

2022-2023

$$\text{Exo 1} \quad \mathcal{X} = L^\infty(\Omega, \mathbb{P}, \mathbb{P}) \quad X < Y \iff \mathbb{E}[f(X)] \leq \mathbb{E}[f(Y)]$$

$\forall f \nearrow$  et concave.

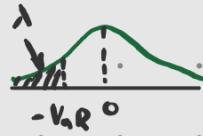
$$(1) \text{ Vérifier } X, Z \in \mathcal{X} \rightarrow X < \mathbb{E}(X|Z) \text{ (par sens)}.$$

$$(2) \quad X \in \mathcal{X}, \lambda \in [0, 1] \quad \text{AVaR}_\lambda(X) = \frac{1}{\lambda} \mathbb{E}[(q-X)^+] - q \quad \text{où } q \text{ est un quantile d'ordre } \lambda. \quad \text{d.l.q. } \text{AVaR}_\lambda(X) = \frac{1}{\lambda} \inf_{r \geq q} (\mathbb{E}[(r-X)^+] - \lambda r)$$

On peut supposer que  $X$  admet une densité  $p(x)$

$$(\rightarrow q^+ = q^- = q(r))$$

$$\begin{aligned} \text{VaR}_\lambda(X) &= \inf \{m : \mathbb{P}(X + m < 0) \leq \lambda\} \\ \text{AVaR}_\lambda(X) &= \frac{1}{\lambda} \int_0^\lambda \text{VaR}_z(X) dz \end{aligned}$$



$$\text{dVaR}_\lambda(X) := \frac{1}{\lambda} \int_0^\lambda \text{VaR}_z(X) dz = \frac{1}{\lambda} \int_0^\lambda q^+(z) dz = \frac{1}{\lambda} \underbrace{\mathbb{E}[-q^+(U)]}_{\lambda = \mathbb{P}(U < \lambda)} \underbrace{| U < \lambda}_{X < q} =$$

$$= \frac{1}{\lambda} \int_0^q -x p(x) dx = \frac{1}{\lambda} \int_0^q (q-x) p(x) dx - \frac{q}{\lambda} \underbrace{\int_0^q p(x) dx}_{1} = \frac{1}{\lambda} \mathbb{E}[(q-X)^+] - q$$

$$S(r) = \underbrace{\frac{1}{\lambda} \mathbb{E}[(r-X)^+]}_{\mathbb{P}(X < r)} - r \quad \text{- fonction convexe}$$

$$S'(r) = \frac{1}{\lambda} \mathbb{E}[\mathbf{1}_{\{X < r\}}] - 1 = 0 \rightarrow r^* \quad \mathbb{P}(X < r^*) = \mathbb{P}(X \leq r^*) = \lambda \rightarrow$$

Pour  $r = \text{VaR}_\lambda$ ,  $S'(r) = 0 \rightarrow \square$

$$(3) \quad X \rightarrow Y \rightarrow \forall \lambda \in [0, 1] \quad \text{AVaR}_\lambda(X) \geq \text{AVaR}_\lambda(Y)$$

$x \mapsto (r-x)^+$  est convexe  $\rightarrow \forall r$

$$\frac{1}{\lambda} (\mathbb{E}[(r-X)^+] - \lambda r) \geq \frac{1}{\lambda} (\mathbb{E}[(r-Y)^+] - \lambda r) \stackrel{\inf}{\geq} \text{AVaR}_\lambda(Y)$$

$$AVaR_\lambda(Y) \in \inf \left\{ \frac{1}{\lambda} (\mathbb{E}(r-X)^+ - \lambda r) \right\} = AVaR_\lambda^*(X)$$

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(4)  $\rho$  convexe et invariante en loi

$$\begin{aligned} \rho \text{ convexe monétaire} \rightarrow \rho(x) &= \max \{ \mathbb{E}[-X] - L_{\min}(\Omega) \} \\ \Omega \in M_{1,f}(\mathbb{P}) \quad \xrightarrow{\Omega \in \text{Ell}_{1,f} = \{ \mu \text{ et } \sigma \geq 0 \text{ tel que } \mathbb{E}[\Omega] = 1 \}} \\ \text{pour } X \in L^\infty(\mathbb{P}) \quad L_{\min}(\Omega) &= \sup_{X \in \mathcal{A}_\rho \cap \{X : \rho(X) \leq 0\}} \mathbb{E}_\Omega[-X] \end{aligned}$$

$\rho$  convexe monétaire invariante en loi  $\rightarrow$

$$\rho(x) = \sup_{\mu \in M_1([0,1])} \{ \int AVaR_\lambda(X) \mu(d\lambda) - \beta_{\min}(\mu) \}$$

$$\text{où } \beta_{\min}(\mu) = \sup_{X \in \mathcal{A}_\rho \cap [0,1]} \int AVaR(X) \mu(d\lambda) = \sup_{X \in L^\infty([0,1])} \{ \int AVaR_\lambda(X) d\lambda - \rho(X) \}$$

En déduire  $X \preceq Y \rightarrow \rho(X) \geq \rho(Y)$

ne dépend pas de  $X$

$$\begin{aligned} \text{Par (3)} \quad \int_{[0,1]} AVaR_\lambda(X) \mu(d\lambda) - \beta_{\min}(\mu) &\geq \int_{[0,1]} AVaR_\lambda(Y) \mu(d\lambda) - \beta_{\min}(\mu) \\ \text{Ainsi} \quad \rho(X) &\leq \rho(Y) \end{aligned}$$

Maintenant,  $\rho$  convexe invariante en loi  $\rho(X) \geq VaR_{\lambda_0}(X)$   
normalisée

$$(5) \quad \varepsilon > 0 \quad A_\varepsilon = \{-X \geq VaR_{\lambda_0}(X) - \varepsilon\}, \quad Y_\varepsilon = \mathbb{E}[X | \mathbb{I}_{A_\varepsilon}] = \mathbb{I}_{A_\varepsilon^c} X + \mathbb{I}_{A_\varepsilon} \mathbb{E}[X | A_\varepsilon]$$

$$\text{u.q. } \mathbb{P}(Y_\varepsilon < \mathbb{E}[X | A_\varepsilon]) = 0 \text{ et } \mathbb{P}(Y_\varepsilon \leq \mathbb{E}[X | A_\varepsilon]) > \lambda_0$$

$$\begin{aligned} \mathbb{P}(Y_\varepsilon < \mathbb{E}[X | A_\varepsilon]) &= \mathbb{P}(A_\varepsilon^c, X < \frac{\mathbb{E}[X | A_\varepsilon]}{\mathbb{P}(A_\varepsilon)}) \leq \mathbb{P}(x) = 0 \\ \{-X < VaR_{\lambda_0}(X) - \varepsilon\} &\xrightarrow{\parallel} \end{aligned}$$

$$\text{Mais sur } A_\varepsilon \quad X < -VaR_{\lambda_0}(X) + \varepsilon \rightarrow \{X < \mathbb{E}[X | A_\varepsilon]\} \subset \{X \leq -VaR_{\lambda_0}(X) + \varepsilon\}$$

$$\mathbb{P}(Y_E = \mathbb{E}[X|A_E]) \geq \mathbb{P}(A_E) = \mathbb{P}(X \leq -\text{VaR}_{\lambda_0}(X) + \varepsilon) > \lambda_0 \quad \text{par déf. de VaR}$$

$$\left( \begin{array}{l} \text{VaR}_{\lambda_0}(x) = \inf \{m : \mathbb{P}(X+m \leq 0) \leq \lambda_0\} \\ \text{Si } \mathbb{P}(X \leq -\text{VaR}_{\lambda_0}(X) + \varepsilon) = \lambda_0 \rightarrow ?! \end{array} \right)$$

(c) En déduire que  $\rho(x) \geq -\mathbb{E}[X|A_E]$

$$\text{Par (5), } q_Y^+(\lambda_0) = \mathbb{E}[X|A_E] \Rightarrow \text{VaR}_{\lambda_0}(Y) = -\mathbb{E}[X|A_E]$$

$$\text{Par (1)} \quad \rho(x) \geq \rho(\mathbb{E}[X|A_E]) = \rho(Y_E) \geq \text{VaR}_{\lambda_0}(Y_E) = -\mathbb{E}[X|A_E]$$

(7)  $X \in \mathcal{X}$  t.q.  $\mathbb{P}(X=x)=0 \quad \forall x$  M.q.  $\rho(x) \geq \text{AVaR}_{\lambda_0}(x)$

$$\begin{aligned} \text{AVaR}_{\lambda_0}(x) &= \frac{1}{\lambda_0} \int_0^{\lambda_0} q_Y^+(z) dz = \mathbb{E}[-q_Y^+(U) | U < \lambda_0] = \mathbb{E}[X | X < -\text{VaR}_{\lambda_0}(X)] = \\ &= \frac{\mathbb{E}[-X | \{X < -\text{VaR}_{\lambda_0}(X)\}]}{\mathbb{P}(X < -\text{VaR}_{\lambda_0}(X))} = \frac{\lim_{\varepsilon \downarrow 0} \mathbb{E}[-X | A_\varepsilon]}{\lim_{\varepsilon \downarrow 0} \mathbb{P}(A_\varepsilon)} = \lim_{\varepsilon \downarrow 0} \mathbb{E}[-X | A_\varepsilon] \leq \rho(x) \end{aligned}$$

(8)  $X, V \in \mathcal{X}$ ,  $V \sim \mathcal{U}(0,1)$   $V \perp\!\!\!\perp X$  et  $\eta > 0$ . M.q.  $\mathbb{P}(X+\eta V = x) = 0$ .  $\forall x$

En déduire que  $\rho(x) \geq \text{AVaR}_{\lambda_0}(x)$

$$\mathbb{P}(X+\eta V = x) = \int \mathbb{P}(y + \eta V = x) dF_X(y) = \int \mathbb{P}\left(V = \frac{x-y}{\eta}\right) dF_X(y) = 0$$

Par (7),  $\rho(X+\eta V) \geq \text{AVaR}_{\lambda_0}(X+\eta V)$

$\downarrow \eta \downarrow 0$  les metric sont Lipschitz en  $\|\cdot\|_\infty$   
 $\|X+\eta V - X\|_\infty = \eta \rightarrow 0$

$\rho(x) \geq \text{AVaR}_{\lambda_0}(x)$

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(1) Axioms de mesure de risque monétaire normalisée  $\rho(0) = 0$

1)  $X \geq Y \rightarrow p(X) \leq p(Y)$  - *Monotonicité*

2)  $c$ -const.  $\rightarrow p(X+c) = p(X) - c$  - *Invariance par translation*

$p(0) = 0$  - *normalisée*

**Cohérent** = convexe + positivement homogène

$B$  est dit étoilé si  $\forall X \in B, \lambda \in [0,1], \lambda X \in B$



$p$  est étoilée si  $\forall X \in \mathcal{X}, \forall \lambda > 1, p(\lambda X) \geq \lambda p(X)$

(2) M.q.  $p$  est étoiléessi  $\forall Y \in \mathcal{X} \quad \forall \lambda \in [0,1], p(\lambda Y) \leq \lambda p(Y)$

En déduire que  $p$  convexe  $\rightarrow$  étoilée

$$\forall X \in \mathcal{X} \quad \forall \lambda > 1 \quad p(\lambda X) \geq \lambda p(X) \Leftrightarrow \forall Y \in \mathcal{X} \quad \forall \lambda > 1 \quad \frac{1}{\lambda} p(Y) \geq p\left(\frac{Y}{\lambda}\right) \Leftrightarrow$$

$$\Leftrightarrow \forall Y \in \mathcal{X} \quad \forall \lambda \in [0,1] \quad p(Y) \geq p(\lambda Y).$$

$p$  convexe  $\rightarrow p(\lambda X + (1-\lambda) \cdot 0) \leq \lambda p(X) \quad \forall \lambda \in [0,1] \Rightarrow$  étoilée

(3)  $p$  est étoiléessi  $\partial_p$  est étoilé

$$\partial_p = \{X \in \mathcal{X} : p(X) \leq 0\}$$

$$\Rightarrow X \in \partial_p \quad \forall \lambda \in [0,1] \quad p(\lambda X) \stackrel{(2)}{\leq} \lambda p(X) \leq 0 \Rightarrow \lambda X \in \partial_p$$

$$\Leftarrow \partial_p \text{ étoilé. } \forall Y \in \mathcal{X} \quad Y + p(Y) \in \partial_p \Rightarrow \lambda Y + \lambda p(Y) \in \partial_p \Rightarrow$$

$$\begin{aligned} &\rightarrow p(\lambda Y + \lambda p(Y)) \leq 0 \\ &\quad \parallel \\ &\quad p(\lambda Y) \leq \lambda p(Y) \end{aligned} \quad \left\{ \rightarrow p(\lambda Y) \leq \lambda p(Y) \right.$$

(4)  $(p_\theta)_{\theta \in \Theta}$  une famille de mesures de risque étoilées

Est-ce que  $\begin{cases} p_v(X) = \sup_{\theta \in \mathbb{H}} p_\theta(X) \\ p_u(X) = \inf_{\theta \in \mathbb{H}} p_\theta(X) \end{cases}$  sont des m.d.r. étoilées?

Ils sont des mesures de risque monétaire (on peut vérifier les propriétés)

$$\forall \lambda \in [0,1] \quad \downarrow p(X)$$

$$p_v(\lambda X) = \sup_{\theta \in \mathbb{H}} p_\theta(\lambda X) \leq \lambda p_v(X)$$

$$p_u(\lambda X) = \inf_{\theta \in \mathbb{H}} p_\theta(\lambda X)$$

$$p_\theta(\lambda X) \leq \lambda p_\theta(X)$$

$$\inf_{\theta \in \mathbb{H}} p_\theta(\lambda X) \leq \lambda p_\theta(X)$$

$$p_u(\lambda X) \leq \lambda p_v(X) \rightarrow \text{Qui pour les deux.}$$

(5) M.q.  $p$  étoilée + sous-additive  $\rightarrow$  cohérente

$$\text{sous-add.} \quad \text{étoilée}$$

$$1) \text{ Convexité: } \forall \lambda \in [0,1] \quad p(\lambda X + (1-\lambda)Y) \leq p(\lambda X) + p((1-\lambda)Y) \leq \lambda p(X) + (1-\lambda)p(Y)$$

2) Positivement homogène:

$$\text{sous-add}$$

$$p\left(\frac{\lambda}{2}X + \frac{\lambda}{2}X\right) \leq \lambda p(X) \quad \Rightarrow \quad p(\lambda X) = \lambda p(X) \quad \text{pour } \lambda > 1 \quad X \in \mathcal{X}$$

$$p(\lambda X) \geq \lambda p(X) \quad \Rightarrow \quad \frac{1}{\lambda}p(Y) = p\left(\frac{Y}{\lambda}\right) \quad \forall \lambda > 1 \quad Y \in \mathcal{X}$$

$\tilde{C}$  étoilée

$$\text{Si } \lambda > 1 \quad \Rightarrow \quad p(\mu Y) = \sum p(Y) \quad \forall \mu < 1 \quad \forall Y \in \mathcal{X}$$

$p$  étoilée. Pour  $Y \in \mathcal{X}$   $\mathcal{A}_Y = \left\{ \lambda(Y + p(Y)) + z : \lambda \in [0,1], z \in \mathcal{X}, z \geq 0 \right\}$

(6) M.q.  $\mathcal{A}_Y \subset \mathcal{A}_p$  convexe. En déduire  $\forall X \in \mathcal{X}, p(X) \leq \inf_{Y \in \mathcal{X}} p_{\mathcal{A}_Y}(X)$

$$\lambda \in [0,1], z \geq 0$$

$$\rho(\lambda(Y + \rho(Y)) + Z) \stackrel{\text{def}}{\geq} \rho(\lambda(Y + \rho(Y))) \leq \lambda \rho(Y + \rho(Y)) = \lambda (\rho(Y) - \rho(Y)) = 0$$

$$\rightarrow A_Y \subset A_{\rho}$$

Convexe?  $\lambda_1, \lambda_2 \in [0,1]$ ,  $Z_1, Z_2 \geq 0$ ,  $y \in S_{0,1}[$

Déf  $A$  est un ensemble de positions acceptables si:

- $A \neq \emptyset$  et  $\inf\{m \in \mathbb{R} : m \in A\} = 0$
- Si  $X \in A$ ,  $Y \geq X \rightarrow Y \in A$

$$\begin{aligned} & g(\lambda_1(Y + \rho(Y)) + Z_1) + (1-g)(\lambda_2(Y + \rho(Y)) + Z_2) = \\ & = (\underbrace{g\lambda_1 + (1-g)\lambda_2}_{= \tilde{\lambda} \in (0,1]})(Y + \rho(Y)) + \underbrace{\lambda_1 Z_1 + (1-\lambda_1)Z_2}_{= \tilde{Z} \geq 0} \in A_Y \end{aligned}$$

↳  $\rho_A(X)$  est une mesure de risque monétaire

On a vu dans le cours que  $\rho(X) = \rho_{A_\rho}(X) = \inf\{m : \rho(X+m) \in A_\rho\}$

Par définition de  $\rho_A$ , si  $A_1 \subset A_2 \rightarrow \rho_{A_2}(X) \leq \rho_{A_1}(X)$

Donc  $\rho(X) = \rho_{A_\rho}(X) \leq \rho_{A_Y}(X) \quad \forall Y \rightarrow \rho(X) \leq \inf_{Y \in \mathcal{X}} \rho_{A_Y}(X)$

(7) M.Q.  $\rho(X) = \rho_{A_\rho}(X)$  et en déduire  $\rho(X) = \inf_{Y \in \mathcal{X}} \rho_{A_Y}(X)$   
C'est à dire il faut prendre  $Y=X$  dans (6)

$$\rho(X) = \inf\{m : X+m \in A_\rho\}$$

$$\{m : X+m \in A_\rho\} = \{m : m \geq \rho(X)\} \rightarrow \inf\{m : m \geq \rho(X)\} = \rho(X)$$

" "

$$\{m : \rho(X)-m \leq 0\}$$

(8) Pour  $X \in \mathcal{X}$  on considère  $(X_n)$  t.q.  $X_n \geq X_{n+1}$  et  $X_n \rightarrow X$  p.s.

M.Q.  $\lim_n \rho_{A_Y}(X_n) = \rho_{A_Y}(X)$ . En déduire que

$$\rho_{A_Y}(X) = \sup_{\Theta \in \mathcal{M}_1(\mathbb{P})} \{E^\Theta[-X] - L^*(\Theta)\} \quad \text{où} \quad L^*(\Theta) = \sup_{X \in A_Y} E^\Theta[-X]$$

$$\rho_{A_Y}(X) \geq \rho_{A_Y}(X_{n+1}) \geq \rho_{A_Y}(X_n) \rightarrow \lim \rho_{A_Y}(X_n) = \rho_{A_Y}(X) \leq \rho(X)$$

$$\text{On a } p_{\lambda Y}(X_n) + X_n = \lambda_n(Y + p(Y)) + Z_n$$

$$\lambda_n \in (0,1) \Rightarrow \exists n_k: \lambda_{n_k} \rightarrow \lambda_\infty \in (0,1]$$

$$Z_{n_k} = X_{n_k} + p_{\lambda Y}(X_{n_k}) - \lambda_n(Y + p(Y))$$

VI  
O  
↓ p.s. ↓ ↓

$$Z_\infty = X + p_\infty - \lambda_\infty(Y + p(Y)) \geq 0 \Rightarrow p_{\lambda Y}(x) \leq p_\infty \Rightarrow$$

$$\Rightarrow p_{\lambda Y}(x) = p_\infty$$

$\diamond$  Donc  $p$  est continue par au dessus  $\Rightarrow p(x) = \sup_{\Phi \in \mathcal{M}_1(\mathbb{P})} \{E_\Phi[-X] - L_{\min}(\Phi)\}$  Thm de cours

Théorème  $p$  monétaire convexe. On pose  $L_{\min}(\Phi) = \sup_{K \in \mathcal{C}_p} E_\Phi^K[-X]$ ,  $\Phi \in \mathcal{M}_1(\mathbb{P})$

Les conditions suivantes sont équivalentes:

a)  $p$  est sc.i. pour la topologie faible \*  $\sigma(L^\infty, L^1)$

b)  $p(x) = \sup_{\Phi \in \mathcal{M}_1(\mathbb{P})} \{E_\Phi[-X] - L_{\min}(\Phi)\}$ ,  $x \in L^\infty$

c)  $p$  est continue par au dessus: si  $X_n \downarrow X$  ps. alors  $p(X_n) \uparrow p(X)$

d)  $p$  satisfait la propriété de Fatou:  $X_n \in L^\infty$  t.q.  $X_n \rightarrow X$  ps.,  $\exists M: \|X_n\| \leq M$

Alors  $p(X) \leq \liminf_{n \rightarrow \infty} p(X_n)$ .

Représentation de  $\text{AVaR}_\lambda(X) = \max_{\Phi \in \mathcal{L}_\lambda} E_\Phi^{\Phi}[-X]$  où

$$\mathcal{L}_\lambda = \left\{ \Phi \text{-proba: } \frac{d\Phi}{d\mathbb{P}} \leq \frac{1}{\lambda} \text{ ps.} \right\}$$

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objectif

$$p_\lambda(X) = \sup_{\Phi \in \mathcal{M}_1(\mathbb{P})} E_\Phi^{\Phi}[-X] = \text{AVaR}_\lambda(X)$$

$\frac{d\Phi}{d\mathbb{P}} \leq \frac{1}{\lambda}$

(1)  $\beta$  est cohérente ( $=$  mesure de risque + convexe + positivement homogène)

- Monotonie  $X \leq Y \rightarrow E_{\beta}[-Y] \leq E_{\beta}[-X] \forall \beta \in M_1(P) \rightarrow$   
 $\sup_{\beta} \beta(Y) \leq \sup_{\beta} \beta(X)$

- Invariance par transl.  $c \in \mathbb{R}$

$$\beta_{\lambda}(X+c) - \sup_{\beta} \{E_{\beta}[-X] - \lambda\} = \beta(X) - \lambda$$

- Positivement hom.  $\alpha > 0$

$$\beta_{\lambda}(\alpha X) = \sup_{\beta} \alpha E_{\beta}[-X] = \alpha \beta_{\lambda}(X)$$

- Convexité:  $\lambda \in [0, 1]$

$$\beta_{\lambda}(\lambda X + (1-\lambda)Y) = \sup_{\beta} \{ \lambda E_{\beta}[-X] + (1-\lambda)E_{\beta}[-Y] \} \leq \lambda \sup_{\beta} E_{\beta}[-X] + (1-\lambda) \sup_{\beta} E_{\beta}[-Y] = \lambda \beta(X) + (1-\lambda) \beta(Y)$$

(2) M.Q.  $\beta_{\lambda}(X) = \sup_{Z \in L^1(P)} E[-XZ]$  puis que  
 $0 \leq Z \leq \frac{1}{\lambda}$   
 $EZ = 1$

$$\beta_{\lambda}(X) = \sup_{\Psi: \mathbb{R} \rightarrow [0, \frac{1}{\lambda}] \text{ mes.}} E[-X\Psi(X)]$$

$$E[\Psi(X)] = 1$$

$$\sup_{Z \in L^1(P)} E[-XZ] = \sup_Z E[-X \underbrace{E[Z|X]}_{G(X)-\text{mesurable}}] = \sup_{\Psi} E[-X\Psi(X)]$$

$\rightarrow \exists \Psi \text{-mes.}: E[Z|X] - \Psi(X) \in [0, \frac{1}{\lambda}]$

$E[\Psi(X)] - E[Z] = 1$

Inverse:  $E[-X \underbrace{\Psi(X)}_{Z \in L^1(P)}]$

$0 \leq Z \leq \frac{1}{\lambda} \quad EZ = 1$

(3)  $q$ -quantile d'ordre  $\lambda$  de  $X$  ( $P(X < q) \leq \lambda$ ,  $P(X \leq q) \geq \lambda$ )

$$\Psi_0(x) = \frac{1}{\lambda} \mathbb{I}_{\{X < q\}} + \infty \mathbb{I}_{\{X = q\}}$$

$$\text{de } = \begin{cases} 1 - \frac{P(X < q)}{\lambda} & \text{si } P(X = q) > 0 \\ 0 & \text{sinon} \end{cases}$$

(a) Vérifier que  $\Psi_0(x) \in [0, 1]$   $\forall x \in \mathbb{R}$

Il suffit de montrer  $\infty \leq \frac{1}{\lambda}$  i.e.  $\lambda - P(X < q) \leq P(X = q)$

C'est vrai car  $P(X \leq q) \geq \lambda$

(b) Calculer  $E[\Psi_0(x)]$  et  $E[-X\Psi_0(x)]$  le cas  $\neq 0$

$$E[\Psi_0(x)] = \frac{1}{\lambda} P(X < q) + \frac{1 - \frac{P(X < q)}{\lambda}}{P(X = q)} P(X = q) = 1$$

$$E[-X\Psi_0(x)] = -\frac{1}{\lambda} E[X \mathbb{I}_{\{X < q\}}] - q \left(1 - \frac{P(X < q)}{\lambda}\right) =$$

$$= -q - E[\mathbb{I}_{\{X < q\}}(X - q)] \cdot \frac{1}{\lambda} = \frac{1}{\lambda} E[(q - X)^+] - q = AVaR_\lambda(x)$$

$$\frac{1}{\lambda} \int_0^\lambda (q - q^+(u)) du - q = -\frac{1}{\lambda} \int_0^\lambda q^+(u) du$$

$$AVaR_\lambda(x) = \frac{1}{\lambda} \int_0^\lambda V_a R_\lambda(x) du = \frac{1}{\lambda} \int_{x+q}^\lambda q^+(u) du = \frac{1}{\lambda} \int_0^\lambda (q - q^+(u)) du - q =$$

$$= \frac{1}{\lambda} E[(q - X)^+] - q$$

$$\uparrow q > q^+(u) \Leftrightarrow u < \lambda$$

(c) Vérifier que  $\forall \Psi: \mathbb{R} \rightarrow [0, \frac{1}{\lambda}]$  mes.  $\Psi(x) \leq \Psi_0(x) \Leftrightarrow x < q$  et

$\Psi(x) \geq \Psi_0(x)$  si  $x > q$ . En déduire que  $0 \leq E[-X(\Psi_0(x) - \Psi(x))] \leq \infty$

$$\mathbb{E}[-X(\Psi_0(x) - \Psi(x))] = \mathbb{E}\left[\underbrace{-X(\Psi_0(x) - \Psi(x))}_{\geq 0} \mathbb{I}_{\{X < q\}}\right] +$$

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$$+ \mathbb{E}\left[\underbrace{-X(\Psi_0(x) - \Psi(x))}_{\leq 0} \mathbb{I}_{\{X > q\}}\right] + \mathbb{E}[-X(\infty - \Psi(x)) \mathbb{I}_{\{X = q\}}] \geq$$

$$\geq -q \mathbb{E}[(\Psi_0(x) - \Psi(x)) \mathbb{I}_{\{X < q\}}] + q \mathbb{E}[(\Psi(x) - \Psi_0(x)) \mathbb{I}_{\{X > q\}}] - q \mathbb{E}[(\Psi_0(x) - \Psi(x)) \mathbb{I}_{\{X = q\}}]$$

$$= -q \mathbb{E}[\Psi_0(x) - \Psi(x)] = 0$$

$$(4) \text{ Par (2) et (3c), } p_\lambda(x) = \mathbb{E}[-X\Psi_0(x)] \stackrel{(3b)}{=} \mathcal{V}_0 R_\lambda(x)$$

$$\text{Construire } \Theta_0 \in \mathcal{M}_1(\mathbb{P}) \quad \frac{d\Theta_0}{d\mathbb{P}} \leq \frac{1}{\lambda} \quad \mathbb{E}_{\Theta_0}[-X] = p_\lambda(x)$$

$$\Psi_0(x) = \frac{1}{\lambda} \mathbb{I}_{\{X < q\}} + \infty \mathbb{I}_{\{X = q\}} = Z = \frac{d\Theta_0}{d\mathbb{P}}$$

$$d\Theta_0 = \left( \frac{1}{\lambda} \mathbb{I}_{\{X < q\}} + \infty \mathbb{I}_{\{X = q\}} \right) \cdot d\mathbb{P}$$

(4) (a) A partir de (1),

$$p_\lambda(x) = \sup_{\Theta \in \mathcal{M}_1(\mathbb{P}) : \frac{d\Theta}{d\mathbb{P}} \leq \frac{1}{\lambda}} \left\{ \sup_{Z : Z = \frac{d\Theta}{d\mathbb{P}}} \mathbb{E}[-ZX] \right\}$$

évident //

$$p_\lambda(x) \stackrel{(1)}{=} \sup_{\Theta} \mathbb{E}\left[-X \cdot \frac{d\Theta}{d\mathbb{P}}\right] \quad \textcircled{1}$$

Pour  $\Theta$  fixé.  $\forall Z \sim \frac{d\Theta}{d\mathbb{P}} \exists \tilde{\Theta} : \frac{d\tilde{\Theta}}{d\mathbb{P}} = Z \quad Z \in [0, \frac{1}{\lambda}], \mathbb{E}Z = \mathbb{E}\frac{d\Theta}{d\mathbb{P}} = 1 \rightarrow$

$$\rightarrow \tilde{\Theta} \in \mathcal{M}_1(\mathbb{P}) \rightarrow \textcircled{2}$$

$$(6) \quad p_\lambda(x) = \sup_{\Theta \in \mathcal{M}_1(\mathbb{P})} \left\{ \int V_0 R_u(x) \cdot \left(-q + \frac{d\Theta}{d\mathbb{P}}(u)\right) du \right\}$$

$$\frac{d\Theta}{d\mathbb{P}} \leq \frac{1}{\lambda}$$

$$\sup_{\substack{X \sim \tilde{\gamma} \\ Y \sim \gamma}} \mathbb{E}[XY] = \int_0^1 q_x^+(u) q_y^+(u) du \xrightarrow{\text{Lemme 1}} \sup_{\substack{X \sim \tilde{\gamma} \\ Y \sim \gamma}} \mathbb{E}[-2X] = \int_0^1 q_x^+(u) q_{-\frac{d\Phi}{dP}}^+(u) du \xrightarrow{\text{Théo Jalabert}} -V_\alpha R_u(X)$$

(c) Monotonie de  $u \in [0,1] \mapsto V_\alpha R_u(X) = -q_x^+(u)$

$$\text{et } u \mapsto -q_{-\frac{d\Phi}{dP}}^+$$

$$\text{et la valeur } \int_0^1 -q_{-\frac{d\Phi}{dP}}^+(u) du = -\mathbb{E}\left[q_{-\frac{d\Phi}{dP}}^+(Y)\right] = \mathbb{E}\left[\frac{d\Phi}{dP}\right] = 1$$

$$\text{M.g. pour } \Phi \in \mathcal{U}_1(P): \frac{d\Phi}{dP} \leq \frac{1}{\lambda} \quad \int_0^1 V_\alpha R_u(X) \left(-q_{-\frac{d\Phi}{dP}}^+(u)\right) du \leq \frac{1}{\lambda} \int_0^1 V_\alpha R_u(X) du$$

Construire le proba pour laquelle on a égalité ( $\rightarrow (1) = V_\alpha R_\lambda$ )

Pour  $y \in [0,1]$

$$\text{Si } -\frac{1}{\lambda} \leq Z \leq 0 \rightarrow q_Z^+(y) = \inf\{x : P(Z \leq x) \geq y\} \leq 0$$

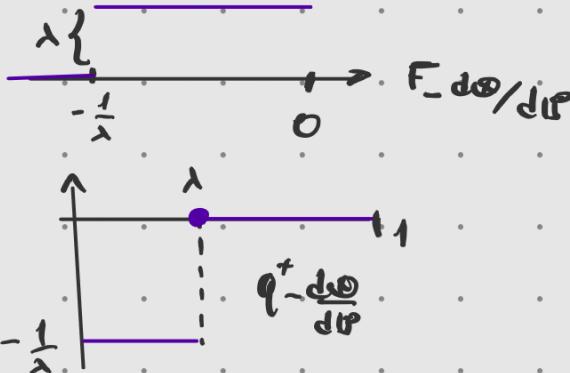
$$\text{et } q_Z^+(y) \geq -\frac{1}{\lambda} \text{ car } F(Z \leq -\frac{1}{\lambda} - \frac{1}{n}) = 0 \quad \forall n$$

$$\text{Donc } -q_{-\frac{d\Phi}{dP}}^+ \leq \frac{1}{\lambda} \rightarrow \text{inégalité}$$

Égalité si  $-q_{-\frac{d\Phi}{dP}}^+(u) = \begin{cases} \frac{u}{\lambda}, & \text{pour } u < \lambda \\ 0, & u \geq \lambda \end{cases} \rightarrow$

$$\frac{d\Phi}{dP} = \frac{1}{\lambda} \mathbf{1}_{u < \lambda}$$

$$\mathbb{E}\left[\frac{d\Phi}{dP}\right] = 1$$



2019-2020

$\beta_1, \beta_2$  mesures de risque monétaires t.q.  $\inf_{Z \in \mathcal{X}} \{\beta_1(-Z) + \beta_2(Z)\} \geq -\infty$

Inf-convolution  $\beta_1 \square \beta_2(X) = \inf_{Z \in \mathcal{X}} \{\beta_1(X-Z) + \beta_2(Z)\}$

(1) M.Q.  $\beta_1 \square \beta_2(X) > -\infty$  pour  $X \in \mathcal{X}$  puis que  $\beta_1 \square \beta_2$  est une mesure de risque monétaire. Vérifier que  $\beta_1 \square \beta_2 = \beta_2 \square \beta_1$ .

$\beta_1$  est  $\| \cdot \|_\infty$ -Lip.  $\rightarrow \beta_1(z) - \beta_1(x-z) \leq \|x\|_\infty \rightarrow \beta_1(x-z) \geq \beta_1(-z) - \|x\|_\infty$

$$\rightarrow \beta_1 \square \beta_2(x) \geq \inf_{z \in \mathcal{X}} \left\{ \beta_1(-z) + \beta_2(z) \right\} - \|x\|_\infty \geq -\infty$$

Montonic + cash invariance grâce à cette propriété de  $\beta(\cdot - z)$

$$\inf_{z \in \mathcal{X}} \left\{ \beta_1(x-z) + \beta_2(z) \right\} = \inf_{y \in \mathcal{X}} \left\{ \beta_1(y) + \beta_2(y-x) \right\} = \beta_2 \square \beta_1(x)$$

(2)  $\beta_1, \beta_2$  est convexe si  $\beta_1$  et  $\beta_2$  sont convexes

$$\begin{aligned} \beta_1 \square \beta_2(\lambda X + (1-\lambda)Y) &= \inf_z \left\{ \beta_1(\lambda X + (1-\lambda)Y - z) + \beta_2(z) \right\} \leq \\ &\leq \lambda \inf_z \left\{ \beta_1(X-z) + \beta_2(z) \right\} + (1-\lambda) \inf_z \left\{ \beta_2(Y-z) + \beta_2(z) \right\} \end{aligned}$$

$\forall \varepsilon > 0 \exists z', z'':$

$$\begin{aligned} &\lambda (\beta_1(X-z') + \beta_2(z')) \leq \inf_z \dots + \frac{\varepsilon}{2} \\ &+ (1-\lambda) (\beta_1(Y-z'') + \beta_2(z'')) \leq \inf_z \dots + \frac{\varepsilon}{2} \\ &\lambda \beta_1 \square \beta_2(X) + (1-\lambda) \beta_1 \square \beta_2(Y) \\ &\quad \inf_z \dots + \inf_z \dots + \varepsilon \geq (\lambda \beta_1(X-z') + (1-\lambda) \beta_1(Y-z'')) + (\lambda \beta_2(z') + (1-\lambda) \beta_2(z'')) \geq \end{aligned}$$

Jensen

$$\geq \beta_1(\lambda X + (1-\lambda)Y - \underbrace{(\lambda z' + (1-\lambda)z'')}_{z}) + \beta_2(\lambda z' + (1-\lambda)z'') \geq \beta_1 \square \beta_2(\lambda X + (1-\lambda)Y)$$

$$(3) (\beta_1 \square \beta_2)^*(Y) = \sup_{X \in \mathcal{X}} \{ \mathbb{E}(XY) - \beta_1 \square \beta_2(X) \} =$$

$$= \sup_{X \in \mathcal{X}} \left\{ \mathbb{E}(XY) - \inf_z \left\{ \beta_1(X-z) + \beta_2(z) \right\} \right\} = \sup_{X, z} \left\{ \mathbb{E}[\underbrace{\mathbb{E}[(X-z)Y]}_W + \mathbb{E}[ZY]] - \beta_1(X-z) \right\}$$

$$-\mathbb{P}_2(Z) = \sup_{W, Z} \{\mathbb{E}[WY] - \mathbb{P}_1(W) + \mathbb{E}[ZY] - \mathbb{P}_2(Z)\} = \mathbb{P}_1^*(Y) + \mathbb{P}_2^*(Y)$$

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(4) Entropie:  $\text{Ent}_s(X) = \frac{1}{s} \log(\mathbb{E}[e^{-sX}])$ ,  $X \in \mathcal{X}$

M.Q.  $\text{Ent}_s(X)$  est monétaire, convexe, invariant en loi

- $X \leq Y \rightarrow \mathbb{E}[e^{-sX}] \geq \mathbb{E}[e^{-sY}]$

$$\text{Ent}_s(X) \geq \text{Ent}_s(Y)$$

$\mathbb{E}[e^{-sX}]$  - transformée de Laplace qui caractérise la loi

→ s.c.i pour la topologie faible \*

- $\frac{d}{ds} \log(\mathbb{E}[e^{-s(X+c)}]) = \frac{1}{s} \log(\mathbb{E}[e^{-sc} e^{-sX}]) = -c + \text{Ent}_s(X)$

- Convexité:  $\text{Ent}_s(\lambda X + (1-\lambda)Y) - \frac{1}{s} \log(\mathbb{E}[e^{-s\lambda X} \cdot e^{-s(1-\lambda)Y}]) \leq$

$$\leq \left\{ \int f g dx \leq (\int f^p dx)^{1/p} (\int g^q dx)^{1/q}, \text{ Hölder avec } p = \frac{1}{\lambda}, q = \frac{1}{1-\lambda} \right\} \leq$$

$$\leq \frac{1}{s} \log \mathbb{E}[e^{-sX}]^\lambda + \frac{1}{s} \log \mathbb{E}[e^{-sY}]^{1-\lambda} = \lambda \text{Ent}_s(X) + (1-\lambda) \text{Ent}_s(Y)$$

(5)  $\text{Ent}_s(X) = \sup_{\Phi \in M_1(\mathbb{R})} \{\mathbb{E}_{\Phi}[-X] - \mathbb{L}_s(\Phi)\}$ .  $\mathbb{L}_s \Leftrightarrow \text{Ent}_s^*$

grâce à s.c.i pour la top. faible \*.

$$\mathbb{L}_s(\Phi) = \sup_{X \in \mathcal{X}} \{\mathbb{E}_{\Phi}[-X] - \text{Ent}_s(X)\} = \sup_{X \in \mathcal{X}} \{\mathbb{E}_{\Phi}[-X]\}$$

$$\sup_{X \in \mathcal{X}} \left\{ \mathbb{E}_{\Phi} \left[ X \cdot \left( -\frac{d\Phi}{dP} \right) \right] - \text{Ent}_s(X) \right\} = \text{Ent}_s^* \left( -\frac{d\Phi}{dP} \right)$$

$$\text{Ent}_s^*(y) = \begin{cases} +\infty & \text{si } Y \notin \{Y \in \mathbb{N}: Y \leq 0 \text{ et } Y = -1\} \\ \mathbb{L}_s(\Phi) & \text{sinon où } \frac{d\Phi}{dP} = -y \end{cases}$$

(5) On admet que  $\mathbb{L}_s(\Phi) = \frac{1}{s} H(\Phi | P) = \frac{1}{s} \mathbb{E} \left[ \frac{d\Phi}{dP} \log \left( \frac{d\Phi}{dP} \right) \right]$

$$\text{M.Q. pour } s, t > 0 \quad (\text{Ent}_s \square \text{Ent}_t)^{*+} = \text{Ent}_{\frac{st}{s+t}}$$

$$(\text{Ent}_s \square \text{Ent}_t)^{*+} = \text{Ent}_s^{*+} + \text{Ent}_t^{*+}$$

$$(\text{Ent}_s \square \text{Ent}_t)^{*+} = \sup_{Y \in L'} \{ \mathbb{E}[XY] - \text{Ent}_s^*(Y) - \text{Ent}_t^*(Y) \} = \\ Y \leq 0 \quad \mathbb{E}[Y] = 0$$

$$= \sup_{\Phi \in M_1(P)} \left\{ \mathbb{E}\left[-X \frac{d\Phi}{dP}\right] - \underbrace{\lambda_s(\Phi) + \lambda_t(\Phi)}_{-\frac{t+s}{st} H(\Phi||P)} \right\} = \sup_{\Phi \in M_1(P)} \left\{ \mathbb{E}\left[-X\right] - \frac{s}{s+t} \lambda_s(\Phi) \right\} = \text{Ent}_{\frac{st}{s+t}}^{(s)}$$

(7) Trouver  $\lambda \in [0, 1]$  tq.  $\text{Ent}_s(\lambda X) + \text{Ent}_t((1-\lambda)X) = \text{Ent}_{\frac{st}{s+t}}$   
 En déduire la valeur de  $\text{Ent}_s \square \text{Ent}_t$  et déterminer la séparation optimale.

$$\text{Ent}_s(\lambda X) + \text{Ent}_t((1-\lambda)X) = \frac{1}{s} \log(\mathbb{E}[e^{-s\lambda X}]) + \frac{1}{t} \log(\mathbb{E}[e^{-t(1-\lambda)X}]) = \\ = \log\left(\mathbb{E}[e^{-\lambda s X}]^{\frac{1}{s}} \mathbb{E}[e^{-(1-\lambda)t X}]^{\frac{1}{t}}\right) = \\ = \frac{t+s}{st} \log\left(\mathbb{E}[e^{-\lambda s X}]^{\frac{t}{t+s}} \mathbb{E}[e^{-(1-\lambda)t X}]^{\frac{s}{t+s}}\right) = \frac{t+s}{st} \log\left(\mathbb{E}[e^{-\frac{st}{t+s} X}]\right) = \text{Ent}_{\frac{st}{t+s}}(X)$$

$$\text{On choisit } \lambda \text{ tq. } \lambda s = (1-\lambda)t \rightarrow \lambda = \frac{t}{s+t}$$

$$\text{On sait que } \text{Ent}_t \square \text{Ent}_s \geq (\text{Ent}_t \square \text{Ent}_s)^{*+} = \text{Ent}_{\frac{s+t}{s+t}}(X) = \text{Ent}_s\left(\frac{t}{s+t} X\right) + \text{Ent}_t\left(\frac{s}{s+t} X\right)$$

$$\inf_Z \{ \text{Ent}_s(X-Z) + \text{Ent}_t(Z) \} \rightarrow Z^* = \frac{s}{s+t} X \rightsquigarrow \text{Ent}_t \\ X - Z^* = \frac{t}{s+t} X \rightsquigarrow \text{Ent}_s$$