

A simple two parameter distribution for modelling neuronal activity and capturing neuronal association

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Abstract

Recent developments in electrophysiological technology have lead to an increase in the size of electrophysiological datasets. Consequently, there is a requirement for new analysis techniques that can make use of these new datasets, while remaining easy to use in practice. In this work, we fit the Conway-Maxwell-binomial distribution to spiking data read from a mouse exposed to visual stimuli.

1 Introduction

Motivate by pointing out how much computational power it can require to calculate n th order correlations.

Point out that we don't necessarily need to measure correlations anyway.

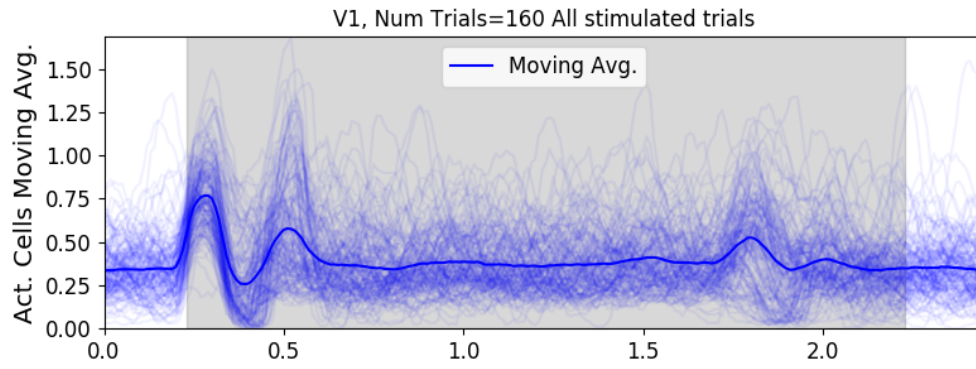


Figure 1: A moving average of the number of active neurons in V1. Averages were taken across 100ms windows split into 100 bins. The midpoint of the time interval for each window is used as the timepoint (x-axis point) for that window. The shaded area indicates the presence of a visual stimulus. Translucent lines indicate the moving averages for each of the 160 trials, the opaque line is an average across trials. We can see a transient increase in the average number of active neurons, followed by a fluctuation and another increase.

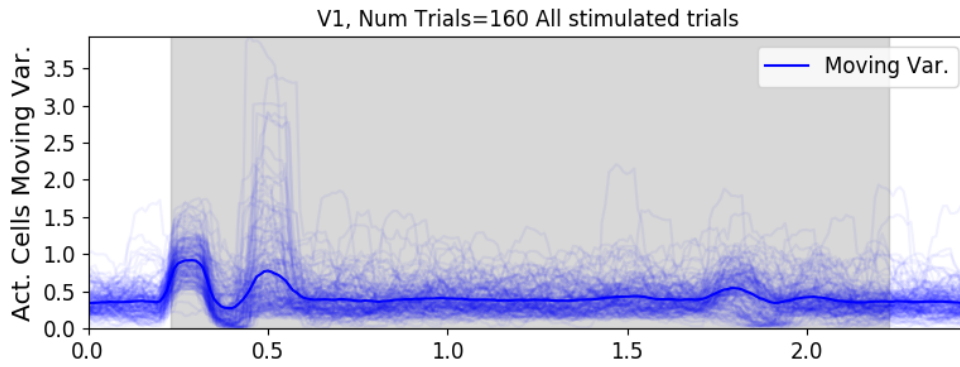


Figure 2: A moving variance of the number of active neurons in V1. Variances were taken across 100ms windows split into 100 bins. The midpoint of the time interval for each window is used as the timepoint (x-axis point) for that window. The shaded area indicates the presence of a visual stimulus. Translucent lines indicate the moving variances for each of the 160 trials, the opaque line is an average across trials. We can see a transient increase in the variance of the number of active neurons, followed by a fluctuation and another increase.

2 Results

2.1 Transient increases in mean number of active neurons and variance in number of active neurons at stimulus onset

2.2 Conway-Maxwell-binomial distribution fits better than binomial or beta-binomial

2.3 Conway-Maxwell-binomial distribution captures changes in association at stimulus onset

2.4 Reproduced stimulus related quenching of neural variability

3 Discussion

4 Data

We used data collected by Nick Steinmetz and his lab ‘CortexLab at UCL’ [9]. The data can be found online ¹ and are free to use for research purposes.

¹<http://data.cortexlab.net/dualPhase3/>

Two ‘Phase3’ Neuropixels [4] electrode arrays were inserted into the brain of an awake, head-fixed mouse for about an hour and a half. These electrode arrays recorded 384 channels of neural data each at 30kHz and less than $7\mu\text{V}$ RMS noise levels. The sites are densely spaced in a ‘continuous tetrode’-like arrangement, and a whole array records from a 3.8mm span of the brain. One array recorded from visual cortex, hippocampus, and thalamus, the other array recorded from motor cortex and striatum. The data were spike-sorted automatically by Kilosort and manually by N. Steinmetz using Phy. In total 831 well-isolated individual neurons were identified.

4.1 Experimental protocol

The mouse was shown a visual stimulus on three monitors placed around the mouse at right angles to each other, covering about ± 135 degrees azimuth and ± 35 degrees elevation.

The stimulus consisted of sine-wave modulated full-field drifting gratings of 16 drift directions ($0^\circ, 22.5^\circ, \dots, 337.5^\circ$) with 2Hz temporal frequency and 0.08 cycles/degree spatial frequency displayed for 2 seconds plus a blank condition. Each of these 17 conditions were presented 10 times in a random order across 170 different trials. There were therefore 160 trials with a visual stimulus present, and 10 trials without a visual stimulus.

5 Methods

Details about all kinds of things here.

5.1 Binning data

We converted the spike times for each cell into spike counts by putting the spike times into time bins of a given ‘width’ (in seconds). We used time bins of 1ms, 5ms, and 10ms. We used different time bin widths to assess the impact of choosing a bin width.

5.2 Number of *active* neurons

To count the number of active neurons in each neuronal ensemble, we split the time interval for each trial into bins of a given width. We counted the number of spikes fired by each cell in each bin. If a cell fired *at least* one spike in a given bin, we regarded that cell as active in that bin. We recorded the number of active cells in every bin, and we recorded each cell's individual spike counts.

It should be noted that when we used a bin width of 1ms, the maximum number of spikes in any bin was 1. For the wider time bins, some bins had spike counts greater than 1. Consequently when using a bin width of 1ms, the number of active neurons and the total spike count of a given bin were identical. But for wider bin widths, the total spike count was greater than the number of active neurons.

So for the 1ms bin width, the activity of a neuron and the number of spikes fired by that neuron in any bin can be modelled as a Bernoulli variable. But for wider time bins, only the activity can be modelled in this way.

5.3 Moving windows for measurements

When taking measurements (e.g. moving average over the number of active neurons) or fitting distributions (eg. the beta binomial distribution) we slid a window containing a certain number of bins across the data, and made our measurements at each window position. For example, when analysing 1ms bin data, we used a window containing 100 bins, and we slid the window across the time interval for each trial moving 10 bins at a time. So that for 2560ms of data, we made 246 measurements.

For the 5ms bin width data, we used windows containing 40 bins, and slid the window 2 bins at a time when taking measurements.

For the 10ms bin width data, we used windows containing 40 bins, and slid the window 1 bin at a time when taking measurements.

By continuing to use windows containing 40 bins, we retained statistical power but sacrificed the number of measurements taken.

There was an interval between each trial with a grey image in place of the moving of the moving bar stimulus. This interval varied in time. But we included some of this interval when recording the data for each trial. We started recording the number of active neurons, and the number of spikes from each neuron from 280ms before

each trial until 280ms after each trial. This way, we could see the change in our measurements at the onset of a stimulus.

The actual measurements we took in each window were as follows:

Number of active neurons The number of neurons firing a spike in each bin. Most of the other measurements are aggregations of these measurements, or models fitted to these measurements.

Spike counts for each cell The number of spikes fired by each cell in each bin.

Moving average The average number of active cells in each window.

Moving variance The variance of the number of active cells in each window.

Average correlation We measured the correlation between the spike counts of each pair of cells in the ensemble, and took the average of these measurements.

Binomial p We fitted a binomial distribution to the data in each window and recorded the fitted probability of success, p in each case.

Beta-binomial α, β We fitted a beta-binomial distribution to the data in each window, and recorded the values of the fitted shape parameters, α and β , of each distribution.

Conway-Maxwell-binomial distribution p, ν We fitted a Conway-Maxwell-binomial distribution to the data in each window, and recorded the fitted values of p and ν for each distribution.

Log-likelihoods We also recorded the log-likelihood of each of the fitted distributions for each window.

5.4 Fano factor

The *Fano factor* of a random variable is defined as the ratio of the variable's variance to its mean.

$$F = \frac{\sigma^2}{\mu} \quad (1)$$

We measured the Fano factor of the spike count of a given cell by measuring the mean and variance of the spike count across trials, and taking the ratio of those two quantities. When calculated in this way the Fano factor can be used as a measure of neural variability. This is similar to the calculation used in [2].

5.5 Probability Distributions suitable for modelling ensemble activity

We present here three different probability distributions that could be suitable to model the number of active neurons in an ensemble. Each distribution has the set $\{0, \dots, n\}$ as its support, where n is the number of neurons in the ensemble. These are simple distributions with either two or three parameters each. However, we regard n as known when using these distributions for modelling, so in effect each distribution has either one or two free parameters.

5.5.1 Association

Association between random variables is similar to the correlation between random variables but is more general in concept. The correlation is a measure of association; and association doesn't have a mathematical definition like correlation does. Essentially, the association between two random variables is their tendency to take the same or similar values. Positively associated variables tend to take the same value, and negatively associated variables tend to take different values. In this research, we work with probability distributions of the number of successes in a set of Bernoulli trials. These Bernoulli variables may or may not be associated.

A probability distribution over the number of successes in n Bernoulli trials, where the Bernoulli variables may be associated, could constitute a good model for the number of active neurons in an ensemble of n neurons.

5.5.2 Binomial distribution

The binomial distribution is a two parameter discrete probability distribution that can be thought of as a probability distribution the number of successes from n independent Bernoulli trials, each with the same probability of success. The parameters of the binomial distribution are n , and $0 \leq p \leq 1$, the probability of success for each of these trials. A random variable with the binomial distribution can take values from $\{0, \dots, n\}$. The probability mass function of the distribution is

$$P(k; n, p) = \binom{n}{k} p^k (1 - p)^{n-k} \quad (2)$$

133 As model for the activity of a neuronal ensemble, the main problem with the bi-
 134 nomial distribution is that it treats each neuron, represented as a Bernoulli trial, as
 135 independent. It is well known that neurons are not independent, and that correlated
 136 behaviour between neurons is vital for representing sensory information. The bino-
 137 mial distribution falls short in this regard, but it is useful as performance benchmark
 138 when assessing the performance of other models.

139 5.5.3 Beta-binomial distribution

140 The beta distribution is the conjugate distribution of the binomial distribution. The
 141 beta-binomial distribution is the combination of the beta distribution and the bino-
 142 mial distribution, in that the probability of success for the binomial distribution is
 143 sampled from the beta distribution. This allows the beta-binomial distribution to
 144 capture some over dispersion relative to the binomial distribution.

The beta-binomial distribution is a three parameter distribution, n the number
 of Bernoulli trials, and $\alpha \in \mathbb{R}_{>0}$ and $\beta \in \mathbb{R}_{>0}$ the shape parameters of the beta
 distribution. The probability mass function for the beta-binomial distribution is

$$P(k; n, \alpha, \beta) = \binom{n}{k} \frac{B(k + \alpha, n - k + \beta)}{B(\alpha, \beta)} \quad (3)$$

145 where $B(\alpha, \beta)$ is the beta function.

This probability distribution can be reparametrised in a number of ways. One of
 which defines new parameters π and ρ by

$$\pi = \frac{\alpha}{\alpha + \beta} \quad (4)$$

$$\rho = \frac{1}{\alpha + \beta + 1} \quad (5)$$

146 This reparametrisation is useful because π acts as a location parameter analogous to
 147 the p parameter of a binomial distribution. A value of $\rho > 0$ indicates over-dispersion
 148 relative to a binomial distribution.

149 As a model for the activity of a neuronal ensemble, the beta-binomial distribution
 150 is more suitable than a binomial distribution because the over-dispersion of the beta-
 151 binomial distribution can be used to model positive association between the neurons.
 152 An extreme example of this over-dispersion/positive association can be seen in figure

153 3b. In this figure, the neurons are positively associated and so tend to take the same
 154 value, consequently the probability mass of the beta-binomial distribution builds up
 155 close to $k = 0$ and $k = n$. It is worth noting that the location parameter for each
 156 distribution has the same value, $p = \pi = 0.5$.

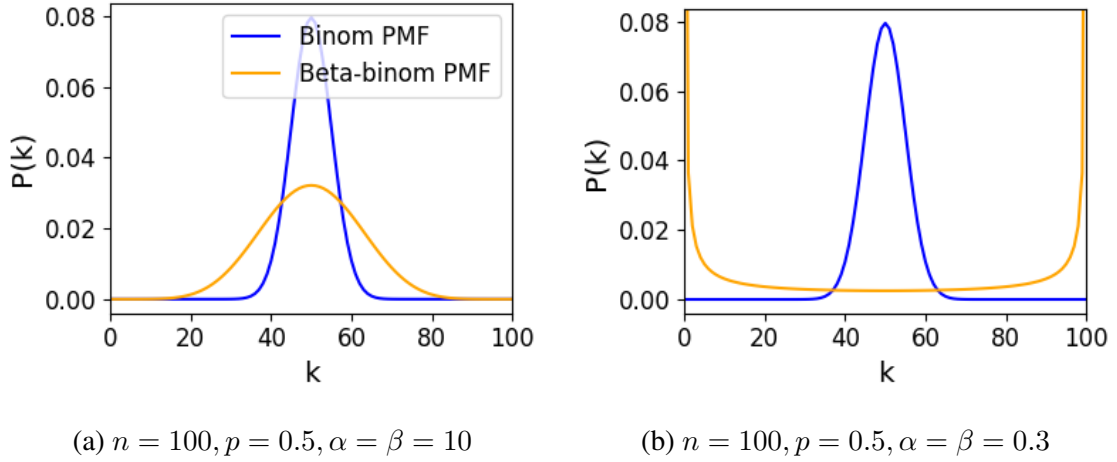


Figure 3: Figures showing the over-dispersion possible for a beta-binomial distribution relative to a binomial distribution. Parameters are shown in the captions.

157 5.5.4 Conway-Maxwell-binomial distribution

158 The Conway-Maxwell-binomial distribution (COMb distribution) is a three param-
 159 eter generalisation of the binomial distribution that allows for over dispersion and
 160 under dispersion relative to the binomial distribution. The parameters of the distri-
 161 bution are n the number of Bernoulli trials, $0 \leq p \leq 1$, the location parameter, and
 162 $\nu \in \mathbb{R}$ the shape parameter.

The probability mass function of the COMb distribution is

$$P(k; n, p, \nu) = \frac{1}{S(n, p, \nu)} \binom{n}{k}^\nu p^k (1-p)^{n-k} \quad (6)$$

where

$$S(n, p, \nu) = \sum_{j=0}^n \binom{n}{j}^\nu p^j (1-p)^{n-j} \quad (7)$$

163 The only difference between this PMF and the PMF for the standard binomial is
 164 the introduction of ν and the consequent introduction of the normalising function
 165 $S(n, p, \nu)$.

Indeed, if $\nu = 1$ the COMb distribution is identical to the binomial distribution with the same values for n and p . We can see in figure 4d that the KL-divergence $D_{KL}(P_{COMb}(n, p, \nu), P_{Bin}(n, p)) = 0$ along the line where $\nu = 1$.

If $\nu < 1$ the COMb distribution will exhibit over-dispersion relative to the binomial distribution. If $p = 0.5$ and $\nu = 0$ the COMb distribution is the discrete uniform distribution, and if $\nu < 0$ the mass of the COMb distribution will tend to build up near $k = 0$ and $k = n$. This over-dispersion represents positive association in the Bernoulli variables. An example of this over-dispersion can be seen in figure 4b.

If $\nu > 1$ the COMb distribution will exhibit under-dispersion relative to the binomial distribution. The larger the value of ν the more probability mass will build up at $n/2$ for even n , or at $\lfloor n/2 \rfloor$ and $\lceil n/2 \rceil$ for odd n . This under-dispersion represents negative association in the Bernoulli variables. An example of this under-dispersion can be seen in figure 4a.

It should be noted that the shape parameter of the COMb distribution, p , does not correspond to the mean of the distribution, as is the case for the binomial and beta-binomial distributions. An illustration of this can be seen in figure 4c. This is because an interaction between the p and ν parameters skews the mean. There is no analytical expression for the mean of the COMb distribution.

ν	Relative dispersion	Associaton between neurons/variables
< 1	over	positive
0	none	none
> 1	under	negative

Since the COMb distribution has the potential to capture positive and negative associatons between the neurons/Bernoulli variables, it should be an excellent candidate for modelling the number of active neurons in a neuronal ensemble.

We wrote a dedicated Python package to enable easy creation and fitting of COMb distribution objects. The format of the package imitates the format of other distribution objects from the `scipy.stats` Python package. The COMb package can be found here:

https://github.com/thomasjdelaney/Conway_Maxwell_Binomial_Distribution

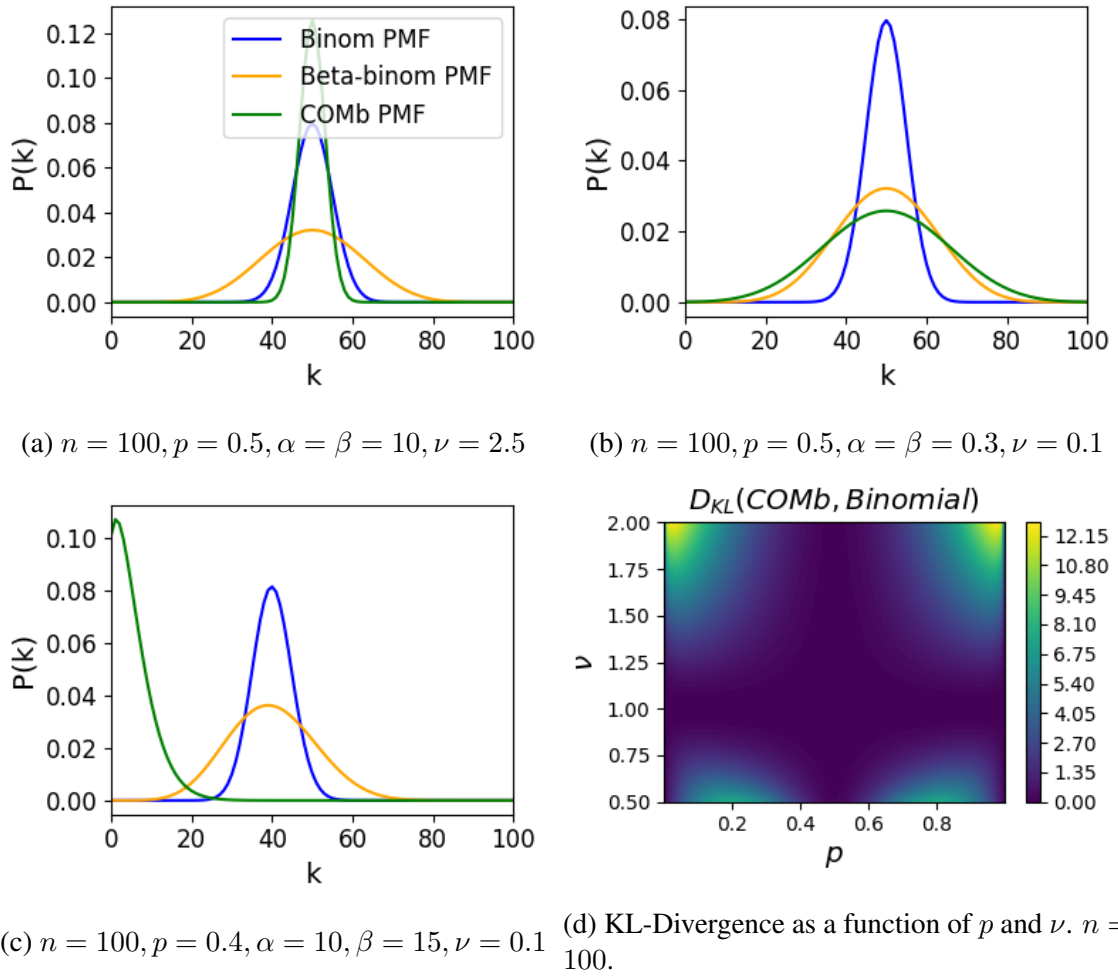


Figure 4: Figures showing (a) the over-dispersion and (b) under-dispersion permitted by the COMb distribution. (c) illustrates that the shape parameter of the COMb distribution, p , does not correspond to the mean of the distribution, as it does for the binomial and beta-binomial distributions. (d) shows a heatmap for the value of the Kullback-Liebler divergence between the COMb distribution and the standard binomial distribution with same value for n , as a function of p and ν . Parameters are shown in the captions.

5.6 Fitting

We fitted a binomial, beta-binomial, and Conway-Maxwell-binomial (COMb) distribution to the neural activity in each of the overlapping windows covering each trial. To fit the distributions we minimised the appropriate negative log likelihood function using the data from the window.

There is an analytical solution for the binomial distribution.

$$\hat{p} = \frac{1}{n} \sum_{i=1}^N k_i \quad (8)$$

We minimised the negative log likelihood function of the beta-binomial distribution numerically.

The log likelihood function of the COMb distribution is

$$\ell(p, \nu | k_1, \dots, k_N) = N [n \log(1 - p) - \log S(n, p, \nu)] \quad (9)$$

$$+ \log \frac{p}{1 - p} \sum_{i=1}^N k_i \quad (10)$$

$$+ \nu \sum_{i=1}^N \log \binom{n}{k_i} \quad (11)$$

We minimised the negation of this function using numerical methods.

More specifically, we used the `minimize` function of the `scipy.optimize` Python package.

5.7 Goodness-of-fit

After fitting, we measured the goodness-of-fit of each model/distribution with their log likelihood. We calculated this directly using the `logpmf` functions of the distribution objects in Python.

6 Discussion

Point out that the Conway-Maxwell-binomial distribution could be used to measure activity and association without having to sort the voltage traces into spikes. That does defeat the purpose slightly, however.

References

- [1] Patricia M. E. Altham, *Two Generalizations of the Binomial Distribution*. Journal of the Royal Statistical Society 27, 162-167, (1978)
- [2] Mark M Churchland, Byron M Yu, John P Cunningham, Leo P Sugrue, Marlene R Cohen, Greg S Corrado, William T Newsome, Andrew M Clark, Paymon Hosseini, Benjamin B Scott, David C Bradley, Matthew A Smith, Adam Kohn, J Anthony Movshon, Katherine M Armstrong, Tirin Moore, Steve W Chang, Lawrence H Snyder, Stephen G Lisberger, Nicholas J Priebe, Ian M Finn, David Ferster, Stephen I Ryu, Gopal Santhanam, Maneesh Sahani, Krishna V Shenoy, *Stimulus onset quenches neural variability: A widespread cortical phenomenon*. Nature Neuroscience 13, 369-378, (2010)
- [3] Fraser Daly, Robert E. Gaunt, *The Conway-Maxwell-Poisson distribution: Distributional theory and approximation*. ALEA 13, 635-658, (2016)
- [4] James J. Jun, Nicholas A. Steinmetz, Joshua H. Siegle, Daniel J. Denman 5, Marius Bauza, Brian Barbarits, Albert K. Lee, Costas A. Anastassiou, Alexandru Andrei, Çağatay Aydın, Mladen Barbic, Timothy J. Blanche, Vincent Bonin, João Couto, Barundeb Dutta, Sergey L. Gratiy, Diego A. Gutnisky, Michael Häusser, Bill Karsh, Peter Ledochowitsch, Carolina Mora Lopez, Catalin Mite-lut, Silke Musa, Michael Okun, Marius Pachitariu, Jan Putzeys, P. Dylan Rich, Cyrille Rossant, Wei-lung Sun, Karel Svoboda, Matteo Carandini, Kenneth D. Harris, Christof Koch, John O’Keefe, Timothy D. Harris, *Fully integrated silicon probes for high-density recording of neural activity*. Nature 551, 232-236, (2017)
- [5] Joseph B. Kadane, *Sums of Possibly Associated Bernoulli Variables: The Conway–Maxwell–Binomial Distribution*. Bayesian Analysis 11, 403-420, (2016)
- [6] Joseph B. Kadane, Galit Shmueli, Thomas P. Minka, Sharad Borle, Peter Boatwright, *Conjugate Analysis of the Conway-Maxwell-Poisson Distribution*. Bayesian Analysis 1, (2006)
- [7] R. L. Prentice, *Binary Regression Using an Extended Beta-Binomial Distribution, With Discussion of Correlation Induced by Covariate Measurement Errors*. Journal of American Statistical Association 81, 321-327, (1986)

- 245 [8] Galit Shmueli, Thomas P. Minka, Joseph B. Kadane, Sharad Borle, Peter
246 Boatwright, *A useful distribution for fitting discrete data: revival of the Con-*
247 *way–Maxwell–Poisson distribution*. Applied Statistics 54, 127-142, (2005)
- 248 [9] Nick Steinmetz, Matteo Carandini, Kenneth D. Harris, *"Single*
249 *Phase3" and "Dual Phase3" Neuropixels Datasets*. figshare, Dataset:
250 <https://doi.org/10.6084/m9.figshare.7666892.v2> (2019)
- 251 [10] Joseph S. Verducci, Michael E. Mack, Morris H. DeGroot, *Estimating multiple*
252 *rater agreement for a rare diagnosis*. Journal of Multivariate Analysis 27, 512-
253 535, (1988)