Introduction to Sampling Theory

Lecture 21 Regression Method of Estimation



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Slides can be downloaded from http://home.iitk.ac.in/~shalab/sp



The ratio method of estimation uses the auxiliary information which is correlated with the study variable to improve the precision which results in the improved estimators when the regression of *Y* on *X* is linear and passes through origin.

When the regression of Y on X is linear, it is not necessary that the line should always pass through origin.

Under such conditions, it is more appropriate to use the regression type estimators to estimate the population mean.

In ratio method, the conventional estimator sample mean \overline{y} was improved by multiplying it by a factor $\frac{\overline{X}}{\overline{x}}$ where \overline{x} is an unbiased estimator of population mean \overline{X} which is chosen as population mean of auxiliary variable.

Now we consider another idea based on difference.

Consider an estimator $(\overline{x} - \overline{X})$ for which

$$E(\overline{x} - \overline{X}) = 0.$$

Consider an improved estimator of \overline{Y} as

$$\hat{\overline{Y}}^* = \overline{y} + \mu(\overline{x} - \overline{X})$$

which is an unbiased estimator of \overline{Y} and μ is any constant.

Now find μ such that the $Var(\hat{Y}^*)$ is minimum.

$$Var(\hat{\overline{Y}}^*) = Var(\overline{y}) + \mu^2 Var(\overline{x}) + 2\mu Cov(\overline{x}, \overline{y})$$

$$\frac{\partial Var(\overline{Y}^*)}{\partial \mu} = 0$$

$$\Rightarrow \mu = -\frac{Cov(\overline{x}, \overline{y})}{Var(\overline{x})}$$

$$= -\frac{\frac{N-n}{Nn}S_{XY}}{\frac{N-n}{Nn}S_X^2}$$

$$=-\frac{S_{XY}}{S_X^2}$$

where
$$S_{XY} = \frac{1}{N-1} \sum_{i=1}^{N} (X_i - \overline{X})(Y_i - \overline{Y}), \quad S_X^2 = \frac{1}{N-1} \sum_{i=1}^{N} (X_i - \overline{X}).$$

Consider a linear regression model $y = \beta_0 + \beta_1 x + \varepsilon$ where y is the dependent variable, x is the independent variable and ε is the random error component which takes care of the difference arising due to lack of exact relationship between x and y.

So we can write the model for each observation as

$$y_{i} = \beta_{0} + \beta_{1}x_{i} + \varepsilon_{i}, i = 1, 2, ..., n.$$

Minimize $\sum_{i=1}^{n} \varepsilon_i^2$ for given n paired data set (x_i, y_i) , i = 1, 2, ..., n.

Minimizes the sum of squares

$$S(\beta_0, \beta_1) = \sum_{i=1}^{n} \varepsilon_i^2 = \sum_{i=1}^{n} (y_i - \beta_0 - \beta_1 x_i)^2$$

with respect to β_0 and β_1 .

$$\frac{\partial S(\beta_0, \beta_1)}{\partial \beta_0} = -2\sum_{i=1}^n (y_i - \beta_0 - \beta_1 x_i)$$

$$\frac{\partial S(\beta_0, \beta_1)}{\partial \beta_1} = -2\sum_{i=1}^n (y_i - \beta_0 - \beta_1 x_i) x_i$$

The solutions of eta_0 and eta_1 are obtained by setting

$$\frac{\partial S(\beta_0, \beta_1)}{\partial \beta_0} = 0$$
$$\frac{\partial S(\beta_0, \beta_1)}{\partial \beta_1} = 0.$$

We obtain the solutions b_0 of β_0 and b_1 of β_1

$$b_0 = \overline{y} - b_1 \overline{x}$$

$$b_1 = \frac{S_{xy}}{S_{xx}}$$

where

$$S_{xy} = \sum_{i=1}^{n} (x_i - \overline{x})(y_i - \overline{y}), \quad S_{xx} = \sum_{i=1}^{n} (x_i - \overline{x})^2, \quad \overline{x} = \frac{1}{n} \sum_{i=1}^{n} x_i, \quad \overline{y} = \frac{1}{n} \sum_{i=1}^{n} y_i.$$

$$b_1 = \frac{S_{xy}}{S_{xx}}$$
 corresponds to $\beta = \frac{Cov(x, y)}{Var(x)} = \frac{S_{xy}}{S_x^2}$.

Thus the optimum value of μ is same as the regression coefficient of y on x with a negative sign, i.e.,

$$\mu = -\beta$$
.

So the estimator $\hat{\vec{Y}}^*$ with optimum value of μ is

$$\hat{\overline{Y}}_{reg} = \overline{y} + b_1(\overline{X} - \overline{x})$$

which is the estimator of \overline{Y} and the procedure of estimation is called as the regression method of estimation.

Regression estimates with pre-assigned β :

If value of β is known as β_0 (say) then the regression estimator is

$$\hat{\overline{Y}}_{reg} = \overline{y} + \beta_0 (\overline{X} - \overline{x})$$

Bias of $\hat{\overline{Y}}_{reg}$

Now, assuming that the random sample (x_i, y_i) , i = 1, 2, ..., n is drawn by SRSWOR,

$$E(\hat{\overline{Y}}_{reg}) = E(\overline{y}) + \beta_0 \left[\overline{X} - E(\overline{x}) \right]$$
$$= \overline{Y} + \beta_0 \left[\overline{X} - \overline{X} \right]$$
$$= \overline{Y}.$$

Thus $\hat{\overline{Y}}_{reg}$ is an unbiased estimator of \overline{Y} when β is known.

Regression estimates with pre-assigned $oldsymbol{eta}$: Variance of $\hat{\overline{Y}}_{reg}$

$$Var(\hat{\overline{Y}}_{reg}) = E \left[\hat{\overline{Y}}_{reg} - E(\hat{\overline{Y}}_{reg}) \right]^{2}$$

$$= E \left[\overline{y} + \beta_{0}(\overline{X} - \overline{x}) - \overline{Y} \right]^{2}$$

$$= E \left[(\overline{y} - \overline{Y}) - \beta_{0}(\overline{x} - \overline{X}) \right]^{2}$$

$$= E \left[(\overline{y} - \overline{Y})^{2} + \beta_{0}^{2}(\overline{x} - \overline{X}) - 2\beta_{0}E(\overline{x} - \overline{X})(\overline{y} - \overline{Y}) \right]$$

$$= Var(\overline{y}) + \beta_{0}^{2}Var(\overline{x}) - 2\beta_{0}Cov(\overline{x}, \overline{y})$$

$$= \frac{f}{n} \left[S_{Y}^{2} + \beta_{0}^{2}S_{X}^{2} - 2\beta_{0}S_{XY} \right]$$

$$= \frac{f}{n} \left[S_{Y}^{2} + \beta_{0}^{2}S_{X}^{2} - 2\beta_{0}\rho S_{X}S_{Y} \right]$$

where

$$f = \frac{N-n}{N}$$
; $S_X^2 = \frac{1}{N-1} \sum_{i=1}^{N} (X_i - \overline{X})^2$; $S_Y^2 = \frac{1}{N-1} \sum_{i=1}^{N} (Y_i - \overline{Y})^2$

 ρ : Correlation coefficient between X and Y.

Regression estimates with pre-assigned β :

Comparing $Var(\hat{Y}_{reg})$ with $Var(\bar{y})$, we note that

$$Var(\hat{\overline{Y}}_{reg}) < Var(\overline{y})$$

If
$$\beta_0^2 S_x^2 - 2\beta_0 S_{xy} < 0$$

$$\beta_0 S_X^2 \left(\beta_0 - \frac{2S_{XY}}{S_X^2} \right) < 0$$

which is possible when

either

$$eta_0 < 0 ext{ and } \left(eta_0 - rac{2S_{XY}}{S_X^2}
ight) > 0 \Longrightarrow rac{2S_{XY}}{S_X^2} < eta_0 < 0$$

or
$$\beta_0 > 0 \text{ and } \left(\beta_0 - \frac{2S_{XY}}{S_X^2}\right) < 0 \Rightarrow 0 < \beta_0 < \frac{2S_{XY}}{S_X^2}.$$

Optimal value of β :

Choose β such that $Var(\hat{Y}_{reg})$ is minimum.

So
$$\frac{\partial Var(\hat{Y}_{reg})}{\partial \beta} = \frac{\partial}{\partial \beta} \left[S_Y^2 + \beta^2 S_X^2 - 2\beta \rho S_X S_Y \right] = 0$$
$$\Rightarrow \beta = \rho \frac{S_Y}{S_X} = \frac{S_{XY}}{S_X^2}.$$

The minimum value of variance of $\hat{\overline{Y}}_{reg}$ with optimum value of

$$\beta_{opt} = \frac{\rho S_{Y}}{S_{X}}$$
 is

$$Var_{min}(\hat{Y}_{reg}) = \frac{f}{n} \left[S_Y^2 + \rho^2 \frac{S_Y^2}{S_X^2} S_X^2 - 2\rho \frac{S_Y}{S_X} \rho S_X S_Y \right]$$
$$= \frac{f}{n} S_Y^2 (1 - \rho^2).$$

Optimal value of β :

We see from

$$Var_{min}(\hat{\overline{Y}}_{reg}) = \frac{f}{n}S_Y^2(1-\rho^2).$$

Since $-1 \le \rho \le 1$, so

$$Var(\hat{\overline{Y}}_{reg}) \leq Var_{SRS}(\overline{y})$$

which always holds true. So the regression estimator is always better than the sample mean under SRSWOR.

Departure from β :

If β_0 is the preassigned value of regression coefficient, then

$$Var_{\min}(\hat{\overline{Y}}_{reg}) = \frac{f}{n} \Big[S_Y^2 + \beta_0^2 S_X^2 - 2\beta_0 \rho S_X S_Y \Big]$$

$$= \frac{f}{n} \Big[S_Y^2 + \beta_0^2 S_X^2 - 2\rho \beta_0 S_X S_Y - \rho^2 S_Y^2 + \rho^2 S_Y^2 \Big]$$

$$= \frac{f}{n} \Big[(1 - \rho^2) S_Y^2 + \beta_0^2 S_X^2 - 2\beta_0 S_X^2 \beta_{opt} + \beta_{opt}^2 S_X^2 \Big]$$

$$= \frac{f}{n} \Big[(1 - \rho^2) S_Y^2 + (\beta_0 - \beta_{opt})^2 S_X^2 \Big]$$

where
$$\beta_{opt} = \frac{\rho S_Y}{S_X}$$
.

Estimate of Variance:

An unbiased sample estimate of $Var(\hat{\overline{Y}}_{reg})$ is

$$\widehat{Var}(\widehat{Y}_{reg}) = \frac{f}{n(n-1)} \sum_{i=1}^{n} \left[(y_i - \overline{y}) - \beta_0 (x_i - \overline{x}) \right]^2$$
$$= \frac{f}{n} (s_y^2 + \beta_0^2 s_x^2 - 2\beta_0 s_{xy})$$

Note that the variance of \hat{Y}_{reg} increases as the difference between eta_0 and eta_{opt} increases.