q_h^{(t)}}\{\|\mathbf{y}_p - \Phi_p \mathbf{h}\|^2\} = \mathbf{y}_p^H \mathbf{y}_p - \mathbf{y}_p^H \Phi_p \boldsymbol{\mu}_h - \boldsymbol{\mu}_h^H \Phi_p^H \mathbf{y}_p + \boldsymbol{\mu}_p^H \Phi_p^H \Phi_p \boldsymbol{\mu}_p + \text{tr}(\Phi_p^H \Phi_p \boldsymbol{\Sigma}_h) \quad (86)$$

$$= \|\mathbf{y}_p - \Phi_p \boldsymbol{\mu}_h\|^2 + \text{tr}(\Phi_p^H \Phi_p \boldsymbol{\Sigma}_h) \quad (87)$$

Therefore,

$$\ln q^{(t+1)}(\alpha) \propto Z \ln(\alpha) - \alpha (\|\mathbf{y}_p - \Phi_p \boldsymbol{\mu}_h\|^2 + \text{tr}(\Phi_p^H \Phi_p \boldsymbol{\Sigma}_h)) + \ln p(\alpha) \quad (88)$$

$$\propto Z \ln(\alpha) - \alpha (\|\mathbf{y}_p - \Phi_p \boldsymbol{\mu}_h\|^2 + \text{tr}(\Phi_p^H \Phi_p \Sigma_h)) + (a-1) \ln(\alpha) - b\alpha \quad (89)$$

$$\propto (a+Z-1) \ln(\alpha) - \alpha (b + \|\mathbf{y}_p - \Phi_p \boldsymbol{\mu}_h\|^2 + \text{tr}(\Phi_p^H \Phi_p \Sigma_h)) \quad (90)$$

Therefore,

$$a^{(t+1)} = a^{(t)} + z \quad (91)$$

$$b^{(t+1)} = b^{(t)} + \|\mathbf{y}_p - \Phi_p \boldsymbol{\mu}_h^{(t)}\|^2 + \text{tr}(\Phi_p^H \Phi_p \Sigma_h^{(t)}) \quad (92)$$

where $\Sigma_h^{(t)}$ and $\boldsymbol{\mu}_h^{(t)}$ are the posterior covariance matrix and the posterior mean vector of $h^{(t)}$, which are both adjusted after updating $q(h)$. The mean of α is

$$\hat{\alpha}^{(t+1)} = \frac{a^{(t+1)}}{b^{(t+1)}} \quad (93)$$

2) Update $q(h)$

$$q^{(t+1)}(h) \propto \exp(\mathbb{E}_{q_{-h}^{(t)}}[\ln p(y, \Theta)]) \quad (94)$$

$$\ln q^{(t+1)}(h) \propto \mathbb{E}_{q_{-h}^{(t)}}\{\ln p(y, \Theta)\} \quad (95)$$

$$\propto \mathbb{E}_{q_{-h}^{(t)}}\{\ln p(y|h, \alpha)\} + \mathbb{E}_{q_{-h}^{(t)}}\{\ln p(h|\gamma^{(t)})\} \quad (96)$$

$$\propto -\mathbb{E}_{q_{-h}^{(t)}}\{\alpha\}\|\mathbf{y}_p - \Phi_p h\|^2 + \mathbb{E}_{q_{-h}^{(t)}}\{\gamma^{(t)}\}\|h\|^2 \quad (97)$$

$$\propto -\hat{\alpha}^{(t+1)}\|\mathbf{y}_p - \Phi_p h\|^2 - h^H \text{diag}(\gamma^{-1(t)})^{-1} h \quad (98)$$

$$\propto -\hat{\alpha}^{(t+1)}\|\mathbf{y}_p - \Phi_p h\|^2 - h^H \text{diag}(\gamma^{(t)}) h \quad (99)$$

$$\propto \hat{\alpha}^{(t+1)} h^H \Phi_p^H \mathbf{y}_p + \hat{\alpha}^{(t+1)} \mathbf{y}_p^H \Phi_p h - \hat{\alpha}^{(t+1)} h^H \Phi_h^H \Phi_h h - h^H \text{diag}(\gamma^{(t)}) h + \text{const} \quad (100)$$

$$\propto \hat{\alpha}^{(t+1)} h^H \Phi_p^H \mathbf{y}_p + \hat{\alpha}^{(t+1)} \mathbf{y}_p^H \Phi_p h - h^H (\hat{\alpha}^{(t+1)} \Phi_h^H \Phi_h + \text{diag}(\gamma^{(t)})) h + \text{const} \quad (101)$$

As in the assumption, h follows Guassian distribution, i.e.,

$$q^{(t+1)}(h) = \mathcal{CN}\left(h | \boldsymbol{\mu}_h^{(t+1)}, \Sigma_h^{(t+1)}\right) \quad (102)$$

$$\ln q^{(t+1)}(h) \propto -(h - \boldsymbol{\mu}_h^{(t+1)})^H \Sigma_h^{(t+1)^{-1}} (h - \boldsymbol{\mu}_h^{(t+1)}) \quad (103)$$

$$\propto \underbrace{h^H \Sigma_h^{(t+1)^{-1}} \boldsymbol{\mu}_h^{(t+1)}}_{\text{linear}} + \underbrace{\boldsymbol{\mu}_h^{(t+1)^H} \Sigma_h^{(t+1)^{-1}} h}_{\text{quadratic}} + \text{const} \quad (104)$$

Here, we can see that the covariance is

$$\Sigma_h^{(t+1)} = (\hat{\alpha}^{(t+1)} \Phi_h^H \Phi_h + \text{diag}(\gamma^{(t)}))^{-1} \quad (105)$$

Therefore, (101) can be written as

$$\begin{aligned} \ln q^{(t+1)}(h) &\propto h^H \Sigma_h^{(t+1)^{-1}} \left(\hat{\alpha}^{(t+1)} \Sigma_h^{(t+1)} \Phi_p^H \mathbf{y}_p \right) + \\ &\quad \left(\hat{\alpha}^{(t+1)} \Sigma_h^{(t+1)} \Phi_p^H \mathbf{y}_p \right)^H \Sigma_h^{(t+1)^{-1}} h - \\ &\quad h^H \Sigma_h^{(t+1)^{-1}} h + \text{const} \end{aligned} \quad (106)$$

Please note that the Hermitian matrix $\Sigma_h^{(t+1)^H} = \Sigma_h^{(t+1)}$. Therefore, the mean is

$$\boldsymbol{\mu}^{(t+1)} = \hat{\alpha}^{(t+1)} \Sigma_h^{(t+1)} \Phi_p^H \mathbf{y}_p$$

3) Update $q(\gamma)$

$$\ln q^{(t+1)}(\gamma) \propto \mathbb{E}_{q_{-\gamma}^{(t)}}\{\ln p(y, \Theta)\} \quad (107)$$

$$\propto \mathbb{E}_{q_{-\gamma}^{(t)}} \{ \ln p(y|h, \alpha) + \ln p(h|\gamma) + \ln p(\gamma) + \ln p(\alpha) \} \quad (108)$$

$$\propto \mathbb{E}_{q_{-\gamma}^{(t)}} \{ \ln p(h|\gamma) + \ln p(\gamma) \} \quad (109)$$

$$\propto \sum_{l=0}^{P_{\max}-1} \left(\ln(\gamma_l) - \gamma_l \|h_l^{(t+1)}\|^2 + (c-1)\ln\gamma_l - d\gamma_l \right) \quad (110)$$

$$\propto \sum_{l=0}^{P_{\max}-1} \left(\underbrace{(c+1-1)\ln(\gamma_l) - \gamma_l}_{c^{t+1}} \underbrace{(d + \|h_l^{(t+1)}\|^2)}_{d^{(t+1)}} \right) \quad (111)$$

where $\|h_l^{(t+1)}\|^2 = \Sigma_{h_l}^{t+1} + |\mu_{h_l}^{t+1}|^2$. Therefore,

$$\gamma = \text{diag}\left(\frac{c_0^{t+1}}{d_0^{(t+1)}}, \dots, \frac{c_{P_{\max}}^{t+1}}{d_{P_{\max}}^{(t+1)}}\right) \quad (112)$$

4) Initial values

$$a^{(0)} = b^{(0)} = 1 \quad (113)$$

$$c^{(0)} = d^{(0)} = 1 \quad (114)$$

4.1.1 OTFS Channel Estimation using VB known channel positions

4.1.2 Simplified OTFS Channel Estimation using VB

If we assume the noise is known to us, the estimation process can be simplified as,

1) Update $q(h)$

$$q^{(t+1)}(h) \propto \exp(\mathbb{E}_{q_h^{(t)}} [\ln p(y, \Theta)]) \quad (115)$$

$$\ln q^{(t+1)}(h) \propto \mathbb{E}_{q_h^{(t)}} \{ \ln p(y, \Theta) \} \quad (116)$$

$$\propto \mathbb{E}_{q_h^{(t)}} \{ \ln p(y|h, \alpha) \} + \mathbb{E}_{q_h^{(t)}} \{ \ln p(h|\gamma^{(t)}) \} \quad (117)$$

$$\propto \alpha \mathbf{h}^H \Phi_p^H \mathbf{y}_p + \alpha \mathbf{y}_p^H \Phi_p \mathbf{h} - \mathbf{h}^H (\alpha \Phi_h^H \Phi_h + \text{diag}(\boldsymbol{\gamma}^{(t)})) \mathbf{h} + \text{const} \quad (118)$$

As in the assumption, \mathbf{h} follows Guassian distribution, i.e.,

$$q^{(t+1)}(\mathbf{h}) = \mathcal{CN}\left(\mathbf{h} | \boldsymbol{\mu}_h^{(t+1)}, \boldsymbol{\Sigma}_h^{(t+1)}\right) \quad (119)$$

$$\ln q^{(t+1)}(\mathbf{h}) \propto -(\mathbf{h} - \boldsymbol{\mu}_h^{(t+1)})^H \boldsymbol{\Sigma}_h^{(t+1)^{-1}} (\mathbf{h} - \boldsymbol{\mu}_h^{(t+1)}) \quad (120)$$

$$\propto \underbrace{\mathbf{h}^H \boldsymbol{\Sigma}_h^{(t+1)^{-1}} \boldsymbol{\mu}_h^{(t+1)} + \boldsymbol{\mu}_h^{(t+1)^H} \boldsymbol{\Sigma}_h^{(t+1)^{-1}} \mathbf{h}}_{\text{linear}} + \underbrace{\mathbf{h}^H \boldsymbol{\Sigma}_h^{(t+1)^{-1}} \mathbf{h}}_{\text{quadratic}} + \text{const} \quad (121)$$

Here, we can see that the covariance is

$$\boldsymbol{\Sigma}_h^{(t+1)} = (\alpha \Phi_h^H \Phi_h + \text{diag}(\boldsymbol{\gamma}^{(t)}))^{-1} \quad (122)$$

$$\boldsymbol{\mu}^{(t+1)} = \alpha \boldsymbol{\Sigma}_h^{(t+1)} \Phi_p^H \mathbf{y}_p \quad (123)$$

2) Update $q(\gamma)$

$$\ln q^{(t+1)}(\gamma) \propto \mathbb{E}_{q_{-\gamma}^{(t)}} \{ \ln p(y, \Theta) \} \quad (124)$$

$$\propto \mathbb{E}_{q_{-\gamma}^{(t)}} \{ \ln p(y|h, \alpha) + \ln p(h|\gamma) + \ln p(\gamma) + \ln p(\alpha) \} \quad (125)$$

$$\propto \mathbb{E}_{q_{-\gamma}^{(t)}} \{ \ln p(h|\gamma) + \ln p(\gamma) \} \quad (126)$$

$$\propto \sum_{l=0}^{P_{\max}-1} \left(\ln(\gamma_l) - \gamma_l \|h_l^{(t+1)}\|^2 + (c-1)\ln\gamma_l - d\gamma_l \right) \quad (127)$$

$$\propto \sum_{l=0}^{P_{\max}-1} \left(\underbrace{(c+1-1) \ln(\gamma_l)}_{c^{t+1}} - \gamma_l \underbrace{(d + \|h_l^{(t+1)}\|^2)}_{d^{(t+1)}} \right) \quad (128)$$

where $\|h_l^{(t+1)}\|^2 = \Sigma_{h_l}^{t+1} + |\mu_{h_l}^{t+1}|^2$. Therefore,

$$\gamma = \text{diag}\left(\frac{c_0^{t+1}}{d_0^{(t+1)}}, \dots, \frac{c_{P_{\max}}^{t+1}}{d_{P_{\max}}^{(t+1)}}\right) \quad (129)$$

4.1.3 Enhanced Simplified OTFS Channel Estimation using VB

Now, we assume $p(h_i) \sim \mathcal{CN}(\beta_i, \gamma_i^{-1})$ and $p(\beta_i) \sim \mathcal{CN}(\mu_{\beta_i}, \sigma_{\beta_i}^2)$. Therefore,

$$p(y, \Theta) = p(y|h, \alpha)p(y|\beta, \gamma)p(\beta)p(\gamma)$$

1. Update $q(h)$

$$q^{(t+1)}(h) \propto \exp(\mathbb{E}_{q_{-h}^{(t)}}[\ln p(y, \Theta)]) \quad (130)$$

$$\ln q^{(t+1)}(h) \propto \mathbb{E}_{q_{-h}^{(t)}}\{\ln p(y, \Theta)\} \quad (131)$$

$$\propto \mathbb{E}_{q_{-h}^{(t)}}\{\ln p(y|h, \alpha)\} + \mathbb{E}_{q_{-h}^{(t)}}\{\ln p(h|\beta^t, \gamma^{(t)})\} \quad (132)$$

$$\propto -\alpha\|\mathbf{y}_p - \Phi_p h\|^2 - (\mathbf{h} - \beta^{(t)})^H \text{diag}(\boldsymbol{\gamma}^{(t)}) (\mathbf{h} - \beta^{(t)}) \quad (133)$$

$$(134)$$

First, we open the left

$$-\alpha\|\mathbf{y}_p - \Phi_p h\|^2 \propto \alpha \mathbf{h}^H \Phi_p^H \mathbf{y}_p + \alpha \mathbf{y}_p^H \Phi_p \mathbf{h} - \alpha \mathbf{h}^H \Phi_h^H \Phi_h \mathbf{h} \quad (135)$$

Then, we open the right

$$(\mathbf{h} - \beta^{(t)})^H \text{diag}(\boldsymbol{\gamma}^{(t)}) (\mathbf{h} - \beta^{(t)}) \propto \boldsymbol{\beta}^{(t)H} \text{diag}(\boldsymbol{\gamma}^{(t)}) \mathbf{h} + \mathbf{h}^H \text{diag}(\boldsymbol{\gamma}^{(t)}) \boldsymbol{\beta}^{(t)} - \mathbf{h}^H \text{diag}(\boldsymbol{\gamma}^{(t)}) \mathbf{h} \quad (136)$$

Therefore,

$$q^{(t+1)}(h) \propto h^H (\alpha \Phi_p^H \mathbf{y}_p + \text{diag}(\boldsymbol{\gamma}^{(t)}) \boldsymbol{\beta}^{(t)}) + \mathbf{h}^H (\alpha \Phi_h^H \Phi_h + \text{diag}(\boldsymbol{\gamma}^{(t)})) \mathbf{h} \quad (137)$$

Here, we can see that the covariance is

$$\boldsymbol{\Sigma}_h^{(t+1)} = (\alpha \Phi_h^H \Phi_h + \text{diag}(\boldsymbol{\gamma}^{(t)}))^{-1} \quad (138)$$

$$\boldsymbol{\mu}^{(t+1)} = \boldsymbol{\Sigma}_h^{(t+1)} (\alpha \Phi_p^H \mathbf{y}_p + \text{diag}(\boldsymbol{\gamma}^{(t)}) \boldsymbol{\mu}_{\beta}^{(t)}) \quad (139)$$

2. Update $q(\gamma)$

$$\ln q^{(t+1)}(\gamma) \propto \ln p(h|\beta, \gamma) + \ln p(\gamma) \quad (140)$$

$$\propto \sum_{l=0}^{P_{\max}} \left(\ln(\gamma_l) - \gamma_l \|h_l^{(t+1)} - \beta_l\|^2 + (c-1) \ln \gamma_l - d \gamma_l \right) \quad (141)$$

$$\propto \sum_{l=0}^{P_{\max}} (c+1-1) \ln \gamma_l - \gamma_l (d + \|h_l^{(t+1)} - \beta_l\|^2) \quad (142)$$

$$\propto \sum_{l=0}^{P_{\max}} (c+1-1) \ln \gamma_l - \gamma_l (d + \|h_l^{(t+1)}\|^2 - \beta_l^H h - h_l^H \beta_l + \|\beta_l\|^2) \quad (143)$$

$$\propto \sum_{l=0}^{P_{\max}} (c+1-1) \ln \gamma_l - \gamma_l (d + \|h_l^{(t+1)}\|^2 - \mu_{\beta_l}^H \mu_l - \mu_l^H \mu_{\beta_l} + \|\beta_l\|^2) \quad (144)$$

$$(145)$$

where $\|h_l^{(t+1)}\|^2 = \Sigma_{h_l}^{t+1} + |\mu_{h_l}^{t+1}|^2$, $\|\beta_l\|^2 = \sigma_{\beta_l}^2 + |\mu_{\beta_l}^{t+1}|^2$

3. Update $q(\beta)$

$$\ln q^{(t+1)}(\beta) \propto \ln p(h|\beta, \gamma) + \ln p(\beta) \quad (146)$$

$$\propto \sum_{l=0}^{P_{\max}} \left(-\gamma_l \|h_l^{(t+1)} - \beta_l\|^2 - \sigma_{\beta_l}^{-2} \|\beta_l - \mu_{\beta_l}\|^2 \right) \quad (147)$$

$$\propto \sum_{l=0}^{P_{\max}} -(\gamma_l + \sigma_{\beta_l}^{-2}) \beta_l^H \beta_l + \beta_l^H (\gamma_l \mu_{h_l} + \sigma_{\beta_l}^{-2} \mu_{\beta_l}) \quad (148)$$

$$\quad (149)$$

Therefore,

$$\sigma_{\beta_l}^{(t+1)^2} = (\gamma_l + \sigma_{\beta_l}^{-2})^{-1} \quad (150)$$

$$\mu_{\beta_l} = \sigma_{\beta_l}^{(t+1)^2} (\gamma_l \mu_{h_l} + \sigma_{\beta_l}^{-2} \mu_{\beta_l}) \quad (151)$$

4.2 Symbol Detection

In this section, we assume the channel is known. The target is to find the channel and the symbol to maximize the posterior, i.e.,

$$p(x; y, H, \sigma^2) = \arg \max_{x \in \Omega} p(y|x; H, \sigma^2) p(x) \quad (152)$$

Here, we use a

4.3 Joint Method

We shall look at some examples to solve this problem

$$\|y_p - \phi_p h\|^2 = (y_p - \phi_p h)^H (y_p - \phi_p h) \quad (153)$$

$$= y_p^H y_p - y_p^H \phi_p h - (\phi_p h)^H y_p + (\phi_p h)^H (\phi_p h) \quad (154)$$

$$= y_p^H y_p - y_p^H \phi_p h - (\phi_p h)^H y_p + h^H \phi_p^H \phi_p h \quad (155)$$

Here,

$$(y_p^H \phi_p h)^H = (\phi_p h)^H y_p$$

Therefore,

$$y_p^H \phi_p h + (\phi_p h)^H y_p = 2\text{Re}\{y_p^H \phi_p h\}$$

$$\|y_p - \phi_p h\|^2 = y_p^H y_p - 2\text{Re}\{y_p^H \phi_p h\} + h^H \phi_p^H \phi_p h$$

Now, we do the expectation for $\|y_p - \phi_p h\|^2$ on h ,

$$<\|y_p - \phi_p h\|^2>_h = y_p^H y_p - 2\text{Re}\{y_p^H \phi_p u_h\} + < h^H \phi_p^H \phi_p h >_h \quad (156)$$

$$= y_p^H y_p - 2\text{Re}\{y_p^H \phi_p u_h\} + < \text{tr}(h^H \phi_p^H \phi_p h) >_h \quad (157)$$

$$= y_p^H y_p - 2\text{Re}\{y_p^H \phi_p u_h\} + < \text{tr}(\phi_p^H \phi_p h h^H) >_h \quad (158)$$

$$= y_p^H y_p - 2\text{Re}\{y_p^H \phi_p u_h\} + \text{tr}(\phi_p^H \phi_p < h h^H >_h) \quad (159)$$

$$\quad (160)$$

The covariance of h is,

$$\Sigma_h = < h h^H > - u_h u_h^H$$

$$< h h^H > = \Sigma_h + u_h u_h^H$$

Therefore,

$$<|y_p - \phi_p h|^2>_h = y_p^H y_p - 2Re\{y_p^H \phi_p u_h\} + tr(\phi_p^H \phi_p (\Sigma_h + u_h u_h^H)) \quad (161)$$

$$= y_p^H y_p - 2Re\{y_p^H \phi_p u_h\} + tr(\phi_p^H \phi_p \Sigma_h) + tr(\phi_p^H \phi_p u_h u_h^H) \quad (162)$$

$$= y_p^H y_p - 2Re\{y_p^H \phi_p u_h\} + tr(\phi_p^H \phi_p \Sigma_h) + tr(u_h^H \phi_p^H \phi_p u_h) \quad (163)$$

$$= y_p^H y_p - 2Re\{y_p^H \phi_p u_h\} + u_h^H \phi_p^H \phi_p u_h + tr(\phi_p^H \phi_p \Sigma_h) \quad (164)$$

$$= ||y_p - \phi_p u_h||^2 + tr(\phi_p \Sigma_h \phi_p^H) \quad (165)$$

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