

# Week 2

## 1 Independent random variables

Definition: Two random variables  $X_1, X_2$  are said to be independent if,

$$P(X_1 = x_1, X_2 = x_2) = P(X_1 = x_1)P(X_2 = x_2) \quad \forall x_1 \in \mathcal{X}_1 \text{ \& } x_2 \in \mathcal{X}_2$$

$P(X_1 = x_1, X_2 = x_2)$  means that  $X_1 = x_1$  **AND**  $X_2 = x_2$  occur.

### 1.1 Claim

$$\sum_{x_2 \in \mathcal{X}_2} P(X_1 = x_1, X_2 = x_2) = P(X_1 = x_1)$$

**Proof:** Suppose  $A$  &  $B$  are disjoint/mutually exclusive. Then,

$$P(A \cap B) = 0 \quad \& \quad P(A \cup B) = P(A) + P(B)$$

Now the events  $(X_1 = x_1) \cap (X_2 = x_2)$  are disjoint for different values of  $x_2 \in \mathcal{X}_2$ . (if  $x_2 \neq x'_2 \forall x_2 \in \mathcal{X}_2$ )

Thus,

$$\begin{aligned} \sum_{x_2 \in \mathcal{X}_2} P(X_1 = x_1, X_2 = x_2) &= \sum_{x_2 \in \mathcal{X}_2} P((X_1 = x_1) \cap (X_2 = x_2)) \\ &= P\left(\bigcup_{x_2 \in \mathcal{X}_2} (X_1 = x_1) \cap (X_2 = x_2)\right) \end{aligned}$$

Now,

$$\begin{aligned} \bigcup_{x_2 \in \mathcal{X}_2} [(X_1 = x_1) \cap (X_2 = x_2)] &= (X_1 = x_1) \cap \left[ \bigcup_{x_2 \in \mathcal{X}_2} (X_2 = x_2) \right] \\ &= (X_1 = x_1) \cap [x_2 \in \mathcal{X}_2] \end{aligned}$$

As  $[x_2 \in \mathcal{X}_2]$  forms the entire sample space,

$$\begin{aligned} P\left(\bigcup_{x_2 \in \mathcal{X}_2} (X_1 = x_1) \cap (X_2 = x_2)\right) &= P\left((X_1 = x_1) \cap [x_2 \in \mathcal{X}_2]\right) \\ &= P\left((X_1 = x_1) \cap \Omega\right) \\ &= P(X_1 = x_1) \end{aligned}$$

## 2 Entropy

Suppose  $X_1 \in \mathcal{X}_1$  &  $X_2 \in \mathcal{X}_2$  are **independent** random variables. Then,

$$H(X_1, X_2) = H(X_1) + H(X_2)$$

Proof:

$$\begin{aligned} & \sum_{\substack{x_1 \in \mathcal{X}_1 \\ x_2 \in \mathcal{X}_2}} P(X_1 = x_1, X_2 = x_2) \log \left( \frac{1}{P(X_1 = x_1, X_2 = x_2)} \right) \\ = & \sum_{x_1 \in \mathcal{X}_1} \sum_{x_2 \in \mathcal{X}_2} P(X_1 = x_1, X_2 = x_2) \left[ \log \frac{1}{P(X_1 = x_1)} + \log \frac{1}{P(X_2 = x_2)} \right] \\ = & \sum_{x_1 \in \mathcal{X}_1} \log \frac{1}{P(X_1 = x_1)} \left( \sum_{x_2 \in \mathcal{X}_2} P(X_1 = x_1, X_2 = x_2) \right) \\ & + \sum_{x_2 \in \mathcal{X}_2} \log \frac{1}{P(X_2 = x_2)} \left( \sum_{x_1 \in \mathcal{X}_1} P(X_1 = x_1, X_2 = x_2) \right) \end{aligned}$$

From the previous claim,

$$\begin{aligned} H(X_1, X_2) &= \sum_{x_1 \in \mathcal{X}_1} P(X_1 = x_1) \log \frac{1}{P(X_1 = x_1)} \\ &+ \sum_{x_2 \in \mathcal{X}_2} P(X_2 = x_2) \log \frac{1}{P(X_2 = x_2)} \\ &= H(X_1) + H(X_2) \end{aligned}$$

### 2.1 Conditional probability distribution

What if  $X_1$  &  $X_2$  are not independent? Then we would use conditional probability distribution. i.e.

$$P(X_2 = x_2 / X_1 = x_1) := \frac{P(X_2 = x_2, X_1 = x_1)}{P(X_1 = x_1)}, \quad P(X_1 = x_1) \neq 0$$

This definition for conditional probability satisfies the probability axioms and hence it is a valid probability measure.

### 2.2 Conditional entropy

Definition:

$$H(X_2 / X_1) := \sum_{x_1 \in \mathcal{X}_1} P(X_1 = x_1) H(X_2 / X_1 = x_1)$$

where,

$$H(X_2 / X_1 = x_1) := \sum_{x_2 \in \mathcal{X}_2} P(X_2 = x_2 / X_1 = x_1) \log \frac{1}{P(X_2 = x_2 / X_1 = x_1)}$$

### 2.2.1 Support of a function

When  $P(X = x) = 0$ , it is not considered in calculating the entropy.

$$H(X) = - \sum_{\{x \in \mathcal{X} : P(X=x) \neq 0\}} P(X = x) \log(P(X = x))$$

Suppose that  $f: X \rightarrow \mathbb{R}$  is a real-valued function whose domain is an arbitrary set  $X$ . The set-theoretic support of  $f$ , written  $\text{supp}(f)$ , is the set of points in  $X$  where  $f$  is non-zero:

$$\text{supp}(X) = \{x \in X : f(x) \neq 0\}$$

Note:  $P(X = x)$  can be denoted as  $P_X(x)$  or  $P(x)$ .

### 2.3 Important Results

$$H(X) = - \sum_{x \in \text{supp}(P)} P(x) \log(P(x))$$

$$H(X/Y) = \sum_{y \in \text{supp}(P_Y)} P_Y(y) H(X/Y = y)$$

$$H(X/Y = y) = - \sum_{x \in \text{supp}(P_{X/Y})} P_{X/Y}(x/y) \log(P_{X/Y}(x/y))$$

### 2.4 Chain Rule

$$H(X, Y) = H(X) + H(Y/X) = H(Y) + H(X/Y)$$

Proof:

$$\begin{aligned} H(X) + H(Y/X) &= - \sum_{x \in \text{supp}(P)} P(x) \log(P(x)) \\ &\quad + \sum_{x \in \text{supp}(P_X)} P(x) \left( \sum_{y \in \text{supp}(P_Y)} P(y/x) \log \left( \frac{1}{P(y/x)} \right) \right) \\ &= - \sum_{x \in \text{supp}(P_X)} \left( \sum_{y \in \text{supp}(P_Y)} P(x, y) \right) \log(P(x)) \\ &\quad + \sum_{x \in \text{supp}(P_X)} \left( \sum_{y \in \text{supp}(P_Y)} P(x, y) \right) \log \left( \frac{1}{P(y/x)} \right) \\ &= \sum_{x \in \text{supp}(P_X)} \sum_{y \in \text{supp}(P_Y)} P(x, y) \left( \log \frac{1}{P(x)P(y/x)} \right) \\ &= - \sum_x \sum_y P(x, y) \log P(x, y) = H(x, y) \end{aligned}$$

## 2.5 Upper and lower bound of entropy

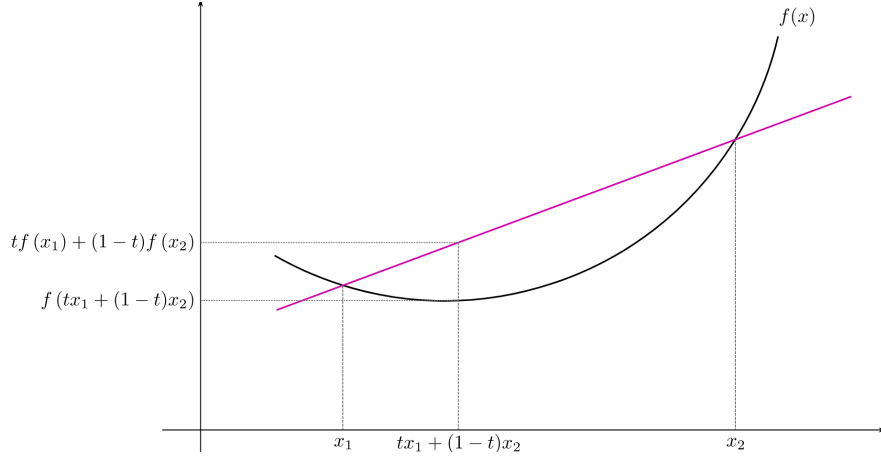
$$0 \leq H(X) \leq |\log \mathcal{X}|$$

Proof:

$$H(X) = \sum_x P(x) \log \frac{1}{P(x)} \rightarrow \text{always positive as } P(x) > 0 \forall x \in \text{supp}(P_X).$$

$$H(X) = 0 \text{ when } |\text{supp}(P_X)| = 1.$$

### 2.5.1 Jensen's inequality (Concave/convex functions)



$x_1 < x_3 < x_2$  such that  $x_3 = tx_1 + (1-t)x_2$

If  $f(x_3) \geq tf(x_1) + (1-t)f(x_2) \rightarrow$  Concave function.

If  $f(tx_1 + (1-t)x_2) \leq tf(x_1) + (1-t)f(x_2) \rightarrow$  Convex function.

Hence, the above graph is a convex function.

We shall use the concavity property of the log function to obtain the proof for  $H(X) \leq |\log \mathcal{X}|$ .

Proof:

$$H(X) = \sum_{x \in \text{supp}(P)} P(x) \log \frac{1}{P(x)}$$

Let us assume  $\lambda_x$  is another notation for  $P(x)$ .

$$H(X) = \sum_{x \in \text{supp}(P)} \lambda_x \log \frac{1}{P(x)}$$

This is a convex combination of  $\log \frac{1}{P(x)} : x \in \text{supp}(P_X)$  as  $\sum_x P(x) = 1$ .

$$\begin{aligned} H(X) &\leq \log \left( \sum_{x \in \text{supp}(P_X)} \lambda_x \frac{1}{P(x)} \right) \leq \log |\text{supp}(P_x)| \\ &\Rightarrow H(X) \leq |\mathcal{X}| \end{aligned}$$

### 2.5.2 When is $H(X)$ is $\log|\mathcal{X}|$ ?

When Jensen's inequality is satisfied with equality, i.e. if  $f(x)$  is a straight line, but the  $\log x$  curve is strictly concave.

Suppose  $f(x)$  is strictly concave(or convex) &  $\lambda_1, \lambda_2 \neq 0$  &  $\lambda_1 + \lambda_2 = 1$ .

If  $f(\lambda_1 x_1 + \lambda_2 x_2) = \lambda_1 f(x_1) + \lambda_2 f(x_2)$ , then  $x_1 = x_2$ .

More generally, if  $\lambda_i \neq 0$ ,  $\sum_{i=1}^n \lambda_i = 1$  and Jensen's inequality holds with equality, then  $x_1 = x_2 = \dots = x_n$ .

Applying this to  $H(X) \leq \log|\mathcal{X}|$ , from the previous proof,

$$H(X) = \sum_{x \in \text{supp}(P)} \lambda_x \log \frac{1}{P(x)} \leq \log |\text{supp}(P_x)|$$

Suppose equality holds, then by the above claim we must have:

$$\begin{aligned} \frac{1}{P(x)} &= \text{const} \quad \forall x \in \text{supp}(P_x) \\ \Rightarrow \text{const. } c = P(x) &= \frac{1}{|\text{supp}(P_x)|} \end{aligned}$$

If  $|\text{supp}(P_x)| = |\mathcal{X}|$ ,  $H(X) = \log |\text{supp}(P_x)|$ . Then  $P(x) = \frac{1}{|\mathcal{X}|} \quad \forall x \in \mathcal{X}$ .

We have just proved that  $H(X) = \log_2 |\mathcal{X}|$ , this can be true only when  $P_x$  is uniform. Thus:

Lemma:  $H(X) = \log_2 |\mathcal{X}|$  iff  $P_x$  is uniform.

## 3 Relative Entropy/ Information Divergence/ Kullback-Leibler Divergence

Suppose there is a random variable  $X$  for which we have two probability distributions  $p_x$  &  $q_x$ .

Then the RE or ID or KL is defined as:

$$D(p_X || q_Y) := \sum_{x \in \text{supp}(p_x)} p(x) \log \frac{p(x)}{q(x)}$$

( $D(p_X || q_Y)$  is a 'kind of' a distance measure between distributions  $p$  &  $q$ .)

### 3.1 Lower limit of relative entropy

$$D(p || q) \geq 0$$

Proof:

$$\begin{aligned} D(p || q) &= - \sum_{x \in \text{supp}(p_x)} p(x) \log \frac{q(x)}{p(x)} \\ &\geq - \log \left( \sum_{x \in \text{supp}(p_x)} p(x) \frac{q(x)}{p(x)} \right) \end{aligned}$$

$$\begin{aligned}
&\geq -\log \left( \sum_{x \in \text{supp}(p_x)} q(x) \right) \\
&\geq 0 \quad \text{as} \quad \sum q(x) \leq 1
\end{aligned}$$

### 3.2 When is RE equal to zero?

By applying Jensen's inequality for equality condition.

$$\begin{aligned}
\frac{p(x)}{q(x)} &= \text{const. } c \quad \forall x \in \text{supp}(p_x) \\
&\Rightarrow p(x) = cq(x) \quad \& \\
\sum_{x \in \text{supp}(p_x)} p(x) &= \sum_{x \in \text{supp}(p_x)} cq(x) \quad \forall x \in \text{supp}(p_x)
\end{aligned}$$

These together mean that  $c = 1$ .

$$\begin{aligned}
&\Rightarrow p(x) = q(x) \quad \forall x \in \text{supp}(p_x) \\
&\Rightarrow D(p||q) = 0 \quad \text{iff} \quad p_x = q_x
\end{aligned}$$

### 3.3 Upper and lower limits of conditional entropy

$$\begin{aligned}
H(X/Y) &= \sum_{y \in \text{supp}(P_Y)} P_Y(y) \sum_{x \in \text{supp}(P_{X/Y})} P_{X/Y}(x/y) \log \frac{1}{P_{X/Y}(x/y)} \\
0 &\leq H(X/Y) \leq H(X)
\end{aligned}$$

Proof:

$H(X/Y) \geq 0$  is true because  $H(X/Y = y) \geq 0$  &  $P(y) \geq 0$ .

$$\begin{aligned}
H(X) - H(X/Y) &= \sum_{x \in \text{supp}(P_x)} P(x) \log \frac{1}{P(x)} - \sum_{x,y} P(x,y) \log \frac{1}{P(x,y)} \\
&= \sum_{x,y} P(x,y) \log \frac{1}{P(x)} - \sum_{x,y} P(x,y) \log \frac{1}{P(x,y)} \\
&= \sum_{x,y} P(x,y) \log \left( \frac{P(x,y)}{P(x)} \right) = \sum_{x,y} P(x,y) \log \left( \frac{P(x,y)}{P(x)P(y)} \right)
\end{aligned}$$

As both  $P(x,y)$  &  $P(x)P(y)$  are valid joint distribution of  $x,y$ . Hence,

$$\begin{aligned}
H(X) - H(X/Y) &= D(P(x,y)||P(x)P(y)) \\
&\geq 0 \quad (\text{as } D(p||q) \geq 0) \\
&\Rightarrow H(X/Y) \leq H(X)
\end{aligned}$$