

# Markov Chains - Summary

Lectured by Prof Geoffrey Grimmett

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**Definition 1** (Markov Chain). Let  $S$  be a set (the *state space*). Let  $\mathbf{X} = (X_n : n = 0, 1, 2, \dots)$  be a sequence of random variables taking values in  $S$ . Then  $X$  is a *Markov chain* if it satisfies the *Markov property*

$$\mathbb{P}(X_{n+1} = i_{n+1} \mid X_0 = i_0, X_1 = i_1, \dots, X_n = i_n) = \mathbb{P}(X_{n+1} = i_{n+1} \mid X_n = i_n) \\ \forall n \in \mathbb{N}, i_0, i_1, \dots, i_{n+1} \in S$$

**Definition 2** (Homogeneous). A Markov chain  $X$  is *homogeneous* if

$$\mathbb{P}(X_{n+1} = j \mid X_n = i)$$

does not depend on  $n$ .

**Definition 3** (Initial distribution). The *initial distribution* is  $\lambda = (\lambda_i : i \in S)$  where  $\lambda_i = \mathbb{P}(X_0 = i)$ .

**Definition 4** (Transition probabilities). The (1-step) *transition probabilities* are  $p_{i,j} = \mathbb{P}(X_{n+1} = j \mid X_n = i)$  and the *transition matrix* is  $P = (p_{i,j})$  for  $i, j \in S$ .

**Theorem 1.** Let  $\lambda$  be a distribution on a state space  $S$  and  $P$  a transition matrix. Then the sequence  $\mathbf{X} = (X_n : n = 0, 1, 2, \dots)$  is a Markov chain with initial distribution  $\lambda$  and transition matrix  $P$  if and only if

$$\mathbb{P}(X_0 = i_0, X_1 = i_1, \dots, X_n = i_n) = \lambda_{i_0} p_{i_0, i_1} p_{i_1, i_2} \dots p_{i_{n-1}, i_n} \quad \forall n \in \mathbb{N}, i_0, \dots, i_n \in S$$

**Theorem 2** (Extended Markov Property). Let  $X$  be a Markov chain. For  $n \geq 0$ , let  $H$  be an event given in terms of the history  $(X_0, X_1, \dots, X_{n-1})$  and let  $F$  be an event given in terms of the future  $(X_{n+1}, X_{n+2}, \dots)$ . Then

$$\mathbb{P}(F \mid X_n = i, H) = \mathbb{P}(F \mid X_n = i) \quad \forall n \in \mathbb{N}, i \in S$$

**Definition 5** ( $n$ -step transition probabilities). Let  $X$  be a Markov chain. Then the  $n$ -step transition probability is  $p_{i,j}(n) = \mathbb{P}(X_n = j \mid X_0 = i)$ .

**Theorem 3** (Chapman-Kolmogorov). Let  $X$  be a Markov chain over a state space  $S$ . Then

$$p_{i,j}(m+n) = \sum_{k \in S} p_{i,k}(m) p_{k,j}(n)$$

**Definition 6** (Communication). Let  $X$  be a homogeneous Markov chain over a state space  $S$ . Let  $i, j \in S$ .

We say  $i$  leads to  $j$  if  $p_{i,j}(n) > 0$  for some  $n \geq 0$  and write  $i \rightarrow j$ .

We say  $i$  and  $j$  communicate if  $i \rightarrow j$  and  $j \rightarrow i$  and write  $i \leftrightarrow j$ .

An equivalence class of  $(S, \leftrightarrow)$  is called a *communicative class*.

**Definition 7** (Reducibility). Let  $S$  be a state space with an associated Markov chain. The chain is *irreducible* if  $S$  is the single equivalence (communicative) class. It is *reducible* otherwise.

**Definition 8** (Closure and absorption). A subset  $C$  of a state space  $S$  is *closed* if  $i \in C, i \rightarrow j \implies j \in C$ .

If  $C = \{i\}$  is closed, the state  $i$  is called *absorbing*.

**Notation.** We write  $\mathbb{P}(\cdot \mid X_0 = i) = \mathbb{P}_i(\cdot)$  and  $\mathbb{E}(\cdot \mid X_0 = i) = \mathbb{E}_i(\cdot)$ .

**Definition 9** (First-passage time). The *first-passage time* to a state  $j$  is

$$T_j = \inf \{n \geq 1 : X_n = j\}.$$

**Definition 10** (First-passage probabilities). The *first-passage probabilities* are  $f_{i,j}(n) = \mathbb{P}_i(T_j = n)$ . We also write  $f_{i,j} = \sum_{n \geq 1} f_{i,j}(n)$ .

**Definition 11** (Recurrence and transience). A state  $i$  is called *recurrent* (or *persistent*) if  $f_{i,i} = \sum_{n \geq 1} f_{i,i}(n)$  satisfies  $f_{i,i} = 1$ . A non-recurrent state is called *transient*.

**Theorem 4.** A state  $i$  is recurrent if and only if  $\sum_n p_{i,i}(n) = \infty$ .

**Theorem 5.** Let  $C$  be a communicating class. Then

- (i) either every state in  $C$  is recurrent or every state in  $C$  is transient;
- (ii) if  $C$  contains some recurrent state then  $C$  is closed.

**Theorem 6.** Suppose the state space  $S$  is finite. Then

- (i) there exists at least one recurrent state.
- (ii) if the chain is irreducible, every state is recurrent.

**Theorem 7** (Polya). A random walk  $X$  on  $\mathbb{Z}^d$  is recurrent if  $d = 1$  or  $2$ , and transient if  $d \geq 3$ .

**Definition 12** (Hitting times and probabilities). The hitting time for an event  $A \subseteq S$  is

$$H^A = \inf \{n \geq 0 : X_n \in A\}.$$

(Note that if  $X_0 \in A$  then  $H^A = 0$ .)

The hitting probability for an event  $A$  starting in a state  $i$  is

$$h_i^A = \mathbb{P}_i(H^A < \infty)$$

i.e. the probability that the Markov chain ever reaches  $A$  given that it starts in state  $i$ .

The expected hitting time is

$$k_i^A = \mathbb{E}_i(H^A) \quad [= \infty \text{ if } h_i^A < 1].$$

**Theorem 8.** The vector  $h^A = (h_i^A : i \in S)$  of hitting probabilities is the *minimal* non-negative solution of the equations

$$h_i^A = \begin{cases} 1 & \text{if } i \in A \\ \sum_{j \in S} p_{ij} h_j^A & \text{if } i \notin A \end{cases}$$

where *minimal* means that if  $x = (x_i : i \in S)$  is any non-negative solution to the above then  $h_i^A \leq x_i$  for all  $i \in S$ .

**Definition 13** (Stopping time). A random variable  $T : \Omega \rightarrow \{0, 1, 2, \dots\} \cup \{\infty\}$  is a *stopping time* if for all  $n \geq 0$  the event  $\{T = n\}$  is given in terms of  $X_0, X_1, \dots, X_n$ .

**Theorem 9** (Strong Markov Property). Let  $X$  be a Markov chain with transition matrix  $P$  and let  $T$  be a stopping time for  $X$ . Given that  $T < \infty$  and that  $X_T = i$ , the sequence  $Y_k = X_{T+k}$  ( $k \geq 0$ ) is a Markov chain with transition matrix  $P$ . In addition, given  $T < \infty$ ,  $Y = (Y_k)$  is independent of the  $X_0, X, \dots, X_{T-1}$ .

**Definition 14** (Return visits). Suppose  $X_0 = i$ , then  $V = |\{n \geq 1 : X_n = i\}|$  is the number of return visits of  $X$  to state  $i$ .

Note that  $\mathbb{P}_i(V \geq 1) = f_{i,i}$  and so inductively by the Strong Markov Property,  $\mathbb{P}_i(V \geq k) = f_{i,i}^k$ .

**Definition 15** (Mean recurrence time). The *mean recurrence time* of a state  $i \in S$  is

$$\mu_i = \mathbb{E}_i(T_i) = \begin{cases} \infty & \text{if } i \text{ is transient} \\ \sum_{n=1}^{\infty} n f_{i,i}(n) & \text{if } i \text{ is recurrent} \end{cases}$$

**Definition 16** (Null and positive recurrence). Suppose state  $i$  is recurrent. Then  $i$  is *null* recurrent if  $\mu_i = \infty$ . Conversely,  $i$  is *positive* (or *non-null*) recurrent if  $\mu_i < \infty$ .

**Definition 17** (Period). The *period*  $d_i$  of a state  $i \in S$  is

$$d_i = \gcd \{n \geq 1 : p_{i,i}(n) > 0\}.$$

We say that  $i$  is *aperiodic* if  $d_i = 1$ .

**Definition 18** (Ergodic). A state  $i$  is called *ergodic* if it is positive, recurrent and aperiodic.

**Definition 19** (Invariant distribution). Let  $X$  be a Markov chain with transition matrix  $P$ . The vector  $\pi = (\pi_i : i \in S)$  is an *invariant distribution* if

- (i)  $\pi_i \geq 0$  for all  $i$
- (ii)  $\sum_i \pi_i = 1$
- (iii)  $\pi = \pi P$

**Theorem 10.** Let  $X$  be an *irreducible* Markov chain. Then

- (i) There exists an invariant distribution if some state is positive recurrent.
- (ii) If there exists an invariant distribution  $\pi$  then every state is positive recurrent, and  $\pi_i = \frac{1}{\mu_i}$  for each  $i \in S$ . In particular, there is a unique invariant distribution.

**Theorem 11.** Let  $X = (X_n)$  be a Markov chain which is irreducible, positive recurrent and aperiodic. Then  $p_{i,j}(n) \rightarrow \pi_j$  as  $n \rightarrow \infty$  where  $\pi = (\pi_j)$  is the unique invariant distribution.

**Theorem 12.** Denote  $V_i(n) = |\{1 \leq m \leq n : X_m = i\}|$  (number of visits to  $i$  up to time  $n$ ).

For an irreducible, positive recurrent Markov chain,  $\frac{V_i(n)}{n} \Rightarrow \pi_i$  as  $n \rightarrow \infty$  where  $Z_n \Rightarrow c$  as  $n \rightarrow \infty$  means ‘convergence in probability’,  $\mathbb{P}(|Z_n - c| > \epsilon) \rightarrow 0 \forall \epsilon > 0$ .

**Theorem 13.** Let  $X = (X_0, X_1, \dots, X_n)$  be an irreducible Markov chain with invariant distribution  $\pi$ . Suppose  $X_0$  has distribution  $\pi$ . Define  $Y_k = X_{n-k}$  for  $0 \leq k \leq N$ . Then  $Y$  is a Markov chain with transition matrix  $\hat{p}_{ij} = \frac{\pi_j}{\pi_i} p_{ji}$  and invariant distribution  $\pi$ .

**Definition 20** (Reversible Markov chain). Let  $X$  and  $Y$  be defined as in the theorem above. Then  $X$  is reversible if  $Y$  and  $X$  have the same transition matrix, i.e. if

$$\pi_i p_{ij} = \pi_j p_{ji} \quad \text{for } i, j \in S$$

These are known as the *detailed balance equations*.

**Theorem 14.** Let  $X$  be an irreducible Markov chain with transition matrix  $P$ . Let  $\pi$  be a distribution satisfying the detailed balance equations  $\pi_i p_{ij} = \pi_j p_{ji}$  for  $i, j \in S$ . Then  $\pi$  is the unique invariant distribution.