

# Mathematical Notes MSc Thesis

Lucas Paul Unterweger

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## 1 Important Prior Distributions

### Normal Distribution

$$f(x|\mu, \sigma) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}$$

### Gamma Distribution

$$f(x|\alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\beta x}, \quad x > 0$$

where  $\alpha$  is the shape parameter,  $\beta$  is the rate parameter, and  $\Gamma(\cdot)$  is the gamma function.

### Beta Distribution

$$f(x|a, b) = \frac{x^{a-1}(1-x)^{b-1}}{B(a, b)}$$

where  $x \in [0, 1]$ ,  $a > 0$ ,  $b > 0$ , and  $B(a, b)$  is the Beta function defined as:

$$B(a, b) = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}$$

where  $\Gamma$  is the gamma function.

## 2 Other Functions

**Confluent Hyper-Geometrics Function of the second kind (Tricomi's (confluent hypergeometric) function)**

$$U(a, b, z) = \frac{1}{\Gamma(a)} \int_0^\infty e^{-zt} t^{a-1} (1+t)^{b-a-1} dt$$

**Marginal Prior with  $\phi^\xi = \frac{2c^\xi}{\kappa_B^2 a^\xi}$**

$$\begin{aligned} p(\sqrt{\theta_j}|\phi^\xi, a^\xi, c^\xi) &= \frac{\Gamma(c^\xi + \frac{1}{2})}{\sqrt{2\pi\phi^\xi} \cdot B(a^\xi, c^\xi)} \cdot U\left(c^\xi + \frac{1}{2}, \frac{3}{2} - a^\xi, \frac{\theta_j}{2\phi^\xi}\right) \\ &\propto U\left(c^\xi + \frac{1}{2}, \frac{3}{2} - a^\xi, \frac{\theta_j}{2\phi^\xi}\right) \end{aligned}$$

### 3 Bayesian Regression

Consider a Bayesian regression model where the relationship between the independent variable  $x$  and the dependent variable  $y$  is modeled as:

$$y = \beta_0 + \beta_1 x + \epsilon$$

where  $\epsilon$  is the error term assumed to follow a normal distribution with mean 0 and variance  $\sigma^2$ .

We assume the following conjugate prior distributions for the parameters:

$$\begin{aligned}\beta_0 &\sim \mathcal{N}(\mu_0, \tau_0^2) \\ \beta_1 &\sim \mathcal{N}(\mu_1, \tau_1^2) \\ \sigma^2 &\sim \text{Inv-Gamma}(\alpha, \beta)\end{aligned}$$

The likelihood function for the observed data  $(x_i, y_i)$ ,  $i = 1, \dots, n$ , is given by:

$$L(\beta_0, \beta_1, \sigma^2 | x, y) = \prod_{i=1}^n f(y_i | \beta_0, \beta_1, x_i, \sigma^2)$$

where  $f(y_i | \beta_0, \beta_1, x_i, \sigma^2)$  is the probability density function (PDF) of the normal distribution.

The posterior distribution of the parameters, denoted as  $\pi(\beta_0, \beta_1, \sigma^2 | x, y)$ , is proportional to the product of the likelihood function and the prior distributions:

$$\pi(\beta_0, \beta_1, \sigma^2 | x, y) \propto L(\beta_0, \beta_1, \sigma^2 | x, y) \times \pi(\beta_0) \times \pi(\beta_1) \times \pi(\sigma^2)$$

The maximum a posteriori (MAP) estimation involves maximizing the log-posterior function:

$$\ell(\beta_0, \beta_1, \sigma^2 | x, y) = \sum_{i=1}^n \log f(y_i | \beta_0, \beta_1, x_i, \sigma^2) + \log \pi(\beta_0) + \log \pi(\beta_1) + \log \pi(\sigma^2)$$

with respect to the parameters  $\beta_0$ ,  $\beta_1$ , and  $\sigma^2$ .

The estimates of the parameters, denoted as  $\hat{\beta}_0$ ,  $\hat{\beta}_1$ , and  $\hat{\sigma}^2$ , can be obtained by maximizing the log-posterior function.

## 4 Derviation Idea

$$post(\beta, \sigma^2|x, y) \propto L(\beta, \sigma^2|x, y) \times prior(\beta, \sigma^2)$$