FIT2086 Lecture 8 Model Selection and Penalized Regression

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Outline

- Model Selection Criteria
 - Overfitting and Underfitting
 - Model Selection Methods

- Penalized Regression Techniques
 - Statistical Stability
 - Penalized Regression



Revision from last week (1)

A logistic regression models the conditional log-odds as

$$\log \left(\frac{\mathbb{P}(Y_i = 1 \mid x_{i,1}, \dots, x_{i,p})}{\mathbb{P}(Y_i = 0 \mid x_{i,1}, \dots, x_{i,p})} \right) = \beta_0 + \sum_{j=1}^p \beta_j x_{i,j} \equiv \eta_i$$

Logistic regression model of conditional probability

$$\mathbb{P}(Y_i = 1 \mid x_{i,1}, \dots, x_{i,p}) = \frac{1}{1 + \exp(-\eta_i)}$$

- Estimated using maximum likelihood
- Performance measures for classification
 - Classification error
 - Sensitivity and specificity
 - Area-under-the-curve (AUC)
 - Logarithmic loss



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Supervised Learning – recap

- Imagine we have measured p+1 variables on n individuals (people, objects, things)
- One variable is our target
- ullet We would like to predict our target using the remaining p variables (predictors)
- If the variable we are predicting is categorical, we are performing classification
- If the variable we are predicting is numerical, we are performing regression

Underfitting/Overfitting Example (1)

- We often have many measured predictors
 - Should we use them all, and if not, why not?
- Omitting important predictors
 - Called underfitting
 - Leads to systematic error ("bias") in predicting the target
- Including spurious predictors
 - Called overfitting
 - Leads our model to "learn" noise and random variation
 - Poorer ability to predict to new, unseen data from our population

Underfitting/Overfitting Example (2)

- ullet Example: we observe x and y data and want to build a prediction model for y using x
 - Data looks nonlinear so we use polynomial regression
 - We take $x, x^2, x^3, \dots, x^{20} \Rightarrow$ very flexible model
 - How many terms to include?
- For example, do we use

$$y = \beta_0 + \beta_1 x + \beta_2 x^2 + \varepsilon$$

or

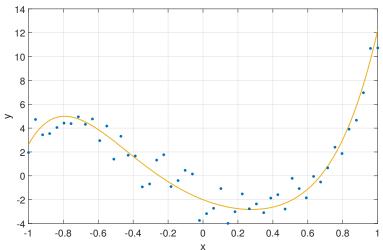
$$y = \beta_0 + \beta_1 x + \beta_2 x^2 + \beta_3 x^3 + \beta_4 x^4 + \beta_5 x^5 + \varepsilon$$

or another model with some other number of polynomial terms.

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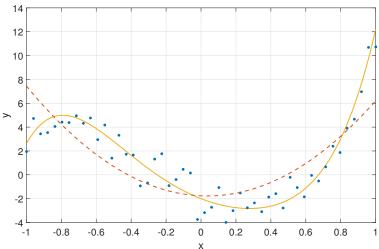
Underfitting/Overfitting Example (3)

• Example dataset of 50 samples



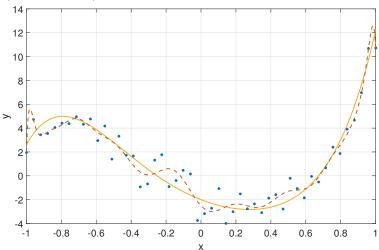
Underfitting/Overfitting Example (4)

• Use (x, x^2) , too simple – underfitting



Underfitting/Overfitting Example (5)

• Use (x, x^2, \dots, x^{20}) , too complex – overfitting

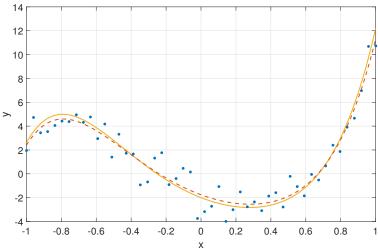




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Underfitting/Overfitting Example (6)

• (x, x^2, \dots, x^5) seems "just right". But how to find this model?



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Mean-Squared Prediction Error (1)

- Introduce the idea of a true, underlying (population) model
 this is the model that "generated" the data
- For regression:

$$\mathbb{E}\left[Y\,|\,x_1,\ldots,x_p\right] = f(x_1,\ldots,x_p)$$

• For binary classification:

$$\log\left(\frac{\mathbb{P}(Y=1\,|\,x_1,\ldots,x_p)}{1-\mathbb{P}(Y=1\,|\,x_1,\ldots,x_p)}\right) = f(x_1,\ldots,x_p)$$

- ullet In real life, $f(\cdot)$ could be any function
- For simplicity, let's focus on regression

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Mean-Squared Prediction Error (2)

- ullet We receive a sample of data from our population ${f y}=(y_1,\ldots,y_n)$
 - This is our training data
- ullet We use this to build a model $\hat{\mathcal{M}}(\mathbf{y}) \equiv \hat{\mathcal{M}}$ to estimate $f(\cdot)$
- ullet In linear regression, a model ${\mathcal M}$ is a particular linear combination of
 - predictors x_1, \ldots, x_p , and
 - non-linear transformations of the predictors
- We would like our model to predict well onto new data
 - We call this the model's ability to generalise
- How can we measure this?

Mean-Squared Prediction Error (3)

- ullet Imagine we receive new data $\mathbf{y}'=(y_1',\dots,y_m')$ from our population
 - This is called testing data
- ullet Once we have a model, $\hat{\mathcal{M}}(\mathbf{y})$, we can make predictions

$$\hat{y}_i(\hat{\mathcal{M}}(\mathbf{y})) \equiv \hat{y}(x'_{i,1}, \dots, x'_{i,p}, \hat{\mathcal{M}}(\mathbf{y}))$$

for the value of y_i' , given predictors $x_{i,1}',\ldots,x_{i,p}'$:

We can measure how well our model predicts the new data using

$$MSPE(\hat{\mathcal{M}}(\mathbf{y})) = \frac{1}{m} \sum_{i=1}^{m} (y_i' - \hat{y}_i(\hat{\mathcal{M}}(\mathbf{y})))^2$$

which is the mean-squared prediction error on our testing data

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Mean-Squared Prediction Error (4)

- Remember that our predictions $\hat{y}(x_1,\ldots,x_p,\hat{\mathcal{M}}(\mathbf{y}))$ are produced by a model that we learned using the training data \mathbf{y} \Rightarrow they depend on the training data that we have observed
- If we sampled different training data from our population, we would learn a different model, $\hat{\mathcal{M}}(\mathbf{y})$, and get different predictions $\hat{y}_i(\hat{\mathcal{M}}(\mathbf{y}))$
- To remove this dependency, we can average the $\mathrm{MSPE}(\hat{\mathcal{M}}(\mathbf{y}))$ over all the different models we could learn from all the different possible training samples we could draw from our population
- We can write this expected mean-squared prediction error as

$$EMSPE = \mathbb{E}\left[MSPE(\hat{\mathcal{M}}(\mathbf{y}))\right] = bias^2 + variance$$

where

- bias measures systematic error of our predictions;
- variance measures the error incurred through random variation

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Underfitting/Overfitting (1)

- We can demonstrate these ideas on our example
- The true model for this example was

$$\mathbb{E}[y \mid x] = f(x) = 9.7 x^5 + 0.8 x^3 + 9.4 x^2 - 5.7 x - 2$$

- To generate a sample from this true model:
 - I random sampled $\mathbf{x} = (x_1, \dots, x_{50})$ uniformly from (-1, 1)
 - ullet For each x_i value, I generated the samples y_i using

$$y_i = f(x_i) + \varepsilon$$

where $\varepsilon_i \sim N(0,1)$.

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Underfitting and Bias (1)

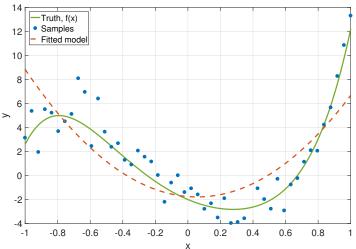
- When underfitting, our model is too simple to capture the truth
- In linear regression, this means we are omitting important predictors
- For our example, imagine we fit a quadratic polynomial to our data
 - The "truth" is a fith-order polynomial
- Regardless of how much training data from the population we see, we cannot make a quadratic look like a fifth order polynomial
- \bullet The inability of our model to fit the truth well leads to systematic error in predicting $f(\cdot)$
- The increase in EMSPE introduced by underfitting is therefore primarily due to bias

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17 / 71

Underfitting and Bias (2)

• Using $\mathcal{M} = \{x, x^2\}$ on a new sample – too simple

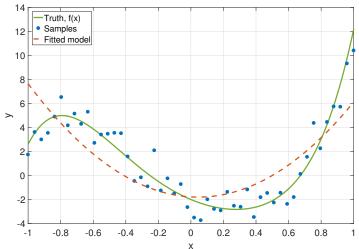


18 / 71

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Underfitting and Bias (3)

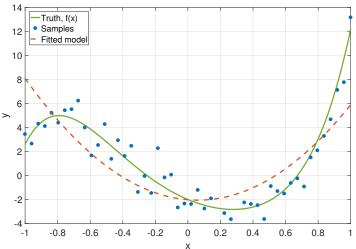
• Again, using $\mathcal{M} = \{x, x^2\}$ – again, too simple



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Underfitting and Bias (4)

• Again, using $\mathcal{M} = \{x, x^2\}$ – once more, too simple



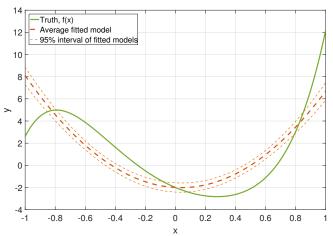
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Underfitting and Bias (5)

- Bias and variance were determined by simulation
- Generated 10,000 samples of size n=50 from our true model (our training data)
- ullet For each sample, fitted a 2nd order polynomial, i.e., x,x^2 terms
- Produced predictions for 5,000 equally spaced x' values in (-1,1) (our testing data) for each of the 10,000 different models
 - Calculated the average prediction at each x' value
 - ullet Calculated the variance of predictions at each x' value
 - Used these to get bias² and variance
- For visualisation:
 - ullet Calculated the average prediction at each x' value
 - ullet Calculated 2.75 and 97.5 percentiles of predictions at each x'

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Underfitting and Bias (6)



Fitting 2nd order polynomial to data generated from a 5th order polynomial; average fit is poor (high bias) but does not vary greatly (low variance). $EMSPE = 3.33, bias^2 = 3.27, variance = 0.058$

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22 / 71

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Overfitting and Variance (1)

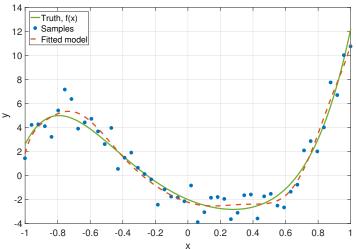
- When overfitting, our model is more complex than the truth
- In linear regression, this means we are including unassociated predictors
- The extra predictor introduces a relationship between predictor and target that is not really there
 - ⇒ we learn random "noise" from our training data
- In our example, imagine we fitted a 7th order polynomial to our data
 - The "truth" is a 5th order polynomial, so the coefficients for the 6th and 7th order polynomial terms should be zero
 - But when we learn them from our training data, due to random chance, they will never be exactly equal to zero
- The increase in EMSPE introduced by overfitting is therefore primarily due to variance

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Overfitting and Variance (2)

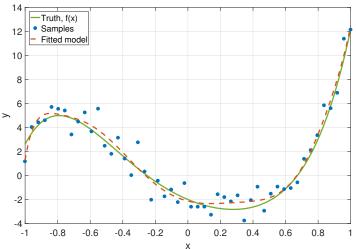
• Using (x, x^2, \dots, x^7) on a new sample – too complex



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Overfitting and Variance (3)

• Again, using (x, x^2, \dots, x^7) – too complex

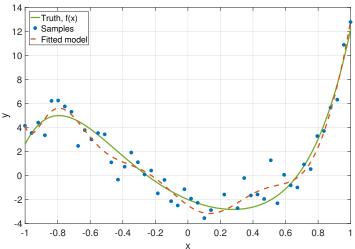


25 / 71

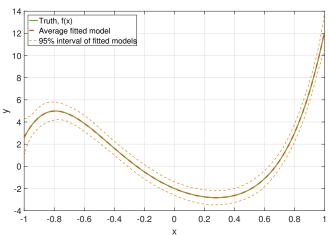
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Overfitting and Variance (4)

• Again, using (x, x^2, \dots, x^7) – too complex



Overfitting and Variance (5)

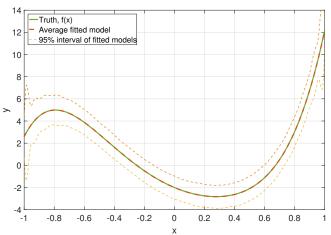


Fitting 7th order polynomial to data generated from a 5th order polynomial; average fit is good (no bias) but does varies more than 2nd order polynomial. $\underline{ \text{EMSPE}} = 0.1470, \underline{ \text{bias}}^2 = 0, \underline{ \text{variance}} = 0.1470$

27 / 71

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Overfitting and Variance (6)



Fitting 20th order polynomial to data generated from a 5th order polynomial; average fit is good (no bias) but variation is higher, especially near boundaries. EMSPE = 0.4860, $bias^2 = 0$, variance = 0.4860

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28 / 71

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Underfitting/Overfitting (1)

- In general:
 - Simpler models have higher bias, lower variance
 - More complex models have lower bias, higher variance
 - ullet Variance decreases with increasing n, bias does not
 - Variance increases with increasing number of predictors
- Recall EMSPE can be written as:

$$EMSPE = bias^2 + variance$$

- Model selection for prediction is about trading these off
 ⇒ select the right level of complexity for the data
- In linear/logistic regression, "complexity" is increased by including more predictors, or transformations of predictors
 - Omitting a predictor will reduce variance, but may lead to bias if predictor is important

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Underfitting/Overfitting (2)

- Model selection is quite subtle
- In linear models, including all predictors that are associated with target may not be enough to remove bias
 - If relationship between target and predictor is nonlinear, also need to include the right transformations ...
- You can overfit and underfit at the same time
 - Omit important predictor, include unassociated predictor ...
- Sometimes can do better by omitting associated predictors!
 - If association is weak and noise is large, may be too hard to get a good estimate of the coefficient β_j
 - Ergo, may be better to omit it
- Combining tools with real world knowledge/thinking can often improve predictions

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Hypothesis Testing to Select Predictors (1)

- Recap: can select predictors using hypothesis testing \implies we know that a predictor j is unimportant if $\beta_j = 0$
- So we can test the hypothesis:

$$H_0$$
 : $\beta_j = 0$ vs
$$H_A$$
 : $\beta_j \neq 0$

which, in this setting is a variant of the t-test (see Ross, Chapter 9 and Studio 6)

- Imagine we tried fitting each predictor by itself to the data
 - ullet Could choose to include predictor if $p ext{-value}$ was less than 0.05

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Hypothesis Testing to Select Predictors (2)

- \bullet If we were testing one hypothesis, this would guarantee probability of false discovery to be 5%
 - \bullet If predictor is unassociated, we would incorrectly conclude it to be associated for 5% of possible samples from our population
- ullet But if we have p predictors, we need to do p tests
- So, even if all predictors were unassociated, would expect 0.05p tests to have p-values < 0.05 just by chance
- If p is large, this can be big, i.e., if p=1000, we would falsely believe 50 predictors were associated just by chance on average
- This can easily happen if we include lots of transformations; e.g., trying all squares, logarithms, interactions of predictors
- Sometimes called data dredging

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Hypothesis Testing to Select Predictors (3)

- This is called the multiple testing problem
- We examine one way to resolve it: the Bonferroni procedure
- ullet Let lpha be our significance level
- This says we should instead reject null hypothesis only if

$$p\text{-value} < \frac{\alpha}{p}$$

• So if $\alpha=0.05$ and p=1000, we would need to see p-values of $0.05/1000=5\times 10^{-5}$ to believe that a predictor is associated with the target

Hypothesis Testing to Select Predictors (4)

- \bullet Bonferroni guarantees the probability of making at least one false discovery is 5%
- Very conversative approach, but can be confident that predictors we identify are very likely truly associated
- Related method called false discovery rate (FDR)
 ⇒ more power to discover associated predictors

34 / 71

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Likelihood

- The likelihood $p(\mathbf{y} | \hat{\mathcal{M}})$ is a general measure of goodness of fit
- The likelihood quantifies how much probability our model assigns to the data we have observed
 - \implies more probability, the more compatible the data is with our model
- Given a probability distribution $p(y \mid \theta)$ and estimates $\hat{\theta}$ for the parameters, compute likelihood by:
 - 1 Plug our estimated parameters $\hat{\theta}$ into our probability distribution
 - 2 Evaluate the probability assigned to each y_1, \ldots, y_n by this model, i.e.,

$$p(\mathbf{y} \mid \hat{\theta}) = \prod_{i=1}^{n} p(y_i \mid \hat{\theta})$$

- We have seen likelihoods of
 - Simple Gaussian, Poisson, Bernoulli models (Lecture 3, Studio 3)
 - Linear regression models (Lecture 6)
 - Logistic regression models (Lecture 7)

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35 / 71

Model Selection Criteria (1)

- We can use adjusted likelihood to guide selection of models
 can control complexity through model selection criteria
- A model selection criterion assigns a score to a model which
 - takes into account how well the model fits the data (likelihood)
 - takes into account how complex the model is
- Search for the model that minimises model selection criterion
 ⇒ in linear/logistic regression, find good set of predictors
- Let
 - \mathcal{M} denote a model (set of predictors to use);
 - $\hat{\mathcal{M}}$ denote the model fitted to data \mathbf{y} ;
 - $L(\mathbf{y} \,|\, \hat{\mathcal{M}})$ denote the minimised negative log-likelihood for the model \mathcal{M}
- Neg-log-likelihood always decreases with more predictors

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Model Selection Criteria (2)

- Penalize neg-log-likelihood of model by a complexity penalty
- The Akaike information criterion (AIC):

$$AIC(\hat{\mathcal{M}}) = L(\mathbf{y} | \hat{\mathcal{M}}) + k_{\mathcal{M}}$$

where $k_{\mathcal{M}}$ is number of predictors in model \mathcal{M}

• The Kullback-Leibler information criterion (KIC):

$$KIC(\hat{\mathcal{M}}) = L(\mathbf{y} | \hat{\mathcal{M}}) + 3/2 k_{\mathcal{M}}$$

- ullet Both of these tend to overfit, even for large n
 - Moderate overfitting less damaging for large n;
 - Both rarely underfit
- In practice, KIC often performs better than AIC

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Model Selection Criteria (3)

• The Bayesian information criterion (BIC):

$$BIC(\hat{\mathcal{M}}) = L(\mathbf{y} | \hat{\mathcal{M}}) + k_{\mathcal{M}}/2 \log n$$

where n is the number of samples used to fit $\hat{\mathcal{M}}$

- ullet As n increases, probability of overfitting goes to zero
- ullet More conservative; for small n can sometimes underfit
- The Risk inflation information criterion (RIC):

$$RIC(\hat{\mathcal{M}}) = L(\mathbf{y} | \hat{\mathcal{M}}) + k_{\mathcal{M}} \log p$$

where p is total number of possible predictors

- Similar to multiple testing idea; can be quite conservative
- ullet Recommended over AIC/KIC if p is very large
- Many refined criteria exist: MML, MDL, corrected AIC/KIC
 ⇒ often better for small n

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38 / 71

Cross validation (1)

- Cross-validation (CV) is a very general strategy
- Motivation: ideally, we would select model that minimises prediction error on future data y'_1, \ldots, y'_m

$$MSPE(\hat{\mathcal{M}}(\mathbf{y})) = \frac{1}{m} \sum_{i=1}^{m} (y_i' - \hat{y}_i(\hat{\mathcal{M}}(\mathbf{y})))^2$$

- ⇒ requires knowing our testing data ahead of time ...
- \bullet We don't have future testing data, but we know the data ${\bf y}$ we used to fit the model is generated by the same process
- So we can try and estimate the MSPE

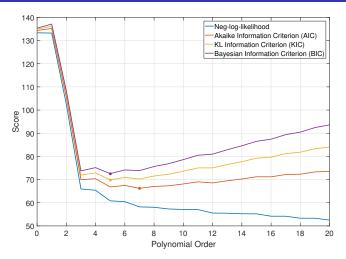
Cross validation (2)

- To estimate MSPE using cross validation
 - \bullet We randomly partition data y we have into two sets:
 - A training set y_{train},
 - ullet and a testing set $\mathbf{y}_{\mathrm{test}}$
 - 2 Fit a model ${\cal M}$ to the training data $y_{\rm train}$
 - $oldsymbol{\circ}$ Compute the MSPE of this model on the testing data \mathbf{y}_{test}
 - Repeat this procedure multiple times and average the MSPE
- This gives us an estimate of how well our model would predict future data from the same population
- Do this for all our different models, and choose the one that has the smallest CV error

Cross validation (3)

- K-fold CV
 - ullet split into K equal sized random partitions;
 - ullet train on K-1 partitions, test on the remaining partition
 - ullet repeat this m times
- The bigger m, the more stable the estimate
- Leave-one-out CV (LOO CV):
 - ullet train on n-1 of the samples, test on the remaining sample
 - ullet do this for all n possible partitions
- Cross-validation is:
 - simple to implement, very flexible;
 - potential slow if model fitting is computationally expensive
- ullet LOO CV is similar to AIC for large samples n
 - ⇒ can be very different for small samples

Example: fitting a polynomial



Model selection scores for fitting zero-order through twentieth-order polynomials to data generated by our example fifth-order polynomial. The likelihood always decreases with extra polynomial terms. The dots signify the models selected by each criterion.

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 - Overfitting and Underfitting
 - Model Selection Methods

- Penalized Regression Techniques
 - Statistical Stability
 - Penalized Regression



Conventional Model Selection Methods (1)

- All-subsets selection:
 - Try all combination of predictors to model with smallest model selection criterion score
- Forward selection algorithm:
 - Start with the empty model;
 - Find the predictor that reduces info criterion by most
 - If no predictor improves model, end.
 - Add this predictor to the model
 - Return to Step 2
- Backwards selection is related algorithm
 - Start with the full model and remove predictors

Conventional Model Selection Methods (2)

- Problems with traditional methods
- Statistically unstable
 - Small changes in data ⇒ big changes in model
- Multiple hypothesis testing problem
 - Potential issues with false positives
- All-subsets selection
 - ullet Infeasible for moderate to large number of predictors p
- Stepwise selection
 - Affected adversely by correlation in predictors
 - ullet Slow for large p



Statistical Instability (1)

- Example dataset (n = 354, p = 10)
 - Target: Y, measure of diabetes progression
 - Predictors:
 - AGE, SEX, BMI
 - BP: blood pressure
 - S1,...,S6: blood serum measurements
- Experiment
 - $oldsymbol{0}$ Removed one of the 354 samples at random
 - Pound "best" combination of predictors
 - Checked all possible (1024) combinations of predictors
 - Selected combination with smallest CV
- Repeated this four times



Statistical Instability (2)

Test	AGE	SEX	BMI	BP	S1	S2	S3	S4	S5	S6
1		✓	✓	✓	✓	✓	✓		✓	
2		\checkmark	✓	\checkmark	\checkmark	\checkmark	\checkmark			
3	\checkmark									
4	\checkmark		\checkmark							

Predictors chosen by selecting model that minimises cross-validation score using all-subsets selection method. For each test, a single one of the n=354 samples was removed at random.

• The "all-or-nothing" nature of this type of model selection (either exclude or include predictor) means that the "best" model can change substantially with only a small change to the data.

Penalized Regression (1)

- Alternative: apply a "penalty" to the coefficients
- For linear regression, we have penalized least-squares:

$$(\hat{\beta}_0, \hat{\beta}_{\lambda}) = \underset{\beta_0, \beta}{\operatorname{arg \, min}} \left\{ \operatorname{RSS}(\beta_0, \beta) + \lambda \sum_{j=1}^p g(\beta_j) \right\}$$

- In comparison regular least squares only minimises goodness-of-fit (RSS)
- How to choose the penalty function $g(\cdot)$?
 - \bullet Chosen to increase with increasing magnitude of β_j
- Generalises easily to logistic regression, i.e., penalized maximum likelihood

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Penalized Regression (2)

- Intuition:
 - Large coefficients imply strong effects
 - Coefficient of zero implies no association
- For this to make sense we need to standardise our predictors before fitting, i.e., ensure that

$$\sum_{i=1}^{n} x_{i,j} = 0 \text{ and } \sum_{i=1}^{n} x_{i,j}^{2} = n$$

- This can be easily done by subtracting mean and dividing by standard deviation
- Each predictor is then on the same (arbitrary) scale \Rightarrow implies that larger β_j means stronger association

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Penalized Regression (3)

- ullet The penalty includes a user-chosen hyperparameter λ
- ullet λ controls how strong the penalty is
- Generally,
 - ullet As λ goes to zero, the estimates of eta are closer to least-squares
 - ullet As λ gets larger, the estimates of eta become smaller
- ullet Choosing λ sets the "complexity" of the regression model
- ullet By varying λ we can generate a "path" of regression models
 - The models become more complex as we move along the path

Penalized Regression (4)

- Advantages of penalized regression:
 - Increased statistical stability
 - \bullet Can be applied even when the sample size is smaller than the number of predictors (the n < p setting)
 - Handles correlation between predictors
- Multi-collinearity occurs when at least two of the predictors are highly correlated
 - This means it is difficult to distinguish between their relationship with the target
 - In traditional least-squares this leads to increased variance
 - In stepwise methods, this dramatically increases instability

Ridge Regression (1)

- The first penalized method introduced
- Ridge regression solves

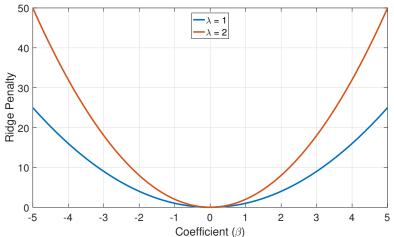
$$(\hat{\beta}_0, \hat{\beta}_{\lambda}) = \underset{\beta_0, \beta}{\operatorname{arg\,min}} \left\{ \operatorname{RSS}(\beta_0, \beta) + \lambda \sum_{j=1}^p \beta_j^2 \right\}$$

- Strengths:
 - \bullet Very quick to run, even for large n and p
 - Produces very stable, low variance estimates
 - Very appropriate for multi-collinearity
- Weaknesses:
 - Cannot estimate coefficients to be exactly zero
 - i.e., no variable selection



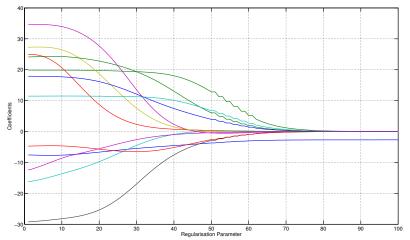
Ridge Regression (2)

• Penalty is the square of the coefficient (ℓ_2 penalty)



Ridge Regression (3)

• Example ridge regression coefficient path



Lasso Regression (1)

- A "sparse" estimator
- Lasso regression solves

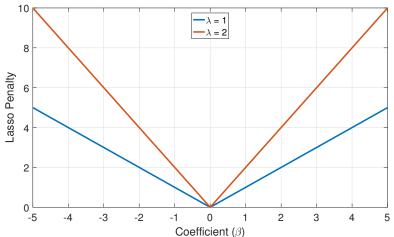
$$(\hat{\beta}_0, \hat{\beta}_{\lambda}) = \operatorname*{arg\,min}_{\beta_0, \beta} \left\{ \operatorname{RSS}(\beta_0, \beta) + \lambda \sum_{j=1}^p |\beta_j| \right\}$$

- Strengths:
 - Efficient algorithms exist
 - Can estimate coefficients to be exactly zero
 - i.e., can perform variable selection
- Weaknesses:
 - Produces biased estimates for large coefficients
 - Correlated predictors can be problematic
 - Can overfit (include unassociated predictors)

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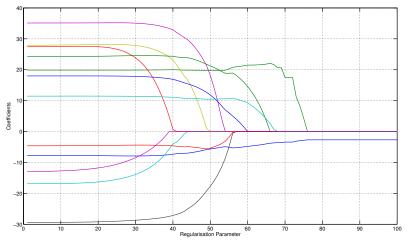
Lasso Regression (2)

• Penalty is the absolute value of the coefficient (ℓ_1 penalty)



Lasso Regression (3)

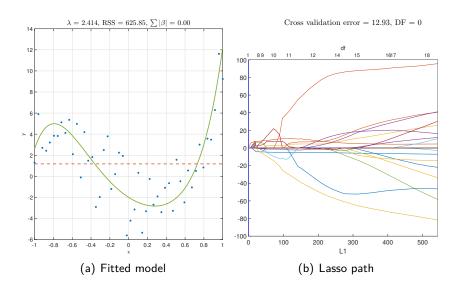
• Example lasso regression coefficient path

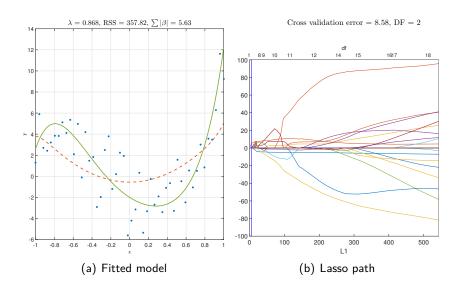


Selecting a Model Using Penalized Regression

- ullet How to choose the hyperparameter λ
- Standard procedure:
 - **1** Vary λ over some grid of values
 - 2 For each λ , use cross-validation to estimate prediction error
 - 3 Select λ with smallest cross-validation error
 - **1** Use that λ to estimate our final model from all the data
- Can also use methods like AIC, BIC, etc.
- Implementations:
 - MATLAB lasso() and lassoglm() functions
 - R, use the glmnet package

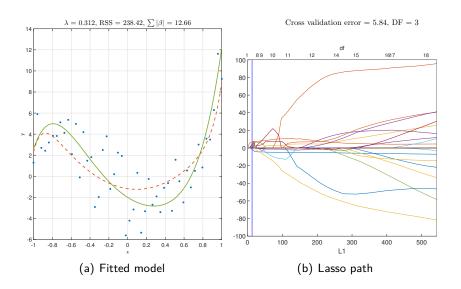
- ullet Example: we observe x and y data and want to build a prediction model for y using x
 - Data looks nonlinear so we use polynomial regression
 - We take $x, x^2, x^3, \dots, x^{20} \Rightarrow$ very flexible model
 - Which terms to include?
- For example, our model could be:
 - (x, x^2, x^3) ; or
 - (x, x^4, x^{16}, x^{17}) ; or
 - $(x^2, x^3, x^5, x^6, x^9)$; and so on ...
- Let us use Lasso regression

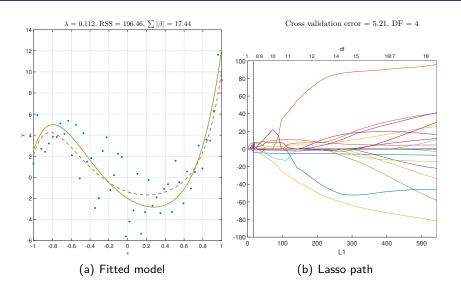




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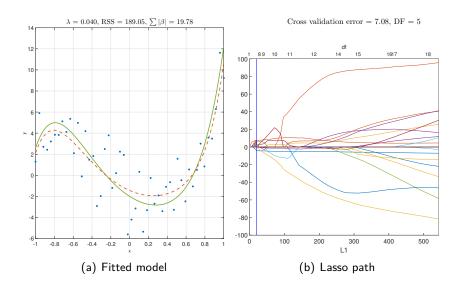
61 / 71

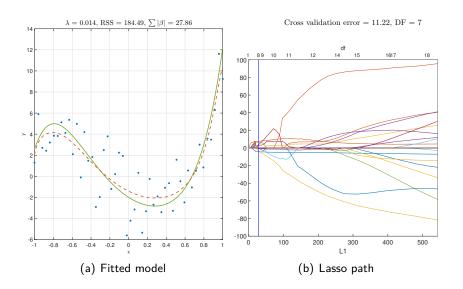


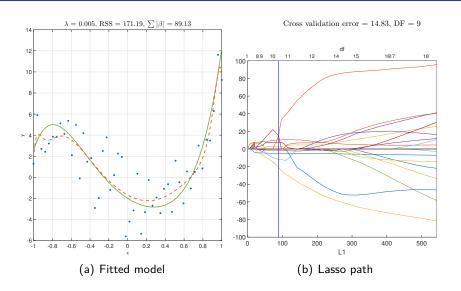


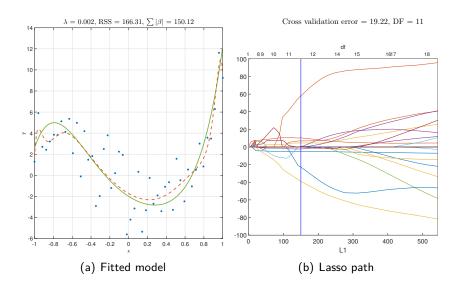
63 / 71

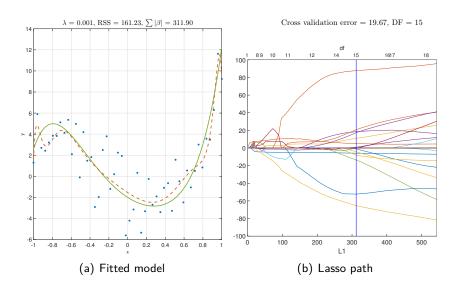
(Monash University) September 10, 2019

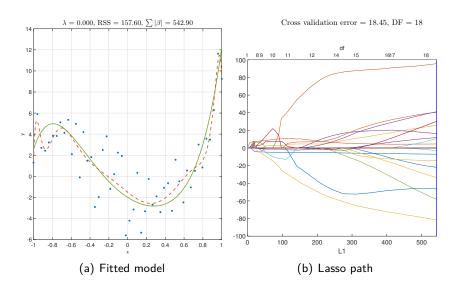












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69 / 71

Bias/Variance Trade-off

- Penalized regression has a bias/variance motivation
- Recall that the EMSPE of a model can be written as

$$EMSPE = bias^2 + variance$$

- If all relevant predictors included, least-squares is unbiased
 ⇒ however, LS variance can be high
- Penalized regression introduce some bias but reduces variance
 - They shrink the coefficients towards zero
- ullet Choosing a good λ means we can reduce the overall EMSPE

Reading/Terms to Revise

- Terms you should know:
 - Underfitting and bias
 - Overfitting and variance
 - Expected mean-squared prediction error
 - Multiple hypothesis testing, Bonferroni procedure
 - AIC, BIC
 - Cross-validation
 - Statistical instability
 - Penalized regression
 - Ridge/lasso regression
- Next week: non-linear machine learning algorithms for classification and regression (k nearest neighbours, decision trees, random forests).