202C-hw2

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Question 1

Solution:

a)
$$x = \sqrt{z}$$
, $\frac{dx}{dx} = \pm \frac{1}{2\sqrt{z}}$

b)
$$x = log(z)$$
, $\frac{dx}{dz} = \frac{1}{z}$

$$c) \quad x = \frac{1}{z}, \quad \frac{dx}{dz} = -\frac{1}{z^2}$$

d)
$$x = \frac{e^z}{1 + e^z}$$
, $\frac{dx}{dz} = \frac{e^z}{(1 + e^z)^2}$

Question 2

Solution:

$$x \sim Gamma(a, b)$$

$$f(x) = \frac{\beta^{\alpha}}{\Gamma(\alpha)} x^{\alpha - 1} e^{-x\beta}, \quad x = \frac{1}{z}$$

$$Jacobian: \quad \left| \frac{dx}{dz} \right| = \left| -\frac{1}{z^2} \right| = z^{-2}$$

$$f(z) = \frac{b^a}{\Gamma(a)} z^{-a-1} e^{-\frac{b}{z}} \sim Inver.gam(a, b), \quad z > 0$$

Question 3

Solution:

a)

$$\begin{split} FIP &= \frac{-\frac{d^2}{d\lambda^2}logp(\lambda)}{-\frac{d^2}{d\lambda^2}(logp(y|\lambda) + logp(\lambda))} \\ &= \left[\frac{a-1}{a-1 + \sum yi}\right] \end{split}$$

b)

by weight average

$$E(y|\lambda) = \frac{a + \sum yi}{b+n}$$
$$= \frac{b}{b+n} \frac{a}{b} + \frac{n}{b+n} \overline{y}$$

c)

From hw 1 we got:

$$\begin{split} f(y|a,b) &= \frac{b^a}{\Gamma(a)y!} \frac{\Gamma(y+a)}{(b+1)^{y+a}} \\ &= \frac{(y+a-1)!}{(a-1)!y!} \frac{b^a}{(b+1)^{y+a}} \\ &= \binom{y+a-1}{y} b^a (b+1)^{-y-a} \\ &= \binom{y+a-1}{y} \left(\frac{b}{1+b}\right)^a \left(\frac{1}{1+b}\right)^y \end{split}$$

Which
$$f(y|a,b) \sim NB(a, \frac{b}{1+b})$$

 $prior \ data \ mean = \frac{a}{b}$ $prior \ sample \ size = b \ (from \ weight \ average)$

Question 4

Solution:

$$y = e^{-zi}I_{0 < zi < 1}, \quad \left| \frac{dy}{dz} \right| = \left| -e^{-zi} \right| = e^{-zi}I_{0 < zi < 1}$$

$$f(zi|\theta) \propto \theta exp(-zi(\theta - 1))e^{-zi}I_{0 < zi < 1}$$

$$= \theta exp(-zi\theta)I_{0 < zi < 1}$$

$$= \frac{\theta}{\Gamma(1)}zi^{1-1}e^{-zi(\theta)}I_{0 < zi < 1}$$

$$= \theta e^{-zi(\theta)} \sim Gamma(1, \theta), 0 < zi < 1$$

Also because alpha = 1,

$$\theta e^{-zi(\theta)} \sim exp(\theta), 0 < zi < 1$$

Question 5

Solution:

a)

$$\begin{split} Lf(y|\theta) &= \theta^n e^{\theta - 1log(\sum yi)} I_{0 < yi < 1} \\ & f(\theta|z) \propto \theta^{a-1} e^{-\theta b} \\ f(\theta|y) \propto \theta^{a+n-1} e^{-\theta(b - \sum log(yi)} \sim gamma(a+n,b - \sum log(yi)) \end{split}$$

From posterior distribution:

$$mean = \frac{a+n}{b-\sum log(yi)}$$

$$variance = \frac{a+n}{(b-\sum log(yi))^2}$$

$$mode = \frac{a+n-1}{b-\sum log(yi)}$$

b)

$$\begin{split} FIP &= \frac{-\frac{d^2}{d\theta^2}logp(\theta)}{-\frac{d^2}{d\theta^2}(logp(y|\theta) + logp(\theta))} \\ evaluated \ at \ mode &= \frac{a-1}{a+n-1} \end{split}$$

c)

$$FIP = \left[\frac{-\frac{d^2}{d\theta^2}logp(\theta)}{-\frac{d^2}{d\theta^2}(logp(y|\theta) + logp(\theta))}\right], is \ better$$

Easier to compute by using this formula

 \mathbf{d}

From wikipedia, Pareto distribution and Zeta distribution.

Question 6

a)

from hw 1 Q2

$$p(\lambda|y) \propto \lambda^{\sum (yi)+a-1} e^{-\lambda(b+n)}, \lambda > 0$$

 $posterior\ follows\ gamma\ distribution\ \sim Gamma(\sum (yi) + a, b + n)$

first normal approximation, by using mean and variance from posterior distribution.

$$mean = \frac{\sum (yi) + a}{b+n}, \quad variance = \frac{\sum (yi) + a}{(b+n)^2}$$

we have
$$f(y|\lambda) \sim N(\frac{\sum (yi) + a}{b+n}, \frac{\sum (yi) + a}{(b+n)^2})$$

second normal approximation, by using fisher information and mode of lambda.

$$\begin{split} mode\ of\ \lambda &= \frac{\sum (yi) + a - 1}{b + n} \\ I &= \frac{(b+n)^2}{\sum (yi) + a - 1} \\ we\ have\ &= f(y|\lambda) \sim N(\frac{\sum (yi) + a - 1}{b+n}, \frac{\sum (yi) + a - 1}{(b+n)^2}) \end{split}$$

b)

i: when the parameter from posterior distribution is large then we can say the normal approximation(s) to the gamma posterior are good.

ii: when the parameter from posterior distribution is close to 0 then we can say the normal approximation(s) to the gamma posterior are not good.

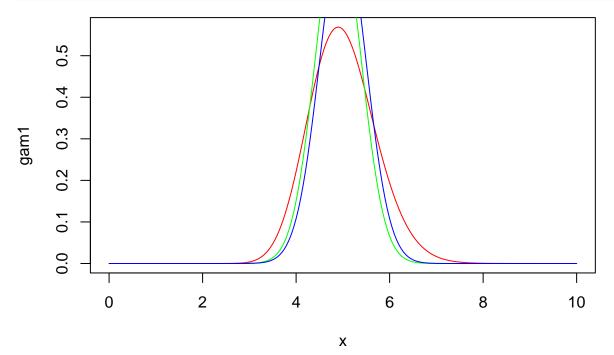
c)

Example when postrior are good, I chose sum of yi = 10, alpha = 40, beta = 10

```
x <- seq(0, 10, by = 0.05)

gam1 <- dgamma(x, shape=50, rate = 10)
nor1 <- dnorm(x,mean = 49/10, 49/10^2)
nor12 <- dnorm(x,mean = 50/10, 50/10^2)

plot(x, gam1, type = "l", col = "red")
lines(x, nor1, type = "l", col = "green")
lines(x, nor12, type = "l", col = "blue")</pre>
```

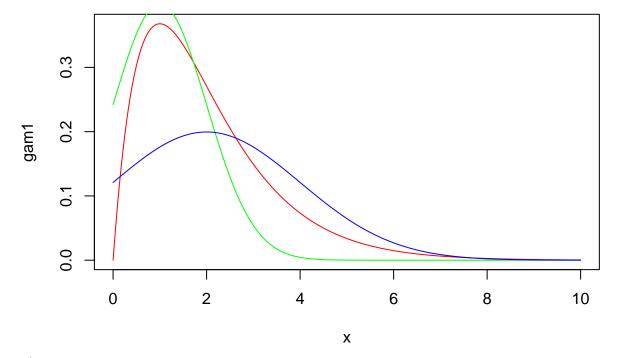


Example when posterior are not good, I chose sum of yi = 10, alpha = 2, beta = 1

```
x <- seq(0, 10, by = 0.05)

gam1 <- dgamma(x, shape=2, rate = 1)
nor1 <- dnorm(x,mean = 1/1, 1/1^2)
nor12 <- dnorm(x,mean = 2/1, 2/1^2)

plot(x, gam1, type = "l", col = "red")
lines(x, nor1, type = "l", col = "green")
lines(x, nor12, type = "l", col = "blue")</pre>
```



d)

When alpha is large we get good normal approximation, when alpha is small we get poor normal approximation.