Whither the Binomial...

- · Recall example of sending bit string over network
 - n = 4 bits sent over network where each bit had independent probability of corruption p = 0.1
 - X = number of bit corrupted. X ~ Bin(4, 0.1)
 - In real networks, send large bit strings (length $n \approx 10^4$)
 - Probability of bit corruption is very small $p \approx 10^{-6}$
 - X ~ Bin(10⁴, 10⁻⁶) is unwieldy to compute
- Extreme *n* and *p* values arise in many cases
 - # bit errors in file written to disk (# of typos in a book)
 - # of elements in particular bucket of large hash table
 - # of servers crashes in a day in giant data center
 - # Facebook login requests that go to particular server

Binomial in the Limit

· Recall the Binomial distribution

$$P(X = i) = \frac{n!}{i!(n-i)!} p^{i} (1-p)^{n-i}$$

• Let $\lambda = np$ (equivalently: $p = \lambda/n$)

$$P(X=i) = \frac{n!}{i!(n-i)!} \left(\frac{\lambda}{n}\right)^{i} \left(1 - \frac{\lambda}{n}\right)^{n-i} = \frac{n(n-1)...(n-i+1)}{n^{i}} \frac{\lambda^{i}}{i!} \frac{(1 - \lambda/n)^{n}}{(1 - \lambda/n)^{i}}$$

When n is large, p is small, and λ is "moderate":

$$\frac{n(n-1)...(n-i+1)}{n^{i}} \approx 1 \qquad (1-\lambda/n)^{n} \approx e^{-\lambda} \qquad (1-\lambda/n)^{i} \approx 1$$

• Yielding: $P(X=i) \approx 1 \frac{\lambda^i}{i!} \frac{e^{-\lambda}}{1} = \frac{\lambda^i}{i!} e^{-\lambda}$

Poisson Random Variable

- X is a <u>Poisson</u> Random Variable: X ~ Poi(λ)
 - X takes on values 0, 1, 2...
 - and, for a given parameter $\lambda > 0$,
 - has distribution (PMF):

$$P(X=i) = e^{-\lambda} \frac{\lambda^i}{i!}$$

- Note Taylor series: $e^{\lambda} = \frac{\lambda^0}{0!} + \frac{\lambda^1}{1!} + \frac{\lambda^2}{2!} + \dots = \sum_{i=0}^{\infty} \frac{\lambda^i}{i!}$
- So: $\sum_{i=0}^{\infty} P(X=i) = \sum_{i=0}^{\infty} e^{-\lambda} \frac{\lambda^i}{i!} = e^{-\lambda} \sum_{i=0}^{\infty} \frac{\lambda^i}{i!} = e^{-\lambda} e^{\lambda} = 1$

Sending Data on Network Redux

- Recall example of sending bit string over network
 - Send bit string of length $n = 10^4$
 - Probability of (independent) bit corruption $p = 10^{-6}$
 - $X \sim Poi(\lambda = 10^4 * 10^{-6} = 0.01)$
 - What is probability that message arrives uncorrupted?

$$P(X=0) = e^{-\lambda} \frac{\lambda^{i}}{i!} = e^{-0.01} \frac{(0.01)^{0}}{0!} \approx 0.990049834$$

• Using Y ~ Bin(10^4 , 10^{-6}):

$$P(Y = 0) \approx 0.990049829$$

Caveat emptor: Binomial computed with built-in function in R software package, so some approximation may have occurred. Approximation are closer to you than they may appear in some software packages.

Simeon-Denis Poisson

 Simeon-Denis Poisson (1781-1840) was a prolific French mathematician



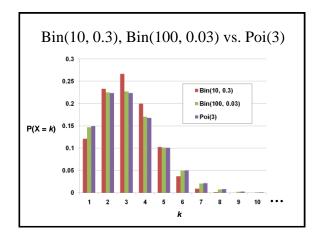


- Published his first paper at 18, became professor at 21, and published over 300 papers in his life
 - He reportedly said "Life is good for only two things, discovering mathematics and teaching mathematics."
- · Definitely did not look like Charlie Sheen

Poisson Random is Binomial in Limit

- Poisson approximates Binomial where n is large, p is small, and λ = np is "moderate"
- · Different interpretations of "moderate"
 - n > 20 and p < 0.05
 - n > 100 and p < 0.1
- Really, Poisson is Binomial as

$$n \rightarrow \infty$$
 and $p \rightarrow 0$, where $np = \lambda$



Tender (Central) Moments with Poisson

- Recall: Y ~ Bin(n, p)
 - E[Y] = np
 - Var(Y) = np(1 p)
- $X \sim Poi(\lambda)$ where $\lambda = np$ $(n \rightarrow \infty \text{ and } p \rightarrow 0)$
 - $E[X] = np = \lambda$
 - $Var(X) = np(1-p) = \lambda(1-0) = \lambda$
 - Yes, expectation and variance of Poisson are same
 It brings a tear to my eye...
 - Recall: Var(X) = E[X²] (E[X])²
 - $E[X^2] = Var(X) + (E[X])^2 = \lambda + \lambda^2 = \lambda(1 + \lambda)$

It's Really All About Raisin Cake



- · Bake a cake using many raisins and lots of batter
- · Cake is enormous (in fact, infinitely so...)
 - Cut slices of "moderate" size (w.r.t. # raisins/slice)
 - Probability p that a particular raisin is in a certain slice is very small (p = 1/# cake slices)
- Let X = number of raisins in a certain cake slice
- X ~ Poi(λ), where $\lambda = \frac{\text{\# raisins}}{\text{\# cake slices}}$

CS = Baking Raisin Cake With Code

- · Hash tables
 - strings = raisins
 - buckets = cake slices
- · Server crashes in data center
 - servers = raisins
 - list of crashed machines = particular slice of cake
- Facebook login requests (i.e., web server requests)
 - requests = raisins
 - server receiving request = cake slice

Defective Chips

- · Computer chips are produced
 - p = 0.1 that a chip is defective
 - Consider a sample of n = 10 chips
 - What is P(sample contains ≤ 1 defective chip)?
 - Using Y ~ Bin(10, 0.1):

$$P(Y \le 1) = {10 \choose 0} (0.1)^0 (1 - 0.1)^{10} + {10 \choose 1} (0.1)^1 (1 - 0.1)^9 \approx 0.7361$$

• Using X ~ Poi(
$$\lambda$$
 = (0.1)(10) = 1)

$$P(X \le 1) = e^{-1} \frac{1^0}{0!} + e^{-1} \frac{1^1}{1!} = 2e^{-1} \approx 0.7358$$

Efficiently Computing Poisson

- Let X ~ Poi(λ)
 - Want to compute P(X = i) for multiple values of i
 - E.g., Computing $P(X \le a) = \sum_{i=1}^{n} P(X = i)$
- Iterative formulation:
 - Compute P(X = i + 1) from P(X = i)

$$\frac{P(X=i+1)}{P(X=i)} = \frac{e^{-\lambda} \lambda^{i+1} / (i+1)!}{e^{-\lambda} \lambda^{i} / i!} = \frac{\lambda}{i+1}$$

• Use recurrence relation:

$$P(X = 0) = e^{-\lambda} \frac{\lambda^0}{0!} = e^{-\lambda}$$
$$P(X = i+1) = \frac{\lambda}{i+1} P(X = i)$$

Approximately Poisson Approximation

- Poisson can still provide good approximation even when assumptions "mildly" violated
- · "Poisson Paradigm"
- · Can apply Poisson approximation when...
 - "Successes" in trials are not entirely independent
 Example: # entries in each bucket in large hash table
 - · Probability of "Success" in each trial varies (slightly)
 - Small relative change in a very small p
 - Example: average # requests to web server/sec. may fluctuate slightly due to load on network

Birthday Problem Redux

- What is the probability that of n people, none share the same birthday (regardless of year)?
 - $\binom{n}{2}$ trials, one for each pair of people (x, y), $x \neq y$
 - Let $E_{x,y} = x$ and y have same birthday (trial success)
 - $P(E_{x,y}) = p = 1/365$ (note: all $E_{x,y}$ not independent)

• X ~ Poi(
$$\lambda$$
) where $\lambda = \binom{n}{2} \frac{1}{365} = \frac{n(n-1)}{730}$

$$P(X=0) = e^{-n(n-1)/730} \frac{(n(n-1)/730)^0}{0!} = e^{-n(n-1)/730}$$

P(X=0) = eO!
Solve for smallest integer n, s.t.: $e^{-n(n-1)/730} \le 0.5$

 $\ln(e^{-n(n-1)/730}) \le \ln(0.5) \to n(n-1) \ge -730\ln(0.5) \to n \ge 23$

Same as before!

Poisson Processes

- · Consider "rare" events that occur over time
 - Earthquakes, radioactive decay, hits to web server, etc.
 - Have time interval for events (1 year, 1 sec, whatever...)
 - Events arrive at rate: λ events per interval of time
- Split time interval into $n \rightarrow \infty$ sub-intervals
 - Assume at most one event per sub-interval
 - Event occurrences in sub-intervals are independent
 - With many sub-intervals, probability of event occurring in any given sub-interval is small
- N(t) = # events in original time interval ~ Poi(λ)

Web Server Load

- · Consider requests to a web server in 1 second
 - In past, server load averages 2 hits/second
 - X = # hits server receives in a second
 - What is P(X = 5)?
- Model
 - Assume server cannot acknowledge > 1 hit/msec.
 - 1 sec = 1000 msec. (= large n)
 - P(hit server in 1 msec) = 2/1000 (= small p)
 - X ~ Poi(λ = 2)

$$P(X=5) = e^{-2} \frac{2^5}{5!} \approx 0.0361$$

Geometric Random Variable

- X is <u>Geometric</u> Random Variable: X ~ Geo(p)
 - X is number of independent trials until first success
 - p is probability of success on each trial
 - X takes on values 1, 2, 3, ..., with probability:

$$P(X = n) = (1-p)^{n-1} p$$

- E[X] = 1/p $Var(X) = (1 p)/p^2$
- Examples:
 - Flipping a fair (p = 0.5) coin until first "heads" appears.
 - Urn with N black and M white balls. Draw balls (with replacement, p = N/(N + M)) until draw first black ball.
 - Generate bits with P(bit = 1) = p until first 1 generated

Negative Binomial Random Variable

- X is <u>Negative Binomial</u> RV: X ~ NegBin(r, p)
 - X is number of independent trials until *r* successes
 - p is probability of success on each trial
 - X takes on values r, r + 1, r + 2..., with probability:

$$P(X = n) = {n-1 \choose r-1} p^r (1-p)^{n-r}$$
, where $n = r, r+1,...$

- E[X] = r/p $Var(X) = r(1 p)/p^2$
- Note: Geo(p) ~ NegBin(1, p)
- · Examples:
 - # of coin flips until r-th "heads" appears
 - # of strings to hash into table until bucket 1 has r entries

Hypergeometric Random Variable

- X is <u>Hypergeometric</u> RV: X ~ HypG(n, N, m)
 - Urn with N balls: (N-m) black and m white.
 - Draw n balls without replacement
 - X is number of white balls drawn

$$P(X = i) = \frac{\binom{m}{i} \binom{N - m}{n - i}}{\binom{N}{n}}, \text{ where } i = 0, 1, ..., n$$

- E[X] = n(m/N) $Var(X) = [nm(N-n)(N-m)]/[N^2(N-1)]$
- Let p = m/N (probability of drawing white on 1st draw)
- Note: $HypG(n, N, m) \rightarrow Bin(n, m/N)$
 - As $N \rightarrow \infty$ and m/N remains constant

Endangered Species

- Determine *N* = how many of some species remain
 - Randomly tag m of species (e.g., with white paint)
 - Allow animals to mix randomly (assuming no breeding)
 - Later randomly observe another *n* of the species
 - X = number of tagged animals in observed group of n
 - X ~ HypG(n, N, m)
- "Maximum Likelihood" estimate
 - Set *N* to be value that maximizes: $P(X = i) = \frac{\binom{i}{N} n i}{\binom{N}{N}}$

for the value *i* of X that you observed $\Rightarrow \hat{N} = mn/i$

• Similar to assuming: i = E[X] = nm/N