



1. of interest

out of c components, at least t of them will last h hours.

step 1: sample $x_1, x_2, \dots, x_c \stackrel{iid}{\sim} \exp(\lambda)$

step 2: $Y = \begin{cases} 1 & \text{if at least } t \text{ of them } > h \\ 0 & \text{o/w.} \end{cases}$ of interest p .

$\therefore p = E(Y)$

\therefore repeat step 1 and 2. 得到 n 个 Y .

step 3: 根据 MLE 估计 $\hat{p} \rightarrow E(Y)$

\therefore 求 $E(Y)$, 从而得到 p .

\therefore step 1 是要 generate random variables.

那么, 有哪些方法可以生成呢?

① probability integral transformation.

X continuous. w/ cdf $F_X(x)$

def $Y = F_X(X)$. then

$Y \sim \text{Unif}(0,1) \Rightarrow P(Y \leq y) = y$.

$P(Y \leq y) = P(F_X(X) \leq y)$

$= P(F_X^{-1}(F_X(X)) \leq F_X^{-1}(y))$

$= P(X \leq F_X^{-1}(y))$

$= F_X(F_X^{-1}(y))$

$= y$.

$\Rightarrow Y \sim \text{Unif}(0,1)$

Proof. Y has cdf $F_Y(y)$. $Y \sim \exp(\lambda)$

$F_Y(Y) = U \Rightarrow U \sim \text{Unif}(0,1)$

$\therefore P(U \leq u) = P(F_Y(Y) \leq u)$

$= P(Y \leq F_Y^{-1}(u))$

$= u$.

$\therefore Y = F_Y^{-1}(U) \sim F_Y$

$P(Y \leq y) = P(F_Y^{-1}(U) \leq y)$

$= P(U \leq F_Y(y))$

$= F_Y(y)$

$$\Rightarrow \begin{cases} F_Y(Y) = U \\ Y = F_Y^{-1}(U) \end{cases}$$

$f_Y(y) = \lambda e^{-\lambda y}$

$F_Y(y) = \int_0^y \lambda e^{-\lambda t} dt$

$-e^{-\lambda t} \Big|_0^y = 1 - e^{-\lambda y}$

\therefore 求 F_Y^{-1} *parameterization 方法不同, 结果会不同*

$y = 1 - e^{-\lambda t}$

$1 - y = e^{-\lambda t}$

$\log(1-y) = -\lambda t$

$-2 \log(1-y) = t(y)$

$Y = -2 \log(1-U) \sim \exp(2)$

\therefore 这样, step 1 就 decompose 为产生 U , 再产生 $Y \sim \exp(\lambda)$

exponential-uniform transformation.

$X \sim \text{Unif}(0,1)$ $Y = -\log X \Rightarrow X = e^{-Y}$ $\frac{dY}{dX} = -e^{-X}$

$\therefore f_Y(y) = -e^{-y} \sim \exp(1)$

example: if $U_j \sim \text{Unif}(0,1)$ i.i.d.

$Y_j = -\lambda \log(U_j) \Rightarrow \frac{dY_j}{dU_j} = \frac{1}{U_j} \Rightarrow \frac{dY_j}{dU_j} = -\frac{1}{U_j} e^{-Y_j/\lambda} \sim \exp(1/\lambda)$

$Y = -2(\log U_1 + \log U_2 + \dots + \log U_n)$

$= -2 \log U_1 + (-2) \log U_2 + \dots + (-2) \log U_n$

$\sim \exp(2)$

$\therefore \sum_{j=1}^n \exp(2) \sim \text{Gamma}(n, 2) \Rightarrow \chi^2(2n)$

$Y = -\beta \sum_{j=1}^n \log(U_j) \sim \text{Gamma}(n, \beta)$ *generalization.*

$Y = \frac{-\sum_{j=1}^a \log(U_j)}{-\sum_{j=1}^b \log(U_j)} \xrightarrow{\text{exp(1)}} \frac{\text{Gamma}(a, 1)}{\text{Gamma}(a, 1) + \text{Gamma}(b, 1)} = \text{Beta}(a, b)$

$X_1 \sim \text{Gamma}(a, 1)$

$\therefore f(x_1, x_2) = f(x_1) f(x_2)$

$X_2 \sim \text{Gamma}(b, 1)$

$\therefore Y = \frac{X_1}{X_1 + X_2} \sim \text{Beta}(a, b)$

② acceptance/rejection method. (Setup)

$P(Y \leq y) = P(V \leq y | U < \frac{1}{c} f_Y(V)) \Rightarrow$ we see connections between

U, V indep $Y \leq y$ & V, U

\therefore we can use unif U, V

to generate Y .

V 都是 R.V., 所以要对两个 R.V. 都积分.

分母: $\int_0^1 \int_0^{\frac{1}{c} f_Y(V)} f_{U,V}(u,v) du dv$ 分子: 正确.

$\int_0^1 \int_0^{\frac{1}{c} f_Y(V)} 1 du dv$

$\int_0^1 \frac{1}{c} f_Y(V) dv$

$= \frac{1}{c}$

(a) generate $U, V \stackrel{iid}{\sim} \text{Unif}(0,1)$

(b) if $U < \frac{1}{c} f_Y(V)$ $c = \max_y f_Y(y)$

set $Y = V$. o/w, return to step (a)



to generate R.V. $Y \sim \text{Beta}(2, 3)$

(a) generate $U \sim \text{Unif}(0,1)$, $V \sim f_V \text{Beta}(2, 3)$

(b) if $U < \frac{1}{M} \frac{f_Y(V)}{f_V(V)}$, set $Y = V$. o/w, return to step (a)

\hookrightarrow the more Y and V "look like"

the more $U < \frac{1}{M} \frac{f_Y(V)}{f_V(V)}$, the more likely to accept.

$f(Y \leq y) = f(V \leq y | U < \frac{1}{M} \frac{f_Y(V)}{f_V(V)})$ *只有满足这条件的才接受, 且这样接受的点 则是 Y 的 pdf.*

$$= \frac{f(V \leq y, U \leq \frac{1}{M} \frac{f_Y(V)}{f_V(V)})}{f(U < \frac{1}{M} \frac{f_Y(V)}{f_V(V)})}$$

$$= \frac{\int_0^y \int_0^{\frac{1}{M} \frac{f_Y(V)}{f_V(V)}} 1 du dv}{\int_0^1 \int_0^{\frac{1}{M} \frac{f_Y(V)}{f_V(V)}} 1 du dv}$$

$$= \int_0^y f_Y(v) dv \quad \text{p. 253 SI.}$$

$$M = \sup_y \frac{f_Y(y)}{f_V(y)} < \infty$$

③

$X \sim \text{Unif}(0,1) \rightarrow -\lambda \log(X) \sim \text{Gamma}$

$X \sim \text{Beta}(a, 1) \rightarrow -\log(X) \sim \exp(\frac{1}{a})$