ASHR SVAR Models

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Setup

Load packages & functions

```
rm(list=ls())
require(tinytex) #LaTeX
require(ggplot2) #plots
require(AEC) #JP-Renne functions
require(AER) #NW formula
require(forecast) #time series stuff
require(expm) #matrix exponents
require(here) #directory finder
require(stringr) # analysis of strings, important for the detection in tweets
require(dplyr) #data management
require(lubridate) #data dates management
require(zoo) #for lagging
require(jtools) #tables
require(huxtable) #tables
require(lmtest) #reg tests
require(vroom) #for loading data
require(data.table) #for data filtering
require(sysid) #for ARMA-X modeling
require(sandwhich) #regression errors
require(stargazer) #nice req tables
require(tidytext) #text mining
require(textstem) #lemmatization
require(quanteda) #tokenization
require(texreg) #arima tables
require(vars) #VAR models
require(xts) #time series objects
require(tseries) #includes adf test
require(quantmod)
require(TSA)
require(aTSA)
require(tibble)
require(FinTS)
require(kableExtra)
require(writexl)
require(purrr)
getwd()
#setwd("...") -> set wd at base repo folder
#load helper functions
source(here("helperfunctions/data loaders.R"))
source(here("helperfunctions/date_selector.R"))
source(here("helperfunctions/plotters.R"))
source(here("helperfunctions/quick_arma.R"))
source(here("helperfunctions/r.vol_calculators.R"))
source(here("helperfunctions/truths cleaning function.R"))
source(here("helperfunctions/armax_functions.R"))
source(here("helperfunctions/var_irf.R"))
```

Load Data

```
#load final dataset
source(here("helperfunctions/full_data.R"))

#select timeframe
Vdata = filter(data, between(timestamp, as.Date('2014-01-01'), as.Date('2025-05-07')))
```

Some SVAR estimations

Note that this is not an exhaustive list of our VAR estimations, you can find more by going on /modeling/VAR/VAR_SPY_FULLMODELS.rmd or VAR_ASHR_FULLMODELS.rmd or VAR_VGK_FULLMODELS.rmd).

Dummy variable

Here we use a dummy variable which equal to one if Trump has made a post or 0 otherwise, taking into account the closed hour market posts.

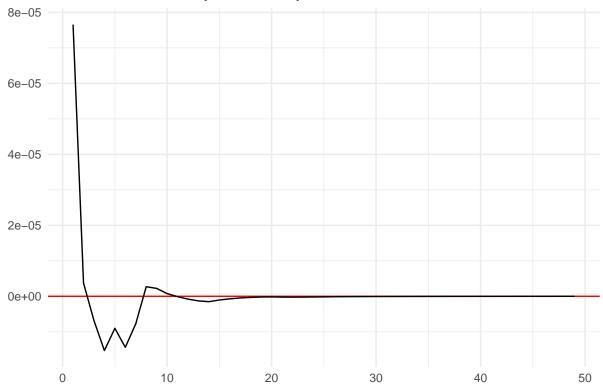
```
y = cbind(Vdata$dummy, Vdata$ASHR_vol)
colnames(y)[1:2] <- c("dummy", "vol")</pre>
est.VAR <- VAR(y,p=6)
#extract results
mod_vol = est.VAR$varresult$vol
f = formula(mod_vol)
d = model.frame(mod_vol)
lm_clean = lm(f, data = d)
#apply Newey-West
nw_vcov = NeweyWest(lm_clean, lag=6)
nw_se = sqrt(diag(nw_vcov))
\#t\text{-}stats
coef = coef(lm clean)
t_stat = coef/nw_se
#recalculate p-values
robust = 2*(1-pt(abs(t_stat), df = df.residual(lm_clean)))
            <- nw_se[names(coef(lm_clean))]</pre>
nw_se
robust
            <- robust[names(coef(lm_clean))]</pre>
#table
screenreg(lm_clean, override.se = nw_se, override.pvalues = robust, digits = 6)
```

```
-0.000006 ***
## dummy.l1
##
                 (0.000001)
## vol.11
                 0.282482 ***
                 (0.023326)
##
## dummy.12
                 -0.000005 ***
##
                 (0.000001)
## vol.12
                 0.072926 *
                 (0.031382)
##
## dummy.13
                -0.000006 ***
                 (0.000001)
##
## vol.13
                  0.047892 ***
##
                 (0.008643)
                 -0.000004 **
## dummy.14
                 (0.000001)
##
## vol.14
                 0.056084 ***
##
                 (0.014109)
## dummy.15
                -0.000006 ***
                 (0.000001)
## vol.15
                 0.059763 ***
##
                 (0.017855)
## dummy.16
                -0.000005 ***
                 (0.000001)
## vol.16
                0.109466 ***
##
                 (0.023870)
## const
                 0.000095 ***
                 (0.000009)
## --
## R^2
                  0.270711
## Adj. R^2
                  0.270236
## Num. obs. 19965
## ==========
## *** p < 0.001; ** p < 0.01; * p < 0.05
#extract results
mod_post = est.VAR$varresult$dummy
ff = formula(mod_post)
dd = model.frame(mod_post)
lm_clean_post = lm(ff, data= dd)
#apply Newey-West
nw_vcov_post = NeweyWest(lm_clean_post, lag=6)
nw_se_post = sqrt(diag(nw_vcov_post))
#t-stats
coef_post = coef(lm_clean_post)
t_stat_post = coef_post/nw_se_post
\#recalculate\ p-values
robust_post = 2*(1-pt(abs(t_stat_post), df = df.residual(lm_clean_post)))
nw se post
               <- nw se post[names(coef(lm clean post))]</pre>
robust_post
               <- robust_post[names(coef(lm_clean_post))]</pre>
#table
```

```
##
## ===========
            Model 1
## -----
## dummy.11
               -0.092434 ***
##
              (0.003468)
## vol.l1
             -24.484296
             (31.437508)
##
## dummy.12
              -0.084188 ***
##
              (0.003669)
## vol.12
             -95.442834
##
              (57.815084)
## dummy.13
             -0.077451 ***
##
              (0.004069)
               5.890426
## vol.13
##
              (81.556094)
## dummy.14
               -0.075739 ***
##
               (0.003714)
## vol.14
            -153.454700 ***
              (41.958279)
##
## dummy.15
             -0.087766 ***
##
               (0.003737)
## vol.15
               -7.311264
             (32.226447)
              -0.099812 ***
## dummy.16
              (0.004184)
             181.502758 **
## vol.16
##
              (62.005728)
## const
               1.724904 ***
               (0.040648)
## -----
## R^2
               0.154875
## Adj. R^2
               0.154324
## Num. obs. 19965
## ==========
## *** p < 0.001; ** p < 0.01; * p < 0.05
#HO test whether there is NOT heteroscedasticity. if less by alpha, then there is heteroscedasticity
bptest(lm_clean)
##
## studentized Breusch-Pagan test
## data: lm_clean
## BP = 160.42, df = 12, p-value < 2.2e-16
#Recreate a Robust Omega Matrix
U = residuals(est.VAR)
T = nrow(U)
L = 6 #number of lag
```

```
Omega = matrix(0, ncol(U), ncol(U))
for(l in 0:L) {
  weight = 1 - 1/(L+1)
  Gamma_l = t(U[(1+1):T, , drop=FALSE]) %*% U[1:(T-1), , drop=FALSE] /T
  if (1 == 0){
    Omega = Omega + Gamma_1_
  } else {
    Omega = Omega + weight*(Gamma_l_ + t(Gamma_l_))
  }
}
#make the B matrix
loss <- function(param){</pre>
  #Define the restriction
  B \leftarrow matrix(c(param[1], param[2], 0, param[3]), ncol = 2)
  #Make BB' approximatively equal to omega
  X <- Omega - B %*% t(B)
  #loss function
 loss <- sum(X^2)
  return(loss)
}
res.opt \leftarrow optim(c(1, 0, 1), loss, method = "BFGS")
B.hat <- matrix(c(res.opt$par[1], res.opt$par[2], 0, res.opt$par[3]), ncol = 2)
print(cbind(Omega,B.hat %*% t(B.hat)))
##
                 dummy
                                 vol
## dummy 1.024704e+01 2.442181e-04 1.024704e+01 2.450982e-04
         2.442181e-04 1.462361e-07 2.450982e-04 4.466506e-05
B.hat
##
                 [,1]
                              [,2]
## [1,] 3.201099e+00 0.000000000
## [2,] 7.656688e-05 0.006682754
nb.sim = 7*7
#get back the coefficient of est.VAR
phi <- Acoef(est.VAR)</pre>
PHI = make.PHI(phi)
#take the constant
constant <- sapply(est.VAR$varresult, function(eq) coef(eq)["const"])</pre>
c=as.matrix(constant)
#Simulate the IRF
p <- length(phi)</pre>
n <- dim(phi[[1]])[1]</pre>
```

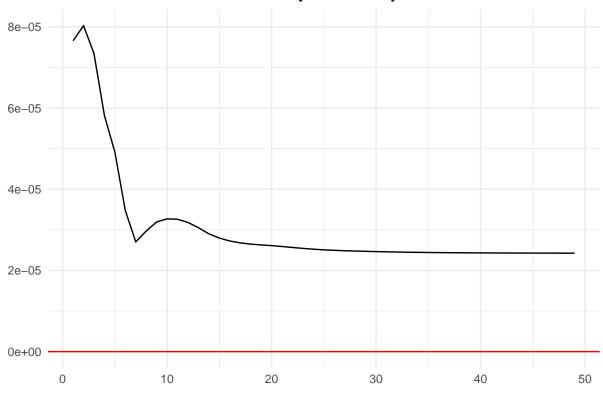
ASHR IRF of Dummy on Volatility



```
ggplot(Yd,aes(x=period, y=cumsum(response))) +
  geom_hline(yintercept = 0, color="red") +
  geom_line() +
  theme_light() +
  ggtitle("ASHR Cumulality IRF of Dummy on Volatility") +
  ylab("")+
```

```
xlab("") +
theme_minimal()
```

ASHR Cumulalitve IRF of Dummy on Volatility



Post Counts

```
y2 = cbind(Vdata$N, Vdata$ASHR_vol)
colnames(y2)[1:2] <- c("N", "vol")
est.VAR2 <- VAR(y2,p=6)

#extract results
mod_vol2 = est.VAR2$varresult$vol
f2 = formula(mod_vol2)
d2 = model.frame(mod_vol2)
lm_clean2 = lm(f2, data= d2)

#apply Newey-West
nw_vcov2 = NeweyWest(lm_clean2, lag=6)
nw_se2 = sqrt(diag(nw_vcov2))

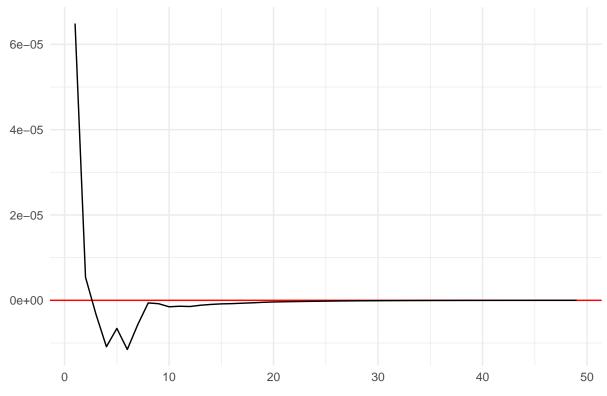
#t-stats
coef2 = coef(lm_clean2)
t_stat2 = coef2/nw_se2</pre>
```

```
#recalculate p-values
robust2 = 2*(1-pt(abs(t_stat2), df = df.residual(lm_clean2)))
          <- nw_se2[names(coef(lm_clean2))]</pre>
          <- robust2[names(coef(lm_clean2))]</pre>
robust2
#table
screenreg(lm_clean2, override.se = nw_se2, override.pvalues = robust2, digits = 6)
##
## ==========
            Model 1
             -0.000001 ***
## N.11
              (0.000000)
##
              0.282497 ***
## vol.11
##
               (0.023442)
## N.12
              -0.000001 ***
               (0.000000)
##
                0.072640 *
## vol.12
##
              (0.031158)
## N.13
              -0.000002 ***
                (0.000000)
##
## vol.13
               0.047738 ***
                (0.008436)
##
## N.14
                -0.000001 *
##
               (0.000000)
## vol.14
                0.056237 ***
                (0.014097)
              -0.000002 ***
## N.15
##
                (0.000000)
## vol.15
               0.059528 ***
                (0.017745)
## N.16
                -0.000001 ***
               (0.000000)
                0.109380 ***
## vol.16
                (0.023887)
                0.000081 ***
## const
##
                (0.000009)
## ---
             ______
## R^2
                 0.268901
## Adj. R^2
                0.268425
## Num. obs. 19965
## =============
## *** p < 0.001; ** p < 0.01; * p < 0.05
#extract results
mod_post2 = est.VAR2$varresult$N
ff2 = formula(mod_post2)
dd2 = model.frame(mod_post2)
lm_clean_post2 = lm(ff2, data = dd2)
#apply Newey-West
```

```
nw_vcov_post2 = NeweyWest(lm_clean_post2, lag=6)
nw_se_post2 = sqrt(diag(nw_vcov_post2))
#t-stats
coef_post2 = coef(lm_clean_post2)
t_stat_post2 = coef_post2/nw_se_post2
#recalculate p-values
robust_post2 = 2*(1-pt(abs(t_stat_post2), df = df.residual(lm_clean_post2)))
nw_se_post2
                - nw_se_post2[names(coef(lm_clean_post2))]
                <- robust_post2[names(coef(lm_clean_post2))]</pre>
robust_post2
#table
screenreg(lm_clean_post2, override.se = nw_se_post2, override.pvalues = robust_post2, digits = 6)
##
## ==========
##
           Model 1
## -----
## N.l1
             -0.042465 ***
##
               (0.003696)
             -129.076390
## vol.11
             (98.105219)
## N.12
               -0.038882 ***
               (0.004629)
## vol.12
              -22.372568
             (276.983664)
## N.13
              -0.026617 ***
##
               (0.004632)
              -26.972939
## vol.13
##
             (204.041589)
               -0.021873 ***
## N.14
##
               (0.005161)
## vol.14
             -433.230826 **
             (146.247539)
##
## N.15
              -0.040310 ***
##
               (0.004097)
## vol.15
              -27.699216
##
             (90.152967)
## N.16
               -0.047981 ***
##
               (0.005136)
             548.279747 **
## vol.16
              (199.229337)
##
## const
                3.519787 ***
##
               (0.095735)
## -----
## R^2
                0.100612
## Adj. R^2
                0.100026
## Num. obs. 19965
## ============
## *** p < 0.001; ** p < 0.01; * p < 0.05
```

```
#Recreate a Robust Omega Matrix
U2 = residuals(est.VAR2)
T2 = nrow(U2)
Omega2 = matrix(0, ncol(U2), ncol(U2))
for(l in 0:L) {
  weight = 1 - 1/(L+1)
  Gamma_1_2 = t(U2[(1+1):T2, , drop=FALSE]) %*% U2[1:(T2-1), , drop=FALSE] /T2
  if (1 == 0){
    Omega2 = Omega2 + Gamma_1_2
 } else {
    Omega2 = Omega2 + weight*(Gamma_1_2 + t(Gamma_1_2))
}
#make the B matrix
loss2 <- function(param2){</pre>
  #Define the restriction
  B2 \leftarrow matrix(c(param2[1], param2[2], 0, param2[3]), ncol = 2)
  #Make BB' approximatively equal to omega
  X2 <- Omega2 - B2 %*% t(B2)
  #loss function
 loss2 \leftarrow sum(X2^2)
  return(loss2)
}
res.opt2 \leftarrow optim(c(1, 0, 1), loss2, method = "BFGS")
B.hat2 \leftarrow matrix(c(res.opt2$par[1], res.opt2$par[2], 0, res.opt2$par[3]), ncol = 2)
print(cbind(Omega2,B.hat2 %*% t(B.hat2)))
##
                             lov
       86.36287757 5.984100e-04 8.636288e+01 6.030108e-04
## vol 0.00059841 1.452053e-07 6.030108e-04 2.564515e-05
B.hat2
##
                 [,1]
                              [,2]
## [1,] 9.293163e+00 0.000000000
## [2,] 6.488758e-05 0.005063688
#get back the coefficient of est.VAR
phi2 <- Acoef(est.VAR2)</pre>
PHI2 = make.PHI(phi2)
#take the constant
constant2 <- sapply(est.VAR2$varresult, function(eq) coef(eq)["const"])</pre>
c2=as.matrix(constant2)
#Simulate the IRF
p2 <- length(phi2)
```

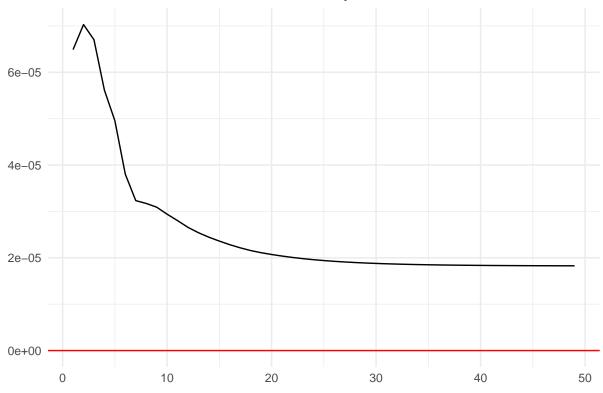
ASHR IRF of N on Volatility



```
ggplot(Yd2,aes(x=period, y=cumsum(response))) +
  geom_hline(yintercept = 0, color="red") +
  geom_line() +
  theme_light() +
  ggtitle("ASHR Cumulality IRF of N on Volatility") +
```

```
ylab("")+
xlab("") +
theme_minimal()
```

ASHR Cumulalitve IRF of N on Volatility



Res.Df	\mathbf{Df}	\mathbf{F}	$\Pr(>F)$
2e + 04			
2e + 04	-6	13.9	8.02e-16

grangertest(y2[,c("vol", "N")], order = 6)

Res.Df	\mathbf{Df}	${f F}$	$\Pr(>F)$
2e+04			
2e+04	-6	3.17	0.00413

Trade Mention

```
y4 = cbind(Vdata$trade, Vdata$ASHR_vol)
colnames(y4)[1:2] <- c("trade", "vol")</pre>
est.VAR4 \leftarrow VAR(y4,p=6)
#extract results
mod_vol4 = est.VAR4$varresult$vol
f4 = formula(mod_vol4)
d4 = model.frame(mod_vol4)
lm_clean4 = lm(f4, data = d4)
#apply Newey-West
nw_vcov4 = NeweyWest(lm_clean4, lag=6)
nw_se4 = sqrt(diag(nw_vcov4))
#t-stats
coef4 = coef(lm_clean4)
t_stat4 = coef4/nw_se4
#recalculate p-values
robust4 = 2*(1-pt(abs(t_stat4), df = df.residual(lm_clean4)))
           <- nw_se4[names(coef(lm_clean4))]</pre>
nw_se4
            <- robust4[names(coef(lm_clean4))]</pre>
robust4
#table
screenreg(lm_clean4, override.se = nw_se4, override.pvalues = robust4, digits = 6)
```

```
##
## ==========
##
        Model 1
## -----
             -0.000025 ***
## trade.l1
##
               (0.000006)
              0.281371 ***
## vol.l1
              (0.022717)
## trade.12
               0.000012
##
               (0.000019)
## vol.12
               0.071544 *
##
               (0.030181)
## trade.13
               -0.000019 *
               (0.000009)
##
## vol.13
               0.045746 ***
##
               (0.008280)
## trade.14
               -0.000009
##
               (0.000008)
## vol.14
               0.056738 ***
               (0.013529)
##
## trade.15
               -0.000013 ***
               (0.000004)
##
## vol.15
               0.056226 **
               (0.017213)
##
```

```
## trade.16 -0.000013 ***
##
              (0.000003)
             0.109845 ***
## vol.16
             (0.022503)
##
## const
               0.000062 ***
##
              (0.000009)
## -----
## R^2
              0.266464
## Adj. R^2 0.265986
## Num. obs. 19965
## ==========
## *** p < 0.001; ** p < 0.01; * p < 0.05
#Table for the effect of volatility on posts for variable trade
#extract results
mod_post4 = est.VAR4$varresult$trade
ff4 = formula(mod_post4)
dd4 = model.frame(mod_post4)
lm_clean_post4 = lm(ff4, data = dd4)
#apply Newey-West
nw_vcov_post4 = NeweyWest(lm_clean_post4, lag=6)
nw_se_post4 = sqrt(diag(nw_vcov_post4))
#t-stats
coef_post4 = coef(lm_clean_post4)
t_stat_post4 = coef_post4/nw_se_post4
#recalculate p-values
robust_post4 = 2*(1-pt(abs(t_stat_post4), df = df.residual(lm_clean_post4)))
<- robust_post4[names(coef(lm_clean_post4))]</pre>
robust_post4
screenreg(lm_clean_post4, override.se = nw_se_post4, override.pvalues = robust_post4, digits = 6)
## ===========
          Model 1
## -----
## trade.l1
            0.026737
              (0.018212)
##
## vol.11
            -9.226534 **
##
             (3.417063)
## trade.12
              0.020039 *
##
              (0.008617)
## vol.12
              4.799208
             (6.150288)
## trade.13
            0.022165
##
              (0.016504)
            25.623669
## vol.13
##
            (25.460008)
## trade.14
             0.024458
```

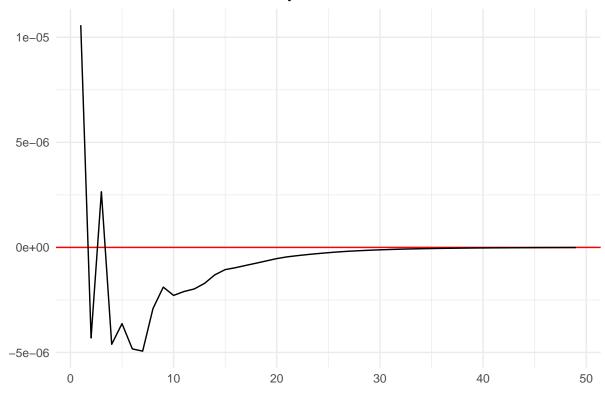
```
##
                (0.012912)
## vol.14
               -14.855768 *
##
                (7.295870)
## trade.15
                 0.018788
##
                 (0.010879)
## vol.15
                -0.433456
                 (4.696217)
## trade.16
                 0.013759
##
                 (0.010389)
## vol.16
                -2.276427
                (2.921722)
                 0.029769 ***
## const
                 (0.002707)
## -----
## R^2
                  0.018726
## Adj. R^2
                  0.018087
## Num. obs. 19965
## ===========
## *** p < 0.001; ** p < 0.01; * p < 0.05
#Recreate a Robust Omega Matrix
U4 = residuals(est.VAR4)
T4 = nrow(U4)
Omega4 = matrix(0, ncol(U4), ncol(U4))
for(l in 0:L) {
  weight = 1 - 1/(L+1)
  Gamma_1_4 = t(U4[(1+1):T4, , drop=FALSE]) %*% U4[1:(T4-1), , drop=FALSE] /T4
  if (1 == 0){
    Omega4 = Omega4 + Gamma_1_4
  } else {
    Omega4 = Omega4 + weight*(Gamma_l_4 + t(Gamma_l_4))
}
#make the B matrix
loss4 <- function(param4){</pre>
  #Define the restriction
 B4 <- matrix(c(param4[1], param4[2], 0, param4[3]), ncol = 2)
  #Make BB' approximatively equal to omega
  X4 <- Omega4 - B4 %*% t(B4)
  #loss function
  loss4 \leftarrow sum(X4^2)
  return(loss4)
}
res.opt4 \leftarrow optim(c(1, 0, 1), loss4, method = "BFGS")
B.hat4 <- matrix(c(res.opt4\$par[1], res.opt4\$par[2], 0, res.opt4\$par[3]), ncol = 2)
print(cbind(Omega4,B.hat4 %*% t(B.hat4)))
```

trade vol

```
## trade 8.201947e-02 3.026524e-06 8.201858e-02 3.028763e-06
        3.026524e-06 1.426699e-07 3.028763e-06 1.210986e-05
B.hat4
##
                [,1]
                             [,2]
## [1,] 0.2863888688 0.0000000000
## [2,] 0.0000105757 0.003479906
#get back the coefficient of est.VAR
phi4 <- Acoef(est.VAR4)</pre>
PHI4 = make.PHI(phi4)
#take the constant
constant4 <- sapply(est.VAR4$varresult, function(eq) coef(eq)["const"])</pre>
c4=as.matrix(constant4)
#Simulate the IRF
p4 <- length(phi4)
n4 <- dim(phi4[[1]])[1]
Y4 <- simul.VAR(c=c4, Phi = phi4, B = B.hat4, nb.sim ,y0.star=rep(0, n4*p4),
                  indic.IRF = 1, u.shock = c(1,0))
#Plot the IRF
Yd4 = data.frame(
  period = 1:nrow(Y4),
 response = Y4[,2])
ggplot(Yd4,aes(x=period, y=response)) +
  geom_hline(yintercept = 0, color="red") +
  geom_line() +
  theme_light() +
  ggtitle("ASHR IRF of Trade on Volatility") +
```

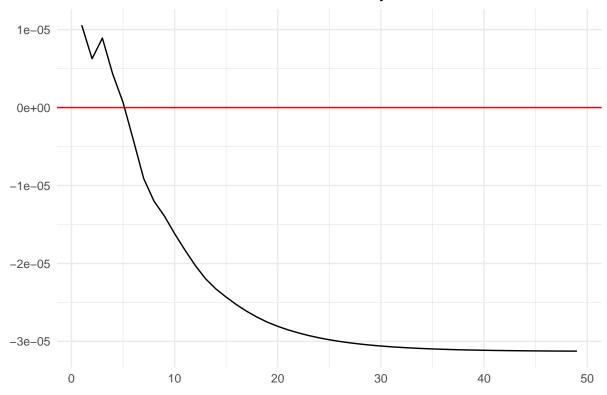
ylab("")+
xlab("") +
theme_minimal()

ASHR IRF of Trade on Volatility



```
ggplot(Yd4,aes(x=period, y=cumsum(response))) +
  geom_hline(yintercept = 0, color="red") +
  geom_line() +
  theme_light() +
  ggtitle("ASHR Cumulalitve IRF of trade on Volatility") +
  ylab("")+
  xlab("") +
  theme_minimal()
```





grangertest(y4[,c("vol","trade")], order = 6)

Res.Df	Df	F	Pr(>F)
2e+04			
2e+04	-6	5.36	1.55e-05

grangertest(y4[,c("trade", "vol")], order = 6)

Res.Df	Df	\mathbf{F}	$\Pr(>F)$
2e+04			
2e+04	-6	2.79	0.0103

China Mention

```
ychina = cbind(Vdata$china, Vdata$ASHR_vol)
colnames(ychina)[1:2] <- c("china", "vol")</pre>
est. VARchina <- VAR(ychina, p=6)
#extract results
mod_volchina = est.VARchina$varresult$vol
fchina = formula(mod_volchina)
dchina = model.frame(mod_volchina)
lm_cleanchina = lm(fchina, data= dchina)
#apply Newey-West
nw_vcovchina = NeweyWest(lm_cleanchina, lag=6)
nw_sechina = sqrt(diag(nw_vcovchina))
#t-stats
coefchina = coef(lm_cleanchina)
t_statchina = coefchina/nw_sechina
#recalculate p-values
robustchina = 2*(1-pt(abs(t_statchina), df = df.residual(lm_cleanchina)))
               <- nw_sechina[names(coef(lm_cleanchina))]</pre>
nw_sechina
robustchina
              <- robustchina[names(coef(lm_cleanchina))]</pre>
#table
screenreg(lm_cleanchina, override.se = nw_sechina, override.pvalues = robustchina, digits = 6)
##
## ===========
##
             Model 1
## china.l1
                -0.000005
               (0.000003)
## vol.l1
               0.280637 ***
                 (0.022879)
              -0.000004
## china.12
                (0.000004)
## vol.12
                 0.072261 *
                (0.030357)
## china.13
                -0.000011 ***
                 (0.000002)
## vol.13
                 0.045298 ***
                 (0.008336)
##
## china.14
                -0.000007 *
                 (0.000003)
##
## vol.14
                 0.056264 ***
##
                 (0.013530)
## china.15
                -0.000007 *
                 (0.000003)
##
## vol.15
                 0.056857 ***
##
                 (0.017212)
## china.16
                -0.000010 ***
                 (0.000003)
##
```

```
## vol.16
            0.109272 ***
##
                (0.023313)
## const
               0.000063 ***
               (0.000009)
##
## -----
## R^2
                0.266273
## Adj. R^2
             0.265795
## Num. obs. 19965
## ==========
## *** p < 0.001; ** p < 0.01; * p < 0.05
#Table for the effect of volatility on posts for variable china
#extract results
mod_postchina = est.VARchina$varresult$china
ffchina = formula(mod_postchina)
ddchina = model.frame(mod_postchina)
lm_clean_postchina = lm(ffchina, data= ddchina)
#apply Newey-West
nw_vcov_postchina = NeweyWest(lm_clean_postchina, lag=6)
nw_se_postchina = sqrt(diag(nw_vcov_postchina))
#t-stats
coef_postchina = coef(lm_clean_postchina)
t_stat_postchina = coef_postchina/nw_se_postchina
#recalculate p-values
robust_postchina = 2*(1-pt(abs(t_stat_postchina), df = df.residual(lm_clean_postchina)))
                   <- nw_se_postchina[names(coef(lm_clean_postchina))]</pre>
nw_se_postchina
                   <- robust_postchina[names(coef(lm_clean_postchina))]</pre>
robust_postchina
screenreg(lm_clean_postchina, override.se = nw_se_postchina, override.pvalues = robust_postchina, digit
## ==========
            Model 1
## -----
## china.l1 0.075875 * (0.034050)
## vol.l1
            -10.102292 ***
               (2.724874)
##
## china.12
               0.044876 *
##
               (0.021535)
## vol.12
                3.212025
               (7.103060)
                0.007296
## china.13
               (0.010495)
## vol.13
             10.034793
               (7.822179)
##
## china.14
              0.026711 **
```

(0.010004)

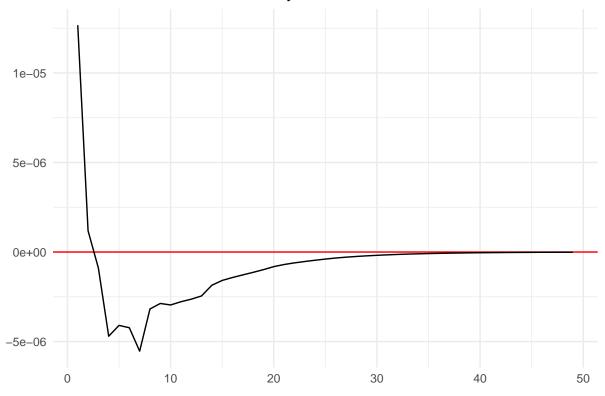
vol.14 -14.106424 ***

```
##
                 (4.258786)
                 0.048839
## china.15
##
                 (0.034024)
## vol.15
                 -4.365003
##
                (3.514978)
## china.16
                 0.054840
                 (0.048399)
## vol.16
                 3.977833
##
                 (4.376596)
## const
                 0.047264 ***
                 (0.006139)
## ----
## R^2
                  0.034788
## Adj. R^2
                  0.034159
## Num. obs. 19965
## ==========
## *** p < 0.001; ** p < 0.01; * p < 0.05
#Recreate a Robust Omega Matrix
Uchina = residuals(est.VARchina)
Tchina = nrow(Uchina)
Omegachina = matrix(0, ncol(Uchina), ncol(Uchina))
for(l in 0:L) {
  weight = 1 - 1/(L+1)
 Gamma_l_china = t(Uchina[(1+1):Tchina, , drop=FALSE]) %*% Uchina[1:(Tchina-l), , drop=FALSE] /Tchina
 if (1 == 0){
   Omegachina = Omegachina + Gamma_l_china
 } else {
   Omegachina = Omegachina + weight*(Gamma_l_china + t(Gamma_l_china))
  }
}
#make the B matrix
losschina <- function(paramchina){</pre>
  #Define the restriction
 Bchina <- matrix(c(paramchina[1], paramchina[2], 0, paramchina[3]), ncol = 2)
  #Make BB' approximatively equal to omega
 Xchina <- Omegachina - Bchina %*% t(Bchina)</pre>
  #loss function
 losschina <- sum(Xchina^2)</pre>
  return(losschina)
}
res.optchina <- optim(c(1, 0, 1), losschina, method = "BFGS")</pre>
B.hatchina <- matrix(c(res.optchina$par[1], res.optchina$par[2], 0, res.optchina$par[3]), ncol = 2)
print(cbind(Omegachina, B. hatchina %*% t(B. hatchina)))
                china
                               vol
## china 1.939575e-01 5.702559e-06 1.939570e-01 5.579026e-06
## vol 5.702559e-06 1.425509e-07 5.579026e-06 1.809801e-05
```

B.hatchina

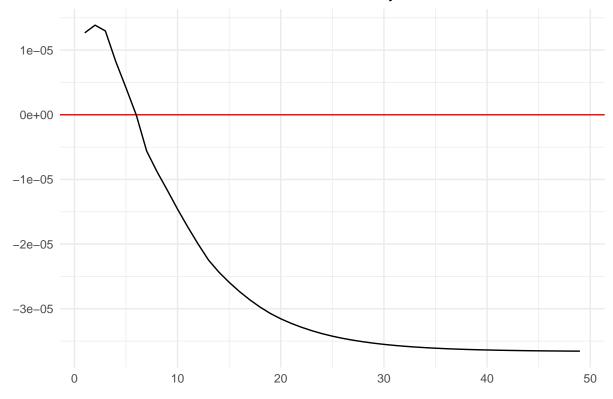
```
[,2]
##
                 [,1]
## [1,] 4.404055e-01 0.000000000
## [2,] 1.266793e-05 0.004254156
#get back the coefficient of est.VAR
phichina <- Acoef(est.VARchina)</pre>
PHIchina = make.PHI(phichina)
#take the constant
constantchina <- sapply(est.VARchina$varresult, function(eq) coef(eq)["const"])</pre>
cchina=as.matrix(constantchina)
#Simulate the IRF
pchina <- length(phichina)</pre>
nchina <- dim(phichina[[1]])[1]</pre>
Ychina <- simul.VAR(c=cchina, Phi = phichina, B = B.hatchina, nb.sim, y0.star=rep(0, nchina*pchina),
                  indic.IRF = 1, u.shock = c(1,0))
#Plot the IRF
Ydchina = data.frame(
  period = 1:nrow(Ychina),
  response = Ychina[,2])
ggplot(Ydchina,aes(x=period, y=response)) +
  geom_hline(yintercept = 0, color="red") +
  geom_line() +
  theme_light() +
  ggtitle("ASHR IRF of China on Volatility") +
  ylab("")+
  xlab("") +
  theme_minimal()
```

ASHR IRF of China on Volatility



```
ggplot(Ydchina,aes(x=period, y=cumsum(response))) +
  geom_hline(yintercept = 0, color="red") +
  geom_line() +
  theme_light() +
  ggtitle("ASHR Cumulalitve IRF of China on Volatility") +
  ylab("")+
  xlab("") +
  theme_minimal()
```





grangertest(ychina[,c("vol", "china")], order = 6)

Res.Df	Df	F	Pr(>F)
2e+04			
2e+04	-6	1.08	0.372

grangertest(ychina[,c("china", "vol")], order = 6)

Res.Df	Df	\mathbf{F}	$\Pr(>F)$
2e+04			
2e+04	-6	1.92	0.0733

Split Terms

Here we look for the first and second mandate effect of posts. We will use the tariff variable as a proxy for the posts.

```
# First and Second Mandate

#first term
Vdata_f = filter(data,between(timestamp, as.Date('2017-01-20'), as.Date('2021-01-20')))

#second term
Vdata_s = filter(data,between(timestamp, as.Date('2025-01-20'), as.Date('2025-05-07')))
```

First mandate

```
y_f_d = cbind(Vdata_f$dummy, Vdata_f$ASHR_vol)
colnames(y_f_d)[1:2] <- c("dummy", "vol")</pre>
est.VAR_f_d \leftarrow VAR(y_f_d,p=6)
#extract results
mod_vol_f_d = est.VAR_f_d$varresult$vol
f_f_d = formula(mod_vol_f_d)
d_f_d = model.frame(mod_vol_f_d)
lm_clean_f_d = lm(f_f_d, data = d_f_d)
#apply Newey-West
nw_vcov_f_d = NeweyWest(lm_clean_f_d, lag=6)
nw_se_f_d = sqrt(diag(nw_vcov_f_d))
#t-stats
coef f d = coef(lm clean f d)
t_stat_f_d = coef_f_d/nw_se_f_d
#recalculate p-values
robust_f_d = 2*(1-pt(abs(t_stat_f_d), df = df.residual(lm_clean_f_d)))
              <- nw_se_f_d[names(coef(lm_clean_f_d))]</pre>
nw se f d
              <- robust_f_d[names(coef(lm_clean_f_d))]</pre>
robust_f_d
#table
screenreg(lm_clean_f_d, override.se = nw_se_f_d, override.pvalues = robust_f_d, digits = 6)
======= Model 1
               - dummy.l1 -0.000003 (0.000001)
vol.l1 0.241914 (0.041504)
dummy.l2 -0.000003 (0.000001)
vol.12 0.073400 (0.019082)
dummy.l3 -0.000005 (0.000001)
vol.l3 0.065194 (0.023639)
dummy.l4 -0.000004 (0.00001)
vol.l4 0.064584 (0.020922)
dummy.l5 -0.000005 (0.000001)
vol.l5 0.090550 (0.022232)
dummy.l6 -0.000005 (0.000001)
vol.l6 0.176515 (0.059500)
const 0.000061 (0.000009)
```

```
Adj. R^2 0.349005
Num. obs. 7036
```

```
#Table for the effect of volatility on posts for variable dummy
#extract results
mod_post_f_d = est.VAR_f_d$varresult$dummy
ff_f_d = formula(mod_post_f_d)
dd_f_d = model.frame(mod_post_f_d)
lm_clean_post_f_d = lm(ff_f_d, data= dd_f_d)
#apply Newey-West
nw_vcov_post_f_d = NeweyWest(lm_clean_post_f_d, lag=6)
nw_se_post_f_d = sqrt(diag(nw_vcov_post_f_d))
#t-stats
coef_post_f_d = coef(lm_clean_post_f_d)
t_stat_post_f_d = coef_post_f_d/nw_se_post_f_d
#recalculate p-values
robust_post_f_d = 2*(1-pt(abs(t_stat_post_f_d), df = df.residual(lm_clean_post_f_d)))
                    <- nw_se_post_f_d[names(coef(lm_clean_post_f_d))]</pre>
nw_se_post_f_d
                    <- robust_post_f_d[names(coef(lm_clean_post_f_d))]</pre>
robust_post_f_d
#table
screenreg(lm_clean_post_f_d, override.se = nw_se_post_f_d, override.pvalues = robust_post_f_d, digits =
##
## ==========
##
             Model 1
## -----
              -0.130197 ***
## dummy.l1
##
                (0.006089)
## vol.l1
              68.970633
             (134.059501)
               -0.106720 ***
## dummy.12
##
                (0.005931)
## vol.12
             -251.984638
##
              (131.858784)
```

-0.090644 ***

-0.091087 ***

-0.115207 ***

(0.006611)

(0.006437)

-244.070250 (131.284080)

-316.430144 * (153.259332)

dummy.13

dummy.14

vol.14

dummy.15

vol.13

##

##

##

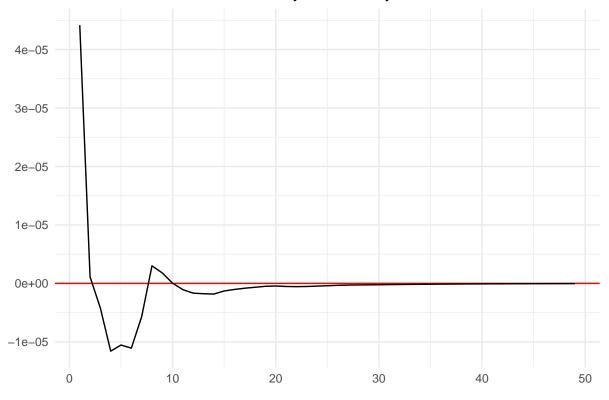
##

```
## dummy.16
              -0.138102 ***
##
               (0.007732)
## vol.16
              515.688924 *
             (223.552268)
##
## const
                1.980841 ***
##
               (0.063785)
## R^2
                0.184354
## Adj. R^2
                0.182844
## Num. obs. 7036
## ==========
## *** p < 0.001; ** p < 0.01; * p < 0.05
#Construct the Robust Omega Matrix
U_f_d = residuals(est.VAR_f_d)
T_f_d = nrow(U_f_d)
Omega_f_d = matrix(0, ncol(U_f_d), ncol(U_f_d))
for(l in 0:L) {
  weight = 1 - 1/(L+1)
 if (1 == 0){
   Omega_f_d = Omega_f_d + Gamma_l__f_d
 } else {
   Omega_f_d = Omega_f_d + weight*(Gamma_l__f_d + t(Gamma_l__f_d))
}
#make the B matrix
loss_f_d <- function(param_f_d){</pre>
  #Define the restriction
 B_f_d \leftarrow matrix(c(param_f_d[1], param_f_d[2], 0, param_f_d[3]), ncol = 2)
  #Make BB' approximatively equal to omega
 X_f_d \leftarrow Omega_f_d - B_f_d %*% t(B_f_d)
 #loss function
 loss_f_d <- sum(X_f_d^2)</pre>
 return(loss_f_d)
res.opt_f_d \leftarrow optim(c(1, 0, 1), loss_f_d, method = "BFGS")
B.hat_f_d \leftarrow matrix(c(res.opt_f_dpar[1], res.opt_f_dpar[2], 0, res.opt_f_dpar[3]), ncol = 2)
print(cbind(Omega_f_d,B.hat_f_d %*% t(B.hat_f_d)))
##
               dummy
                             vol
## dummy 9.9149258293 1.389427e-04 9.9149251733 1.390914e-04
## vol 0.0001389427 4.024686e-08 0.0001390914 2.429784e-06
B.hat_f_d
               [,1]
                          [,2]
##
```

```
## [1,] 3.148797e+00 0.00000000
## [2,] 4.417287e-05 0.00155815
```

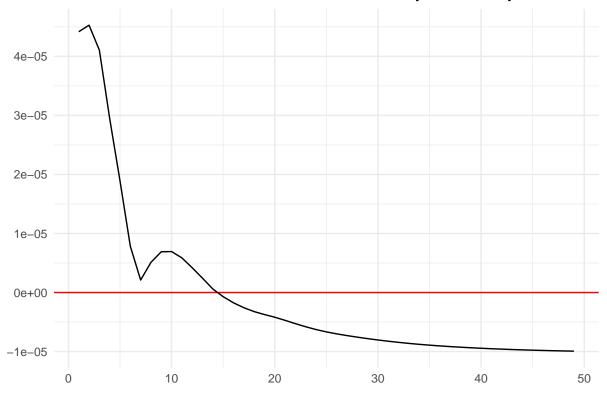
```
#get back the coefficient of est.VAR
phi_f_d <- Acoef(est.VAR_f_d)</pre>
PHI_f_d = make.PHI(phi_f_d)
#take the constant
constant_f_d <- sapply(est.VAR_f_d$varresult, function(eq) coef(eq)["const"])</pre>
c_f_d=as.matrix(constant_f_d)
#Simulate the IRF
p_f_d <- length(phi_f_d)</pre>
n_f_d <- dim(phi_f_d[[1]])[1]</pre>
Y_f_d \leftarrow simul.VAR(c=c_f_d, Phi = phi_f_d, B = B.hat_f_d, nb.sim ,y0.star=rep(0, n_f_d*p_f_d),
                   indic.IRF = 1, u.shock = c(1,0))
#Plot the IRF
Yd_f_d = data.frame(
  period = 1:nrow(Y_f_d),
 response = Y_f_d[,2])
ggplot(Yd_f_d,aes(x=period, y=response)) +
  geom_hline(yintercept = 0, color="red") +
  geom_line() +
  theme_light() +
  ggtitle("ASHR IRF of First Term Dummy on Volatility") +
  ylab("")+
  xlab("") +
  theme_minimal()
```

ASHR IRF of First Term Dummy on Volatility



```
ggplot(Yd_f_d,aes(x=period, y=cumsum(response))) +
  geom_hline(yintercept = 0, color="red") +
  geom_line() +
  theme_light() +
  ggtitle("ASHR Cumulalitve IRF of First Mandate Dummy on Volatility") +
  ylab("")+
  xlab("") +
  theme_minimal()
```

ASHR Cumulalitye IRF of First Mandate Dummy on Volatility



#does vol granger cause dummy
grangertest(y_f_d[,c("vol","dummy")], order =6)

Res.Df	Df	${f F}$	$\Pr(>F)$
7.02e+03			
7.03e+03	-6	4.09	0.000418

#does dummy granger cause vol
grangertest(y_f_d[,c("dummy", "vol")], order =6)

Res.Df	\mathbf{Df}	${f F}$	$\Pr(>F)$
7.02e+03			
7.03e+03	-6	15.1	2.67e-17

Second Mandate

```
y_s_d = cbind(Vdata_s$dummy, Vdata_s$ASHR_vol)
colnames(y_s_d)[1:2] \leftarrow c("dummy", "vol")
est.VAR_s_d \leftarrow VAR(y_s_d,p=6)
#extract results
mod_vol_s_d = est.VAR_s_d$varresult$vol
f_s_d = formula(mod_vol_s_d)
d_s_d = model.frame(mod_vol_s_d)
lm_clean_s_d = lm(f_s_d, data= d_s_d)
#apply Newey-West
nw_vcov_s_d = NeweyWest(lm_clean_s_d, lag=6)
nw_se_s_d = sqrt(diag(nw_vcov_s_d))
#t-stats
coef_s_d = coef(lm_clean_s_d)
t_stat_s_d = coef_s_d/nw_se_s_d
#recalculate p-values
robust_s_d = 2*(1-pt(abs(t_stat_s_d), df = df.residual(lm_clean_s_d)))
nw_se_s_d <- nw_se_s_d[names(coef(lm_clean_s_d))]</pre>
robust_s_d <- robust_s_d[names(coef(lm_clean_s_d))]
#table
screenreg(lm_clean_s_d, override.se = nw_se_s_d, override.pvalues = robust_s_d, digits = 6)
======== Model 1
           ——- dummy.ll -0.000002
(0.000003)
vol.l1 0.453713 ** (0.139092)
dummy.l2 -0.000003
(0.000002)
vol.l2 0.073292
(0.094132)
dummy.l3 -0.000005 (0.000002)
vol.13 -0.063668
(0.032167)
dummy.l4 -0.000002
(0.000002)
vol.l4 0.082677
(0.089570)
dummy.l5 -0.000006 (0.000002)
vol.15 -0.003739
(0.037991)
dummy.l6 -0.000006 ** (0.000002)
vol.16 0.111940
(0.073132)
const 0.000065 (0.000016)
 Adj. R^2 0.369822
Num. obs. 512
```

```
#Table for the effect of volatility on posts for variable dummy
#extract results
mod_post_s_d = est.VAR_s_d$varresult$dummy
ff_s_d = formula(mod_post_s_d)
dd_s_d = model.frame(mod_post_s_d)
lm_clean_post_s_d = lm(ff_s_d, data= dd_s_d)
#apply Newey-West
nw_vcov_post_s_d = NeweyWest(lm_clean_post_s_d, lag=6)
nw_se_post_s_d = sqrt(diag(nw_vcov_post_s_d))
#t-stats
coef_post_s_d = coef(lm_clean_post_s_d)
t_stat_post_s_d = coef_post_s_d/nw_se_post_s_d
#recalculate p-values
robust_post_s_d = 2*(1-pt(abs(t_stat_post_s_d), df = df.residual(lm_clean_post_s_d)))
                     <- nw_se_post_s_d[names(coef(lm_clean_post_s_d))]</pre>
nw_se_post_s_d
                     <- robust_post_s_d[names(coef(lm_clean_post_s_d))]</pre>
robust_post_s_d
#table
screenreg(lm_clean_post_s_d, override.se = nw_se_post_s_d, override.pvalues = robust_post_s_d, digits =
##
## =========
##
             Model 1
## -----
              -0.215778 ***
## dummy.l1
##
                (0.021070)
## vol.11
             -212.214190
              (241.235993)
## dummy.12
               -0.204705 ***
                (0.020905)
##
## vol.12
              -377.569865
              (221.982942)
## dummy.13
               -0.208529 ***
##
                (0.022295)
              527.218353
## vol.13
##
              (421.359653)
## dummy.14
                -0.213563 ***
##
                (0.022187)
## vol.14
              -450.277869 *
              (193.512623)
##
## dummy.15
                -0.214362 ***
##
                (0.021926)
## vol.15
              -151.169780
##
              (262.979738)
## dummy.16
               -0.200965 ***
```

##

const

vol.16

(0.022394)

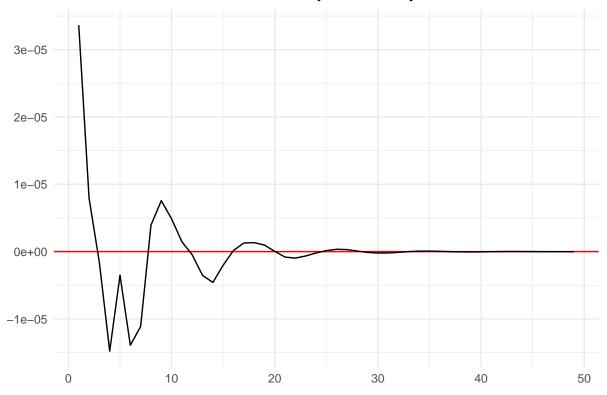
3.172941 ***

222.174258 (306.804126)

```
##
               (0.258307)
## -----
## R^2
               0.297405
## Adj. R^2
               0.279101
## Num. obs. 512
## =========
## *** p < 0.001; ** p < 0.01; * p < 0.05
#Construct the Robust Omega Matrix
U_s_d = residuals(est.VAR_s_d)
T_s_d = nrow(U_s_d)
Omega_s_d = matrix(0, ncol(U_s_d), ncol(U_s_d))
for(1 in 0:L) {
  weight = 1 - 1/(L+1)
 if (1 == 0){
   Omega_s_d = Omega_s_d + Gamma_l_s_d
 } else {
   Omega_s_d = Omega_s_d + weight*(Gamma_l_s_d + t(Gamma_l_s_d))
 }
}
#make the B matrix
loss_s_d <- function(param_s_d){</pre>
  #Define the restriction
 B_s_d \leftarrow matrix(c(param_s_d[1], param_s_d[2], 0, param_s_d[3]), ncol = 2)
 #Make BB' approximatively equal to omega
 X_s_d \leftarrow Omega_s_d - B_s_d %*% t(B_s_d)
 #loss function
 loss_s_d \leftarrow sum(X_s_d^2)
 return(loss s d)
}
res.opt_s_d <- optim(c(1, 0, 1), loss_s_d, method = "BFGS")
B.hat_s_d <- matrix(c(res.opt_s_dpar[1], res.opt_s_dpar[2], 0, res.opt_s_dpar[3]), ncol = 2)
print(cbind(Omega_s_d,B.hat_s_d %*% t(B.hat_s_d)))
##
              dummy
                            vol
## dummy 9.8009866129 1.052217e-04 9.800984276 1.05317e-04
## vol
        0.0001052217 3.718029e-08 0.000105317 3.20576e-05
B.hat s d
                          [,2]
##
               [,1]
## [1,] 3.130652e+00 0.000000000
## [2,] 3.364058e-05 0.005661844
```

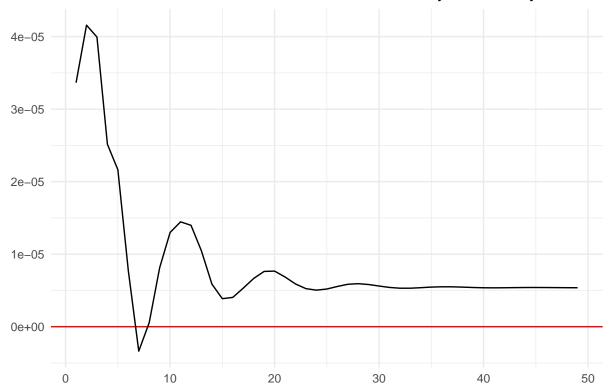
```
\#get\ back\ the\ coefficient\ of\ est.VAR
phi_s_d <- Acoef(est.VAR_s_d)</pre>
PHI_s_d = make.PHI(phi_s_d)
#take the constant
constant_s_d <- sapply(est.VAR_s_d$varresult, function(eq) coef(eq)["const"])</pre>
c_s_d=as.matrix(constant_s_d)
#Simulate the IRF
p_s_d <- length(phi_s_d)</pre>
n_s_d <- dim(phi_s_d[[1]])[1]</pre>
Y_s_d <- simul.VAR(c=c_s_d, Phi = phi_s_d, B = B.hat_s_d, nb.sim ,y0.star=rep(0, n_s_d*p_s_d),
                   indic.IRF = 1, u.shock = c(1,0))
#Plot the IRF
Yd_s_d = data.frame(
 period = 1:nrow(Y_s_d),
 response = Y_s_d[,2])
ggplot(Yd_s_d,aes(x=period, y=response)) +
  geom_hline(yintercept = 0, color="red") +
  geom_line() +
  theme_light() +
  ggtitle("ASHR IRF of Second Term Dummy on Volatility") +
  ylab("")+
 xlab("") +
 theme_minimal()
```

ASHR IRF of Second Term Dummy on Volatility



```
ggplot(Yd_s_d,aes(x=period, y=cumsum(response))) +
  geom_hline(yintercept = 0, color="red") +
  geom_line() +
  theme_light() +
  ggtitle("ASHR Cumulalitve IRF of Second Mandate Dummy on Volatility") +
  ylab("") +
  xlab("") +
  theme_minimal()
```

ASHR Cumulalitve IRF of Second Mandate Dummy on Volatility



#does vol granger cause dummy
grangertest(y_s_d[,c("vol","dummy")], order =6)

Res.Df	Df	${f F}$	$\Pr(>F)$
499			
505	-6	0.26	0.955

#does dummy granger cause vol
grangertest(y_s_d[,c("dummy", "vol")], order =6)

Res.Df	\mathbf{Df}	${f F}$	$\Pr(>F)$
499			
505	-6	1.01	0.418