

Gaussian Distribution

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I. Multivariate Gaussian Distribution:

Proof that Multivariate Gaussian Distribution is normalize:

First, we have the PDF of the Gaussian Distribution is:

$$p(x | \mu, \sigma^2) = \frac{1}{(2\pi)^{\frac{D}{2}} |\Sigma|^{\frac{1}{2}}} \times e^{-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)}$$

The Gaussian Distribution is normalize

$$\iff \int_{-\infty}^{+\infty} p(x | \mu, \sigma^2) = 1$$

where μ is a D-dimensional mean vector, Σ is a D x D covariance matrix, and $|\Sigma|$ denotes the determinant of Σ

Set

$$\begin{aligned} \Delta^2 &= \frac{-1}{2}(x - \mu)^T \Sigma^{-1}(x - \mu) \\ &= \frac{-1}{2}x^T \Sigma^{-1}x + x^T \Sigma^{-1}\mu + constant \end{aligned}$$

Consider eigenvalues and eigenvectors of Σ we have:

$$\Sigma u_i = \lambda_i u_i, i = 1, \dots, D$$

Because Σ is a real, symmetric matrix, its eigenvalues will be real and its eigenvectors form an orthonormal set.

Proof:

1. its eigenvalues will be real

Example:

$$\begin{pmatrix} \sigma_1^2 & \sigma_{1,2} \\ \sigma_{2,1} & \sigma_2^2 \end{pmatrix} \iff \text{The equation to find the eigenvalues is :}$$

$$(\sigma_1^2 - \lambda) \times (\sigma_2^2 - \lambda) - (\sigma_{1,2})^2 = 0$$

$$\iff (\sigma_1^2 - \lambda) \times (\sigma_2^2 - \lambda) = (\sigma_{1,2})^2$$

$\implies \lambda$ must be a real number

With $\lambda = \lambda_1$:

$$\begin{pmatrix} \sigma_1^2 - \lambda_1 & cov(\sigma_{1,2}) \\ cov(\sigma_{1,2}) & \sigma_2^2 - \lambda_1 \end{pmatrix}$$

$$(\sigma_1^2 - \lambda_1)x_1 + (\sigma_{1,2})x_2 = 0 \quad (1)$$

$$(\sigma_{1,2})x_1 + (\sigma_2^2 - \lambda_1)x_2 = 0 \quad (2)$$

From (1) we have:

$$x_1 = \frac{-y \times cov(\sigma_{1,2})}{\sigma_1^2 - \lambda_1}$$

$$x_2 = x_2$$

So the eigenvector in this case is: $\begin{pmatrix} \frac{-cov(\sigma_{1,2})}{\sigma_1^2 - \lambda_1} \\ 1 \end{pmatrix}$

With $\lambda = \lambda_2$:

$$\begin{pmatrix} \sigma_1^2 - \lambda_2 & cov(\sigma_{1,2}) \\ cov(\sigma_{1,2}) & \sigma_2^2 - \lambda_2 \end{pmatrix}$$

$$(\sigma_1^2 - \lambda_2)x_1 + (\sigma_{1,2})x_2 = 0 \quad (3)$$

$$(\sigma_{1,2})x_1 + (\sigma_2^2 - \lambda_2)x_2 = 0 \quad (4)$$

From (3) we have:

$$x_1 = \frac{-y \times cov(\sigma_{1,2})}{\sigma_1^2 - \lambda_2}$$

$$x_2 = x_2$$

So the eigenvector in this case is: $\begin{pmatrix} \frac{-cov(\sigma_{1,2})}{\sigma_1^2 - \lambda_2} \\ 1 \end{pmatrix}$

$$\text{And: } \begin{pmatrix} \frac{-cov(\sigma_{1,2})}{\sigma_1^2 - \lambda_2} \\ 1 \end{pmatrix}^T \times \begin{pmatrix} \frac{-cov(\sigma_{1,2})}{\sigma_1^2 - \lambda_2} \\ 1 \end{pmatrix} = 1$$

So its eigenvectors form an orthonormal set.

$$\Sigma = \Sigma_{i=1}^D \lambda_i u_i (u_i)^T \longrightarrow \Sigma^{-1} = \Sigma_{i=1}^D \frac{1}{\lambda_i} u_i u_i^T$$

So that:

$$\begin{aligned} \Delta^2 &= \frac{-1}{2} (x - \mu)^T \Sigma^{-1} (x - \mu) \\ &= \sum_{i=1}^D \frac{1}{\lambda_i} (x - \mu)^T u_i u_i^T (x - \mu) \end{aligned}$$

Let:

$$\begin{aligned} y_i &= u_i^T (x - \mu) \\ \longrightarrow \Delta^2 &= \sum_{i=1}^D \frac{y_i^2}{\lambda_i} \\ |\Sigma|^{1/2} &= \prod_{j=1}^D \lambda_j^{1/2} \end{aligned}$$

Now, we have:

$$\begin{aligned} p(x | \mu, \sigma^2) &= p(y) = \prod_{j=1}^D \frac{1}{(2\pi\lambda_j)^{1/2}} e^{-\frac{(y_j)^2}{2\lambda_j}} \\ \Longleftrightarrow \int_{-\infty}^{+\infty} p(y) dy &= \prod_{j=1}^D \int_{-\infty}^{+\infty} \frac{1}{(2\pi\lambda_j)^{1/2}} e^{-\frac{(y_j)^2}{2\lambda_j}} \end{aligned}$$

But in the last homework we have proofed that:

$$\frac{1}{\sqrt{2\pi}} \int_{-\infty}^{+\infty} e^{-\frac{y^2}{2}} dy = 1$$

So:

$$\begin{aligned} \int_{-\infty}^{+\infty} \frac{1}{(2\pi\lambda_j)^{1/2}} e^{-\frac{(y_j)^2}{2\lambda_j}} &= 1 \\ \Longleftrightarrow \prod_{j=1}^D \int_{-\infty}^{+\infty} \frac{1}{(2\pi\lambda_j)^{1/2}} e^{-\frac{(y_j)^2}{2\lambda_j}} &= 1 \end{aligned}$$