

```

race: 1=white, 2=black, 3=Asian
gender: 1=male, 0=female

```

Mean systolic blood pressures (using `adjmean`) are calculated for six categories: white males, white females, black males, black females, Asian males, and Asian females. Had the outcome been a dichotomous variable for hypertension (1=hypertensive, 0=normal), the probability of being hypertensive (using `adjprop`) would be calculated for each of the six categories.

An overall test of differences is printed, and additionally, a test for interaction. The names for the dummy variables for *xvar1* are `X12`, `X13`, `X14`, etc., where the first number ("1") specifies dummy variables representing *xvar1*, and the second number designates which dummy variable. Dummy variables for *xvar2* are `X22`, `X23`, `X24`, etc. Interaction terms are created by multiplying combinations for the *xvar1* and *xvar2* dummy variables in the model. An *F* test for a linear model or the likelihood ratio test for a logistic model is printed to test an interaction effect.

`adjust(covlist)` specifies an optional list of covariates. If this option is not specified, unadjusted estimates are reported. Values for the covariates default to their means. Alternatively, some (or all) covariates can be set to any value for prediction after estimation of the model. For example, the following statement estimates the systolic blood pressure means for each race category adjusted for age, gender, and cholesterol level:

```
. adjmean sbp, by(race) adjust(age gender chl)
```

If we wanted to calculate these means for someone who is 50 years old, male, with a cholesterol level of 250, we would write

```
. adjmean sbp, by(race) adjust(age=50 gender=1 chl=250)
```

`model` displays the output from `regress` or `logistic`. If this option is not requested, the model is not shown.

`level(#)` specifies the confidence level in percent for the confidence intervals. The default is 95%.

`graph` displays points for each adjusted mean or probability (*y*-axis) by *xvar1* (*x*-axis). If only *xvar1* is requested, confidence intervals are shown also. If *xvar2* is specified, the points are separated by categories of *xvar2*, with *xvar1* still defining the *x*-axis. This sounds more confusing than it is—see examples which follow.

`bar` displays a bar graph rather than point estimates when specified with `graph`.

`graph_options` are any of the options allowed with `graph`, `twoway`.

Examples

To illustrate how these programs work, let's take a look at a study of the cost of treating patients with acute low back pain (Carey et al. 1995). All patients (*n* = 1406) had low back pain (LBP) at entry into the study, and were followed for 6 months.

Contains data from LBP.DTA				
obs:	1,406			Low Back Pain: Outcomes of Care
vars:	10			20 Jan 1998 13:28

1. id	long	%9.0g		Patient ID
2. provtype	byte	%9.0g	praclbl	Provider Type
3. chiromd	byte	%9.0g	chirlbl	Chiropractor vs. MD
4. age	byte	%8.0g		Patient Age in Years
5. gender	byte	%8.0g	genlbl	Patient Gender
6. cost	int	%10.0g		Total Cost of LBP Treatment
7. status4w	byte	%9.0g	func4	Low Back Pain at 4 Weeks
8. severity	byte	%8.0g		Severity of LBP
9. severe3	byte	%9.0g	sevlbl	Severity of LBP
10. sciatica	byte	%8.0g	yesno	Sciatica Present

Example 1

Does the cost of care for an episode of low back pain differ by the type of provider seen? Provider types are 1 for primary care physician, 2 for chiropractor, 3 for orthopedist, and 4 for HMO physician. To answer this, we would look at the mean cost for each provider type (`provtype`).

```

. adjmean cost, by(provtype)
*Unadjusted* Means and 95% Confidence Intervals
Outcome:      Total Cost of LBP Treatment -- cost
Nominal X:    Provider Type -- provtype
Covariates:   (none)

```

```

provtype      numobs      mean      se      lower      upper
1:PC          538      398.0911    26.1085    346.9193    449.2628
2:Chiro       514      626.3891    26.71108    574.0363    678.7419
3:Ortho       162      627.5185    47.57903    534.2653    720.7717
4:HMO         192      340.5104    43.70413    254.8519    426.1689

Test for difference of 4 means:
F(3, 1402) = 19.38
Prob > F < 0.0001

```

We see that the “unadjusted” mean costs differ by type of provider seen, with patients of chiropractors and orthopedists incurring the largest costs (\$626 and \$628), and HMO patients the least (\$341). The overall F test shows a significant ($p < 0.0001$) difference for the means. There are a number of other ways we could have gotten these same unadjusted means quite simply in Stata, e.g.,

```

. tabulate provtype, summarize(cost)
. oneway cost provtype, tabulate

```

Example 2

An argument might be made that the cost of treating LBP is higher for chiropractors and orthopedist because their patients have more severe pain. We rerun Example 1, but this time we will adjust for variables such as age, severity (a scale from 0 to 23, where 0 means no impairment and 23 means complete incapacitation), and sciatica (another measure of severity, where 1 means sciatica present, 0 means no sciatica). In addition to requesting the adjusted means, we request the model to be printed and a graph of the results.

```

. adjmean cost, by(provtype) adjust(age severity sciatica) model graph ylabel
xlabel(1,2,3,4)

```

Source	SS	df	MS	Number of obs =	1401
Model	89651393.4	6	14941898.9	F(6, 1394) =	48.04
Residual	433565958	1394	311022.925	Prob > F =	0.0000
				R-squared =	0.1713
				Adj R-squared =	0.1678
				Root MSE =	557.69

cost	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
X2	248.1997	34.51434	7.191	0.000	180.4941 315.9054
X3	200.9192	50.19444	4.003	0.000	102.4544 299.384
X4	-24.06906	47.30386	-0.509	0.611	-116.8635 68.72537
age	-3.698749	1.149891	-3.217	0.001	-5.954453 -1.443045
severity	26.00893	2.204246	11.799	0.000	21.68494 30.33293
sciatica	238.0384	36.73184	6.480	0.000	165.9827 310.094
_cons	201.8666	57.91298	3.486	0.001	88.26064 315.4726


```

*Adjusted* Means and 95% Confidence Intervals
Outcome:      Total Cost of LBP Treatment -- cost
Nominal X:    Provider Type -- provtype
Covariates:   age severity sciatica

```

provtype	numobs	adjmean	se	lower	upper
1:PC	536	386.3189	24.11297	339.0583	433.5794
2:Chiro	513	634.5186	24.69707	586.1132	682.924
3:Ortho	161	587.238	44.0196	500.9612	673.5149
4:HMO	191	362.2498	40.61471	282.6465	441.8532

```

Test for difference of 4 means:
F(3, 1394) = 22.43
Prob > F < 0.0001
( See Figure 1 below)

```

The variable, X2, is the dummy variable which compares the adjusted mean cost for chiropractors (a `provtype` of 2) versus primary care physicians (a `provtype` of 1), X3 compares orthopedists (a `provtype` of 3) to primary care, and X4 compares HMO's (a `provtype` of 4) to primary care. The orthopedists' mean costs decreases from \$628 (Example 1 result) to \$587 after adjustment for age, severity of low back pain, and sciatica. Primary care physician-adjusted mean costs are lower as well, while chiropractors and HMO-adjusted costs increase. There is still an overall statistical difference in these means. Chiropractors and orthopedists charge more, even after adjusting for age and measures of severity. A graph of the adjusted means and 95% confidence intervals is in Figure 1. Had we preferred a bar graph, we could have typed

```

. adjmean cost, by(provtype) adjust(age severity sciatica) model ylabel graph bar
(output omitted)
( See Figure 2 below)

```

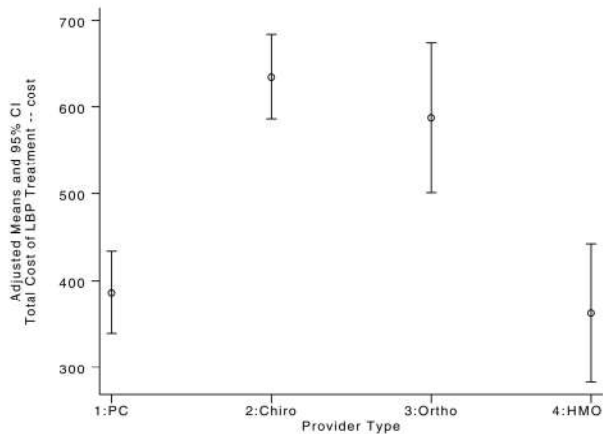


Figure 1. Adjusted means and 95% CI.

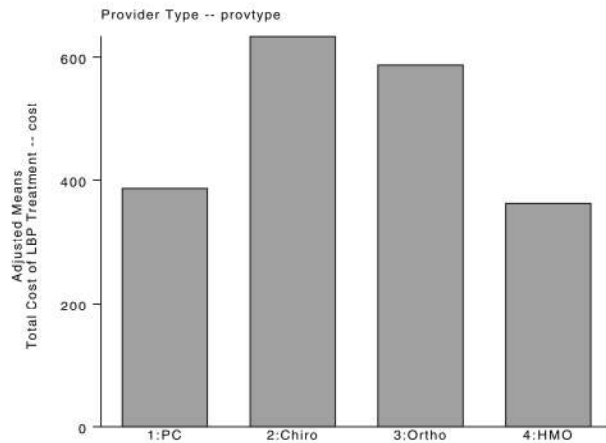


Figure 2. Adjusted means bar graph.

If you would like to test the difference of any pairs of means (not in the model), you can do so with the `test` statement after running `adjmean`. For example, suppose we wanted a test for a comparison between mean costs for chiropractors and orthopedists:

```
. test _b[X2] = _b[X3]
      (1) X2 - X3 = 0.0
           F( 1, 1394) =    0.87
           Prob > F =    0.3499
```

Example 3

Suppose this time we want to compare costs for chiropractors versus any type of physician (`provtype` = 1, 3, and 4 combined), adjusted for age, severity, and sciatica. The variable is called `chiomd` (1=chiropractor, 0=MD). In addition, after fitting the model, let's calculate the predicted mean cost by chiropractor versus MD for someone who is 40 years old, has a low severity score (2), and no sciatica (0). The program does this by estimating the betas for the regression equation, then calculating the adjusted means by substituting specified values for the covariates into the equation.

```
. adjmean cost, by(chiomd) adjust(age=40 severity=2 sciatica=0)
*Adjusted* Means and 95% Confidence Intervals
Outcome:      Total Cost of LBP Treatment -- cost
Nominal X:    Chiropractor vs. MD -- chiomd
Covariates:   age=40 severity=2 sciatica=0

chiomd      numobs   adjmean      se      lower      upper
0:MD         888    134.7998    27.6567    80.59362    189.0059
1:Chiro       513    351.9856    31.48407    290.2779    413.6932

Test for difference of 2 means:
F(1, 1396) =    48.28
Prob > F <    0.0001
```

Overall, the mean costs are lower for these patients with less severe back pain, though we see that chiropractors have significantly higher costs (\$352) than MD's (\$135). Let's repeat Example 3 for patients who are the same age (40), but have more severe low back pain, say a high severity score of 20 and the presence of sciatica (1).

```
. adjmean cost, by(chiomd) adjust(age=40 severity=20 sciatica=1)
*Adjusted* Means and 95% Confidence Intervals
Outcome:      Total Cost of LBP Treatment -- cost
Nominal X:    Chiropractor vs. MD -- chiomd
Covariates:   age=40 severity=20 sciatica=1

chiomd      numobs   adjmean      se      lower      upper
0:MD         888    847.096    37.21108    774.1641    920.0288
1:Chiro       513    1064.282    40.35933    985.1794    1143.385

Test for difference of 2 means:
F(1, 1396) =    48.28
Prob > F <    0.0001
```

Not surprisingly, the cost of care increases substantially when the patients have more severe low back pain. Chiropractors continue to have significantly higher mean costs (\$1,064) than MD's (\$847).

Example 4

This example estimates the mean costs by categories of two nominal variables: `chiromd` (1=chiropractor, 0=MD) and `sciatica` (1=sciatica present, 0=no sciatica), adjusted for age and severity. The model will be printed, and a graph is requested, defaulting to points which will be connected.

```
. adjmean cost, by(chiromd sciatica) adjust(age severity) model graph ylabel xlabel(0,1) c(11)
```

Source	SS	df	MS			
Model	83919668.0	5	16783933.6	Number of obs =	1401	
Residual	439297683	1395	314908.733	F(5, 1395) =	53.30	
				Prob > F =	0.0000	
				R-squared =	0.1604	
				Adj R-squared =	0.1574	
Total	523217351	1400	373726.679	Root MSE =	561.17	

	cost	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
X12		215.2707	35.58975	6.049	0.000	145.4555	285.086
X22		236.85	47.37293	5.000	0.000	143.9201	329.7798
I22		8.35065	74.12591	0.113	0.910	-137.0596	153.7609
age		-3.800583	1.151852	-3.300	0.001	-6.060132	-1.541035
severity		26.23992	2.218988	11.825	0.000	21.88701	30.59283
_cons		234.9213	54.74507	4.291	0.000	127.5297	342.3128


```
*Adjusted* Means and 95% Confidence Intervals
Outcome:      Total Cost of LBP Treatment -- cost
Nominal X1:   Chiropractor vs. MD -- chiromd
Nominal X2:   Sciatica Present -- sciatica
Interaction:   chiromd * sciatica
Covariates:   age severity
```

chiromd	sciatica	numobs	adjmean	se	lower	upper
0:MD	0:No	705	365.1132	21.21326	323.5359	406.6904
0:MD	1:Yes	183	601.9631	42.1138	519.4216	684.5047
1:Chiro	0:No	387	580.3839	28.6184	524.2928	636.475
1:Chiro	1:Yes	126	825.5845	50.12164	727.348	923.8211

```
Test for difference of 4 means:
F(3, 1395) = 31.83
Prob > F < 0.0001

Test for interaction of chiromd * sciatica:
F(1, 1395) = 0.01
Prob > F = 0.9103
( See Figure 3 below)
```

The dummy variable for `chiromd` is `X12` and the dummy variable for `sciatica` is `X22`, with an interaction variable (`I22`) created between them. (Note: This would have been simpler to read if the variables had retained their original names, but often there are more than two categories per variable, and naming gets complicated. Most of the time, we're not interested in seeing the model, anyway. The default summary statistics usually give all the information we want.)

The highest mean cost is for patients with sciatica seeing chiropractors (\$826), but costs are higher also for sciatica patients seeing physicians (\$602). There is no evidence of interaction ($p = 0.9013$) which we can see from the regression table for `I22`, or from the F statistic at the bottom of the output. The graph (see Figure 3) confirms this. Although it may not be appropriate to connect the points of a nominal variable, it's done here to demonstrate that the `connect` option can work, as well as to show that two lines couldn't be much more parallel (a clear indication of no interaction).

Had we preferred a bar graph, we could have typed

```
. adjmean cost, by(chiromd sciatica) adjust(age severity) model graph bar ylabel
( See Figure 4 below)
```

(Graphs on next page)

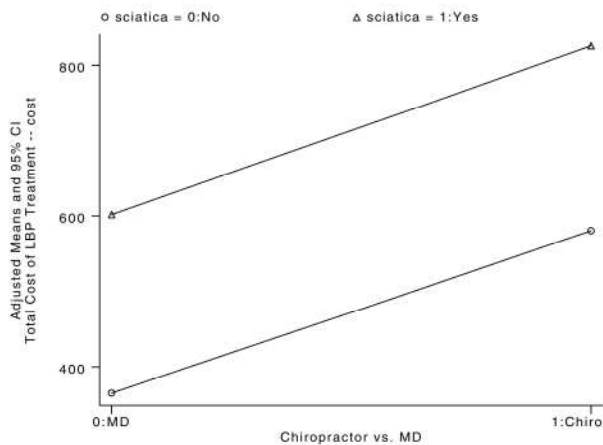


Figure 3. Mean cost of chiromd & sciatica adjusted for age & severity.

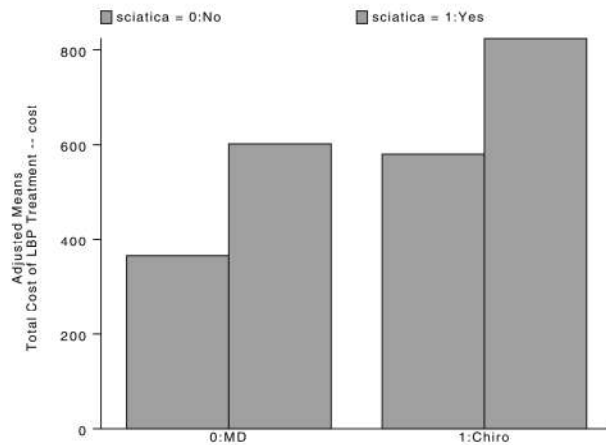


Figure 4. Bar graph version of Figure 3.

Example 5

Let's look at an example where there is significant interaction. When we examine the adjusted mean costs for categories of chiromd by gender (1=male, 0=female), we get the following results (model not printed):

```
. adjmean cost, by(chiromd gender) adjust(age severity sciatica) graph bar ylabel
(output omitted)
*Adjusted* Means and 95% Confidence Intervals
Outcome:      Total Cost of LBP Treatment -- cost
Nominal X1:    Chiropractor vs. MD -- chiromd
Nominal X2:    Patient Gender -- gender
Interaction:   chiromd * gender
Covariates:    age severity sciatica
```

chiromd	gender	numobs	adjmean	se	lower	upper
0:MD	0:Female	182	304.1542	41.50648	222.803	385.5054
0:MD	1:Male	705	446.4079	20.99842	405.2517	487.5641
1:Chiro	0:Female	60	881.6437	72.07227	740.3847	1022.903
1:Chiro	1:Male	453	601.9128	26.28687	550.3915	653.4341

```
Test for difference of 4 means:
F(3, 1393) = 23.94
Prob > F < 0.0001

Test for interaction of chiromd * gender:
F(1, 1393) = 22.19
Prob > F < 0.0001
```

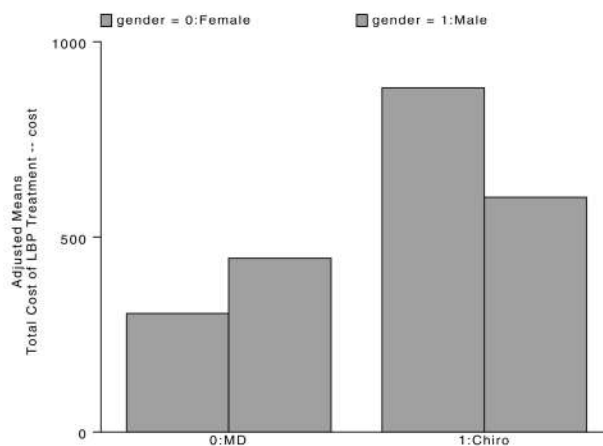


Figure 5. Mean costs by chiromd and gender.

Within chiropractors, mean costs are higher for females (\$882) than males (\$602), whereas within physicians, the relationship is reversed—females: \$304, males: \$446. The test for interaction is highly significant ($p < 0.0001$), and the bar graph (see Figure 5) illustrates the relationship visually.

Example 6

If we repeat the last example using our original four categories for provider (**provtype**), we find that females have higher mean costs than males if they visit chiropractors, but have about the same or lower costs than males for primary care, orthopedic, or HMO physicians. This is easiest to see in Figure 6.

```
. adjmean cost, by(provtype gender) adjust(age severity sciatica) graph bar ylabel
(output omitted)

*Adjusted* Means and 95% Confidence Intervals

Outcome:      Total Cost of LBP Treatment -- cost
Nominal X1:    Provider Type -- provtype
Nominal X2:    Patient Gender -- gender
Interaction:   provtype * gender
Covariates:    age severity sciatica
```

provtype	gender	numobs	adjmean	se	lower	upper
1:PC	0:Female	89	268.8453	58.88354	153.4357	384.2549
1:PC	1:Male	447	409.9181	26.25605	358.4572	461.379
2:Chiro	0:Female	60	882.098	71.71287	741.5433	1022.653
2:Chiro	1:Male	453	601.6091	26.15592	550.3444	652.8737
3:Ortho	0:Female	25	569.0792	110.9323	351.656	786.5024
3:Ortho	1:Male	136	590.6166	47.57316	497.3749	683.8583
4:HMO	0:Female	68	254.8119	67.46792	122.5772	387.0466
4:HMO	1:Male	122	419.2326	50.38327	320.4832	517.982

```
Test for difference of 8 means:
F(7, 1389) = 12.94
Prob > F < 0.0001

Test for interaction of provtype * gender:
F(3, 1389) = 7.34
Prob > F < 0.0001
```

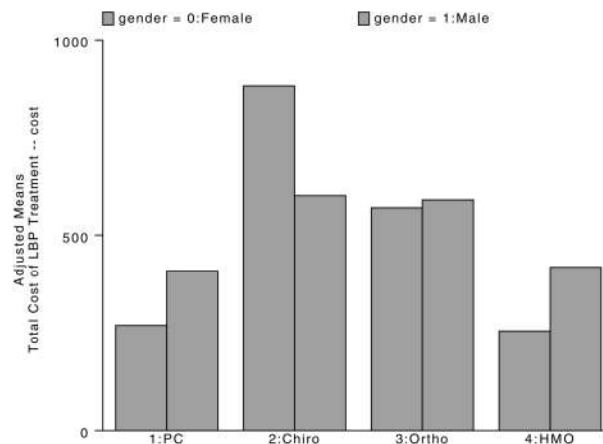


Figure 6. Mean costs adjusted by provider type and gender.

Example 7

In the previous examples, we looked at adjusted means (costs), so let's look at a couple of examples where the outcome is dichotomous rather than continuous, and we want to estimate adjusted probabilities using logistic regression (**adjprop**). Another outcome we considered in this study was how quickly patients got over their back pain. Of course, each provider type was sure their patients would get better quicker. We created a variable, **status4w**, which measured whether patients still had back pain at four weeks (**status4w**: 1=not better, 0=better). So, our logistic regression model predicts the probability of still having low back pain after four weeks. Since patients with less pain at baseline may have improved more quickly, we will adjust for **age**, **severity**, and **sciatica**. The following statement includes a request to display the logistic regression model results and a bar graph by provider type:

```
. adjprop status4w, by(provtype) adj(age severity sciatica) model graph bar ylabel
Logit Estimates
Number of obs = 1401
chi2(6) = 170.75
Prob > chi2 = 0.0000
Pseudo R2 = 0.1279
Log Likelihood = -582.29855
```