

Expectations Formation with Fat-Tailed Processes: Evidence and Theory*

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January 25, 2026

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Abstract

Using a large sample of forecasts of 22 firm-level variables, we document three facts: (i) the relationship between forecast revisions and future forecast errors is strongly non-linear, (ii) the distributions of the underlying processes have fat tails, and (iii) extreme realizations tend to mean-revert. Next, we formally show that these three facts are consistent with a model in which the underlying process is non-Gaussian, but forecasters fail to recognize this. Finally, we provide additional evidence in support of our theory by showing that it can explain evidence from an online forecasting experiment, that the non-linear relationship between errors and revisions is also present in macroeconomic forecasts, and that it provides an explanation for the presence of non-linearity in the momentum of stock returns.

*We are grateful to Xavier Gabaix for very helpful feedback and Mauro Cazzaniga for outstanding research assistance. We also thank Nick Barberis, Eben Lazarus, Stefan Nagel, Larry Schmidt, Kelly Shue, and seminar participants at CFM, Columbia, MIT, Yale SOM, NYU Stern, University of Toronto, Stanford GSB, Wharton, Chicago Booth, NBER Behavioral Finance, and NBER Forecasting & Empirical Methods for helpful comments and suggestions.

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The recent literature on expectations formation has accumulated evidence that subjective forecast errors are predictable. To explain this predictability, theories of expectations formation have been proposed with a variety of ingredients such as information frictions (Coibion and Gorodnichenko, 2015), sluggish updating (Bouchaud et al., 2019; Ma et al., 2024), representativeness bias (Bordalo et al., 2019, 2020a), and imperfect memory (Afrouzi et al., 2023). A common feature of these theories is that they assume that the underlying data-generating process (DGP) is Gaussian. However, in many economic applications, tails are “fat”, meaning that extreme outcomes occur much more frequently than in a Gaussian distribution. For example, fat tails are present in the cross-sectional distributions of many economic variables, including population size, income, wealth, and firm size (Gabaix, 2009). More importantly, fat tails are also common in dynamic processes about which agents need to continuously update their expectations, such as stock returns (Fama, 1965), firm growth (Stanley et al., 1996), and income growth (Guvenen et al., 2014). Therefore, understanding how expectations are formed in the presence of fat tails is important for studying many dynamic decisions, such as corporate investment, household consumption-savings, or portfolio choice.

In this paper, we study expectations formation in the presence of fat tails, both empirically and theoretically. We begin with an empirical analysis of a large dataset of subjective forecasts made by professional analysts about over twenty firm outcomes, including sales growth, earnings per share, dividend payouts, and profitability ratios. The large size of the dataset is useful because it allows us to study the tails of the distribution. Using these data, we document three robust empirical facts: (i) the relationship between forecast revisions and future forecast errors—the variables used in Coibion and Gorodnichenko (2015) regressions—is strongly non-linear for all of the firm-level outcomes; (ii) the distributions of the underlying processes have fat tails; and (iii) the conditional expectation of future realizations is non-linear in current realizations, with persistence in the bulk and mean reversion in the tails. Next, we build an expectations formation model that connects these facts. The key ingredients in our model are that the underlying process is non-Gaussian, but forecasters fail to recognize this and incorrectly fit a linear model to the data. Finally, we provide additional evidence in support of our theory by showing that it can explain expectations formation in an online forecasting experiment, that the non-linear relationship between errors and revisions is also present in macroeconomic forecasts, and that it provides an explanation for the presence of non-linearity in the momentum of stock returns. Overall, our analysis suggests that forecasters partially ignore the presence of fat tails and rely too much on linear models, which is an insight that could be leveraged in models of household and firm behavior with non-Gaussian DGPs (Guvenen et al., 2021; Boar et al., 2025).

The data that we use in our empirical analysis are a large panel of analyst forecasts and realizations of 22 firm-level variables from IBES that cover the most important financial statement items, such

as revenue, earnings, cash flow, and debt levels. The main benefit of these data is that the sample is large, with over 3 million observations covering more than 20 years. Using these data, the first and most important fact that we document is that the relationship between current forecast revisions and future forecast errors is strongly non-linear. In some papers, revisions linearly and *positively* predict forecast errors, a feature commonly interpreted as evidence of underreaction (Coibion and Gorodnichenko, 2015; Bouchaud et al., 2019). In others, revisions linearly and *negatively* predict forecast errors, which is evidence of overreaction (Bordalo et al., 2019, 2020a). In our panel, we show that both coexist. For large revisions, there is a *negative* relationship between revisions and future errors, indicating overreaction to news. For more moderate revisions, the relationship between revisions and future errors becomes more positive, most of the time indicating underreaction to news (or a much milder amount of overreaction). We show that this non-linear relation between errors and revisions is present for all 22 firm-level variables in our sample.

This non-linear relationship between errors and revisions is robust to a battery of robustness checks that we conduct. For example, we show that this non-linearity is present in both individual *and* consensus forecasts and, therefore, is not driven by forecast aggregation (Bordalo et al., 2020a). Additionally, we show that the non-linearity is driven by cross-sectional rather than aggregate variation and does not vary with analysts' forecasting experience. These findings suggest that it is unlikely to be driven by a slow convergence of learning that can occur with one short time series (Bianchi et al., 2022; Farmer et al., 2024).

Our second empirical fact is that the realizations of all 22 firm-level variables in our sample have fat tails. To detect the presence of fat tails, we measure the kurtosis of each variable and the extent to which it is well-approximated by a power law in the tails of the distribution (Gabaix, 2009). We find that most variables in our sample have kurtosis above six and power law tail parameters below three (with R^2 's of log-log regressions above 0.99), both of which indicate the presence of fat tails relative to a normal distribution. These fat tails do not arise from heterogeneity in volatility across firms (Wyart and Bouchaud, 2003) and do not arise from time-varying aggregate volatility.

Our third and final fact is that the conditional expectation of future realizations given the current realization is non-linear in a way that is reminiscent of the non-linearity in our first fact. In particular, we find that the relationship between current and one-year-ahead realizations is increasing and linear in the bulk of the distribution for most of the variables in our sample. However, in the tails of the distribution, this persistence is less strong for all variables and often changes sign, indicating that extreme values tend to mean-revert. As described below, this non-linearity helps inform the exact way the DGP in our model deviates from that of standard Gaussian models.

In the next part of the paper, we develop a model of expectations formation that is designed to connect the above three facts, while also being simple and portable. In the model, the DGP for the forecasting variable contains a persistent and transitory component, which is common in models of dividend growth (Bansal and Yaron, 2004; Lettau and Wachter, 2007) and income dynamics (Guvenen et al., 2014). Our two key assumptions are that the transitory component follows a power law distribution, a tractable way of characterizing processes with fat tails (Gabaix, 2009), and that the persistent component has a thinner tail than the transitory component.¹

We show that this model of the DGP is consistent with our second and third facts. The model generates fat tails by assumption, consistent with our second fact. To show that it generates the non-linear conditional expectation of future realizations conditional on current realizations—our third fact—we leverage a result from empirical Bayes theory known as Tweedie’s formula (Efron, 2012). Although this expectation cannot be expressed in closed-form due to the non-normality, this result allows us to characterize it as a function of the observable density of realizations. We show that this result implies that the conditional expectation is locally linear when the density is locally Gaussian, as in the bulk of the distribution of most variables. However, when the density is locally non-Gaussian, as in the tails, the conditional expectation is no longer linear and is asymptotically decreasing in the current realization. Intuitively, very large realizations are likely due to the transitory component and, hence, are not likely to persist.

Given this DGP, consistent with our second and third facts, we can also explain the non-linear relationship between forecast errors and revisions—our first fact—with a single assumption: agents construct their forecasts ignoring fat tails. Formally, we assume that agents form forecasts according to the Kalman filter, which would be the rational expectation in our model if the transitory and persistent components were both normally distributed, and is equivalent to assuming that forecasters use the linear forecast with the lowest possible mean-squared error. Using results from the theory of regularly varying functions (Bingham et al., 1987; Geluk and de Vries, 2006), we prove that this assumption is enough to generate overreaction in the tails and underreaction in the bulk. This is because large revisions are driven by large realizations that asymptotically come from the transitory component of the DGP. While a rational forecaster would recognize that these extreme shocks are unlikely to persist, our agents who ignore fat tails do not and, therefore, overreact. However, because our agents are unbiased unconditionally, overreaction in the tails has to be compensated by underreaction in the bulk, consistent with our first fact.

While we do not provide a deep psychological microfoundation for why agents ignore fat tails,

¹The latter assumption is consistent with our third fact: if the persistent component had the thicker tail, we would not observe mean reversion in the tails of the distribution. Additionally, our results generalize to the case in which the persistent component is regularly varying (Bingham et al., 1987).

we view it as a form of bounded rationality in which forecasters use a misspecified model, as in Fuster et al. (2010) and Gabaix (2019). To illustrate this point, we show that our assumption that forecasters use the Kalman filter is analogous to assuming that forecasters are insensitive to the strength of the signal that they observe. The latter assumption is consistent with laboratory and field evidence in Augenblick et al. (2024) and Ba et al. (2024), and suggests that insensitivity to signal strength in our time series setting is a possible reason why forecasters ignore the presence of fat tails.

We conclude by providing three pieces of evidence from different domains that support our theory that forecasters do not fully recognize the presence of fat tails. First, we run an online forecasting experiment similar to Afrouzi et al. (2023), but where the underlying process has fat tails. The key benefit of this experiment is that it allows us to test our theory of expectations formation directly by experimentally varying the features of the data-generating process. When we run the experiment using our estimated DGP, we find that the relationship between errors and revisions is non-linear, consistent with our first fact. In contrast, when forecasters forecast a similar process with no fat tails (as in Afrouzi et al. 2023), we find no evidence of a non-linear relationship between errors and revisions. These findings provide direct evidence that the non-linear relationship between errors and revisions in the data is driven by the fat tails of the DGP, which is the key result of our theory.

Second, we show that the non-linear relationship between forecast errors and revisions—our first fact—is also present in forecasts of macroeconomic outcomes from the Survey of Professional Forecasters (1968–2025). Of the 14 variables that are available, we find that the relationship between individual-level forecast errors and revisions is non-linear, with a more negative slope in the tails of the distribution for all but one variable. For most variables, the slope is positive in the bulk of the distribution and negative in the tails of the distribution, similar to our findings from firm-level forecasts. These findings suggest that the prevalence of more overreaction in the tails of the distribution is more general than our sample of firm-level forecasts.

Finally, we show that our model makes predictions for return momentum that are supported by the data. We translate expectations of sales growth, one of the variables in our firm-level sample, into returns by applying the Campbell (1991) return decomposition with a constant subjective discount rate (as in Bouchaud et al. 2019 and Nagel and Xu 2022). Our model predicts that the relationship between past and future returns should be positive in the bulk of the distribution, where underreaction to news is dominant. In contrast, it predicts mean-reversion of returns in the tails, where agents fail to recognize that the extreme shocks are not persistent. We find some support for this prediction in the universe of smaller stocks: for these stocks, momentum tends to mean-revert for extreme losers and winners.

Related literature. Economists have long recognized the benefits of using survey data on expectations to study economic choice models (Manski, 2004). But the recent literature has shifted the emphasis to the study of forecast error predictability (e.g. Coibion and Gorodnichenko 2015). This paper contributes to this branch in two ways. First, we contribute to the empirical literature that has documented under- and overreaction across many forecasting variables and horizons. Broadly speaking, this literature tends to find evidence of underreaction when looking at shorter-term or consensus forecasts (Coibion and Gorodnichenko, 2015; Bouchaud et al., 2019), and overreaction when looking at longer-term or individual forecasts (Bordalo et al., 2019, 2020a; Wang, 2021). Relative to this empirical literature, our contributions are to provide field evidence of both under- and overreaction within the *same* forecasting variable and horizon, and lab evidence that the degree of overreaction can vary within a sample depending on the tail parameter in the DGP.

Second, we contribute to the literature that proposes models of belief formation that can generate under- and overreaction, including constant-gain learning (Nagel and Xu, 2022), selective recall (Bordalo et al., 2020b), and biased perceptions of autocorrelation (Wang, 2021). Most of this literature works with models in which data-generating processes are Gaussian, implying that conditional expectations are linear. In contrast, our empirical analyses highlight how this relationship can be quite non-linear, and we provide a theory that links this non-linearity to the presence of under- and overreaction within a given process. Our focus on non-Gaussian dynamics is similar to Kozlowski et al. (2020), but we focus on forecasters ignoring these dynamics rather than learning.² With forecasts of a single time series, rational learning can converge relatively slowly (Farmer et al., 2024), generating an in-sample relationship between errors and revisions even under full rationality (Singleton, 2021; Bianchi et al., 2022). In contrast, one benefit of having the large cross-section in our data is that, under fairly mild assumptions, it averages out the portion of the error-revision predictability that is driven by a short time series. Our assumption that forecasters (partially) ignore non-Gaussian dynamics is inspired by the literature on bounded rationality, which argues that economic agents use simplified models to minimize computation costs or because of cognitive constraints (Fuster et al., 2010; Gabaix, 2019; Molavi et al., 2024). However, it is possible that this assumption could be microfounded via Bayesian learning about the tail parameter of the process.

Three closely related papers are Kwon and Tang (2025), Augenblick et al. (2024), and Graeber et al. (2025). Kwon and Tang (2025) also provide a model of belief formation with non-Gaussian dynamics. In their model, news events belong to categories with different power law distributions, and forecasters have diagnostic expectations, causing them to overreact to news from categories with fatter tails and underreact to news from categories with thinner tails. This prediction is similar

²See Dew-Becker et al. (2024) for a characterization of Bayesian learning with arbitrary non-Gaussian dynamics.

to that of our model. An important difference is that this provides a theory of why over- and underreaction would vary depending on the category from which a realization is drawn, while our model provides a theory of why over- and underreaction would vary even *within* a category. [Augenblick et al. \(2024\)](#) propose a model in which forecasters incorrectly perceive signal quality and shrink it to a default, which leads to overreaction to weak signals and underreaction to strong ones. As we discuss in the paper, insensitivity to signal strength is a possible microfoundation for ignoring fat tails. Finally, [Graeber et al. \(2025\)](#) analyze the S-shaped relationship between returns and earnings surprises, which loosely map into revisions and forecast errors. In their theory, the strong sensitivity between returns and surprises around zero comes from overreaction that occurs at a category boundary, while the lower sensitivity away from zero comes from dampening within a category due to noisy perceptions. Our main fact—underreaction in the bulk and overreaction in the tails—points to a different mechanism in our setting.

Through our model of the DGP with fat tails, we connect the expectations formation literature with the literature on power laws. The omnipresence of power laws ([Gabaix, 2009](#)) suggests that the misperception of fat tails that we document is likely important for understanding subjective forecasts in other settings. The facts that we document about the data-generating process of firm-level realizations are consistent with the literature on firm dynamics. For example, we find that sales growth (rather than its level) has fat tails, which is a violation of Gibrat’s law that was first documented by [Stanley et al. \(1996\)](#) and recently emphasized by [Boar et al. \(2025\)](#). Our third fact is consistent with [Jaimovich et al. \(2025\)](#), who find that revenue is more persistent in the bulk of the distribution than in the tails and exhibits fat tails.

Outline. Section 1 describes our data on forecasts of 22 firm-level variables. Section 2 documents our three main facts. Section 3 describes the model that we build to connect and explain these facts. Section 4 provides additional evidence in support of our theory from an online forecasting experiment, macroeconomic forecasts, and data on stock returns.

1 Firm-Level Forecast Data

1.1 Data Source and Sample

Our analysis primarily relies on an annual panel between 2000 and 2023 of forecasts issued by over 38,000 security analysts who follow around 36,000 U.S. and international firms. These forecasts cover 22 different annual firm-level variables at different horizons. We obtain these analyst-level forecasts from LSEG IBES Unadjusted Detail files (henceforth “Detail”). In robustness analysis, we

also study consensus analyst forecasts, which come from the LSEG IBES Summary Statistics files (henceforth “Summary”). Both of these files contain current estimates as of the third Wednesday of each month. We extract the forecasts issued in the third month of each fiscal year $t + 1$, which ensures that the results for fiscal year t have been released. We focus on forecasts for two different horizons: fiscal years $t + 1$ and $t + 2$. Along with one and two-year forecasts, LSEG IBES also provides data on the realizations of the forecasted variables, which the provider ensures are measured consistently with the forecasts made by analysts.

1.2 Definition of Forecasting Variables

We study forecasts and realizations of all the variables available in the Summary and Detail files. The list of these variables is provided in [Table 1](#). These variables fall into one of three categories: accounting items in local currency (e.g., Revenue), items in local currency on a “per-share” basis (e.g., Earnings Per Share), and accounting ratios in percentage points (e.g., Return on Assets). To avoid imposing unnecessary sample filters, we choose to include all 22 variables regardless of their data coverage, which we discuss below.

The only transformation we apply to the variables of interest is done to ensure they are stationary. Out of the 22 variables, three are accounting ratios that are already in percentage points, which are presumably already stationary and therefore not transformed. The remaining 19 variables are accounting items in local currency or on a per-share basis. To make them stationary, we convert them from levels to growth rates, as in [Bordalo et al. \(2020a\)](#). More specifically, let X_{it+h} denote the realization of one of these variables for firm i at $t + h$. We transform this variable into a log growth rate according to $g_{it+h} = \log X_{it+h} - \log X_{it+h-1}$. To simplify exposition, for the remaining three accounting ratios, we define g_{it+h} as the untransformed realization for firm i at $t + h$. Moving forward, we refer directly to g_{it+h} as our realization of interest.

1.3 Definitions of Forecast Errors and Revisions

For each firm i and each year t , we denote $F_t^j X_{it+h}$ as the forecast of X_{it+h} made in year t by analyst j that comes from the Detail files. We denote the consensus forecast by $F_t X_{it+h}$, which is defined as the mean forecast across analysts from the Summary files. For both the analyst-level and consensus, we use forecasts extracted three months after the end of fiscal year t , which ensures that the results for fiscal year t have been released by the time these forecasts are made. Similarly, we denote $F_{t-1}^j X_{it+h}$ and $F_{t-1} X_{it+h}$ as the analyst-level and consensus forecasts extracted three months after the end of fiscal year $t - 1$.

For forecasts of the three untransformed accounting ratios, we can work with our raw forecasts directly, so $F_t^j g_{it+h} = F_t^j X_{it+h}$. For the remaining 19 variables, we define forecasts of growth rates as $F_t^j g_{it+h} = \log F_t^j X_{it+h} - \log F_t^j X_{it+h-1}$, with the convention that $F_t^j X_{it} = X_{it}$. For consensus forecasts, variables are constructed analogously without the j superscript. Note that the way we translate from forecasts of levels to growth rates implicitly ignores a Jensen's inequality term because $\log F_t^j X_{it+h} \neq F_t^j \log X_{it+h}$. We ignore this adjustment for simplicity, but we do not think it materially affects our analysis for two reasons. First, a constant Jensen's term would simply shift the unconditional level of forecasts, while our analysis focuses on the conditional properties of forecast errors. Second, we show that our results are robust to using percent (instead of log) growth.

Having defined $F_t^j g_{it+h}$ for each of the 22 variables, we then construct forecast errors and revisions following the literature on expectations formation (Coibion and Gorodnichenko, 2015; Bouchaud et al., 2019):

$$ERR_{it+1}^j = g_{it+1} - F_t^j g_{it+1}, \quad REV_{it}^j = F_t^j g_{it+1} - F_{t-1}^j g_{it+1}.$$

For consensus forecasts, we again use the analogous definitions without the j superscript.

1.4 Summary Statistics

Table 1 shows summary statistics on the realizations, analyst-level forecast errors, and analyst-level forecast revisions for each of our 22 variables. **Table IA.1** shows the same table for consensus-level forecasts. For the sample of each variable in the table, we remove outliers by dropping the top and bottom one percent of each distribution. The number of observations reported in the table is the number of firm-year observations for which we can compute all of the variables of interest. The size of our sample varies a lot across variables, but is in general quite large. Earnings Per Share (EPS), which is studied by most of the literature on analyst forecasts, is one of the most covered variables. However, many other variables have good coverage, such as Net Sales, which is even more covered than EPS with over 407k observations. ROE and Net Income are nearly as covered as EPS. These four variables, as well as EBITDA, EBIT, and Pre-Tax Profit, have more than 200k observations. 13 variables have more than 100k observations, and 19 more than 50k. Two variables, EPS Before Goodwill and Funds From Operations Per Share, have fewer than 10k observations. In the following section, we report analysis results for all variables, ranking them by their number of observations.

In the main text of this paper, we focus on analyst-level forecasts to avoid the possibility that our results are affected by aggregating across forecasters (Bordalo et al., 2020a), but also to be

Table 1: Summary Statistics: Analyst-Level Forecasts

| Variable | Transformation | Realizations | | Errors | | Revisions | | # of Analysts | # of Obs. |
|---------------------------------|----------------|--------------|-----------|-------------|-----------|-------------|-----------|---------------|-----------|
| | | Mean (2) | SD (3) | Mean (4) | SD (5) | Mean (6) | SD (7) | | |
| Net Sales | Growth | 0.06 | 0.16 | -0.00 | 0.10 | -0.02 | 0.13 | 6.33 | 399809 |
| Earnings Per Share | Growth | 0.06 | 0.44 | -0.04 | 0.31 | -0.07 | 0.34 | 6.35 | 360011 |
| Return On Equity | Level | 0.13 | 0.13 | -0.01 | 0.07 | -0.01 | 0.07 | 5.17 | 345082 |
| Net Income | Growth | 0.06 | 0.43 | -0.03 | 0.31 | -0.06 | 0.34 | 6.24 | 325230 |
| EBITDA | Growth | 0.06 | 0.29 | -0.02 | 0.20 | -0.04 | 0.23 | 5.82 | 280111 |
| Pre-tax Profit | Growth | 0.06 | 0.41 | -0.04 | 0.30 | -0.06 | 0.32 | 5.71 | 269883 |
| EBIT | Growth | 0.06 | 0.37 | -0.02 | 0.27 | -0.07 | 0.31 | 6.32 | 231455 |
| GAAP EPS | Growth | 0.06 | 0.55 | -0.08 | 0.40 | -0.03 | 0.41 | 4.85 | 171523 |
| Book Value Per Share | Growth | 0.05 | 0.17 | -0.01 | 0.16 | -0.02 | 0.13 | 3.82 | 170561 |
| Return On Assets | Level | 0.05 | 0.06 | -0.01 | 0.03 | -0.01 | 0.03 | 3.47 | 157063 |
| Gross Margin | Level | 0.41 | 0.21 | -0.00 | 0.04 | -0.01 | 0.03 | 4.56 | 150220 |
| Net Asset Value | Growth | 0.06 | 0.18 | -0.00 | 0.17 | -0.02 | 0.14 | 3.12 | 119885 |
| Cash Flow Per Share | Growth | 0.04 | 0.48 | -0.03 | 0.41 | -0.07 | 0.48 | 3.78 | 112323 |
| Net Debt | Growth | 0.04 | 0.41 | 0.07 | 0.41 | 0.10 | 0.45 | 3.69 | 90126 |
| Enterprise Value | Growth | 0.04 | 0.32 | 0.04 | 0.29 | 0.02 | 0.26 | 2.98 | 88791 |
| Dividend Per Share | Growth | 0.06 | 0.37 | -0.02 | 0.31 | -0.04 | 0.32 | 2.82 | 63840 |
| Capital Expenditures | Growth | 0.04 | 0.41 | -0.06 | 0.45 | 0.10 | 0.52 | 2.89 | 62099 |
| Operating Profit | Growth | 0.06 | 0.40 | -0.05 | 0.28 | -0.05 | 0.30 | 3.57 | 57756 |
| EBITDA Per Share | Growth | 0.05 | 0.31 | -0.02 | 0.21 | -0.05 | 0.25 | 1.44 | 16824 |
| Cash Earnings Per Share | Growth | 0.05 | 0.31 | -0.02 | 0.27 | -0.06 | 0.28 | 1.85 | 11578 |
| Funds From Operations Per Share | Growth | 0.03 | 0.14 | -0.00 | 0.08 | -0.04 | 0.11 | 4.74 | 9279 |
| EPS Before Goodwill | Growth | 0.09 | 0.36 | -0.01 | 0.27 | -0.06 | 0.30 | 2.19 | 2558 |

Notes: This table reports the mean, standard deviation, and number of observations for each one of the 22 variables forecasted by analysts in our sample. The sample includes all observations for which an analyst's forecast error and revision are available for a given variable. Variables are ranked by decreasing number of observations. 19 variables are transformed into log growth as explained in the text; 3 are left untransformed. Column (8) reports the average number of analysts forecasts available per firm-year, and column (9) reports the total number of observations.

consistent with our experimental evidence in Section 4.1. However, after presenting our main results in Section 2.1, we discuss the results from using consensus forecasts, which are presented in the Appendix. Additionally, we also conduct a series of robustness checks that are also presented in the Appendix, including showing that our results do not depend on our particular log growth transformation.

2 Three Facts

In this section, we document the three key facts that motivate the model of expectations formation that we develop in Section 3.

2.1 Fact #1: Non-Linear Relationship Between Forecast Errors and Revisions

Our first and central empirical fact is based on a regression of forecast errors on forecast revisions. This regression was introduced by [Coibion and Gorodnichenko \(2015\)](#) (CG) and takes the following form, for forecaster j , firm i and year t :

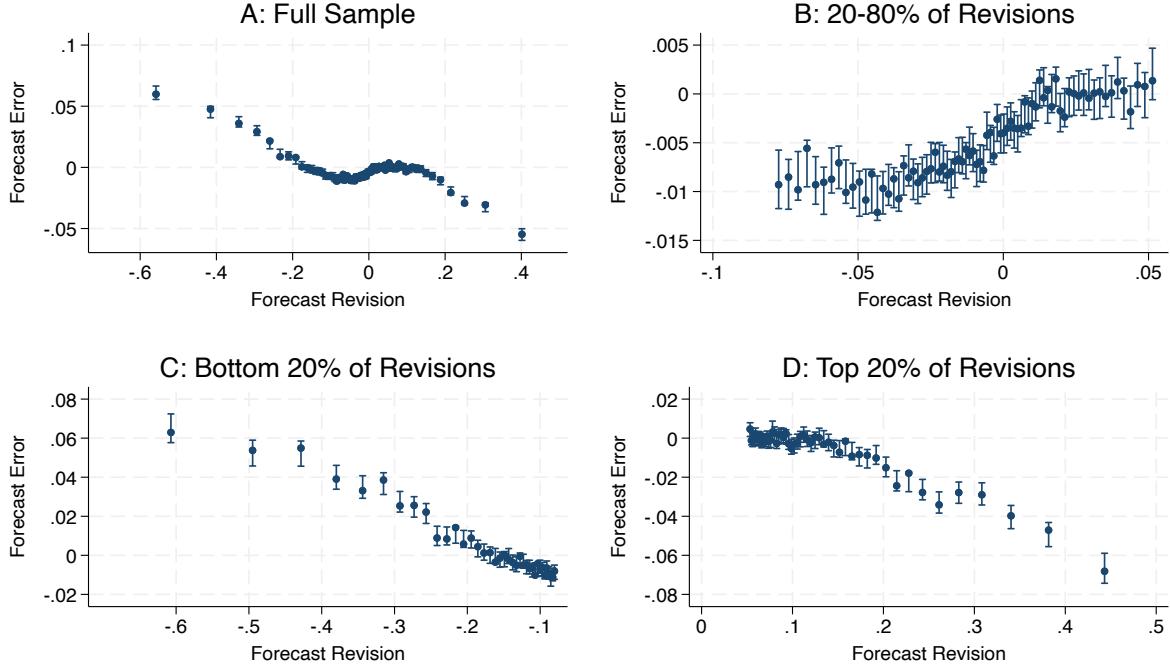
$$ERR_{it+1}^j = \alpha + \beta REV_{it}^j + e_{it+1}^j \quad (1)$$

where the forecast error ERR_{it+1}^j and revision REV_{it}^j for each variable are defined in Section 1. The slope coefficient is typically used to distinguish between different models of expectations formation, requiring only panel data on expectations and realizations. The rational expectations hypothesis predicts $\beta = 0$, since forecast errors cannot be predicted using data in the forecaster's information set. When this regression is run using consensus forecasts, $\beta > 0$ is typically interpreted as evidence of information frictions, as in models of sticky or noisy information ([Coibion and Gorodnichenko, 2015](#)). With individual forecasts, $\beta > 0$ is interpreted as non-Bayesian underreaction ([Bouchaud et al., 2019](#)), while $\beta < 0$ is interpreted as overreaction ([Bordalo et al., 2020a](#)). A key feature of prior literature is that it restricts analysis to linear functional forms, as in equation (1). While this is a natural starting point, especially in settings with small sample sizes, this section uses our large sample of firm-level expectations to provide evidence that this relationship is non-linear.

Graphical evidence. [Figure 1](#) provides a non-parametric look at the relationship between forecast errors and revisions for Net Sales, the best-populated variable with more than 400k observations. The figure shows that the relationship between forecast errors ERR_{it+1}^j and forecast revisions REV_{it}^j is highly non-linear. Panel A shows a binned scatterplot on the full sample with 100 bins; Panels B, C, and D show the same plot separately for observations in the 20-80th, 0-20th, and 80-100th percentiles of revisions. For revisions in the bulk of the distribution, Panel B shows that errors are increasing in revisions, consistent with forecasters underreacting to news that causes moderately-sized revisions in forecasts. This finding is not novel to our paper: it is consistent with the underreaction in analysts' EPS forecasts for US firms ([Bouchaud et al., 2019](#)), as well as managers' revenue forecasts in the US and Italy ([Ma et al., 2024](#)).

Panels C and D of [Figure 1](#) show that for large (positive or negative) revisions, the positive relationship between forecast errors and revisions reverses and becomes negative: large positive (negative) forecast revisions are predictive of negative (positive) future forecast errors. Unlike the relationship in the bulk of the distribution, which is consistent with underreaction, this finding is consistent with *overreaction* to news. In other words, forecasters appear to overreact in response to

Figure 1: Non-Linear Error-Revision Relationship: Analyst-Level Net Sales Forecasts



Notes: This figure shows binned scatterplots of analyst-level forecast errors on forecast revisions for Net Sales. Panel A shows the results for the entire sample; Panel B restricts the sample to the 20-80th percentiles of revisions; Panel C restricts the sample to below the 20th percentile of revisions; Panel D restricts the sample to above the 80th percentile of revisions. Vertical bars represent 95% confidence intervals, assuming the relationship is piecewise linear and continuous (option “ci(1 1)” in Stata command “binsreg”).

news that generates large revisions, while underreacting to more moderate news.

Regression evidence. While the results in Figure 1 are specific to Net Sales, we now show that this same non-linearity is present for almost all of the other variables in Table 1. To do so, we estimate the following variant of the regression in (1) separately for each variable with OLS:

$$ERR_{it+1}^j = \alpha + \beta REV_{it}^j + \gamma REV_{it}^j \times (P_{ijt}^+ + P_{ijt}^-) + \delta^+ P_{ijt}^+ + \delta^- P_{ijt}^- + e_{it+1}^j \quad (2)$$

where P_{ijt}^+ (resp. P_{ijt}^-) is equal to 1 if REV_{it}^j belongs to the top (resp. bottom) 20% of all revisions for this variable. In this specification, β is the error-revision coefficient for revisions between the 20th and 80th percentiles of the distribution of revisions, while γ captures the change in this slope in the tails of the distribution (i.e., below the 20th and above the 80th percentiles). Equation (2) crucially allows intercepts to differ in the two tails, as is the case in Figure 1, but estimates a pooled slope in the tails of the distribution. We choose to estimate a pooled slope to improve

Table 2: Non-Linear Error-Revision Relationship: Analyst-Level Forecasts

| Variable | Error _{it+1} | | Revision _{it} | | Revision _{it} × Top/Bottom 20% of Revision _{it} | | Top 20% of Revision _{it} | | Bottom 20% of Revision _{it} | | # of Obs. |
|---------------------------------|-----------------------|--------|------------------------|--------|---|--------|-----------------------------------|--------|--------------------------------------|-----|-----------|
| | $\hat{\beta}$ | t-stat | $\hat{\gamma}$ | t-stat | $\hat{\delta}^+$ | t-stat | $\hat{\delta}^-$ | t-stat | | | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | | (9) | |
| Net Sales | 0.11 | 6.31 | -0.26 | -14.99 | 0.02 | 5.86 | -0.02 | -5.09 | 399809 | | |
| Earnings Per Share | 0.08 | 2.97 | -0.26 | -9.44 | 0.02 | 1.48 | -0.07 | -5.09 | 360011 | | |
| Return On Equity | 0.23 | 9.75 | -0.27 | -11.70 | 0.00 | 2.56 | -0.02 | -5.99 | 345082 | | |
| Net Income | 0.04 | 1.53 | -0.25 | -11.08 | 0.02 | 1.63 | -0.06 | -4.91 | 325230 | | |
| EBITDA | -0.04 | -1.51 | -0.15 | -5.28 | 0.01 | 1.72 | -0.02 | -2.00 | 280111 | | |
| Pre-tax Profit | 0.01 | 0.47 | -0.22 | -9.05 | 0.03 | 2.38 | -0.05 | -3.07 | 269883 | | |
| EBIT | -0.05 | -1.60 | -0.22 | -8.12 | 0.03 | 2.94 | -0.03 | -2.06 | 231455 | | |
| GAAP EPS | -0.02 | -0.73 | -0.14 | -4.99 | 0.01 | 0.31 | -0.04 | -1.91 | 171523 | | |
| Book Value Per Share | -0.06 | -1.52 | -0.42 | -8.53 | 0.01 | 2.44 | -0.04 | -6.90 | 170561 | | |
| Return On Assets | 0.41 | 10.35 | -0.45 | -10.15 | -0.01 | -5.32 | -0.01 | -8.49 | 157063 | | |
| Gross Margin | 0.09 | 3.90 | -0.15 | -6.23 | 0.00 | 1.70 | -0.01 | -6.62 | 150220 | | |
| Net Asset Value | 0.04 | 0.99 | -0.56 | -11.26 | 0.03 | 7.75 | -0.03 | -4.61 | 119885 | | |
| Cash Flow Per Share | -0.28 | -15.86 | -0.06 | -3.27 | 0.05 | 4.05 | -0.00 | -0.36 | 112323 | | |
| Net Debt | -0.17 | -8.31 | -0.12 | -5.46 | 0.12 | 10.06 | 0.12 | 9.14 | 90126 | | |
| Enterprise Value | -0.20 | -5.82 | -0.26 | -8.25 | 0.11 | 11.91 | -0.01 | -0.82 | 88791 | | |
| Dividend Per Share | -0.12 | -4.54 | -0.17 | -6.15 | -0.01 | -1.13 | -0.03 | -3.08 | 63840 | | |
| Capital Expenditures | -0.36 | -21.93 | -0.16 | -8.53 | 0.06 | 5.20 | -0.05 | -4.32 | 62099 | | |
| Operating Profit | 0.12 | 3.07 | -0.35 | -7.83 | 0.04 | 3.40 | -0.06 | -4.54 | 57756 | | |
| EBITDA Per Share | -0.09 | -2.84 | -0.13 | -3.14 | 0.02 | 2.24 | -0.03 | -1.85 | 16824 | | |
| Cash Earnings Per Share | -0.19 | -4.95 | -0.20 | -3.83 | 0.03 | 1.98 | -0.02 | -1.01 | 11578 | | |
| Funds From Operations Per Share | -0.01 | -0.25 | -0.09 | -1.72 | 0.01 | 1.61 | -0.02 | -2.00 | 9279 | | |
| EPS Before Goodwill | -0.13 | -2.29 | -0.11 | -1.36 | 0.04 | 2.82 | -0.09 | -3.21 | 2558 | | |

Notes: This table reports the OLS estimates of equation (2), for each one of the 22 variables. We report the estimates of β (column 1), γ (column 3), δ^+ (column 5), and δ^- (column 7), along with the corresponding t-statistics with standard errors clustered by firm and time. Variables are ranked by number of observations.

power and because our focus is on the non-linearity of the error-revision relationship rather than any asymmetries.³

Table 2 shows the estimated coefficients in (2) for all of the 22 variables in our sample. Starting with Net Sales in the first line, we find $\hat{\beta} = 0.1$ ($t = 6.3$) and $\hat{\gamma} = -0.26$ ($t = -15$). This is consistent with **Figure 1**: in the bulk of the distribution, errors are positively related to revisions with slope $\hat{\beta} = 0.1$, consistent with underreaction, while in the tails of the distribution, errors are negatively related to revisions with slope $\hat{\beta} + \hat{\gamma} = 0.1 - 0.26 = -0.16$, consistent with overreaction.

³Table IA.2 shows the results of a specification that allows for asymmetric slopes. While the strong non-linearity shown here is present for all variables, some variables exhibit asymmetry.

The remaining rows in [Table 2](#) show that the non-linear relationship between forecast errors and revisions that we find for sales is present for all 22 variables in our data. For 20 out of 22 variables, we find that $\hat{\gamma}$ is negative and significant at the 1% level. $\hat{\beta}$ varies more across variables, with some variables having underreaction in the bulk of the distribution (e.g., Net Sales, Earnings Per Share, Return On Equity), while others have overreaction (e.g., Enterprise Value, Dividend Per Share). Overall, across all variables available in our data, these results show that the error-revision relationship is highly non-linear.

Robustness. We now discuss a battery of robustness checks that we conduct to ensure that the non-linear relationship between forecast errors and revisions is robust. We discuss each one of these robustness checks in a different paragraph below. For each robustness check, we produce an analogous version of [Table 2](#) that is referenced and shown in the Appendix.

First, we check that the relationship continues to hold for consensus forecasts. A potential issue with analyst-level forecasts is that they are sensitive to analyst-level noise, which introduces a negative bias in the estimate of β in (1) ([de Silva and Thesmar, 2024](#)). Therefore, our results could be driven by the fact that expectation noise is larger for bigger revisions, inducing a stronger bias in the tails. However, [Table IA.3](#) shows that this is not the case. In particular, when working with consensus forecasts, we find that $\hat{\gamma}$ is statistically significantly negative for all but 5 variables, with magnitudes that are similar to the analyst-level results.

Second, we show that our results are not sensitive to our log growth transformation. One concern with this transformation is that we ignore a Jensen's inequality term when constructing $F_t^j g_{it+h}$ (the log of a forecast is not necessarily the forecast of the log). To address this concern, we redefine growth as $g_{it+1} = \frac{X_{it+1} - X_{it}}{X_{it}}$ and forecasted growth at horizon h as $F_t^j g_{it+h} = \frac{F_t^j X_{it+h} - F_t^j X_{it+h-1}}{F_t^j X_{it+h-1}}$ (for the 19 variables that are transformed). The benefit of this approach is that our construction of $F_t^j g_{it+1}$ is no longer affected by Jensen's inequality.⁴ [Table IA.4](#) shows that, when using percent growth, $\hat{\gamma}$ is negative for all 19 transformed variables, and significantly negative for 16 variables with similar magnitudes.

Third, we argue later in the paper that the driver of the non-linear error-revision relationship is the presence of fat tails. One concern with this argument is that fat tails may instead reflect a mixture of thin-tailed processes with heterogeneous variances ([Wyart and Bouchaud, 2003](#)). To address this concern, we restrict our sample to firms for which we have at least 10 observations and compute the

⁴We still need to ignore a Jensen's inequality term when computing $F_{t-1}^j g_{it+1}$, which is needed to compute forecast revisions. However, the fact that our results are quantitatively unchanged when using percent growth to remove this effect on $F_t^j g_{it+1}$ makes us less concerned that this is an important issue for our results.

standard deviation of the realization of each variable. We then normalize realizations and forecasts by this firm-specific standard deviation. [Table IA.5](#) shows that, even after adjusting for firm-specific volatility, our results are quantitatively similar with $\hat{\gamma}$ estimated to be significantly negative for all but one variable. In [Table IA.6](#), we perform a similar analysis but instead normalize by the annual cross-sectional standard deviation of each realization. Again, we find quantitatively similar results, suggesting our findings are not driven by time-varying aggregate volatility.

Finally, part of the literature on expectation formation has emphasized that rational learning can occur slowly ([Farmer et al., 2024](#)), which can generate a correlation between errors and revisions in short time series that may not be present out of sample ([Bianchi et al., 2022](#)). Our use of cross-sectional variation alleviates this concern to the extent that our results are driven by firm-specific variation, but not if they are driven by aggregate variation. [Table IA.7](#) shows that our results are indeed driven by the former: they are unchanged after absorbing aggregate shocks by removing the cross-sectional mean of errors and revisions in each year. Additionally, we directly test whether learning drives the non-linear error revision relationship by splitting our sample into four quartiles of analyst experience, measured by the number of firms followed by each analyst up to the current year. [Table IA.8](#) shows that the non-linear pattern is quantitatively similar in all quartiles. This finding suggests that learning is not a first-order driver of our results, which are likely driven by our use of extensive cross-sectional variation.

2.2 Fact #2: Fat Tails in Distribution of Realizations

The strong evidence in [Section 2.1](#) that the relationship between forecast errors and revisions is non-linear points towards a different treatment of large versus small shocks. This evidence is hard to square with established models of expectation formation, which feature linear DGPs (e.g., AR1s) and linear expectations models. Next, we document two additional facts about the realizations of the variables being forecasted that guide our theory by informing our model of the DGP.

Our second fact is that the distributions of realizations have thick tails. It is well known that many financial and economic variables have tails that are fatter than a normal distribution ([Gabaix, 2009](#)). For example, the distribution of firm sizes follows a power law distribution with a tail coefficient of one ([Axtell, 2001](#)), which is known as Zipf's law, and the distribution of growth rates in COMPUSTAT follows a Laplace distribution, which has fatter tails than a normal distribution ([Stanley et al., 1996; Bottazzi and Secchi, 2006](#)). To measure the extent to which our variables are Gaussian, we use two metrics on the set of realizations at the firm-year level. The first is kurtosis, which would be three if the variables are normal. The second is the slope coefficient in a log rank-log size regression, which is a standard way of estimating the tail parameter of a power law

distribution that would be infinity for a normal distribution (Gabaix, 2009). Specifically, for each variable g_{it} , we regress the log rank of $|g_{it}|$ on $\log |g_{it}|$ using observations between the 90th and 99th percentiles of $|g_{it}|$. We then report the negative of the slope coefficient, which corresponds to our estimate of the tail parameters.

[Table 3](#) shows that the realizations of all 22 forecasting variables in our sample have tails that are much fatter than those of a Gaussian distribution. Apart from Return on Assets and Gross Margin, all variables have a tail coefficient that is below three. These imply very fat tails: a tail parameter of two or lower implies that the distribution has an infinite variance. Apart from Capital Expenditures, Enterprise Value, and Gross Margin, all variables have a kurtosis of five or more. These estimates are very far from normal, and formal tests based on kurtosis reject the null of a normal distribution at conventional significance levels.

Robustness. In the Appendix, we show that the presence of fat tails documented in [Table 3](#) is robust to the alternative variable definitions discussed in the robustness checks in Section 2.1. First, [Table IA.9](#) shows that fat tails are still present after dividing by firm-specific estimates of volatility, which suggests that they do not arise from a mixture of thin-tailed processes with heterogeneous variances ([Wyart and Bouchaud, 2003](#)). Second, [Table IA.10](#) shows that adjusting for time-varying aggregate volatility does not change the results. Finally, [Table IA.11](#) shows that tail coefficients remain low when using percent instead of log growth.

2.3 Fact #3: Non-Linear Conditional Expectation

The third and final fact that informs the specification of the data-generating process in our model is that the conditional expectation of the current realization is a non-linear function of the lagged realization. [Figure 2](#) illustrates this point graphically for Net Sales, the best-populated variable. In Panel A, we plot the relationship between Net Sales at $t + 1$ and Net Sales at t using a binned scatter plot. As is evident from the figure, the conditional expectation of Net Sales conditional on its lag exhibits significant non-linearity reminiscent of the non-linearity in the relationship between forecast errors and revisions in [Figure 1](#). Panels B, C, and D of [Figure 2](#) zoom in on the three parts of the distribution of past growth. These panels show that Net Sales is approximately increasing and linear in the bulk of the distribution, while the slope is severely attenuated, or even reversed, in the tails of the distribution. This evidence is consistent with [Jaimovich et al. \(2025\)](#), who find that sales growth is more persistent in the bulk of the distribution than in the tails and exhibits fat tails.

While the evidence in [Figure 2](#) is specific to Net Sales, we now show that this non-linear

Table 3: Fat Tails in the Distribution of Realizations

| Variable | (1) | (2) | Transformation | Kurtosis | Log-Log Regression | | # of Obs. |
|---------------------------------|--------|------|----------------|----------------|--------------------|--------|-----------|
| | | | | Tail Parameter | R ² | (4) | |
| Net Sales | Growth | 5.86 | | 2.67 | 0.99 | 140482 | |
| Earnings Per Share | Growth | 6.41 | | 2.39 | 0.99 | 129735 | |
| Return On Equity | Level | 9.49 | | 2.61 | 1.00 | 127268 | |
| Net Income | Growth | 6.38 | | 2.40 | 0.99 | 113862 | |
| Pre-tax Profit | Growth | 6.25 | | 2.45 | 0.99 | 102386 | |
| EBITDA | Growth | 6.16 | | 2.52 | 0.99 | 99731 | |
| Book Value Per Share | Growth | 8.60 | | 2.09 | 0.99 | 89884 | |
| EBIT | Growth | 6.44 | | 2.42 | 0.99 | 82100 | |
| Return On Assets | Level | 6.46 | | 3.40 | 0.99 | 79479 | |
| GAAP EPS | Growth | 5.91 | | 2.56 | 0.99 | 74210 | |
| Net Asset Value | Growth | 7.92 | | 2.16 | 0.99 | 71494 | |
| Gross Margin | Level | 2.44 | | 8.00 | 0.98 | 64483 | |
| Cash Flow Per Share | Growth | 5.56 | | 2.63 | 0.99 | 62446 | |
| Dividend Per Share | Growth | 7.07 | | 2.45 | 0.99 | 49402 | |
| Net Debt | Growth | 6.57 | | 2.33 | 0.99 | 48338 | |
| Enterprise Value | Growth | 4.58 | | 3.00 | 0.99 | 48140 | |
| Capital Expenditures | Growth | 4.98 | | 2.82 | 0.99 | 44623 | |
| Operating Profit | Growth | 6.23 | | 2.54 | 0.99 | 33033 | |
| EBITDA Per Share | Growth | 6.26 | | 2.41 | 0.99 | 15343 | |
| Cash Earnings Per Share | Growth | 6.04 | | 2.40 | 0.99 | 10013 | |
| Funds From Operations Per Share | Growth | 8.39 | | 2.38 | 0.99 | 3399 | |
| EPS Before Goodwill | Growth | 6.25 | | 2.87 | 0.99 | 2413 | |
| Normal Distribution | . | . | | 3 | ∞ | . | . |

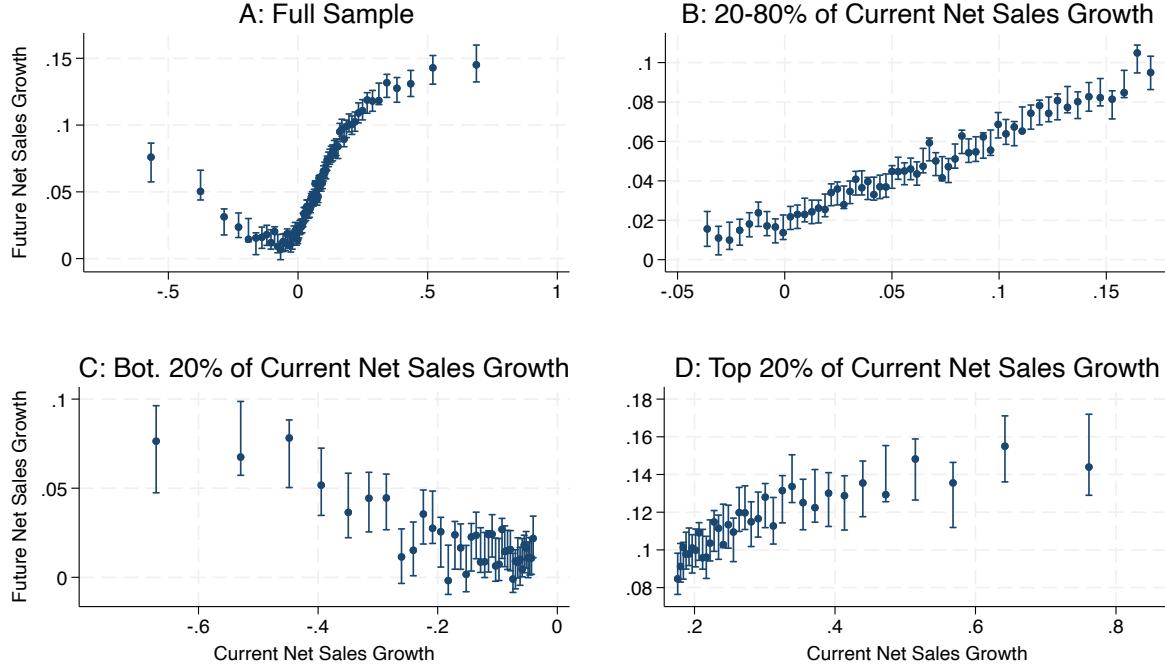
Notes: This table reports the kurtosis and tail parameters of the 22 variables in our data. Tail parameters, reported in column (3), are obtained through the following procedure. For each variable g_{it} , we regress the log rank of $|g_{it}|$ on $\log |g_{it}|$, restricting ourselves to the centiles of $|g_{it}|$ between 90 and 99. We report the R^2 of these regressions in column (4). We restrict ourselves to realizations available in LSEG IBES. Variables are ranked by number of observations in the final column.

relationship holds for all other 22 variables. For each variable, g_{it} , we estimate the following regression with OLS:

$$g_{it+1} = \alpha_g + \beta_g g_{it} + \gamma_g g_{it} \times (M_{it}^+ + M_{it}^-) + \delta_g^+ M_{it}^+ + \delta_g^- M_{it}^- + e_{it+1} \quad (3)$$

where M_{it}^+ (resp. M_{it}^-) is equal to 1 if g_{it} belongs to the top (resp. bottom) 20% of its distribution. This specification is analogous to equation (2), but for realizations instead of forecasts: β_g is the coefficient for the relationship between g_{it+1} and g_{it} in the bulk of the distribution, while γ_g captures the change in this slope in the tails. As in (2), we allow the intercepts in the two tails to differ, but impose that the slopes are equal.

Figure 2: Non-Linear Conditional Expectation of Realizations: Net Sales



Notes: This figure shows binned scatter plots of current Net Sales on past Net Sales. Panel A shows the results for the entire sample; Panel B restricts the sample to the 20-80th percentiles of past Net Sales; Panel C restricts the sample to below the 20th percentile of past Net Sales; Panel D restricts the sample to above the 80th percentile of past Net Sales.

Table 4 shows that $\widehat{\beta}_g > 0$ in 20 out of 22 cases, and is statistically significant in 18 cases. Additionally, we find that $\widehat{\gamma}_g < 0$ in all cases, and it is statistically significant in all but one case. Overall, this evidence shows that the relationship between current and lagged values for our variables is highly non-linear, with persistence in the bulk and more mean-reversion in the tails of the distribution.

Robustness. In the Appendix, we show that this non-linear conditional expectation is robust to the alternative variable definitions discussed in the robustness checks in Section 2.1: **Table IA.12** adjusts by firm-level volatility; **Table IA.13** adjusts by aggregate volatility; and **Table IA.14** uses percent instead of log growth. In all cases, the non-linearity is strong and pervasive.

3 Model

In this section, we develop a parsimonious model that ties the non-linear relationship between forecast errors and revisions, our Fact #1, to Facts #2 and #3, which are about the data-generating

Table 4: Non-Linear Conditional Expectation of Realizations

| Variable | g_{it+1} | | g_{it} | | $g_{it} \times$ Top/Bottom 20% of g_{it} | | Top 20% of g_{it} | | Bottom 20% of g_{it} | | # of Obs. (9) |
|---------------------------------|----------------------------|---------------|-----------------------------|---------------|--|---------------|-------------------------------|---------------|---------------------------|--|------------------|
| | $\widehat{\beta}_g$ (1) | t-stat (2) | $\widehat{\gamma}_g$ (3) | t-stat (4) | $\widehat{\delta}_g^+$ (5) | t-stat (6) | $\widehat{\delta}_g^-$ (7) | t-stat (8) | | | |
| | | | | | | | | | | | |
| Net Sales | 0.43 | 12.40 | -0.44 | -8.12 | 0.10 | 6.28 | 0.00 | 0.23 | 140482 | | |
| Earnings Per Share | 0.20 | 6.18 | -0.49 | -13.11 | 0.23 | 9.63 | -0.03 | -1.19 | 129735 | | |
| Return On Equity | 0.86 | 68.21 | -0.28 | -16.55 | 0.08 | 14.23 | 0.01 | 3.44 | 127268 | | |
| Net Income | 0.18 | 5.60 | -0.47 | -11.48 | 0.24 | 9.39 | -0.02 | -0.66 | 113862 | | |
| Pre-tax Profit | 0.22 | 5.92 | -0.48 | -11.61 | 0.23 | 8.96 | 0.00 | 0.03 | 102386 | | |
| EBITDA | 0.23 | 7.03 | -0.40 | -10.66 | 0.15 | 8.65 | -0.00 | -0.15 | 99731 | | |
| Book Value Per Share | 0.47 | 13.87 | -0.53 | -13.11 | 0.08 | 10.37 | -0.03 | -6.04 | 89884 | | |
| EBIT | 0.25 | 6.48 | -0.46 | -10.67 | 0.18 | 6.85 | -0.01 | -0.55 | 82100 | | |
| Return On Assets | 0.90 | 89.07 | -0.28 | -18.41 | 0.04 | 13.39 | 0.00 | 6.29 | 79479 | | |
| GAAP EPS | 0.01 | 0.28 | -0.36 | -7.56 | 0.22 | 8.84 | 0.01 | 0.44 | 74210 | | |
| Net Asset Value | 0.46 | 12.83 | -0.44 | -11.91 | 0.08 | 13.88 | -0.03 | -3.76 | 71494 | | |
| Gross Margin | 0.96 | 209.91 | -0.05 | -4.53 | 0.03 | 4.77 | 0.01 | 4.58 | 64483 | | |
| Cash Flow Per Share | -0.18 | -5.43 | -0.14 | -4.80 | 0.14 | 9.39 | 0.05 | 2.84 | 62446 | | |
| Dividend Per Share | 0.44 | 7.82 | -0.66 | -9.91 | 0.17 | 8.37 | -0.03 | -1.67 | 49402 | | |
| Net Debt | 0.09 | 4.35 | -0.27 | -8.79 | 0.16 | 9.73 | 0.00 | 0.13 | 48338 | | |
| Enterprise Value | -0.05 | -0.98 | -0.10 | -1.69 | 0.06 | 2.34 | -0.01 | -0.52 | 48140 | | |
| Capital Expenditures | 0.05 | 1.49 | -0.22 | -5.51 | 0.07 | 4.21 | -0.06 | -3.45 | 44623 | | |
| Operating Profit | 0.24 | 4.56 | -0.51 | -8.79 | 0.25 | 9.60 | -0.01 | -0.69 | 33033 | | |
| EBITDA Per Share | 0.13 | 3.28 | -0.27 | -4.76 | 0.08 | 3.12 | 0.00 | 0.05 | 15343 | | |
| Cash Earnings Per Share | 0.12 | 2.83 | -0.36 | -5.06 | 0.10 | 3.64 | -0.00 | -0.19 | 10013 | | |
| Funds From Operations Per Share | 0.37 | 3.48 | -0.35 | -2.78 | 0.04 | 2.46 | -0.02 | -1.54 | 3399 | | |
| EPS Before Goodwill | 0.13 | 2.02 | -0.34 | -7.87 | 0.13 | 4.13 | -0.03 | -1.50 | 2413 | | |

Notes: This table reports the OLS estimates of equation (3), for each one of the 22 variables. We report the estimates of β_g (column 1), γ_g (column 3), δ_g^+ (column 5), and δ_g^- (column 7), along with the corresponding t-statistics with standard errors clustered by firm and time. Variables are ranked by number of observations. The sample includes one observation per firm-year.

process. We start by describing a model of the DGP that is consistent with our latter two facts, and then turn to a model of belief formation that, given this model of the DGP, generates the first fact. We then show the model can quantitatively replicate our facts, and a potential microfoundation based on [Augenblick et al. \(2024\)](#).

3.1 Data-Generating Process

The first piece of the model is the data-generating process (DGP) for g_t , our forecasting variable of interest. Our model does not have any firm-specific heterogeneity, and we therefore omit the firm

index i for brevity.⁵ Without loss of generality, we normalize the unconditional mean of g_t to zero. We then assume that the DGP for g_t takes the following form

$$g_{t+1} = g_{t+1}^* + \sigma_\epsilon \epsilon_{t+1}, \quad \epsilon_t \sim f_\epsilon(\cdot), \quad (4)$$

$$g_{t+1}^* = \rho g_t^* + \sigma_u u_{t+1}, \quad u_t \sim f_u(\cdot), \quad (5)$$

where ϵ_t and u_t are IID shocks with a zero mean and unit variance. The DGP for g_t therefore consists of two components: (i) a persistent component, g_t^* , that follows an AR1 process with shocks, u_t that have PDF $f_u(\cdot)$, and (ii) a transitory shock, ϵ_t , with PDF $f_\epsilon(\cdot)$. Throughout, we assume that g_t^* is an *unobservable* latent state. We denote its unconditional variance by $\sigma_{g^*}^2 = \frac{\sigma_u^2}{1-\rho^2}$, and denote its PDF by $\Phi(\cdot)$. We also denote the marginal PDF of g_t by $h(\cdot)$, which is given by:

$$h(g) = \int_{-\infty}^{+\infty} \Phi(g - \sigma_\epsilon \epsilon) f_\epsilon(\epsilon) d\epsilon.$$

and denote $\sigma_g^2 = \sigma_\epsilon^2 + \sigma_{g^*}^2$ its unconditional variance.

At this point, if we assumed that ϵ_t and u_t were Gaussian, our model of the DGP would be identical to the models of dividend growth in [Bansal and Yaron \(2004\)](#) and [Lettau and Wachter \(2007\)](#). However, unlike these models and existing literature on belief formation, we instead assume the PDF of ϵ_t , $f_\epsilon(\cdot)$, has heavy tails in the sense that it is well approximated by a power law with tail parameter ν for large values of ϵ_t . Formally, we make the following assumption, where we use the notation $a(x) \sim b(x)$ to denote asymptotic equivalence in the sense that $\lim_{x \rightarrow \infty} \frac{a(x)}{b(x)} = 1$.⁶

Assumption 1. *The distribution of ϵ_t is symmetric and follows a power law:*

$$P(\epsilon_t > x) = \int_x^\infty f_\epsilon(\epsilon) d\epsilon \sim \Theta x^{-\nu} \text{ as } \epsilon \rightarrow \infty$$

where $\nu > 2$ is the tail index and Θ is a constant of proportionality.

An example of a case in which Assumption 1 holds would be if $f_\epsilon(\cdot)$ is the PDF of a t -distribution, where ν would be the degrees of freedom of the distribution. Finally, we need to make an assumption about the distribution of u_t . In all of our theoretical results, we assume that u_t has “thinner” tails than ϵ_t , which is formalized in the following assumption.

⁵The lack of firm-specific heterogeneity in the model is consistent with the fact that our results are unchanged when standardizing at the firm-level.

⁶We prove all of our results in Appendix B under a weaker assumption that ϵ_t is a regularly varying random variable such that $P(\epsilon_t > x) \sim x^{-\nu} L(x)$, where $L(x)$ is a slowly varying function.

Assumption 2. *The distribution of u_t is symmetric and exhibits tail behavior that is asymptotically negligible compared to ϵ_t :*

$$P(u_t > x) = \int_x^\infty f_u(u) du = o(x^{-\nu}).$$

For our theoretical results that characterize the properties of beliefs, Assumptions 1 and 2 are sufficient. However, to characterize the properties of the data-generating process for g_t in closed form, we will make a stronger assumption than Assumption 2 that u_t is normally distributed.

3.2 Replicating Facts #2 and #3: Fat Tails and Non-Linear $E(g_{t+1}|g_t)$

Before turning to our model of beliefs, we show that our model of the DGP can generate our second and third facts. Our second fact makes clear that we need a model of the DGP with fat tails. The fact that our model generates this follows from the properties of power laws: the tail parameter of a sum of independent random variables is the minimum of the tail parameters (Jessen and Mikosch, 2006). Given that g_t^* is the sum of random variables that have thinner tails than ϵ_t and g_t is the sum of g_t^* and ϵ_t , the tail parameter of g_t will be the same as that of ϵ_t , which we denoted by ν .

Showing our model replicates our third fact—the non-linearity in $E(g_{t+1}|g_t)$ —is more difficult because there is no closed-form expression for $E(g_{t+1}|g_t)$. Nevertheless, the following proposition leverages a result from empirical Bayes theory known as Tweedie’s formula (Efron, 2012) to characterize it.

Proposition 1 (Tweedie’s formula). *If u_t is normally distributed, the expectation of future growth conditional on current growth takes the following form:*

$$E(g_{t+1}|g_t) = -\rho\sigma_{g^*}^2 \frac{d}{dg} \log h(g_t). \quad (6)$$

In particular, it is a function of the **observable** distribution of growth, g_t .

Proof. See Appendix B.1. □

Proposition 1 shows that $E(g_{t+1}|g_t)$ is a function of $h(\cdot)$, the marginal PDF of g_t . This result is useful because $h(\cdot)$ is observable, which means we can characterize the shape of $E(g_{t+1}|g_t)$ even without a closed form. For example, suppose that this distribution is approximately Gaussian

in the bulk of the distribution: $\log h(g) \approx -\frac{g^2}{2\Sigma_g} + \text{constant}$, which [Table IA.15](#) shows is a good approximation, especially for Net Sales. Given this approximation, [Proposition 1](#) implies

$$E(g_{t+1}|g_t) \approx \rho \frac{\sigma_{g^*}^2}{\Sigma_g} g_t, \quad (7)$$

which is linear in g_t . This is consistent with our third fact: in the bulk of the distribution, the expectation of g_{t+1} is linearly increasing in g_t . Intuitively, in the bulk of the distribution, g_t is dominated by g_t^* and g_t^* is persistent, implying that g_t is a good approximation of g_{t+1} . Note that if g_t were globally Gaussian, then this equation would hold with equality, and we would recover the well-known result that the rational expectation given a signal of a state with additive Gaussian noise is linear. The benefit of [Proposition 1](#) is that it allows us to extend this result to distributions that are locally Gaussian.

In the tails of the distribution, [Table 3](#) shows that the growth distribution is well approximated by a power law, consistent with [Assumption 1](#). Under [Assumption 1](#), we can use the properties of regularly varying functions to derive the following result that characterizes $E(g_{t+1}|g_t)$.

Proposition 2. *If u_t is normally distributed, [Assumption 1](#) implies that:*

$$E(g_{t+1}|g_t) \sim \frac{\rho\sigma_{g^*}^2(\nu + 1)}{g_t} \text{ as } |g_t| \rightarrow \infty. \quad (8)$$

Proof. See [Appendix B.2](#). □

[Proposition 2](#) shows that, in the tails of the growth distribution, the expectation of g_{t+1} is *decreasing* in g_t , unlike in the bulk of the distribution. The intuition is that extreme values of g_t are likely driven by ϵ_t . However, because ϵ_t is not persistent, this value of g_t is unlikely to persist at $t+1$. In the limit of $|g_t| \rightarrow \infty$, g_t reflects ϵ_t with probability one, and $E(g_{t+1}|g_t)$ converges to the unconditional expectation of zero.⁷ While it is clear that our model of the DGP needs to have fat tails, [Proposition 2](#) also explains why an alternative model in which the distribution of permanent shocks, u_t , has fatter tails than that of transitory shocks, ϵ_t , would be inconsistent with our third fact.⁸ With fatter-tailed permanent shocks, large shocks would be more likely to persist, and the conditional expectation of g_{t+1} given g_t would be *increasing* in g_t in the tails of the distribution.

⁷A related result is developed in [Chambers and Healy \(2011\)](#), who show that when a distribution has sufficiently fat tails such that MLRP does not hold, a positive signal may generate a negative update about the underlying parameter.

⁸Our specification of the DGP is reminiscent of the recent literature on income dynamics ([Guvenen et al., 2014, 2021](#)), which models income processes as the sum of persistent and transitory components. However, this literature emphasizes the importance of non-normal persistent shocks, while the key ingredient in our model is non-normal *transitory* shocks. Another difference is that our focus is primarily on generating the kurtosis in the data rather than skewness. The latter is a pervasive feature of income data due to extreme negative events like job loss.

3.3 Expectations Formation

Given our model of the DGP, we now describe our model of subjective expectation formation. We assume that expectations do not have “full information,” in the sense that forecasters only observe realizations of g_t but not g_t^* . However, this assumption alone cannot explain our first fact, given it would imply that forecast errors should not be predictable by revisions, which are in forecasters’ information set. Therefore, we also assume that forecasts are not rational given this information set. In particular, our core assumption is that forecasters incorrectly perceive the distributions of ϵ_t and u_t to be Gaussian, inconsistent with Assumptions 1 and 2. We do not microfound this misperception, but we view it as consistent with the idea that economic agents use simplified, or “sparse”, models of reality to form their beliefs (Fuster et al., 2010; Gabaix, 2019). Because of this model misspecification, forecast errors will be (conditionally) predictable.

Since only past values of g_t are observable, agents have to solve a filtering problem to compute their expectations about g_t^* given g_0, \dots, g_t , which they in turn use to forecast future growth. Under the assumption that agents perceive ϵ_t and u_t as Gaussian, the solution to this filtering problem implies that their expectations will be characterized by the Kalman filter. Given the DGP, our assumption on belief formation is equivalent to assuming that forecasters use the *linear* forecast with the lowest possible mean squared error. For our theoretical results, we assume that agents are in a steady state in the sense that the posterior variance of the Kalman filter and, hence, the Kalman gain is constant. Denoting this steady state Kalman gain as K , agents’ expectations at horizon k , $F_t g_{t+k}$, are characterized by:

$$F_t g_{t+k} = \rho^k K \sum_{s \geq 0} \rho^s (1 - K)^s g_{t-s}. \quad (9)$$

Our assumption that forecasters use exactly the Kalman filter is not crucial for our results. For example, the results in the next section about the non-linear error-revision relationship would be present if forecasters linearly extrapolated current observations. What is important is that forecasters use a forecasting rule that does not fully capture the presence of fat tails. Aside from its tractability, we choose to use the Kalman filter specifically because it would correspond to rational expectations under the assumption that ϵ_t and u_t were normal. This clarifies that our analytical results are entirely about ignoring the presence of fat tails, rather than an arbitrary assumption about the history or functional form that forecasts take.

3.4 Replicating Fact #1: Non-Linear Error-Revision Relationship

We now show that our model of belief formation, combined with our model of the DGP that replicates Facts #2 and #3, also generates Fact #1.

Proposition 3. *Define forecast errors and forecast revisions as follows:*

$$\begin{aligned} ERR_{t+1} &= g_{t+1} - F_t g_{t+1}, \\ REV_t &= F_t g_{t+1} - F_{t-1} g_{t+1}. \end{aligned}$$

Then, under Assumptions 1 and 2, forecast errors are asymptotically linear in forecast revisions:

$$E(ERR_{t+1} | REV_t) \sim C \times REV_t \quad \text{as } |REV_t| \rightarrow \infty, \quad (10)$$

where $C < 0$.

Proof. See Appendix B.3. □

Proposition 3 shows that in the tails of the distribution of revisions, the relationship between forecast errors and revisions is linear and negative. This is consistent with our first fact in Figure 1: there is overreaction in the tails of the distribution of revisions. The proof of this result relies on showing that revisions are driven by changes in g_t , which are asymptotically large (in absolute value) for one of two reasons: (i) a large realization of ϵ_t or (ii) a large realization of past ϵ_{t-h} . The intuition for why these large revisions reflect overreaction is that the forecaster does not realize that they are less likely to be persistent because they are driven by the fat-tailed component of the process, ϵ_t , which is transitory. If ϵ_t and u_t were Gaussian, there would still be transitory shocks, but the change in g_t would not be informative about the relative size of transitory and persistent shocks. In contrast, when ϵ_t has fatter tails than u_t , large values of $|g_t|$ asymptotically only reflect ϵ_t .

The definition of the Kalman filter implies that it is the minimum linear mean squared error predictor. Combining this with Proposition 3 implies the following result.

Corollary 1. *Under Assumptions 1 and 2, there exists an $\bar{R} > 0$ such that:*

$$\begin{aligned} E\left(ERR_{t+1} \times REV_t \mid |REV_t| > \bar{R}\right) &< 0 \\ E\left(ERR_{t+1} \times REV_t \mid |REV_t| < \bar{R}\right) &> 0 \end{aligned}$$

Proof. See Appendix B.4. □

Corollary 1 shows that our model can generate the non-linear error-revision relationship in Figure 1. While Proposition 3 shows there is overreaction in the tails, Corollary 1 shows that errors and revisions are positively correlated *on average* in the bulk of the distribution of revisions. The intuition for why forecasters underreact in the bulk is similar to the intuition for overreaction in the tails: the forecasters do not realize that intermediate values of g_t , which generate smaller revisions, are more likely to reflect u_t than ϵ_t , and hence are more likely to be persistent. Note, however, that while Corollary 1 shows that errors and revisions are positively correlated *on average* in the bulk where $|REV_t| < \bar{R}$, we cannot show that this is true for all $|REV_t| < \bar{R}$ without further assumptions on $f_\epsilon(\cdot)$ and $f_u(\cdot)$. Additionally, we cannot characterize the exact value of \bar{R} at which the switch between underreaction and overreaction occurs.

3.5 Quantitative Fit

We now assess whether our model can quantitatively account for our three main empirical facts. Because of computational constraints, we focus on the one variable that has the largest sample, Net Sales.

Simulation procedure. To simulate from our model, we make three additional assumptions. First, we make the stronger version of Assumption 2 that u_t is normally distributed with unit variance, which is motivated by the evidence in Table IA.15. Second, we assume that ϵ_t is distributed according to a t -distribution with $\nu > 2$ degrees of freedom normalized to have a unit variance. The t -distribution is asymptotically a power law with tail parameter ν , and has the nice property of converging to a normal distribution as $\nu \rightarrow \infty$. Finally, we relax the assumption that the Kalman filter updating equations are applied using a constant Kalman gain, which would only apply in a steady state. Instead, we use the following updating equations to compute subjective forecasts, which follow from applying standard Kalman filter results to equations (4) and (5) under the (incorrect) assumption that $\epsilon_t \sim N(0, 1)$:

$$\begin{aligned} F_t g_{t+h} &= \rho^h F_t g_t^*, \\ F_t g_t^* &= (1 - K_t) F_{t-1} g_t^* + K_t g_t, \quad F_0 g_0^* = g_0^*, \\ K_t &= \frac{\Sigma_t}{\Sigma_t + \sigma_\epsilon^2}, \\ \Sigma_{t+1} &= \rho^2 (1 - K_t) \Sigma_t + \sigma_u^2, \quad \Sigma_0 = 0. \end{aligned} \tag{11}$$

In our simulations, we sample time series of g_t according to equations (4) and (5) with 100,000 observations. We repeat this simulation 100 times, where g_0^* is drawn from its stationary distribution;

the length of the simulation burn-in period is 50 observations for each series. This gives us a total of 10 million simulated observations.⁹

Fit on Facts #2 and #3. We start by estimating the four parameters of the DGP for g_t in (4) and (5): ρ , the persistence of g_t^* , σ_ϵ , the scale parameter for ϵ_t , ν , the tail parameter of ϵ , and the innovation volatility, σ_u . We estimate these parameters using Simulated Minimum Distance (SMD), minimizing the difference between a set of statistics computed from simulated data in the model and the corresponding value of those statistics in the data for Net Sales with the inverse-diagonal of the sample covariance matrix as the weighting matrix.¹⁰ We use the following six statistics: our estimates of β_g and γ_g from Table 4; the tail parameter from Table 3; the 10-90 percentile difference of g_t , $g_{t+1} - g_t$, and $g_{t+3} - g_t$. The first two statistics, β_g and γ_g , capture Fact #3: the non-linearity in the conditional expectation of g_{t+1} given g_t . These regression coefficients jointly identify ρ , σ_u , and ν , but they do not separately identify each of these parameters. The tail parameter that captures Fact #2 identifies ν separately. The dispersion of g_t helps identify the scale of the DGP, in particular σ_u and σ_ϵ . Finally, the dispersion of one- and three-year changes in g_t is useful for separately identifying ρ and σ_u because the extent to which the three-year changes tend to be larger than one-year changes is affected by ρ much more than other parameters. Our choice of one- and three-year changes, specifically, follows the literature on income dynamics, which uses the same moments to identify persistence and volatility in similar data-generating processes (e.g., Guvenen et al., 2014, 2021).

Our SMD estimates for the DGP parameters are shown in the first four columns of Table 5. Consistent with Fact #2, we estimate $\nu = 2.411$. Combined with our estimate of a relatively persistent process for g_t^* , $\rho = 0.784$, this generates Fact #3. The first six rows of Table 6 show the fit of the model on the targeted statistics at these estimated parameters. The results show that the model is able to fit the data very well, replicating all of the targeted statistics almost exactly. In Section 4.3, we will examine the out-of-sample prediction that this estimated model makes for return predictability.

Fit on Fact #1. We now turn to the model's fit of Fact #1. Assuming expectations are formed according to the Kalman filter in (11), the final column panel of Table 6 shows the model's prediction

⁹We choose this simulation size to be as large as possible without exceeding the RAM of our GPU. The simulation itself is not memory or computationally intensive—these constraints only become binding when we compute the rational expectation using a particle filter described below.

¹⁰We perform this optimization using a two-step procedure in which we first search on a quasi-random grid of 15,000 points, and then run local Nelder-Mead optimizations using the top 5 points as starting points. Our result is then the parameter vector that has the lowest objective function from any of these Nelder-Mead optimizations.

Table 5: Simulated Minimum Distance Parameter Estimates

| | ρ | σ_u | σ_ϵ | ν | λ |
|------------|--------|------------|-------------------|-------|-----------|
| Estimate | 0.784 | 0.336 | 1.058 | 2.411 | 0.197 |
| Std. Error | 0.020 | 0.018 | 0.046 | 0.037 | 0.024 |

Notes: This table shows the parameter estimates and standard errors from the two-step Simulated Minimum Distance estimation described in the main text. Standard errors are computed using the covariance matrix of the estimation moments and the gradient of the moment vector computed using central differencing with a step size equal to 1% of the estimated parameter value. We adjust our standard error for λ to account for the first step of the estimation using the procedure in [Murphy and Topel \(1985\)](#).

Table 6: Fit of Model on Estimation Targets

| Moment | Data | Model | Model ($\lambda = 1$) |
|-------------------------|--------|--------|-------------------------|
| $\hat{\beta}_g$ | 0.429 | 0.425 | . |
| $\hat{\gamma}_g$ | -0.438 | -0.437 | . |
| Tail Parameter | 2.670 | 2.670 | . |
| P90-P10 g_t | 2.146 | 2.147 | . |
| P90-P10 $g_{t+1} - g_t$ | 2.595 | 2.594 | . |
| P90-P10 $g_{t+3} - g_t$ | 2.844 | 2.842 | . |
| $\hat{\beta}$ | 0.106 | 0.144 | 1.816 |
| $\hat{\gamma}$ | -0.265 | -0.255 | -2.861 |

Notes: This table shows the fit of the estimated model on the statistics used to estimate it, where the top half are the statistics used to estimate the four DGP parameters and the bottom half are the two statistics used to estimate λ . The first column shows their values in the data. The second column shows their value in the estimated model computed from simulations at the estimated parameter values shown in [Table 5](#). The final column shows values at the same model parameters, changing only λ to be set to 1. The set of statistics that correspond to the data-generating process are not reported because they do not depend on λ .

β and γ from [Table 2](#). Comparing this with the data in the first column, we see that the model qualitatively generates the non-linear error-revision relationship in the data, with underreaction in the bulk of the distribution and overreaction in the tails, consistent with our theoretical results. However, the degree of this non-linearity is an order of magnitude too large. The fact that our model cannot fit the data is not surprising given that our model of beliefs has no additional degrees of freedom, given our estimates of the DGP parameters.

To attenuate the bias of our linear forecasting model, we follow [Fuster et al. \(2010\)](#) and [Gabaix \(2019\)](#) and assume that expectations are formed by shrinking away from the rational expectation towards a default model, which we consider as the Kalman filter:

$$F_t^\lambda g_{t+h} = \lambda F_t g_{t+h} + (1 - \lambda) E_t g_{t+h} \quad (12)$$

where E_t is the rational expectation given the full history of g_t and λ is the weight that is placed on the Kalman filter forecast. $\lambda = 1$ corresponds to the case in the final column of [Table 6](#), while

$\lambda = 0$ corresponds to the case of rational expectations in which error-revision coefficients would always be zero. We then estimate λ to assess whether the anchored model in (12) can match our first fact quantitatively. Doing so requires computing the rational expectation, $E_t g_{t+1}$, which we do using the particle filter from [Fernandez-Villaverde and Rubio-Ramirez \(2007\)](#) as described in Appendix C. Having computed the rational expectation, we estimate λ using SMD with β and γ as statistics and the inverse-diagonal of the covariance matrix as a weighting matrix.¹¹

Our estimated value of λ , shown in the last column of [Table 5](#), implies that forecasters place 20% of their weight on the Kalman filter forecast, and 80% of their weight on the true forecast. The second column of [Table 6](#) shows that with $\lambda = 0.197$, the model is able to quantitatively replicate the non-linear error-revision relationship in the data.

3.6 A Microfoundation Based on Insensitivity to Signal Strength

In this section, we show that our assumption that forecasters ignore fat tails and use the Kalman filter is analogous to assuming that forecasters are insensitive to the strength of the signals that they observe each period, as in the theory of [Augenblick et al. \(2024\)](#) (ALT).

In order to make this point, we first note that rational expectations, which cannot be computed in closed form, are well-approximated by the following state-dependent Kalman filter:¹²

$$\begin{aligned} F_t^{SD} g_{t+h} &= \rho^h F_t^{SD} g_t^*, \\ F_t^{SD} g_t^* &= F_{t-1} g_t + \tilde{K}(v_t) v_t, \quad v_t = g_t - F_{t-1}^{SD} g_t \\ \tilde{K}(v_t) &= K \frac{\gamma_0}{1 + \exp(|v_t| - \gamma_1)}, \end{aligned} \tag{13}$$

where K is the steady state Kalman gain and (γ_0, γ_1) are parameters. (13) corresponds to a state-dependent version of the steady state Kalman filter, where $\tilde{K}(\cdot)$ tilts away from the steady state Kalman gain based on the time t forecast error, v_t .¹³ We estimate $\gamma_0 = 4.55$ and $\gamma_1 = 0.89$ by minimizing the mean squared difference between the one-period-ahead forecasts in (13) and the rational expectation in our full sample of simulated data. We find that this state-dependent filter provides a very close approximation to the particle filter with an R^2 of 99.4% and a slope coefficient that is nearly one.

¹¹In principle, we could have estimated λ jointly with the DGP parameters. However, this is too computationally-intensive because it requires rerunning the particle filter for each different set of DGP parameters. Nevertheless, we adjust our standard error for λ to account for the two-step estimation procedure following [Murphy and Topel \(1985\)](#).

¹²We thank Stefan Nagel for suggesting the idea of using a state-dependent Kalman gain.

¹³We tilt away from the steady state Kalman gain because convergence occurs after only three periods.

Given the quasi-rational expectation in (13), the theory in ALT provides a microfoundation for forecasters ignoring fat tails. To see this, observe that our assumption that forecasters use a constant Kalman gain implies that forecasters place too much weight on v_t when it is large and too little weight when it is small. For extremely large forecast errors, the quasi-rational weight on recent information should be $\lim_{v \rightarrow \infty} \tilde{K}(v) = 0$, while the constant Kalman gain uses a positive weight. For small errors, the quasi-rational weight on v_t is $\tilde{K}(0) = K \frac{4.55}{1 + \exp(-0.89)}$, which is larger than the constant Kalman gain. Therefore, using a constant Kalman gain, as we assume in our model, is analogous to not recognizing that different signals have different strengths and giving too much weight to large uninformative shocks but too little to small informative ones. As a result, forecasters overreact to imprecise signals (larger errors) and underreact to precise ones (smaller errors), just like in ALT.

An alternative way to see the connection between underreaction versus overreaction and ALT is the following result.

Proposition 4. *If u_t is normally distributed, Assumption 1 implies that:*

$$\text{var}(g_{t+1}|g_t) = \sigma_g^2 + \rho^2 \sigma_{g^*}^4 \frac{d^2}{dg^2} \log h(g_t).$$

Proof. See Appendix B.5. □

Proposition 4 shows that the conditional variance of g_{t+1} given g_t depends on the second derivative of the log-density of g_t . Given our model of the DGP, $\log h(\cdot)$ is concave in the bulk and convex in the tails. Using Proposition 4, these patterns imply that g_t is a more precise signal of g_{t+1} for smaller values of past growth, while it is less precise in the tails of the distribution. Therefore, our model delivers a similar result to ALT: forecasters are underreacting in the parts of the distribution where signals are strong, and overreacting when signals are weak.¹⁴

4 Additional Evidence from an Experiment, Macro Forecasts, and Stock Returns

This section presents additional evidence in support of our model from three different domains. The first comes from an online forecasting experiment, the second comes from macro forecasts, and

¹⁴This result is also consistent with the model in Ba et al. (2024) (BBI), which (under parameter restrictions) makes the same prediction for the relationship between underreaction and overreaction and signal strength as ALT. While both models in ALT and BBI make this same prediction, the mechanisms are quite different. This leads them to make other different predictions, such as how overreaction responds to complexity. See Appendix C of BBI for additional details.

the third comes from data on stock returns.

4.1 Evidence from a Forecasting Experiment

Experimental design. The design of our online forecasting experiment is taken from Afrouzi et al. (2023) (AKLMT). Participants are asked to predict the outcome of a process, and their compensation depends on the accuracy of their forecasts. We recruit participants on Amazon MTurk, and they are reasonably representative of the general population. The experiment does not require participants to have any prior knowledge of statistics, and participants do not know the DGP (AKLMT show that this is not important). The interface is graphical and user-friendly: participants click with their mice to provide their forecasts at one- and two-period-ahead horizons. Prior to making their first forecast, they see 40 prior realizations of the process, and then sequentially provide both forecasts in 40 periods, seeing the realization of the process between each period. We refer the reader to AKLMT for further details about this design; see Figure IA.1 for an example of the interface our participants see for these parameters.

Starting with this design, we set the DGP of the process being forecasted to be:

$$g_{t+1} = g_{t+1}^* + 12.16\epsilon_{t+1}, \quad \epsilon_t \sim t(2.533) \quad (14)$$

$$g_{t+1}^* = 0.529g_t^* + 12.62u_{t+1}, \quad u_t \sim N(0, 1) \quad (15)$$

This is the DGP in our theoretical model (equations (4)-(5)) with parameters selected to approximate our data on Net Sales.¹⁵ We set σ_u and σ_ϵ , which simply scale the DGP, to be larger numbers so that values are easier for participants to interpret. As shown in Figure IA.1, the tail of our DGP is thick enough that participants typically observe a few large movements in the DGP in the past 40 realizations they are initially given. We are not in a situation where large shocks of ϵ rarely happen.

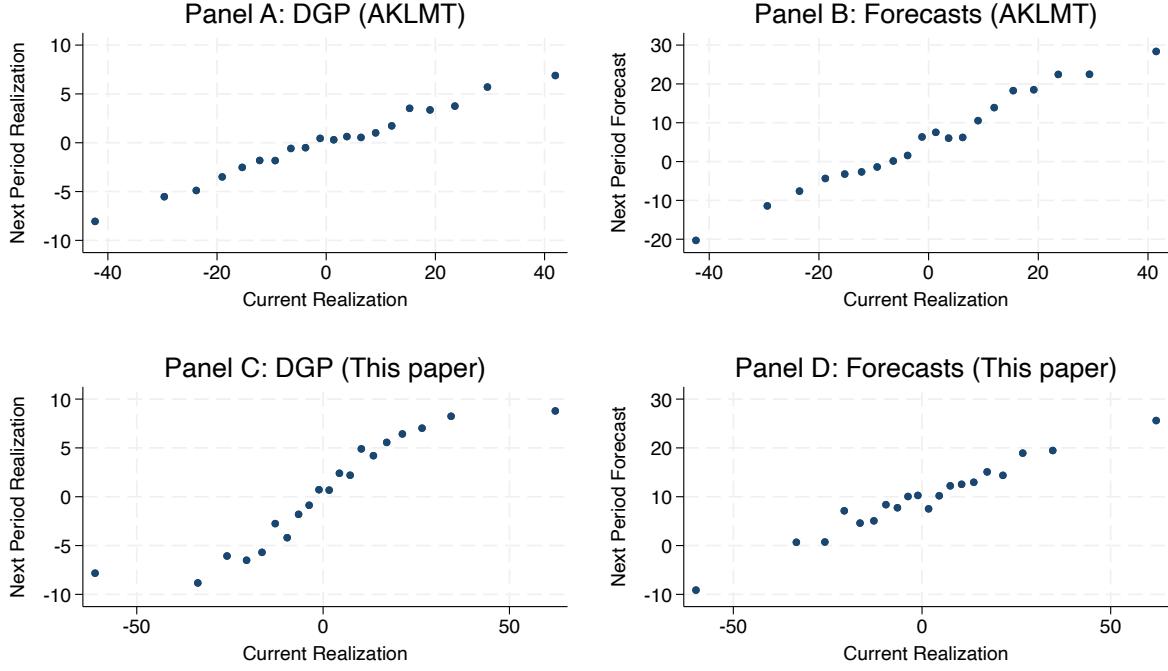
We ran the experiment in two waves with 201 participants on March 17, 2025 and 202 participants on April 29, 2025. Given that each participant makes 40 forecasts, our data have 16,120 observations. We compare our results using the DGP in (14) and (15) with those from a condition in AKLMT in which the process is an AR1 with Gaussian shocks and persistence parameter of 0.2.¹⁶

Before describing the results, we briefly discuss the relative benefits and costs of this experiment relative to our evidence in Section 2. The main benefit of running an experiment is that it allows us to directly control the parameters of the DGP. However, this benefit comes with several downsides.

¹⁵These parameters are from an estimation in a prior draft that used a smaller sample of data on Net Sales.

¹⁶We use 0.2 because it is the closest experimental condition in AKLMT to the regression coefficient of g_{it+1} and g_{it} when simulating from (14) and (15).

Figure 3: Data-Generating Process and Forecasts in Experimental Data



Notes: Panels A and C show binned scatterplots of g_{it+1} versus g_{it} in experimental data, where i is the participant and t the round of forecasting. Panels B and D show binned scatterplots of $F_{it}g_{it+1}$ against g_{it} . Panels A and B use data from Afrouzi et al. (2023), where the DGP is a Gaussian AR1 process with a persistence parameter of 0.2. Panels C and D use the experimental data described in the main text, where the DGP is fitted to our data – see equations (14) and (15).

First, participants in the experiment are fundamentally different from the forecasters in our data, who are professional equity analysts with incentives aside from accuracy. Second, the professional forecasters in our data are likely to have a better understanding of the DGP, while we cannot control the priors of the participants in our experiment. Finally, the environment of forecasting a variable in our experiment is obviously quite different from forecasting in the real world. Despite these differences, we still view this experiment as useful because, by comparing our results with those in AKLMT, we can perform a direct test of our theory's key prediction that transitory fat-tailed shocks create a non-linear relationship between errors and revisions.

Results. Panel A of Figure 3 shows the relationship between g_{it+1} and g_{it} in the data from AKLMT, where i is the participant and t is the round of forecasting. As expected, this relationship is linear. In contrast, Panel C shows the same relationship in our experimental data, where the DGP is in (14) and (15). Although our experimental data have one order of magnitude fewer observations than our firm-level data, these data still replicate the non-linear relationship in Fact #3.

Panels B and D of [Figure 3](#) provide evidence that, despite the differences in the DGP, participants' one-period ahead forecasts, $F_{it}g_{it+1}$, are still linear in g_{it} , unlike the true DGP. This is consistent with our theory that forecasters ignore the fact that large shocks are likely to be transitory. Comparing Panels A and C with B and D also illustrates that forecasters substantially overestimate the persistence of the process: $E(F_{it}g_{it+1} | g_{it})$ is steeper than $E(g_{it+1} | g_{it})$ at all points. The combination of this extrapolation (a core result of AKLMT) and the likely presence of expectation noise ([de Silva and Thesmar, 2024](#)) leads to a strong negative correlation between errors and revisions on average. Columns (1) and (3) of [Table 7](#) show that the relationship between forecast errors and revisions is significantly negative in our experiment and in AKLMT, with a similar magnitude.

Despite the presence of overreaction on average, the fact that $E(F_{it}g_{it+1} | g_{it})$ is linear in Panel D while $E(g_{it+1} | g_{it})$ is non-linear in Panel C provides preliminary evidence that forecasters are likely to overreact more in the tails of the distribution, consistent with our theory. We test this statistically by estimating (2), which is the same regression that we estimated on analyst forecast data in [Section 2.1](#). For the data from AKLMT, whose DGP is fully linear and Gaussian, column (2) shows no difference in the amount of overreaction that is present in the bulk of the distribution and the tails: the extra slope for the top and bottom quintile of past revisions is small and not statistically different from zero. In contrast, column (4) shows that, in our experimental data, this extra slope is three times larger and has a t -stat of 3.3, which is lower than in the analyst data but consistent with the fact that the sample is an order of magnitude smaller. Collectively, these findings provide direct evidence for our theory's key prediction that transitory fat-tailed shocks create a non-linear relationship between g_{t+1} and g_t and, as a result, between errors and revisions.

4.2 Evidence from SPF Macro Forecasts

We next study forecasts of macroeconomic outcomes from the [Survey of Professional Forecasters \(1968–2025\)](#) (SPF), which have been studied extensively in the literature on belief formation ([Bordalo et al., 2020a](#)). While we find qualitatively similar results to our findings at the firm-level, we interpret these results as suggestive given the much smaller sample size. In constructing our sample, we attempt to follow the exact choices made in the replication package of [Bordalo et al. \(2020a\)](#) (BGMS) as closely as possible, and refer the reader to BGMS for additional details.

Forecasting variables. The SPF is conducted around the end of the second month in each quarter and collects forecasts from individual forecasters that are anonymized and identified by forecaster IDs. We study the same set of outcome variables in the SPF as in BGMS, extending their sample

Table 7: Non-Linear Forecast Error-Revision Relationship: Experimental Data

| | Dependent Variable: Error_{it+1} | | | | | |
|---|---|-------------------|-------------------|------------------|------------------|-----|
| | Afrouzi et al. (2023) | New Experiment | (1) | (2) | (3) | (4) |
| Revision_{it} | -0.42 (-21.56) | -0.40 (-10.24) | -0.35 (-20.51) | -0.27 (-6.73) | | |
| $\text{Revision}_{it} \times \text{Top/Bottom 20\% of Revision}_{it}$ | | -0.06 (-1.17) | | -0.17 (-3.28) | | |
| Top 20% of Revision_{it} | | | -2.37 (-0.98) | | 7.64 (3.55) | |
| Bottom 20% of Revision_{it} | | | -7.57 (-2.94) | | -3.96 (-1.53) | |
| Constant | -6.98 (-5.33) | -5.22 (-4.03) | -8.40 (-7.86) | -9.14 (-7.40) | | |
| # of Obs. | 6859 | 6859 | 15003 | 15003 | | |

Notes: This table shows results from regressions of forecast errors onto revisions. In columns (1) and (2), we use data from Afrouzi et al. (2023), where the DGP is a Gaussian AR1 process. In columns (3) and (4), we use the experimental data described in the main text, where the DGP is fitted to our data – see equations (14) and (15). Columns (1) and (3) regress errors on revisions. Columns (2) and (4) estimate equation (2). All standard errors are clustered at the participant level. These regressions have fewer than 16,120 observations because computing forecast revisions loses one observation per subject and because we trim the top and bottom 1% of errors and revisions (as in Table 2).

forward to the second quarter in 2025.¹⁷ Following BGMS, for variables that are not stationary in levels (e.g., GDP or CPI), we look at annual growth rates from the most recent realized quarter to three quarters in the future. For the remaining variables, we look at levels three quarters in the future. While macroeconomic variables are released quarterly, they are often subsequently revised. Therefore, we follow BGMS and focus on initial releases from the [Real-Time Data Set for Macroeconomists \(1968–2025\)](#). As an example, actual GDP growth from quarter $t - 1$ to $t + 3$ is defined as the initial release at $t + 4$ of the GDP level in $t + 3$ divided by the contemporaneous release of the GDP level at $t - 1$. For financial variables, the actual outcomes are not revised, and we use historical data from the Federal Reserve Bank of St. Louis.

Forecast errors and revisions. The SPF contains forecasts for the levels of variables, and we focus on an annual forecast horizon. Forecasters report forecasts of outcomes in the current and next four quarters. Given the timing of this survey, forecasts made in quarter t are made with knowledge of the realized values of variables with quarterly releases in quarter $t - 1$ and the realized values of variables with monthly releases up to the previous month. For the variables that we study in growth

¹⁷The only variable that we do not include is housing starts, which is only available in a monthly vintage and has many missing observations.

rates, we transform forecasts of the levels of these variables into growth rates by dividing the level forecasts by the most recent realized level from the previous quarter. For the remaining variables, we use the level forecasts directly. Denoting y_t as a realized outcome in quarter t and $F_{jt}y_{t+h}$ as the forecast made in quarter t for quarter $t+h$ by forecaster j , we construct the forecast errors as $ERR_{t}y_{t+4} = y_{t+4} - F_{jt}y_{t+4}$, and the quarterly forecast revision as $REV_{t}y_{t+4} = F_{jt}y_{t+4} - F_{jt-1}y_{t+4}$. We winsorize outliers by removing, for each forecast horizon in a given quarter, forecasts that are more than 10 interquartile ranges away from the median. We also require that the forecasters in our sample make at least five forecasts.

Results. Table 8 performs the same analysis as Table 2 by estimating equation (2) at the forecaster-level for each of the variables in our SPF sample. As shown in the table, the size of our sample is much smaller, which reflects the fact that we only have one time-series for each variable. Nevertheless, we find patterns that are similar to our first fact at the firm-level: the error-revision relationship is non-linear with more overreaction in the tails of the distribution. Specifically, we find that $\hat{\beta} > 0$ for most variables, but $\hat{\gamma} < 0$ for all but one variable. These findings, coupled with our experimental evidence in Section 4.1, suggest that the non-linear error-revision relationship is more general than our sample of firm-level forecasts.

4.3 Evidence from Stock Return Predictability

This section examines and tests the predictions that our model of belief formation makes for return predictability. We focus on one type of predictability: momentum, the fact that past returns predict future returns (Jegadeesh and Titman, 2011).

Model prediction: non-linearity in momentum. To derive predictions about returns, we make two simplifying assumptions: (i) earnings growth is a constant fraction of sales growth, and (ii) the subjective discount rate is constant (as in Bouchaud et al. 2019 and Nagel and Xu 2022), which is consistent with evidence in De la O and Myers (2021). As shown in Appendix B.6, these assumptions coupled with the Campbell (1991) decomposition allow us to link returns with changes in expectations of our measure of Net Sales, g_t , using the following expression:

$$\log(1 + R_t) \approx \log(1 + R_f + \pi) + (g_t - F_{t-1}g_t) + \sum_{k=1}^{\infty} c^k REV_t g_{t+k}, \quad (16)$$

where R_t is the net return, π is the (constant) expected return, R_f is the gross risk-free rate, c is a linearization constant, and $REV_t g_{t+k}$ denotes investors' subjective revisions about future growth

Table 8: Non-Linear Error-Revision Relationship: SPF Macro Forecasts

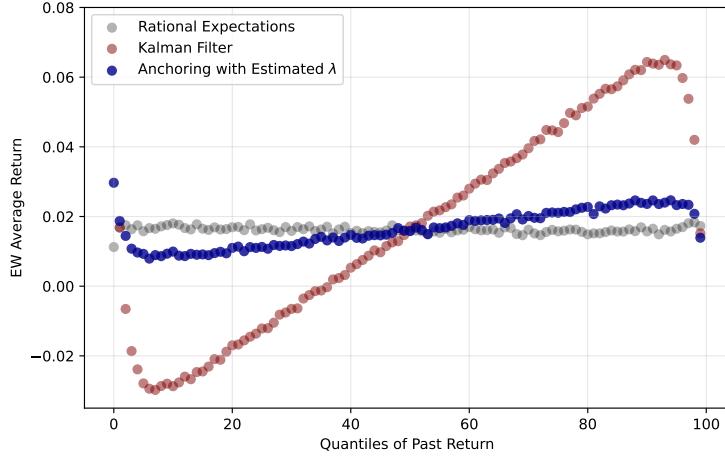
| Error _{it+1} | Revision _{it} | | Revision _{it} × Top/Bottom 20% of Revision _{it} | | Top 20% of Revision _{it} | | Bottom 20% of Revision _{it} | | # of Obs. | |
|---------------------------------|--------------------------|---------------|---|---------------|-----------------------------------|---------------|--------------------------------------|---------------|-----------|--|
| | | | | | | | | | | |
| | $\widehat{\beta}$ (1) | t-stat (2) | $\widehat{\gamma}$ (3) | t-stat (4) | $\widehat{\delta}^+$ (5) | t-stat (6) | $\widehat{\delta}^-$ (7) | t-stat (8) | | |
| Nominal GDP | 0.29 | 2.01 | -0.81 | -3.90 | 0.01 | 3.50 | -0.00 | -1.39 | 6026 | |
| Real GDP | 0.43 | 2.49 | -0.96 | -4.32 | 0.01 | 2.72 | -0.01 | -1.75 | 6045 | |
| Real Consumption | 0.07 | 0.42 | -0.75 | -2.66 | 0.01 | 3.62 | -0.00 | -0.60 | 4425 | |
| Unemployment | 0.48 | 2.55 | -0.98 | -3.46 | 0.46 | 2.35 | -0.52 | -4.14 | 6161 | |
| GDP Price Index | 0.42 | 2.02 | -0.70 | -2.93 | 0.01 | 4.03 | -0.00 | -0.82 | 5995 | |
| CPI | 0.32 | 1.56 | -0.65 | -3.04 | 0.01 | 2.38 | -0.00 | -1.94 | 4525 | |
| Industrial Production | 0.16 | 0.76 | -0.66 | -2.35 | 0.01 | 2.07 | -0.01 | -1.87 | 5567 | |
| Real Non-Residential Investment | 0.47 | 2.23 | -1.02 | -3.34 | 0.02 | 2.55 | -0.02 | -1.98 | 4308 | |
| Real Residential Inv. | 0.32 | 1.61 | -0.92 | -3.38 | 0.03 | 2.50 | -0.05 | -3.30 | 4302 | |
| Real Federal Govt. Consump. | -0.62 | -3.51 | 0.05 | 0.21 | -0.01 | -1.63 | -0.01 | -2.30 | 4158 | |
| Real State/Local Govt. Consump. | -0.15 | -0.83 | -0.43 | -2.23 | 0.00 | 1.41 | -0.00 | -1.92 | 4158 | |
| 3-Month Treasury Rate | 0.80 | 4.68 | -0.76 | -3.34 | 0.28 | 1.95 | -0.56 | -3.07 | 4491 | |
| 10-Year Treasury Rate | 0.14 | 0.67 | -0.46 | -1.81 | 0.28 | 2.78 | -0.12 | -0.91 | 3773 | |
| AAA Corporate Bond Rate | -0.15 | -1.15 | -0.13 | -0.78 | 0.08 | 0.91 | -0.09 | -0.77 | 3655 | |

Notes: This table reports the OLS estimates of equation (2), for each one of the variables in our SPF sample that is constructed following [Bordalo et al. \(2020a\)](#). We report the estimates of β (column 1), γ (column 3), δ^+ (column 5), and δ^- (column 7), along with the corresponding t -statistics with standard errors clustered by forecaster and time. Variables are ordered following [Bordalo et al. \(2020a\)](#).

between $t - 1$ and t . This expression is intuitive: given that discount rates are fixed, returns are driven by earnings surprises and revisions of future growth. As detailed in Appendix B.6, we use (16) to generate a panel of simulated returns based on the simulated panel of realized and expected sales growth from our estimated model.

[Figure 4](#) shows the prediction that our model makes for momentum by showing a binned scatterplot of current returns r_t against past returns r_{t-1} . The model predicts momentum in the bulk of past returns, but mean-reversion in the tails. This is consistent with our expectations formation model, which predicts overreaction to large news, and underreaction to intermediate ones. As the figure also shows, the predictability is stronger for the more “biased” expectation: Kalman filter expectations (blue dots) generate the strongest predictability, while forecasts anchored to rational expectations (red dots) generate weaker predictability. Predictability disappears completely when forecasts are rational (gray dots) or when ϵ is Gaussian, in which case the Kalman filter is rational.

Figure 4: Model-Implied Relationship Between Current and Future Returns



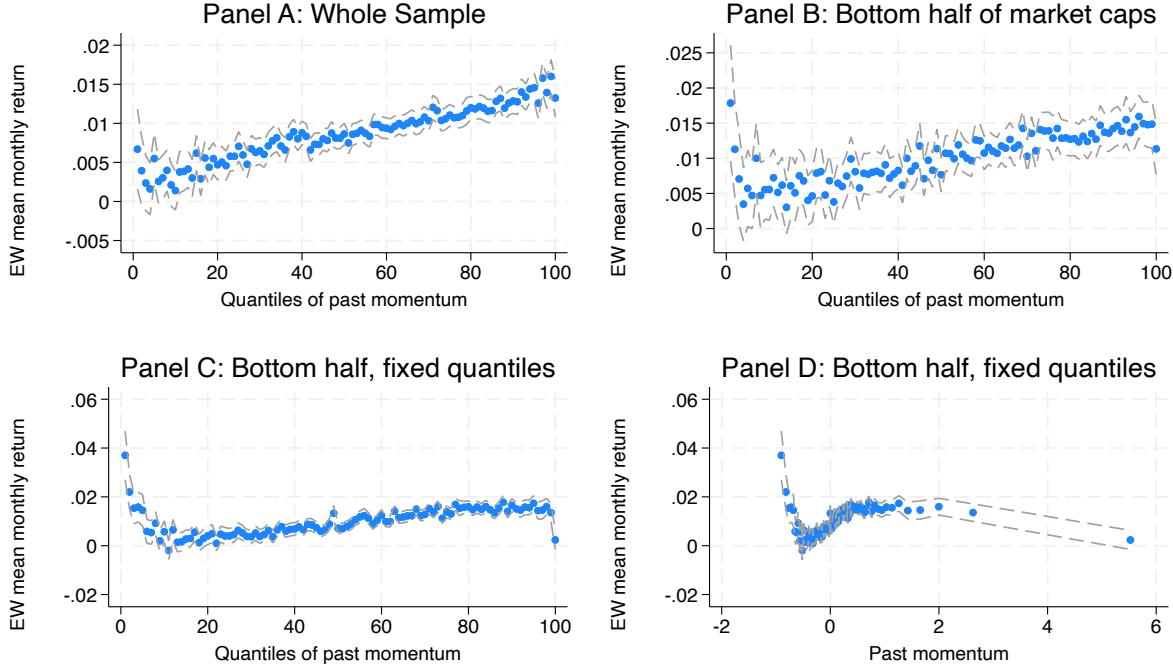
Notes: This figure shows a binned scatterplot of future returns against current returns in our simulated model. Given a set of earnings growth expectations, we compute returns using (16), as described in Appendix B.6. The three sets of points on the graph correspond to three cases in which beliefs are set equal to (i) the rational expectation computed with the particle filter, (ii) the Kalman filter, and (iii) the combination of the former two as in (12) with the value of λ in Table 5. We conduct this simulation assuming $R_f = 1.01$, $\pi = 5.5\%$, $c = 0.96$, and ρ is set to the estimated value in Table 5. We set the constant of proportionality between sales and earnings growth such that the standard deviation of returns is 15%.

Non-linearity in momentum in the data. We now examine whether momentum indeed exhibits mean-reversion in the tails in the data. We use CRSP monthly returns from 1927 to 2023 adjusted for delisting, and restrict to the sample of firms listed on NYSE, AMEX, and NASDAQ. Our measure of momentum follows the literature (Jegadeesh and Titman, 2011): every month t , momentum is the cumulative return between months $t - 12$ and $t - 1$ (thereby excluding the last month of past returns, which exhibit reversals). We define firm size as the market capitalization at $t - 12$.

We begin by plotting returns against centiles of momentum that are redefined each month, as in standard asset pricing tests. Panel A of Figure 5 shows the results for the entire CRSP sample: the line is upward sloping, consistent with the presence of momentum. However, we also find that there is a bit of mean-reversion for “super losers”, but not for “extreme winners”. In Panel B, we show that the non-linearity is more pronounced for smaller firms, which we define as those with below-median size. This is to be expected, as smaller stocks are more expensive to trade and therefore display more predictability (Novy-Marx and Velikov, 2016).

In Panel C of Figure 5, we show results when the centiles of past returns are defined on the entire sample rather than within each month, as we do within our model in Figure 4. While this is less standard in asset pricing tests because it implies the portfolios are not tradeable, it allows for the possibility that some of the non-linearity in momentum returns may come from time-series variation, which is more aligned with our model that does not distinguish between time-series and

Figure 5: Relationship Between Current and Past Returns in the Data



Notes: This figure shows binned scatterplots of monthly returns, r_t , as a function of past annual returns excluding month $t - 1$, $r_{t-12,t-1}$. There are 100 bins in each panel. Panels A-C use the standard convention of using bins of past returns $r_{t-12,t-1}$ as the x-axis. Panel D uses past returns, $r_{t-12,t-1}$, directly as the x-axis. Panel A uses the entire CRSP sample of stocks traded on AMEX, NASDAQ and NYSE. Panel B uses only the bottom half of stocks ranked by 12-month lagged market cap. In contrast to Panels A and B, quantiles of returns are defined on the entire sample in Panels C and D, rather than separately for each month. Dashed lines are 95% confidence bands based on standard errors assuming that returns are independent.

cross-sectional variation. Panel C shows that, using fixed centiles and focusing on firms below the 50% size cutoff, the non-linear relationship starts to appear significantly with a shape that is remarkably similar to that of our estimated model in [Figure 4](#). The differences relative to Panel B suggest that the time series of momentum returns generates some mean-reversion in the tails, consistent with the literature on momentum crashes ([Santa-Clara and Barroso, 2015](#)): when volatility is high, losers are more likely to overperform, and winners more likely to underperform. Panel D of [Figure 5](#) reproduces the analysis of Panel C, except that the x-axis is now rescaled to the average values of past returns. This last graph is less consistent with the practice of forming portfolios, as is typically done in asset pricing tests. Nevertheless, there is still significant non-linearity in the return-momentum relationship, in line with the prediction of our model.

5 Conclusion

In this paper, we argue that recognizing the complexity of the underlying data generating process, in particular its fat tails and non-Gaussian dynamics, is crucial for understanding the properties of subjective forecasts. We document three facts using data on forecasts of 22 firm-level variables by equity analysts: (i) the relationship between forecast revisions and future forecast errors—the variables used in [Coibion and Gorodnichenko \(2015\)](#) regressions—is strongly non-linear; (ii) the distributions of the underlying processes have fat tails; and (iii) the conditional expectation of future realizations is non-linear in current realizations, with persistence in the bulk and mean reversion in the tails. Next, we build a forecasting model that connects these facts. The key ingredients in our model are that the underlying process is non-Gaussian, but forecasters fail to recognize this and incorrectly fit a linear model to the data. Finally, we provide additional evidence in support of our theory by showing that it can explain evidence from an online forecasting experiment, that the non-linear relationship between errors and revisions is also present in macroeconomic forecasts, and that it provides an explanation for the presence of non-linearity in the momentum of stock returns.

Our paper raises several questions for further work. First, our model of belief formation is reduced-form. It would be fruitful to try to provide a microfoundation for why forecasters ignore fat tails in data-generating processes, which would allow us to study how this bias would manifest in other settings. One possible microfoundation that we emphasize is that forecasters may be insensitive to signal strength, as in [Augenblick et al. \(2024\)](#) and [Ba et al. \(2024\)](#). Second, our paper raises the question of how learning occurs in an environment with fat-tailed processes, which is likely to occur much more slowly and be more difficult. Finally, many of the driving processes in dynamic models of firm and household behavior have fat tails, such as revenue growth ([Stanley et al., 1996](#); [Boar et al., 2025](#); [Jaimovich et al., 2025](#)) and income growth ([Guvenen et al., 2014](#)). Our findings suggest that allowing agents to, at least partially, ignore the presence of the fat tails might improve the ability of these models to explain firm and household behavior.

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Online Appendix to “Expectations Formation with Fat-Tailed Processes: Evidence and Theory”

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A Additional Tables and Figures

Table IA.1: Summary Statistics: Consensus-Level Forecasts

| Variable | Transformation | Realizations | | Errors | | Revisions | | # of Obs. |
|---------------------------------|----------------|--------------|-----------|-------------|-----------|-------------|-----------|-----------|
| | | Mean (2) | SD (3) | Mean (4) | SD (5) | Mean (6) | SD (7) | |
| Net Sales | Growth | 0.06 | 0.18 | -0.02 | 0.14 | -0.00 | 0.12 | 144567 |
| Return On Equity | Level | 0.12 | 0.15 | -0.01 | 0.09 | -0.01 | 0.06 | 128459 |
| Earnings Per Share | Growth | 0.05 | 0.47 | -0.09 | 0.39 | -0.02 | 0.30 | 126341 |
| Net Income | Growth | 0.06 | 0.47 | -0.08 | 0.39 | -0.01 | 0.31 | 112501 |
| Pre-tax Profit | Growth | 0.05 | 0.45 | -0.09 | 0.38 | -0.01 | 0.28 | 106043 |
| Book Value Per Share | Growth | 0.05 | 0.19 | -0.03 | 0.19 | -0.01 | 0.13 | 99542 |
| EBITDA | Growth | 0.06 | 0.31 | -0.05 | 0.25 | -0.00 | 0.20 | 98311 |
| Return On Assets | Level | 0.05 | 0.06 | -0.01 | 0.04 | -0.01 | 0.03 | 89610 |
| Net Asset Value | Growth | 0.06 | 0.18 | -0.00 | 0.19 | -0.02 | 0.13 | 89572 |
| Dividend Per Share | Growth | 0.06 | 0.31 | -0.01 | 0.28 | -0.01 | 0.22 | 89522 |
| EBIT | Growth | 0.05 | 0.40 | -0.07 | 0.34 | -0.01 | 0.26 | 83365 |
| GAAP EPS | Growth | 0.05 | 0.57 | -0.11 | 0.47 | 0.01 | 0.39 | 82230 |
| Capital Expenditures | Growth | 0.04 | 0.54 | -0.12 | 0.61 | 0.19 | 0.61 | 78556 |
| Gross Margin | Level | 0.40 | 0.22 | -0.01 | 0.07 | -0.01 | 0.04 | 77380 |
| Cash Flow Per Share | Growth | 0.05 | 0.52 | -0.05 | 0.45 | -0.02 | 0.45 | 70384 |
| Net Debt | Growth | 0.04 | 0.42 | 0.09 | 0.43 | 0.07 | 0.38 | 44476 |
| Enterprise Value | Growth | 0.05 | 0.33 | 0.06 | 0.31 | 0.02 | 0.21 | 43661 |
| Operating Profit | Growth | 0.05 | 0.40 | -0.08 | 0.34 | -0.00 | 0.27 | 42824 |
| Cash Earnings Per Share | Growth | 0.04 | 0.32 | -0.04 | 0.30 | -0.03 | 0.26 | 22621 |
| EBITDA Per Share | Growth | 0.04 | 0.31 | -0.04 | 0.26 | -0.02 | 0.22 | 22411 |
| Funds From Operations Per Share | Growth | 0.01 | 0.16 | -0.02 | 0.12 | -0.03 | 0.11 | 3547 |
| EPS Before Goodwill | Growth | 0.08 | 0.35 | -0.01 | 0.29 | -0.04 | 0.25 | 2616 |

Notes: This table reports the mean, standard deviation, and number of observations for each one of the 22 firm variables forecasted by analysts in our sample. The sample includes all observations for which the consensus forecast error and revision are available for a given variable. Variables are ranked by decreasing number of observations. 19 variables are transformed into log growth as explained in the text; 3 are left untransformed. [Click here to go back to main text.](#)

Table IA.2: Non-Linear Error-Revision Relationship: Allowing for Asymmetric Slopes

| Variable | Error _{i,t+1} | | Revision _{i,t} | | Revision _{i,t} × Top 20% of Revision _{i,t} | | Revision _{i,t} × Bottom 20% of Revision _{i,t} | | Top 20% of Revision _{i,t} | | Bottom 20% of Revision _{i,t} | |
|---------------------------------|------------------------|--------|-------------------------|--------|--|--------|---|-------|------------------------------------|-------|---------------------------------------|--|
| | $\widehat{\beta}$ | | t-stat | | $\widehat{\gamma}^+$ | | t-stat | | $\widehat{\delta}^+$ | | t-stat | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) | (11) | |
| Net Sales | 0.11 | 6.31 | -0.26 | -11.33 | -0.27 | -12.74 | 0.02 | 12.25 | -0.02 | -5.72 | 399809 | |
| Earnings Per Share | 0.08 | 2.97 | -0.24 | -6.65 | -0.27 | -9.64 | 0.01 | 1.48 | -0.08 | -7.58 | 360011 | |
| Return On Equity | 0.23 | 9.75 | -0.37 | -8.59 | -0.22 | -11.03 | 0.01 | 11.30 | -0.01 | -6.94 | 345082 | |
| Net Income | 0.04 | 1.53 | -0.21 | -6.16 | -0.27 | -12.16 | 0.01 | 1.00 | -0.07 | -8.05 | 325230 | |
| EBITDA | -0.04 | -1.51 | -0.11 | -2.29 | -0.17 | -6.55 | 0.00 | 1.43 | -0.03 | -3.58 | 280111 | |
| Pre-tax Profit | 0.01 | 0.47 | -0.16 | -4.03 | -0.27 | -11.90 | 0.00 | 1.01 | -0.07 | -6.67 | 269883 | |
| EBIT | -0.05 | -1.60 | -0.14 | -3.09 | -0.26 | -11.11 | 0.00 | 0.56 | -0.05 | -4.61 | 231455 | |
| GAAP EPS | -0.02 | -0.73 | -0.10 | -2.84 | -0.18 | -7.87 | -0.02 | -1.01 | -0.07 | -3.97 | 171523 | |
| Book Value Per Share | -0.06 | -1.52 | -0.49 | -8.97 | -0.37 | -7.56 | 0.02 | 5.10 | -0.03 | -5.08 | 170561 | |
| Return On Assets | 0.41 | 10.35 | -0.70 | -11.25 | -0.35 | -9.63 | -0.00 | -0.77 | -0.01 | -6.80 | 157063 | |
| Gross Margin | 0.09 | 3.90 | -0.23 | -6.44 | -0.12 | -4.28 | 0.00 | 5.18 | -0.00 | -3.61 | 150220 | |
| Net Asset Value | 0.04 | 0.99 | -0.41 | -6.85 | -0.68 | -13.71 | 0.01 | 3.11 | -0.05 | -9.43 | 119885 | |
| Cash Flow Per Share | -0.28 | -15.86 | 0.04 | 1.47 | -0.15 | -6.36 | -0.01 | -0.94 | -0.07 | -5.97 | 112323 | |
| Net Debt | -0.17 | -8.31 | 0.13 | 5.69 | -0.54 | -21.33 | -0.06 | -5.67 | -0.05 | -4.02 | 90126 | |
| Enterprise Value | -0.20 | -5.82 | -0.01 | -0.12 | -0.44 | -11.09 | 0.02 | 1.80 | -0.07 | -7.76 | 88791 | |
| Dividend Per Share | -0.12 | -4.54 | -0.18 | -4.51 | -0.16 | -5.66 | -0.01 | -0.61 | -0.03 | -2.89 | 63840 | |
| Capital Expenditures | -0.36 | -21.93 | -0.14 | -5.98 | -0.19 | -6.85 | 0.05 | 3.53 | -0.07 | -5.62 | 62099 | |
| Operating Profit | 0.12 | 3.07 | -0.31 | -5.97 | -0.37 | -7.95 | 0.02 | 3.59 | -0.07 | -6.42 | 57756 | |
| EBITDA Per Share | -0.09 | -2.84 | -0.15 | -2.37 | -0.12 | -2.72 | 0.03 | 3.45 | -0.02 | -1.67 | 16824 | |
| Cash Earnings Per Share | -0.19 | -4.95 | -0.15 | -2.12 | -0.24 | -4.17 | 0.01 | 0.79 | -0.04 | -2.11 | 11578 | |
| Funds From Operations Per Share | -0.01 | -0.25 | -0.03 | -0.46 | -0.11 | -1.90 | 0.00 | 1.01 | -0.02 | -2.09 | 9279 | |
| EPS Before Goodwill | -0.13 | -2.29 | -0.18 | -1.46 | -0.06 | -0.71 | 0.06 | 2.61 | -0.06 | -3.33 | 2558 | |

Notes: This table reproduces Table 2 allowing for asymmetric slopes. Instead of a single interaction term γ for both top and bottom 20% of revisions, we estimate separate interaction coefficients γ^+ for the top 20% and γ^- for the bottom 20%. We report the estimates of β (column 1), γ^+ (column 3), γ^- (column 5), δ^+ (column 7), and δ^- (column 9), along with the corresponding t-statistics with standard errors clustered by firm and time. Variables are ranked by number of observations.

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Table IA.3: Non-Linear Error-Relationship: Consensus Forecasts

| Error _{it+1} | Revision _{it} | | Revision _{it} × Top/Bottom 20% of Revision _{it} | | Top 20% of Revision _{it} | | Bottom 20% of Revision _{it} | | # of Obs. | |
|---------------------------------|--------------------------|---------------|---|---------------|-----------------------------------|---------------|--------------------------------------|---------------|-----------|--|
| | | | | | | | | | | |
| | $\widehat{\beta}$ (1) | t-stat (2) | $\widehat{\gamma}$ (3) | t-stat (4) | $\widehat{\delta}^+$ (5) | t-stat (6) | $\widehat{\delta}^-$ (7) | t-stat (8) | | |
| Net Sales | 0.23 | 5.34 | -0.40 | -8.31 | 0.01 | 0.92 | -0.04 | -11.76 | 145225 | |
| Return On Equity | 0.40 | 7.63 | -0.44 | -6.86 | 0.00 | 2.01 | -0.03 | -10.09 | 128761 | |
| Earnings Per Share | 0.02 | 0.49 | -0.22 | -4.82 | -0.03 | -2.58 | -0.07 | -6.14 | 126893 | |
| Net Income | -0.08 | -1.79 | -0.12 | -2.77 | -0.03 | -1.76 | -0.05 | -4.69 | 112990 | |
| Pre-tax Profit | -0.07 | -1.38 | -0.13 | -2.69 | -0.02 | -1.63 | -0.04 | -3.18 | 106455 | |
| Book Value Per Share | -0.21 | -7.02 | -0.28 | -8.62 | -0.01 | -3.15 | -0.04 | -10.06 | 99986 | |
| EBITDA | -0.03 | -1.29 | -0.18 | -5.02 | -0.01 | -0.92 | -0.05 | -3.94 | 98819 | |
| Net Asset Value | -0.31 | -9.97 | -0.23 | -6.44 | 0.01 | 2.89 | -0.01 | -1.11 | 89954 | |
| Dividend Per Share | -0.00 | -0.02 | -0.30 | -4.20 | -0.00 | -0.22 | -0.04 | -4.75 | 89927 | |
| Return On Assets | 0.46 | 11.50 | -0.51 | -11.36 | -0.01 | -9.72 | -0.02 | -14.44 | 89807 | |
| EBIT | -0.16 | -4.82 | -0.08 | -2.13 | -0.02 | -1.83 | -0.03 | -2.57 | 83770 | |
| GAAP EPS | -0.19 | -4.19 | 0.01 | 0.23 | -0.02 | -1.12 | -0.00 | -0.30 | 82578 | |
| Capital Expenditures | -0.40 | -25.03 | -0.23 | -9.78 | 0.10 | 6.04 | -0.09 | -5.38 | 78871 | |
| Gross Margin | 0.08 | 2.39 | -0.27 | -6.58 | 0.00 | 3.16 | -0.01 | -7.67 | 77518 | |
| Cash Flow Per Share | -0.33 | -23.51 | 0.03 | 1.50 | 0.01 | 1.05 | 0.02 | 2.94 | 70787 | |
| Net Debt | -0.23 | -7.75 | -0.00 | -0.13 | 0.11 | 10.77 | 0.14 | 13.20 | 44693 | |
| Enterprise Value | -0.10 | -1.50 | -0.36 | -5.79 | 0.09 | 7.82 | -0.01 | -1.13 | 43855 | |
| Operating Profit | -0.01 | -0.15 | -0.32 | -4.52 | -0.00 | -0.36 | -0.07 | -3.93 | 43025 | |
| Cash Earnings Per Share | -0.39 | -12.04 | 0.01 | 0.12 | -0.02 | -1.99 | -0.02 | -0.90 | 22777 | |
| EBITDA Per Share | -0.09 | -1.91 | -0.16 | -2.56 | -0.00 | -0.26 | -0.03 | -2.78 | 22528 | |
| Funds From Operations Per Share | -0.00 | -0.04 | -0.23 | -2.29 | -0.00 | -0.13 | -0.04 | -4.77 | 3548 | |
| EPS Before Goodwill | -0.29 | -1.90 | 0.08 | 0.48 | 0.00 | 0.30 | -0.02 | -0.71 | 2646 | |

Notes: This table reproduces Table 2 using consensus-level forecasts. See Table 2 for additional notes. [Click here to go back to main text.](#)

Table IA.4: Non-Linear Error-Relationship: Percent Growth

| Error _{it+1} | Revision _{it} | | Revision _{it} × Top/Bottom 20% of Revision _{it} | | Top 20% of Revision _{it} | | Bottom 20% of Revision _{it} | | # of Obs. | |
|---------------------------------|------------------------|---------------|---|---------------|-----------------------------------|---------------|--------------------------------------|---------------|-----------|--|
| | | | | | | | | | | |
| | $\hat{\beta}$ (1) | t-stat (2) | $\hat{\gamma}$ (3) | t-stat (4) | $\hat{\delta}^+$ (5) | t-stat (6) | $\hat{\delta}^-$ (7) | t-stat (8) | | |
| Net Sales | 0.10 | 5.67 | -0.27 | -13.42 | 0.03 | 7.86 | -0.02 | -3.65 | 399832 | |
| Earnings Per Share | 0.00 | 0.05 | -0.23 | -7.17 | 0.06 | 4.11 | -0.05 | -2.06 | 364593 | |
| Net Income | -0.04 | -1.67 | -0.22 | -7.25 | 0.06 | 4.14 | -0.04 | -1.89 | 329340 | |
| EBITDA | -0.08 | -3.25 | -0.13 | -3.94 | 0.02 | 2.65 | -0.01 | -0.73 | 281746 | |
| Pre-tax Profit | -0.07 | -3.18 | -0.18 | -6.45 | 0.05 | 3.26 | -0.01 | -0.52 | 273285 | |
| EBIT | -0.12 | -4.30 | -0.14 | -3.88 | 0.04 | 3.11 | 0.00 | 0.24 | 234200 | |
| GAAP EPS | -0.11 | -5.92 | -0.15 | -4.91 | 0.07 | 2.89 | -0.02 | -0.74 | 173552 | |
| Book Value Per Share | -0.11 | -2.77 | -0.40 | -8.73 | 0.02 | 3.55 | -0.04 | -7.12 | 170925 | |
| Net Asset Value | -0.02 | -0.60 | -0.45 | -8.97 | 0.04 | 6.39 | -0.02 | -2.17 | 120417 | |
| Cash Flow Per Share | -0.30 | -17.77 | -0.02 | -1.12 | 0.00 | 0.37 | 0.00 | 0.07 | 113717 | |
| Net Debt | -0.09 | -4.25 | -0.04 | -1.78 | 0.09 | 6.20 | 0.25 | 16.00 | 96338 | |
| Enterprise Value | -0.17 | -4.60 | -0.04 | -0.89 | 0.05 | 5.12 | 0.06 | 4.64 | 88700 | |
| Dividend Per Share | -0.19 | -6.65 | -0.11 | -3.04 | -0.00 | -0.27 | -0.01 | -0.61 | 64154 | |
| Capital Expenditures | -0.38 | -25.04 | -0.46 | -24.24 | 0.29 | 15.58 | -0.21 | -15.90 | 62076 | |
| Operating Profit | 0.02 | 0.79 | -0.28 | -6.18 | 0.06 | 3.32 | -0.03 | -1.76 | 58369 | |
| EBITDA Per Share | -0.15 | -5.08 | -0.12 | -2.86 | 0.04 | 3.14 | -0.02 | -1.33 | 16905 | |
| Cash Earnings Per Share | -0.22 | -7.03 | -0.15 | -2.73 | 0.01 | 0.95 | -0.01 | -0.27 | 11651 | |
| Funds From Operations Per Share | 0.01 | 0.18 | -0.14 | -3.17 | 0.01 | 2.32 | -0.02 | -2.23 | 9324 | |
| EPS Before Goodwill | -0.18 | -2.33 | -0.18 | -2.28 | 0.10 | 4.41 | -0.11 | -1.75 | 2560 | |

Notes: This table reproduces Table 2 where we change our transformation of the variables that are converted into growth rates from log growth rates to percent growth rates. We compute percent growth rates as $g_{it} = \frac{X_{it} - X_{it-1}}{X_{it-1}}$, and redefine forecasts for these variables as $F_t^j g_{it+h} = \frac{F_t^j X_{it+h} - F_t^j X_{it+h-1}}{F_t^j X_{it+h-1}}$, with the convention that $F_t^j X_{it} = X_{it}$. See Table 2 for additional notes. [Click here to go back to main text.](#)

Table IA.5: Non-Linear Error-Relationship: Adjusting for Firm-Level Volatility

| Error _{it+1} | Revision _{it} | | Revision _{it} × Top/Bottom 20% of Revision _{it} | | Top 20% of Revision _{it} | | Bottom 20% of Revision _{it} | | # of Obs. | |
|---------------------------------|--------------------------|---------------|---|---------------|-----------------------------------|---------------|--------------------------------------|---------------|-----------|--|
| | | | | | | | | | | |
| | $\widehat{\beta}$ (1) | t-stat (2) | $\widehat{\gamma}$ (3) | t-stat (4) | $\widehat{\delta}^+$ (5) | t-stat (6) | $\widehat{\delta}^-$ (7) | t-stat (8) | | |
| Net Sales | 0.11 | 6.81 | -0.33 | -16.89 | 0.20 | 9.95 | -0.16 | -9.31 | 303623 | |
| Earnings Per Share | 0.03 | 1.49 | -0.26 | -7.67 | 0.15 | 7.12 | -0.12 | -3.85 | 272283 | |
| Return On Equity | 0.18 | 11.53 | -0.31 | -15.35 | 0.13 | 6.05 | -0.17 | -8.09 | 248053 | |
| Net Income | -0.02 | -1.11 | -0.24 | -7.07 | 0.17 | 7.12 | -0.10 | -2.90 | 242091 | |
| EBITDA | -0.03 | -1.67 | -0.29 | -11.04 | 0.19 | 7.35 | -0.13 | -4.85 | 205224 | |
| Pre-tax Profit | -0.00 | -0.18 | -0.32 | -11.37 | 0.20 | 8.94 | -0.13 | -3.43 | 196802 | |
| EBIT | -0.06 | -2.21 | -0.41 | -10.22 | 0.30 | 12.92 | -0.21 | -4.35 | 159223 | |
| Book Value Per Share | -0.12 | -2.92 | -0.40 | -8.27 | 0.16 | 5.64 | -0.17 | -5.36 | 119693 | |
| GAAP EPS | -0.03 | -1.06 | -0.20 | -5.21 | 0.15 | 4.30 | -0.06 | -1.97 | 110490 | |
| Return On Assets | 0.23 | 8.95 | -0.39 | -11.66 | 0.00 | 0.07 | -0.38 | -9.18 | 95610 | |
| Gross Margin | 0.11 | 5.23 | -0.25 | -11.33 | 0.12 | 5.29 | -0.17 | -9.62 | 85574 | |
| Cash Flow Per Share | -0.27 | -16.27 | -0.21 | -9.07 | 0.27 | 11.26 | -0.11 | -2.68 | 74669 | |
| Net Asset Value | -0.07 | -2.03 | -0.62 | -10.91 | 0.41 | 16.04 | -0.24 | -4.08 | 70815 | |
| Net Debt | -0.15 | -5.53 | -0.22 | -5.64 | 0.38 | 8.35 | 0.25 | 6.64 | 58422 | |
| Enterprise Value | -0.15 | -3.65 | -0.46 | -11.48 | 0.55 | 16.25 | -0.11 | -2.99 | 43748 | |
| Dividend Per Share | -0.21 | -8.59 | -0.15 | -4.54 | 0.12 | 4.02 | -0.07 | -2.22 | 38038 | |
| Capital Expenditures | -0.33 | -16.59 | -0.13 | -5.00 | 0.10 | 1.66 | -0.08 | -2.57 | 33460 | |
| Operating Profit | 0.15 | 2.85 | -0.46 | -7.14 | 0.26 | 6.52 | -0.12 | -2.47 | 22761 | |
| Funds From Operations Per Share | 0.01 | 0.18 | -0.11 | -1.86 | 0.11 | 3.03 | -0.11 | -2.27 | 6941 | |
| Cash Earnings Per Share | -0.25 | -5.81 | -0.26 | -4.16 | 0.23 | 3.90 | -0.28 | -3.31 | 4483 | |
| EBITDA Per Share | -0.04 | -1.08 | -0.36 | -6.07 | 0.25 | 4.43 | -0.24 | -5.92 | 4068 | |

Notes: This table reproduces Table 2 after normalizing forecasts and realizations by estimates of the firm-level standard deviation of the realizations. The sample here consists of firms for which we have at least 10 observations. See Table 2 for additional notes. This table does not include EPS Before Goodwill because we do not have enough observations. [Click here to go back to main text.](#)

Table IA.6: Non-Linear Error-Relationship: Adjusting for Aggregate Volatility

| Error _{it+1} | Revision _{it} | | Revision _{it} × Top/Bottom 20% of Revision _{it} | | Top 20% of Revision _{it} | | Bottom 20% of Revision _{it} | | # of Obs. | |
|---------------------------------|--------------------------|---------------|---|---------------|-----------------------------------|---------------|--------------------------------------|---------------|-----------|--|
| | | | | | | | | | | |
| | $\widehat{\beta}$ (1) | t-stat (2) | $\widehat{\gamma}$ (3) | t-stat (4) | $\widehat{\delta}^+$ (5) | t-stat (6) | $\widehat{\delta}^-$ (7) | t-stat (8) | | |
| Net Sales | 0.11 | 5.46 | -0.27 | -12.06 | 0.06 | 4.90 | -0.09 | -7.87 | 399555 | |
| Earnings Per Share | 0.08 | 2.53 | -0.25 | -10.00 | 0.02 | 1.26 | -0.11 | -5.65 | 359817 | |
| Return On Equity | 0.21 | 8.94 | -0.22 | -8.29 | 0.00 | 1.05 | -0.03 | -5.32 | 344222 | |
| Net Income | 0.03 | 1.42 | -0.24 | -13.95 | 0.02 | 1.25 | -0.10 | -5.90 | 325134 | |
| EBITDA | -0.04 | -1.34 | -0.14 | -5.99 | 0.02 | 1.18 | -0.05 | -2.64 | 279943 | |
| Pre-tax Profit | 0.00 | 0.09 | -0.21 | -10.90 | 0.03 | 1.79 | -0.08 | -3.97 | 269692 | |
| EBIT | -0.04 | -1.37 | -0.22 | -9.49 | 0.04 | 2.50 | -0.06 | -2.93 | 231398 | |
| GAAP EPS | -0.02 | -0.91 | -0.13 | -4.67 | -0.00 | -0.03 | -0.06 | -2.21 | 171482 | |
| Book Value Per Share | -0.07 | -1.71 | -0.41 | -7.92 | 0.03 | 2.49 | -0.12 | -6.63 | 170367 | |
| Return On Assets | 0.50 | 8.27 | -0.52 | -8.59 | -0.08 | -4.80 | -0.14 | -8.02 | 156935 | |
| Gross Margin | 0.10 | 4.23 | -0.15 | -5.62 | 0.00 | 1.33 | -0.02 | -5.75 | 150124 | |
| Net Asset Value | 0.02 | 0.62 | -0.55 | -11.89 | 0.11 | 8.67 | -0.11 | -5.41 | 119865 | |
| Cash Flow Per Share | -0.28 | -15.30 | -0.08 | -4.21 | 0.08 | 4.39 | -0.01 | -0.52 | 112303 | |
| Net Debt | -0.17 | -8.09 | -0.12 | -5.44 | 0.20 | 10.25 | 0.20 | 8.31 | 90095 | |
| Enterprise Value | -0.19 | -5.20 | -0.28 | -8.53 | 0.26 | 12.27 | -0.03 | -1.03 | 88775 | |
| Dividend Per Share | -0.14 | -5.46 | -0.19 | -7.40 | 0.00 | 0.01 | -0.08 | -4.87 | 63788 | |
| Capital Expenditures | -0.37 | -23.67 | -0.16 | -6.75 | 0.12 | 4.93 | -0.10 | -4.19 | 62079 | |
| Operating Profit | 0.12 | 3.29 | -0.35 | -8.27 | 0.07 | 3.57 | -0.11 | -4.80 | 57718 | |
| EBITDA Per Share | -0.09 | -3.02 | -0.13 | -3.47 | 0.04 | 1.88 | -0.06 | -1.98 | 16816 | |
| Cash Earnings Per Share | -0.18 | -4.54 | -0.19 | -3.41 | 0.04 | 1.20 | -0.04 | -0.90 | 11563 | |
| Funds From Operations Per Share | 0.05 | 1.62 | -0.15 | -2.72 | 0.03 | 1.19 | -0.06 | -1.67 | 9231 | |
| EPS Before Goodwill | -0.11 | -2.23 | -0.10 | -1.19 | 0.06 | 1.26 | -0.13 | -2.30 | 2552 | |

Notes: This table reproduces Table 2 after normalizing forecasts and realizations by estimates of the cross-sectional standard deviation of the realizations in each year. See Table 2 for additional notes. [Click here to go back to main text.](#)

Table IA.7: Non-Linear Error-Relationship: Removing Time Fixed Effects

| Error _{it+1} | Revision _{it} | | Revision _{it} × Top/Bottom 20% of Revision _{it} | | Top 20% of Revision _{it} | | Bottom 20% of Revision _{it} | | # of Obs. | |
|---------------------------------|--------------------------|---------------|---|---------------|-----------------------------------|---------------|--------------------------------------|---------------|-----------|--|
| | | | | | | | | | | |
| | $\widehat{\beta}$ (1) | t-stat (2) | $\widehat{\gamma}$ (3) | t-stat (4) | $\widehat{\delta}^+$ (5) | t-stat (6) | $\widehat{\delta}^-$ (7) | t-stat (8) | | |
| Net Sales | 0.08 | 3.35 | -0.25 | -11.65 | 0.02 | 6.84 | -0.02 | -6.34 | 399799 | |
| Earnings Per Share | 0.07 | 2.31 | -0.24 | -8.60 | 0.01 | 0.84 | -0.06 | -6.39 | 359967 | |
| Return On Equity | -0.00 | -0.00 | -0.24 | -0.66 | 0.10 | 1.04 | -1.31 | -1.42 | 340601 | |
| Net Income | 0.02 | 0.87 | -0.22 | -10.67 | 0.01 | 0.90 | -0.05 | -5.34 | 325194 | |
| EBITDA | -0.03 | -1.23 | -0.15 | -6.43 | 0.01 | 1.25 | -0.02 | -3.11 | 280094 | |
| Pre-tax Profit | -0.01 | -0.30 | -0.19 | -7.03 | 0.02 | 1.87 | -0.04 | -3.76 | 269840 | |
| EBIT | -0.04 | -1.39 | -0.23 | -9.58 | 0.03 | 3.25 | -0.02 | -2.40 | 231446 | |
| GAAP EPS | -0.05 | -0.98 | -0.10 | -2.08 | -0.01 | -0.94 | -0.04 | -2.81 | 171544 | |
| Book Value Per Share | -0.10 | -2.11 | -0.39 | -7.23 | 0.02 | 5.64 | -0.03 | -6.55 | 170543 | |
| Return On Assets | 0.21 | 3.08 | -0.24 | -3.20 | -0.00 | -1.61 | -0.01 | -4.42 | 157075 | |
| Gross Margin | 0.09 | 4.65 | -0.14 | -6.09 | 0.00 | 2.75 | -0.00 | -6.62 | 150215 | |
| Net Asset Value | -0.03 | -0.65 | -0.49 | -8.02 | 0.04 | 11.91 | -0.02 | -3.21 | 119887 | |
| Cash Flow Per Share | -0.25 | -16.94 | -0.09 | -5.15 | 0.05 | 4.70 | -0.00 | -0.30 | 112347 | |
| Net Debt | -0.15 | -6.46 | -0.14 | -6.44 | 0.11 | 10.47 | 0.11 | 11.70 | 90108 | |
| Enterprise Value | -0.20 | -5.78 | -0.26 | -7.32 | 0.09 | 10.94 | -0.02 | -2.10 | 88772 | |
| Dividend Per Share | -0.16 | -4.84 | -0.15 | -4.92 | -0.00 | -0.59 | -0.03 | -3.87 | 63847 | |
| Capital Expenditures | -0.37 | -21.85 | -0.16 | -7.53 | 0.04 | 4.03 | -0.06 | -6.36 | 62094 | |
| Operating Profit | 0.04 | 0.98 | -0.26 | -6.43 | 0.04 | 3.97 | -0.04 | -5.15 | 57758 | |
| EBITDA Per Share | -0.04 | -0.91 | -0.16 | -3.31 | 0.01 | 0.80 | -0.02 | -1.36 | 16820 | |
| Cash Earnings Per Share | -0.21 | -5.63 | -0.18 | -3.36 | 0.04 | 2.56 | -0.02 | -1.01 | 11578 | |
| Funds From Operations Per Share | -0.01 | -0.15 | -0.10 | -1.74 | 0.01 | 1.45 | -0.01 | -1.32 | 9280 | |
| EPS Before Goodwill | -0.13 | -2.12 | -0.07 | -0.92 | 0.01 | 0.65 | -0.06 | -2.40 | 2558 | |

Notes: This table reproduces Table 2 after removing the cross-sectional mean from forecast errors and forecast revisions, separately, for each year. See Table 2 for additional notes. [Click here to go back to main text.](#)

Table IA.8: Non-Linear Error-Relationship: Split by Analyst Experience

| Quartile of Past Experience | 1 | | 2 | | 3 | | 4 | |
|---------------------------------|-----------------------|---------------|-----------------------|---------------|-----------------------|---------------|-----------------------|---------------|
| | $\hat{\gamma}$ (1) | t-stat (2) | $\hat{\gamma}$ (3) | t-stat (4) | $\hat{\gamma}$ (5) | t-stat (6) | $\hat{\gamma}$ (7) | t-stat (8) |
| Net Sales | -0.28 | -16.28 | -0.28 | -14.70 | -0.25 | -11.09 | -0.25 | -7.54 |
| Earnings Per Share | -0.27 | -8.91 | -0.25 | -6.12 | -0.27 | -7.83 | -0.26 | -6.78 |
| Return On Equity | -0.27 | -9.27 | -0.30 | -12.03 | -0.27 | -9.72 | -0.22 | -4.86 |
| Net Income | -0.24 | -8.63 | -0.24 | -6.00 | -0.26 | -10.26 | -0.24 | -8.32 |
| EBITDA | -0.16 | -5.06 | -0.15 | -4.81 | -0.13 | -3.53 | -0.13 | -4.18 |
| Pre-tax Profit | -0.26 | -7.28 | -0.20 | -5.70 | -0.25 | -8.57 | -0.19 | -6.45 |
| EBIT | -0.23 | -7.12 | -0.19 | -6.82 | -0.21 | -4.97 | -0.24 | -6.00 |
| GAAP EPS | -0.13 | -3.66 | -0.12 | -3.58 | -0.14 | -4.52 | -0.15 | -3.32 |
| Book Value Per Share | -0.40 | -8.37 | -0.40 | -6.14 | -0.42 | -7.13 | -0.46 | -5.52 |
| Return On Assets | -0.48 | -9.55 | -0.43 | -6.74 | -0.43 | -7.70 | -0.37 | -7.01 |
| Gross Margin | -0.16 | -4.52 | -0.15 | -2.76 | -0.14 | -2.91 | -0.17 | -3.91 |
| Net Asset Value | -0.59 | -14.20 | -0.57 | -8.99 | -0.49 | -8.11 | -0.59 | -6.65 |
| Cash Flow Per Share | -0.03 | -0.71 | -0.05 | -1.96 | -0.08 | -3.10 | -0.10 | -2.18 |
| Net Debt | -0.13 | -3.98 | -0.08 | -2.41 | -0.11 | -2.15 | -0.18 | -5.92 |
| Enterprise Value | -0.23 | -4.56 | -0.26 | -5.86 | -0.30 | -8.44 | -0.26 | -5.61 |
| Dividend Per Share | -0.22 | -4.31 | -0.15 | -3.49 | -0.13 | -2.44 | -0.16 | -2.90 |
| Capital Expenditures | -0.15 | -6.30 | -0.13 | -3.79 | -0.10 | -2.82 | -0.26 | -6.44 |
| Operating Profit | -0.33 | -7.17 | -0.32 | -4.20 | -0.36 | -5.86 | -0.39 | -5.49 |
| EBITDA Per Share | -0.22 | -2.89 | -0.02 | -0.29 | -0.14 | -1.97 | -0.11 | -1.70 |
| Cash Earnings Per Share | -0.23 | -3.43 | -0.12 | -0.86 | -0.24 | -1.94 | -0.15 | -3.34 |
| Funds From Operations Per Share | 0.03 | 0.33 | -0.14 | -1.80 | -0.09 | -1.08 | -0.16 | -1.34 |
| EPS Before Goodwill | -0.02 | -0.22 | -0.03 | -0.17 | -0.06 | -0.22 | -0.30 | -2.31 |

Notes: Each pair of columns in this table reports the OLS estimates of equation (2), for each one of the 22 variables on a different subsample based on analyst experience. For each variable, the experience of analyst j is defined at date t as the number of firm-years for which analyst j has issued a forecast for this variable prior to t . The table reports estimates of γ , along with the corresponding t -statistics with standard errors clustered by firm and time. Variables are ranked by number of observations. [Click here to go back to main text.](#)

Table IA.9: Fat Tails in Realizations: Adjusting for Firm-Level Volatility

| Variable | Transformation | Kurtosis | Log-Log regression | | # of Obs. |
|---------------------------------|----------------|----------|--------------------|----------------|-----------|
| | | | Tail Parameter | R ² | |
| (1) | (2) | (3) | (4) | (5) | |
| Net Sales | Growth | 3.40 | 4.14 | 0.99 | 87500 |
| Earnings Per Share | Growth | 3.70 | 4.17 | 0.98 | 80346 |
| Return On Equity | Level | 3.76 | 4.15 | 1.00 | 73100 |
| Net Income | Growth | 3.60 | 4.25 | 0.98 | 69126 |
| Pre-tax Profit | Growth | 3.66 | 4.13 | 0.98 | 59902 |
| EBITDA | Growth | 3.40 | 4.40 | 0.98 | 59770 |
| Book Value Per Share | Growth | 3.64 | 3.94 | 0.99 | 50930 |
| EBIT | Growth | 3.49 | 4.46 | 0.98 | 44622 |
| Return On Assets | Level | 4.01 | 3.92 | 0.99 | 38226 |
| GAAP EPS | Growth | 3.47 | 4.59 | 0.98 | 36745 |
| Cash Flow Per Share | Growth | 3.23 | 4.58 | 0.98 | 33233 |
| Net Asset Value | Growth | 3.32 | 3.86 | 0.99 | 33099 |
| Gross Margin | Level | 5.38 | 3.17 | 0.99 | 27309 |
| Net Debt | Growth | 3.80 | 3.76 | 0.98 | 24395 |
| Dividend Per Share | Growth | 4.19 | 3.93 | 0.99 | 23469 |
| Enterprise Value | Growth | 2.80 | 5.07 | 0.98 | 18837 |
| Capital Expenditures | Growth | 2.80 | 5.26 | 0.98 | 17345 |
| Operating Profit | Growth | 3.60 | 4.24 | 0.98 | 10142 |
| EBITDA Per Share | Growth | 3.19 | 4.92 | 0.99 | 3089 |
| Cash Earnings Per Share | Growth | 3.10 | 5.08 | 0.97 | 2975 |
| Funds From Operations Per Share | Growth | 4.02 | 3.88 | 0.97 | 2250 |
| Normal Distribution | . | 3 | ∞ | . | . |

Notes: This table reproduces Table 3 after normalizing forecasts and realizations by estimates of the firm-level standard deviation of the realizations. The sample here consists of firms for which we have at least 10 observations. See Table 3 for additional notes. [Click here to go back to main text.](#)

Table IA.10: Fat Tails in Realizations: Adjusting for Aggregate Volatility

| Variable | Transformation | Kurtosis | Log-Log regression | | # of Obs. |
|---------------------------------|----------------|----------|--------------------|----------------|-----------|
| | | | Tail Parameter | R ² | |
| (1) | (2) | (3) | (4) | (5) | |
| Net Sales | Growth | 5.84 | 2.70 | 0.99 | 140481 |
| Earnings Per Share | Growth | 6.34 | 2.42 | 0.99 | 129735 |
| Return On Equity | Level | 7.62 | 2.76 | 1.00 | 127268 |
| Net Income | Growth | 6.28 | 2.42 | 0.99 | 113861 |
| Pre-tax Profit | Growth | 6.18 | 2.47 | 0.99 | 102386 |
| EBITDA | Growth | 6.07 | 2.53 | 0.99 | 99731 |
| Book Value Per Share | Growth | 8.38 | 2.13 | 1.00 | 89882 |
| EBIT | Growth | 6.36 | 2.48 | 0.99 | 82099 |
| Return On Assets | Level | 6.02 | 3.16 | 0.99 | 79477 |
| GAAP EPS | Growth | 5.88 | 2.57 | 0.99 | 74212 |
| Net Asset Value | Growth | 7.84 | 2.17 | 0.99 | 71494 |
| Gross Margin | Level | 2.47 | 7.21 | 0.98 | 64481 |
| Cash Flow Per Share | Growth | 5.52 | 2.64 | 0.99 | 62446 |
| Dividend Per Share | Growth | 6.87 | 2.50 | 0.99 | 49401 |
| Net Debt | Growth | 6.57 | 2.33 | 0.99 | 48338 |
| Enterprise Value | Growth | 4.44 | 3.12 | 0.99 | 48140 |
| Capital Expenditures | Growth | 4.96 | 2.82 | 0.99 | 44623 |
| Operating Profit | Growth | 6.08 | 2.54 | 0.99 | 33033 |
| EBITDA Per Share | Growth | 6.19 | 2.38 | 0.99 | 15344 |
| Cash Earnings Per Share | Growth | 6.01 | 2.44 | 0.99 | 10012 |
| Funds From Operations Per Share | Growth | 7.11 | 2.33 | 1.00 | 3399 |
| EPS Before Goodwill | Growth | 5.70 | 3.07 | 0.98 | 2413 |
| Normal Distribution | . | 3 | ∞ | . | . |

Notes: This table reproduces Table 3 after normalizing realizations by the yearly cross-sectional standard deviation of the realizations. See Table 2 for additional notes. [Click here to go back to main text.](#)

Table IA.11: Fat Tails in Realizations: Percent Growth

| Variable | Transformation | Kurtosis | Log-Log regression | | # of Obs. |
|---------------------------------|----------------|----------|--------------------|----------------|-----------|
| | | | Tail Parameter | R ² | |
| (1) | (2) | (3) | (4) | (5) | |
| Net Sales | Growth | 8.27 | 2.35 | 1.00 | 140482 |
| Earnings Per Share | Growth | 25.66 | 1.29 | 1.00 | 129736 |
| Net Income | Growth | 27.02 | 1.29 | 1.00 | 113862 |
| Pre-tax Profit | Growth | 23.34 | 1.36 | 1.00 | 102386 |
| EBITDA | Growth | 14.73 | 1.89 | 1.00 | 99730 |
| Book Value Per Share | Growth | 9.80 | 2.26 | 0.99 | 89884 |
| EBIT | Growth | 19.63 | 1.52 | 1.00 | 82100 |
| GAAP EPS | Growth | 28.02 | 1.16 | 1.00 | 74221 |
| Net Asset Value | Growth | 11.92 | 2.08 | 1.00 | 71494 |
| Cash Flow Per Share | Growth | 23.00 | 1.31 | 1.00 | 62446 |
| Dividend Per Share | Growth | 15.61 | 1.66 | 0.99 | 49401 |
| Net Debt | Growth | 28.51 | 1.17 | 1.00 | 48339 |
| Enterprise Value | Growth | 9.01 | 2.06 | 1.00 | 48140 |
| Capital Expenditures | Growth | 16.21 | 1.60 | 1.00 | 44624 |
| Operating Profit | Growth | 19.50 | 1.55 | 1.00 | 33033 |
| EBITDA Per Share | Growth | 13.53 | 2.01 | 0.99 | 15344 |
| Cash Earnings Per Share | Growth | 11.62 | 2.05 | 0.99 | 10013 |
| Funds From Operations Per Share | Growth | 7.36 | 2.29 | 1.00 | 3399 |
| EPS Before Goodwill | Growth | 23.33 | 1.63 | 0.99 | 2414 |
| Normal Distribution | . | 3 | ∞ | . | . |

Notes: This table reproduces Table 3 where we change our transformation of the variables that are converted into growth rates from log growth rates to percent growth rates. We compute percent growth rates as $g_{it} = \frac{X_{it} - X_{it-1}}{X_{it-1}}$. We do not report results for the three “level variables” which are not transformed anyway. See Table 3 for additional notes. [Click here to go back to main text.](#)

Table IA.12: Non-Linear Conditional Expectation of Realizations: Adjusting for Firm-Level Volatility

| Variable | g_{it+1} | | g_{it} | | $g_{it} \times$ Top/Bottom 20% of g_{it} | | Top 20% of g_{it} | | Bottom 20% of g_{it} | | # of Obs. (9) |
|---------------------------------|----------------------------|---------------|-----------------------------|---------------|--|---------------|-------------------------------|---------------|---------------------------|--|------------------|
| | $\widehat{\beta}_g$ (1) | t-stat (2) | $\widehat{\gamma}_g$ (3) | t-stat (4) | $\widehat{\delta}_g^+$ (5) | t-stat (6) | $\widehat{\delta}_g^-$ (7) | t-stat (8) | | | |
| | | | | | | | | | | | |
| Net Sales | 0.43 | 11.87 | -0.36 | -6.89 | 0.63 | 6.90 | 0.10 | 3.80 | | | 87500 |
| Earnings Per Share | 0.30 | 9.77 | -0.53 | -12.58 | 0.77 | 12.90 | 0.10 | 2.42 | | | 80346 |
| Return On Equity | 0.89 | 74.58 | -0.03 | -2.00 | 0.22 | 2.73 | 0.13 | 6.02 | | | 73100 |
| Net Income | 0.30 | 10.46 | -0.51 | -11.63 | 0.75 | 10.10 | 0.12 | 2.72 | | | 69126 |
| Pre-tax Profit | 0.30 | 9.28 | -0.54 | -13.47 | 0.79 | 12.29 | 0.09 | 2.39 | | | 59902 |
| EBITDA | 0.28 | 9.90 | -0.40 | -10.64 | 0.64 | 9.13 | 0.10 | 3.04 | | | 59770 |
| Book Value Per Share | 0.51 | 17.75 | -0.45 | -13.20 | 0.76 | 14.56 | 0.16 | 6.28 | | | 50930 |
| EBIT | 0.33 | 10.13 | -0.53 | -10.57 | 0.81 | 9.82 | 0.08 | 2.00 | | | 44622 |
| Return On Assets | 0.87 | 51.63 | 0.06 | 3.60 | -0.20 | -2.41 | 0.06 | 2.29 | | | 38226 |
| GAAP EPS | 0.12 | 3.74 | -0.46 | -11.98 | 0.72 | 11.63 | 0.11 | 2.33 | | | 36745 |
| Cash Flow Per Share | -0.13 | -3.89 | -0.16 | -4.44 | 0.36 | 9.43 | 0.14 | 5.13 | | | 33233 |
| Net Asset Value | 0.49 | 15.54 | -0.40 | -9.22 | 0.74 | 10.05 | 0.06 | 1.71 | | | 33099 |
| Gross Margin | 1.00 | 293.41 | -0.01 | -2.07 | 0.07 | 2.78 | 0.21 | 6.66 | | | 27309 |
| Net Debt | 0.12 | 4.18 | -0.23 | -6.82 | 0.34 | 9.23 | 0.09 | 2.23 | | | 24395 |
| Dividend Per Share | 0.44 | 10.73 | -0.51 | -9.41 | 0.85 | 9.68 | 0.15 | 4.85 | | | 23469 |
| Enterprise Value | -0.03 | -0.66 | -0.16 | -2.64 | 0.28 | 3.48 | 0.02 | 0.31 | | | 18837 |
| Capital Expenditures | 0.03 | 0.87 | -0.11 | -2.27 | 0.15 | 2.10 | -0.07 | -1.05 | | | 17345 |
| Operating Profit | 0.22 | 4.03 | -0.48 | -6.28 | 0.77 | 7.66 | 0.06 | 1.03 | | | 10142 |
| EBITDA Per Share | 0.30 | 4.89 | -0.29 | -2.81 | 0.55 | 2.64 | 0.22 | 1.88 | | | 3089 |
| Cash Earnings Per Share | 0.12 | 1.97 | -0.29 | -2.53 | 0.52 | 2.88 | 0.19 | 1.81 | | | 2975 |
| Funds From Operations Per Share | 0.39 | 5.53 | -0.39 | -3.50 | 0.51 | 3.84 | -0.10 | -1.01 | | | 2250 |

Notes: This table reproduces Table 4 after normalizing forecasts and realizations by estimates of the firm-level standard deviation of the realizations. The sample here consists of firms for which we have at least 10 observations. This explains why “EPS Before Goodwill” is not reported here (there are not enough observations). See Table 4 for additional notes. [Click here to go back to main text.](#)

Table IA.13: Non-Linear Conditional Expectation of Realizations: Adjusting for Aggregate Volatility

| Variable | g_{it+1} | | g_{it} | | $g_{it} \times$ Top/Bottom 20% of g_{it} | | Top 20% of g_{it} | | Bottom 20% of g_{it} | | # of Obs. (9) |
|---------------------------------|----------------------------|---------------|-----------------------------|---------------|--|---------------|-------------------------------|---------------|---------------------------|--------|------------------|
| | $\widehat{\beta}_g$ (1) | t-stat (2) | $\widehat{\gamma}_g$ (3) | t-stat (4) | $\widehat{\delta}_g^+$ (5) | t-stat (6) | $\widehat{\delta}_g^-$ (7) | t-stat (8) | | | |
| | Net Sales | 0.42 | 12.19 | -0.41 | -8.60 | 0.32 | 6.58 | 0.01 | 0.56 | 140481 | |
| Earnings Per Share | 0.20 | 5.75 | -0.47 | -15.52 | 0.32 | 10.57 | -0.03 | -1.04 | 129735 | | |
| Return On Equity | 0.44 | 2.30 | 0.08 | 0.41 | 0.02 | 0.31 | -0.03 | -0.68 | 127268 | | |
| Net Income | 0.19 | 5.96 | -0.45 | -13.33 | 0.33 | 9.35 | -0.02 | -0.63 | 113861 | | |
| Pre-tax Profit | 0.19 | 5.25 | -0.43 | -10.98 | 0.34 | 10.26 | -0.01 | -0.28 | 102386 | | |
| EBITDA | 0.22 | 7.38 | -0.39 | -11.49 | 0.31 | 8.89 | -0.01 | -0.38 | 99731 | | |
| Book Value Per Share | 0.47 | 14.84 | -0.51 | -13.69 | 0.23 | 10.52 | -0.08 | -5.45 | 89882 | | |
| EBIT | 0.24 | 5.58 | -0.44 | -10.61 | 0.30 | 8.28 | -0.01 | -0.36 | 82099 | | |
| Return On Assets | 0.64 | 4.03 | -0.04 | -0.20 | 0.10 | 0.72 | 0.04 | 0.77 | 79477 | | |
| GAAP EPS | 0.02 | 0.34 | -0.37 | -8.35 | 0.28 | 10.25 | 0.00 | 0.01 | 74212 | | |
| Net Asset Value | 0.47 | 13.30 | -0.45 | -12.15 | 0.24 | 12.54 | -0.08 | -4.83 | 71494 | | |
| Gross Margin | 0.90 | 16.42 | -0.23 | -1.65 | 0.69 | 1.73 | 0.53 | 1.80 | 64481 | | |
| Cash Flow Per Share | -0.18 | -6.21 | -0.13 | -4.34 | 0.18 | 7.33 | 0.07 | 3.56 | 62446 | | |
| Dividend Per Share | 0.43 | 7.51 | -0.65 | -11.34 | 0.35 | 9.69 | -0.06 | -2.07 | 49401 | | |
| Net Debt | 0.08 | 4.18 | -0.27 | -8.77 | 0.24 | 10.35 | 0.00 | 0.18 | 48338 | | |
| Enterprise Value | -0.06 | -1.19 | -0.09 | -1.43 | 0.13 | 2.35 | -0.03 | -0.98 | 48140 | | |
| Capital Expenditures | 0.06 | 1.67 | -0.22 | -5.86 | 0.11 | 4.71 | -0.10 | -3.42 | 44623 | | |
| Operating Profit | 0.24 | 4.77 | -0.49 | -10.77 | 0.43 | 10.51 | -0.02 | -0.46 | 33033 | | |
| EBITDA Per Share | 0.12 | 2.67 | -0.24 | -3.11 | 0.13 | 2.12 | 0.01 | 0.09 | 15344 | | |
| Cash Earnings Per Share | 0.08 | 2.12 | -0.34 | -5.68 | 0.21 | 4.36 | -0.03 | -0.86 | 10012 | | |
| Funds From Operations Per Share | 0.46 | 4.25 | -0.46 | -3.72 | 0.22 | 2.09 | -0.09 | -1.08 | 3399 | | |
| EPS Before Goodwill | 0.21 | 2.09 | -0.54 | -3.97 | 0.40 | 3.45 | -0.08 | -1.00 | 2413 | | |

Notes: This table reproduces Table 4 after normalizing realizations by yearly cross-sectional dispersion of realizations. See Table 4 for additional notes. [Click here to go back to main text.](#)

Table IA.14: Non-Linear Conditional Expectation of Realizations: Percent Growth

| Variable | g_{it+1} | | g_{it} | | $g_{it} \times$ Top/Bottom 20% of g_{it} | | Top 20% of g_{it} | | Bottom 20% of g_{it} | | # of Obs. (9) |
|---------------------------------|----------------------------|---------------|-----------------------------|---------------|--|---------------|-------------------------------|---------------|---------------------------|--------|------------------|
| | $\widehat{\beta}_g$ (1) | t-stat (2) | $\widehat{\gamma}_g$ (3) | t-stat (4) | $\widehat{\delta}_g^+$ (5) | t-stat (6) | $\widehat{\delta}_g^-$ (7) | t-stat (8) | | | |
| | Net Sales | 0.44 | 12.74 | -0.39 | -8.69 | 0.11 | 7.31 | 0.02 | 2.17 | 140482 | |
| Earnings Per Share | 0.14 | 4.50 | -0.17 | -5.25 | 0.14 | 7.03 | 0.35 | 8.50 | 129736 | | |
| Net Income | 0.14 | 5.26 | -0.17 | -6.14 | 0.15 | 7.09 | 0.37 | 8.16 | 113862 | | |
| Pre-tax Profit | 0.19 | 6.40 | -0.20 | -7.08 | 0.15 | 7.04 | 0.33 | 7.33 | 102386 | | |
| EBITDA | 0.21 | 6.68 | -0.24 | -6.75 | 0.12 | 7.32 | 0.13 | 5.53 | 99730 | | |
| Book Value Per Share | 0.45 | 14.10 | -0.52 | -15.55 | 0.10 | 14.80 | -0.01 | -2.31 | 89884 | | |
| EBIT | 0.23 | 6.87 | -0.25 | -6.98 | 0.13 | 5.89 | 0.23 | 5.69 | 82100 | | |
| GAAP EPS | -0.05 | -1.54 | 0.01 | 0.40 | 0.08 | 3.16 | 0.60 | 9.23 | 74221 | | |
| Net Asset Value | 0.46 | 14.08 | -0.45 | -13.88 | 0.10 | 18.03 | -0.01 | -1.19 | 71494 | | |
| Cash Flow Per Share | -0.20 | -6.26 | 0.16 | 5.62 | 0.01 | 0.79 | 0.51 | 15.81 | 62446 | | |
| Dividend Per Share | 0.36 | 7.92 | -0.45 | -8.44 | 0.17 | 9.34 | 0.13 | 4.57 | 49401 | | |
| Net Debt | 0.07 | 2.47 | -0.10 | -2.97 | 0.11 | 6.62 | 0.30 | 11.81 | 48339 | | |
| Enterprise Value | -0.06 | -1.19 | 0.00 | 0.03 | 0.05 | 1.99 | 0.09 | 6.29 | 48140 | | |
| Capital Expenditures | 0.04 | 1.37 | -0.10 | -3.37 | 0.09 | 8.06 | 0.11 | 4.57 | 44624 | | |
| Operating Profit | 0.19 | 4.49 | -0.22 | -5.24 | 0.17 | 7.36 | 0.25 | 7.34 | 33033 | | |
| EBITDA Per Share | 0.11 | 3.09 | -0.13 | -2.72 | 0.07 | 2.51 | 0.13 | 2.90 | 15344 | | |
| Cash Earnings Per Share | 0.08 | 2.01 | -0.20 | -3.90 | 0.10 | 4.60 | 0.14 | 5.47 | 10013 | | |
| Funds From Operations Per Share | 0.34 | 3.53 | -0.32 | -2.54 | 0.05 | 2.45 | -0.01 | -0.39 | 3399 | | |
| EPS Before Goodwill | 0.13 | 2.37 | -0.16 | -2.37 | 0.09 | 1.43 | 0.17 | 3.53 | 2414 | | |

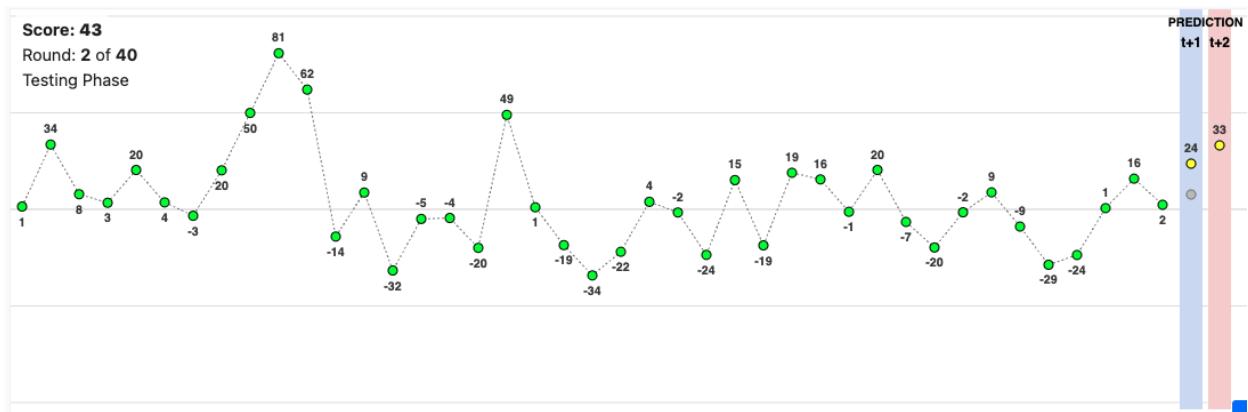
Notes: This table reproduces [Table 4](#) where we change our transformation of the variables that are converted into growth rates from log growth rates to percent growth rates. We compute percent growth rates as $g_{it} = \frac{X_{it} - X_{it-1}}{X_{it-1}}$. We do not report results for the three “level variables” which are not transformed anyway. See [Table 4](#) for additional notes. [Click here to go back to main text.](#)

Table IA.15: Fit of a Gaussian Approximation in the Bulk of the Distribution

| Variable | R^2 (1) | # of Obs. (2) |
|---------------------------------|--------------|------------------|
| Net Sales | 0.954 | 87500 |
| Earnings Per Share | 0.939 | 80345 |
| Return On Equity | 0.831 | 73100 |
| Net Income | 0.949 | 69126 |
| Pre-tax Profit | 0.940 | 59903 |
| EBITDA | 0.951 | 59771 |
| Book Value Per Share | 0.950 | 50929 |
| EBIT | 0.941 | 44620 |
| Return On Assets | 0.717 | 38226 |
| GAAP EPS | 0.908 | 36745 |
| Cash Flow Per Share | 0.929 | 33233 |
| Net Asset Value | 0.902 | 33099 |
| Gross Margin | 0.508 | 27309 |
| Net Debt | 0.894 | 24396 |
| Dividend Per Share | 0.893 | 23469 |
| Enterprise Value | 0.776 | 18838 |
| Capital Expenditures | 0.826 | 17344 |
| Operating Profit | 0.771 | 10142 |
| EBITDA Per Share | 0.631 | 3089 |
| Cash Earnings Per Share | 0.600 | 2976 |
| Funds From Operations Per Share | 0.583 | 2249 |

Notes: This table tests whether the central portion of the distribution of realizations is approximately Gaussian. For each variable, x , we keep one observation per firm-year. We restrict the sample to observations between the 20th and 80th percentiles. We then compute the empirical log-PDF for each centile bin c as $p_c = \log\left(\frac{1}{100} \frac{1}{x_c^+ - x_c^-}\right)$, where x_c^+ is the upper bound of c and x_c^- is the lower bound of c . We compute \hat{x}_c as the mean value of x in centile c . We then regress p_c on \hat{x}_c and \hat{x}_c^2 . For a Gaussian distribution, the R^2 of this regression should equal to 1. Variables are ranked by number of observations. [Click here to go back to main text.](#)

Figure IA.1: Screenshot of the Experiment Interface



Notes: This figure shows a screenshot of our experiment. [Click here to go back to main text.](#)

B Proofs and Additional Derivations

In this appendix, we prove all of the results in the main text. We prove our results under a weaker version of Assumption 1, stated below.

Assumption IA.1. *The distribution of ϵ_t is symmetric and regularly varying:*

$$P(\epsilon_t > x) = \int_x^\infty f_\epsilon(\epsilon) d\epsilon \sim x^{-\nu} L(x) \text{ as } \epsilon \rightarrow \infty,$$

where $\nu > 2$ is the index and $L(x)$ is slowly varying in the sense that, for all c , $\frac{L(cx)}{L(x)} \xrightarrow{x \rightarrow \infty} 1$.

B.1 Proof of Proposition 1

Proof. From the definition of the process, we obtain:

$$E(g_{t+1}|g_t) = \rho E(g_t^*|g_t)$$

We now apply Tweedie's formula (Robbins, 1956; Efron, 2012), which states that, for $y = x + \eta$, where x is Gaussian and x and η are independent, then:

$$E(x|y) = -\sigma_x^2 \frac{d \log h(y)}{dy}$$

where h is the marginal distribution of y and σ_x^2 the variance of x . Combining the last two equations with $x = g_t^*$ and $\eta = \epsilon_t$ delivers the result. \square

B.2 Proof of Proposition 2

Proof. By the definition of a conditional expectation, we have:

$$E(g^*|g) = \frac{\int x f_\epsilon\left(\frac{g-x}{\sigma_\epsilon}\right) \Phi(x) dx}{\int f_\epsilon\left(\frac{g-x}{\sigma_\epsilon}\right) \Phi(x) dx} = \frac{\int x \frac{f_\epsilon\left(\frac{g-x}{\sigma_\epsilon}\right)}{f_\epsilon\left(\frac{g}{\sigma_\epsilon}\right)} \Phi(x) dx}{\int \frac{f_\epsilon\left(\frac{g-x}{\sigma_\epsilon}\right)}{f_\epsilon\left(\frac{g}{\sigma_\epsilon}\right)} \Phi(x) dx}. \quad (\text{IA.1})$$

Given Assumption IA.1, we have

$$\frac{f_\epsilon\left(\frac{g-x}{\sigma_\epsilon}\right)}{f_\epsilon\left(\frac{g}{\sigma_\epsilon}\right)} = \frac{(g-x)^{-\nu-1} L_1\left(\frac{g-x}{\sigma_\epsilon}\right)}{g^{-\nu-1} L_1\left(\frac{g}{\sigma_\epsilon}\right)} \quad (\text{IA.2})$$

$$= \left(1 - \frac{x}{g}\right)^{-\nu-1} \frac{L_1\left(\frac{g-x}{\sigma_\epsilon}\right)}{L_1\left(\frac{g}{\sigma_\epsilon}\right)}. \quad (\text{IA.3})$$

where the first equality follows from the assumption that ϵ is regularly varying for some slowly varying $L_1(\cdot)$. Now we will derive some bounds on each component of the previous expression, with the goal of deriving bounds on $\frac{f_\epsilon\left(\frac{g-x}{\sigma_\epsilon}\right)}{f_\epsilon\left(\frac{g}{\sigma_\epsilon}\right)}$. By the binomial expansion, we have:

$$\left(1 - \frac{x}{g}\right)^{-\nu-1} = 1 + (\nu+1)\frac{x}{g} + O\left(\frac{x^2}{g^2}\right).$$

When $|x| \leq |g|^\beta$ and $\beta \in (0, 1/2)$, this remainder term is $O(g^{2\beta-2}) = o\left(\frac{1}{g}\right)$. Next, using the Uniform Convergence Theorem for slowly varying functions (Bingham et al., 1987, Theorem 1.2.1), we have

$$\lim_{|g| \rightarrow \infty} \sup_{|x| \leq g^\beta} \left| \frac{L_1\left(\frac{g-x}{\sigma_\epsilon}\right)}{L_1\left(\frac{g}{\sigma_\epsilon}\right)} \right| = 1.$$

Multiplying the previous two results, we have, for $|x| \leq |g|^\beta$,

$$\frac{f_\epsilon\left(\frac{g-x}{\sigma_\epsilon}\right)}{f_\epsilon\left(\frac{g}{\sigma_\epsilon}\right)} = 1 + (\nu+1)\frac{x}{g} + o\left(\frac{1}{g}\right).$$

Next, consider the region $|g/2| > |x| > |g|^\beta$. Taylor's theorem implies that, for some constant C ,

$$\left| \left(1 - \frac{x}{g}\right)^{-\nu-1} - 1 - (\nu+1)\frac{x}{g} \right| \leq C \left(\frac{x}{g}\right)^2$$

Since L_1 is slowly varying, the Uniform Convergence Theorem implies that

$$\sup_{|x| \leq |g|/2} \left| \frac{L_1\left(\frac{g-x}{\sigma_\epsilon}\right)}{L_1\left(\frac{g}{\sigma_\epsilon}\right)} - 1 \right| \longrightarrow 0 \quad \text{as } |g| \rightarrow \infty.$$

Hence, uniformly for $|x| \leq |g|/2$,

$$\frac{L_1\left(\frac{g-x}{\sigma_\epsilon}\right)}{L_1\left(\frac{g}{\sigma_\epsilon}\right)} = 1 + o(1).$$

Combining these results, we obtain

$$\frac{f_\epsilon\left(\frac{g-x}{\sigma_\epsilon}\right)}{f_\epsilon\left(\frac{g}{\sigma_\epsilon}\right)} = \left(1 - \frac{x}{g}\right)^{-\nu-1} \frac{L_1\left(\frac{g-x}{\sigma_\epsilon}\right)}{L_1\left(\frac{g}{\sigma_\epsilon}\right)}$$

$$= 1 + (\nu + 1) \frac{x}{g} + O\left(\frac{x^2}{g^2}\right) + o(1),$$

where the $o(1)$ term is uniform over $|x| \leq |g|/2$. Therefore, for $|g|^\beta < |x| < |g|/2$,

$$\left| \frac{f_\epsilon\left(\frac{g-x}{\sigma_\epsilon}\right)}{f_\epsilon\left(\frac{g}{\sigma_\epsilon}\right)} - 1 - (\nu + 1) \frac{x}{g} \right| \leq C' \frac{x^2}{g^2} + o(1),$$

for some constant $C' > 0$.

Finally, we consider the region $|x| \geq |g|/2$. In this region, Taylor's theorem does not give a uniform bound because the remainder term approaches ∞ as $|x| \rightarrow |g|$. However, since ϵ is regularly varying and has a bounded density, we have

$$\frac{f_\epsilon\left(\frac{g-x}{\sigma_\epsilon}\right)}{f_\epsilon\left(\frac{g}{\sigma_\epsilon}\right)} \leq B|g|^{\nu+1},$$

for some $B > 0$.

Next, return to the integrals in (IA.1). Focusing on the denominator, we have

$$\begin{aligned} \int \frac{f_\epsilon\left(\frac{g-x}{\sigma_\epsilon}\right)}{f_\epsilon\left(\frac{g}{\sigma_\epsilon}\right)} \Phi(x) dx &= \int \left(1 + (\nu + 1) \frac{x}{g}\right) \Phi(x) dx + \int \left(\frac{f_\epsilon\left(\frac{g-x}{\sigma_\epsilon}\right)}{f_\epsilon\left(\frac{g}{\sigma_\epsilon}\right)} - 1 - (\nu + 1) \frac{x}{g} \right) \Phi(x) dx \\ &= \int \left(1 + (\nu + 1) \frac{x}{g}\right) \Phi(x) dx + \int_{|x| \leq g^\beta} \left(\frac{f_\epsilon\left(\frac{g-x}{\sigma_\epsilon}\right)}{f_\epsilon\left(\frac{g}{\sigma_\epsilon}\right)} - 1 - (\nu + 1) \frac{x}{g} \right) \Phi(x) dx \\ &\quad + \int_{|x| > g^\beta} \left(\frac{f_\epsilon\left(\frac{g-x}{\sigma_\epsilon}\right)}{f_\epsilon\left(\frac{g}{\sigma_\epsilon}\right)} - 1 - (\nu + 1) \frac{x}{g} \right) \Phi(x) dx \end{aligned}$$

where the final equality just breaks up the integral. Focusing on the first integral, we have

$$\begin{aligned} \left| \int_{|x| \leq g^\beta} \left(\frac{f_\epsilon\left(\frac{g-x}{\sigma_\epsilon}\right)}{f_\epsilon\left(\frac{g}{\sigma_\epsilon}\right)} - 1 - (\nu + 1) \frac{x}{g} \right) \Phi(x) dx \right| &\leq \sup_{|x| \leq g^\beta} \left(\frac{f_\epsilon\left(\frac{g-x}{\sigma_\epsilon}\right)}{f_\epsilon\left(\frac{g}{\sigma_\epsilon}\right)} - 1 - (\nu + 1) \frac{x}{g} \right) \int_{|x| \leq g^\beta} \Phi(x) dx \\ &= o\left(\frac{1}{g}\right) \end{aligned}$$

where the second equality follows from the bounds above. Turning to the second integral, we split it into two regions based on the bounds above. The first of these two regions is $|x| \in (g^\beta, |g|/2) \equiv \mathcal{R}$. In this region,

we have

$$\begin{aligned}
\left| \int_{|x| \in \mathcal{R}} \left(\frac{f_\epsilon \left(\frac{g-x}{\sigma_\epsilon} \right)}{f_\epsilon \left(\frac{g}{\sigma_\epsilon} \right)} - 1 - (\nu + 1) \frac{x}{g} \right) \Phi(x) dx \right| &\leq \int_{|x| \in \mathcal{R}} \left| \frac{f_\epsilon \left(\frac{g-x}{\sigma_\epsilon} \right)}{f_\epsilon \left(\frac{g}{\sigma_\epsilon} \right)} - 1 - (\nu + 1) \frac{x}{g} \right| \Phi(x) dx \\
&\leq \frac{C'}{g^2} \int_{|x| \in \mathcal{R}} x^2 \Phi(x) dx + o(1) \int_{|x| \in \mathcal{R}} \Phi(x) dx \\
&= o\left(\frac{1}{g}\right)
\end{aligned}$$

where the first line follows from the triangle inequality, the second inequality follows from the bound established above, and the final line follows because a Gaussian PDF implies $\int_{|x| > g^\beta} x^k \Phi(x) dx \rightarrow 0$ for any finite k as $|g| \rightarrow \infty$.

For the second region, $|x| \geq |g|/2$, we have

$$\begin{aligned}
\left| \int_{|x| \geq |g|/2} \left(\frac{f_\epsilon \left(\frac{g-x}{\sigma_\epsilon} \right)}{f_\epsilon \left(\frac{g}{\sigma_\epsilon} \right)} - 1 - (\nu + 1) \frac{x}{g} \right) \Phi(x) dx \right| &\leq \int_{|x| \geq |g|/2} \left(B|g|^{\nu+1} + 1 + (\nu + 1) \frac{|x|}{|g|} \right) \Phi(x) dx \\
&= O\left(|g|^{\nu+1} e^{-g^2/(8\sigma_{g^*}^2)}\right) \\
&= o\left(\frac{1}{g}\right)
\end{aligned}$$

where the first line follows from the bounds above, the second line follows from Gaussian tail bounds, and the third line follows from the fact that exponential decay dominates polynomial growth.

Putting together the expressions for these two integrals, we have

$$\int \frac{f_\epsilon \left(\frac{g-x}{\sigma_\epsilon} \right)}{f_\epsilon \left(\frac{g}{\sigma_\epsilon} \right)} \Phi(x) dx = \int \left(1 + (\nu + 1) \frac{x}{g} \right) \Phi(x) dx + o\left(\frac{1}{g}\right).$$

We can then use the same arguments to show that

$$\int x \frac{f_\epsilon \left(\frac{g-x}{\sigma_\epsilon} \right)}{f_\epsilon \left(\frac{g}{\sigma_\epsilon} \right)} \Phi(x) dx = \int x \left(1 + (\nu + 1) \frac{x}{g} \right) \Phi(x) dx + o\left(\frac{1}{g}\right).$$

Combining the previous two results and using the fact that $\int \Phi(x) dx = 1$, $\int x \Phi(x) dx = 0$, and $\int x^2 \Phi(x) dx = \sigma_{g^*}^2$, we have, as $|g| \rightarrow \infty$,

$$E(g^* | g) \sim \frac{(\nu + 1) \sigma_{g^*}^2}{g}.$$

Applying the data-generating process delivers the desired result. \square

B.3 Proof of Proposition 3

Proof. By definition of subjective expectations:

$$ERR_{t+1} = \frac{1-K}{K} REV_t + \underbrace{(g_{t+1} - \rho g_t)}_{\sigma_\epsilon \epsilon_{t+1} + \sigma_u u_{t+1} - \rho \sigma_\epsilon \epsilon_t} \quad (\text{IA.4})$$

This implies:

$$E(ERR_{t+1} | REV_t) = \frac{1-K}{K} REV_t - \rho \sigma_\epsilon E(\epsilon_t | REV_t) \quad (\text{IA.5})$$

Now, we focus on the second term on the right-hand side. Given the definition of forecasts in equation (9), we can express forecast revisions using the following orthogonal decomposition:

$$REV_t = K \rho \sigma_\epsilon \left[\underbrace{\epsilon_t - K \rho \sum_{s \geq 0} ((1-K)\rho)^s \epsilon_{t-s-1}}_{\equiv E_t} + \underbrace{\frac{\sigma_u}{\sigma_\epsilon} \sum_{s \geq 0} ((1-K)\rho)^s u_{t-s}}_{\equiv U_t} \right]$$

Under Assumption IA.1, we can make use of the following result for the asymptotic behavior of regularly varying random variables to characterize the tail behavior of E_t :

Lemma IA.1 (Geluk and de Vries 2006, Theorem 1). *Suppose X_0, X_1, \dots are IID regularly varying random variables and c_0, c_1, \dots are weights that represent a convergent geometric series. Then*

$$P \left(\sum_{s \geq 0} c_i X_i > x \right) \sim \sum_{s \geq 0} P(c_i X_i > x) \text{ as } x \rightarrow \infty$$

Applying Lemma IA.1 to the definition of E_t , we have, as $E \rightarrow +\infty$,

$$\begin{aligned} P(E_t > E) &= P \left(K \rho \sum_{s \geq 0} ((1-K)\rho)^s \epsilon_{t-s-1} > E \right) \\ &\sim \sum_{s \geq 0} P(K \rho ((1-K)\rho)^s \epsilon_{t-s-1} > E) \\ &\sim \sum_{s \geq 0} L(E) \left(\frac{K \rho ((1-K)\rho)^s}{E} \right)^\nu \\ &= L(E) \frac{1}{E^\nu} \times \frac{K^\nu \rho^\nu}{1 - \rho^\nu (1-K)^\nu} \\ &\sim P(\epsilon_t > E) \times \frac{K^\nu \rho^\nu}{1 - \rho^\nu (1-K)^\nu} \end{aligned}$$

The final expression shows that E_t is also regularly varying with index ν but a different scale given by the

formula. Given this result, we can characterize the $E(\epsilon_t | REV_t)$ for large revisions as follows:

$$\begin{aligned} E(\epsilon_t | REV_t = r) &= E(\epsilon_t | K\rho\sigma_\epsilon(\epsilon_t - E_t + U_t) = r) \\ &= E\left(\epsilon_t | \epsilon_t - E_t + U_t = \frac{r}{K\rho\sigma_\epsilon}\right) \\ &\sim E\left(\epsilon_t | \epsilon_t - E_t = \frac{r}{K\rho\sigma_\epsilon}\right) \text{ as } r \rightarrow +\infty \end{aligned}$$

where the last line follows from the fact that E_t has the same tail parameter as ϵ_t , which is thicker than that of U_t by Assumption 2. To compute the expectations of one regularly varying variable conditional on the sum of two, one needs to know the relative scale of these two variables. As shown above, E_t has the same index as ϵ_t , but its counter-CDF differs by a scale of $\frac{K^\nu\rho^\nu}{1-\rho^\nu(1-K)^\nu}$ in the tails. Applying Proposition 4.9 from Denuit et al. (2025) along with the fact that ϵ_t (and hence E_t) is symmetric, we have, as $r \rightarrow +\infty$:

$$\begin{aligned} E\left(\epsilon_t | \epsilon_t - E_t = \frac{r}{K\rho\sigma_\epsilon}\right) &\sim \frac{L(\frac{r}{K\rho\sigma_\epsilon})}{L(\frac{r}{K\rho\sigma_\epsilon}) + L(\frac{r}{K\rho\sigma_\epsilon})\frac{K^\nu\rho^\nu}{1-\rho^\nu(1-K)^\nu}} \frac{r}{K\rho\sigma_\epsilon} \\ &= \frac{1}{1 + \frac{K^\nu\rho^\nu}{1-\rho^\nu(1-K)^\nu}} \frac{r}{K\rho\sigma_\epsilon} \end{aligned}$$

Now, we plug this result into equation (IA.5) and obtain that:

$$E(ERR_{t+1} | REV_t) \sim \underbrace{\left[\frac{1-K}{K} - \frac{1}{K} \frac{1}{1 + \frac{K^\nu\rho^\nu}{1-(1-K)^\nu\rho^\nu}} \right]}_{\equiv C} REV_t.$$

Repeating the analogous argument for the case in which $REV_t \rightarrow -\infty$ and using the fact that ϵ_t and u_t are symmetric delivers the first desired result.

To show that $C < 0$, we use the definition of the steady state Kalman gain, K , and prior variance, P :

$$\begin{aligned} K &= \frac{P}{P + \sigma_\epsilon^2} \\ P(1 - \rho^2(1 - K)) &= \sigma_u^2 \end{aligned}$$

whose combination shows that

$$1 - K = \frac{K}{\frac{\sigma_u^2}{\sigma_\epsilon^2} + K\rho^2}.$$

For $C < 0$, we need

$$\frac{1-K}{K} < \frac{1}{K} \cdot \frac{1}{1 + \frac{K^\nu\rho^\nu}{1-(1-K)^\nu\rho^\nu}}.$$

Multiplying both sides by K and cross-multiplying yields

$$(1 - K)K^\nu \rho^\nu + K(1 - K)^\nu \rho^\nu < K.$$

Dividing by K and factoring

$$(1 - K)\rho^\nu [K^{\nu-1} + (1 - K)^{\nu-1}] < 1.$$

Since $\nu > 2$, we have $\nu - 1 > 1$. For any $x \in (0, 1)$, this implies $x^{\nu-1} < x$. Applying this to both K and $1 - K$:

$$K^{\nu-1} + (1 - K)^{\nu-1} < K + (1 - K) = 1.$$

Therefore:

$$(1 - K)\rho^\nu [K^{\nu-1} + (1 - K)^{\nu-1}] < (1 - K)\rho^\nu < 1$$

where the final inequality comes from the fact that $\rho, K \in (0, 1)^2$. This is the desired inequality two lines above, which concludes the proof. \square

B.4 Proof of Corollary 1

Proof. The definition of the Kalman filter implies that it is the minimum linear mean squared error predictor. Combined with the law of iterated expectations, this implies that:

$$\begin{aligned} E(ERR_{t+1}REV_t) &= E\left[E(ERR_{t+1}REV_t | |REV_t| \geq r) P(|REV_t| \geq r) \right. \\ &\quad \left. + E(ERR_{t+1}REV_t | |REV_t| < r) P(|REV_t| < r)\right] = 0. \end{aligned}$$

By Proposition 3, there exists an \bar{R} such that $E(ERR_{t+1}REV_t | |REV_t| \geq \bar{R}) < 0$. Therefore, the fact that $E(ERR_{t+1}REV_t) = 0$ implies that $E(ERR_{t+1}REV_t | |REV_t| < \bar{R}) > 0$, which is the desired result. \square

B.5 Proof of Proposition 4

Proof. By the definition of the process, we have:

$$\text{var}(g_{t+1} | g_t) = \rho^2 \text{var}(g_t^* | g_t) + \sigma_u^2 + \sigma_\epsilon^2$$

Differentiating the log density twice gives:

$$\frac{d^2}{dg^2} \log h(g) = \frac{h''(g)}{h(g)} - \left(\frac{h'(g)}{h(g)}\right)^2.$$

Proposition 1 implies that $\frac{h'(g)}{h(g)} = -\frac{E(g_t^*|g)}{\sigma_{g^*}^2}$, which in turn gives:

$$\frac{h''(g)}{h(g)} = -\frac{1}{\sigma_{g^*}^2} + \frac{E((g_t^*)^2|g)}{\sigma_{g^*}^4}.$$

Combining the previous two expressions gives

$$\begin{aligned}\frac{d^2}{dg^2} \log h(g) &= -\frac{1}{\sigma_{g^*}^2} + \frac{E((g_t^*)^2|g)}{\sigma_{g^*}^4} - \left(\frac{E(g_t^*|g)}{\sigma_{g^*}^2}\right)^2 \\ &= \frac{1}{\sigma_{g^*}^4} (-\sigma_{g^*}^2 + \text{var}(g_t^*|g_t)).\end{aligned}$$

Using the first formula above for $\text{var}(g_t^*|g_t)$ gives the desired result. □

B.6 Additional Details on Asset Pricing Model in Section 4.3

We first derive (16). Define the return on a stock as $R_{t+1} = \frac{P_{t+1}+D_{t+1}}{P_t} - 1$, which has sales S_t , earnings \mathcal{E}_t , and a payout ratio $DE_t = \frac{D_t}{\mathcal{E}_t}$. Consistent with our variable definitions and transformations in Section 1, denote the (log) growth rate of sales as $g_t = \log S_t - \log S_{t-1}$. Assume that earnings growth is a constant fraction of sales growth: $\log\left(\frac{\mathcal{E}_t}{\mathcal{E}_{t-1}}\right) = \gamma g_t$. Denoting lower-case letters as logs and following [Campbell and Shiller \(1988\)](#), we can approximate the price-earnings ratio to first order around the mean price-dividend ratio as:

$$p_t - e_t = \kappa + \gamma g_{t+1} - r_{t+1} + (1 - c)(d_{t+1} - e_{t+1}) + c(p_{t+1} - e_{t+1}),$$

where $c = \frac{e^{\bar{pd}}}{1+e^{\bar{pd}}}$, \bar{pd} is the mean of the log price-dividend ratio, and κ is a constant. Assuming a constant log payout ratio of \bar{de} , we can iterate the first equation forward to obtain:

$$p_t - e_t = \frac{\kappa}{1 - c} + \bar{de} + \sum_{k=1}^{\infty} c^{k-1} (\gamma g_{it+k} - r_{it+k}) + \lim_{k \rightarrow \infty} c^k (p_{it+k} - e_{it+k}).$$

Imposing the usual transversality condition, the previous equation becomes:

$$p_t - e_t = \frac{\kappa}{1 - c} + \bar{de} + \sum_{k=1}^{\infty} c^{k-1} (\gamma g_{it+k} - r_{it+k}).$$

Now, rearranging the first equation above, we have

$$r_{t+1} = \kappa + \gamma g_{t+1} + (1 - c)\bar{de} + c(p_{t+1} - e_{t+1}) - (p_t - e_t).$$

Letting F_t denote investors' subjective beliefs, the previous two equations imply:

$$F_{t+1}r_{t+1} - F_t r_{t+1} = \gamma F_{t+1}g_{t+1} - \gamma F_t g_{t+1} + c[(p_{t+1} - e_{t+1}) - F_t(p_{t+1} - e_{t+1})].$$

Rewriting, we obtain:

$$r_{t+1} - F_t r_{t+1} = \gamma(F_{t+1} - F_t) \sum_{k=0}^{\infty} c^k g_{t+1+k} - (F_{t+1} - F_t) \sum_{k=1}^{\infty} c^k r_{t+1+k}.$$

To derive predictions based on our model of subjective expectations, we follow [Bouchaud et al. \(2019\)](#) and [Nagel and Xu \(2022\)](#) and assume that subjective risk premia are constant and equal to the risk-free rate, r_f , plus a constant risk premium, π . Under this assumption, the final term in the previous equation is zero and we obtain (16):

$$r_{t+1} = \log(r_f + \pi) + \underbrace{\gamma(F_{t+1} - F_t) \sum_{k=0}^{\infty} c^k g_{t+1+k}}_{=\sum_{k=1}^{\infty} c^k REV_t g_{t+k}}.$$

Given a panel of simulated earnings growth expectations, we can compute the final two terms in the previous equation by recognizing that our model of belief formation implies:

$$F_t \sum_{k=0}^{\infty} c^k g_{t+1+k} = F_t g_{t+1} \sum_{k=0}^{\infty} c^k \rho^k = \frac{F_t g_{t+1}}{1 - c\rho}.$$

Using this relationship, we can now simulate a path of return realizations using the following relationship:

$$r_{t+1} = \log(1 + R_f + \pi) + \frac{\gamma}{1 - c\rho} (g_{t+1} - F_t g_{t+1}).$$

We then set γ to generate a volatility of R_t equal to 15%, after winsorizing all returns that are above 100%.

C Computing Rational Expectations Using the Particle Filter

Constructing the forecasts in (12) requires computing the rational expectation, $E(g_{t+1} | g_0, \dots, g_t)$. In the case where ϵ_t is normally distributed, this corresponds to the Kalman filter and takes the simple closed form in (11). However, outside of this special case, this expectation cannot be computed in closed form and instead must be computed using sequential Monte Carlo methods. We choose to compute this expectation using the particle filtering algorithm from [Fernandez-Villaverde and Rubio-Ramirez \(2007\)](#) (also known as sequential importance sampling), which is described in Algorithm 1.

The goal of this algorithm is to estimate the posterior distribution of a latent state process, $\{g_t^*\}$, given a sequence of observed data, $\{g_t\}_{t=1}^T$. The algorithm proceeds by approximating the filtering distribution $p(g_t^* | g_1, \dots, g_t)$ using a collection of P particles, each representing a possible realization of the latent state. At each time step t , these particles are propagated forward via the state transition equation for g_t^* , incorporating stochastic innovations sampled from the distribution of u_t . The new particles are then evaluated against the observed data using a likelihood function $f_\epsilon(\cdot)$, the density function of the distribution of ϵ_t . Particles are assigned weights according to this likelihood, and a resampling step is used to adjust the distribution of the particles based on their posterior probabilities. Under mild conditions, [Fernandez-Villaverde and Rubio-Ramirez \(2007\)](#) show that the expectations computed in Step 5, which depend on P , converge to their population counterparts as $P \rightarrow \infty$. This is the sense in which the algorithm “works.”

In the case where $f_\epsilon(\cdot)$ is the density function of a normal distribution, the expectations computed in Step 5 converge to those of the Kalman filter as $P \rightarrow \infty$. However, in the presence of non-Gaussian shocks, as in our case, these two solutions will differ. The key step in this algorithm is Step 6, where each particle’s weight is computed using $f_\epsilon(\cdot)$. Intuitively, this weighting step incorporates the non-normality of $f_\epsilon(\cdot)$ by “tilting” the contribution of each particle to the posterior distribution of the latent state according to how well it explains the data under the true distribution. As a result, particles that better align with the observed data under the correct distribution are favored during the resampling step (Step 7), allowing the particle approximation to capture features—like skewness or fat tails—that the Kalman filter necessarily misses.

We run the particle filter on each of our 100 simulations of length 100,000 with P set to 10,000.¹⁸ The most computationally intensive part of this Algorithm 1 is Step 7, which requires resampling a large number of particles from a non-uniform distribution.

¹⁸We choose this particular value of P because it is the largest value of P such that all the particles fit in the memory of our GPU. We have found very similar results with P set to 5,000, which makes us confident that 10,000 is sufficiently large that the particle filter recovers the true conditional expectations.

Algorithm 1 Particle Filter

- 1: **Fix** a number of particles, P , and a time series of length T , $\{g_t\}_{t=1}^T$.
- 2: **Initialize** particles, $\{\xi_0^p\}_{p=1}^P$, at $\xi_0^p = g_0^*$ for all p , and set $t = 1$.
- 3: **Sample** $\{u_t^p\}_{p=1}^P$ from $N(0, \sigma_u^2)$.
- 4: **Update** particles according to:

$$\{\xi_t^p\}_{p=1}^P = \{\rho \xi_{t-1}^p + u_t^p\}_{p=1}^P.$$

- 5: **Compute** and store:

$$E(g_t | g_1, \dots, g_{t-1}) = E(g_t^* | g_1, \dots, g_{t-1}) = \frac{1}{P} \sum_{p=1}^P \xi_t^p$$

- 6: **Define** $f_\epsilon(\cdot)$ as the PDF of ϵ_t and compute:

$$q_t^p = \frac{f_\epsilon\left(\frac{g_t - \xi_t^p}{\sigma_\epsilon}\right)}{\sum_{p=1}^P f_\epsilon\left(\frac{g_t - \xi_t^p}{\sigma_\epsilon}\right)},$$

- 7: **Resample** $\{\xi_t^p\}_{p=1}^P$ from $\{\xi_t^p\}_{p=1}^P$ with replacement and sampling weights $\{q_t^p\}$.
 - 8: **if** $t = T$ **then**
 - 9: **Stop.**
 - 10: **else**
 - 11: **Set** $t \leftarrow t + 1$ and go to Step 3.
 - 12: **end if**
-