```
import numpy as np
import scipy.stats
import matplotlib.pyplot as plt
```

# Solutions to random variables and probability distributions

# Inverse transform sampling

The PDF is

$$f(x) = \left\{ egin{array}{ll} 0 & x < 0 \ rac{1}{2} {\sin x} & ext{otherwise} \ 0 & x > \pi \end{array} 
ight.$$

The CDF is

$$F(x) = \begin{cases} 0 & x < 0\\ \frac{1}{2} - \frac{1}{2}\cos x & \text{otherwise} \\ 1 & x > \pi \end{cases},$$

with its inverse

$$F^{-1}(u) = \cos^{-1}(1-2u), \ u \in [0,\pi]$$
 .

```
def sin_distr_pdf(x):
    return np.piecewise(x, [(0 <= x) & (x <= np.pi)], [lambda x: 0.5*np.sin(x), 0])

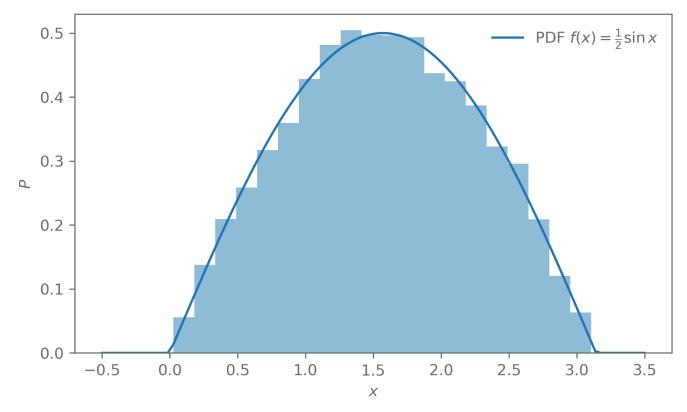
def sin_distr_cdf(x):
    return np.piecewise(x, [x < 0, (0 <= x) & (x <= np.pi)], [0, 0.5 - 0.5*np.cos(x),

def sin_distr_cdf_inv(u):
    if np.any((u < 0) | (1 < u)):
        raise ValueError("u outside domain of (0, 1)")
    return np.arccos(1- 2*u)</pre>
```

```
n = 10000
u = np.random.uniform(size=n)

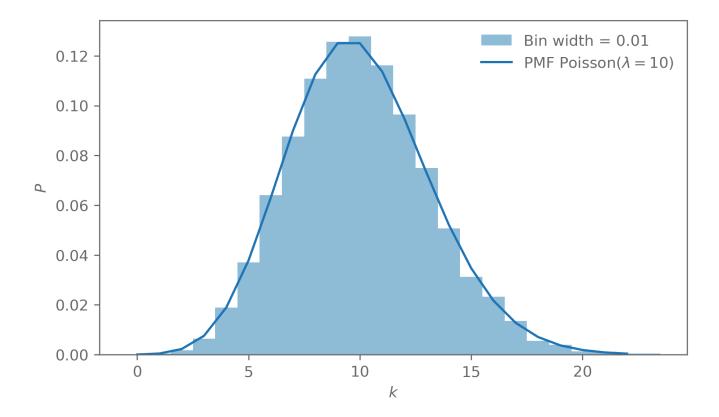
grid = np.linspace(-0.5, 3.5, 100)

plt.hist(sin_distr_cdf_inv(u), bins=20, density=True, alpha=0.5)
plt.plot(grid, sin_distr_pdf(grid), c="C0", label=r"PDF $f(x)=\frac{1}{2}\sin x$")
plt.legend(frameon=False)
plt.xlabel("$x$")
plt.ylabel("$P$");
```

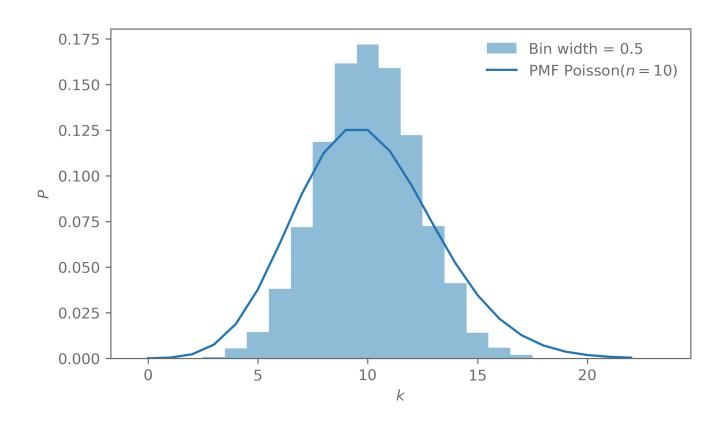


## Poisson distribution

## Uniform samples in small bins



The approximation breaks down when the bin width is too broad:



#### Derivation of Poisson distribution from the binomial distribution

Let  $\lambda$  be the rate of events. This is, in an interval of T=1, we expect  $\lambda$  events to occur.

Devide the interval into n sub-intervals, such that the probability of more than one event happening per sub-interval goes to 0.

The probability of k events happening in the n sub-intervals is then given by a binomial distribution, with  $p=\frac{\lambda}{n}$ :

$$\Pr(X = k) = \binom{n}{k} p^k (1 - p)^{n - k} = \frac{n!}{k! (n - k)!} \left(\frac{\lambda}{n}\right)^k \left(1 - \frac{\lambda}{n}\right)^{n - k}$$

As  $n o \infty$ , we have

$$rac{n!}{(n-k)!n^k} = rac{n(n-1)\dots(n-k+1)}{n^k} 
ightarrow 1 \ , \ \left(1-rac{\lambda}{n}
ight)^{-k} 
ightarrow 1 \ , \ \left(1-rac{\lambda}{n}
ight)^n 
ightarrow \mathrm{e}^{-\lambda} \ .$$

Which leaves

$$\Pr(X=k) o rac{\lambda^k}{k!} \mathrm{e}^{-\lambda} \ (n o \infty) \ ,$$

the PMF of the Poisson distribution

## Change of variables

#### **Sum of Gaussians**

The sum of two Gaussian RVs is also Gaussian distributed. We also know that a Gaussian is completely charaterised by its mean and variance. Let  $X \sim \mathcal{N}(\mu_X, \sigma_X)$ ,  $Y \sim \mathcal{N}(\mu_Y, \sigma_Y)$ , and Z = X + Y. Then

$$\mu_Z = E[Z] = E[X] + E[Y] = \mu_X + \mu_Y$$

and

$$\sigma_Z^2 = \mathrm{Var}[Z] = \mathrm{Var}[X] + \mathrm{Var}[Y] = \sigma_X^2 + \sigma_Y^2 \;.$$

#### Sum of RVs

Let  $X \sim p_X$  and  $Y \sim p_Y$ . What is the distribution of Z = X + Y?

One way to solve this is to write the distribution of Z as the marginal of the joint distribution  $p_{XZ}$  of X and Z:

$$p_Z(z) = \int p_{XZ}(x,z) \mathrm{d}x \;.$$

We find  $p_{XZ}$  using a change of variables:  $(x,y)=u\left((x,z)\right)$ 

$$p_{XZ}(x,z) = p_{XY}\left(u\left((x,z)
ight)
ight) |J_u| \mathrm{d}x\mathrm{d}y$$

The determinant of the Jacobian  $J_u$  is 1, so

$$p_{XZ}(x,z) = p_{XY}(x=x,y=z-x) = p_X(x)p_Y(y=z-x) \; ,$$

where we used that X and Y are independent. Put together we get

$$p_Z(z) = \int p_X(x) p_Y(z-x) \mathrm{d}x \;.$$

The distribution of the sum of two independent RVs is a convolution of their respective distributions.

### Chi-squared

 $X \sim \mathcal{N}(0,1)$ ,  $Z = X^2$ , with z = g(x). Note that g(x) is not monotonic.

$$p_Z(z) = \sum_i p_X(g_i^{-1}(z)) \left| rac{\mathrm{d} g_i^{-1}(z)}{\mathrm{d} z} 
ight| \, ,$$

with 
$$g_i^{-1}(z)=\pm\sqrt{z}$$
 and  $\left|rac{\mathrm{d}g_i^{-1}(z)}{\mathrm{d}z}
ight|=rac{1}{2\sqrt{z}}.$ 

Which leaves

$$p_Z(z) = 2rac{1}{\sqrt{2\pi}} \mathrm{e}^{-rac{1}{2}z} rac{1}{2\sqrt{z}} = rac{1}{\sqrt{2\pi z}} \mathrm{e}^{-rac{1}{2}z}$$