

RECAP

- ▶ We looked at n -length random vectors \mathbf{X} which are essentially multivariate random variables.
- ▶ Its CDF/pdf is simply joint CDF/pdf of components
- ▶ $E[\mathbf{X}] = [E[X_1], \dots, E[X_n]]$. $C_{\mathbf{X}}$ is the covariance matrix.

Let $\mathbf{Y} = G(\mathbf{X})$ where $G : \mathbb{R}^n \rightarrow \mathbb{R}^n$, continuous invertible with continuous partial derivatives. Let H denote its inverse. Then $f_{\mathbf{Y}}(\mathbf{y}) = f_{\mathbf{X}}(H(\mathbf{y}))|J|$ where J is the determinant of the Jacobian matrix.

- ▶ If $\mathbf{Y} = \mathbf{A}\mathbf{X} + \mathbf{b}$, then $C_{\mathbf{Y}} = \mathbf{A}C_{\mathbf{X}}\mathbf{A}^T$ and

$$f_{\mathbf{Y}}(\mathbf{y}) = \frac{1}{|\det(\mathbf{A})|} f_{\mathbf{X}}(\mathbf{A}^{-1}(\mathbf{y} - \mathbf{b}))$$

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- ▶ A random vector \mathbf{Z} is called as a standard normal vector if its components Z_i are independent and standard normal.

$$f_{\mathbf{Z}}(\mathbf{z}) = \frac{1}{(2\pi)^{n/2}} e^{\{-\frac{1}{2}\mathbf{z}^T \mathbf{z}\}}$$

- ▶ For $\mathbf{X} = A\mathbf{Z} + \boldsymbol{\mu}$. and $E[\mathbf{X}] = \boldsymbol{\mu}$ and $\Sigma := C_{\mathbf{X}} = AA^T$ and

$$f_{\mathbf{X}}(\mathbf{x}) = \frac{1}{(2\pi)^{n/2} \sqrt{\det(C_{\mathbf{X}})}} e^{\{-\frac{1}{2}(\mathbf{x}-\boldsymbol{\mu})^T C_{\mathbf{X}}^{-1}(\mathbf{x}-\boldsymbol{\mu})\}}$$

- ▶ We say that $\mathbf{X} \sim \mathcal{N}(\boldsymbol{\mu}, \Sigma)$

A random vector \mathbf{X} is Gaussian iff for some A and $\boldsymbol{\mu}$, it can be written as $\mathbf{X} = A\mathbf{Z} + \boldsymbol{\mu}$



Affine transformations preserve Gaussianity

- ▶ Suppose $\mathbf{X} \sim \mathcal{N}(\boldsymbol{\mu}, \Sigma)$, then what can we say about $\mathbf{Y} = \mathbf{A}\mathbf{X} + \mathbf{b}$? Is it a Gaussian vector ?
- ▶ Easy to see that $E[\mathbf{Y}] = AE[\mathbf{X}] + b$ and $C_{\mathbf{Y}} = A\Sigma A^T$.
- ▶ Like in the univariate case, we can use MGF (for multivariate MGF see Bertsekas) to show the following

Suppose $\mathbf{X} \sim \mathcal{N}(\boldsymbol{\mu}, \Sigma)$. Now consider $\mathbf{Y} = \mathbf{A}\mathbf{X} + \mathbf{b}$. Then we have $\mathbf{Y} \sim \mathcal{N}(AE[\mathbf{X}] + b, A\Sigma A^T)$.

Equivalent definitions of a Gaussian vector

The following are equivalent definitions (without proof)

$\mathbf{X} \sim \mathcal{N}(\mu, \Sigma)$ iff for some A and μ , it can be written as
 $\mathbf{X} = A\mathbf{Z} + \mu$

\mathbf{X} is Gaussian with mean vector μ and covariance matrix Σ (denoted by $\mathbf{X} \sim \mathcal{N}(\mu, \Sigma)$) iff it has the pdf

$$f_{\mathbf{X}}(\mathbf{x}) = \frac{1}{(2\pi)^{n/2} \sqrt{\det(C_{\mathbf{X}})}} e^{\{-\frac{1}{2}(\mathbf{x}-\mu)^T C_{\mathbf{X}}^{-1}(\mathbf{x}-\mu)\}}$$

$\mathbf{X} \sim \mathcal{N}(\mu, \Sigma)$ iff for all vectors $\mathbf{a} \in \mathbb{R}^n$, it turns out that $\mathbf{a}^T \mathbf{X}$ is univariate Gaussian $\mathcal{N}(\mathbf{a}^T \mu, \mathbf{a}^T \Sigma \mathbf{a})$.

Marginalization and Conditioning (without proof)

- ▶ Let $\mathbf{X} \sim \mathcal{N}(\boldsymbol{\mu}, \Sigma)$ and partition \mathbf{X} as $[\mathbf{X}_1, \mathbf{X}_2]^T$ where \mathbf{X}_1 is $m \times 1$ and \mathbf{X}_2 is $(n - m) \times 1$.
- ▶ We similarly have $\boldsymbol{\mu} = [\boldsymbol{\mu}_1, \boldsymbol{\mu}_2]^T$ and $\Sigma = \begin{bmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{bmatrix}$ where Σ_{11} is $m \times m$ matrix and so on ..

Marginalization property: The m -dimensional marginal distribution of \mathbf{X}_1 is $\mathcal{N}(\boldsymbol{\mu}_1, \Sigma_{11})$ and \mathbf{X}_2 is $\mathcal{N}(\boldsymbol{\mu}_2, \Sigma_{22})$

Conditioning property: The m -dimensional conditional distribution of \mathbf{X}_1 given $\mathbf{X}_2 = \mathbf{x}_2$ is

$$\mathcal{N}(\boldsymbol{\mu}_1 + \Sigma_{12}\Sigma_{22}^{-1}(\mathbf{x}_2 - \boldsymbol{\mu}_2), \Sigma_{11} - \Sigma_{12}\Sigma_{22}^{-1}\Sigma_{21}).$$

Note the decrease in variance which does not depend on \mathbf{x}_2 .

Towards Bivariate Gaussians

- ▶ Suppose $X_1 \sim \mathcal{N}(\mu_1, \sigma_1^2)$ and $X_2 \sim \mathcal{N}(\mu_2, \sigma_2^2)$. Also suppose X_1 and X_2 are independent. Then is $\mathbf{X} = [X_1, X_2]^T$ bivariate Gaussian ? What is the mean vector and Covariance matrix ?
- ▶ For $\mathbf{x} = [x_1, x_2]^T$, we have

$$\begin{aligned} f_{\mathbf{X}}(\mathbf{x}) &= f_{X_1}(x_1)f_{X_2}(x_2) \\ &= \frac{1}{2\pi\sigma_1\sigma_2} e^{-\left(\frac{(x_1-\mu_1)^2}{2\sigma_1^2} + \frac{(x_2-\mu_2)^2}{2\sigma_2^2}\right)} \\ &= \frac{1}{(2\pi)\sqrt{\det(\Sigma)}} e^{\{-\frac{1}{2}(\mathbf{x}-\boldsymbol{\mu})^T \Sigma^{-1}(\mathbf{x}-\boldsymbol{\mu})\}} \end{aligned}$$

where $\Sigma = \begin{bmatrix} \sigma_1^2 & 0 \\ 0 & \sigma_2^2 \end{bmatrix}$ and hence \mathbf{X} is bivariate Gaussian.

- ▶ In general, a vector composed of independent Gaussians is a Gaussian vector. The converse is not true: a vector of dependent Gaussian components need not be Gaussian vector (EX 5.35 in probabilitycourse.com).

Bivariate Gaussians

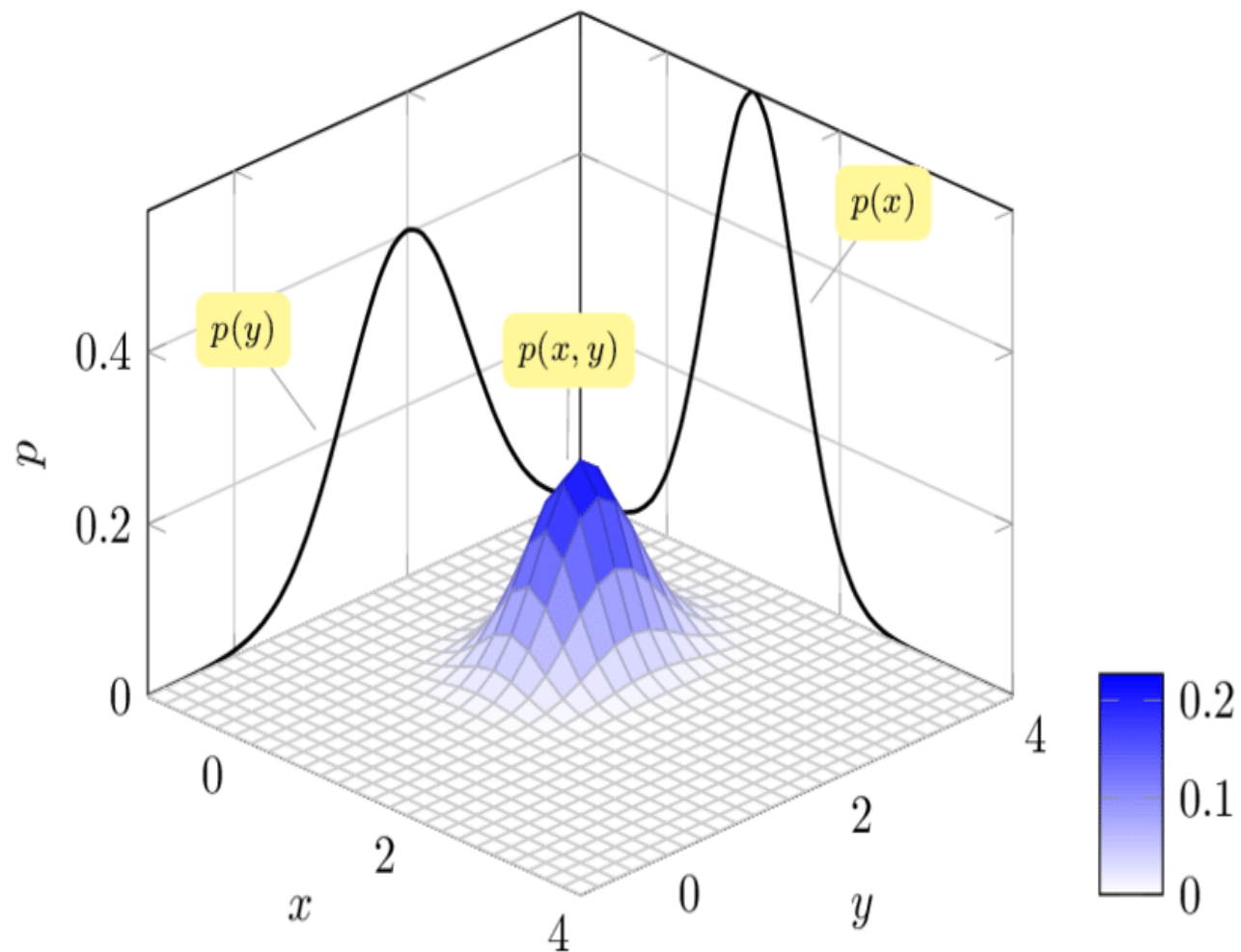
- ▶ In general, X_1 and X_2 need not be independent in which case we have a general bivariate Gaussian

$\mathbf{X} = [X_1, X_2]^T \sim \mathcal{N}(\boldsymbol{\mu}, \Sigma)$ where $\boldsymbol{\mu} = [\mu_1, \mu_2]^T$ and

$$\Sigma = \begin{bmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{bmatrix}$$

- ▶ Show that Bivariate Gaussian is closed under marginalization, i.e., $X_1 \sim \mathcal{N}(\mu_1, \Sigma_{11})$ and $X_2 \sim \mathcal{N}(\mu_2, \Sigma_{22})$.
- ▶ Show that Bivariate Gaussian is closed under conditioning. (For proof see Theorem 5.4, probabilitycourse.com).
- ▶ This means that Given $X_2 = x_2$, one can show that $f_{X_1|X_2}(x_1|x_2)$ is Gaussian.
- ▶ These two properties make multivariate Gaussians as efficient modelling tools and the handy in Gaussian processes and Bayesian optimization.

Some Bivariate gaussian pdfs



Some Bivariate gaussian pdfs

