

CLT proof

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WLLN X_1, \dots, X_n iid $E X_i = \mu$, $Var X_i = \sigma^2 < \infty$

define $\bar{X} = \frac{1}{n} \sum X_i$ (average of a R.V.)

then $\lim_{n \rightarrow \infty} P(|\bar{X}_n - \mu| > \epsilon) = 0$

$\bar{X}_n \xrightarrow{P} \mu$ *

proof. chebychev's

$$P(|\bar{X}_n - \mu| > \epsilon) = P(|\bar{X}_n - \mu|^2 > \epsilon^2) \leq \frac{E(\bar{X}_n - \mu)^2}{\epsilon^2}$$

$$= \frac{Var(\bar{X}_n)}{\epsilon^2} = \frac{\sigma^2}{n\epsilon^2} \rightarrow 0 \text{ as } n \rightarrow \infty$$

consistency of s^2 $P(|S_n^2 - \sigma^2| > \epsilon) \leq \frac{E(|S_n^2 - \sigma^2|)}{\epsilon^2}$

$$S_n^2 \xrightarrow{P} \sigma^2 \text{ if } Var(S_n^2) \rightarrow 0 = \frac{Var(S_n^2)}{\epsilon^2}$$

CLT Thm. sample mean \sim Normal dist

X_1, X_2, \dots, X_n iid R.V.s finite μ & σ^2 , MGF exist. *

$$\textcircled{1} \sum X_i \sim N(n\mu, n\sigma^2) \Rightarrow \frac{\sum X_i - n\mu}{\sqrt{n}\sigma} \sim N(0, 1)$$

$$\textcircled{2} \frac{1}{n} \sum X_i = \bar{X}_n \sim N(\mu, \frac{\sigma^2}{n}) \Rightarrow \frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \sim N(0, 1)$$

$\textcircled{1}$ & $\textcircled{2}$ equivalent

in practice, we only observe one realization \rightarrow one \bar{X}_n .

Proof. show $\frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \xrightarrow{d} N(0, 1)$ MGF = $e^{\frac{t^2}{2}}$ $N(0, 1)$ MGF. * standard normal easier to work with.

$$\begin{aligned} M_{\frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma}}(t) &= M_{\sum_{i=1}^n Y_i / \sqrt{n}}(t) \quad Y_i = \frac{X_i - \mu}{\sigma} \Rightarrow \sum Y_i = \frac{\sum X_i - n\mu}{\sigma} \\ &= [M_{Y_i(\frac{t}{\sqrt{n}})}]^n \quad \text{sub in} \\ M_{\frac{\sum Y_i}{\sqrt{n}}}(t) &= [M_{\frac{Y_i}{\sqrt{n}}}(t)]^n = [M_{Y_i(\frac{t}{\sqrt{n}})}]^n \quad \frac{\sqrt{n}(\frac{1}{n} \sum X_i - \mu)}{\sigma} \downarrow \frac{\frac{1}{\sqrt{n}} \sum X_i - \sqrt{n}\mu}{\sigma} = \frac{\sum X_i - n\mu}{\sqrt{n}\sigma} \\ M_{\frac{Y_i}{\sqrt{n}}}(t) &= E(e^{t \frac{Y_i}{\sqrt{n}}}) = M_{Y_i}(\frac{t}{\sqrt{n}}) \end{aligned}$$

2) $M_{Y_i}(\frac{t}{\sqrt{n}})$ expand around 0. \Rightarrow 我觉得是因为在0处展开导数 这个 Moment 有表达式, $M_{Y_i}^{(1)}(0) = EX$, $M_{Y_i}^{(2)}(0) = EX^2$

$$f(x) = f(x_0) + (x-x_0)f'(x_0) + \frac{(x-x_0)^2}{2!}f''(x_0) + \frac{(x-x_0)^3}{3!}f'''(x_0) + \dots$$

$$= \sum_{k=0}^{\infty} f^{(k)}(x_0) \frac{(x-x_0)^k}{k!} \quad 2) f^{(0)}(x_0) = f(x_0) \text{ 0次导为原函数}$$

$$\begin{aligned} M_{Y_i}(\frac{t}{\sqrt{n}}) &= M_{Y_i}(0) + \frac{t}{\sqrt{n}} \cdot M_{Y_i}^{(1)}(0) + \frac{(\frac{t}{\sqrt{n}})^2}{2!} M_{Y_i}^{(2)}(0) + \dots \\ &= e^{\frac{t^2}{2}} \Big|_{t=0} + \frac{(\frac{t}{\sqrt{n}})^2}{2!} EX^2 + \dots \end{aligned}$$

$$= 1 + \frac{(\frac{t}{\sqrt{n}})^2}{2} + \text{Remainders.} \quad \sum_{k=3}^{\infty} M_{Y_i}^{(k)}(0) \cdot \frac{(\frac{t}{\sqrt{n}})^k}{k!} = 0 \text{ 得补充}$$

$$\begin{aligned} [M_{Y_i}(\frac{t}{\sqrt{n}})]^n &= (1 + \frac{t^2}{2n})^n \\ &= (1 + \frac{t^2/2}{n})^n \end{aligned}$$

$$\therefore \lim_{n \rightarrow \infty} [M_{Y_i}(\frac{t}{\sqrt{n}})]^n = \lim_{n \rightarrow \infty} (1 + \frac{t^2/2}{n})^n = e^{\frac{t^2}{2}}$$

Slutsky's Theorem.

$$X_n \xrightarrow{D} X$$

$$Y_n \xrightarrow{P} a \text{ (constant)}$$

$$\Rightarrow X_n Y_n \xrightarrow{D} aX$$

$$\Rightarrow X_n + Y_n \xrightarrow{D} X + a.$$

eg. by CLT $\frac{\sqrt{n}(\bar{X} - \mu)}{\sigma} \xrightarrow{D} N(0, 1)$

$$\boxed{\text{if } \lim_{n \rightarrow \infty} Var S_n^2 \rightarrow 0} \rightarrow \boxed{S_n^2 \rightarrow \sigma^2} \rightarrow \boxed{\frac{\sigma}{S_n} \rightarrow 1}$$

$$P(|S_n^2 - \sigma^2| > \epsilon) \leq \frac{Var(S_n^2)}{\epsilon^2} \rightarrow 0 = 0.$$

$$\therefore \textcircled{1} \frac{\sigma}{S_n} \xrightarrow{P} 1 \text{ (by defn)}$$

$$\textcircled{2} \frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \xrightarrow{D} N(0, 1)$$

Slutsky's

$$\frac{\sqrt{n}(\bar{X}_n - \mu)}{S_n} \xrightarrow{D} N(0, 1)$$

这让我想到了正态分布 $\bar{X} \sim N$, $S_n \sim \chi^2$

$$\frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \sim N(0, 1)$$

$$S^2/\sigma^2 \sim \chi^2(n-1)$$

Scaled?