

Question 1. *To evaluate a new test for detecting Hansen's disease, a group of people 5% of which are known to have Hansen's disease are tested. The test finds Hansen's disease among 98% of those with the disease and 3% of those who don't. What is the probability that someone testing positive for Hansen's disease under this new test actually has it?*

My Answer 1.

A: event that someone is known to have Hansen's disease.

B: event that someone is known to have no Hansen's disease.

X: event that someone has positive result with Hansen's disease.

$$P(A) = 0.05$$

$$P(B) = 0.95$$

$$P(H/A) = 0.98$$

$$P(H/B) = 0.03$$

So, the probability that someone testing positive for Hansen's disease under this new test actually is:

$$P(A/H) = \frac{P(A) * P(H/A)}{P(H)}$$

$$P(A/H) = \frac{P(A) * P(H/A)}{P(H/A) * P(A) + P(H/B) * P(B)}$$

$$P(A/H) = \frac{0.098 * 0.05}{0.098 * 0.05 + 0.03 * 0.95}$$

$$P(A/H) \approx 0.632$$

Question 2. *Proof the following distributions are normalized then calculate the mean and standard deviation of these distributions:*

1. *Univariate normal distribution.*
2. *(Optional) Multivariate normal distribution.*

My Answer 2.

To prove that the univariate Gaussian distribution is normalized, we should show that it is normalized for a zero-mean Gaussian and extend that result to show that $\mathcal{N}(x|\mu, \sigma^2)$ is normalized.

The formula Gaussian distribution is given by:

$$\varphi(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{1}{2\sigma^2}x^2\right) \quad -\infty < x < \infty. \quad (1)$$

To prove that the above expression is normalized, we have to show that

$$\int_{-\infty}^{\infty} \exp\left(-\frac{1}{2\sigma^2}x^2\right) dx = \sqrt{2\pi\sigma^2} \quad (2)$$

Proof. Let

$$I = \int_{-\infty}^{\infty} \exp\left(-\frac{1}{2\sigma^2}x^2\right) dx \quad (3)$$

Squaring the above expression,

$$I^2 = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \exp\left(-\frac{1}{2\sigma^2}x^2 - \frac{1}{2\sigma^2}y^2\right) dx dy$$

To integrate this expression we make the transformation from Cartesian coordinates (x, y) to polar coordinates (r, θ) , which is defined by:

$$x = r \cos \theta \quad (4)$$

$$y = r \sin \theta \quad (5)$$

and using the trigonometric identity $\cos^2 \theta + \sin^2 \theta = 1$, we have $x^2 + y^2 = r^2$. Also the Jacobian of the change of variables is given by,

$$\begin{aligned} \frac{\partial (x, y)}{\partial (r, \theta)} &= \begin{vmatrix} \frac{\partial (x)}{\partial (r)} & \frac{\partial (x)}{\partial (\theta)} \\ \frac{\partial (y)}{\partial (r)} & \frac{\partial (y)}{\partial (\theta)} \end{vmatrix} \\ &= \begin{vmatrix} \cos \theta & -r \sin \theta \\ \sin \theta & r \cos \theta \end{vmatrix} \\ &= r \cos^2 \theta + r \sin^2 \theta \\ &= r \end{aligned}$$

using the same trigonometric identity $\cos^2 \theta + \sin^2 \theta = 1$. Thus equation (4) can be rewritten as

$$I^2 = \int_0^{2\pi} \int_0^{\infty} \exp\left(-\frac{r^2}{2\sigma^2}\right) r dr d\theta \quad (6)$$

$$= 2\pi \int_0^{\infty} \exp\left(-\frac{r^2}{2\sigma^2}\right) r dr \quad (7)$$

$$= 2\pi \int_0^{\infty} \exp\left(-\frac{u}{2\sigma^2}\right) \frac{1}{2} du \quad (8)$$

$$= \pi \left[\exp\left(-\frac{u}{2\sigma^2}\right) (-2\sigma^2) \right]_0^{\infty} \quad (9)$$

$$= 2\pi \sigma^2 \quad (10)$$

where we have used the change of variables $r^2 = u$. Thus

$$I = (2\pi \sigma^2)^{1/2}.$$

Finally to prove that $\mathcal{N}(x|\mu, \sigma^2)$ is normalized, we make the transformation $y = x - \mu$ so that,

$$\begin{aligned} \int_{-\infty}^{\infty} \mathcal{N}(x|\mu, \sigma^2) dx &= \frac{1}{(2\pi \sigma^2)^{1/2}} \int_{-\infty}^{\infty} \exp\left(-\frac{y^2}{2\sigma^2}\right) dy \\ &= \frac{I}{(2\pi \sigma^2)^{1/2}} \\ &= 1 \end{aligned}$$

as required.

My Answer 3.

CALCULATING VARIANCE

The normalization condition of the univariate Gaussian distribution is given by:

$$\begin{aligned}\int_{-\infty}^{\infty} \mathcal{N}(x | \mu, \sigma^2) dx &= 1 \\ \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^{\infty} \exp\left\{-\frac{1}{2\sigma^2}(x - \mu)^2\right\} dx &= 1 \\ \int_{-\infty}^{\infty} \exp\left\{-\frac{1}{2\sigma^2}(x - \mu)^2\right\} dx &= (2\pi\sigma^2)^{1/2}\end{aligned}$$

Differentiating both sides with respect to σ^2 ,

$$\begin{aligned}\int_{-\infty}^{\infty} \exp\left\{-\frac{1}{2\sigma^2}(x - \mu)^2\right\} \left\{\frac{(x - \mu)^2}{2}\right\} \left\{\frac{1}{(\sigma^2)^2}\right\} dx &= \left(\frac{1}{2}\right) (2\pi\sigma^2)^{-1/2} (2\pi) \\ \int_{-\infty}^{\infty} \exp\left\{-\frac{1}{2\sigma^2}(x - \mu)^2\right\} (x - \mu)^2 dx &= \sigma^2 \sqrt{2\pi\sigma^2} \\ \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^{\infty} \exp\left\{-\frac{1}{2\sigma^2}(x - \mu)^2\right\} (x - \mu)^2 dx &= \sigma^2\end{aligned}$$

It follows directly that

$$\mathbb{E}[(x - \mu)^2] = \text{Varr}[x] = \sigma^2$$

My Answer 4.

CALCULATING MEAN

The mean of the univariate Gaussian distribution is given by:

$$\mathbb{E}[x] = \int_{-\infty}^{\infty} \left(\frac{1}{2\pi\sigma^2}\right)^{1/2} \exp\left\{-\frac{1}{2\sigma^2}(x - \mu)^2\right\} x dx$$

Changing the variables using $y = x - \mu$, the above expression becomes

$$\begin{aligned}\mathbb{E}[x] &= \int_{-\infty}^{\infty} \left(\frac{1}{2\pi\sigma^2}\right)^{1/2} \exp\left(-\frac{1}{2\sigma^2}y^2\right) (y + \mu) dy \\ &= \int_{-\infty}^{\infty} \left(\frac{1}{2\pi\sigma^2}\right)^{1/2} \exp\left(-\frac{1}{2\sigma^2}y^2\right) y dy \\ &\quad + \int_{-\infty}^{\infty} \left(\frac{1}{2\pi\sigma^2}\right)^{1/2} \exp\left(-\frac{1}{2\sigma^2}y^2\right) \mu dy\end{aligned}$$

$$\begin{aligned}
 & \int_{-\infty}^{\infty} \left(\frac{1}{2\pi\sigma^2} \right)^{1/2} \exp\left(-\frac{1}{2\sigma^2}y^2\right) y \, dy \\
 &= \int_{-\infty}^0 \left(\frac{1}{2\pi\sigma^2} \right)^{1/2} \exp\left(-\frac{1}{2\sigma^2}y^2\right) y \, dy \\
 &\quad + \int_0^{\infty} \left(\frac{1}{2\pi\sigma^2} \right)^{1/2} \exp\left(-\frac{1}{2\sigma^2}y^2\right) y \, dy \\
 &= \int_{-\infty}^0 \left(\frac{1}{2\pi\sigma^2} \right)^{1/2} \left(\frac{1}{2} \right) \exp\left(-\frac{1}{2\sigma^2}u\right) du \\
 &\quad + \int_0^{\infty} \left(\frac{1}{2\pi\sigma^2} \right)^{1/2} \left(\frac{1}{2} \right) \exp\left(-\frac{1}{2\sigma^2}u\right) du \\
 &= \left(\frac{1}{2\pi\sigma^2} \right)^{1/2} \left(\frac{1}{2} \right) (-2\sigma^2) \\
 &\quad \left\{ \left[\exp\left(-\frac{1}{2\sigma^2}u\right) \right]_{-\infty}^0 + \left[\exp\left(-\frac{1}{2\sigma^2}u\right) \right]_0^{\infty} \right\} \\
 &= \left(\frac{1}{2\pi\sigma^2} \right)^{1/2} \left(\frac{1}{2} \right) (-2\sigma^2) \left\{ \left[\frac{1}{e^0} - \frac{1}{e^{\infty}} \right] + \left[\frac{1}{e^{\infty}} - \frac{1}{e^0} \right] \right\} \\
 &= 0
 \end{aligned}$$

where we have used the change of variables $y^2 = u$ and hence $2y \, dy = du$.
 So, we have:

$$\begin{aligned}
 \mathbb{E}[x] &= \int_{-\infty}^{\infty} \left(\frac{1}{2\pi\sigma^2} \right)^{1/2} \exp\left(-\frac{1}{2\sigma^2}y^2\right) (y + \mu) \, dy \\
 &= \int_{-\infty}^{\infty} \left(\frac{1}{2\pi\sigma^2} \right)^{1/2} \exp\left(-\frac{1}{2\sigma^2}y^2\right) y \, dy \\
 &\quad + \int_{-\infty}^{\infty} \left(\frac{1}{2\pi\sigma^2} \right)^{1/2} \exp\left(-\frac{1}{2\sigma^2}y^2\right) \mu \, dy \\
 &= 0 + \mu * \int_{-\infty}^{\infty} \left(\frac{1}{2\pi\sigma^2} \right)^{1/2} \exp\left(-\frac{1}{2\sigma^2}y^2\right) dy \\
 &= \mu
 \end{aligned}$$