SPY SVAR Models

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Setup

Load packages & functions

```
rm(list=ls())
require(tinytex) #LaTeX
require(ggplot2) #plots
require(AEC) #JP-Renne functions
require(AER) #NW formula
require(forecast) #time series stuff
require(expm) #matrix exponents
require(here) #directory finder
require(stringr) # analysis of strings, important for the detection in tweets
require(dplyr) #data management
require(lubridate) #data dates management
require(zoo) #for lagging
require(jtools) #tables
require(huxtable) #tables
require(lmtest) #reg tests
require(vroom) #for loading data
require(data.table) #for data filtering
require(sysid) #for ARMA-X modeling
require(sandwhich) #regression errors
require(stargazer) #nice req tables
require(tidytext) #text mining
require(textstem) #lemmatization
require(quanteda) #tokenization
require(texreg) #arima tables
require(vars) #VAR models
require(xts) #time series objects
require(tseries) #includes adf test
require(quantmod)
require(TSA)
require(aTSA)
require(tibble)
require(FinTS)
require(kableExtra)
require(writexl)
require(purrr)
getwd()
#setwd("...") -> set wd at base repo folder
#load helper functions
source(here("helperfunctions/data loaders.R"))
source(here("helperfunctions/date_selector.R"))
source(here("helperfunctions/plotters.R"))
source(here("helperfunctions/quick_arma.R"))
source(here("helperfunctions/r.vol_calculators.R"))
source(here("helperfunctions/truths cleaning function.R"))
source(here("helperfunctions/armax_functions.R"))
source(here("helperfunctions/var_irf.R"))
```

Load Data

```
#load final dataset
source(here("helperfunctions/full_data.R"))

#select timeframe
Vdata = filter(data,between(timestamp, as.Date('2014-01-01'), as.Date('2025-05-07')))
```

Some SVAR estimations

Note that this is not an exhaustive list of our VAR estimations, you can find more by going on /modeling/VAR/VAR_SPY_FULLMODELS.rmd or VAR_ASHR_FULLMODELS.rmd or VAR_VGK_FULLMODELS.rmd).

Dummy variable

Here we use a dummy variable which equal to one if Trump has made a post or 0 otherwise, taking into account the closed hour market posts.

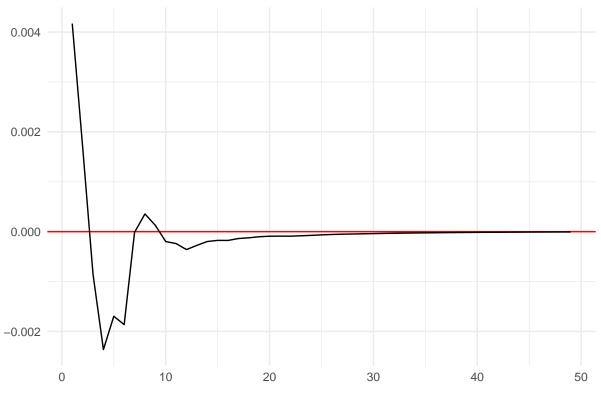
```
y = cbind(Vdata$dummy, Vdata$SPY_vol)
colnames(y)[1:2] <- c("dummy", "vol")</pre>
est.VAR <- VAR(y,p=6)
#extract results
mod_vol = est.VAR$varresult$vol
f = formula(mod_vol)
d = model.frame(mod_vol)
lm_clean = lm(f, data = d)
#apply Newey-West
nw_vcov = NeweyWest(lm_clean, lag=6)
nw_se = sqrt(diag(nw_vcov))
\#t\text{-}stats
coef = coef(lm clean)
t_stat = coef/nw_se
#recalculate p-values
robust = 2*(1-pt(abs(t_stat), df = df.residual(lm_clean)))
            <- nw_se[names(coef(lm_clean))]</pre>
nw_se
robust
            <- robust[names(coef(lm_clean))]</pre>
#table
screenreg(lm_clean, override.se = nw_se, override.pvalues = robust, digits = 6)
```

```
0.000083
## dummy.l1
##
                 (0.000201)
## vol.11
                 0.344511 ***
                 (0.103790)
##
## dummy.12
                 -0.000473 ***
##
                 (0.000071)
## vol.12
                 0.023714
                 (0.042739)
##
## dummy.13
                 -0.000804 ***
                 (0.000088)
##
## vol.13
                 0.082941 ***
##
                 (0.007496)
## dummy.14
                 -0.000546 ***
                 (0.000088)
##
## vol.14
                 0.096948
##
                 (0.059298)
## dummy.15
                 -0.000579 ***
                 (0.000147)
## vol.15
                 0.022887 ***
##
                 (0.006876)
## dummy.16
                 -0.000099
                 (0.000101)
## vol.16
                0.164034 ***
##
                 (0.047379)
## const
                 0.008726 ***
                 (0.001609)
## --
## R^2
                  0.325745
## Adj. R^2
                  0.325306
## Num. obs. 19965
## ===========
## *** p < 0.001; ** p < 0.01; * p < 0.05
#extract results
mod_post = est.VAR$varresult$dummy
ff = formula(mod_post)
dd = model.frame(mod_post)
lm_clean_post = lm(ff, data= dd)
#apply Newey-West
nw_vcov_post = NeweyWest(lm_clean_post, lag=6)
nw_se_post = sqrt(diag(nw_vcov_post))
#t-stats
coef_post = coef(lm_clean_post)
t_stat_post = coef_post/nw_se_post
\#recalculate\ p\mbox{-}values
robust_post = 2*(1-pt(abs(t_stat_post), df = df.residual(lm_clean_post)))
nw se post
               <- nw se post[names(coef(lm clean post))]</pre>
robust_post
               <- robust_post[names(coef(lm_clean_post))]</pre>
#table
```

```
##
## ===========
            Model 1
## -----
## dummy.11
               -0.093610 ***
##
               (0.003418)
## vol.l1
               0.410764
               (0.303756)
##
## dummy.12
               -0.085915 ***
##
               (0.003447)
## vol.12
             -0.558043
##
               (0.306897)
## dummy.13
               -0.076436 ***
##
               (0.003627)
               -0.269652
## vol.13
##
               (0.350524)
## dummy.14
               -0.077830 ***
##
               (0.003667)
## vol.14
             -0.719476 *
               (0.300555)
##
## dummy.15
               -0.087761 ***
##
               (0.003688)
## vol.15
               -0.276823
               (0.169338)
               -0.096838 ***
## dummy.16
               (0.003939)
              0.973172
## vol.16
##
               (0.525025)
               1.721194 ***
## const
               (0.039256)
## -----
## R^2
               0.155107
## Adj. R^2
               0.154556
## Num. obs. 19965
## ==========
## *** p < 0.001; ** p < 0.01; * p < 0.05
#HO test whether there is NOT heteroscedasticity. if less by alpha, then there is heteroscedasticity
bptest(lm_clean)
##
## studentized Breusch-Pagan test
## data: lm_clean
## BP = 386.02, df = 12, p-value < 2.2e-16
#Recreate a Robust Omega Matrix
U = residuals(est.VAR)
T = nrow(U)
L = 6 #number of lag
```

```
Omega = matrix(0, ncol(U), ncol(U))
for(l in 0:L) {
  weight = 1 - 1/(L+1)
  Gamma_l = t(U[(1+1):T, , drop=FALSE]) %*% U[1:(T-1), , drop=FALSE] /T
  if (1 == 0){
    Omega = Omega + Gamma_1_
  } else {
    Omega = Omega + weight*(Gamma_l_ + t(Gamma_l_))
  }
}
#make the B matrix
loss <- function(param){</pre>
  #Define the restriction
  B \leftarrow matrix(c(param[1], param[2], 0, param[3]), ncol = 2)
  #Make BB' approximatively equal to omega
  X <- Omega - B %*% t(B)
  #loss function
 loss <- sum(X^2)
  return(loss)
}
res.opt \leftarrow optim(c(1, 0, 1), loss, method = "BFGS")
B.hat <- matrix(c(res.opt$par[1], res.opt$par[2], 0, res.opt$par[3]), ncol = 2)
print(cbind(Omega,B.hat %*% t(B.hat)))
##
                dummy
                              vol
## dummy 10.24415441 0.013344869 10.24415377 0.013344867
          0.01334487 0.005298555 0.01334487 0.005297573
## vol
B.hat
##
                [,1]
                             [,2]
## [1,] 3.200648961 0.00000000
## [2,] 0.004169425 -0.07266491
nb.sim = 7*7
#get back the coefficient of est.VAR
phi <- Acoef(est.VAR)</pre>
PHI = make.PHI(phi)
#take the constant
constant <- sapply(est.VAR$varresult, function(eq) coef(eq)["const"])</pre>
c=as.matrix(constant)
#Simulate the IRF
p <- length(phi)</pre>
n <- dim(phi[[1]])[1]</pre>
```

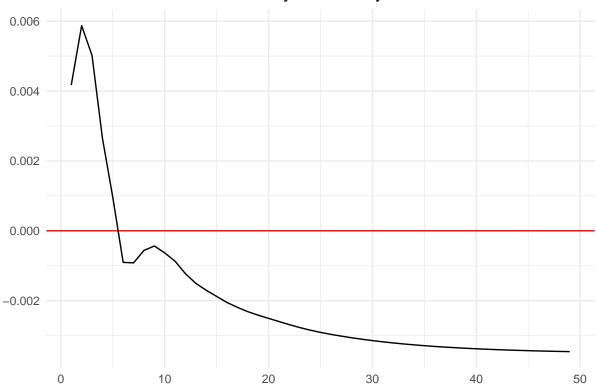
S&P IRF of Dummy on Volatility



```
ggplot(Yd,aes(x=period, y=cumsum(response))) +
  geom_hline(yintercept = 0, color="red") +
  geom_line() +
  theme_light() +
  ggtitle("S&P Cumulality IRF of Dummy on Volatility") +
  ylab("")+
```

```
xlab("") +
theme_minimal()
```

S&P Cumulalitye IRF of Dummy on Volatility



Post Counts

```
y2 = cbind(Vdata$N, Vdata$SPY_vol)
colnames(y2)[1:2] <- c("N", "vol")
est.VAR2 <- VAR(y2,p=6)

#extract results
mod_vol2 = est.VAR2$varresult$vol
f2 = formula(mod_vol2)
d2 = model.frame(mod_vol2)
lm_clean2 = lm(f2, data= d2)

#apply Newey-West
nw_vcov2 = NeweyWest(lm_clean2, lag=6)
nw_se2 = sqrt(diag(nw_vcov2))

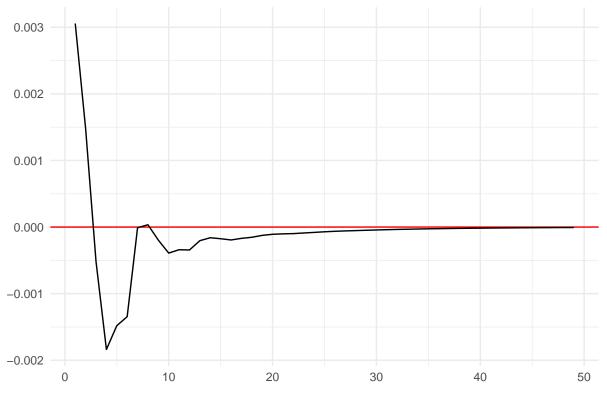
#t-stats
coef2 = coef(lm_clean2)
t_stat2 = coef2/nw_se2</pre>
```

```
#recalculate p-values
robust2 = 2*(1-pt(abs(t_stat2), df = df.residual(lm_clean2)))
          <- nw_se2[names(coef(lm_clean2))]</pre>
          <- robust2[names(coef(lm_clean2))]</pre>
robust2
#table
screenreg(lm_clean2, override.se = nw_se2, override.pvalues = robust2, digits = 6)
##
## ==========
           Model 1
             0.000045
## N.l1
              (0.000037)
##
## vol.11
             0.345011 ***
##
              (0.104492)
              -0.000116 ***
## N.12
               (0.000023)
##
                0.023575
## vol.12
##
              (0.043816)
## N.13
              -0.000213 ***
                (0.000028)
##
## vol.13
              0.082525 ***
##
                (0.008145)
## N.14
               -0.000147 ***
##
               (0.000021)
## vol.14
                0.096739
              (0.060827)
              -0.000119 **
## N.15
##
                (0.000041)
## vol.15
               0.022593 **
               (0.006952)
## N.16
                0.000000
               (0.000028)
               0.164442 ***
## vol.16
                (0.049763)
                0.007587 ***
## const
##
                (0.001578)
## ---
             _____
## R^2
                0.325324
                0.324885
## Adj. R^2
## Num. obs. 19965
## ===========
## *** p < 0.001; ** p < 0.01; * p < 0.05
#extract results
mod_post2 = est.VAR2$varresult$N
ff2 = formula(mod_post2)
dd2 = model.frame(mod_post2)
lm_clean_post2 = lm(ff2, data= dd2)
#apply Newey-West
```

```
nw_vcov_post2 = NeweyWest(lm_clean_post2, lag=6)
nw_se_post2 = sqrt(diag(nw_vcov_post2))
#t-stats
coef_post2 = coef(lm_clean_post2)
t_stat_post2 = coef_post2/nw_se_post2
#recalculate p-values
robust_post2 = 2*(1-pt(abs(t_stat_post2), df = df.residual(lm_clean_post2)))
nw_se_post2
               - nw_se_post2[names(coef(lm_clean_post2))]
robust_post2 <- robust_post2[names(coef(lm_clean_post2))]</pre>
#table
screenreg(lm_clean_post2, override.se = nw_se_post2, override.pvalues = robust_post2, digits = 6)
##
## ===========
##
           Model 1
## N.l1
              -0.043917 ***
##
                (0.003632)
               0.898680
## vol.l1
               (0.753418)
## N.12
                -0.039092 ***
                (0.004198)
               -0.990623
## vol.12
                (0.742365)
## N.13
               -0.025847 ***
##
                (0.004288)
## vol.13
               -1.241543
##
               (0.872900)
                -0.023845 ***
## N.14
##
               (0.004990)
## vol.14
               -1.830204 *
                (0.745455)
##
## N.15
                -0.040694 ***
##
                (0.003975)
## vol.15
                -0.683162
                (0.469187)
##
## N.16
                -0.045066 ***
##
                (0.004790)
                2.727809
## vol.16
                (1.511411)
##
## const
                3.531650 ***
##
                (0.092709)
## -----
## R^2
                 0.100807
## Adj. R^2
                0.100221
## Num. obs. 19965
## ============
## *** p < 0.001; ** p < 0.01; * p < 0.05
```

```
#Recreate a Robust Omega Matrix
U2 = residuals(est.VAR2)
T2 = nrow(U2)
Omega2 = matrix(0, ncol(U2), ncol(U2))
for(l in 0:L) {
  weight = 1 - 1/(L+1)
  Gamma_1_2 = t(U2[(1+1):T2, , drop=FALSE]) %*% U2[1:(T2-1), , drop=FALSE] /T2
  if (1 == 0){
   Omega2 = Omega2 + Gamma_1_2
 } else {
    Omega2 = Omega2 + weight*(Gamma_1_2 + t(Gamma_1_2))
}
#make the B matrix
loss2 <- function(param2){</pre>
  #Define the restriction
  B2 \leftarrow matrix(c(param2[1], param2[2], 0, param2[3]), ncol = 2)
  #Make BB' approximatively equal to omega
  X2 <- Omega2 - B2 %*% t(B2)
  #loss function
 loss2 \leftarrow sum(X2^2)
  return(loss2)
}
res.opt2 \leftarrow optim(c(1, 0, 1), loss2, method = "BFGS")
B.hat2 \leftarrow matrix(c(res.opt2$par[1], res.opt2$par[2], 0, res.opt2$par[3]), ncol = 2)
print(cbind(Omega2,B.hat2 %*% t(B.hat2)))
##
                           vol
## N 86.4367215 0.028419903 86.43672134 0.028419567
## vol 0.0284199 0.005294799 0.02841957 0.005293894
B.hat2
##
                [,1]
                           [,2]
## [1,] 9.297135115 0.00000000
## [2,] 0.003056809 0.07269491
#get back the coefficient of est.VAR
phi2 <- Acoef(est.VAR2)</pre>
PHI2 = make.PHI(phi2)
#take the constant
constant2 <- sapply(est.VAR2$varresult, function(eq) coef(eq)["const"])</pre>
c2=as.matrix(constant2)
#Simulate the IRF
p2 <- length(phi2)
```

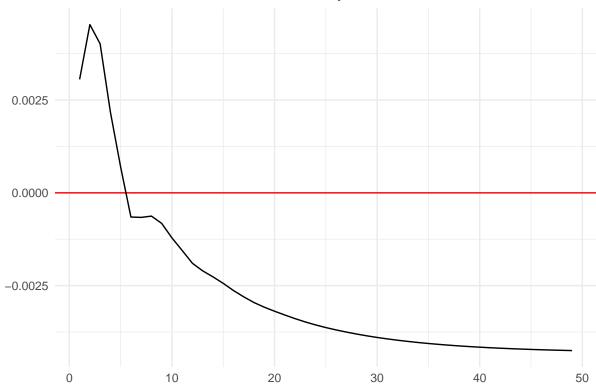
S&P IRF of N on Volatility



```
ggplot(Yd2,aes(x=period, y=cumsum(response))) +
  geom_hline(yintercept = 0, color="red") +
  geom_line() +
  theme_light() +
  ggtitle("S&P Cumulalitye IRF of N on Volatility") +
```

```
ylab("")+
xlab("") +
theme_minimal()
```

S&P Cumulalitve IRF of N on Volatility



Res.Df	Df	\mathbf{F}	$\Pr(>F)$
2e+04			
2e+04	-6	3.92	0.000646

grangertest(y2[,c("vol", "N")], order = 6)

Res.Df	Df	\mathbf{F}	$\Pr(>F)$
2e+04			
2e+04	-6	3.89	0.000688

Trade Mention

```
y4 = cbind(Vdata$trade, Vdata$SPY_vol)
colnames(y4)[1:2] <- c("trade", "vol")</pre>
est.VAR4 \leftarrow VAR(y4,p=6)
#extract results
mod_vol4 = est.VAR4$varresult$vol
f4 = formula(mod_vol4)
d4 = model.frame(mod_vol4)
lm_clean4 = lm(f4, data = d4)
#apply Newey-West
nw_vcov4 = NeweyWest(lm_clean4, lag=6)
nw_se4 = sqrt(diag(nw_vcov4))
#t-stats
coef4 = coef(lm_clean4)
t_stat4 = coef4/nw_se4
#recalculate p-values
robust4 = 2*(1-pt(abs(t_stat4), df = df.residual(lm_clean4)))
            <- nw_se4[names(coef(lm_clean4))]</pre>
{\tt nw\_se4}
            <- robust4[names(coef(lm_clean4))]</pre>
robust4
#table
screenreg(lm_clean4, override.se = nw_se4, override.pvalues = robust4, digits = 6)
##
```

```
## ==========
##
        Model 1
## -----
## trade.l1
             0.003399
##
              (0.003747)
              0.346107 ***
## vol.11
              (0.101918)
## trade.12
               0.005600
##
              (0.004809)
## vol.12
               0.022949
##
               (0.041538)
## trade.13
              -0.003904 *
               (0.001726)
##
## vol.13
               0.081148 ***
##
               (0.008258)
## trade.14
               0.000725
##
               (0.003458)
## vol.14
               0.095797
               (0.057082)
##
## trade.15
              -0.002363
##
               (0.001901)
## vol.15
               0.023502 **
               (0.007162)
##
```

```
## trade.16 -0.001543
##
               (0.001228)
## vol.16
              0.165323 ***
              (0.049319)
##
## const
               0.005939 ***
##
               (0.001536)
## -----
## R^2
               0.325134
## Adj. R^2 0.324695
## Num. obs. 19965
## ==========
## *** p < 0.001; ** p < 0.01; * p < 0.05
#Table for the effect of volatility on posts for variable trade
#extract results
mod_post4 = est.VAR4$varresult$trade
ff4 = formula(mod_post4)
dd4 = model.frame(mod_post4)
lm_clean_post4 = lm(ff4, data = dd4)
#apply Newey-West
nw_vcov_post4 = NeweyWest(lm_clean_post4, lag=6)
nw_se_post4 = sqrt(diag(nw_vcov_post4))
#t-stats
coef_post4 = coef(lm_clean_post4)
t_stat_post4 = coef_post4/nw_se_post4
#recalculate p-values
robust_post4 = 2*(1-pt(abs(t_stat_post4), df = df.residual(lm_clean_post4)))
             <- nw_se_post4[names(coef(lm_clean_post4))]</pre>
nw_se_post4
               <- robust_post4[names(coef(lm_clean_post4))]</pre>
robust_post4
screenreg(lm_clean_post4, override.se = nw_se_post4, override.pvalues = robust_post4, digits = 6)
## ===========
           Model 1
## -----
## trade.l1
              0.025465
               (0.018216)
##
## vol.l1
             0.023385
               (0.038889)
##
## trade.12
               0.019892 *
##
               (0.008655)
## vol.12
               0.140914 *
              (0.057926)
## trade.13
             0.022679
##
               (0.016472)
             -0.077822 ***
## vol.13
##
              (0.009836)
              0.021687
## trade.14
```

```
##
                 (0.012853)
## vol.14
                -0.037637 **
                 (0.013028)
##
## trade.15
                 0.019364
##
                 (0.010830)
## vol.15
                -0.019644
                 (0.020558)
## trade.16
                 0.012565
##
                 (0.010365)
## vol.16
                 0.107182
                 (0.070029)
                 0.027409 ***
## const
                 (0.002344)
## -----
## R^2
                  0.020264
## Adj. R^2
                  0.019626
## Num. obs. 19965
## ===========
## *** p < 0.001; ** p < 0.01; * p < 0.05
#Recreate a Robust Omega Matrix
U4 = residuals(est.VAR4)
T4 = nrow(U4)
Omega4 = matrix(0, ncol(U4), ncol(U4))
for(l in 0:L) {
  weight = 1 - 1/(L+1)
  Gamma_1_4 = t(U4[(1+1):T4, , drop=FALSE]) %*% U4[1:(T4-1), , drop=FALSE] /T4
  if (1 == 0){
    Omega4 = Omega4 + Gamma_1_4
  } else {
    Omega4 = Omega4 + weight*(Gamma_l_4 + t(Gamma_l_4))
}
#make the B matrix
loss4 <- function(param4){</pre>
  #Define the restriction
 B4 <- matrix(c(param4[1], param4[2], 0, param4[3]), ncol = 2)
  #Make BB' approximatively equal to omega
  X4 <- Omega4 - B4 %*% t(B4)
  #loss function
  loss4 \leftarrow sum(X4^2)
  return(loss4)
}
res.opt4 \leftarrow optim(c(1, 0, 1), loss4, method = "BFGS")
B.hat4 <- matrix(c(res.opt4\$par[1], res.opt4\$par[2], 0, res.opt4\$par[3]), ncol = 2)
print(cbind(Omega4,B.hat4 %*% t(B.hat4)))
```

vol

##

trade

```
## trade 8.203633e-02 6.209714e-05 8.203534e-02 6.210009e-05
        6.209714e-05 5.281703e-03 6.210009e-05 5.280741e-03
B.hat4
                [,1]
                            [,2]
## [1,] 0.2864181131 0.00000000
## [2,] 0.0002168162 0.07266838
#get back the coefficient of est.VAR
phi4 <- Acoef(est.VAR4)</pre>
PHI4 = make.PHI(phi4)
#take the constant
constant4 <- sapply(est.VAR4$varresult, function(eq) coef(eq)["const"])</pre>
c4=as.matrix(constant4)
#Simulate the IRF
p4 <- length(phi4)
n4 <- dim(phi4[[1]])[1]
Y4 <- simul.VAR(c=c4, Phi = phi4, B = B.hat4, nb.sim ,y0.star=rep(0, n4*p4),
                  indic.IRF = 1, u.shock = c(1,0))
#Plot the IRF
Yd4 = data.frame(
```

period = 1:nrow(Y4),
response = Y4[,2])

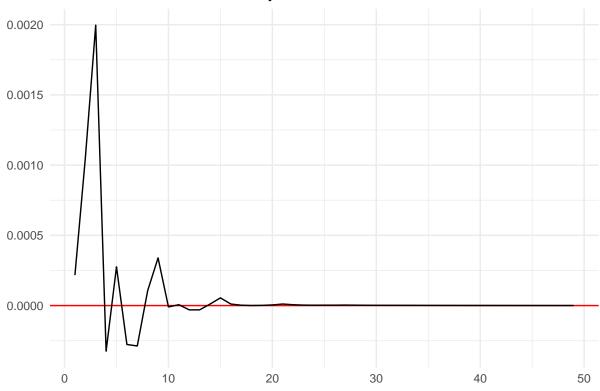
geom_line() +
theme_light() +

ylab("")+
xlab("") +
theme_minimal()

ggplot(Yd4,aes(x=period, y=response)) +
 geom_hline(yintercept = 0, color="red") +

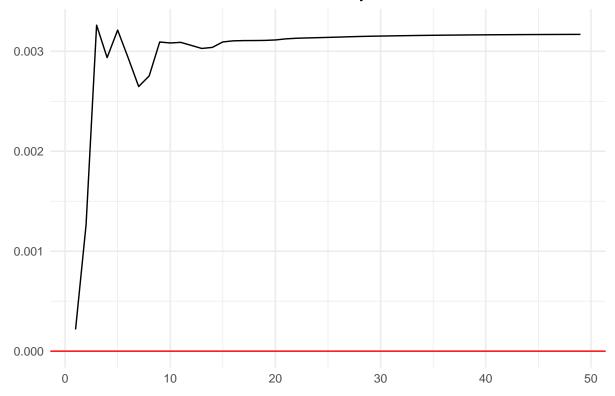
ggtitle("S&P IRF of Trade on Volatility") +

S&P IRF of Trade on Volatility



```
ggplot(Yd4,aes(x=period, y=cumsum(response))) +
  geom_hline(yintercept = 0, color="red") +
  geom_line() +
  theme_light() +
  ggtitle("S&P Cumulalitve IRF of trade on Volatility") +
  ylab("")+
  xlab("") +
  theme_minimal()
```

S&P Cumulalitve IRF of trade on Volatility



grangertest(y4[,c("vol","trade")], order = 6)

Res.Df	Df	F	Pr(>F)
2e+04			
2e+04	-6	10.6	9.06e-12

grangertest(y4[,c("trade", "vol")], order = 6)

Res.Df	Df	\mathbf{F}	$\Pr(>F)$
2e+04			
2e+04	-6	2.98	0.00655

China Mention

```
ychina = cbind(Vdata$china, Vdata$SPY_vol)
colnames(ychina)[1:2] <- c("china", "vol")</pre>
est. VARchina <- VAR(ychina, p=6)
#extract results
mod_volchina = est.VARchina$varresult$vol
fchina = formula(mod_volchina)
dchina = model.frame(mod_volchina)
lm_cleanchina = lm(fchina, data= dchina)
#apply Newey-West
nw_vcovchina = NeweyWest(lm_cleanchina, lag=6)
nw_sechina = sqrt(diag(nw_vcovchina))
#t-stats
coefchina = coef(lm_cleanchina)
t_statchina = coefchina/nw_sechina
#recalculate p-values
robustchina = 2*(1-pt(abs(t_statchina), df = df.residual(lm_cleanchina)))
              <- nw_sechina[names(coef(lm_cleanchina))]</pre>
nw_sechina
robustchina
              <- robustchina[names(coef(lm_cleanchina))]</pre>
#table
screenreg(lm_cleanchina, override.se = nw_sechina, override.pvalues = robustchina, digits = 6)
##
## ===========
##
             Model 1
## china.l1
                0.006729
               (0.006694)
## vol.l1
               0.344512 ***
                (0.097994)
## china.12
               0.002778
                (0.004067)
## vol.12
                 0.024149
                (0.043585)
## china.13
                -0.004652 *
                (0.002066)
## vol.13
                0.081646 ***
                 (0.009192)
##
## china.14
                -0.002442 *
                 (0.001084)
##
## vol.14
                 0.094919
##
                 (0.058821)
## china.15
                -0.000607
                 (0.000970)
##
## vol.15
                 0.022961 **
                 (0.007678)
##
## china.16
                 0.000596
                (0.000981)
##
```

```
## vol.16
            0.166695 **
##
               (0.054194)
## const
               0.005857 ***
##
               (0.001612)
## -----
## R^2
                0.326344
## Adj. R^2
             0.325905
## Num. obs. 19965
## ==========
## *** p < 0.001; ** p < 0.01; * p < 0.05
#Table for the effect of volatility on posts for variable china
#extract results
mod_postchina = est.VARchina$varresult$china
ffchina = formula(mod_postchina)
ddchina = model.frame(mod_postchina)
lm_clean_postchina = lm(ffchina, data= ddchina)
#apply Newey-West
nw_vcov_postchina = NeweyWest(lm_clean_postchina, lag=6)
nw_se_postchina = sqrt(diag(nw_vcov_postchina))
#t-stats
coef_postchina = coef(lm_clean_postchina)
t_stat_postchina = coef_postchina/nw_se_postchina
#recalculate p-values
robust_postchina = 2*(1-pt(abs(t_stat_postchina), df = df.residual(lm_clean_postchina)))
                   <- nw_se_postchina[names(coef(lm_clean_postchina))]</pre>
nw_se_postchina
                   <- robust_postchina[names(coef(lm_clean_postchina))]</pre>
robust_postchina
screenreg(lm_clean_postchina, override.se = nw_se_postchina, override.pvalues = robust_postchina, digit
## ==========
            Model 1
## -----
## china.l1
               0.074471 *
              (0.034456)
## vol.l1
             0.028311
               (0.039275)
##
## china.12
             0.044301 *
##
               (0.021635)
## vol.12
                0.138351
               (0.165247)
## china.13
                0.005765
               (0.010638)
              0.140252
## vol.13
##
               (0.202759)
## china.14
              0.024636 **
```

(0.009414)

-0.108103

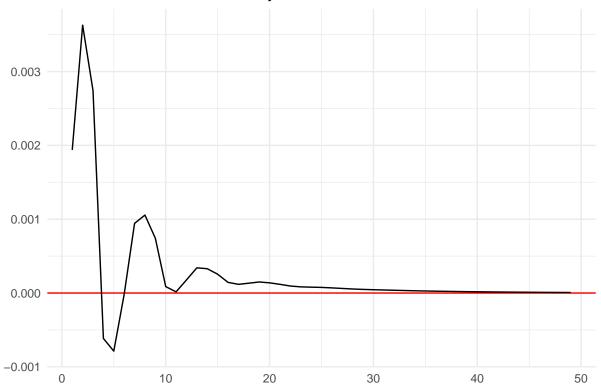
vol.14

```
##
                 (0.064202)
                 0.048992
## china.15
##
                 (0.034173)
## vol.15
                -0.061820 **
##
                (0.022604)
## china.16
                 0.055998
                 (0.048651)
## vol.16
                 0.044227
##
                 (0.048698)
## const
                 0.041625 ***
                 (0.004912)
## ----
## R^2
                  0.036628
## Adj. R^2
                  0.036000
## Num. obs. 19965
## ==========
## *** p < 0.001; ** p < 0.01; * p < 0.05
#Recreate a Robust Omega Matrix
Uchina = residuals(est.VARchina)
Tchina = nrow(Uchina)
Omegachina = matrix(0, ncol(Uchina), ncol(Uchina))
for(l in 0:L) {
  weight = 1 - 1/(L+1)
 Gamma_l_china = t(Uchina[(1+1):Tchina, , drop=FALSE]) %*% Uchina[1:(Tchina-l), , drop=FALSE] /Tchina
 if (1 == 0){
   Omegachina = Omegachina + Gamma_l_china
 } else {
   Omegachina = Omegachina + weight*(Gamma_l_china + t(Gamma_l_china))
  }
}
#make the B matrix
losschina <- function(paramchina){</pre>
  #Define the restriction
 Bchina <- matrix(c(paramchina[1], paramchina[2], 0, paramchina[3]), ncol = 2)
  #Make BB' approximatively equal to omega
 Xchina <- Omegachina - Bchina %*% t(Bchina)</pre>
  #loss function
 losschina <- sum(Xchina^2)</pre>
  return(losschina)
}
res.optchina <- optim(c(1, 0, 1), losschina, method = "BFGS")
B.hatchina <- matrix(c(res.optchina$par[1], res.optchina$par[2], 0, res.optchina$par[3]), ncol = 2)
print(cbind(Omegachina, B. hatchina %*% t(B. hatchina)))
                china
                               vol
## china 0.1936352468 0.0008527492 0.1936342473 0.0008527453
## vol 0.0008527492 0.0052690830 0.0008527453 0.0052683426
```

B.hatchina

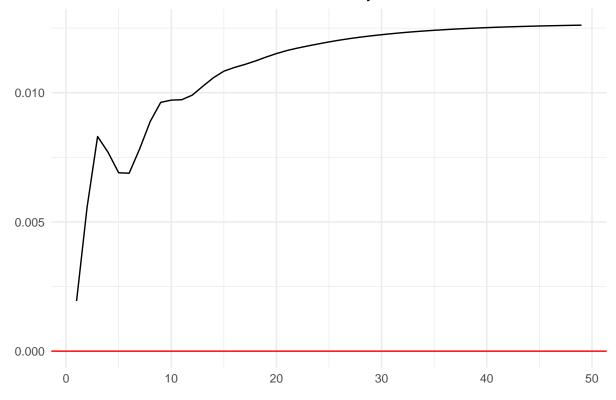
```
[,2]
##
                [,1]
## [1,] 0.440038916 0.00000000
## [2,] 0.001937886 0.07255748
\#get\ back\ the\ coefficient\ of\ est.VAR
phichina <- Acoef(est.VARchina)</pre>
PHIchina = make.PHI(phichina)
#take the constant
constantchina <- sapply(est.VARchina$varresult, function(eq) coef(eq)["const"])</pre>
cchina=as.matrix(constantchina)
#Simulate the IRF
pchina <- length(phichina)</pre>
nchina <- dim(phichina[[1]])[1]</pre>
Ychina <- simul.VAR(c=cchina, Phi = phichina, B = B.hatchina, nb.sim, y0.star=rep(0, nchina*pchina),
                  indic.IRF = 1, u.shock = c(1,0))
#Plot the IRF
Ydchina = data.frame(
  period = 1:nrow(Ychina),
  response = Ychina[,2])
ggplot(Ydchina,aes(x=period, y=response)) +
  geom_hline(yintercept = 0, color="red") +
  geom_line() +
  theme_light() +
  ggtitle("S&P IRF of China on Volatility") +
  ylab("")+
  xlab("") +
  theme_minimal()
```

S&P IRF of China on Volatility



```
ggplot(Ydchina,aes(x=period, y=cumsum(response))) +
  geom_hline(yintercept = 0, color="red") +
  geom_line() +
  theme_light() +
  ggtitle("S&P Cumulalitve IRF of China on Volatility") +
  ylab("")+
  xlab("") +
  theme_minimal()
```

S&P Cumulalitve IRF of China on Volatility



grangertest(ychina[,c("vol", "china")], order = 6)

Res.Df	Df	F	Pr(>F)
2e+04			
2e+04	-6	7.43	5.74e-08

grangertest(ychina[,c("china", "vol")], order = 6)

Res.Df	Df	\mathbf{F}	$\Pr(>F)$
2e+04			
2e+04	-6	8.96	8.55e-10

Split Terms

Here we look for the first and second mandate effect of posts. We will use the tariff variable as a proxy for the posts.

```
# First and Second Mandate

#first term
Vdata_f = filter(data,between(timestamp, as.Date('2017-01-20'), as.Date('2021-01-20')))

#second term
Vdata_s = filter(data,between(timestamp, as.Date('2025-01-20'), as.Date('2025-05-07')))
```

First mandate

```
y_f_d = cbind(Vdata_f$dummy, Vdata_f$SPY_vol)
colnames(y_f_d)[1:2] <- c("dummy", "vol")</pre>
est.VAR_f_d \leftarrow VAR(y_f_d,p=6)
#extract results
mod_vol_f_d = est.VAR_f_d$varresult$vol
f_f_d = formula(mod_vol_f_d)
d_f_d = model.frame(mod_vol_f_d)
lm_clean_f_d = lm(f_f_d, data = d_f_d)
#apply Newey-West
nw_vcov_f_d = NeweyWest(lm_clean_f_d, lag=6)
nw_se_f_d = sqrt(diag(nw_vcov_f_d))
#t-stats
coef f d = coef(lm clean f d)
t_stat_f_d = coef_f_d/nw_se_f_d
\#recalculate\ p-values
robust_f_d = 2*(1-pt(abs(t_stat_f_d), df = df.residual(lm_clean_f_d)))
              <- nw_se_f_d[names(coef(lm_clean_f_d))]</pre>
nw se f d
              <- robust_f_d[names(coef(lm_clean_f_d))]</pre>
robust_f_d
#table
screenreg(lm_clean_f_d, override.se = nw_se_f_d, override.pvalues = robust_f_d, digits = 6)
======== Model 1
               - dummy.ll -0.000478 (0.000133)
vol.l1 0.541944 (0.074062)
dummy.l2 -0.000184 ** (0.000069)
vol.12 -0.113920 ** (0.038762)
dummy.l3 -0.000693 (0.000160)
vol.l3 \ 0.058050
(0.027504)
dummy.l4 -0.000564 (0.000166)
vol.l4 0.188383
(0.132562)
dummy.l5 -0.000435 (0.000113)
vol.l5 -0.088758
(0.091453)
```

```
dummy.l6 0.000118
(0.000118)
vol.l6 0.336662 (0.049019)
const 0.004020 (0.000669)
                 - R^2 0.687909
Adj. R^2 0.687331
Num. obs. 7036
#Table for the effect of volatility on posts for variable dummy
#extract results
mod_post_f_d = est.VAR_f_d$varresult$dummy
ff_f_d = formula(mod_post_f_d)
dd_f_d = model.frame(mod_post_f_d)
lm_clean_post_f_d = lm(ff_f_d, data= dd_f_d)
#apply Newey-West
nw_vcov_post_f_d = NeweyWest(lm_clean_post_f_d, lag=6)
nw_se_post_f_d = sqrt(diag(nw_vcov_post_f_d))
#t-stats
coef_post_f_d = coef(lm_clean_post_f_d)
t_stat_post_f_d = coef_post_f_d/nw_se_post_f_d
#recalculate p-values
robust_post_f_d = 2*(1-pt(abs(t_stat_post_f_d), df = df.residual(lm_clean_post_f_d)))
                   <- nw_se_post_f_d[names(coef(lm_clean_post_f_d))]</pre>
nw_se_post_f_d
                   <- robust_post_f_d[names(coef(lm_clean_post_f_d))]</pre>
robust_post_f_d
#table
screenreg(lm_clean_post_f_d, override.se = nw_se_post_f_d, override.pvalues = robust_post_f_d, digits =
##
## ==========
##
            Model 1
## -----
               -0.136890 ***
## dummy.l1
##
               (0.006701)
## vol.11
               6.918718 **
##
               (2.346819)
              -0.105079 ***
## dummy.12
##
               (0.005805)
## vol.12
              -4.399079 **
               (1.426496)
##
## dummy.13
               -0.089466 ***
               (0.006149)
##
## vol.13
               -1.329415
##
               (0.681253)
## dummy.14
               -0.087428 ***
##
               (0.006320)
```

vol.14

##

-2.004975 **

(0.762777)

```
## dummy.15
               -0.110469 ***
##
               (0.006633)
               -0.823577
## vol.15
##
               (0.894915)
## dummy.16
               -0.135443 ***
##
               (0.007492)
## vol.16
               3.883219 **
##
               (1.365831)
## const
               1.941260 ***
##
               (0.060551)
## R^2
                0.190525
                0.189027
## Adj. R^2
## Num. obs. 7036
## *** p < 0.001; ** p < 0.01; * p < 0.05
#Construct the Robust Omega Matrix
U_f_d = residuals(est.VAR_f_d)
T_f_d = nrow(U_f_d)
Omega_f_d = matrix(0, ncol(U_f_d), ncol(U_f_d))
for(l in 0:L) {
  weight = 1 - 1/(L+1)
 if (1 == 0){
   Omega_f_d = Omega_f_d + Gamma_l_f_d
 } else {
   Omega_f_d = Omega_f_d + weight*(Gamma_l__f_d + t(Gamma_l__f_d))
}
#make the B matrix
loss f d <- function(param f d){</pre>
  #Define the restriction
 B_f_d \leftarrow matrix(c(param_f_d[1], param_f_d[2], 0, param_f_d[3]), ncol = 2)
  #Make BB' approximatively equal to omega
 X_f_d \leftarrow Omega_f_d - B_f_d %*% t(B_f_d)
  #loss function
 loss_f_d <- sum(X_f_d^2)</pre>
 return(loss_f_d)
}
res.opt_f_d \leftarrow optim(c(1, 0, 1), loss_f_d, method = "BFGS")
B.hat_f_d <- matrix(c(res.opt_f_dpar[1], res.opt_f_dpar[2], 0, res.opt_f_dpar[3]), ncol = 2)
print(cbind(Omega_f_d,B.hat_f_d %*% t(B.hat_f_d)))
##
                           vol
             dummy
## dummy 9.80314631 0.0091248204 9.803145998 0.0091245935
## vol 0.00912482 0.0007597912 0.009124594 0.0007587832
```

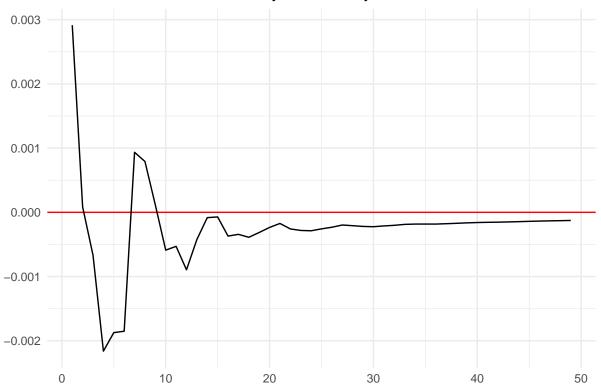
B.hat_f_d ## [,1] [,2] **##** [1,] 3.130997604 0.00000000 ## [2,] 0.002914277 0.02739143 #get back the coefficient of est.VAR phi_f_d <- Acoef(est.VAR_f_d)</pre> PHI_f_d = make.PHI(phi_f_d) #take the constant constant_f_d <- sapply(est.VAR_f_d\$varresult, function(eq) coef(eq)["const"])</pre> c_f_d=as.matrix(constant_f_d) #Simulate the IRF p_f_d <- length(phi_f_d)</pre> n_f_d <- dim(phi_f_d[[1]])[1]</pre> $Y_f_d \leftarrow simul.VAR(c=c_f_d, Phi = phi_f_d, B = B.hat_f_d, nb.sim ,y0.star=rep(0, n_f_d*p_f_d),$ indic.IRF = 1, u.shock = c(1,0)) #Plot the IRF Yd_f_d = data.frame(period = 1:nrow(Y_f_d), response = Y_f_d[,2]) ggplot(Yd_f_d,aes(x=period, y=response)) + geom_hline(yintercept = 0, color="red") + geom_line() +

theme_light() +

ylab("")+
xlab("") +
theme_minimal()

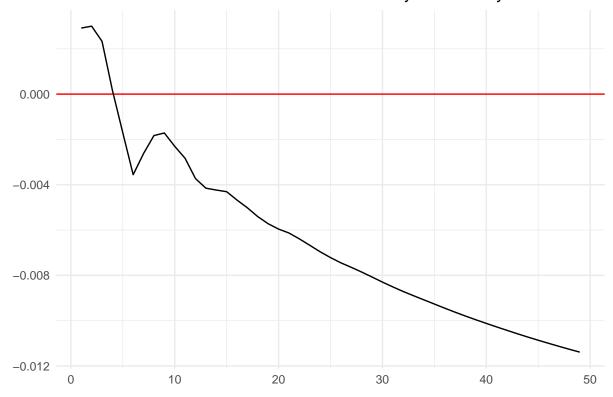
ggtitle("S&P IRF of First Term Dummy on Volatility") +

S&P IRF of First Term Dummy on Volatility



```
ggplot(Yd_f_d,aes(x=period, y=cumsum(response))) +
  geom_hline(yintercept = 0, color="red") +
  geom_line() +
  theme_light() +
  ggtitle("S&P Cumulalitve IRF of First Mandate Dummy on Volatility") +
  ylab("") +
  xlab("") +
  theme_minimal()
```

S&P Cumulalitve IRF of First Mandate Dummy on Volatility



#does vol granger cause dummy
grangertest(y_f_d[,c("vol","dummy")], order =6)

Res.Df	\mathbf{Df}	\mathbf{F}	$\Pr(>F)$
7.02e+03			
7.03e+03	-6	13	9.77e-15

#does dummy granger cause vol
grangertest(y_f_d[,c("dummy", "vol")], order =6)

Res.Df	Df	${f F}$	$\Pr(>F)$
7.02e+03			
7.03e+03	-6	9.87	7.15e-11

Second Mandate

```
y_s_d = cbind(Vdata_s$dummy, Vdata_s$SPY_vol)
colnames(y_s_d)[1:2] <- c("dummy", "vol")</pre>
est.VAR_s_d \leftarrow VAR(y_s_d,p=6)
#extract results
mod_vol_s_d = est.VAR_s_d$varresult$vol
f_s_d = formula(mod_vol_s_d)
d_s_d = model.frame(mod_vol_s_d)
lm_clean_s_d = lm(f_s_d, data= d_s_d)
#apply Newey-West
nw_vcov_s_d = NeweyWest(lm_clean_s_d, lag=6)
nw_se_s_d = sqrt(diag(nw_vcov_s_d))
#t-stats
coef_s_d = coef(lm_clean_s_d)
t_stat_s_d = coef_s_d/nw_se_s_d
#recalculate p-values
robust_s_d = 2*(1-pt(abs(t_stat_s_d), df = df.residual(lm_clean_s_d)))
nw_se_s_d <- nw_se_s_d[names(coef(lm_clean_s_d))]</pre>
robust_s_d <- robust_s_d[names(coef(lm_clean_s_d))]
#table
screenreg(lm_clean_s_d, override.se = nw_se_s_d, override.pvalues = robust_s_d, digits = 6)
======= Model 1
            —- dummy.ll 0.006569
(0.010145)
vol.l1 0.299398 ** (0.112369)
dummy.l2 -0.003222 ** (0.001039)
vol.l2 0.015406
(0.043748)
dummy.l3 -0.005538 ** (0.001707)
vol.13 0.076169 (0.008464)
dummy.l4 0.002474
(0.004957)
vol.l4 0.084229
(0.067860)
dummy.l5 - 0.008527
(0.003989)
vol.l5 0.013424 (0.005076)
dummy.16 - 0.003594
(0.003191)
vol.16 0.126612 *
(0.050031)
const 0.072524 ** (0.024156)
              -- R^2 0.244117
Adj. R^2 0.224424
Num. obs. 512
```

```
#Table for the effect of volatility on posts for variable dummy
#extract results
mod_post_s_d = est.VAR_s_d$varresult$dummy
ff_s_d = formula(mod_post_s_d)
dd_s_d = model.frame(mod_post_s_d)
lm_clean_post_s_d = lm(ff_s_d, data= dd_s_d)
#apply Newey-West
nw_vcov_post_s_d = NeweyWest(lm_clean_post_s_d, lag=6)
nw_se_post_s_d = sqrt(diag(nw_vcov_post_s_d))
#t-stats
coef_post_s_d = coef(lm_clean_post_s_d)
t_stat_post_s_d = coef_post_s_d/nw_se_post_s_d
#recalculate p-values
robust_post_s_d = 2*(1-pt(abs(t_stat_post_s_d), df = df.residual(lm_clean_post_s_d)))
                    <- nw_se_post_s_d[names(coef(lm_clean_post_s_d))]</pre>
nw_se_post_s_d
                    <- robust_post_s_d[names(coef(lm_clean_post_s_d))]</pre>
robust_post_s_d
#table
screenreg(lm_clean_post_s_d, override.se = nw_se_post_s_d, override.pvalues = robust_post_s_d, digits =
##
## =========
##
             Model 1
## -----
              -0.216984 ***
## dummy.l1
##
               (0.020971)
## vol.11
              -0.042631
              (0.082498)
## dummy.12
              -0.208185 ***
              (0.020850)
##
## vol.12
              -0.156762
               (0.090144)
## dummy.13
              -0.205129 ***
##
               (0.021613)
               0.273732 *
## vol.13
               (0.124059)
## dummy.14
              -0.217831 ***
##
               (0.022085)
## vol.14
              -0.204545 ***
##
               (0.058272)
## dummy.15
              -0.214630 ***
##
               (0.021698)
## vol.15
              -0.103301
##
               (0.057821)
## dummy.16
              -0.199256 ***
              (0.022328)
##
```

vol.16

const

##

0.059322

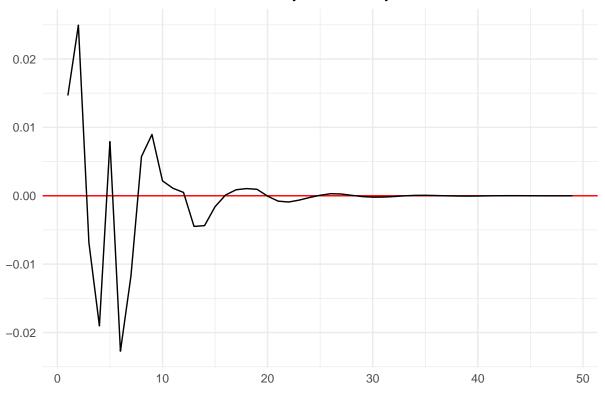
(0.076999)

3.161434 ***

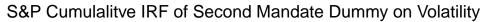
```
##
              (0.255375)
## ----
## R^2
               0.297771
## Adj. R^2
              0.279477
## Num. obs. 512
## =========
## *** p < 0.001; ** p < 0.01; * p < 0.05
#Construct the Robust Omega Matrix
U_s_d = residuals(est.VAR_s_d)
T_s_d = nrow(U_s_d)
Omega_s_d = matrix(0, ncol(U_s_d), ncol(U_s_d))
for(1 in 0:L) {
  weight = 1 - 1/(L+1)
 if (1 == 0){
   Omega_s_d = Omega_s_d + Gamma_l_s_d
 } else {
   Omega_s_d = Omega_s_d + weight*(Gamma_l_s_d + t(Gamma_l_s_d))
 }
}
#make the B matrix
loss_s_d <- function(param_s_d){</pre>
  #Define the restriction
 B_s_d \leftarrow matrix(c(param_s_d[1], param_s_d[2], 0, param_s_d[3]), ncol = 2)
 #Make BB' approximatively equal to omega
 X_s_d \leftarrow Omega_s_d - B_s_d %*% t(B_s_d)
 #loss function
 loss_s_d \leftarrow sum(X_s_d^2)
 return(loss s d)
}
res.opt_s_d <- optim(c(1, 0, 1), loss_s_d, method = "BFGS")
B.hat_s_d \leftarrow matrix(c(res.opt_s_dpar[1], res.opt_s_dpar[2], 0, res.opt_s_dpar[3]), ncol = 2)
print(cbind(Omega_s_d,B.hat_s_d %*% t(B.hat_s_d)))
##
            dummy
                       vol
## dummy 9.7947444 0.0458797 9.79474361 0.04587955
## vol
        0.0458797 0.1852225 0.04587955 0.18522157
B.hat s d
                       [,2]
##
             [,1]
## [1,] 3.12965551 0.000000
## [2,] 0.01465962 -0.430124
```

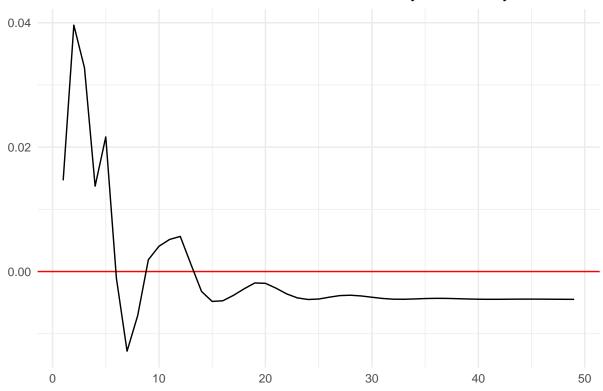
```
\#get\ back\ the\ coefficient\ of\ est.VAR
phi_s_d <- Acoef(est.VAR_s_d)</pre>
PHI_s_d = make.PHI(phi_s_d)
#take the constant
constant_s_d <- sapply(est.VAR_s_d$varresult, function(eq) coef(eq)["const"])</pre>
c_s_d=as.matrix(constant_s_d)
#Simulate the IRF
p_s_d <- length(phi_s_d)</pre>
n_s_d <- dim(phi_s_d[[1]])[1]</pre>
Y_s_d <- simul.VAR(c=c_s_d, Phi = phi_s_d, B = B.hat_s_d, nb.sim ,y0.star=rep(0, n_s_d*p_s_d),
                   indic.IRF = 1, u.shock = c(1,0))
#Plot the IRF
Yd_s_d = data.frame(
 period = 1:nrow(Y_s_d),
 response = Y_s_d[,2])
ggplot(Yd_s_d,aes(x=period, y=response)) +
  geom_hline(yintercept = 0, color="red") +
  geom_line() +
  theme_light() +
  ggtitle("S&P IRF of Second Term Dummy on Volatility") +
  ylab("")+
 xlab("") +
 theme_minimal()
```

S&P IRF of Second Term Dummy on Volatility



```
ggplot(Yd_s_d,aes(x=period, y=cumsum(response))) +
  geom_hline(yintercept = 0, color="red") +
  geom_line() +
  theme_light() +
  ggtitle("S&P Cumulalitve IRF of Second Mandate Dummy on Volatility") +
  ylab("")+
  xlab("") +
  theme_minimal()
```





#does vol granger cause dummy
grangertest(y_s_d[,c("vol","dummy")], order =6)

Res.Df	Df	${f F}$	$\Pr(>F)$
499			
505	-6	0.303	0.935

#does dummy granger cause vol
grangertest(y_s_d[,c("dummy", "vol")], order =6)

Res.Df	\mathbf{Df}	\mathbf{F}	$\Pr(>F)$
499			
505	-6	0.621	0.713