

Stat500(Section002): Homework #8

Due on Dec.01, 2021 at 3:00pm

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Problem 1

Part a

We using the following code to get the model and calculate the AIC from models.

```
library(faraway)
library(quantreg)
library(MASS)
library(nlme)
library(splines)
data(aatemp)
year<-aatemp$year[-(1:3)]; temp<-aatemp$temp[-(1:3)]
out4<-lm(temp ~ poly(year, 4))
out5<-lm(temp ~ poly(year, 5))
out6<-lm(temp ~ poly(year, 6))
choice<-seq(min(year), max(year), length =5);
knots <- c(rep(choice[1],4), choice[2:4], rep(choice[5],4))
bx <- splineDesign(knots, year, outer.ok = TRUE)
gs <- lm(temp ~ bx-1)
```

```
Aic4 = 112*log(sum(out4$residuals^2)/112) + 2*5
Aic5 = 112*log(sum(out5$residuals^2)/112) + 2*6
Aic6 = 112*log(sum(out6$residuals^2)/112) + 2*7
Aic_splne = 112*log(sum(gs$residuals^2)/112) + 2*7
```

Here, we know the spline model does not have a constant, Hence we only use 7 to multiply 2 instead (7+1). And we can get the result aic for all the models:

1. Orthogonal Polynomials to the 4-th degree AIC:82.98333
2. Orthogonal Polynomials to the 5-th degree AIC:78.39366
3. Orthogonal Polynomials to the 6-th degree AIC:79.61436
4. Cubic spline with three internally evenly spaced knots AIC:78.37617

Using the following code to compute BIC with AIC we get:

```
Bic4 = Aic4 + (log(112)-2)*5
Bic5 = Aic5 + (log(112)-2)*6
Bic6 = Aic6 + (log(112)-2)*7
Bic_splne = Aic_splne + (log(112)-2)*7
```

We can get the result:

1. Orthogonal Polynomials to the 4-th degree BIC:96.57582
2. Orthogonal Polynomials to the 5-th degree BIC:94.70465
3. Orthogonal Polynomials to the 6-th degree BIC:98.64385
4. Cubic spline with three internally evenly spaced knots BIC:97.40566

For BIC, we can know the minimal value comes from model with 5 degree polynomials. Hence, I will recommend Orthogonal Polynomials to the 5-th degree model according to BIC. And for AIC, the minimal value comes from model with cubic spline. Hence I will recommend Cubic spline with three internally evenly spaced knots model.

Problem 2

Part a

AIC:

```
library(leaps)
data(prostate)
step(lm(lpsa ~ ., prostate), k=2)
```

We can get the result:

Start: AIC=-58.32

```
lpsa ~ lcavol + lweight + age + lbph + svi + lcp + gleason +
      pgg45
```

	Df	Sum of Sq	RSS	AIC
- gleason	1	0.0412	44.204	-60.231
- pgg45	1	0.5258	44.689	-59.174
- lcp	1	0.6740	44.837	-58.853
<none>			44.163	-58.322
- age	1	1.5503	45.713	-56.975
- lbph	1	1.6835	45.847	-56.693
- lweight	1	3.5861	47.749	-52.749
- svi	1	4.9355	49.099	-50.046
- lcavol	1	22.3721	66.535	-20.567

Step: AIC=-60.23

```
lpsa ~ lcavol + lweight + age + lbph + svi + lcp + pgg45
```

	Df	Sum of Sq	RSS	AIC
- lcp	1	0.6623	44.867	-60.789
<none>			44.204	-60.231
- pgg45	1	1.1920	45.396	-59.650
- age	1	1.5166	45.721	-58.959
- lbph	1	1.7053	45.910	-58.560
- lweight	1	3.5462	47.750	-54.746
- svi	1	4.8984	49.103	-52.037
- lcavol	1	23.5039	67.708	-20.872

Step: AIC=-60.79

```
lpsa ~ lcavol + lweight + age + lbph + svi + pgg45
```

	Df	Sum of Sq	RSS	AIC
- pgg45	1	0.6590	45.526	-61.374
<none>			44.867	-60.789
- age	1	1.2649	46.131	-60.092
- lbph	1	1.6465	46.513	-59.293
- lweight	1	3.5647	48.431	-55.373
- svi	1	4.2503	49.117	-54.009
- lcavol	1	25.4189	70.285	-19.248

Step: AIC=-61.37

```
lpsa ~ lcavol + lweight + age + lbph + svi
```

	Df	Sum of Sq	RSS	AIC
<none>			45.526	-61.374
- age	1	0.9592	46.485	-61.352
- lbph	1	1.8568	47.382	-59.497
- lweight	1	3.2251	48.751	-56.735
- svi	1	5.9517	51.477	-51.456
- lcavol	1	28.7665	74.292	-15.871

Call:

```
lm(formula = lpsa ~ lcavol + lweight + age + lbph + svi, data = prostate)
```

Coefficients:

(Intercept)	lcavol	lweight	age
0.95100	0.56561	0.42369	-0.01489
lbph	svi		
0.11184	0.72095		

From the result we can see that with backward step, the best model is using lcavol, lweight, age, lbph, svi as predictors with $AIC = -61.37$. We delete gleason, lcp and pgg45 3 predictors. And the AIC decrease from -58.32 to -61.37. And we can see that lcavol is always the most influential predictor here, however, with the delete of other predictors, the influence of lcavol might decrease.

Part b

BIC:

```
step(lm(lpsa ~ ., prostate), k=log(97))
```

We can get the result:

Start: AIC=-35.15

```
lpsa ~ lcavol + lweight + age + lbph + svi + lcp + gleason + pgg45
```

	Df	Sum of Sq	RSS	AIC
- gleason	1	0.0412	44.204	-39.634
- pgg45	1	0.5258	44.689	-38.576
- lcp	1	0.6740	44.837	-38.255
- age	1	1.5503	45.713	-36.377
- lbph	1	1.6835	45.847	-36.095
<none>			44.163	-35.149
- lweight	1	3.5861	47.749	-32.151
- svi	1	4.9355	49.099	-29.448
- lcavol	1	22.3721	66.535	0.030

Step: AIC=-39.63

```
lpsa ~ lcavol + lweight + age + lbph + svi + lcp + pgg45
```

	Df	Sum of Sq	RSS	AIC
- lcp	1	0.6623	44.867	-42.766

```

- pgg45      1      1.1920 45.396 -41.627
- age        1      1.5166 45.721 -40.936
- lbph        1      1.7053 45.910 -40.537
<none>                44.204 -39.634
- lweight    1      3.5462 47.750 -36.723
- svi         1      4.8984 49.103 -34.014
- lcavol      1     23.5039 67.708 -2.849

```

Step: AIC=-42.77

```
lpsa ~ lcavol + lweight + age + lbph + svi + pgg45
```

	Df	Sum of Sq	RSS	AIC
- pgg45	1	0.6590	45.526	-45.926
- age	1	1.2649	46.131	-44.644
- lbph	1	1.6465	46.513	-43.844
<none>			44.867	-42.766
- lweight	1	3.5647	48.431	-39.925
- svi	1	4.2503	49.117	-38.561
- lcavol	1	25.4189	70.285	-3.800

Step: AIC=-45.93

```
lpsa ~ lcavol + lweight + age + lbph + svi
```

	Df	Sum of Sq	RSS	AIC
- age	1	0.9592	46.485	-48.478
- lbph	1	1.8568	47.382	-46.623
<none>			45.526	-45.926
- lweight	1	3.2251	48.751	-43.862
- svi	1	5.9517	51.477	-38.583
- lcavol	1	28.7665	74.292	-2.997

Step: AIC=-48.48

```
lpsa ~ lcavol + lweight + lbph + svi
```

	Df	Sum of Sq	RSS	AIC
- lbph	1	1.3001	47.785	-50.377
<none>			46.485	-48.478
- lweight	1	2.8014	49.286	-47.377
- svi	1	5.8063	52.291	-41.636
- lcavol	1	27.8298	74.315	-7.542

Step: AIC=-50.38

```
lpsa ~ lcavol + lweight + svi
```

	Df	Sum of Sq	RSS	AIC
<none>			47.785	-50.377
- svi	1	5.1814	52.966	-44.966
- lweight	1	5.8924	53.677	-43.673
- lcavol	1	28.0445	75.829	-10.160

```
Call:
lm(formula = lpsa ~ lcavol + lweight + svi, data = prostate)
```

Coefficients:

(Intercept)	lcavol	lweight	svi
-0.2681	0.5516	0.5085	0.6662

We can see the result is that with backward step, the best model is using lcavol,lweight,svi as predictors with $BIC = -50.38$, And we delete gleason,lcp,pgg45,age,lbph 5 predictors which is more than situaion in AIC due to the more strict requirement for BIC.

Part c

Adjusted R²: In this part, we will using regsubsets function to get the best model

```
leaps = regsubsets(lpsa ~ ., prostate, method = "backward")
summary(leaps)$which
summary(leaps)$adjr2
```

We get the result

	(Intercept)	lcavol	lweight	age	lbph	svi	lcp	gleason	pgg45
1	TRUE	TRUE	FALSE	FALSE	FALSE	FALSE	FALSE	FALSE	FALSE
2	TRUE	TRUE	TRUE	FALSE	FALSE	FALSE	FALSE	FALSE	FALSE
3	TRUE	TRUE	TRUE	FALSE	FALSE	TRUE	FALSE	FALSE	FALSE
4	TRUE	TRUE	TRUE	FALSE	TRUE	TRUE	FALSE	FALSE	FALSE
5	TRUE	TRUE	TRUE	TRUE	TRUE	TRUE	FALSE	FALSE	FALSE
6	TRUE	TRUE	TRUE	TRUE	TRUE	TRUE	FALSE	FALSE	TRUE
7	TRUE	TRUE	TRUE	TRUE	TRUE	TRUE	TRUE	FALSE	TRUE
8	TRUE	TRUE	TRUE	TRUE	TRUE	TRUE	TRUE	TRUE	TRUE

[1] 0.5345838 0.5771246 0.6143899 0.6208036 0.6245476 0.6258707 0.6272521 0.6233681

We can see that the best model is using lcavol,lweight,age,lbph,svi,lcp,pgg45 as predictors with $AdjR^2 = 0.62725$. We only delete gleason here.

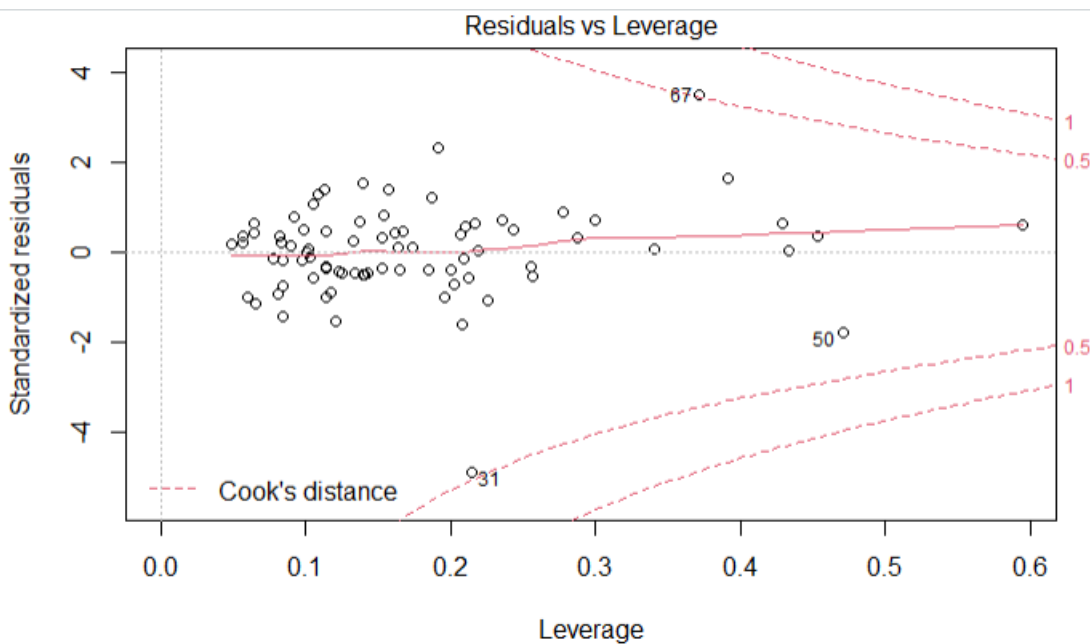
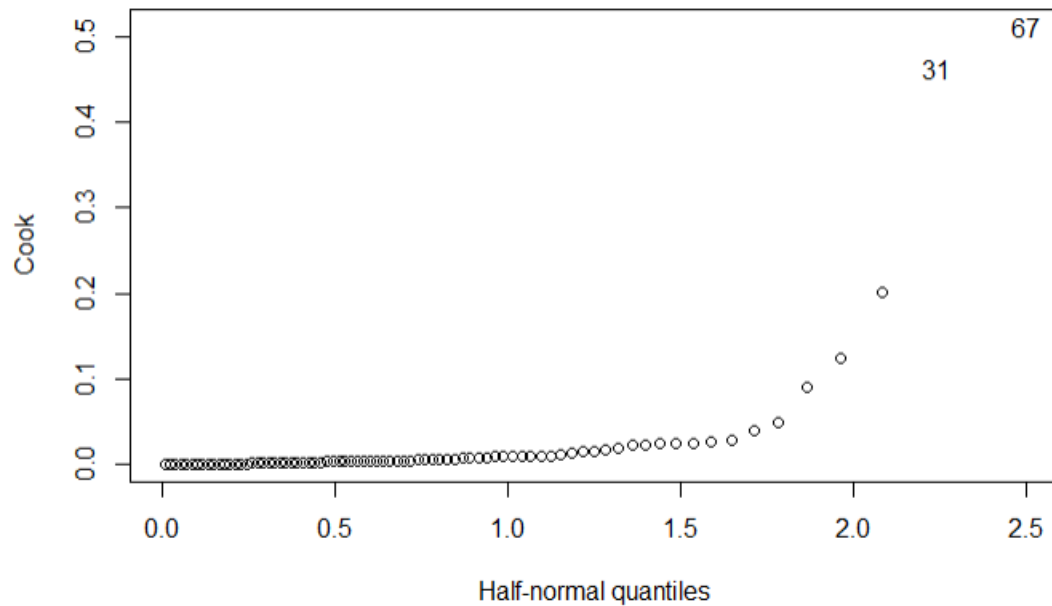
And We can know that with all criteria, best model includes lcavol,lweight,svi as predictors. BIC is the most strict criterion for the number of predictors, AIC is the second, And $AdjR^2$ is the most lenient one with the most predictors stay.

Problem 3

model1: $price \sim bath * size + bed + halfbath + poly(age, 2) + garage + elem$ **Part a**

```
model1 = lm(price ~ bath * size + bed + halfbath + poly(age, 2) + garage + elem, house.dat)
cook = cooks.distance(model1)
halfnorm(cook, ylab="Cook")
plot(model1)
```

We can the result:



We can know that 67,31 is the number of two points which have the largest cook's distances. And from the plot, 67 is greater than 0.5, and should be seen as an influential point for sure. 31 is very close to 0.5 but does not reach it, Hence maybe we should not see it as an influential point. **Part b**

With No.67 31, We will use AIC as our criterion, we using:

```
step(model1)
```

To get result:

```
Start: AIC=602.5
price ~ bath * size + bed + halfbath + poly(age, 2) + garage +
      elem
```

	Df	Sum of Sq	RSS	AIC
– bath:size	2	953.6	106112	599.22
– halfbath	1	802.9	105961	601.10
– poly(age, 2)	2	3505.5	108664	601.12
<none>			105158	602.50
– bed	1	11834.5	116993	609.03
– garage	1	20431.9	125590	614.70
– elem	3	28408.9	133567	615.63

Step: AIC=599.22

```
price ~ bath + size + bed + halfbath + poly(age, 2) + garage +
      elem
```

	Df	Sum of Sq	RSS	AIC
– bath	2	2002.4	108114	596.71
– halfbath	1	697.0	106809	597.74
– poly(age, 2)	2	3839.9	109952	598.06
<none>			106112	599.22
– bed	1	12877.6	118989	606.38
– garage	1	19523.5	125635	610.73
– elem	3	29184.7	135296	612.66
– size	1	23169.9	129282	613.02

Step: AIC=596.71

```
price ~ size + bed + halfbath + poly(age, 2) + garage + elem
```

	Df	Sum of Sq	RSS	AIC
– halfbath	1	338.2	108452	594.96
– poly(age, 2)	2	5436.1	113550	596.64
<none>			108114	596.71
– bed	1	16127.2	124241	605.84
– garage	1	19288.4	127402	607.85
– elem	3	30640.6	138755	610.67
– size	1	23892.8	132007	610.69

Step: AIC=594.96

```
price ~ size + bed + poly(age, 2) + garage + elem
```

	Df	Sum of Sq	RSS	AIC
– poly(age, 2)	2	5104.3	113557	594.64
<none>			108452	594.96
– bed	1	15819.5	124272	603.86
– garage	1	19783.4	128236	606.37
– size	1	23667.8	132120	608.76
– elem	3	30624.4	139077	608.86

Step: AIC=594.64

```
price ~ size + bed + garage + elem
```


	Df	Sum of Sq	RSS	AIC
<none>			113557	594.64
– bed	1	14899	128456	602.51
– garage	1	16102	129658	603.25
– size	1	25001	138557	608.56
– elem	3	52229	165786	618.91

Call:

```
lm(formula = price ~ size + bed + garage + elem, data = house.dat)
```

Coefficients:

(Intercept)	size	bed	garage
194.59	86.24	–21.35	21.81
elemB	elemC	elemD	
–16.83	–64.48	–31.58	

We can see that the best model is using size,bed,garage,elem as the predictors with AIC = 594.64.

Now we delete two points

```
model2 = lm(price~bath*size+bed+halfbath+poly(age,2)+garage+elem,
house.dat,subset = -c(31,67))
step(model2)
```

We get

Start: AIC=533.95

```
price ~ bath * size + bed + halfbath + poly(age, 2) + garage +
elem
```

	Df	Sum of Sq	RSS	AIC
– poly(age, 2)	2	1090	52279	531.60
– halfbath	1	221	51410	532.29
<none>			51189	533.95
– bath:size	2	2823	54013	534.14
– bed	1	5969	57159	540.56
– garage	1	17366	68556	554.74
– elem	3	38318	89507	571.54

Step: AIC=531.6

```
price ~ bath + size + bed + halfbath + garage + elem + bath:size
```

	Df	Sum of Sq	RSS	AIC
– halfbath	1	105	52385	529.75
– bath:size	2	2591	54871	531.37
<none>			52279	531.60
– bed	1	5800	58079	537.80
– garage	1	17185	69464	551.76
– elem	3	58783	111062	584.37

Step: AIC=529.75

```
price ~ bath + size + bed + garage + elem + bath:size
```

	Df	Sum of Sq	RSS	AIC
<none>			52385	529.75
– bath:size	2	2875	55260	529.92
– bed	1	5718	58103	535.83
– garage	1	17546	69931	550.29
– elem	3	58693	111078	582.38

Call:

```
lm(formula = price ~ bath + size + bed + garage + elem + bath:size ,
    data = house.dat, subset = -c(31, 67))
```

Coefficients:

(Intercept)	bath2	bath3	size
428.99	–257.11	–276.74	–40.55
bed	garage	elemB	elemC
–14.26	23.78	–21.49	–71.62
elemD	bath2:size	bath3:size	
–65.62	127.16	137.44	

After removing the two points, we get the best model is $price \sim bath + size + bed + garage + elem + bath : size$ with $AIC = 529.75$. We can know that removing 2 points make AIC become lower and save 2 more predictors.

model2: $price \sim bed + halfbath + age + elem + bath * size$

Part a

```
model_res = lm(price ~ bed + halfbath + age + elem + bath * size, house.dat)
which(model_res$residuals == max(model_res$residuals))
which(model_res$residuals == min(model_res$residuals))
```

We get the No.67 has the largest residual and No.31 has the smallest residual.

```
predict(model_res, house.dat[31,], level = 0.95, interval = "prediction")
predict(model_res, house.dat[67,], level = 0.95, interval = "prediction")
house.dat[c(67, 31),]$price
```

Get the result

	fit	lwr	upr
31	315.2421	221.0477	409.4365
	fit	lwr	upr
67	340.1118	239.2524	440.9713
[1]	441.8	125.9	

We find that the price's real value of two points both not fall into the 95% confidence interval of the prediction. Hence, prices for these two houses do not follow the general regression model that applies to this data set.

Part b

We change age to $\text{poly}(\text{age}, 2)$ here, And do the F-test compare new model and previous model to get:

```
model_res2 = lm(price ~ bed + halfbath + poly(age, 2) + elem + bath * size, house.dat)
```

```
anova(model_res2 , model_res)
```

We get:

```
Model 1: price ~ bed + halfbath + poly(age, 2) + elem + bath * size
Model 2: price ~ bed + halfbath + age + elem + bath * size
  Res.Df    RSS Df Sum of Sq    F Pr(>F)
1      67 125590
2      68 127028 -1    -1438.3 0.7673 0.3842
```

Since the p-value is 0.3842 which is greater than 0.1, We reject to consider the quadratic function of “age”.

Part c

In this part, we change bath2 into reference, and for the difference to bath2 and bath1 in size, we have the hypothesis test: $H_0 : \beta_{bath1:size} = 0, H_A : \beta_{bath1:size} \neq 0$. And we use the following code to get the 95% CI for this parameter. And we will drop No.67 and No.31 here and in next question.

```
new_data = house.dat
new_data$bath = relevel(house.dat$bath, ref="2")
model_new = lm(price~bed + halfbath + age + elem + bath * size ,
new_data , subset=c(31,67))
confint(model_new)
```

We can the result:

	2.5 %	97.5 %
(Intercept)	100.687315	331.71741389
bed	-36.981244	-13.58368931
halfbath	-23.545456	11.69156353
age	-2.680237	5.90137415
elemB	-50.917995	11.96063344
elemC	-92.950503	-42.43537344
elemD	-105.272804	-42.79304665
bath1	30.148424	601.01331500
bath3	-229.738972	122.57906926
size	45.480387	158.80671006
bath1:size	-320.807723	-0.01217422
bath3:size	-59.080744	114.32342012

We can see that 0 is not in the 95% confidence interval of $\beta_{bath1:size}$, And we will reject H_0 , And accept H_1 , which means we think the regression slopes for “size” are different between houses with 1 bathroom and houses with 2-bathrooms.

And for bath2 and bath3, we have the similar hypothesis test: $H_0 : \beta_{bath3:size} = 0, H_A : \beta_{bath3:size} \neq 0$. And we can see 0 is in the 95% confidence interval of $\beta_{bath3:size}$ indeed. Therefore we will not reject H_0 , which means we will think the regression slopes for “size” are the same for houses with 2 bathrooms and houses with 3 bathrooms.

Part d

```
predict(update(model_res , subset=c(67,31)),
data.frame(size=2, age=0 ,elem= "A", bed=3, bath="2", halfbath=1),
level=0.95, interval="prediction")
```

Using the code, we can get the result:

	fit	lwr	upr
1	338.7151	268.5634	408.8668

Since the unit here is 1k dollars, we see that the average price of the house we are interested should be 338715.1 dollars. The 95% confidence interval (268563.408866.8) dollars is the reasonable price range for such a house.