

# Design of Experiment and Analysis of Variance

Repetition

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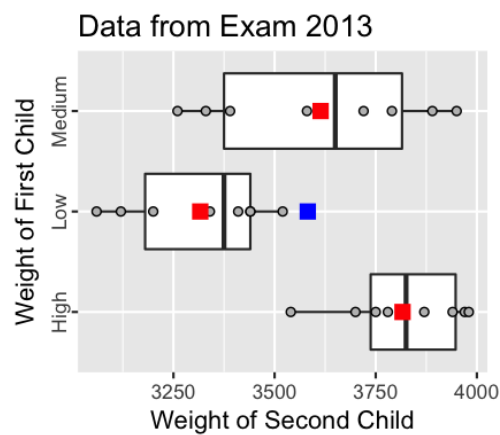


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# ANOVA Model

# ANOVA Model

## ANOVA table



	Df	Sum Sq	Mean Sq	F value
Weight1	2	1012033	506017	12.5
Residuals	21	851962	40570	

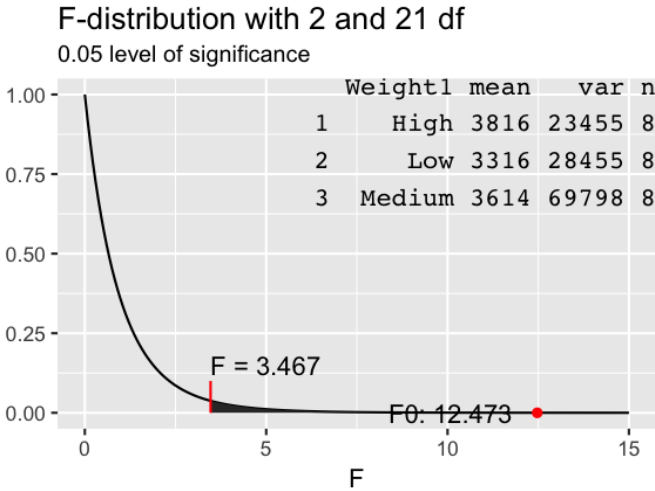
## Hypothesis

$$H_0 : \tau_i = 0 \text{ for } i = 1, 2, 3$$

$$H_1 : \tau_i \neq 0 \text{ for at least one } i$$

## Decision

From **F-table**,  $F_0 > F_c$ .  
So we reject  $H_0$  at 95% confidence level.



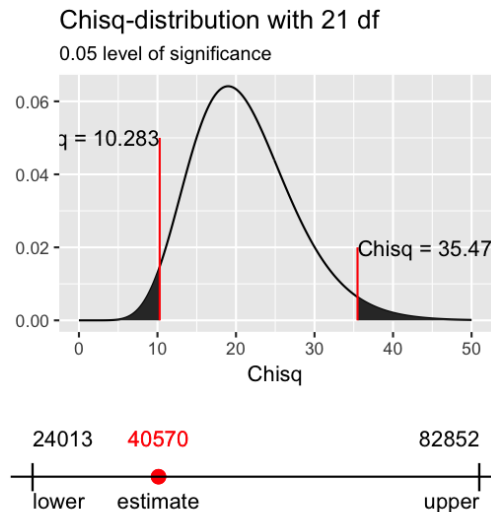
## Questions

a) What is the 95% **confidence interval of error variance**?

b) We found that there is significant effect of first child's weight on second child, but *how different is the effect of High from low and medium?*

# ANOVA Model

## Confidence Interval



	Df	Sum Sq	Mean Sq	F value
Weight1	2	1012033	506017	12.5
Residuals	21	851962	40570	

## Confidence interval of error variance

$$\frac{SSE}{\chi_{\alpha/2, N-a}^2} \leq \sigma^2 \leq \frac{SSE}{\chi_{1-(\alpha/2), N-a}^2}$$

## Exam 2013 Data

Using values in previous slide in above formula,

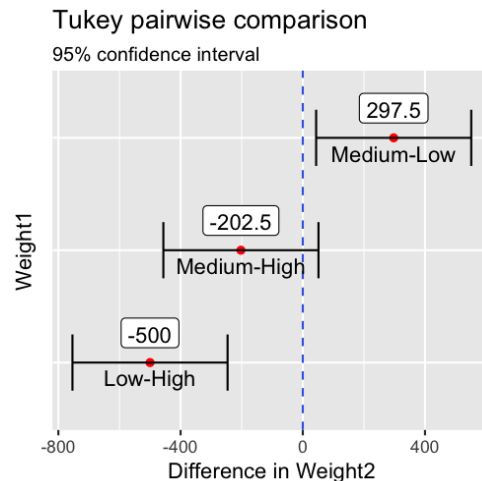
$$\left[ \frac{851962.5}{35.479}, \frac{851962.5}{10.283} \right] = [24013.2, 82852.4]$$

At **95% confidence level**, we can say that the true error variance  $\sigma^2$  lies between the interval (24013.2, 82852.4).

## Exercise: Exam 2011 1(b)

# ANOVA Model

## Post-hoc Test



Think what will happen if you increase the confidence level from 0.05 to 0.1.

## Tukey Pairwise Comparison

For multiple testing, Tukey has suggested **studentized range statistic** as,

$$q = \frac{\bar{y}_{\max} - \bar{y}_{\min}}{\sqrt{\text{MSE}/n}}$$

Compare  $q$  with  $q_{\alpha}(a, f)$  found in [table](#). We declare a pair to be significantly different if  $q > q_{\alpha}(a, f)$ .

## Confidence Interval

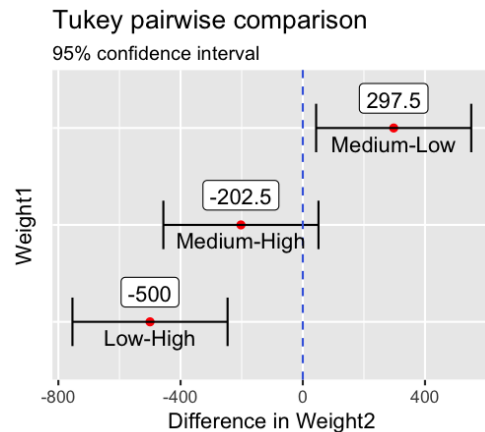
$$(\bar{y}_i - \bar{y}_j) \pm q_{\alpha}(a, f) \sqrt{\frac{\text{MSE}}{n}}$$

$$(\bar{y}_i - \bar{y}_j) - q_{\alpha}(a, f) \sqrt{\frac{\text{MSE}}{n}} \leq \mu_i - \mu_j \leq (\bar{y}_i - \bar{y}_j) + q_{\alpha}(a, f) \sqrt{\frac{\text{MSE}}{n}}$$

You can also check if this interval contains zero.

# ANOVA Model

## Post-hoc Example



```
Weight1 mean n
High      3816 8
Low       3316 8
Medium    3614 8
```

## Exam 2013, 2(d)

From the [Appendix 1](#) and [Appendix 4](#), we have,

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Weight1	2	1012033	506017	12.5	0.000269
Residuals	21	851962	40570	NA	NA

Using [studentized, T-distribution table](#)  $q_{0.05}(3, 21) = 3.565$ . So,

$$T_{0.05} = q_{0.05}(3, 21) \sqrt{\frac{40569.643}{8}} = 253.845$$

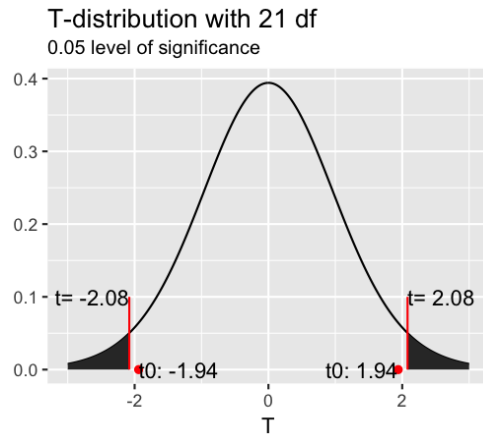
We have three possible pairs for comparison: Low-High, Medium-High and Medium-Low. The difference in their means are: -500, -202.5, 297.5 respectively.

If compared with  $T_{0.05}(3, 21)$ , we see that at 95% confidence level, we find Low-High and Medium-Low differ significantly.

## Exam 2015, 1(e)

# ANOVA Model

## Contrast test



## Use cases

- Comparison of one group with average of other group
- Comparing treatments with control

## Exam 2014: 2(d)

### Hypothesis:

$$H_0 : \frac{1}{2}(\mu_1 + \mu_2) - \frac{1}{2}(\mu_3 + \mu_4) = 0 \text{ vs } H_1 : \frac{1}{2}(\mu_1 + \mu_2) - \frac{1}{2}(\mu_3 + \mu_4) \neq 0$$

### Decision:

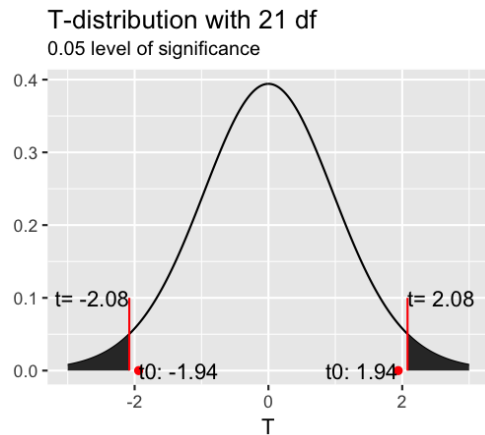
From [Table 8](#), p-value (0.065) > 0.05.

► *What should we conclude?*



# ANOVA Model

## Contrast Calculation



Continue on Exam 2014:  
2(d)

Before any computation, **formulate a hypothesis**.

## Estimate of Contrast

$$\Gamma = \sum_{i=1}^a c_i \mu_i = \frac{1}{2}(\mu_1 + \mu_2) - \frac{1}{2}(\mu_3 - \mu_4) = -5.5$$

Here the contrast coefficients  $c_i$  are:  $1/2, 1/2, -1/2, -1/2$ . Contrast coefficients sum to zero. For estimation of  $\Gamma$ , use respective sample mean. Also use [Table 6](#) for following calculations.

## Standard Error and test statistic

$$SE(\hat{\Gamma}) = \sqrt{\frac{MSE}{n} \sum_{i=1}^a c_i^2} = 2.82 \quad t = \frac{\hat{\Gamma}}{S.E(\hat{\Gamma})} \sim t_{\alpha/2, N-a}$$

# ANOVA Model

## Random Effect Model

### Random factor

$$\tau_i \sim N(0, \sigma_\tau^2)$$

### Fixed factor

$$\sum_{i=1}^a \tau_i = 0$$

## The Model

$$y_{ij} = \mu + \tau_i + \varepsilon_{ij} \text{ where, } \varepsilon_{ij} \sim N(0, \sigma^2) \text{ and } \tau_i \sim N(0, \sigma_\tau^2)$$

## Why Random effect model

*Specific levels are not of interest.* General variation is more important. Levels of a factor are **randomly selected**. For example, if besetning (farm) is not of interest, we randomly sample such farms from a population of farm.

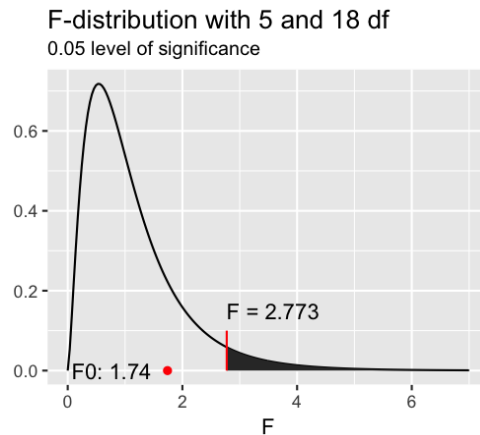
Usually, **blocks** are taken as random factor.

## A comparison with Fixed Effect Model

*Specific levels of a factor is important* than their general variation. These levels are **specifically chosen**. For instance, comparison of new and old drug where a particular drug is used for comparison.

# ANOVA Model

## Random Effect Model



### Exam 2011: 1(a)

#### Hypothesis:

$$H_0 : \sigma_\tau^2 = 0 \text{ vs } H_1 : \sigma_\tau^2 > 0$$

From [Appendix 2](#), we see p-value is larger than  $\alpha = 0.05$ , we can not reject the null hypothesis. So, we claim that **at 95% confidence level** we claim that there is *not significant variation* between the farms.

### Estimates of Variance Components

$$\text{total variation } \text{var}(y_{ij}) = \sigma_\tau^2 + \sigma^2$$

We estimate them as,

$$\hat{\sigma}^2 = \text{MSE} \text{ and } \hat{\sigma}_\tau^2 = \frac{\text{MS}_{\text{treatment}} - \text{MSE}}{n}$$

# ANOVA Model

## Random Effect Model

Exam 2011: 1(c), 1(d)

Exam 2012: 1(c)

Exam 2013: 1(c)

Exam 2014: 1(e)

Exam 2015: 2(c)

## Intraclass Correlation Coefficient

Proportion of variation between groups to total variation.

$$\rho = \frac{\sigma_{\tau}^2}{\sigma_{\tau}^2 + \sigma^2}$$

Using estimates of variance components, we can estimate intraclass correlation coefficient.

## Confidence interval of overall mean

The  $100(1 - \alpha)$  level of confidence interval for overall mean  $\mu$  in case of **random effect model** is,

$$\hat{\mu} \pm t_{\alpha/2, a-1} \sqrt{\frac{MS_{\text{treatment}}}{N}}$$

(Refer to Thore's Lecture on Random effect Model)