

**Statistics 641, Fall 2013**  
**Homework #5**  
**Solutions**

1. Suppose that we have a phase II, single arm trial using a two stage design. The hypotheses of interest are  $H_0: \pi \leq 0.15$  versus  $H_1: \pi \geq 0.4$  where  $\pi$  is the true success rate. Let  $y_k$  be the total number of successes through stage  $k$ ,  $k = 1, 2$ . Note: you can use the functions `dbinom` and `pbinom` in R to calculate binomial probabilities.

- (a) We enroll 16 subjects in stage 1 and stop and accept  $H_0$  if we observe  $a_1 = 3$  or fewer responses ( $y_1 \leq 3$ ), otherwise we continue to stage 2. Find the stopping probabilities under both  $H_0$  and  $H_1$ .

---

Under  $H_0$ , the stopping probability is  $\Pr\{y_1 \leq 3\} = .7899$ , and under  $H_1$  it is .0651.

In R:

```
> pbinom(3, 16, .15)
[1] 0.7898907
> pbinom(3, 16, .40)
[1] 0.06514674
```

---

- (b) At stage 2 we enroll an additional 16 subjects and reject  $H_0$  if  $y_2 > 8$ . Compute the overall probabilities of rejection under both  $H_0$  and  $H_1$  for the two-stage trial.

---

We accept  $H_0$  if  $y_1 \leq 3$  and  $y_2 \leq 8$ . Under  $H_0$  this probability is

$$\Pr\{y_1 \leq 3\} + \sum_{i=4}^8 \Pr\{y_1 = i\} \Pr\{x_2 \leq 8 - i\} = .9659$$

In R:

```
> pbinom(3, 16, .15) + sum(dbinom(4:8,16,.15)*pbinom(4:0, 16, .15))
[1] 0.9658661
```

Under  $H_1$ , this probability is 0.0969.

```
> pbinom(3, 16, .40) + sum(dbinom(4:8,16,.40)*pbinom(4:0, 16, .40))
[1] 0.09691022
```

Therefore, the rejection probabilities are  $1-0.9659=0.0341$  and  $1-0.0969=0.9031$  under  $H_0$  and  $H_1$  respectively.

Alternatively,

```
> 1-pbinom(8,16,.15) - sum(dbinom(4:8,16,.15)*(1-pbinom(4:0, 16, .15)))
[1] 0.03413386
> 1-pbinom(8,16,.40) - sum(dbinom(4:8,16,.40)*(1-pbinom(4:0, 16, .40)))
[1] 0.9030898
```

### or, equivalently, (note that `pbinom(x,...)` is zero if  $x < 0$ )

```
> sum(dbinom(4:16,16,.15)*pbinom(8-4:16, 16, .15,lower=F))
```

```
[1] 0.03413386
> sum(dbinom(4:16,16,.40)*pbinom(8-4:16, 16, .40,lower=F))
[1] 0.9030898
```

- (c) Compute the expected sample sizes for  $\pi = 0.15$  and  $\pi = 0.4$ .

$N$  is either 16 or 32, depending on whether we stop at stage 1. Under  $H_0$ ,  
 $E[N] = 16 \times 0.7899 + 32 \times (1 - 0.7899) = 19.36$  and under  $H_1$ ,  
 $E[N] = 16 \times 0.0651 + 32 \times (1 - 0.0651) = 30.96$ .

- (d) Suppose, instead, we perform a single stage trial with  $N = 32$  subjects and we reject  $H_0$  if we observe more than 8 successes. Find the type I and type II error rates. What is the advantage of the two-stage trial?

Under  $H_0$ , probability of rejection (type I error) is  $\Pr\{y > 8\} = 1 - 0.9587 = .0413$ , and under  $H_1$  the acceptance probability (type II error) is  $\Pr\{y \leq 8\} = 0.0575$ . The type I error rate is slightly smaller for the two-stage trial, but the type II error rate is larger. The advantage of the 2 stage trial is that we can have the potential to stop earlier and discard ineffective treatments sooner.

2. The dataset `data4.csv` contains data collected from a crossover study with 40 subjects per sequence. The variables in the dataset are:

<code>seq</code>	Assigned treatment sequence
<code>y</code>	Response
<code>id</code>	Subject id
<code>period</code>	Period
<code>z</code>	Treatment ("A" or "B")

- (a) Calculate the means within each treatment group separately for periods 1 and 2. Using these means, calculate the estimate of the treatment difference assuming no carryover.

```
> m <- with(data, tapply(y, list(z, period), mean))
> m
      1      2
A  9.6625 22.7025
B 10.1000 22.9200
# difference in "AB" group:
> m[2,2]-m[1,1]
[1] 13.2575
# difference in "BA" group:
> m[2,1]-m[1,2]
[1] -12.6025
# mean difference:
> (m[2,2]-m[1,1] + m[2,1]-m[1,2])/2
[1] 0.3275
```

Note that the estimate of period effect is:

```
> (m[2,2]-m[1,1] - (m[2,1]-m[1,2]))/2
[1] 12.93
```

---

- (b) Fit a regression model that estimates the treatment difference and its standard error.
- 

```
> summary(lm(y ~ z + period + id, data=data))
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-3.409e+00	6.695e-01	-5.091	2.41e-06 ***
zB	3.275e-01	1.412e-01	2.320	0.022947 *
period	1.293e+01	1.412e-01	91.602	< 2e-16 ***
ids02	1.750e+00	8.927e-01	1.960	0.053535 .
...				
ids80	-1.400e+00	8.927e-01	-1.568	0.120878

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

...

The estimate of treatment effect is 0.3275 with SE 0.1412. The *p*-value is 0.0229.

Alternative using mixed-effects model from package lme4:

```
> lmer(y ~ z + period + (1|id), data=data)
```

Linear mixed model fit by REML

...

Random effects:

Groups	Name	Variance	Std.Dev.
id	(Intercept)	9.2589}	3.04284
Residual		0.79697	0.89273

Number of obs: 160, groups: id, 80

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	-3.2125	0.4129	-7.78
zB	0.3275	0.1412	2.32
period	12.9300	0.1412	91.60

Correlation of Fixed Effects:

	(Intr) zB
zB	-0.171
period	-0.513 0.000

---

- (c) Fit a regression model to estimate the effect of treatment using only period 1. (This is equivalent to a parallel group trial in which subjects are assigned only one of “A” or “B.”)

```
> summary(lm(y ~ z, data=data, subset=period==1))
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	9.6625	0.4734	20.412	<2e-16 ***
zB	0.4375	0.6695	0.654	0.515

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

- (d) Comment on the differences between the analyses in parts (b) and (c).

The standard errors of the two estimates are quite different

- Cross-over analysis: SE = 0.1412
- Parallel group (period 1 only): SE = 0.6695

This suggests that there is high correlation between the period 1 and period 2 observations from each subject. Because the cross-over model is based on within-subject differences, the subject-level effects are accounted for and the variability is significantly reduced providing greater power.

Even though the point estimate of the difference is larger in period 1 analysis, the increased variance results in a statistically in-significant difference.