

Tips:

- 1) Distinguish **Variance** and **Standard Deviation** in the question!!!
- 2) Is the **Standard Deviation** *known* when choosing test statistics???

A4 Q5. A digital output voltmeter has an input range of 0 to 30V and displays three significant figures XX.X. The manufacture claims an accuracy of  $\pm 2\%$  of full scale. With a voltage reading of 5 V, what are the percent uncertainties of the reading due to accuracy and resolution?

$$\begin{aligned}\text{Manufacturer Accuracy} &= \pm 2.0\% \text{ of full scale} \\ &= \pm 0.02(30\text{V}) \\ &= \pm 0.6\text{V}\end{aligned}$$

$$\% \text{ uncertainty of accuracy} = \pm (0.6\text{V}/5\text{V})(100) = \pm 12\%$$

The resolution of the device is 0.1 Volts. With a reading of 5V,

$$\% \text{ uncertainty of resolution} = (0.1\text{V}/5\text{V})(100) = 2\% \text{ (or } \pm 1\%)$$

A4 Q12. In using a temperature probe, the following uncertainties were determined:

Hysteresis  $\pm 0.1\text{ C}$

Linearization error  $\pm 0.2\%$  of the reading

Repeatability  $\pm 0.1\text{ C}$

Resolution error  $\pm 0.1\text{ C}$

Zero offset  $\pm 0.1\text{ C}$

Determine the type of these errors (random or systematic) and the total uncertainty due to these effects for a temperature reading of  $100\text{ C}$ .

12)

hysteresis	$\pm 0.1C$	systematic
Lineariz.error	0.2% <i>of reading</i>	systematic
Resolution error	0.10C	random
zero off set	0.1C	systematic
repeatability	$\pm 0.1C$	random

$$B = (0.1^2 + [(.002)(100)]^2 + 0.1^2)^{1/2}$$

$$= 0.24C$$

Assuming that the random errors have been determined with samples > 30,

$$P = (.10^2 + .1^2)^{1/2} = 0.14C$$

So total uncertainty

$$w = [B^2 + p^2]^{1/2} = [B^2 + (tS)^2]^{1/2}$$

$$w = [(0.24)^2 + (0.14)^2]^{1/2}$$

$$w = 0.28C$$

15. In measuring the power of a three phase electrical motor running a pump, the following data were obtained (based on one phase current measurement):

V (volt)	460	459	458	460	461	462	460	459
I (amp)	30.2	31.3	30.4	32.0	31.7	30.7	30.8	31.2
PF	0.78	0.80	0.79	0.82	0.81	0.77	0.78	0.80

For a three phase motor, the power  $P$  is related to the voltage  $V$ , current  $I$  and power factor  $PF$  by the formula

$$P = V \times I \times PF \times \sqrt{3}$$

- Calculate the mean, the standard deviation and the random uncertainty of each parameter,  $V$ ,  $I$ ,  $PF$  and  $P$  at a 95% confidence level
- Calculate the mean value of the power and the random uncertain of the man power at a 95% confidence level
- If all the measured parameter have an accuracy of 1% of the reading of the measuring device, calculate the systematic uncertain in the power measurement at a 95% confidence level
- Calculate the total uncertainty in the power measure at a 95% confidence level

15) (a)

$$\bar{V} = \frac{\sum V_i}{n} = 459.9$$

$$\overline{PF} = \frac{\sum PF_i}{n} = 0.79$$

$$S_V = \left[ \frac{\sum (V_i - \bar{V})^2}{n-1} \right]^{1/2} = 1.25 V$$

$$S_{PF} = \left[ \frac{\sum (PF_i - \overline{PF})^2}{n-1} \right]^{1/2} = 0.02$$

$$\bar{I} = \frac{\sum I_i}{n} = 31.0 \text{ amps}$$

$$\bar{P} = \frac{\sum P_i}{n} = \frac{\sum V_i * I_i * PF_i \sqrt{3}}{n} = 19,629 \text{ Watts}$$

$$S_I = \left[ \frac{\sum (I_i - \bar{I})^2}{n-1} \right]^{1/2} = 0.63 \text{ amp}$$

$$S_P = \left[ \frac{\sum (P_i - \bar{P})^2}{n-1} \right]^{1/2} = 787 \text{ Watts}$$

$$P_V = tS_V = 2.95 V, \quad P_I = tS_I = 1.49 \text{ amp}, \quad P_{PF} = tS_{PF} = 0.05, \text{ and } P_P = tS_P = 1,861 \text{ Watts}$$

where "t" has been obtained from Student - t test Table 6.6 for

$$\nu = n - 1 = 7 \text{ for } 95\% \text{ confidence } \left( \frac{\alpha}{2} = 0.0025 \right) \text{ to be } 2.365$$

$$(b) \quad \bar{P} = \frac{\sum P_i}{n} = \frac{\sum V_i * I_i * PF_i \sqrt{3}}{n} = 19,629 \text{ Watts}$$

$$P_{\bar{P}} = t \frac{S_P}{\sqrt{n}} = 2.365 * \frac{787}{\sqrt{8}} = 658 \text{ Watts} \quad 95\% \text{ confidence}$$

$$(c) \quad B_V = B_{\bar{V}} = 1\% \text{ of } \bar{V} = 0.01 * 459.9 = 4.60 V$$

$$B_I = B_{\bar{I}} = 1\% \text{ of } \bar{I} = 0.01 * 31.0 = 0.31 \text{ amp}$$

$$B_{PF} = B_{\overline{PF}} = 1\% \text{ of } \overline{PF} = 0.01 * 0.79 = 0.79 * 10^{-2}$$

Systematic uncertainty of power (applying Eq. 7.24 and utilizing Eq. 7.6 for simplification),

$$B_P = \left[ \left( \frac{\partial P}{\partial V} B_V \right)^2 + \left( \frac{\partial P}{\partial I} B_I \right)^2 + \left( \frac{\partial P}{\partial PF} B_{PF} \right)^2 \right]^{1/2}$$

$$a. \quad \frac{B_P}{P} = \left[ \left( \frac{B_V}{V} \right)^2 + \left( \frac{B_I}{I} \right)^2 + \left( \frac{B_{PF}}{PF} \right)^2 \right]^{1/2} = (0.01^2 + 0.01^2 + 0.01^2)^{1/2} = 0.017$$

$$B_P = 0.017 * 19629 = 334 \text{ Watts}$$

(d)

$$W_{\bar{P}} = [B_P^2 + P_P^2]^{1/2} = (334^2 + 658^2)^{1/2} = 738 \text{ Watts}$$

(With 95% confidence level)

most of the uncertainty is due to random effects.



13. The yield of a chemical process is being studied. From previous experience with this process the standard deviation of yield is known to be 3. The past 5 days of plant operation have resulted in the following yields: 91.6, 88.75, 90.8, 89.95, and 91.3%. Use a significance level of 0.05.
- Is there evidence that the mean yield is not 90%? Use the P-value approach.
  - What sample size would be required to detect a true mean yield of 85% with a probability of 0.95?



13) a) The parameter of interest is the true mean yield,  $\mu$ .

$$\begin{array}{l} H_0: \mu = 90 \\ H_1: \mu \neq 90 \end{array} \quad \text{Test statistic is } z_0 = \frac{\bar{x} - \mu}{\sigma / \sqrt{n}} \quad \bar{x} = 90.48, \quad \sigma = 3 \quad z_0 = \frac{90.48 - 90}{3 / \sqrt{5}} = 0.36$$

P-value =  $2[1 - \Phi(0.36)] = 2[1 - 0.64058] = 0.7188$ . Since the p-value  $> 0.05$ , we fail to reject  $H_0$  and conclude the yield is not significantly different from 90%.

$$\text{b) } n = \frac{(z_{\alpha/2} + z_{\beta})^2 \sigma^2}{\delta^2} = \frac{(z_{0.025} + z_{0.05})^2 3^2}{(85 - 90)^2} = \frac{(1.96 + 1.65)^2 9}{(-5)^2} = 4.67$$

$$n \cong 5.$$

14. The life in hours of a thermocouple used in a furnace is known to be approximately normally distributed with standard deviation  $\sigma = 20$  hours. A random sample 15 thermocouples resulted in the following data: 553, 552, 567, 579, 550, 541, 537, 553, 552, 546, 538, 553, 581, 539, 529.
- Is there evidence to support the claim that mean life exceeds 540 hours? Use a fixed level test with  $\alpha = 0.05$ ?

14) a) 1) The parameter of interest is the true mean life,  $\mu$ .

2)  $H_0: \mu = 540$

3)  $H_1: \mu > 540$

4)  $\alpha = 0.05$

5)  $z_0 = \frac{\bar{x} - \mu}{\sigma / \sqrt{n}}$

6) Reject  $H_0$  if  $z_0 > z_\alpha$  where  $z_{0.05} = 1.645$

7)  $\bar{x} = 551.33, \sigma = 20$

$$z_0 = \frac{551.33 - 540}{20 / \sqrt{15}} = 2.19$$

8) Since  $2.19 > 1.645$ , reject the null hypothesis and conclude there is sufficient evidence to support the claim the

life exceeds 540 hrs at  $\alpha = 0.05$ .

26. The dynamic modulus of concrete is obtained for two different concrete mixes. For the first mix, an  $n = 33$  sample has a mean of 115.1 and a standard deviation of 0.47 psi. For the second mix, an  $n = 31$  sample has a mean of 114.6 and a standard deviation of 0.38 psi. Test at a 0.05 significance level the null hypothesis of equality of mean dynamic modules versus the two sided alternative.

1. *Null hypothesis*  $H_0 : \mu_1 - \mu_2 = 0$

2. *Level of significance*:  $\alpha = 0.05$ .

*Alternative hypothesis*  $H_1 : \mu_1 - \mu_2 \neq 0$

3. *Criterion*: The null hypothesis specifies  $\delta = \mu_1 - \mu_0 = 0$ . Since the samples are large, we use the large sample statistic where we estimate each population variance by the sample variance.

$$Z = \frac{\bar{X} - \bar{Y} - \delta}{\sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}}$$

The alternative is two-sided so we reject the null hypothesis for  $Z > z_{.025}$  or  $Z < -z_{.025}$

4. *Calculations*: Since  $n_1 = 33$ ,  $n_2 = 31$ ,  $\bar{x} = 115.1$ ,  $\bar{y} = 114.6$ ,  $s_1 = 0.47$ , and  $s_2 = 0.38$

$$\sqrt{\frac{.47^2}{33} + \frac{0.38^2}{31}} = 0.10655$$

$$z = \frac{115.1 - 114.6}{0.10655} = 4.69 > 1.96,$$

5. *Decision*: Because  $4.69 > 1.96$ , we reject the null hypothesis at the .05 level of significance.

The P-value  $2P[Z > 4.69]$  rounds to 0.00000 and gives extremely strong support for rejecting the null hypothesis.



28. Two machines are used for filling plastic bottles with a net volume of 16.0 ounces. The fill volume can be assumed normal, with standard deviation 0.020 and 0.025 for machines 1 and 2, respectively. A member of the quality engineering staff suspects that both machines fill to the same mean net volume, whether or not this volume is 16.0 ounces. A random sample of 10 bottles is taken from the output of each machine, as follows:

Machine 1		Machine 2	
16.03	16.01	16.02	16.03
16.04	15.96	15.97	16.04
16.05	15.98	15.96	16.02
16.05	16.02	16.01	16.01
16.02	15.99	15.99	16.00

- Do you think the engineer is correct? Use the P-value approach.
- If  $\alpha = 0.05$ , what is the power of the test in Part a) for a true difference in means of 0.04?
- Find a 95% CI on the difference in means. Provide a practical interpretation of this interval.
- Assuming equal sample sizes, what sample size should be used to ensure that  $\beta = 0.01$  if the true difference in means is 0.04? Assume that  $\alpha = 0.05$ .



a) The parameter of interest is the difference in fill volume,  $\mu_1 - \mu_2$   $\delta = 0$  The test statistic is

$$H_0: \mu_1 - \mu_2 = 0 \text{ or } \mu_1 = \mu_2 \quad z_0 = \frac{(\bar{x}_1 - \bar{x}_2) - \Delta_0}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}} \quad \bar{x}_1 = 16.015 \quad \bar{x}_2 = 16.005 \quad z_0 = \frac{(16.015 - 16.005) - 0}{\sqrt{\frac{(0.02)^2}{10} + \frac{(0.025)^2}{10}}} = 0.99$$

$$H_1: \mu_1 - \mu_2 \neq 0 \text{ or } \mu_1 \neq \mu_2 \quad \sigma_1 = 0.02 \quad \sigma_2 = 0.025 \quad n_1 = 10 \quad n_2 = 10$$

P-value =  $2(1 - \Phi(0.99)) = 2(1 - 0.8389) = 0.3222$ . Since p-value is greater than 0.05, we do not reject the null hypothesis.

b) Power =  $1 - \beta$ , where

$$\beta = \Phi\left(z_{\alpha/2} - \frac{\Delta - \Delta_0}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}\right) - \Phi\left(-z_{\alpha/2} - \frac{\Delta - \Delta_0}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}\right) = \Phi\left(1.96 - \frac{0.04}{\sqrt{\frac{(0.02)^2}{10} + \frac{(0.025)^2}{10}}}\right) - \Phi\left(-1.96 - \frac{0.04}{\sqrt{\frac{(0.02)^2}{10} + \frac{(0.025)^2}{10}}}\right)$$

$$= \Phi(1.96 - 3.95) - \Phi(-1.96 - 3.95) = \Phi(-1.99) - \Phi(-5.91) = 0.0233 - 0 = 0.0233$$

$$\text{Power} = 1 - 0.0233 = 0.977$$

c)  $(\bar{x}_1 - \bar{x}_2) - z_{\alpha/2} \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}} \leq \mu_1 - \mu_2 \leq (\bar{x}_1 - \bar{x}_2) + z_{\alpha/2} \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}$

$$(16.015 - 16.005) - 1.96 \sqrt{\frac{(0.02)^2}{10} + \frac{(0.025)^2}{10}} \leq \mu_1 - \mu_2 \leq (16.015 - 16.005) + 1.96 \sqrt{\frac{(0.02)^2}{10} + \frac{(0.025)^2}{10}}$$

$$-0.0098 \leq \mu_1 - \mu_2 \leq 0.0298$$

With 95% confidence, we believe the true difference in the mean fill volumes is between  $-0.0098$  and  $0.0298$ . Since 0 is contained in this interval, we can conclude there is no significant difference between the means.

d) Assume the sample sizes are to be equal, use  $\alpha = 0.05$ ,  $\beta = 0.01$ , and  $\Delta = 0.04$

$$n \cong \frac{(z_{\alpha/2} + z_{\beta})^2 (\sigma_1^2 + \sigma_2^2)}{\delta^2} = \frac{(1.96 + 2.33)^2 ((0.02)^2 + (0.025)^2)}{(0.04)^2} = 2.08, \quad n = 11.79, \text{ use } n_1 = n_2 = 12$$

32. The melting points of two alloys used in formulating solder were investigated by melting 21 samples of each material. The sample mean and standard deviation of alloy 1 was  $\bar{x}_1=420.48$ ,  $s_1= 2.34$ , and for alloy 2 they were  $\bar{x}_2=425$ ,  $s_2= 2.5$
- Do the sample data support the claim that both alloys have the same melting point? Use a fixed-level test with  $\alpha = 0.05$  and assume that both populations are normally distributed and have the same standard deviation.
  - Find the P-value for this test.

a) 1) The parameter of interest is the difference in mean melting point,  $\mu_1 - \mu_2$

2)  $H_0: \mu_1 - \mu_2 = 0$  or  $\mu_1 = \mu_2$       5) The test statistic is

3)  $H_1: \mu_1 - \mu_2 \neq 0$  or  $\mu_1 \neq \mu_2$

4)  $\alpha = 0.05$

$$t_0 = \frac{(\bar{x}_1 - \bar{x}_2) - \Delta_0}{s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$$

6) Reject the null hypothesis if  $t_0 < -t_{\alpha/2, n_1+n_2-2}$  where  $-t_{0.025, 40} = -2.021$  or  $t_0 > t_{\alpha/2, n_1+n_2-2}$

where  $t_{0.025, 40} = 2.021$

$$7) \bar{x}_1 = 420.48 \quad \bar{x}_2 = 425, \quad \Delta_0 = 0 \quad s_p = \sqrt{\frac{(n_1 - 1)s_1^2 + (n_2 - 1)s_2^2}{n_1 + n_2 - 2}} = \sqrt{\frac{20(2.34)^2 + 20(2.5)^2}{40}} = 2.42$$

$$s_1 = 2.34$$

$$s_2 = 2.5$$

$$t_0 = \frac{(420.48 - 425) - 0}{2.42 \sqrt{\frac{1}{20} + \frac{1}{20}}} = -5.99$$

$$n_1 = 21$$

$$n_2 = 21$$

8) Since  $-5.99 < -2.021$  reject the null hypothesis and conclude that the data do not support the claim that both alloys have the same melting point at  $\alpha = 0.05$

b) P-value =  $2P(t < -5.424)$     P-value  $< 0.001$