MAFS5130: Time series analysis

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1 Exercises

- 1. Find the ACF and PACF and plot the ACF ρ_k for k = 0, 1, 2, 3, 4 and 5 for each of the following models:
 - (a) $r_t 0.5r_{t-1} = a_t$;
 - (b) $r_t + 0.98r_{t-1} = a_t$;
 - (c) $r_t 1.3r_{t-1} + 0.4r_{t-2} = a_t$.

Solution:

• ACF:

$$E(r_t * r_{t-k})$$

• PACF:

- 2. For each of the following models,
 - (a) $(1 0.9B) (r_t 10) = a_t$;
 - (b) $r_t = 10 0.9a_{t-1} + a_t$;
 - (c) $(1 0.5B)(r_t 10) = a_t 0.9a_{t-1}$.

where $\sigma_a^2 = 2$. Given $r_1 = 1.2$ and $r_2 = 0.1$, find the l-step ahead forecast values and forecast variances for l = 1, 2, 3, 4.

Solution:

For AR(1) model:

• l-step ahead forecast:

$$\hat{r_t}(l) = E(r_{t+l}|\mathcal{F}_t)$$

• 1-step forecast variance:

$$e_t(l) = r_{t+l} - \hat{r}_t(l)$$

$$\operatorname{Var}\left[e_t(l)\right] = \sigma_a^2 \left(1 + \phi^2 + \phi^4 + \cdots\right)$$

- 3. Find the ACF and PACF for k = 0, 1, 2, 3 and 4 for each of the following models:
 - (a) $r_t = (1 0.8B)a_t$;
 - (b) $r_t = (1 1.2B + 0.5B^2) a_t$.

Solution:

- ACF
- PACF

4. Verify whether or not the following models are stationary and / or invertible:

- (a) $(1-B)r_t = (1-1.5B)a_t$;
- (b) $(1 0.8B)r_t = (1 0.5B)a_t$;
- (c) $(1 1.1B + 0.8B^2) r_t = (1 1.7B + 0.72B^2) a_t$
- (d) $(1 0.6B)r_t = (1 1.2B + 0.2B^2) a_t$

Solution:

For AR(1), MA(1), ARMA(1,1),

- Invertibility condition: $|\theta| < 1$
- Stationarity condition: $|\phi| < 1$

For other model, solve equation like $1 - 1.1z + 0.8z^2 = 0$ in (c), if the roots lies outside the unit circle, then it's stationary or invertible; otherwise it's not.

5. Consider the two models:

- (a) $(1 0.43B)(1 B)r_t = a_t$;
- (b) $(1-B)r_t = (1-0.43B)a_t$

where a_t is i.i.d. N(0,1). Given the observations $r_{49} = 33.4$ and $r_{50} = 33.9$, compute their forecasts $r_{50}(l)$, for l = 1, 2, 3, 4, and the corresponding 90% forecast intervals.

Solution:

- Forecast values: $\hat{r}_{50}(1) = E(r_{51}|\mathcal{F}_t) = 34.115$
- Forecast variance: $\operatorname{Var}\left[e_n(l)\right] = \sigma_a^2 \sum_{i=0}^{l-1} \phi^{2i}$
- Forecast intervals: $\left[\widehat{r}_n(l) N_{\frac{\alpha}{2}}\sigma_a\sqrt{\sum_{j=0}^{l-1}\phi^{2j}},\widehat{r}_n(l) + N_{\frac{\alpha}{2}}\sigma_a\sqrt{\sum_{j=0}^{l-1}\phi^{2j}}\right]$ where $\alpha=0.1$ in this case and $N_{\frac{\alpha}{2}}$ is the $\frac{\alpha}{2}$ -quantile of the standard normal distribution.

6. Find the ACF for the following seasonal models:

- (a) $r_t = (1 \theta_1 B) (1 \Theta_1 B) a_t$;
- (b) $(1 \Phi_1 B^s) r_t = (1 \theta_1 B) a_t;$
- (c) $(1 \Phi_1 B^s) (1 \phi_1 B) r_t = a_t$

where a_t is i.i.d. N(0,1).

Solution:

7. Consider the ARCH model:

$$a_t = \eta_t \sigma_t, \sigma_t^2 = \alpha_0 + \alpha_1 a_{t-1}^2$$

Show that the unconditional variance of a_t is $Var(a_t) = \alpha_0/(1 - \alpha_1)$, where $\alpha_0 > 0, 0 \le \alpha_1 < 1$ and η_t is i.i.d. N(0, 1).

Solution: Use conditional expectation to solve the problem.

8. Give the stationarity condition and its representation in terms of $\{\eta_t\}$ for the GARCH model:

$$a_t = \eta_t \sigma_t, \sigma_t^2 = \alpha_0 + \alpha_1 a_{t-1}^2 + \beta_1 \sigma_{t-1}^2$$

where $\alpha_0 > 0, \alpha_1, \beta_1 \geq 0$, and η_t is i.i.d. N(0,1). Futhermore, give Ea_t^4 and the prediction of the conditional variances σ_{t+s}^2

Solution:

- Stationarity condition: $0 \le \alpha_1, \beta_1 \le 1, (\alpha_1 + \beta_1) < 1; [\sum_{i=1}^{\max(m,s)} (\alpha_i + \beta_i) < 1 \text{ for GARCH(m,s)}]$
- Representation in terms of $\{\eta_t\}$: $\sigma_t^2 = \alpha_0 \left[1 + \sum_{j=1}^{\infty} \prod_{i=1}^{j} \left(\alpha_1 \eta_{t-i}^2 + \beta_1\right)\right]$
- 1-step ahead forecast: $\sigma_t^2(1) = E\left(\sigma_{t+1}^2|\mathcal{F}_t\right) = \sigma_{t+1}^2 = \alpha_0 + \alpha_1 Z_t^2 + \beta_1 \sigma_t^2$

- $\bullet \ \ \text{2-step ahead forecast:} \ \sigma_t^2(2) = E\left(\sigma_{t+2}^2|\mathcal{F}_t\right) = \alpha_0 + E\left(\alpha_1\epsilon_{t+1}^2 + \beta_1|\mathcal{F}_t\right)\sigma_{t+1}^2 = \alpha_0 + (\alpha_1+\beta_1)\,\sigma_t^2(1)$
- In general: $\sigma_t^2(\ell) = \alpha_0 + (\alpha_1 + \beta_1) \, \sigma_t^2(\ell 1)$
- 9. Give the stationarity and invertibility conditions, MA and AR representation and ACFs of the seasonal ARMA models:
 - (a) $y_t = \phi y_{t-s} + a_t$;
 - (b) $y_t = \theta a_{t-s} + a_t$

where $\{a_t\}$ is white noise and variance σ_a^2 .

Solution:

- MA representation: $y_t = \frac{1}{1-\phi B^s} a_t = \sum_{j=0}^{\infty} (\phi B^s)^j a_t$
- AR representation: $at = \frac{y_t}{1+\theta B^s} = \sum_{j=0}^{\infty} (-\theta B^s)^j y_t$
- ACFs
- 10. Consider the following EGARCH(1,1) model

$$a_t = \sigma_t \epsilon_t$$
, $(1 - \beta B) \ln (\sigma_t^2) = \alpha_0 + \alpha g (\epsilon_{t-1})$

where $\epsilon_t \sim N(0,1)$ and $E\left(|\epsilon_t|\right) = \sqrt{2/\pi}$ and

$$g(\epsilon_t) = \theta \epsilon_t + [|\epsilon_t| - E(|\epsilon_t|)]$$

Show the representation of $\ln(\sigma_t^2)$ in terms of ϵ_t and give its mean and variance.

Solution:

• Expectation:

$$E(\ln(\sigma_t^2)) = E\left[\frac{1}{1-\beta B}(\alpha_0 + \alpha g(\epsilon_{t-1}))\right]$$

$$= E\left[\frac{\alpha_0}{1-\beta} + \alpha \sum_{i=0}^{\infty} \beta^i B^i g(\epsilon_{t-1})\right]$$

$$= E\left[\frac{\alpha_0}{1-\beta} + \alpha \sum_{i=0}^{\infty} \beta^i g(\epsilon_{t-i-1})\right]$$

$$= \frac{\alpha_0}{1-\beta} + \frac{\alpha}{1-\beta} E\left[g(\epsilon_{t-i-1})\right]$$

$$= \frac{\alpha_0}{1-\beta}$$

• Variance:

$$\operatorname{Var}(\ln(\sigma_t^2)) = \operatorname{Var}\left(\frac{1}{1-\beta B}[\alpha_0 + \alpha g(\epsilon_{t-1})]\right) = \operatorname{Var}\left(\frac{\alpha}{1-\beta B}g(\epsilon_{t-1})\right) = \operatorname{Var}\left(\frac{\alpha}{1-\beta}g(\epsilon_{t-i-1})\right)$$

Since

$$\begin{aligned} \text{Var}[g(\epsilon_{t-1})] &= E[g^2(\epsilon_{t-1})] \\ &= E\left[(\theta \epsilon_t + |\epsilon_t| - E\left(|\epsilon_t|\right))^2 \right] \\ &= E(\theta^2 \epsilon_t^2) + E(|\epsilon_t|^2) + \frac{2}{\pi} + 2E\left(\theta \epsilon_t |\epsilon_t|\right) - 2E\left(|\epsilon_t|\sqrt{\frac{2}{\pi}}\right) - 2E\left(\theta \epsilon_t \sqrt{\frac{2}{\pi}}\right) \\ &= (\theta^2 + 1) + \frac{2}{\pi} - \frac{4}{\pi} = (\theta^2 + 1) - \frac{2}{\pi} \end{aligned}$$

where

$$E\left(\epsilon_t|\epsilon_t|\right) = 0$$

Then

$$Var(\ln(\sigma_t^2)) = \frac{\alpha^2((\theta^2 + 1) - \frac{2}{\pi})}{(1 - \beta)^2}$$

11. Consider the following bivariate VAR model:

$$y_{1t} = 0.3y_{1,t-1} + 0.8y_{2,t-1} + a_{1t}$$

$$y_{2t} = 0.9y_{1,t-1} + 0.4y_{2,t-1} + a_{2t}$$

with $E\left(a_{1t}a_{1\tau}\right)=1$ if $t=\tau$ and 0 otherwise, $E\left(a_{2t}a_{2\tau}\right)=2$ if $t=\tau$ and 0 otherwise, and $E\left(a_{1t}a_{2\tau}\right)=0$ for all t and τ .

- (a) Is this system stationary?
- (b) Calculate the two-step ahead forecast variance for variable $y_{1,t+2}$, that is

$$E[y_{1,t+2} - E(y_{1,t+2}|Y_t, Y_{t-1}, \cdots)]^2$$

where $Y_t = (y_{1t}, y_{2t})'$

Solution:

• Stationarity or invertibility condition:

$$\lambda \left| \begin{array}{cc} 1 & 0 \\ 0 & 1 \end{array} \right| - \Phi_1$$

- VAR(p) model:
 - $\hat{y_{1,t}}(2) = E(y_{1,t+2}|Y_t, Y_{t-1}, \cdots)$
 - Var $(e_t(1)) = \Sigma$
 - Var $(e_t(l)) = \Sigma + \sum_{j=1}^{l-1} \Psi_j \Sigma \Psi_j^T$
 - $\Sigma = E\left[a_t a'_{t+k}\right]$, if k = 0.
- 12. Write down the bivariate system into an VAR model and show that it is not stationary:

$$y_{1t} = \gamma y_{2t} + \varepsilon_{1t}$$

$$y_{2t} = y_{2,t-1} + \varepsilon_{2t}$$

where $\gamma \neq 0, \varepsilon_{1t}$ and ε_{2t} being uncorrelated white noise processes.

Solution:

Stationarity or invertibility condition:

$$\lambda \left| \begin{array}{cc} 1 & 0 \\ 0 & 1 \end{array} \right| - \Phi_1$$

13. Show that the following VAR model

$$\mathbf{y}_t = \sum_{i=1}^p \Phi_i \mathbf{y}_{t-i} + \varepsilon_t$$

can be written as following VCE model:

$$\Phi(B)\mathbf{y_t} = \Phi^*(\mathbf{B})(1 - \mathbf{B})\mathbf{y_t} + \Phi(1)\mathbf{B}\mathbf{y_t}$$

where
$$\Phi^*(B) = I_m - \sum_{i=1}^{p-1} \Phi_i^* B^i$$
 with $\Phi_i^* = -\sum_{j=i+1}^p \Phi_j$.

Solution:

Assume $|\Phi(z)| = \left|I_m - \sum_{i=1}^p \Phi_i z^i\right| = 0$ has d < m unit root and the remaining roots outside the unit circle.

The rank of $\Phi(1) = I_m - \sum_{i=1}^p \Phi_i$ is r and r = m - d.

We can decompose $\Phi(1)$ as

$$\alpha \beta' = -\Phi(1) = -\mathbf{I}_m + \Phi_1 + \ldots + \Phi_p$$

 $\Phi(B)$ can be re-expressed as

$$\Phi(B) = \Phi^*(B)(1 - B) + \Phi(1)B$$

where
$$\Phi^*(B) = \mathbf{I}_m - \sum_{i=1}^{p-1} \Phi_i^* B^i$$
 with $\Phi_i^* = -\sum_{j=i+1}^p \Phi_j$.

14. Consider the two dimensional vector AR(2) model:

$$\begin{bmatrix} Z_{1t} \\ Z_{2t} \end{bmatrix} = \begin{bmatrix} a_{1t} \\ a_{2t} \end{bmatrix} + \begin{bmatrix} -0.2 & 0.1 \\ 0.5 & 0.2 \end{bmatrix} \begin{bmatrix} Z_{1,t-1} \\ Z_{2,t-1} \end{bmatrix} + \begin{bmatrix} 0.8 & 0.7 \\ -0.4 & 0.6 \end{bmatrix} \begin{bmatrix} Z_{1,t-2} \\ Z_{2,t-2} \end{bmatrix}$$

where $\{(a_{1t}, a_{2t})'\}$ is a sequence of i.i.d. standard normal random vectors. Show that it is a partially non-stationary AR model and give its cointegration vector.

Solution:

$$\begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} - \begin{bmatrix} -0.2 & 0.1 \\ 0.5 & 0.2 \end{bmatrix} - \begin{bmatrix} 0.8 & 0.7 \\ -0.4 & 0.6 \end{bmatrix} = \begin{bmatrix} -0.4 & 0.8 \\ 0.1 & -0.2 \end{bmatrix} = \begin{bmatrix} -0.4 \\ 0.1 \end{bmatrix} \begin{bmatrix} 1 & -2 \end{bmatrix} = 0$$

Since there exists unit root, it's non-stationary.

And the cointegrating vector is $\beta = \begin{pmatrix} 1 \\ -2 \end{pmatrix}$.

15. Determine the stationarity and invertibility of the following two dimensional vector models and find their correlation matrix function, ρ_k , for $k = \pm 1, \pm 2, \pm 3$:

(a)
$$(I - \Phi_1 B)$$
 $Z_t = a_t$, where $\Phi_1 = \begin{bmatrix} 0.8 & 0.3 \\ 0.1 & 0.6 \end{bmatrix}$ and $\Sigma = I$;
(b) $(I - \Phi_1 B)$ $Z_t = a_t$, where $\Phi_1 = \begin{bmatrix} 0.4 & 0.2 \\ -0.2 & 0.8 \end{bmatrix}$ and $\Sigma = I$;
(c) $Z_t = (I - \Theta_1 B)$ a_t , where $\Theta_1 = \begin{bmatrix} 0.6 & 1.2 \\ 0.4 & 0.8 \end{bmatrix}$ and $\Sigma = I$;

Solution:

• Stationarity or invertibility condition:

$$\lambda \left| \begin{array}{cc} 1 & 0 \\ 0 & 1 \end{array} \right| - \Phi_1$$

• VAR(1) model:

$$\Sigma = \Gamma(0) - \Phi_1 \Gamma(0) \Phi_1'$$

$$\Gamma(k) = \begin{cases} \Gamma(-1) \Phi_1' + \Sigma, & \text{if } k = 0 \\ \Gamma(k - 1) \Phi_1' = \Gamma(0) \left(\Phi_1'\right)^k, & \text{if } k \ge 1 \end{cases}$$

$$\Gamma(k) = \Gamma'(-k) \ge 0$$

$$\rho(k) = D^{-1} \Gamma(k) D^{-1}$$
 where
$$D = \begin{bmatrix} \sigma_{11}(k) & 0 & \cdots & 0 \\ 0 & \sigma_{22}(k) & \cdots & 0 \\ \vdots & \vdots & \vdots & \vdots \\ 0 & 0 & \cdots & \sigma_{mn}(k) \end{bmatrix}.$$

• solution

$$\begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} = \Gamma(0) - \begin{bmatrix} 0.8 & 0.3 \\ 0.1 & 0.6 \end{bmatrix} \Gamma(0) \begin{bmatrix} 0.8 & 0.1 \\ 0.3 & 0.6 \end{bmatrix}$$

$$\rho(k) = \begin{bmatrix} \sigma_1^{-1} & 0 \\ 0 & \sigma_2^{-1} \end{bmatrix} \Gamma(k) \begin{bmatrix} \sigma^{-1} & 0 \\ 0 & \sigma_2^{-1} \end{bmatrix} = \begin{bmatrix} \sigma_1^{-1} & 0 \\ 0 & \sigma_2^{-1} \end{bmatrix} \Gamma(0) \begin{bmatrix} 0.8^k 0.1^k \\ 0.3^k 0.6^k \end{bmatrix} \begin{bmatrix} \sigma & -1 \\ 0 & \sigma_2^{-1} \end{bmatrix}$$

16. Consider the process

$$Z_{1t} = Z_{1,t-1} + a_{1t} + \theta a_{1,t-1}$$

$$Z_{2t} = \phi Z_{1t} + a_{2t}$$

where $|\phi| < 1$, $|\theta| < 1$ and $a_t = [a_{1t}, a_{2t}]' \sim N(0, \Sigma)$. (a) Write the process in a vector form (b) Is the process $[Z_{1t}, Z_{2t}]'$ stationary and invertible? (c) Write down the model for the vector of the first differences $(I - B)Z_t$, where $Z_t = [Z_{1t}, Z_{2t}]'$. Is the resulting model stationary and invertible?

Solution:

$$\begin{bmatrix} z_{1t} \\ z_{2t} \end{bmatrix} = \begin{bmatrix} 1 & 0 \\ \phi & 0 \end{bmatrix} \begin{bmatrix} z_{1,t-1} \\ z_{2,t-2} \end{bmatrix} + \begin{bmatrix} 1 & 0 \\ \phi & 1 \end{bmatrix} \cdot \begin{bmatrix} a_{1t} \\ a_{2t} \end{bmatrix} + \begin{bmatrix} \theta & 0 \\ \phi & 0 \end{bmatrix} \begin{bmatrix} a_{1,t-1} \\ a_{2,t-1} \end{bmatrix}$$

• (c)
$$\Delta z_{1t} = z_{1t} - z_{1,t-1} = a_1 + \theta a_{1+1}$$

$$\Delta z_{2t} = \phi \Delta z_{1t} + a_{2t} - a_{2,t-1} = \phi a_{1t} + \phi \theta a_{1,t-1} + a_{2t} - a_{2,t-1}$$

$$\begin{bmatrix} \Delta z_{1t} \\ \Delta z_{2t} \end{bmatrix} = \begin{bmatrix} 1 & 0 \\ \phi & 1 \end{bmatrix} \cdot \begin{bmatrix} a_{1t} \\ a_{2t} \end{bmatrix} + \begin{bmatrix} \theta & 0 \\ \phi \theta & -1 \end{bmatrix} \cdot \begin{bmatrix} a_{1,t-1} \\ a_{2,t-1} \end{bmatrix}$$

$$\begin{vmatrix} \lambda - \theta & 0 \\ \phi - \phi \theta & \lambda + 1 \end{vmatrix} = (\lambda - \theta)(\lambda + 1) = 0$$

Since $|\lambda| \ge 1$, so it's not invertible; And it's vector MA model, so it's stationary.

17. Show that the process $y_t = z_t - z_{t-1}$ is weakly stationary, where $z_t = 0.9z_{t-1} + a_t$ and $\{a_t\}$ is white noise series.

Solution:

$$y_t = 0.9z_{t-1} + a_t - 0.9z_{t-2} - a_{t-1}$$
$$= 0.9y_{t-1} + a_t - a_{t-1}$$
$$\Rightarrow (1 - 0.9B)y_t = (1 - B)a_t$$

Since $|\phi_1| = |0.9| < 1$, then $y_t = z_t - z_{t-1}$ is stationary.