## Weak Law of Large Numbers

- Consider I.I.D. random variables X<sub>1</sub>, X<sub>2</sub>, ...
  - $X_i$  have distribution F with  $E[X_i] = \mu$  and  $Var(X_i) = \sigma^2$
  - Let  $\overline{X} = \frac{1}{n} \sum_{i=1}^{n} X_i$
  - For any  $\varepsilon > 0$ :

$$P(|\overline{X} - \mu| \ge \varepsilon) \xrightarrow{n \to \infty} 0$$

$$E[\overline{X}] = E\left[\frac{X_1 + X_2 + \dots + X_n}{n}\right] = \mu \quad \text{Var}(\overline{X}) = \text{Var}\left(\frac{X_1 + X_2 + \dots + X_n}{n}\right) = \frac{\sigma^2}{n}$$

By Chebyshev's inequality:

$$P(|\overline{X} - \mu| \ge \varepsilon) \le \frac{\sigma^2}{n\varepsilon^2} \xrightarrow{n \to \infty} 0$$

## Strong Law of Large Numbers

- Consider I.I.D. random variables X<sub>1</sub>, X<sub>2</sub>, ...
  - X<sub>i</sub> have distribution F with E[X<sub>i</sub>] = μ
  - Let  $\overline{X} = \frac{1}{n} \sum_{i=1}^{n} X_{i}$

$$P\left(\lim_{n\to\infty}\left(\frac{X_1+X_2+\ldots+X_n}{n}\right)=\mu\right)=1$$

- Strong Law ⇒ Weak Law, but not vice versa
- Strong Law implies that for any  $\varepsilon > 0$ , there are only a finite number of values of n such that condition of Weak Law:  $|\overline{X} - \mu| \ge \varepsilon$  holds.

## Intuitions and Misconceptions of LLN

- · Say we have repeated trials of an experiment
  - Let event E = some outcome of experiment
  - Let X<sub>i</sub> = 1 if E occurs on trial i, 0 otherwise
  - Strong Law of Large Numbers (Strong LLN) yields:  $\frac{X_1 + X_2 + \dots + X_n}{X_1 + X_2 + \dots + X_n} \to E[X] = P(E)$

  - Strong LLN justifies "frequency" notion of probability
  - · Misconception arising from LLN:
    - o Gambler's fallacy: "I'm due for a win"
    - o Consider being "due for a win" with repeated coin flips...

#### La Loi des Grands Nombres





- 1713: Weak LLN described by Jacob Bernoulli 1835: Poisson calls it "La Loi des Grands Nombres"
  - That would be "Law of Large Numbers" in French
- 1909: Émile Borel develops Strong LLN for Bernoulli random variables



- 1928: Andrei Nikolaevich Kolmogorov proves Strong LLN in general case
  - 2011: Another year passes in which Charlie Sheen does not make use of LLN



o I'm still holding out hope for 2012...



And now a moment of silence...

...before we present...

...the greatest result of probability theory!

## The Central Limit Theorem (CLT)

- Consider I.I.D. random variables X<sub>1</sub>, X<sub>2</sub>, ...
  - X<sub>i</sub> have distribution F with E[X<sub>i</sub>] = μ and Var(X<sub>i</sub>) = σ<sup>2</sup>

$$\frac{X_1 + X_2 + \ldots + X_n - n\mu}{\sigma\sqrt{n}} \rightarrow N(0,1) \text{ as } n \rightarrow \infty$$

More intuitively:

Demo

- ∘ Let  $\overline{X} = \frac{1}{n} \sum_{i=1}^{n} X_{i}$ ∘ Central Limit Theorem:  $\overline{X} \sim N(\mu, \frac{\sigma^{2}}{n})$  as  $n \to \infty$ ∘ Now let  $Z = \frac{\overline{X} \mu}{\sqrt{\sigma^{2}/n}}$ , noting that  $Z \sim N(0, 1)$ :

$$\overline{X} \sim N(\mu, \frac{\sigma^2}{n}) \Leftrightarrow Z = \frac{\frac{1}{n} \left( \sum_{i=1}^n X_i \right) - \mu}{\sqrt{\sigma^2/n}} = \frac{n \left[ \frac{1}{n} \left( \sum_{i=1}^n X_i \right) - \mu \right]}{n \sqrt{\sigma^2/n}} = \frac{\left( \sum_{i=1}^n X_i \right) - n \mu}{\sigma \sqrt{n}}$$

#### No Limits for Central Limit Theorem

- · History of the Central Limit Theorem
  - 1733: CLT for X ~ Ber(1/2) postulated by Abraham de Moivre





- 1823: Pierre-Simon Laplace extends de Moivre's work to approximating Bin(n, p) with Normal
- 1901: Aleksandr Lyapunov provides precise definition and rigorous proof of CLT





- 2003: Charlie Sheen stars in television series "Two and Half Men"
  - By end of the 7th season (last year), there were 161 episodes
  - Mean quality of subsamples of episodes is Normally distributed (thanks to the Central Limit Theorem)

#### Central Limit Theorem in Real World

- · CLT is why many things in "real world" appear Normally distributed
  - Many quantities are sum of independent variables
  - Exams scores
    - 。Sum of individual problems
  - Election polling
    - Ask 100 people if they will vote for candidate X (p<sub>1</sub> = # "yes"/100)
    - 。Repeat this process with different groups to get p1, ..., pn
    - Have a normal distribution over p<sub>i</sub>
    - Can produce a "confidence interval
      - · How likely is it that estimate for true p is correct
      - · We'll do an example like that soon

## This is Your Midterm on the CLT

- Start with 180 midterm scores: X<sub>1</sub>, X<sub>2</sub>, ..., X<sub>180</sub>
  - $E[X_i] = 68.9$  and  $Var(X_i) = 611.37$
  - Created 18 disjoint samples of size n = 10

$$\ \, Y_1 = \{X_1, \ldots, \, X_{10}\}, \, Y_2 = \{X_{11}, \ldots, \, X_{20}\}, \, Y_i = \{X_{10i\cdot 9}, \ldots, \, X_{10i}\}$$

$$\overline{Y}_i = \frac{1}{10} \sum_{i=1}^{10i} Y_i$$

 $\overline{Y_i}=\frac{1}{10}\sum_{j=10.9}^{10}Y_j$  • Prediction by CLT:  $\overline{Y_i}\sim N(68.9,~611.37/18\approx 33.965)$ 

$$Z_{i} = \frac{\overline{Y_{i}} - E[X_{i}]}{\sqrt{\sigma^{2}/n}} = \frac{\overline{Y_{i}} - 68.9}{\sqrt{611.37/18}} \qquad \overline{Z} = \frac{1}{18} \sum_{i=1}^{18} Z_{i} = 8.2 \times 10^{-16} \qquad \text{Var}(\overline{Z}) = 0.914$$

## **Estimating Clock Running Time**

- · Have new algorithm to test for running time
  - Mean (clock) running time:  $\mu = t \sec$ .
  - Variance of running time: σ<sup>2</sup> = 4 sec<sup>2</sup>.
  - Run algorithm repeatedly (I.I.D. trials), measure time
    - $_{\circ}$  How many trials so estimated time =  $t\pm0.5$  with 95% certainty?
    - ∘  $X_i$  = running time of i-th run (for  $1 \le i \le n$ )
    - By Central Limit Theorem, Z ~ N(0, 1), where:

$$Z_{n} = \frac{\left(\sum_{i=1}^{n} X_{i}\right) - n\mu}{\sigma \sqrt{n}} = \frac{\left(\sum_{i=1}^{n} X_{i}\right) - nt}{2\sqrt{n}}$$

$$P(-0.5 \le \frac{\sum_{i=1}^{n} X_{i}}{n} - t \le 0.5) = P(\frac{-0.5\sqrt{n}}{2} \le \frac{\sqrt{n}}{2} \left(\sum_{i=1}^{n} X_{i}\right) - nt}{2} \le \frac{0.5\sqrt{n}}{2}) = P(\frac{-0.5\sqrt{n}}{2} \le Z_{n} \le \frac{0.5\sqrt{n}}{2})$$

$$= \Phi(\frac{\sqrt{n}}{4}) - \Phi(\frac{-\sqrt{n}}{4}) = \Phi(\frac{\sqrt{n}}{4}) - (1 - \Phi(\frac{\sqrt{n}}{4})) = 2\Phi(\frac{\sqrt{n}}{4}) - 1 \approx 0.95 \implies \Phi(\frac{\sqrt{n^{*}}}{4}) = 0.975$$

$$\bullet \text{ Solve for } n^{*}: \frac{J^{n}}{n} = 1.96 \implies n^{*} = \left[(7.84)^{*}\right] = 62$$

# Estimating Time With Chebyshev

- · Have new algorithm to test for running time
  - Mean (clock) running time:  $\mu = t \sec$ .
  - Variance of running time:  $\sigma^2 = 4 \text{ sec}^2$ .
  - Run algorithm repeatedly (I.I.D. trials), measure time
    - $_{\circ}$  How many trials so estimated time =  $t\pm$  0.5 with 95% certainty?
    - $X_i$  = running time of *i*-th run (for 1 ≤ i ≤ n)
  - What would Chebyshev say?  $P(|X_s \mu_s| \ge k) \le \frac{\sigma_s^2}{k^2}$

$$\mu_{S} = E\left[\sum_{i=1}^{n} \frac{X_{i}}{n}\right] = t \qquad \sigma_{S}^{2} = \operatorname{Var}\left(\sum_{i=1}^{n} \frac{X_{i}}{n}\right) = n \frac{\sigma^{2}}{n^{2}} = \frac{4}{n}$$

$$P(\left|\sum_{i=1}^{n} \frac{X_i}{n} - t\right| \ge 0.5) \le \frac{4/n}{(0.5)^2} = \frac{16}{n} = 0.05 \implies n \ge 320$$

o Thanks for playing Pafnuty...

# Crashing Your Web Site

- Number visitors to web site/minute: X ~ Poi(100)
  - Server crashes if ≥ 120 requests/minute
  - What is P(crash in next minute)?
  - Exact solution:  $P(X \ge 120) = \sum_{i=120}^{\infty} \frac{e^{-100}(100)^i}{i!} \approx 0.0282$
  - Use CLT, where  $Poi(100) \sim \sum_{n=0}^{\infty} Poi(100/n)$  (all I.I.D)

$$P(X \ge 120) = P(X \ge 119.5) = P(\frac{X - 100}{\sqrt{100}} \ge \frac{119.5 - 100}{\sqrt{100}}) = 1 - \Phi(1.95) \approx 0.0256$$

- 。Note: Normal can be used to approximate Poisson
- I'll give you one more chance (one-sided) Chebyshev:  $P(X \ge 120) = P(X \ge E[X] + a) \le \frac{\sigma^2}{\sigma^2 + a^2} = \frac{100}{100 + 20^2} = 0.2$

$$P(X \ge 120) = P(X \ge E[X] + a) \le \frac{\sigma^2}{\sigma^2 + a^2} = \frac{100}{100 + 20^2} = 0.2$$



It's play time!

# Sum of Dice

- You will roll 10 6-sided dice (X1, X2, ..., X10)
  - X = total value of all 10 dice = X<sub>1</sub> + X<sub>2</sub> + ... + X<sub>10</sub>
  - Win if:  $X \le 25$  or  $X \ge 45$
  - Roll!
  - And now the truth (according to the CLT):  $E[X] = 10E[X_i] = 10(3.5) = 35 \qquad \text{Var}(X) = 10 \text{ Var}(X_i) = 10 \frac{35}{12} = \frac{350}{12}$  $1 P(25.5 \le X \le 44.5) = 1 P(\frac{25.5 35}{\sqrt{350/12}} \le \frac{X 35}{\sqrt{350/12}} \le \frac{44.5 35}{\sqrt{350/12}})$  $\approx 1 (2\Phi(1.76) 1) \approx 2(1 0.9608) = 0.0784$
  - If only Chebyshev were right...  $P(|X \mu| \ge k) = P(|X 35| \ge 10) \le \frac{\sigma^2}{k^2} = \frac{350/12}{100} \approx 0.292$