

Chi-square Tests

1. Chi-square Test of Goodness of Fit
2. Chi-square Test of Independence

Do all six colors occur in equal proportion?



1. Chi-square Test of Goodness of Fit

Chi-Square goodness of fit test is used to compare the observed sample distribution with the expected probability distribution.

The chi-square test statistic is of the form

$$\chi^2 = \sum_{i=1}^n \frac{(O_i - E_i)^2}{E_i}$$

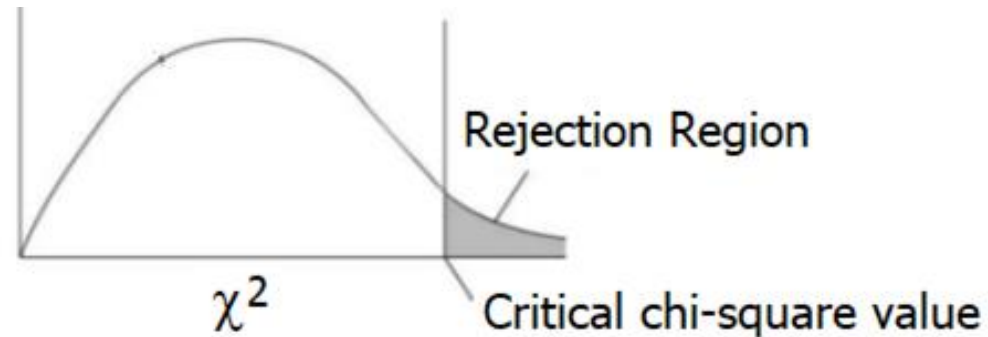
O_i : the observed frequency of the type i .

E_i : the expected frequency for the type i .

Null hypothesis: There is no significant difference between the observed and the expected value.

Alternative hypothesis: There is a significant difference between the observed and the expected value.

The Chi-Square distribution is used in the chi-square tests for goodness of fit.

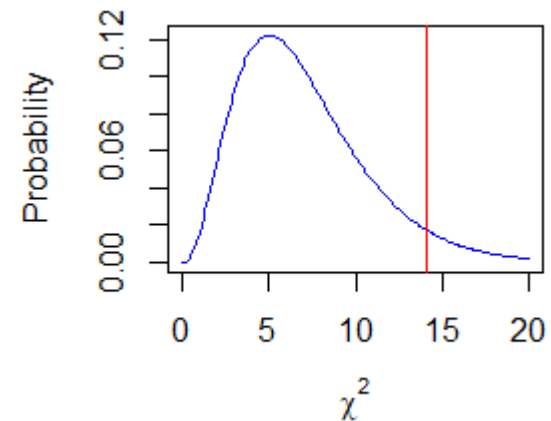


Problem

Find the 95th percentile of the Chi-Squared distribution with 7 degrees of freedom.

Solution

```
> #Critical chi-square
> a=0.05; (qc=qchisq(1-a,df=7))
[1] 14.06714
> #Chi-square distribution
> curve(dchisq(x,df=7),0,20,200,
+       col=4,ylab="Probability",
+       xlab=expression(chi^2))
> abline(v=qc,col=2)
```



Answer

The 95th percentile of the Chi-Squared distribution with 7 degrees of freedom is 14.067.

Students enrolled in an introductory Statistics course at the University of Auckland were asked to complete an online questionnaire. One of the questions asked them to enter their ethnicity. The 727 responses are displayed on the one-way table below.

[Reference] <https://nzmaths.co.nz/category/glossary/one-way-table>

Ethnicity	Chinese	Indian	Korean	Maori	NZ European	Other European	Pacific	Other	Total
Frequency	169	58	56	18	253	45	38	90	727

```
> o = c(169,58,56,18,253,45,38,90)
> tc = c("Frequency")
> tr = c("Chinese","Indian","Korean","Maori","NZ European",
+       "Other European","Pacific","Other")
> mo = matrix(o, dimnames=list(tr,tc))
> as.table(mo)
```

	Frequency
Chinese	169
Indian	58
Korean	56
Maori	18
NZ European	253
Other European	45
Pacific	38
Other	90



chisq.test(x, ...)

This performs chi-squared contingency table tests and goodness-of-fit tests.

Hypotheses

H_0 : Students enrolled in an introductory Statistics course are equally divided in their ethnicity.

H_a : Students enrolled in an introductory Statistics course are not equally divided in their ethnicity.

Degree of freedom: $df = k-1$, k =number of categories

Expected frequency: $E=n/k$, n =total frequency

```
> chisq.test(mo) #chisq.test
      Chi-squared test for given probabilities
data:  mo
X-squared = 494.05, df = 7, p-value < 2.2e-16
```

$P\text{-value} < 0.05$

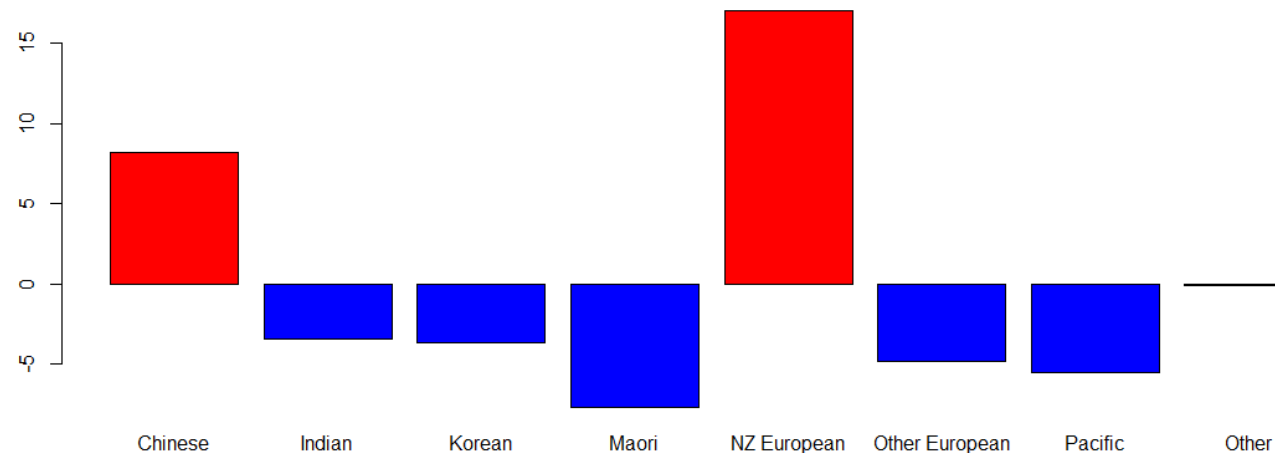
We reject the null hypothesis.

Students are not equally divided in their ethnicity.

• Hanging Chi-Gram

One way to visualize the discrepancies from the null hypothesis is to display them with a hanging chi-gram. This plots category i with a bar of height of the standardized residuals $\frac{O_i - E_i}{\sqrt{E_i}}$

```
> prop.table(mo)
      Frequency
Chinese      0.23246217
Indian       0.07977992
Korean       0.07702889
Maori        0.02475928
NZ European  0.34800550
Other European 0.06189821
Pacific      0.05226960
Other        0.12379642
> n = sum(O); k = 8; E = rep(n/k,k)
> cgram <- (O-E)/sqrt(E)
> barplot(cgram, col=ifelse(cgram>0,"red","blue"),
+         names.arg=tr)
```



We note that the “NZ European” and “Chinese” were greater than expected ($727/8=90.875$). However, the rest 6 ethnicities (“Indian”, ..., “Other”) were fewer than expected.

2. Chi-square Test of Independence

The Chi-Square Test of Independence is commonly used to test the statistical independence between two or more categorical variables. → **Cross-Tabulation Analysis** (교차분석)

Test Statistic

$$\chi^2 = \sum_{i=1}^R \sum_{j=1}^C \frac{(o_{ij} - e_{ij})^2}{e_{ij}}$$

o_{ij} is the observed cell count in the i^{th} row and j^{th} column of the table

e_{ij} is the expected cell count in the i^{th} row and j^{th} column of the table

$$e_{ij} = \frac{\text{row } i \text{ total} * \text{col } j \text{ total}}{\text{grand total}}$$

The quantity $(o_{ij} - e_{ij})$ is sometimes referred to as the *residual* of cell (i, j)

Contingency tables, also known as **two-way tables** or **cross tabulations** are a convenient way to display the frequency distribution from the observations of two categorical variables.

O_{ij} to denote the number of occurrences for which an individual falls into both category A_i and category B_j .

	B_1	B_2	\dots	B_c	total
A_1	O_{11}	O_{12}	\dots	O_{1c}	$O_{1.}$
A_2	O_{21}	O_{22}	\dots	O_{2c}	$O_{2.}$
\vdots	\vdots	\vdots	\ddots	\vdots	\vdots
A_r	O_{r1}	O_{r2}	\dots	O_{rc}	$O_{r.}$
total	$O_{.1}$	$O_{.2}$	\dots	$O_{.c}$	n

$$\chi^2 \text{ statistics} \approx \sum_{i=1}^r \sum_{j=1}^c \frac{(O_{ij} - E_{ij})^2}{E_{ij}} \quad E_{ij} = \frac{O_{i.} O_{.j}}{n}$$

degrees of freedom is $(r - 1) \times (c - 1)$

Example 1: City of residence with favorite baseball team

A cross-tabulation table comparing the two hypothetical variables, Residence City and Favorite Baseball Team, is shown below. Are residence city and being a fan of that city independent?

Residence City	Favorite Baseball Team		
	Toronto Blue Jays	Boston Red Socks	New York Yankees
Boston, MA	11	33	7
Montreal, Canada	23	14	9
Montpellier, VT	22	13	14

```
> dt1 <- array(c(11,23,22, 33,14,13, 7,9,14), dim=c(3,3),
+   dimnames=list("Residence City"=c("Boston","Montreal","Montpellier"),
+     "Favorite Baseball Team"=c("Blue Jays","Red Socks","Yankees")))
> (dt1 <- as.table(dt1))
```

```

      Favorite Baseball Team
Residence City Blue Jays Red Socks Yankees
Boston          11       33       7
Montreal        23       14       9
Montpellier      22       13      14
```

Manual calculation

Given the initial table, we can calculate the expected values.

$$E_{ij} = \frac{O_{i.} \cdot O_{.j}}{n}$$

Residence City	Favorite Baseball Team			Row Total
	Blue Jays	Red Socks	Yankees	
Boston	11	33	7	51
	$56 \cdot 51 / 146$	$60 \cdot 51 / 146$	$30 \cdot 51 / 146$	34.932%
Montreal	23	14	9	46
	$56 \cdot 46 / 146$	$60 \cdot 46 / 146$	$30 \cdot 46 / 146$	31.507%
Montpellier	22	13	14	49
	$56 \cdot 49 / 146$	$60 \cdot 49 / 146$	$30 \cdot 49 / 146$	33.562%
Column Total	56	60	30	146

Manual calculation

Now we can calculate the chi-square statistic.

$$\chi^2 \text{ statistics} \approx \sum_{i=1}^r \sum_{j=1}^c \frac{(O_{ij} - E_{ij})^2}{E_{ij}} = 19.35141$$

Residence City	Favorite Baseball Team			Row Total
	Blue Jays	Red Socks	Yankees	
Boston	11	33	7	51
	19.562	20.959	10.479	
	(11-19.562)^2/19.562	34.932%
Montreal	23	14	9	46
	17.644	18.904	9.452	
	(23-17.644)^2/17.644	31.507%
Montpellier	22	13	14	49
	18.795	20.137	10.068	
	(22-18.795)^2/18.795	33.562%
Column Total	56	60	30	146

Cross-tabulation analysis with gmodels package

```
> library(gmodels)
> CrossTable(dt1, prop.c=FALSE, prop.chisq=FALSE, prop.t=FALSE,
+           expected=TRUE, format="SPSS")
```

Residence City	Favorite Baseball Team			Row Total
	Blue Jays	Red Socks	Yankees	
Boston	11	33	7	51
Expected Values	19.562	20.959	10.479	
Row Percent	21.569%	64.706%	13.725%	34.932%
Montreal	23	14	9	46
	17.644	18.904	9.452	
	50.000%	30.435%	19.565%	31.507%
Montpellier	22	13	14	49
	18.795	20.137	10.068	
	44.898%	26.531%	28.571%	33.562%
Column Total	56	60	30	146

Pearson's Chi-squared test

chi^2 = 19.35141 d.f. = 4 p = 0.0006703343

$$df = (3-1) * (3-1)$$

The cells “Red Socks and Boston”, “Blue Jays and Montreal”, and “Blue Jays and Montpellier” were the three cells where the number of observed respondents were apparently greater than expected.

“Red Socks and Boston” are the most observed fan and city relationship.

```
#Balloon plot
library(gplots)
balloonplot(t(dt1), label=TRUE, show.margins=FALSE,
  main="Balloon Plot for Residence City by Baseball Team")
```

Balloon Plot for Residence City by Baseball Team

Favorite Baseball Team		Blue Jays	Red Socks	Yankees
Residence City	Boston	11	33	7
	Montreal	23	14	9
	Montpellier	22	13	14

■ `chisq.test()`

```
> #Cross-tabulation analysis by chisq.test
> (ct1 <- chisq.test(dt1))

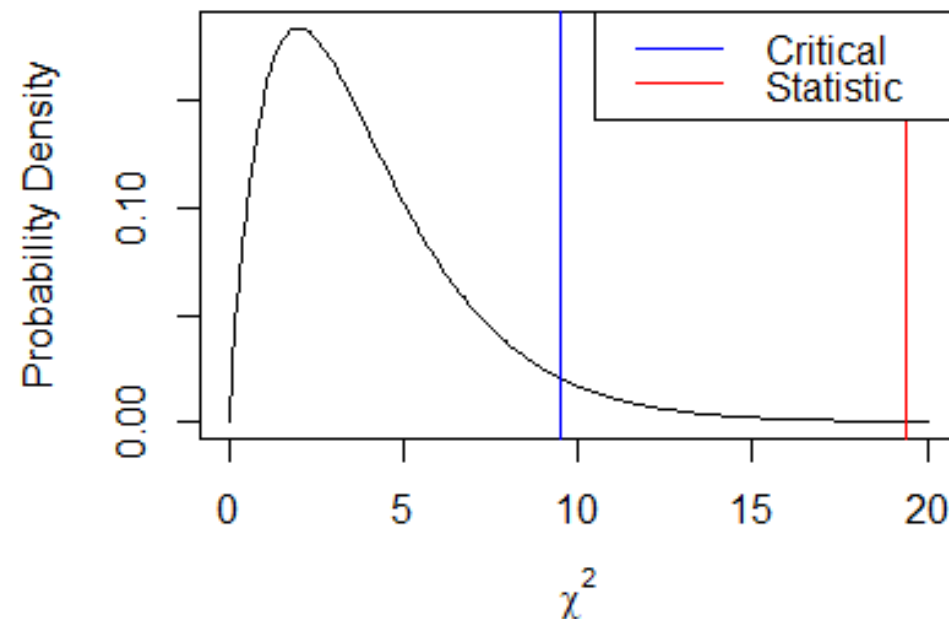
      Pearson's Chi-squared test

data:  dt1
X-squared = 19.351, df = 4, p-value = 0.0006703
```

```
curve(dchisq(x,df=4),0,20,200,xlab= expression(chi^2),
      ylab="Probability Density")
qc <- qchisq(1-0.05,df=4)
abline(v=qc,col=4) #critical chi-square
abline(v=ct1$statistic,col=2) #statistic chi-square
legend("topright",c("Critical","Statistic"),lty=1,
      col=c(4,2))
```

The chi-square value for the table is 19.35, and has an associated probability($p \leq 0.001$) of occurring by chance less than one time in 1000.

We reject the null hypothesis of independence. There is a strong relationship between the “Residence City” and “Favorite Baseball Team” variables.



Example 2: Consumption trend of Y and K university students

Ref: M. H. Huh, Introduction to Statistical Surveys, 3rd ed. (Free Academy, Seoul, 2011) pp.55-56.

```
> # Read data
> dt2 <- read.csv("SurveyData.csv",header=T)
> head(dt2[,1:6],4)
  univ id c1 c2 c3 c4
1    1  1  2  4  4  2
2    1  2  1  2  2  2
3    1  3  2  3  4  3
4    1  4  3  5  5  3
```

```
> #univ : 1 = Y Univ, 2 = K Univ
> University <- factor(dt2$univ,levels=1:2,labels=c("Y","K"))
> #c4 : I accept quickly a new fashion. (negative 1-5 positive)
> FashionAcceptance <- factor(dt2$c4)
> tb2 <- table(University,FashionAcceptance)
> tb2
```

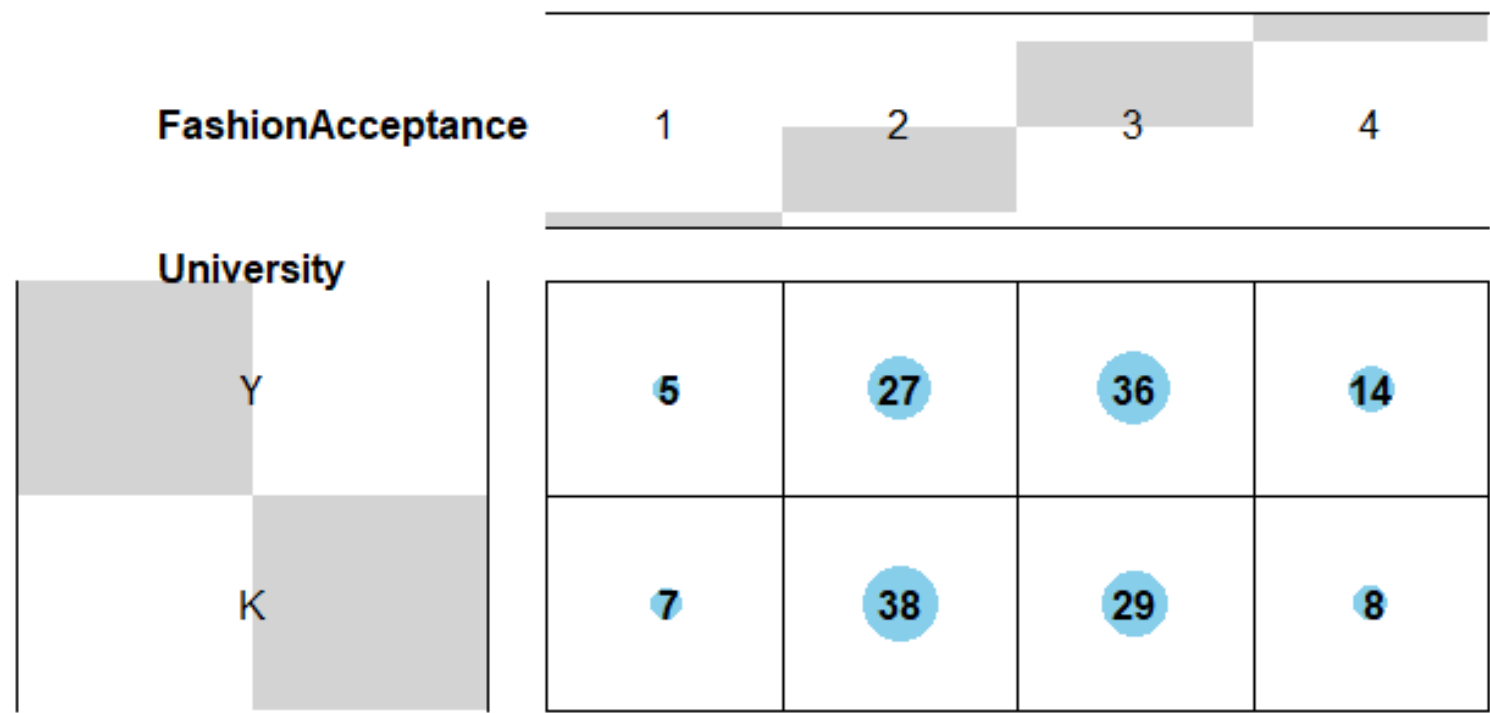
	FashionAcceptance			
University	1	2	3	4
Y	5	27	36	14
K	7	38	29	8

```
> addmargins(tb2)
```

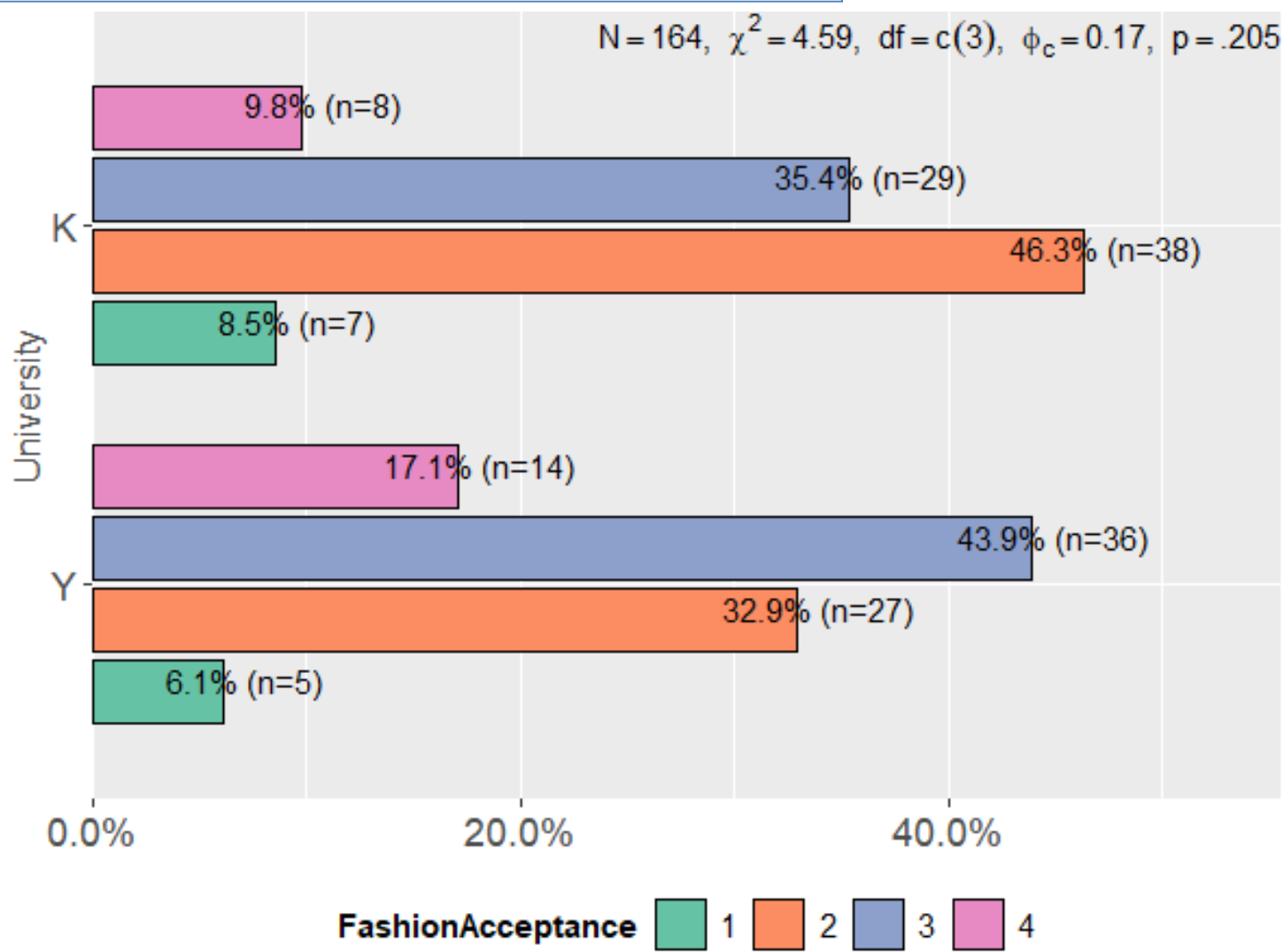
	FashionAcceptance				
University	1	2	3	4	Sum
Y	5	27	36	14	82
K	7	38	29	8	82
Sum	12	65	65	22	164


```
#Balloon plot
library(gplots)
balloonplot(t(tb2), label=TRUE, show.margins=FALSE,
  main="Balloon Plot for Two Universities by FashionAcceptance")
```

Balloon Plot for Two Universities by FashionAcceptance



```
library(sjPlot)
set_theme(geom.label.size=4,axis.textsize=1.1,
          legend.pos="bottom")
sjp.xtab(University,FashionAcceptance,type="bar",y.offset=0.01,
         margin="row",coord.flip=TRUE,wrap.labels=7,
         geom.colors="Set2",show.summary=TRUE)
```



▪ Cross-tabulation analysis with sjPlot package

```
sjt.xtab(University, FashionAcceptance,
  show.col.prc=TRUE, show.row.prc=TRUE)
```

University	FashionAcceptance				Total
	1	2	3	4	
Y	5	27	36	14	82
	6.1 %	32.9 %	43.9 %	17.1 %	100 %
	41.7 %	41.5 %	55.4 %	63.6 %	50 %
K	7	38	29	8	82
	8.5 %	46.3 %	35.4 %	9.8 %	100 %
	58.3 %	58.5 %	44.6 %	36.4 %	50 %
Total	12	65	65	22	164
	7.3 %	39.6 %	39.6 %	13.4 %	100 %
	100 %	100 %	100 %	100 %	100 %

$\chi^2=4.585 \cdot df=3 \cdot \text{Cramer's } V=0.167 \cdot p=0.205$

We observe that the chi-square value for the table is 4.585, and has an associated probability of 0.205.

We retain **the null hypothesis of no difference in the fashion acceptance** between Y and K university students.

Example 3: Boy Scouts and Juvenile Delinquency

This lesson spells out analysis techniques for a **three-way table**.

Boys Scouts and Juvenile Delinquency

Socioeconomic status	Boy scout	Delinquent	
		No	Yes
Low	No	169	42
	Yes	43	11
Medium	No	132	20
	Yes	104	14
High	No	59	2
	Yes	196	8

```
> bs <- read.csv("BoyScout.csv",header=TRUE)
> bs
```

	Socio	Scout	Delinquent	Frequency
1	Low	No	No	169
2	Low	No	Yes	42
3	Low	Yes	No	43
4	Low	Yes	Yes	11
5	Medium	No	No	132
6	Medium	No	Yes	20
7	Medium	Yes	No	104
8	Medium	Yes	Yes	14
9	High	No	No	59
10	High	No	Yes	2
11	High	Yes	No	196
12	High	Yes	Yes	8

Let's think of juvenile delinquency (D) as a response variable. Boy scout status (B) and socioeconomic status (S) are as predictors.

Null hypothesis: D is independent of B and S.

Alternative hypothesis: D is not independent of B and S.

```
> str(bs)
```

```
'data.frame': 12 obs. of 4 variables:
```

```
$ Socio      : Factor w/ 3 levels "High","Low","Medium": 2 2  
$ Scout      : Factor w/ 2 levels "No","Yes": 1 1 2 2 1 1 2 2  
$ Delinquent: Factor w/ 2 levels "No","Yes": 1 2 1 2 1 2 1 2  
$ Frequency  : int 169 42 43 11 132 20 104 14 59 2 ...
```

```
> bs$Socio <- ordered(bs$Socio,  
+                     levels=c("Low","Medium","High"))
```

```
> str(bs)
```

```
'data.frame': 12 obs. of 4 variables:
```

```
$ Socio      : Ord.factor w/ 3 levels "Low"<"Medium"<...: 1 1  
$ Scout      : Factor w/ 2 levels "No","Yes": 1 1 2 2 1 1 2 2  
$ Delinquent: Factor w/ 2 levels "No","Yes": 1 2 1 2 1 2 1 2  
$ Frequency  : int 169 42 43 11 132 20 104 14 59 2 ...
```

```
> #data.frame -> three-way table
```

```
> bs3 <- xtabs(Frequency~Socio+Scout+Delinquent,data=bs)
```

```
> bs3                                     #xtabs creates the contingency table
```

```
, , Delinquent = No
```

	Scout	
Socio	No	Yes
Low	169	43
Medium	132	104
High	59	196

```
, , Delinquent = Yes
```

	Scout	
Socio	No	Yes
Low	42	11
Medium	20	14
High	2	8

```
> ft3 <- ftable(bs3) #ftable prints out the "flat" version of the contingency table
> ft3
      Delinquent  No  Yes
Socio Scout
Low   No        169  42
      Yes        43  11
Medium No       132  20
      Yes       104  14
High  No         59   2
      Yes       196   8

> prop.table(ft3,1) #prop.table calculates the marginal proportions
      Delinquent      No      Yes
Socio Scout
Low   No      0.80094787 0.19905213
      Yes      0.79629630 0.20370370
Medium No      0.86842105 0.13157895
      Yes      0.88135593 0.11864407
High  No      0.96721311 0.03278689
      Yes      0.96078431 0.03921569

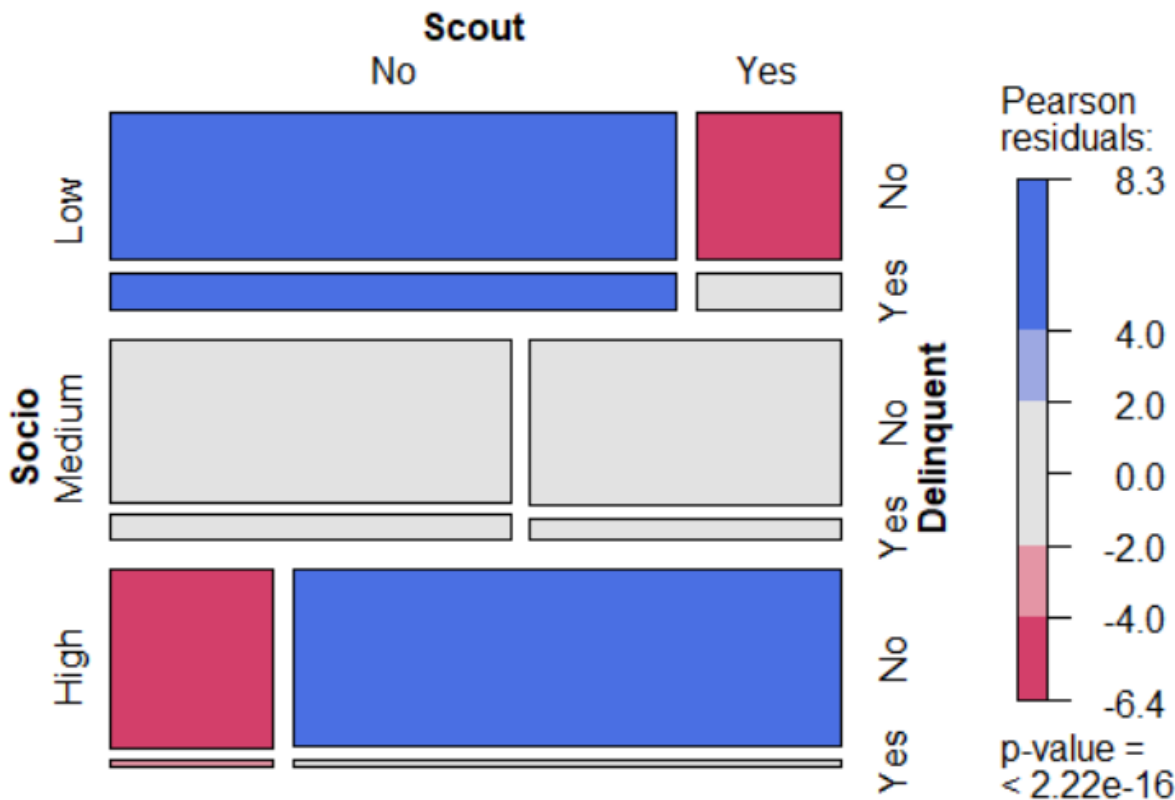
> chisq.test(ft3)
      Pearson's Chi-squared test
data:  ft3
X-squared = 32.958, df = 5, p-value = 3.837e-06
```

• Chi-square statistic ($\chi^2=32.958$) and the p -value($3.837e-06$) <0.05 .
The “Boy scout” and “Socioeconomic status” predictors are not independent of “Delinquent”. The null hypothesis does not hold, thus we reject this model of joint independence.

```
mosaic(x, shade, legend, ... ) {vcd}
Plots (extended) mosaic displays.

#mosaic plot
library(grid); library(vcd)
mosaic(bs3,shade=TRUE,legend=TRUE)
```

Pearson residual $r_{ij} = \frac{o_{ij} - e_{ij}}{\sqrt{e_{ij}}}$



The colors represent the level of the residual for that cell of levels. Blue means there are more observations in that cell than would be expected under the null model (independence). Red means there are fewer observations than would have been expected. You can read this as showing you which cells are contributing to the significance of the chi-squared test result.

The mosaic plot is based on [conditional probabilities](#). The heights and widths of the cells are proportional to the percentages of Socio and Scout categories against Delinquent categories.