

PART 1: averaged casual effect $E(Y^{a=1} = 1) - E(Y^{a=0} = 1)$
 following are the estimators

- 1.1 IPTW:

$$\frac{1}{n} \sum_{i=1}^n \frac{A_i Y_i}{P(A_i = 1 | L_i)} - \frac{1}{n} \sum_{i=1}^n \frac{(1 - A_i) Y_i}{P(A_i = 0 | L_i)}$$

note: after we assign weight to each individual, the number of individuals in the treatment groups is approximately the number of study samples.

- 1.2 normalized IPTW:

$$\left(\sum_{i=1}^n \frac{A_i}{P(A_i = 1 | L_i)} \right)^{-1} \sum_{i=1}^n \frac{A_i Y_i}{P(A_i = 1 | L_i)} - \left(\sum_{i=1}^n \frac{1 - A_i}{P(A_i = 0 | L_i)} \right)^{-1} \sum_{i=1}^n \frac{(1 - A_i) Y_i}{P(A_i = 0 | L_i)}$$

three versions of weight:

- $\frac{1}{P(A_i=1|L_i)}$ and $\frac{1}{P(A_i=0|L_i)}$: note that average weight is 2, so pseudo-population is twice that of the original study samples (page 161)
 - stablized: $\frac{P(A_i=1)}{P(A_i=1|L_i)}$ and $\frac{P(A_i=0)}{P(A_i=0|L_i)}$ narrower 95%CI than nonstablized weights (Hernan & Robin page 162)
 - if censoring is present $\frac{P(A_i=1)}{P(A_i=1|L_i)} \times \frac{P(C_i=0|A=1)}{P(C_i=0|L_i, A_i=1)}$ and $\frac{P(A_i=0)}{P(A_i=0|L_i)} \times \frac{P(C_i=0|A=0)}{P(C_i=0|L_i, A_i=0)}$
- 2.1 direct standardization (g-formula /g-computation/g-standardization)

$$\begin{aligned} E[Y_a] &= E_L[E(Y_a | L)] \\ &= E_L[E(Y_a | A = a, L)] \quad [\text{conditional exchangeability and positivity}] \\ &= E_L[E(Y | A = a, L)] \quad [\text{consistency}] \end{aligned}$$

$$\sum_{\ell} E[Y | A = 1, \ell] P(L = \ell) - \sum_{\ell} E[Y | A = 0, \ell] P(L = \ell)$$

- 2.2 model-based standardization

$$\sum_{\ell} E[Y | A = 1, \ell] P(L = \ell) = E_L(E[Y | A = 1, L])$$

- $E[Y | A = 1, L]$ estimated from regression model $E[Y | A, L]$ (could specify different function forms in the regression model)
- had everyone(n) treated (i.e., A=1), we can compute the estimated $\hat{E}[Y | A = 1, L = \ell]$ for every level of $L = \ell$. page 172 chapter 13.3)

- compute

$$\begin{aligned} & E(E[Y | A = 1, L]) - E(E[Y | A = 0, L]) \\ &= \frac{1}{n} \sum_{i=1}^n \left\{ E[Y_i | A_i = 1, L_i; \hat{\phi}] - E[Y_i | A_i = 0, L_i; \hat{\phi}] \right\} \end{aligned}$$

IPTW and standardization are numerically equivalent with saturated treatment model $P(A|L)$ and outcome model $E(Y|A, L)$

- 3 doubly robust estimator

$$\frac{1}{n} \sum_{i=1}^n \left[E[Y_i | A_i = a, L_i; \hat{\phi}] + \frac{\mathbf{1}_{\{A_i=a\}}}{P(A_i = a | L_i; \hat{\gamma})} (Y_i - E[Y_i | A_i = a, L_i; \hat{\phi}]) \right]$$

- 4 marginal structural model

$$\log \left(\frac{P(Y_a = 1 | \theta)}{1 - P(Y_a = 1 | \theta)} \right) = \theta_0 + \theta_1 a$$

- the procedure is exactly the same as propensity score weighting. However, Rubin proposed propensity score weighting, and Rubin proposed marginal structural model.

PART 2: conditional casual effect/Heterogeneity of the effect

$$E(Y^{a=1} = 1 | V = v) - E(Y^{a=0} = 1 | V = v)$$

- 1. standardization

$$E[Y_a | V = v] = \sum_{\ell} E[Y | A = a, L = \ell, V = v] P(L = \ell | V = v)$$

- add interaction term in the outcome model.

- 2. potential outcome model

$$E(Y_a | V, \theta) = \theta_0 + \theta_1 a + \theta_2 V a + \theta_3 V$$