

Numerical Methods

BINOMIAL EXPANSION

$$\binom{n}{r} = \binom{n}{n-r}; \quad \binom{n}{r} = \binom{n}{r} + \binom{n}{r-1} = \binom{n+1}{r}$$

$$(1+a)^n = 1 + an + \frac{a^2 n(n-1)}{2!} + \frac{a^3 n(n-1)(n-2)}{3!} + \dots$$

Graphical Method - split into linear, quadratic etc graphs and find intersections

$$\text{e.g. } e^x + x - 3 = 0 \Rightarrow e^x = 3 - x \quad \ln x - x^2 + 4 \Rightarrow \ln x = x^2 - 4$$

i. 1 root ii. 2 roots

Change of Sign - to locate roots where $f(x) = 0$ when $f(x)$ is sufficiently well-behaved
find solution in interval (a, b) so works out $f(a)$ and $f(b)$

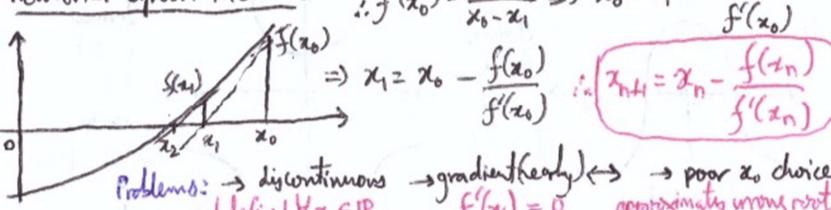
if there is a root in the interval (a, b) as there is a change in sign and $f(x)$ is continuous in this interval

Problems: → curve touches x -axis → several roots → discontinuity in $f(x)$
repeated root close together vertical asymptotes, jump discontin.

fixed Point Iteration

$$\text{e.g. } y = x^3 - 3x + 1, \text{ draw a sketch, pick } x_0 \\ \Rightarrow 3x = x^3 + 1 \therefore x = \frac{1}{3}(x^3 + 1) \\ \underline{x_{n+1} = \frac{1}{3}(x_n^3 + 1)} \\ \Rightarrow x^3 = 3x - 1 \therefore x = \sqrt[3]{3x - 1} \\ \underline{x_{n+1} = \sqrt[3]{3x_n - 1}}$$

Newton-Raphson Method



Problems: → discontinuous → gradient nearly 0 → poor x_0 choice

$$\text{not defined } \forall x \in \mathbb{R} \quad f'(x) = 0 \quad \text{approximates wrong root}$$

(converges of $x_{n+1} = g(x_n)$ to $x = a$ iff $|g'(a)| < 1$ and x_0 picked well.)

Numerical Integration

Trapezium Rule → concave \approx underestimate convex \approx overestimate

$$\text{Area} = \frac{1}{2} \times h \times \sum \text{parallel sides} = \frac{1}{2} h(y_1 + y_2) + \frac{1}{2} h(y_2 + y_3) + \dots$$

$$= \frac{1}{2} \times \text{strip width} \times (\text{first} + \text{last} + 2 \cdot \text{middles})$$

STATISTICS

Measures of Location - gives idea of size of numbers (median)
Measures of Spread - idea of variability (std, range, variance, semi-interquartile range)

Sampling Techniques

- o systematic sampling + representative - data must be listed + unlikely to be biased - practical consideration (data clumping)
- o simple random sampling - every sample of a given size is equally likely to be selected
- o random sampling - every item in a population has a non-zero probability of selection
- o stratified sampling - population split into clear groups / strata. Random samples taken within, which may be proportional.
- o cluster sampling - population is in clusters; which are selected (e.g. choosing several towns to represent one region)
- o quota sampling - in social surveys, quotas on diff. types of people to interview + easy to collect = likely biased
- o opportunity sampling - takes an easily available sample + easy to collect (e.g. delegates at a conference) - not representative, usually unreliable - only those who bothered to attend
- o self-selecting sampling - participants volunteer to take part in survey + easy to collect

Data Processing, interpolate in data range, though not explicitly data - extrapolate outside data range

o cleaning the data → exclude missing values / outliers o use appropriate sampling technique

o present in suitable diagrams o calculate summary measures

Binomial Distribution

$$(p+q=1)$$

→ fixed number of independent trials → two outcomes → constant p (equal likelihood)

$$\text{If } X \sim B(n, p), \text{ then } P(X=x) = \binom{n}{x} p^x (1-p)^{n-x}, E(X) = np, \text{Var}(X) = np(1-p)$$

Normal Distribution

→ continuous w/ symmetry → discrete w/ small intervals vs. 0 → continuous

$$\text{If } Y \sim N(\mu, \sigma^2), \text{ then } P(Y=y) = \frac{f(y)}{\sigma\sqrt{2\pi}} \cdot \exp\left(-\frac{1}{2}\left(\frac{y-\mu}{\sigma}\right)^2\right), E(Y) = \mu, \text{Var}(Y) = \sigma^2$$

PROBLEMS:

→ lack of symmetry: skewed, truncated, lack of tails

→ 3.s.d is unrealistic:

When standardised, $\left(\frac{x-\mu}{\sigma}\right)^2 = x^2$

KEY VALUES

$$1.\text{s.d. } P(-1 < Z < 1) = 68\%$$

$$2.\text{s.d. } P(-2 < Z < 2) = 95\% = \text{outliers}$$

$$3.\text{s.d. } P(-3 < Z < 3) = 99.7\%$$

Logarithm Laws $\log_a(x) = z \Leftrightarrow a^z = x$ Given $a = b^c \Leftrightarrow c = \log_b a$ $\log_a 1 = 0$ $\log_a a = 1$

$$\log_a(MN) = \log_a M + \log_a N \quad \log_a\left(\frac{M}{N}\right) = \log_a M - \log_a N \quad \log_a(M^p) = p \log_a M$$

$$\log_a(b) = \frac{\log_c(b)}{\log_c(a)}$$

$$\log_a(b) = \frac{1}{\log_b(a)}$$

$$\log_a(b)^c = \frac{\log_a(c)}{b}$$

Differentiation Laws $(uv)' = u'v + uv'$ $\left(\frac{u}{v}\right)' = \frac{u'v - uv'}{v^2}$ $(mn)' = nm^{n-1}$ $\frac{dy}{dx} = \frac{dy}{du} \cdot \frac{du}{dx}$

Integration Laws $\int \frac{g'(x)}{g(x)} dx = \ln|g(x)| + C$ $\int x^n dx = \frac{x^{n+1}}{n+1} + C$ $\int u'(u)^r du = \frac{u^{r+1}}{r+1} + C$

$\int \sin x dx = -\cos x + C$ $\int \cos x dx = \sin x + C$ $\int \sec^2 x dx = \tan x + C$ $\int \tan x \sec x dx = \sec x + C$

$\int \csc^2 x dx = -\cot x + C$ $\int \cot x dx = -\ln|\sin x| + C$ $\int_a^b uv' dx = uv \Big|_a^b - \int_a^b u'v dx$

Sequences & Series

$$\text{A.P. } S_n = \frac{n}{2} (a+l) = \frac{n}{2} [a + [a+(n-1)d]]$$

$$\text{G.P. } a_n = a \cdot r^{n-1} \quad S_n = a + ar + ar^2 + \dots + ar^{n-1} \quad \therefore u_n = u_{n+1}, n \rightarrow \infty$$

$$S_{\infty} = \frac{a(1-r^{\infty})}{1-r} \quad S_n(1-r) = a - ar^n = a(1-r^n) \quad \therefore u = \frac{10-3u}{u} \quad \Rightarrow u^2 + 3u - 10 = 0$$

$$S_n = \frac{a(1-r^n)}{1-r} \quad (u+5)(u-2) = 0 \quad \therefore u = -5$$

(convergent, divergent, periodic, oscillating)

MECHANICS

$$v = u + at \quad s = vt - \frac{1}{2}at^2 = ut + \frac{1}{2}et^2 = \frac{1}{2}(u+v)t \quad v^2 = u^2 + 2as$$

$$F = ma \quad F \leq \mu R \quad \mu = \tan \theta \quad W = mg \quad \text{moment} = F \cdot d$$

$$\text{time of flight } T = \frac{2u \sin \theta}{g} \quad \text{horizontal range } R = \frac{u^2 \sin 2\theta}{g} \quad \text{maximum height } h_{\max} = \frac{u^2 \sin^2 \theta}{2g} \quad \text{trajectory equation } y = xt \tan \theta - \frac{x^2 g}{2u^2 \cos^2 \theta}$$

wave-like - join together

Assumptions
→ no A.R.
→ negligible effect of curvature of Earth, rotation

Test Statistic: a statistic that is calculated from sample data in order to test a hypothesis about a population

Bivariate Data: data covers two variables in scatter graph

Random Variable: sum of probabilities of all possibilities = 1

discrete: can only take certain values

continuous: can take any value in a given range

Linear Transformations

$$\rightarrow ax + b = y \Rightarrow \bar{y} = a\bar{x} + b \quad \text{and} \quad \bar{y} = a\bar{x}$$

For standardisation, $X \sim N(\mu, \sigma^2) \Rightarrow X \sim N(0, 1)$:

$$a = \frac{1}{\sigma}, \quad b = -\frac{\mu}{\sigma} \Rightarrow z = \frac{1}{\sigma}x - \frac{\mu}{\sigma} \Rightarrow z = \frac{x-\mu}{\sigma}$$

$$\text{Var}(X) = \frac{\sum (x-\mu)^2}{n} \quad \sigma = \sqrt{\text{Var}(X)} = \sqrt{\frac{\sum x^2 - \mu^2}{n}}$$

Approximating the Binomial Distribution w/ Normal

→ n is large ($n > 50$) → p is close to $\frac{1}{2}$: $(0.1 < p < 0.9)$

Box Plot Advice !
0.5 × IQR outside it = outlier
use a scale
compare IQR, median, outliers

Box Plot Advice !
Use continuity correction when $<$ and \leq

$P(X < z) = 0.8 \Rightarrow P(Z < z) = 0.8$

$\therefore \text{InvNorm}(0.8) \Rightarrow z = 0.8416$

$\therefore z = \frac{x-\mu}{\sigma}$ to find missing μ or σ .

or use simultaneous eq. to find μ and σ for $2P(X=x)$'s.

Reversing Normal Calculations

$P(X < z) = 0.8 \Rightarrow P(Z < z) = 0.8$

$\therefore \text{InvNorm}(0.8) \Rightarrow z = 0.8416$

$\therefore z = \frac{x-\mu}{\sigma}$ to find missing μ or σ .

HYPOTHESIS TESTING

1. define H_0 and H_1 (e.g. $x \neq 0$ or $x > 0$ or $x < 0$), and let X be the r.v. "..."

2. model X as a distribution, state sig. level ($\alpha = ..$) and one/two-tailed

3. calculate relevant values, compared to c.v.

4. either \therefore reject H_0 : sufficient evidence at $\alpha = ..$
 \therefore accept H_0 : insufficient evidence at $\alpha = ..$

P.M.C.C. (r)

\leftrightarrow r correlation

\therefore correlation

\therefore correlation

\therefore lies inside critical region

\therefore lies outside critical region

\therefore acceptance region

\therefore critical region

FURTHER STATISTICS

$$\text{Probability} \quad {}^nC_r = \frac{n!}{r!(n-r)!} \quad {}^nP_r = \frac{n!}{(n-r)!}$$

Discrete Random Variables

X is the r.v. taking values x_i in a discrete distribution w/ $P(X=x_i) = p_i$

$$\mu = E(X) = \sum x_i p_i \quad \sigma^2 = \text{Var}(X) = \sum (x_i - \mu)^2 p_i = \sum x_i^2 p_i - \mu^2$$

$$\text{w.t.p. } \sum (x_i - \mu)^2 = \sum x_i^2 p_i - \mu^2 \quad \text{N.B.: must sum up to 1}$$

$$\text{L.H.S.} = \sum (x_i^2 - 2\mu x_i + \mu^2) p_i = \sum x_i^2 p_i - 2\mu \sum x_i p_i + \mu^2 \sum p_i \\ = \sum x_i^2 p_i - 2\mu^2 + \mu^2 = \sum x_i^2 p_i - \mu^2 \quad \square$$

Linear Transformations

$$E(aX+b) = aE(X) + b \quad \text{Var}(aX+b) = a^2 \text{Var}(X)$$

Binomial Distribution where $X \sim B(n, p)$ → fixed # of independent trials
→ constant p of success
→ two outcomes ($p+q=1$)

$$E(X) = np \quad \text{Var}(X) = np(1-p) \quad P(X=x) = \binom{n}{x} p^x (1-p)^{n-x}$$

w.t.p. $E(X) = np$

$$E(X) = \sum x_i p_i = 0 \cdot \binom{n}{0} \cdot p^0 (1-p)^n + 1 \cdot \binom{n}{1} p^1 (1-p)^{n-1} + 2 \cdot \binom{n}{2} p^2 (1-p)^{n-2} + \dots \\ \dots + (n-1) \cdot \binom{n}{n-1} p^{n-1} (1-p)^1 + n \cdot \binom{n}{n} p^n (1-p)^0 \\ = 0 + np(1-p)^{n-1} + \frac{2n(n-1)}{2!} p^2 (1-p)^{n-2} + \frac{3n(n-1)(n-2)}{3!} p^3 (1-p)^{n-3} + \dots \\ \dots + (n-1)np^{n-1}(1-p)^1 + np^n \\ = np \left[(1-p)^{n-1} + (n-1)p(1-p)^{n-2} + \frac{(n-1)(n-2)}{2!} p^2 (1-p)^{n-3} + \dots \right. \\ \left. \dots + (n-1)p^{n-2} + p^{n-1} \right] \\ = [p(1-p) + p]^{n-1} \times np = np \quad \square$$

Discrete Uniform Distribution where $X \sim U(n)$ over $1, 2, \dots, n$

$$E(X) = \frac{1}{2}(n+1) \quad \text{Var}(X) = \frac{1}{12}(n^2-1) \quad P(X=x) = \frac{1}{n}$$

w.t.p. $E(X) = \frac{1}{2}(n+1)$

$$E(X) = \sum x_i p_i = 1 \times \frac{1}{n} + 2 \times \frac{1}{n} + \dots + n \times \frac{1}{n} = \sum i \times \frac{1}{n} = \left[\frac{n(n+1)}{2} \right] \times \frac{1}{n} = \frac{1}{2}(n+1) \quad \square$$

w.t.p. $\text{Var}(X) = \frac{1}{12}(n^2-1)$

$$\text{Var}(X) = E(X^2) - [E(X)]^2 = \sum x_i^2 p_i - \left(\frac{n+1}{2}\right)^2 = \frac{1}{n}(1^2 + 2^2 + \dots + n^2) - \frac{1}{4}(n+1)^2 \\ = \frac{1}{6}n(n+1)(2n+1) - \frac{1}{4}(n+1)^2 = \frac{n+1}{12} \left[2(2n+1) - 3(n+1) \right] = \frac{(n+1)(n-1)}{12} = \frac{1}{12}(n^2-1)$$

Geometric Distribution, where $X \sim \text{Geo}(p)$ → no upper limit to # of trials
→ constant p of success for each trial
→ independent trials, 2 outcomes

$$E(X) = \frac{1}{p} \quad \text{Var}(X) = \frac{1-p}{p^2} \quad P(X=x) = p(1-p)^{x-1}$$

w.t.p. $E(X) = \frac{1}{p}$

$$E(X) = 1 \times p + 2 \times (1-p)p + 3 \times (1-p)^2 p + \dots$$

$$(1-p)[E(X)] = 1 \times (1-p)p + 2 \times (1-p)^2 p + \dots$$

$$E(X)[1-(1-p)] = p + p(1-p) + p(1-p)^2 + \dots$$

$$\therefore S_{xx} = \frac{a}{1-r} = \frac{p}{1-(1-p)} = 1 \Rightarrow p E(X) = 1 \quad \therefore E(X) = \frac{1}{p} \quad \square$$

Chi-Squared Tests (χ^2)

$$\text{Expected} = E_i = \frac{\text{row total} \times \text{column total}}{\text{grand total}}$$

Contingency Tables with m rows and n columns $\Rightarrow m \times n$ contingency tables

row
column
grand total

Add E_i into cell i in subscript. Then: grand total

$$\chi^2_{\text{calc}} = \sum \frac{(O_i - E_i)^2}{E_i} = \sum \frac{O_i^2}{E_i} - N$$

$$\chi^2_{\text{Yates}} = \sum \frac{(|O_i - E_i| - 0.5)^2}{E_i}$$

Yates' Correction when $v=1$ in contingency tables

Degrees of Freedom (v)

how many values required given Σ 's, to fill out the rest of the table

$$v = (n-1)(m-1)$$

when $m=1$, $v=n-1$

so combine rows/columns accordingly.

Chi-Squared Hypothesis Testing

H_0 : X is independent of Y

H_1 : X is not independent of Y

2. Work out χ^2_{calc} (or χ^2_{Yates}). independent vs. dependent

3. Find χ^2 from c.v. table

4. if $\chi^2_{\text{calc}} > \chi^2$, reject H_0 , sufficient evidence of association

$\chi^2_{\text{calc}} < \chi^2$, accept H_0 , insufficient evidence

Estimating Parameters - removes a degree of freedom

e.g. BINOMIAL value is O_i

GEOMETRIC $E(X) = np = \bar{x} = \frac{\sum x_i f_i}{N}$ $E(X) = \frac{1}{p} = \bar{x} \Rightarrow p = \frac{1}{\bar{x}}$

$$\Rightarrow p = \frac{\sum x_i f_i}{N}$$

Pearson's Product Moment Correlation Coefficient (P.M.C.C.)

Correlation does not imply causation.

$$S_{xx} = \sum (x - \bar{x})^2 = \sum x^2 - \frac{(\sum x)^2}{n}$$

$$S_{xy} = \sum (x - \bar{x})(y - \bar{y}) = \sum xy - \frac{\sum x \sum y}{n}$$

$$S_{yy} = \sum (y - \bar{y})^2 = \sum y^2 - \frac{(\sum y)^2}{n}$$

$$r = \frac{S_{xy}}{\sqrt{S_{xx} S_{yy}}} = \frac{\sum x_i y_i - \frac{\sum x_i \sum y_i}{n}}{\sqrt{(\sum x_i^2 - \frac{(\sum x_i)^2}{n})(\sum y_i^2 - \frac{(\sum y_i)^2}{n})}}$$

Testing Suitability of a Model

1. H_0 : a $B(n, p)$ distribution is suitable

2. H_1 : distribution is not suitable

$$N = \boxed{\quad} \quad \alpha = \boxed{\quad}$$

$$3. \chi^2_{\text{calc}} = \sum \frac{O_i^2}{E_i} - N$$

or equivalent

4. Conclude whether model is suitable

Effect of Coding on P.M.C.C.:

None, as it should be. $r = \frac{a \times b}{\sqrt{a^2 \times b^2}} = 1$