# Least Square Method of Estimation

#### Least Square Method of Estimation

- Since method-of-moments performs poorly for some models, we examine another method of parameter estimation: Least Squares.
- We first consider autoregressive models.
- We assume our time series is stationary (or that the time series has been transformed so that the transformed data can be modeled as stationary).
- ▶ To account for the possibility that the mean is nonzero, we subtract  $\mu$  from each observation and treat  $\mu$  as a parameter to be estimated.

#### LS Estimation for the AR(1) Model

ightharpoonup Consider the mean-centered AR(1) model:

$$\underbrace{Y_t - \mu = \phi(Y_{t-1} - \mu) + \underbrace{\varepsilon_t}_{e_t}}_{=} = \underbrace{Y_t - \mu}_{=} + \underbrace{Y_t - \mu}_{=} - \underbrace{Y_t$$

► The least squares method seeks the parameter values that minimize the sum of squared differences:

$$S_{c}(\phi,\mu) = \sum_{t=2}^{n} [(Y_{t} - \mu) - \phi(Y_{t-1} - \mu)]^{2}$$

This criterion is called the conditional sum-of-squares function (CSS).

$$\frac{1}{3\mu} = 2 \left[ \frac{1}{2} \left( \frac{1}{3} - \mu \right) - \phi \left( \frac{1}{3} - \mu \right) \left[ \frac{1}{3} - \phi \left( \frac{1}{3} - \mu \right) \right] = 0$$

# LS Estimation of $\mu$ for the AR(1) Model

► Taking the derivative of CSS with respect to  $\mu$ , setting equal to 0 and solving for  $\mu$ , we obtain the LS estimator of  $\mu$ :

$$\hat{\mu} = \frac{1}{(n-1)(1-\phi)} \left[ \sum_{t=2}^{n} Y_t - \phi \sum_{t=2}^{n} Y_{t-1} \right]$$

▶ For large n, this  $\hat{\mu} \approx \bar{Y}$ , regardless of the value of  $\phi$ .

# LS Estimation of $\phi$ for the AR(1) Model

▶ Taking the derivative of CSS with respect to  $\phi$ , setting equal to 0 and solving for  $\phi$ , we obtain the LS estimator of  $\phi$ :

$$\hat{\phi} = \frac{\sum_{t=2}^{n} (Y_t - \bar{Y})(Y_{t-1} - \bar{Y})}{\sum_{t=2}^{n} (Y_{t-1} - \bar{Y})^2}$$

- ► This estimator is almost identical to  $r_1$ : it's just missing one term in the denominator,  $(Y_n \bar{Y})^2$ .
- So, especially for large n, the LS and MOM estimators are nearly identical in the AR(1) model.
- In the general AR(p) model, the LS estimators of  $\mu$  and of  $\phi_1, \ldots, \phi_p$  are approximately equal to the MOM estimators, especially for large samples.

#### LS Estimation for Moving Average Models

Consider now the MA(1) model:

$$Y_t = e_t - \theta e_{t-1}$$
 or  $\xi_t - 0 \xi_{t-1}$ 

Recall that this can be written as

$$Y_t = \frac{-\theta Y_{t-1} - \theta^2 Y_{t-2} - \theta^3 Y_{t-3} - \dots + e_t}{e_t}$$

▶ So a least squares estimator of  $\theta$  can be obtained by finding the value of  $\theta$  that minimizes

$$S_c(\theta) = \sum [Y_t + \theta Y_{t-1} + \theta^2 Y_{t-2} + \theta^3 Y_{t-3} + \cdots]^2$$

▶ But this is nonlinear in  $\theta$ , and the infinite series causes technical problems.

# LS Estimation for Moving Average Models

Instead, we proceed by conditioning on one previous value of e<sub>t</sub>. Note that

$$e_t = Y_t + \theta e_{t-1}$$

- If we set  $e_0 = 0$  then we have the set of recursive equations  $e_1 = Y_1$ ,  $e_2 = Y_2 + \theta e_1, \dots, e_n = Y_n + \theta e_{n-1}$ .
- Since we know  $Y_1, Y_2, \ldots, Y_n$  (these are the observed data values) and can calculate the  $e_1, e_2, \ldots, e_n$  recursively, the only unknown quantity here is  $\theta$ .
- We can do a numerical search for the value of  $\theta$  (within the invertible range between -1 and 1) that minimizes  $\sum (e_t)^2$ , conditional on  $e_0 = 0$ .
- A similar approach works for higher-order  $\underline{MA}(q)$  models, except that we assume  $\underline{e_0} = \underline{e_{-1}} = \cdots = \underline{e_{-q}} = 0$  and the numerical search is multidimensional, since we are estimating  $\theta_1, \ldots, \theta_q$ .

# LS Estimation for ARMA Models

▶ With the *ARMA*(1,1) model:

$$Y_t = \phi Y_{t-1} + e_t - \theta e_{t-1},$$

we note that

$$e_t = Y_t - \phi Y_{t-1} + \theta e_{t-1}$$

 $\underbrace{e_t = Y_t - \phi Y_{t-1} + \theta e_{t-1}}_{\text{and minimize } S_c(\phi,\theta) = \sum_{t=2}^n e_t^2; \text{ note that the sum starts at}}$  $\underline{t=2}$  to avoid having to choose an "initial" value  $\underline{Y_0}$ .

- ▶ With the general ARMA(p, q) model, the procedure is similar, except that we assume  $e_p = e_{p-1} = \cdots = e_{p+1-q} = 0$ , and we estimate  $\phi_1, \ldots, \phi_p, \theta_1, \ldots, \theta_q$ .
- For *large <u>samples</u>*, when the parameter sets yield invertible models, the initial values for  $e_p, e_{p-1}, \ldots, e_{p+1-q}$  have little effect on the final parameter estimates.