

Analysis methods

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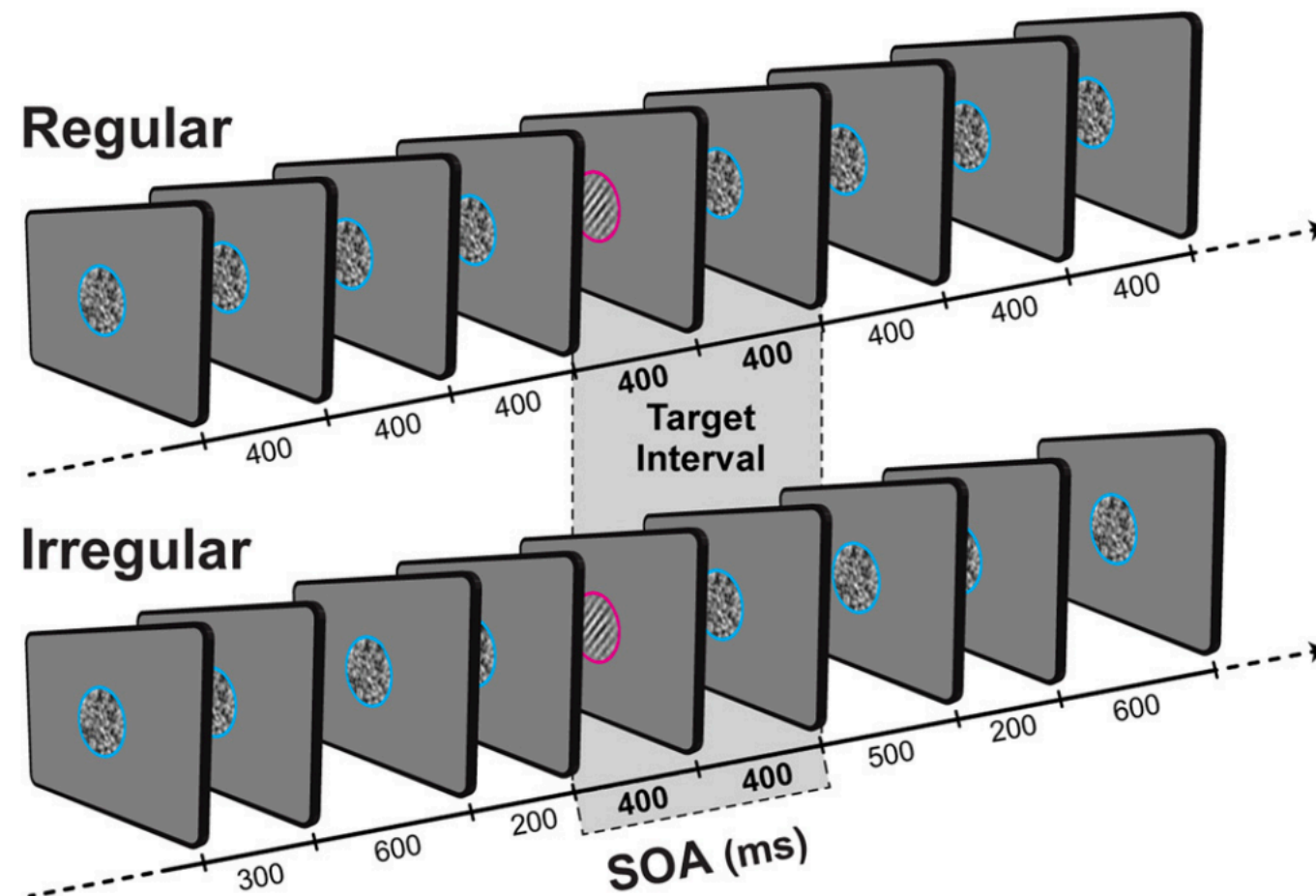


Figure 1. Schematic illustration of the task structure. Each trial consisted of a stream of stimuli presented foveally either with a fixed (regular condition) or jittered (irregular condition) SOA. Notice that the intervals surrounding the targets were exactly the same for the regular and irregular condition. The targets were visual gratings (Gabor) tilted 45° clockwise or counter-clockwise and presented at seven contrast levels to yield a range of performance levels spanning from near chance to near perfect. Participants were asked to respond to the orientation of the target (i.e., left or right) whenever a target was presented. Target presentation was always indicated by a change in the placeholder color to prevent responses to standard stimuli.

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- The psychometric data from each participant and condition were fitted with sigmoidal Weibull functions, each defined by three parameters: threshold α , slope β , and fixed lapse-rate γ (Wichmann and Hill, 2001).
- Guess rates were fixed at 0.5 (i.e., chance level) across all subjects and conditions, and the three parameters were fitted separately for each subject and condition (regular or irregular). Threshold was taken as the predicted contrast level corresponding to 75% accuracy.
- To test for the effect of temporal expectation on target discrimination, fitted threshold and slope values for each participant were submitted to paired t tests. Quality of fit for each subject was assessed by correlating predicted values from the best fitting psychometric function with observed accuracy (mean r-square for group = 0.92, SEM = 0.015; lowest individual r-square = 0.79). The analysis of the psychometric function was performed using the Palamedes toolbox for MATLAB (Prins and Kingdom, 2009).

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- The diffusion model (Palmer et al., 2005) predicts that the psychometric function for accuracy $P_C(x)$ and the chronometric function for the mean response time $t_T(x)$ are functions of stimulus strength x according to the two following equations:

$$P_C(x) = \frac{1}{1 + \exp(-k A x)}$$

$$t_T(x) = \frac{A}{kx} \tanh(k A x) + t_R$$

-
- where k corresponds to the normalized accumulation rate, A to the normalized decision bound, and t_R to the residual time constant. Stimulus strength x follows signal contrast c following a power law whose exponent—which controls the overall slope of both psychometric and chronometric functions as a function of signal contrast—was determined empirically by maximum-likelihood estimation (see below, best-fitting value = 1.97) and fixed across all subjects and conditions.

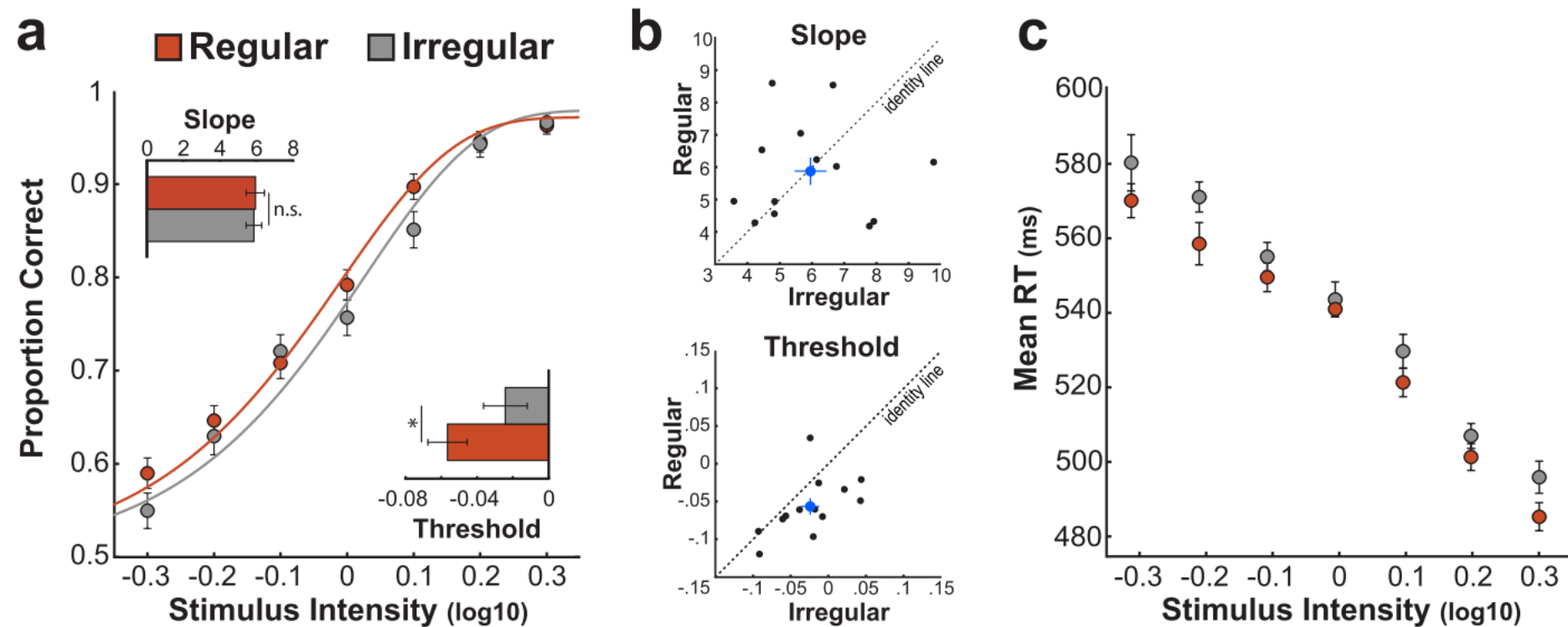
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- We performed constrained maximum-likelihood estimation to recover the best-fitting parameter changes for explaining the effect of temporal expectation (i.e., the difference between regular and irregular conditions) on the group-level psychometric and chronometric functions.
- We used group-level means and standard errors to estimate log-likelihood sums for different sets of parameters, and likelihood-ratio tests to compare between nested models.
- Finally, we computed the model predictions for the effect of temporal expectation on psycho-physical threshold using paired t tests that the observed effect did not differ significantly from model-predicted effects.

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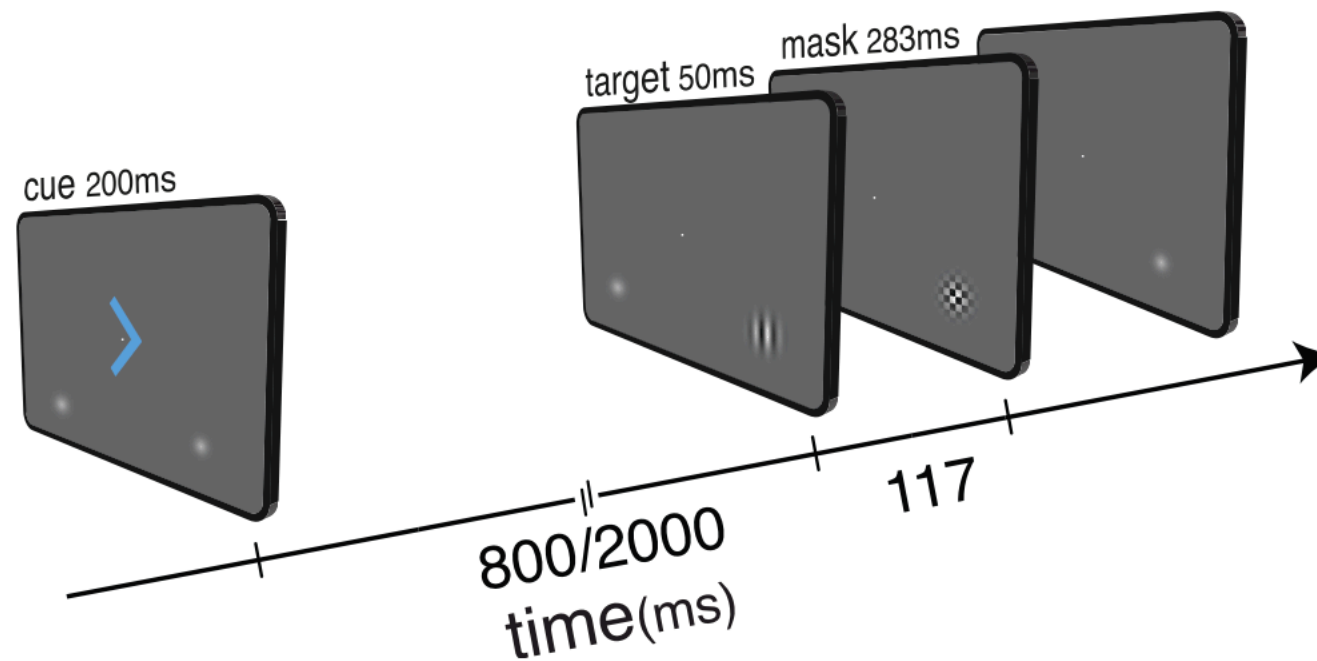


Figure 1. Schematic illustration of the task structure. Foveally presented cues predicted simultaneously where and when target events were more likely to occur. Cue validity for both spatial and temporal expectation was fixed at 80%. Targets consisted of a horizontally or vertically oriented Gabor patch followed by a backward pattern mask. Targets were presented at a fixed contrast, individually adjusted to equate discrimination performance across individuals. Observers responded to the Gabor orientation with their left or right index finger (counterbalanced across subjects).

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- Statistical analysis was performed using MATLAB and SPSS. Where appropriate, the Greenhouse-Geisser correction for nonsphericity was applied. Saccades were detected using a velocity-based algorithm (Engbert & Mergenthaler, 2006). Trials containing saccades (;5% trials) were removed from all analyses.
- Behavioral performance was analyzed using perceptual sensitivity values (d'), proportion of correct responses (P_c) and response times. Sensitivity to stimulus orientation was calculated according the formula:
- $d' = z[P_{cH}] + z[P_{cV}]$
- where P_{cH} and P_{cV} correspond to the observers' proportions of correct responses to horizontal and vertical stimuli respectively, and z to the inverse normal (z score) transformation.

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- Response times were adjusted according to the observers' accuracy in order to account for possible speed-accuracy trade-offs. Similarly to previous studies, this adjustment was done using a measure called inverse efficiency. The inverse efficiency (IE) is calculated by dividing the mean reaction time by the proportion of correct responses.
- Repeated-measures ANOVAs were performed using the Greenhouse-Geisser correction for nonsphericity.
- Additionally, we used nonparametric permutation tests to control for possible biases in the statistics due to difference in power between valid and invalid conditions.

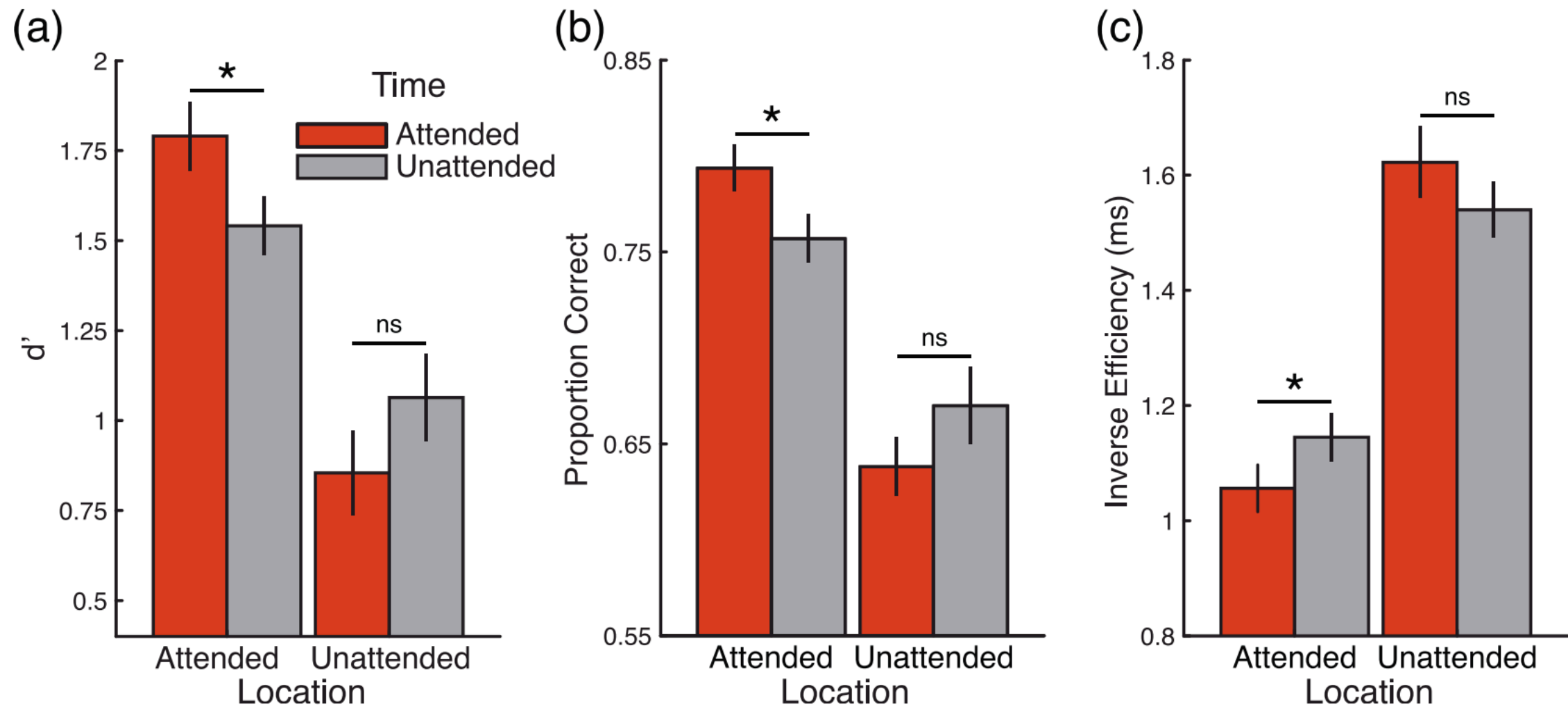
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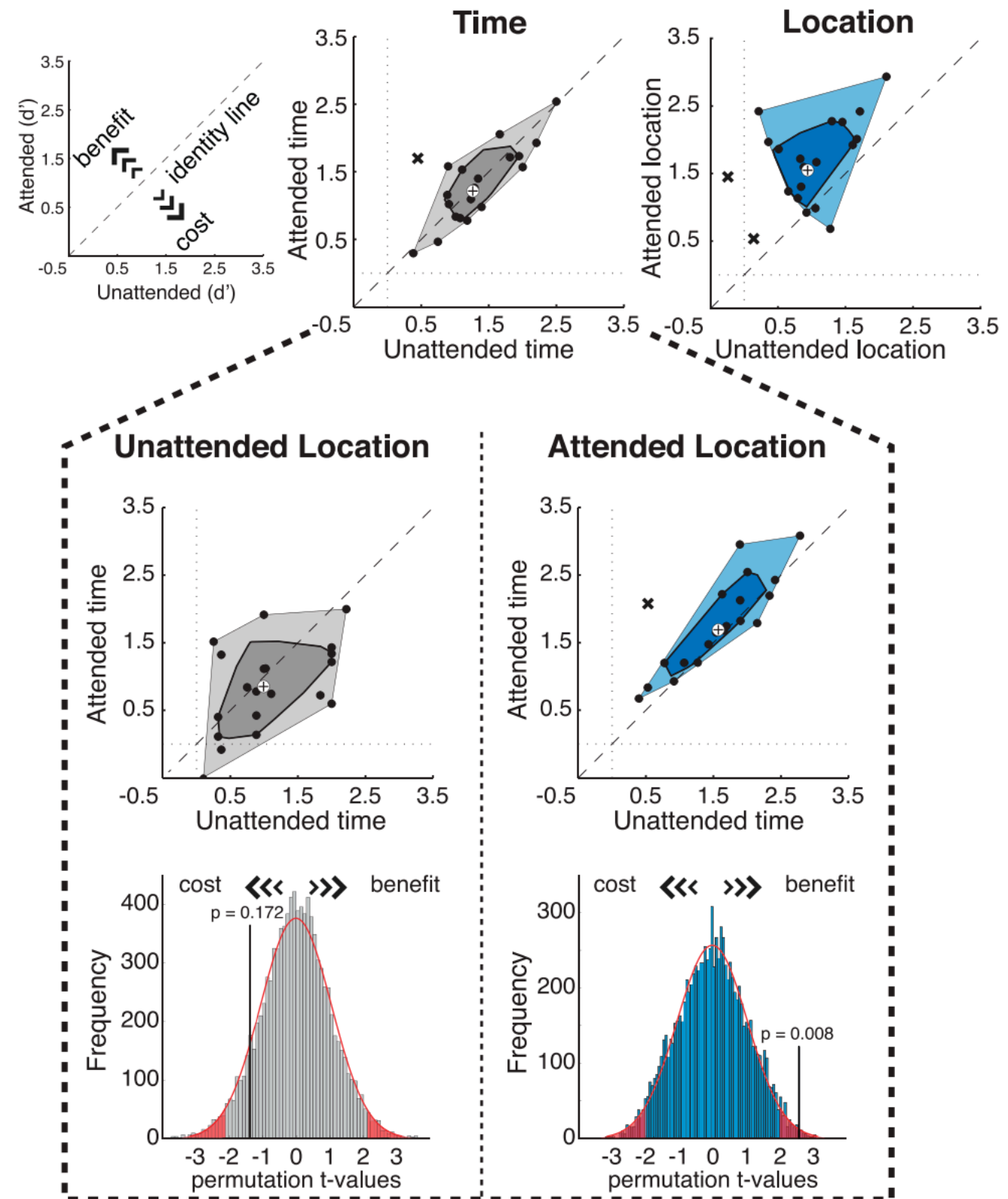
- First, we calculated the mean difference between conditions of interest for each participant. We then performed F tests (ANOVAs) or t tests (for post-hoc comparisons) on these values. Both statistical tests were two-sided and used a critical alpha level of 0.05. We assessed the significance of the observed results by comparison to a null distribution generated via Monte Carlo simulation with 10,000 repetitions.
- This null distribution was generated by randomly shuffling the condition labels within each participant's data in each repetition.
- We then performed the statistical test (F or t tests) on the mean difference between the conditions of interest, and the resulting value was entered into the null distribution.
- The permutation p value was determined as the proportion of random partitions that resulted in a larger test statistic than the observed one.

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Effects of temporal expectations on d' measures at the individual level and permutation tests. Scatter and bagplots showing the main effects of temporal and spatial expectations (Top) and the effect of temporal expectation at the attended and unattended locations (Bottom). Blue indicates the conditions in which there was a significant effect of expectation. The cross at the center of the bagplot represents the center of mass of the bivariate distribution of empirical data, the darker area includes 50% of the data with the largest depth, the lighter polygon contains all other nonoutlier data points, and the Xs represent outliers (Rousseeuw, Ruts, & Tukey, 1999). Outliers were detected using the “Skewness-Adjusted Outlyingness” method using LIBRA toolbox for MATLAB (Verboven & Hubert, 2005; for more details on this method see Hubert & Van der Veen, 2008). Histograms showing the distributions generated from 10,000 random permutations from data sets of 20 participants (see Methods for details). Original t test values (Attended vs. Unattended time at Unattended (left) and Attended (right) Locations) are indicated by the black vertical lines



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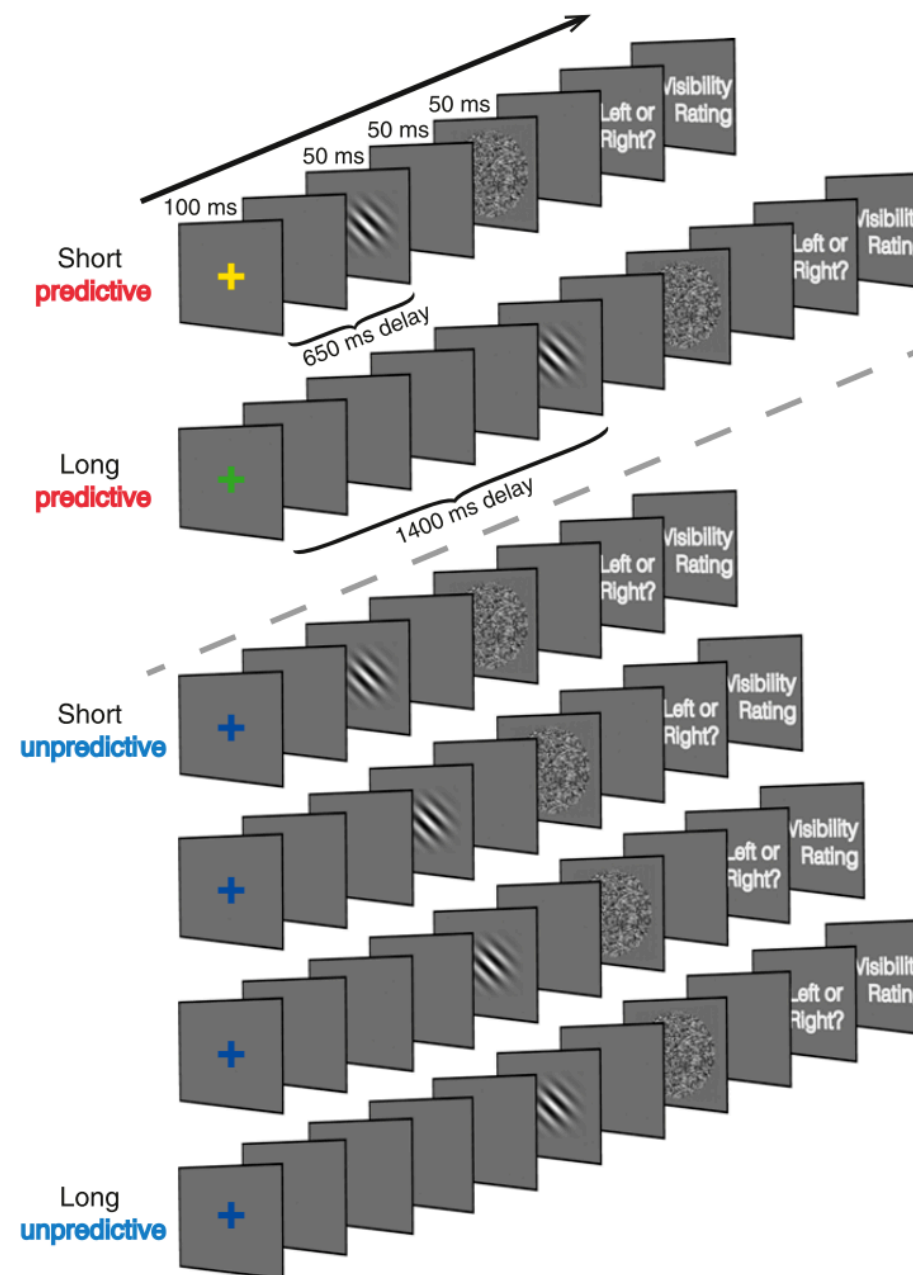


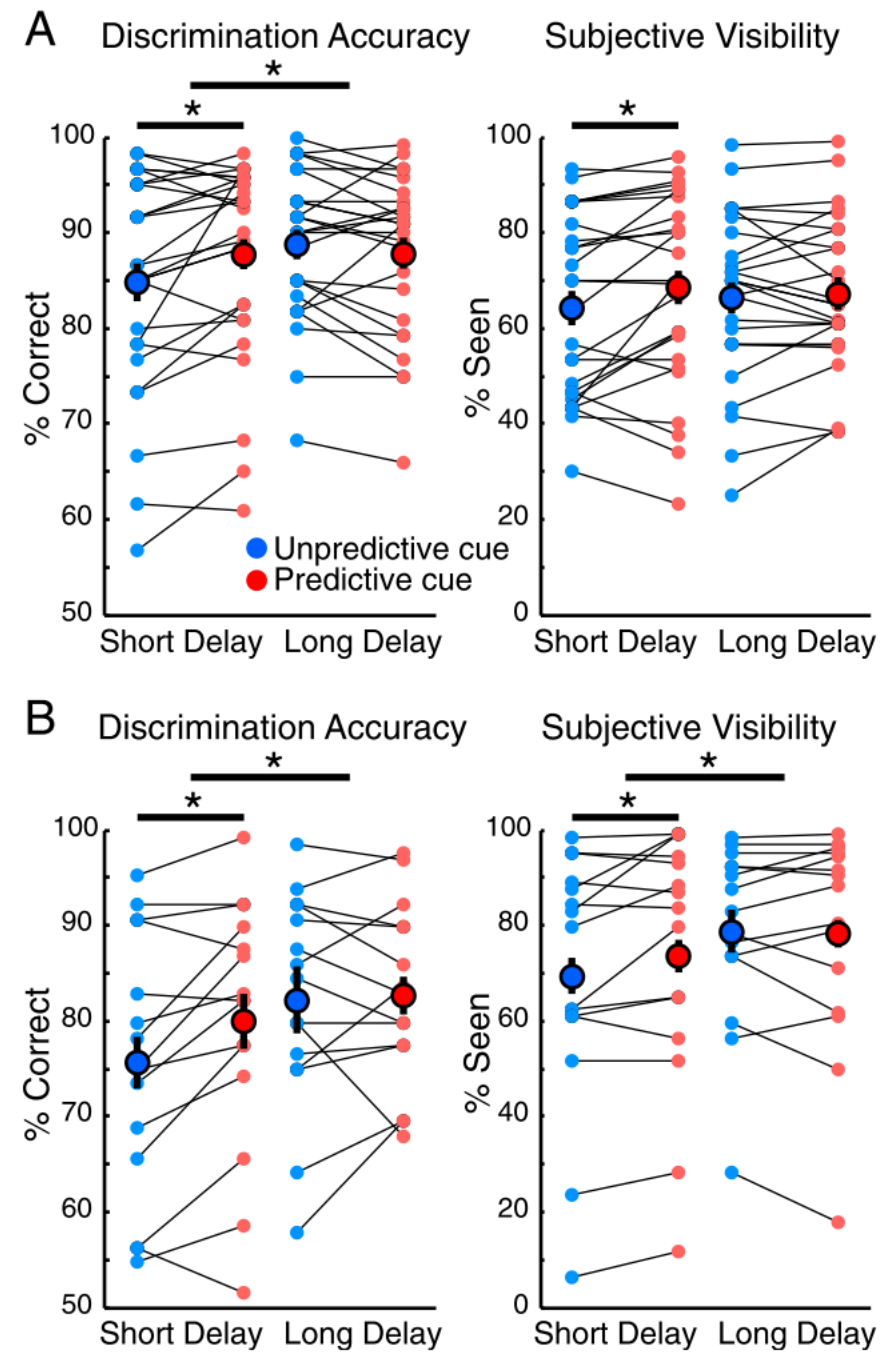
Fig. 1. Schematic of trial types and timing. Colored fixation crosses cued the appearance of target Gabor patches at a short, long, or unpredictable delay (a random selection of one of four delays). Participants provided nonspeeded, two-alternative forced-choice (2AFC) orientation judgments, followed by a seen or guess judgment (experiment 1) or using the perceptual awareness scale (experiment 2).

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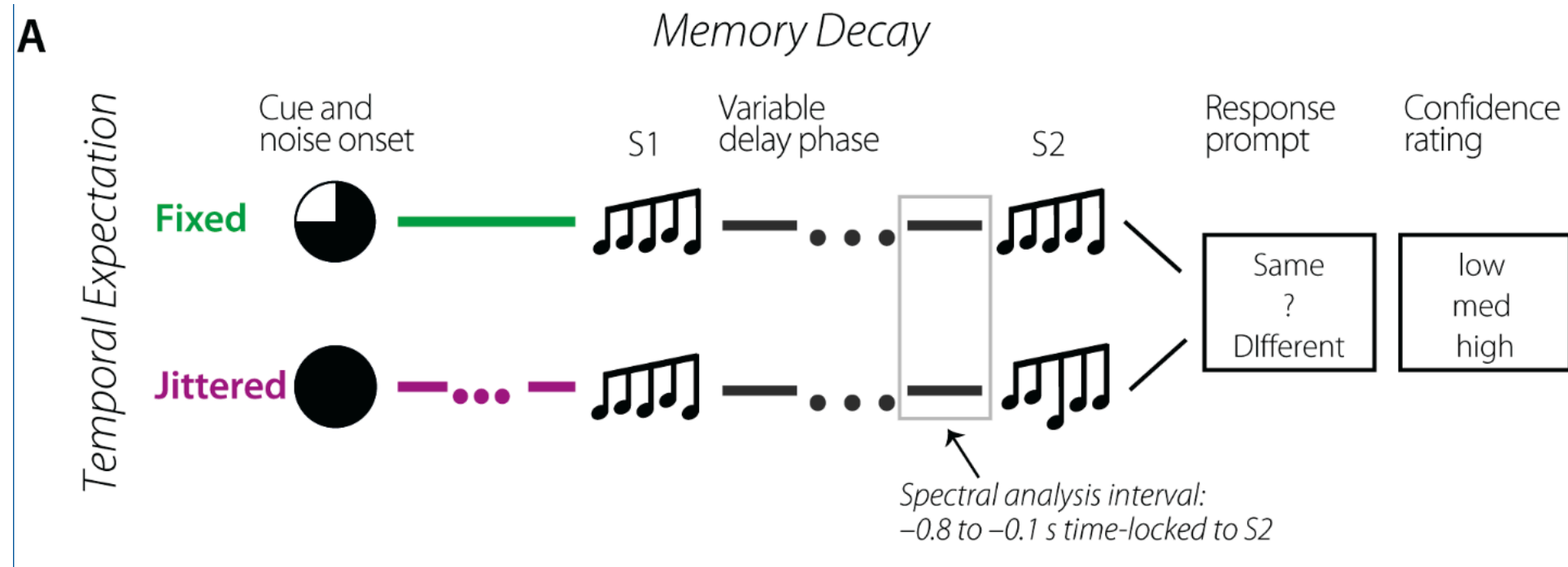
- Accuracy and subjective visibility data from experiment 1 were each submitted to a repeated-measures ANOVA with delay (short, long) and cue type (predictive, unpredictable) as within-subject factors.

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- Experimental design and behavioral performance. A. Experimental design. The upper panel illustrates a “same” trial (S1 and S2 are the same) with a fixed onset time. The lower panel illustrates a “different” trial (S1 and S2 are different) with jittered onset time. The actual durations of the variable delay phases are specified in B and D. The light gray box indicates the interval of the spectral analyses: -0.8 to -0.1 s time-locked to S2.

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- The crucial measure for memory decay was the performance measure A_z , a non-parametric performance measure derived from confidence ratings.
- Hit and false alarm rates at each confidence level were used to construct receiver operating characteristic (ROC) curves (Macmillan and Creelman, 2004) for each condition, and ROC curves were used to derive A_z .
- A_z corresponds to the area under the ROC curve and can be interpreted similarly to proportion correct. A_z was computed for each of the twelve conditions (temporal expectation, 2, \times memory decay, 6), allowing us to estimate memory decay as a function of delay-phase duration separately for fixed and jittered onset times.

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- We fitted Equation 1 (Glass and Mackey, 1988) to Az scores as a function of delay-phase duration:

- $$x(t) = x_0 + e^{-\gamma t} + \frac{\lambda}{\gamma}(1 - e^{-\gamma t}) \quad (1)$$

- where t is equal to time (i.e., delay-phase duration) and x_0 corresponds to the intercept. This specific function contained a term describing decay, γ , and an additional term describing growth, λ . The ratio indicates the function's asymptote.

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- In addition, we also fitted a decay-term-only model (i.e., first term: $x(t) = (x_0 + e^{-\gamma t})$) The decay-only model is more parsimonious and more commonly used to estimate memory decay (Peterson and Peterson, 1959; Wickelgren, 1969).
- To determine which one of these two models represented the memory performance data best, we calculated the Bayesian information criterion (BIC; Schwarz, 1978) for both model fits, as well as for fixed and jittered onset times separately.
- Note that the BIC penalizes for more parameters and allows for an equitable comparison of goodness-of-fit of both models (smaller is better). We averaged the BICs across fixed and jittered onset times separately for each function.
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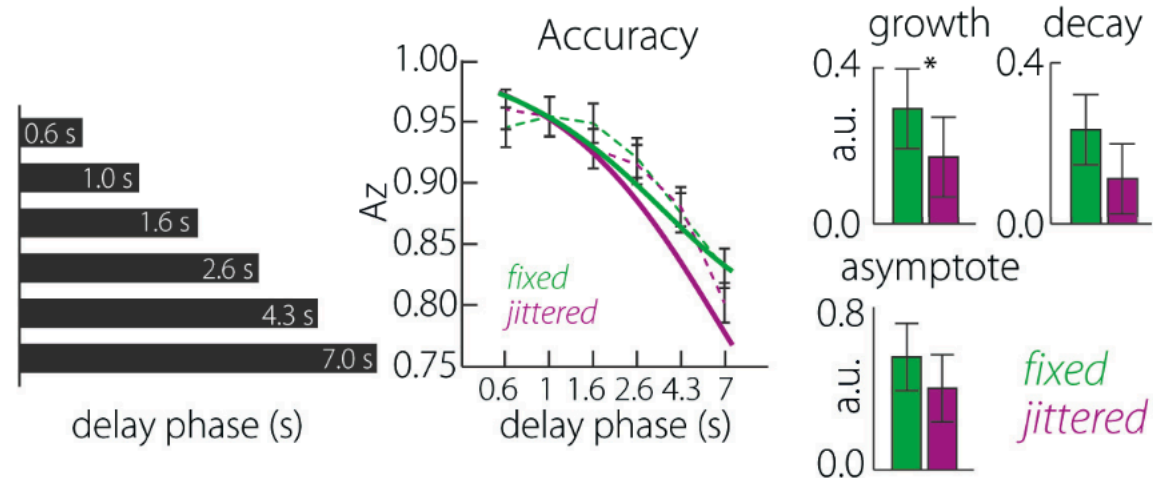
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- 17 out of 18 participants had a lower BIC for the full model (Equation 1) than the decay-only model indicating an overall better fit by the former model. Therefore, all further analyses were conducted on the parameters resulting from the fit of the complete Equation 1. Four of the participants were excluded from subsequent analyses, because R^2 , an indicator for goodness of the model fit, of their fitted models was smaller than 0.3 (see Figure 1C for individual model fits). The average R^2 values for the fixed and jittered conditions, respectively, were 0.66.
- After the fitting of the function, the resulting parameters x_0 , gamma, lambda and for jittered and fixed onset times as dependent variables were assessed with a multivariate ANOVA. This allowed us to test whether there is a global difference between jittered and fixed onset times.
- Subsequently, the parameters x_0 , gamma, lambda were tested for differences between fixed and jittered onset times with univariate repeated-measures ANOVAs, to determine whether memory decay was less strong when S1-onset times were predictable

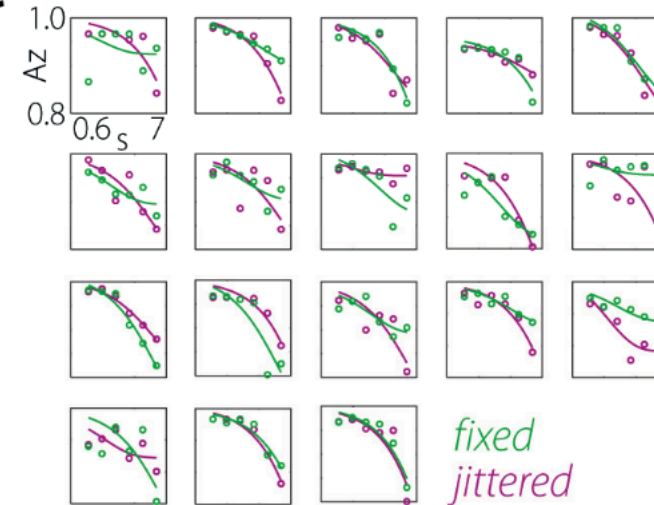
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B



C



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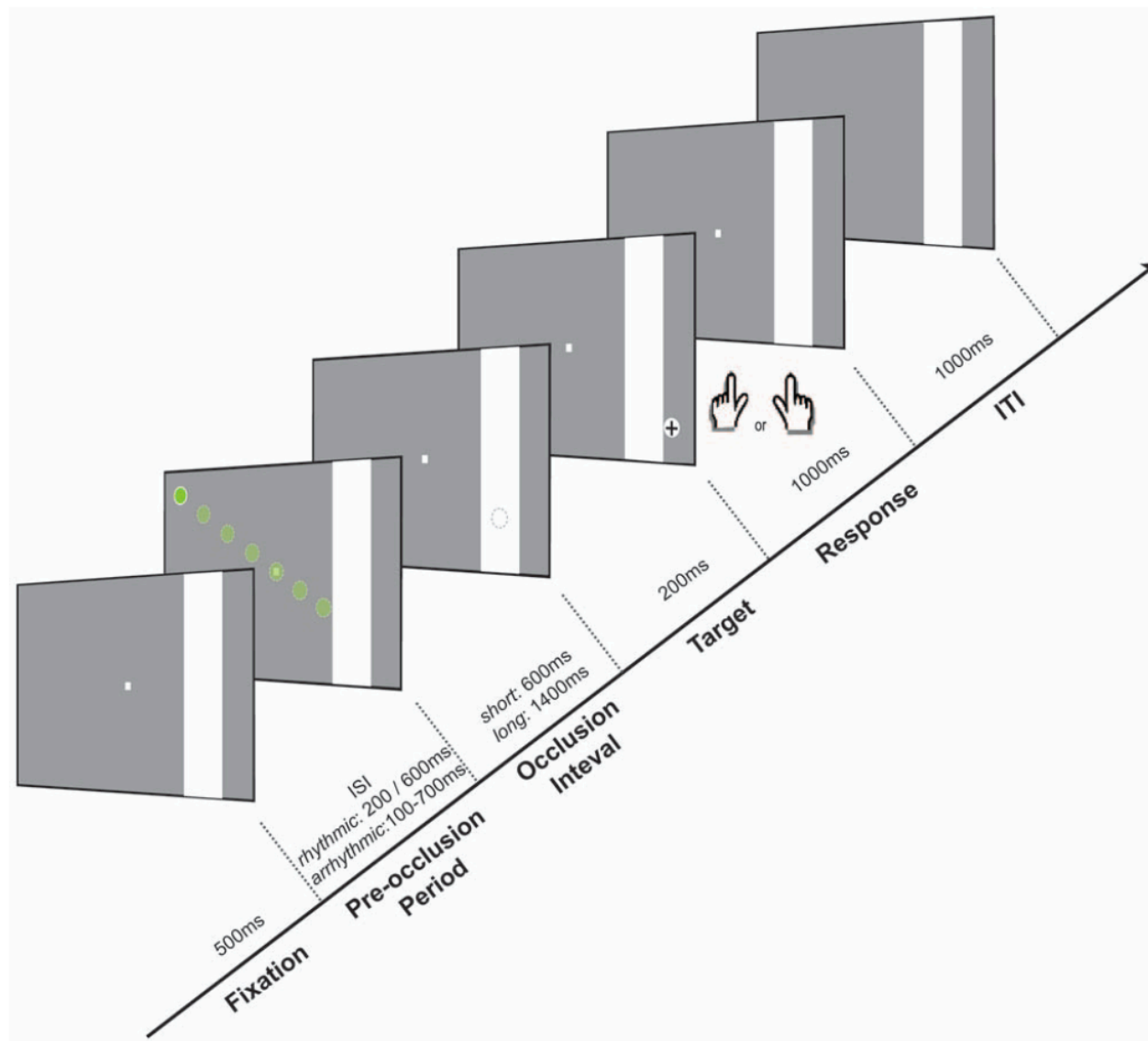


Figure 1. Schematic of the task. A coloured ball appears at the left side of a grey screen and moves across the screen in seven steps. After reaching the occluding band, the ball disappears behind for a single step. When the ball reappears it contains either an upright (50%) or a tilted (50%) cross. Temporal expectations could be induced by the rhythm and/or the colour of the balls preceding the occlusion.
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