

# STA 6106 Homework 5: Randomization test

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## Problem statement

Consider a repeat measure experiment with 3 treatments, which we will label as A, B, and C. Data collected from this experiment is shown below.

Table 1: Data collected from the repeat measure experiment.

Subject	A	B	C
1	10	11	13
2	13	14	17
3	6	8	9

We are interested in whether or not there is a difference in the means of the three populations A, B, and C. That is, we want to test the hypothesis

$$H_o : \mu_A = \mu_B = \mu_C$$

$$H_a : \text{at least one of the means differs}$$

A randomization test will be performed, using systematic data permutations.

## One-way ANOVA

In this problem, we have one treatment with three levels. The omnibus hypothesis test can be conducted using ANOVA. The test statistic is

$$F = \frac{MSB}{MSW} = \frac{\frac{SSB}{df_B}}{\frac{SSW}{df_W}} = \frac{\frac{\sum_{i=1}^k (\bar{y}_{i\bullet} - \bar{y}_{\bullet\bullet})^2}{k-1}}{\frac{\sum_{i=1}^k \sum_{j=1}^{n_i} (\bar{y}_{ij} - \bar{y}_{i\bullet})^2}{N-k}}$$

where  $k = 3$  is the number of treatments,  $n_i = 3$  for  $i = 1, 2, 3$  is the number of observations for the  $i^{th}$  treatment,  $N = \sum_{i=1}^k n_i$  is the total number of observations,  $\bar{y}_{\bullet\bullet} = \frac{1}{N} \sum_{i=1}^k \sum_{j=1}^{n_i} y_{ij}$  represents the grand mean, and  $\bar{y}_{i\bullet} = \frac{1}{n_i} \sum_{j=1}^{n_i} y_{ij}$  is the  $i^{th}$  treatment mean.

If we assume that the observations are independent, and that the underlying populations are normally distributed and homoskedastic across the treatment groups, then under  $H_o$ ,  $F$  has an  $F(df_B, df_W)$ -distribution. However, in a repeated measures experiment we are concerned about a carry-over effect. For example, if a subject has treatment A followed by (after some period of time) treatment B, the measurement taken for treatment B may still be influenced by the earlier treatment A. If this carry-over effect does exist, then the observations will not be independent and our  $F$  statistic will have some unknown distribution.

One way to deal with this carry-over effect is to instead perform a randomization test.

## Randomization test

In a randomization test, we want to consider all of the possible permutations of the data. For a repeated measures experiment, we would go through each row (each subject) and permute the columns (the treatments). For our current problem, in each row there are  $k! = 3! = 6$  possible permutations, and there are  $n = 3$  rows, so there are  $6^3 = 216$  total permutations to consider. For each of these permuted data sets, we calculate the test statistic,  $F_p^*$ . We can then calculate a  $P$ -value by determining the number of statistics that are as extreme as the original one.

$$P - value = \frac{\text{number of } F_p^* \geq F}{\text{number of permutations}}$$

## Optimizing

Consider the test statistic  $F = \frac{SSB}{SSW} \frac{df_B}{df_W}$ . With the permutation test, we are only concerned with the inequality

$$\begin{aligned} F_p^* &\geq F \\ \frac{SSB_p^*}{SSW_p^*} \frac{df_W}{df_B} &\geq \frac{SSB}{SSW} \frac{df_W}{df_B} \end{aligned}$$

where  $F_p^*$  is the statistic calculated on the  $p^{th}$  permuted data set. Now, if there are any constant factors, they can be eliminated. Since the degrees of freedom do not change with permutations, they can be cancelled out.

$$\frac{SSB_p^*}{SSW_p^*} \geq \frac{SSB}{SSW}$$

In addition, recall the identity  $SSTO = SSB + SSW$ , where  $SSTO$  is the total sum of squares. This implies that, if  $SSB$  increases, then  $SSW$  necessarily decreases. Because of this relationship, we know  $SSB$  increases if and only if  $\frac{SSB}{SSW}$  increases. Since we are only interested in whether or not a permuted data set satisfies the above inequality, it is sufficient to only compute  $SSB_p^*$  and compare it to  $SSB$ .

$$SSB_p^* > SSB$$

Furthermore, by rewriting  $SSB$  in its computational formula, we find more constants that can be removed.

$$SSB = \sum_{i=1}^k \left( \frac{1}{n_i} \sum_{j=1}^{n_i} y_{ij} \right)^2 - \sum_{i=1}^k \left( \frac{1}{N} \sum_{l=1}^k \sum_{j=1}^n y_{lj} \right)^2$$

The second term is a constant, so it can be ignored. Also, in this experiment we have  $n_i = n$  for all  $i$ , so the statistic can be reduced to

$$F_r = \sum_{i=1}^k \left( \sum_{j=1}^n y_{ij} \right)^2$$

The notation  $F_r$  will be used to denote this reduced test statistic. For each permuted sample,  $y_{(b)}^*$ , we calculate  $F_{r(b)}^* = \sum_{i=1}^k \left( \sum_{j=1}^n y_{ij}^* \right)^2$ . These will be compared to  $F_r$  computed from the original data set.

```
F.r <- function(data) {
  sum(apply(data, 2, sum)^2)
}
```

## Obtaining systematic permutations

### Algorithm

It is assumed that a method for obtaining all permutations of a single vector is available; most programming languages provide functions for this procedure. In the implementation here, in R, we use the function “permn()” from the library “combinat”. permn() returns a list of all permutations for a given vector.

Now, to permute our original data set, a recursive function is defined. Let  $M$  denote a  $n \times k$  matrix representing the original data set. The  $i^{th}$  row of  $M$  will be denoted by  $M[i, ]$ .

1. Define “permute\_matrix” with two arguments: (matrix “M” and row index “i”), as follows
2.     if  $i$  = number of rows in  $M$
3.         PRINT  $M$
4.     else
5.         SET perm\_list = the list of all permutations of  $M[i, ]$ .
6.         for each permutation in perm\_list
7.             SET  $M[i, ]$  = permutation
8.             CALL permute\_matrix with arguments  $M$  and  $i+1$
- 9.
10. CALL permute\_matrix with arguments  $M$  and  $i = 1$

### Code

```
n <- nrow(data)
k <- ncol(data)
num.permutations <- (factorial(k))^(n - 1)

permute.systematic <- function(data, row.index = 1) {
  ts.list <- NULL
  if(row.index == nrow(data)) {
    return(list(data))
  }
  current.row.perms <- permn(data[row.index, ])
  for(per in current.row.perms) {
    temp <- data
    temp[row.index, ] <- per
    ts.list <- c(ts.list, permute.systematic(temp, row.index + 1))
  }
  return(ts.list)
}

permutation.s <- permute.systematic(data)
```

A	B	C	A	B	C	A	B	C	A	B	C	A	B	C	A	B	C
10	11	13	10	11	13	10	11	13	10	11	13	10	11	13	10	11	13
13	14	17	13	17	14	17	13	14	17	14	13	14	17	13	14	13	17
6	8	9	6	8	9	6	8	9	6	8	9	6	8	9	6	8	9
A	B	C	A	B	C	A	B	C	A	B	C	A	B	C	A	B	C
10	13	11	10	13	11	10	13	11	10	13	11	10	13	11	10	13	11
13	14	17	13	17	14	17	13	14	17	14	13	14	17	13	14	13	17
6	8	9	6	8	9	6	8	9	6	8	9	6	8	9	6	8	9
A	B	C	A	B	C	A	B	C	A	B	C	A	B	C	A	B	C
13	10	11	13	10	11	13	10	11	13	10	11	13	10	11	13	10	11
13	14	17	13	17	14	17	13	14	17	14	13	14	17	13	14	13	17
6	8	9	6	8	9	6	8	9	6	8	9	6	8	9	6	8	9
A	B	C	A	B	C	A	B	C	A	B	C	A	B	C	A	B	C
13	11	10	13	11	10	13	11	10	13	11	10	13	11	10	13	11	10
13	14	17	13	17	14	17	13	14	17	14	13	14	17	13	14	13	17
6	8	9	6	8	9	6	8	9	6	8	9	6	8	9	6	8	9
A	B	C	A	B	C	A	B	C	A	B	C	A	B	C	A	B	C
11	13	10	11	13	10	11	13	10	11	13	10	11	13	10	11	13	10
13	14	17	13	17	14	17	13	14	17	14	13	14	17	13	14	13	17
6	8	9	6	8	9	6	8	9	6	8	9	6	8	9	6	8	9
A	B	C	A	B	C	A	B	C	A	B	C	A	B	C	A	B	C
11	10	13	11	10	13	11	10	13	11	10	13	11	10	13	11	10	13
13	14	17	13	17	14	17	13	14	17	14	13	14	17	13	14	13	17
6	8	9	6	8	9	6	8	9	6	8	9	6	8	9	6	8	9

Figure 1: All 36 permutations, keeping the last row fixed. These will provide all of the possible values for the test statistic

## Obtaining random permutations

As  $n$  or  $k$  become large, the total number of permutations grows exponentially, making the task of systematic permutation computationally intractable. We can instead randomly sample from this set of permutations. We then proceed as in the permutation test on this random sample gives an approximation to the actual  $p$ -value.

### Algorithm

It is assumed that a method for random sampling from a vector without replacement is available. In R, we use the method “sample()”. This procedure will be used to obtain a random permutation of a row in our data matrix.

Let  $M$  denote a  $n \times k$  matrix representing the original data set. The  $i^{th}$  row of  $M$  will be denoted by  $M[i, ]$ . The following algorithm prints out a random permutation of the matrix.

1. for  $i$  from 1 to  $n$
2. SET  $M[i, ]$  = a random permutation of  $M[i, ]$
3. PRINT  $M$

### Code

```
permute.random <- function(data, reps = 100000) {  
  ts.list <- vector("list", reps)  
  n <- nrow(data)  
  m <- ncol(data)  
  for(i in 1:reps) {  
    permutation <- matrix(nrow = n, ncol = m)  
    for(j in 1:(n - 1)) {  
      permutation[j, ] <- sample(data[j, ], m)  
    }  
    permutation[n, ] <- data[n, ]  
    ts.list[[i]] <- permutation  
  }  
  return(ts.list)  
}  
  
permutation.r <- permute.random(data)
```

## Results

First we compute  $F_r$  for the original data.

```
ts <- F.r(data)
```

This gives  $F_r = 3451$ . Now, we calculate the test statistic for each of the permuted data sets, compare them to  $F_r$ , and calculate the percentage that are larger.

```
ts.perm <- sapply(permutation.s, F.r)
p.value.s <- sum(ts.perm >= ts)/length(ts.perm)
```

The  $P$  – value for the permutation test is 0.02778. Let’s now do the same procedure using the set of random permutations.

```
ts.perm <- sapply(permutation.r, F.r)
p.value.r <- sum(ts.perm >= ts)/length(ts.perm)
```

The  $P$  – value for the randomization test is 0.02788. As discussed earlier, the randomization test is an approximation to the permutation test; the relative error of the approximation in this run was 0.368%. But again, since we only have 36 possible permutations for this experiment, the systematic approach is tractable and gives us an exact answer. In either case, we observe a fairly small  $P$  – value and it is reasonable to conclude that at least one of the three treatment means differs.