bivariate regression

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outline

bivariate regression

stat significance (hypothesis testing)

basic measurement

math

- today we start some math
- important you understand it
- memorizing formulas is not enough to pass this class
- again, ask questions early!
- good idea to go over slides again after the class
- \diamond note hats: $\hat{\beta} \vee \beta$
- \diamond instead of $\sum_{i=1}^{n}$ i may just use \sum

looking ahead

- repeat today's material next week
 - · and extend a bit: esp stat significance
 - · and talk about measurment time
 - · over time, esp after midterm, class will get more applied
 - · and we will have more examples

outline

bivariate regression

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basic measurement

bivariate regression 5/56

the idea

- $\diamond Y \leftarrow X$, there is a directional relationship
- like in correlation, but here there is a direction
 - · (almost causality, but to argue causality you need also research design!)
- so we have outcome, or dependent variable predicted or affected by:
 - · independent variable (does not depend on the dependent variable),

bivariate regression 6/56

why regression?

- ols is the most fundamental technique for soc sci
- · and quite powerful
- things like anova, t-test, z-test, chi-sq test, etc are obsolete!
- just run regression! indeed, no studies use these anymore
 the only thing to remember from am1 is descriptive stats
- the only thing to remember from qm1 is descriptive stats, esp graphs

regression

if you want to figure out what predicts something, run

eg what will make you live longer, or which year wine is

examples

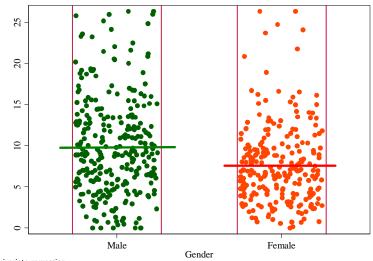
- see some of the useful things you can predict
 - · eg lexp=weighted avg(diet, exercise, smoking, etc)
 - eg lexp=50+2*(veggie serv/day)+3*(hrs at gym)-10*(packs of cigarettes per day) life expectancy http://www.northwesternmutual.com/ learning-center/the-longevity-game.aspx

[*]http://ianayres.yale.edu/prediction-tools

bivariate regression 8/56

conditional mean

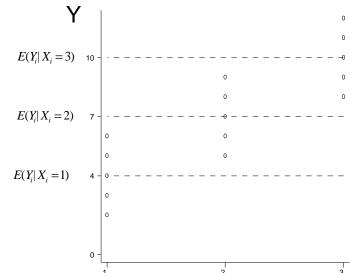
(useful way to start thinking about regression)



bivariate regression

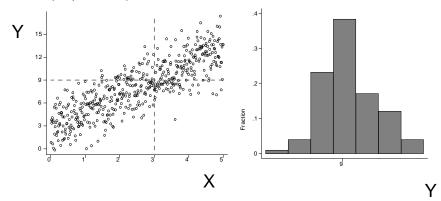
conditional mean of y depends on x

 \diamond for each value of x(1,2,3) E(y) is different



distribution of Y around the Expected Value

 \diamond e.g: E(Y|X=3)=9; Values of Y cluster around 9



bivariate regression 11/56

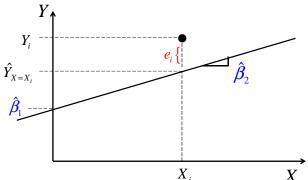
"regression" sounds scary

- regression is easy (yes, we will do all the tedious calculations), but all that regression does it fits a line that
 - minimizes the sum of the squared vertical distances in a scatter plot; hence "OLS" !
 - · sounds complicated but it's easy, too
- that's it! we will be just showing some math that can fit this line

bivariate regression 12/56

regression function

 $\diamond \hat{Y}_i = \hat{\beta}_1 + \hat{\beta}_2 X_i \quad Y_i = \hat{\beta}_1 + \hat{\beta}_2 X_i + e_i$



 \diamond (e_i) are errors of prediction

bivariate regression 13/56

what are the disturbance terms?

$$\diamond Y_i = \beta_1 + \beta_2 X_i + u_i$$

$$\diamond \ u_i = Y_i - \beta_1 - \beta_2 X_i = Y_i - E(Y|X_i)$$

- the combined effect of all other variables not in the model
- random events that affect the outcome
- errors of measurement in Y and X

bivariate regression 14/56

parameters v estimators

parameters (PRF) estimators (SRF)

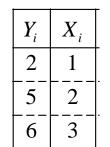
1	,	- /
β_1	$\hat{\beta}_1$	
eta_2	\hat{eta}_2	
μ	$ar{X}$	
p	$\hat{oldsymbol{ ho}}$	
σ	S	
μ_{i}	e_i	

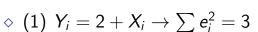
- estimators are based on samples
- parameters are fixed (and usually unknown)
- estimators have sampling distributions

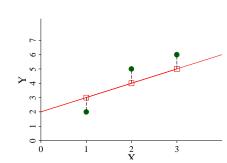
bivariate regression 15/56

first guess





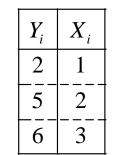


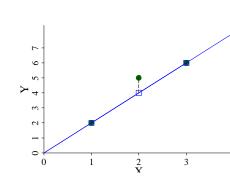


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second guess





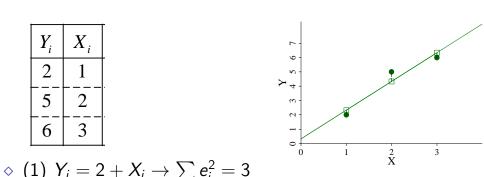


$$(1) Yi = 2 + Xi → ∑ ei2 = 3$$

$$(2) Yi = 0 + 2Xi → ∑ ei2 = 1$$

bivariate regression

example - you cannot beat ols!



♦ (2)
$$Y_i = 0 + 2X_i \rightarrow \sum e_i^2 = 1$$

♦ (3) $Y_i = 0.33 + 2X_i \rightarrow \sum e_i^2 = 0.67$

♦ dofile: guessing; then can use these est to predict like in

ols

$$\diamond Y_i = \hat{\beta}_1 - \hat{\beta}_2 X_i + e_i \rightarrow e_i = Y_i - \hat{\beta}_1 - \hat{\beta}_2 X_i$$

 \diamond chose estimators to minimize $\sum e_i^2 = \sum (Y_i - \hat{\beta}_1 - \hat{\beta}_2 X_i)^2$

* for elaboration and derivations see gujarati...

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intercept

Intercept: $\hat{\beta}_1 = \bar{\mathbf{Y}} - \hat{\beta}_2 \bar{\mathbf{X}}$

Note: sum of the residuals is zero: $\sum_{i=1}^{n} (Y_i - \hat{\beta}_1 - \hat{\beta}_2 X_i)$

bivariate regression 20/56

slope

$$\diamond \quad \hat{\beta}_2 \quad = \frac{\sum_{i=1}^n Y_i X_i - n\bar{X}\bar{Y}}{(\sum_{i=1}^n X_i^2 - n\bar{X}^2)}$$

$$\hat{\beta}_2 = \frac{\sum Y_i X_i - nYX}{\sum X_i^2 - n\bar{X}^2}$$

$$\hat{\beta}_2 = \frac{\sum (Y_i - \bar{Y})(X_i - \bar{X})}{\sum (X_i - \bar{X})^2}$$

$$\diamond \hat{\beta}_2 = \frac{\sum y_i x_i}{\sum x_i^2} \quad y_i = Y_i - \bar{Y} \quad x_i = X_i - \bar{X}$$

♦ Another way to look at the slope coefficient is the covariance of Y and X divided by the variance of X. Since the variance is always positive, the numerator (the covariance) will determine the sign of the slope.

bivariate regression 21/56

solving the problem

	Y_{i}	X_{i}	$ \begin{array}{c} \left(Y_i - \overline{Y}\right) \\ = y_i \end{array} $	$(X_i - \overline{X})$ $= x_i$	y_i^2	x_i^2	$y_i x_i$
	2	1	-2.33	-1	5.53	1	2.33
	5	2	0.67	0	0.45	0	0
	6	3	1.67	1	2.79	1	1.67
Σ	13	6	0	0	8.67	2	4
mean	4.33	2					

$$\hat{\beta}_2 = \frac{\sum y_i x_i}{\sum x_i^2} = \frac{4}{2} = 2$$

$$\hat{\beta}_1 = \bar{Y} - \hat{\beta}_2 \bar{X} = 4.33 - (2)(2) = 0.33$$

bivariate regression 22/56

example: age and fear

- In this example, imagine that we have some sort of survey that measures people's fear of crime, and that our hypothesis is that fear of crime increases with age. Assume the fear measure is an index ranging from 0 to 15.
- ♦ First, we calculate the means. Second, we calculate the deviations from the means and the their squares for each observation, as well as the co-product of the X and Y deviations. Finally, we sum these up.
 - blackboard! all steps!

bivariate regression 23/56

example: age and fear

The Data obs
$$|X_i| |Y_i|$$
 $\overline{X} = \frac{232}{5}$ Deviations from the means obs $|X_i| |Y_i|$ $|X_i| |Y_i|$ $|X_i| |X_i|$ $|X_i|$ $|X_i|$

$$\hat{\beta}_2 = \frac{\sum y_i x_i}{\sum x_i^2} = \frac{342}{1473} = .232$$

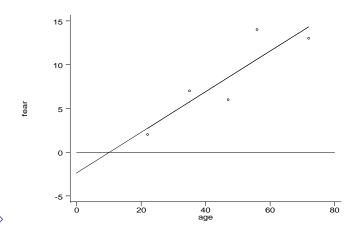
$$\diamond \ \hat{\beta}_1 = \bar{Y} - \hat{\beta}_2 \bar{X} = 8.4 - (.232)(46.4) = -2.365$$

$$\diamond \ \hat{Y}_i = \hat{\beta}_1 + \beta_2 X_i = -2.365 + .232 X_i$$

how would you interpret this?

bivariate regression 24/56

the estimated regression line



bivariate regression 25/56

variance and std error of regression

- ok, we know how to calculate betas and fit the line (that min the sum of the squared resid)
- but there are lines that fit better and lines that fit worse in different samples
 - draw good and bad fits with same betas
- we need a measure of uncertatinty, i.e. how well our line fit the data...
- and the fit is measured by residuals...
- ... so our measure of uncertainty has to do with residuals!

bivariate regression 26/56

variance and std error of regression

again, the mean of the residuals is zero (hence, \bar{e} drops out)

- ♦ why divide by n-2?
- \diamond s^2 and s are measures of the spread of the points around the estimated regression line.
- \diamond they are estimators of the variance and standard deviation of the disturbance terms: σ^2 and σ

bivariate regression 27/56

from \hat{Y} to s (se of reg) to $s_{\hat{\beta}_2}$ (se of slope)

$$\hat{Y}_i$$
 e_i e_i^2 e_i^5

$$\begin{vmatrix} \hat{Y_i} & e_i & e_i^2 \\ 2.739 & -0.739 & 0.546 & \diamond s = \sqrt{\frac{\sum_{i=1}^5 e_i^2}{n-2}} = \sqrt{\frac{21.7}{3}} = 0$$

1.556

6.447

11.377

1.793

21.713

-0.739

-2.539

3.373

-1.339

1.245

5.755

8.539

10.627

14.339

bivariate regression

$$\hat{Y} \mid e \mid e^2$$

27

earlier

 \Diamond

• what is it measuring?

 $\diamond \ \ s_{\hat{\beta}_2} = \frac{s}{\sum_{i=1}^5 x_i^2} = \frac{2.7}{\sqrt{1473}} = .07$

♦ how does it differ from s?

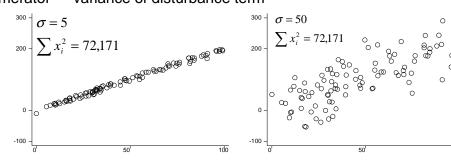
yahts are important! like

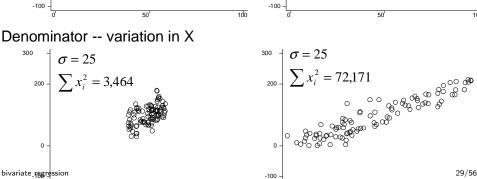
calc yhats and se of beta!!

our lexp we tried to predict

28/56

Standard Error of the Slope Coefficient Numerator -- variance of disturbance term





bivariate regression

ucla: hands-on dofile

- https://stats.idre.ucla.edu/stata/webbooks/reg
- let's just see a first reg output (you'll do it for ps2)
- · what is bivariate regression command?
- where is β_1 and β_2
- excellent for self study!!
- do it at home; and do ask me questions about it if any
- this is especially an excellent resource for final paper

bivariate regression 30/56

outline

bivariate regression

stat significance (hypothesis testing)

basic measurement

basic calculations blackboard; dofile

<u>Y</u>	Χ	У	y2	Х	x2	ху
1	17					
3	13					
5	8					
7	10					
9	2					
Cur						

25 50

$$\bar{Y} = 5$$
 $\bar{X} = 10$

the coefficients-interpretation

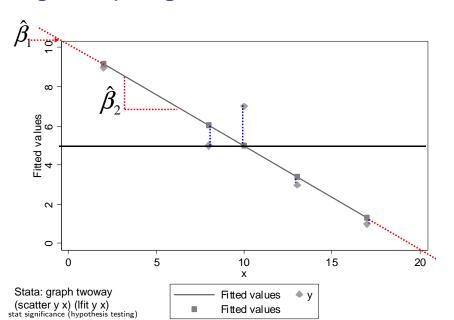
 Beta hat two is the slope coefficient. Thus, a one unit change in X leads to a 0.524 decrease in Y. Beta hat one is the intercept term. It is the predicted value for Y when X is equal to zero.

predicted val and resid blackboard; dofile

$$\diamond \hat{\mathbf{Y}}_i = \hat{\beta}_1 + \hat{\beta}_2 \mathbf{X}_i$$

- ♦ for obs 1:
- $\diamond \ \hat{Y}_1 = 10.24 + (-0.524)(17) = 1.332$
- $\diamond e_1 = 1 1.33 = -0.33$

regression plot again



se of the slope blackboard; dofile

$$\Rightarrow \sum e_i^2 = 5.42$$

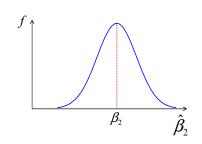
$$\Rightarrow s = \sqrt{\frac{\sum e_i^2}{n-2}} =$$

$$\diamond \ \ s_{\hat{\beta}_2} = \frac{s}{\sqrt{\sum x_i^2}}$$

· it gives us info about reliability (like sd or se) of slope

sampling distribution of the slope

probability distribution of $\hat{\beta}_2$ is centered on the true value of the parameter (i.e. unbiased) and is normally distributed with variance:



$$\diamond \ s_{\hat{\beta}_2}^2 = \frac{s^2}{\sum x_i^2}$$

$$\diamond~~ oldsymbol{s}_{\hat{eta}_2} = \sqrt{rac{s^2}{\sum x_i^2}} = rac{s}{\sqrt{\sum x_i^2}}$$

hypothesis test dofile

- the null is that slope ("the unobserved true parameter")
- · is zero (ie no effect)

$$\diamond H_0: \beta_2 = 0$$

$$\diamond H_A: \beta_2 \neq 0$$

$$\diamond t = \frac{\hat{\beta}_2 - \beta_2}{s_{\hat{\beta}_2}} = \frac{\hat{\beta}_2}{s_{\hat{\beta}_2}}$$

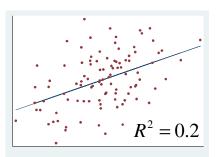
$$\diamond \text{ CI: } \hat{eta}_2 \pm (t_{n-2,\frac{\alpha}{2}})(s_{\hat{\beta}_2})$$

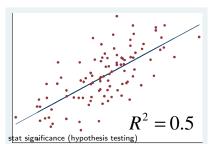
accounting for variation in Y blackboard in 3 colors

- \diamond before regression E[Y] = Y
- TSS total sum of squares $TSS = \sum_{i=1}^{n} (Y_i \bar{Y})^2$
- \diamond after regression $E[Y|X_i] = \hat{Y}_i = \hat{\beta}_1 + \hat{\beta}_2 X_i$

- $ESS = \sum_{i=1}^{n} (\hat{Y}_i \bar{Y})^2$
- · RSS residual sum of squares $RSS = \sum_{i=1}^{n} (Y_i \hat{Y}_i)^2 = \sum_{i=1}^{n} e_i^2$
- \diamond TSS = ESS + RSS

R^2 variation explained





$$\diamond TSS = ESS + RSS$$

$$\diamond \ 1 = \frac{\textit{ESS}}{\textit{TSS}} + \frac{\textit{RSS}}{\textit{TSS}}$$

⋄ R²: the percent of the variance in the dependent variable explained by the model

partitioning variance in Y dofile

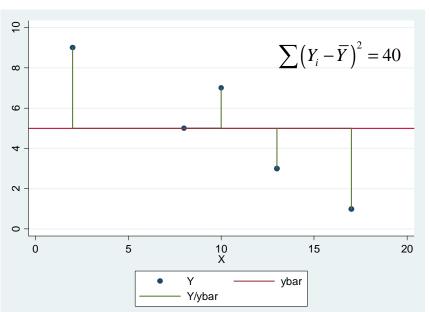
- \diamond before regression $E[Y_i] = Y$
- $TSS = \sum (Y_i \bar{Y})^2 = \sum y_i^2 = 40$
- \diamond after regression $E[Y_i|X_i] = \hat{Y}_i$
- $RSS = \sum (Y_i \hat{Y}_i)^2 = \sum e_i^2 = 5.43$
- ESS = TSS RSS = 40 5.4 = 34.57

$$\diamond R^2 = 1 - \frac{\sum e_i^2}{\sum y_i^2}$$

proportion of the total variance in the Y explained by Xs

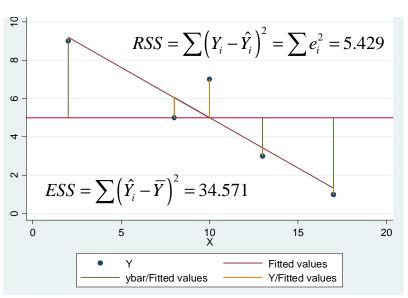
 $\diamond 0 < R^2 < 1$

TSS



stat significance (hypothesis testing)

RSS



exercise 1 dofile

you regressed car's price on its weight

```
price | Coef. Std. Err.
```

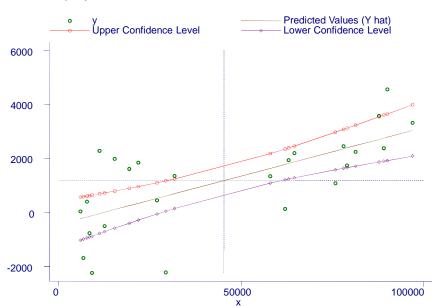
weight | 2.044063 .3768341

- interpret the coefficient
- is it significant ?
- ⋄ calculate 95% CI

reliability of predict val (se of E(Y|X)

- We have discussed the fact that parameter estimates are random variables, and so they have standard errors.
 Predicted values are also random variables because they are linear combinations of the coefficients.
- ♦ The further from the mean of X, the wider the confidence interval around the predicted value.
- leave it to software, no need to know the formula

se of E(Y|X) illustration dofile



anatomy of stata output dofile: outlier

regress DV IV

Source	SS	df	MS	Number of ob	os = n	
Model	$ESS = \sum (\hat{Y}_i - \overline{Y})$	$\binom{1}{2}$ 1		F(1, n-2)	=	
Residual	$RSS = \sum e_i^2$	•	$s^2 = \frac{RSS}{}$	Prob > F	$= \dots$ $= r^2$	
				R-squared Adj R-Square		
Total	$TSS = \sum (Y_i - \overline{Y})$	$\binom{n-1}{n}$	$s_Y^2 = \frac{155}{n-1}$	Root MSE	= s	
DV	Coef. Std.Err.	t	P> t	[95% Conf.	Interval]	
IV	$\hat{oldsymbol{eta}}_2$ $s_{\hat{oldsymbol{eta}}_2}$	$\left(\frac{\hat{eta}_2}{s_{\hat{eta}_2}}\right)$ 1	p val. for H_0 that $\beta_2 = 0$	$\hat{eta}_2 - t_{0.025} s_{\hat{eta}_2} \hat{eta}$	$\hat{\beta}_2 + t_{0.025} s_{\hat{\beta}_2}$	
Intercept	$\hat{oldsymbol{eta}}_{1}$ $S_{\hat{oldsymbol{eta}}_{1}}$	$ \frac{\left(\frac{\hat{\beta}_{1}}{s_{\hat{\beta}_{1}}}\right)}{\left(\frac{s_{\hat{\beta}_{1}}}{s_{\hat{\beta}_{1}}}\right)} 1 $	p val. for H_0 that $\beta_1 = 0$	$\hat{eta}_{\scriptscriptstyle 1} - t_{\scriptscriptstyle 0.025} s_{\hat{eta}_{\scriptscriptstyle 1}}$ $\hat{eta}_{\scriptscriptstyle 1}$	$\hat{\beta}_1 + t_{0.025} s_{\hat{\beta}_1}$	47/56

outline

bivariate regression

stat significance (hypothesis testing)

basic measurement

basic measurement 48/56

intuition

- what happens to betas if we change variables' measurement?
- · millions of dollars as opposed to dollars
- · curved grades (each person gets extra 10 points)
- · proportion of people in poverty v percent in poverty
- ⋄ income per capita v income per 100k people

basic measurement 49/56

add constant c to X or Y (say curved grades)

- ♦ if you add c to each obs, mean of var would change by that much
- but demeaned var doesn't change:

$$\diamond \ x_{i}^{'} = (X_{i}^{'} - \bar{X}^{'}) = [(X_{i} + c) - (\bar{X} + c)] = x_{i} \text{ same for Y}$$

- and nobody cares about intercept anyway, so let's spare our brain

basic measurement 50/

multiply X or Y by constant (say months, not years) think about it, assume some example

- · say year of educ produces \$2 increase in wage
- \diamond how about a month of educ? should be 1/12 of \$2!
- ♦ to convert yr to mo, multiply years by 12, right?
- if a person has 2yr of educ, that's 24mo
- \diamond so if i multiply X by c, say 12, I need to divide β_2 by 12 \diamond what if multiply Y?
- · again, say year of educ produces \$2 increase in wage
- ...or 200 cent increase in wage
 to get cents from dollars, I multiply dollars by 100
- · so if I multiply Y by 100, i get β_2 100x bigger

basic measurement

fun fact1: correlation v bivariate regression

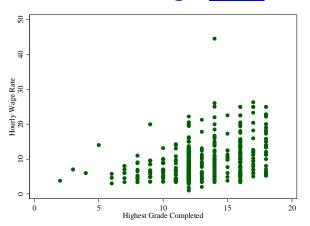
$$\diamond r = \frac{\sum y_i x_i}{\sqrt{(\sum x_i^2)(\sum y_i^2)}} \quad \hat{\beta}_2 = \frac{\sum y_i x_i}{\sum x_i^2}$$

bivariate slope equals corr coef scaled by std dev of Y and X.

$$\hat{\beta}_2 = \frac{\sum y_i x_i}{\sum x_i^2} = r(\frac{s_Y}{s_X})$$

basic measurement 52/56

education and wages dofile



. corr wage educ (obs=534)						
	wage	educ				
wage educ	1.0000 0.3819	1.0000				

. sum wage educ					
Variable	0bs	Mean	Std. Dev.	Min	Max
wage	534	9.023939	5.138876	1	44.5
educ	534	13.01873	2.615373	2	18

basic measurement 53/56

education and wages dofile

. regress wage educ

Source	SS		MS		Number of obs	
Model Residual	12022.2635	1 2053 532 22.5	.22494 982396		F(1, 532) Prob > F R-squared	= 0.0000 = 0.1459
	14075.4884				Adj R-squared Root MSE	
wage		Std. Err.	t	P> t	[95% Conf.	=
educ _cons	.7504488	.07873	9.532 -0.714	0.000 0.476	.5957891 -2.799576	.9051086 1.307678

The estimated regression line:

$$\widehat{wage}_i = \hat{\beta}_1 + \hat{\beta}_2 educ_i = -0.75 + 0.75 educ_i$$

Interpret the coefficients.

basic measurement 54/56

fun fact2: Z scores bivariate regression=correlation

- ⋄ z scores always have a mean of 0 and a variance (and standard deviation) of 1:

$$\hat{\beta}_2 = r_{Z_Y Z_X} \frac{s_{Z_Y}}{s_{Z_X}} = r_{YX}$$

$$\hat{\beta}_1 = \bar{z}_Y - \hat{\beta}_2 \bar{z}_X = 0 - r(0) = 0$$

 Thus, a regression of the z scores of Y on the z scores of X produces a slope equal to the correlation coefficient of X and Y and a zero intercept.

basic measurement 55/56

exercise 2: if no time do at home: see dofile

- confirm the above in stata using our simple data we started today's lecture with
- ⋄ run regression of Y on X
- modify X or Y and check what happened

basic measurement 56/56