

Linear regression *

(2)

Simple regression

Residual sum of squares (RSS) = $e_1^2 + \dots + e_m^2 = (y_1 - \hat{\beta}_0 - \hat{\beta}_1 x_1)^2 + \dots + (y_m - \hat{\beta}_0 - \hat{\beta}_1 x_m)^2$

Least-squares minimize RSS with $\hat{\beta}_1 = \frac{\sum_{i=1}^m (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^m (x_i - \bar{x})^2}$ $\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}$

$\hat{\beta}_1$ and $\hat{\beta}_0$ are unbiased (does not systematically over or under estimate the true parameters).

→ In general sample mean unbiased $\hat{\mu} = \bar{y} = \frac{1}{m} \sum_{i=1}^m y_i$ and $SE(\hat{\mu})^2 = \frac{\sigma^2}{m}$, where σ is the SD of each realization y_i of y . $SE(\hat{\mu})$ tells us how $\hat{\mu}$ differs from actual value of μ .

$$SE(\hat{\beta}_0)^2 = \sigma^2 \left[\frac{1}{m} + \frac{\bar{x}^2}{\sum_{i=1}^m (x_i - \bar{x})^2} \right], \quad SE(\hat{\beta}_1)^2 = \left[\frac{\sigma^2}{\sum_{i=1}^m (x_i - \bar{x})^2} \right] \text{ where } \sigma^2 = \text{Var}(e)$$

Supposing common variance σ^2 and uncorrelated.

Notice in the formula that $SE(\hat{\beta}_1)$ is smaller when the x_i are more spread out; intuitively, we have more leverage to estimate the slope when this is the case. We also see that $SE(\hat{\beta}_0)$ would be the same as $SE(\hat{\mu})$ if \bar{x} were zero in which case $\hat{\beta}_0 = \bar{y}$.

• Unknown \Rightarrow Residual standard error (RSE) = $\sqrt{RSS / (m-2)}$ \Rightarrow average amount the response will deviate from the true regression line. \Rightarrow measure lack of fit of the model.

• A 95% confidence interval is defined as a range of values such that with 95% probability, the range will contain the true unknown value of parameters. $\hat{\beta}_1 \pm 2 \cdot SE(\hat{\beta}_1)$

• H_0 : there is no relationship between x and y : $\beta_1 = 0 \Rightarrow y = \beta_0 + \epsilon$ (null)
 H_a : there is some " " " " " $\beta_1 \neq 0 \Rightarrow y = \beta_0 + \beta_1 x + \epsilon$ (alternative)

We need to test whether $\hat{\beta}_1$ is sufficiently far away from 0. How far is enough? (because if $SE(\hat{\beta}_1)$ is high it has to be farther than if $SE(\hat{\beta}_1)$ is low). So we use

t-statistic $\Rightarrow t = \frac{\hat{\beta}_1 - 0}{SE(\hat{\beta}_1)}$ (the number of SD that $\hat{\beta}_1$ is from 0). with $(m-2)$ degree freedom. (same as gaussian for $n \geq 30$)

We need to compute the probability of observing any number equal to $|t|$ or larger, assuming $\beta_1 = 0 \Rightarrow$ this is the p-value. Small p-value ($< 5\%$ or 1%) means it is unlikely to observe association between predictor and response due to chance \Rightarrow reject null hypothesis

Accuracy of the model (once rejecting the null)

RSE but also R^2 = the proportion of variance explained ($0 \leq R^2 \leq 1$) $R^2 = \frac{TSS - RSS}{TSS} = 1 - \frac{RSS}{TSS}$ where $TSS = \sum (y_i - \bar{y})^2$ is total sum of squares, which is the variability of inherent of the response before the regression is performed).

(TSS - ASS) amount of variability explained by the regression. R^2 measures the proportion of variability in Y that can be explained using X . $R^2 = 1$ good. But hard to determine the right R^2 = domain knowledge).

Multiple linear regression

- $H_0: \beta_1 = \beta_2 = \dots = \beta_p = 0$ H_a : at least one β_j is non zero. This hypothesis is performed using F statistic $F = \frac{(TSS - ASS)/p}{ASS/(n - p - 1)}$. if the linear assumptions are correct $E\{ASS/(n - p - 1)\} = 0^2$ and provided H_0 is true $E\{(TSS - ASS)/p\} = 0^2 \Rightarrow$ if no relationship F is close to 1, ~~else~~ how large does F need to be? if n is big small deviation is strong evidence against H_0 . if n is small we need a strong deviation. But in all case, according to m and p , we look at the p-value of F . (because follows F -distribution).
- test a particular subset of q of the coefficients are zero. we fit a new model without the q and get ASS_0 then $F = \frac{(ASS_0 - ASS)/q}{ASS/(n - p - 1)}$
- In multiple regression, we get t-statistic and p-value for each individual (whether it is related to the response). It is equivalent of the F-statistic that omits that single variable, leaving all the others in.
- It seems that if any one of the individual p-value is small, at least one is related to the predictors, but this is flawed. \Rightarrow i.e. if $p = 100 \Rightarrow \beta_1 = \dots = \beta_{100} = 0$, about 5% will be below 0.05 by chance. But F-statistic doesn't suffer from that as it adjusts from the number of predictors. Hence if H_0 is true, there is only a 5% chance that the F-statistic will result in a p-value below 0.05, regardless the number of predictors. But this approach only when p is small, and small relative to n . otherwise use forward selection as high dimensional method.

- Method in multiple regression \rightarrow got F-statistic
 - P-value big (of F) \rightarrow no relationship
 - P-value of F is small
 - if F is big
 - if p is small we can select individually the most promising (but not the best way).
- Variable selection: we can use Mallows's CP , AIC , BIC , adjusted R^2 (ch 8). But with p predictor we get 2^p different models so there is easier way.
 - Forward selection
begin with a null model of an intercept and no predictor. we then fit p simple linear regression and add to the null model the variable that result in the lowest ASS. we then keep going till a certain threshold.
 - Backward
we start with all variables, and remove the one with largest p-value. And repeat
 - Mixed
Start with nothing, add till p-value for one of the variables \uparrow , then we remove it. we repeat till small p-value for all.
- warning: R^2 always increases (closer to 1) as we add variables, because always reduce the residual sum of squares on the training data. Also $AIC = \sqrt{\frac{1}{n - p - 1}} ASS$ then we can get $\uparrow AIC$ if the decrease in ASS is small relative to the increase in p .